# Linear Model Assignment 4

110024516 統研碩一邱繼賢

2021年11月14日

## Problem 1.

Construct the full model:

$$y = \beta_0 + \beta_1 x_1 + \beta_2 x_2 + \beta_3 x_3 + \beta_4 x_4 + \beta_5 x_5 + \epsilon$$

```
data = read.table("wastes.txt", skip = 1)
names(data) = c("day", "x1", "x2", "x3", "x4", "x5", "y")
g = lm(y ~ x1+x2+x3+x4+x5, data = data)
```

a.

$$CI \; = \; \begin{cases} (\hat{\beta}_3 - t_{(0.025,14)} \; se_{\hat{\beta_3}} \; , \; \hat{\beta}_3 + t_{(0.025,14)} \; se_{\hat{\beta_3}}) \\ (\hat{\beta}_5 - t_{(0.025,14)} \; se_{\hat{\beta_{\rm r}}} \; , \; \hat{\beta}_5 + t_{(0.025,14)} \; se_{\hat{\beta_{\rm r}}}) \end{cases}$$

The results are shown as below:

```
confint(g)[c(4,6),]
```

```
## 2.5 % 97.5 %
## x3 -3.713929e-05 0.0002927368
## x5 -1.652198e-05 0.0002998305
```

b.

$$\begin{split} CI &= ((\hat{\beta}_3 + 2 \ \hat{\beta}_5) - t_{(0.025,14)} \ se_{\hat{\beta}_3 + 2 \ \hat{\beta}_5} \ , \ (\hat{\beta}_3 + 2 \ \hat{\beta}_5) + t_{(0.025,14)} \ se_{\hat{\beta}_3 + 2 \ \hat{\beta}_5}) \\ where \ se_{\hat{\beta}_3 + 2 \ \hat{\beta}_5} &= \sqrt{v\hat{a}r(\hat{\beta}_3) + 2^2 \ v\hat{a}r(\hat{\beta}_5) + 4 \ c\hat{o}v(\hat{\beta}_3 \ , \ \hat{\beta}_5)} \ , \ and \ c\hat{o}v(\hat{\beta}_i \ , \ \hat{\beta}_j) \ = \ (X^TX)_{ij}^{-1} \ \hat{\sigma}^2 \end{split}$$

The result is shown as below:

```
x = model.matrix(g)
xtxi = solve(t(x) %*% x)
sigma = summary(g)$sig
sd_error = sqrt(xtxi[4,4]*sigma^2+4*xtxi[6,6]*sigma^2+4*xtxi[4,6]*sigma^2)
estimate = g$coe[4]+2*g$coe[6]
CI = c(estimate-qt(0.975, g$df)*sd_error, estimate+qt(0.975, g$df)*sd_error)
names(CI) = c("Lower Bound", "Upper Bound")
CI
```

```
## Lower Bound Upper Bound
## 5.898666e-05 7.632279e-04
```

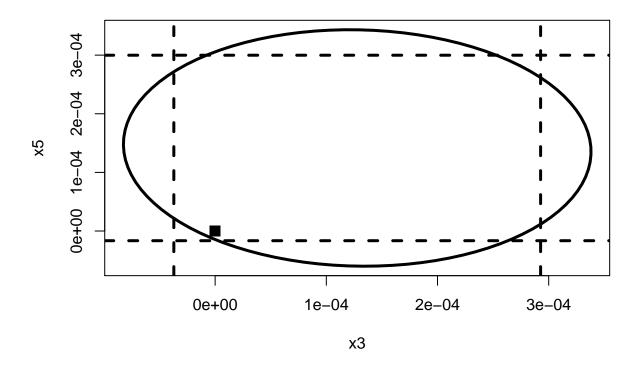
c.

$$\begin{cases} H_0: \beta_3 \ = \ \beta_5 \ = \ 0 \\ H_1: at \ least \ one \ of \ \beta_3 \ or \ \beta_5 \ \neq \ 0 \end{cases}$$

## library(ellipse)

```
##
## Attaching package: 'ellipse'
## The following object is masked from 'package:graphics':
##
## pairs

plot(ellipse(g, c(4,6)), lwd = 3, type = "l")
points(0,0, cex = 1.5, pch = 15)
abline(v=c(confint(g)[4,1], confint(g)[4,2]),lwd=3,lty=2)
abline(h=c(confint(g)[6,1], confint(g)[6,2]),lwd=3,lty=2)
```



From the above ellipse, the origin is in that ellipse, means fail to reject  $H_0$ . Therefore, we do not have enough evidence to show that at least one of  $\beta_3$  or  $\beta_5 \neq 0$  as  $x_1$ ,  $x_2$ ,  $x_4$  are in the model.

**Q.** 探討  $(\beta_1.\beta_2,\beta_3,\beta_4,\beta_5)$  所建構的 95% 聯合信賴區間是否包含 (0,0,0,0,0),同等於在 5% 顯著水準下做以下假設檢定:

$$\begin{cases} H_0: \beta_1 = \beta_2 = \beta_3 = \beta_4 = \beta_5 = 0 \\ H_1: at \ least \ one \ \beta_i \ \neq \ 0 \ , \ i \ = \ 1,...,5 \end{cases}$$

即為進行 full model 的 full test:

#### summary(g)

```
##
## Call:
## lm(formula = y ~ x1 + x2 + x3 + x4 + x5, data = data)
##
## Residuals:
##
        Min
                  1Q
                       Median
                                     3Q
                                             Max
## -0.39447 -0.11847 0.00053 0.08313 0.56232
##
## Coefficients:
##
                 Estimate Std. Error t value Pr(>|t|)
## (Intercept) -2.156e+00 9.135e-01 -2.360
                                                0.0333 *
## x1
               -9.012e-06 5.184e-04 -0.017
                                                0.9864
                1.316e-03 1.263e-03
## x2
                                        1.041
                                                0.3153
## x3
                1.278e-04 7.690e-05
                                        1.662
                                                0.1188
## x4
                7.899e-03 1.400e-02
                                        0.564
                                                0.5815
                1.417e-04 7.375e-05
                                                0.0754 .
## x5
                                        1.921
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
## Residual standard error: 0.2618 on 14 degrees of freedom
## Multiple R-squared: 0.8107, Adjusted R-squared: 0.743
## F-statistic: 11.99 on 5 and 14 DF, p-value: 0.0001184
p-value = 0.0001184 < 0.05 \Rightarrow reject H_0
\therefore the origin (0,0,0,0,0) would lie inside the 95% confidence region for (\beta_1,\beta_2,\beta_3,\beta_4,\beta_5)
e.
檢定非揮發性固體 (x_3-x_4) 對於反映變數是否有線性效應,即判斷模型
```

$$y = \beta_0 + \beta_1 \ x_1 + \beta_2 \ x_2 + \beta_3^* \ (x_3 - x_4) + \beta_4^* \ x_4 + \beta_5 \ x_5 + \epsilon$$

中係數  $\beta_3^*$  是否為 0,而該模型可從原 full model 移項得到

$$y = \beta_0 + \beta_1 x_1 + \beta_2 x_2 + \beta_3 (x_3 - x_4) + (\beta_3 + \beta_4) x_4 + \beta_5 x_5 + \epsilon$$

從此模型可得知,本題所求即為在 full model 下做以下假設檢定

$$\begin{cases} H_0: \beta_3 \ = \ 0 \\ H_1: \beta_3 \ \neq \ 0 \end{cases}$$

而 a. 小題中  $\beta_3$  的信賴區間有包含 0,所以 fail to reject  $H_0$ 

 $\Rightarrow$  We do not have enough evidence to show that non-volatile solids have linear effect on the response under the full model.

### Problem 2.

a.

Construct the full model:

$$lqsa \ = \ \beta_0 + \beta_1 \ lcavol + \beta_2 \ lweight + \beta_3 \ age + \beta_4 \ lbph + \beta_5 \ svi + \beta_6 \ lcp + \beta_7 \ gleason + \beta_8 \ pgg45 + \epsilon$$

```
data2 = read.table("prostate.txt", header = T)
g2 = lm(lpsa ~ lcavol+lweight+age+lbph+svi+lcp+gleason+pgg45, data = data2)
```

i.

$$\begin{cases} 90\% \ CI \ = \ (\hat{\beta}_3 - t_{(0.05,88)} \ se_{\hat{\beta}_3} \ , \ \hat{\beta}_3 + t_{(0.05,88)} \ se_{\hat{\beta}_3}) \\ 95\% \ CI \ = \ (\hat{\beta}_3 - t_{(0.025,88)} \ se_{\hat{\beta}_3} \ , \ \hat{\beta}_3 + t_{(0.025,88)} \ se_{\hat{\beta}_3}) \end{cases}$$

The results are shown as below:

summary(g2)\$coe[4,]

```
confint(g2, level = 0.9)[4,]

## 5 % 95 %

## -0.038210200 -0.001064151

confint(g2)[4,]

## 2.5 % 97.5 %

## -0.041840618 0.002566267
```

```
## Estimate Std. Error t value Pr(>|t|)
## -0.01963718 0.01117272 -1.75759949 0.08229321
```

age 的 90% CI 並沒有包含 0,但 95% CI 則有包含 0,可由此推得 0.05 < p-value < 0.1,觀察  $regression\ summary\$ 報表所呈現的 p-value=0.08229321 的確符合該條件。

ii.

```
Let X_0 = (1.44692, 3.62301, 65, 0.3001, 0, -0.79851, 7, 15)^T, and \beta = (\hat{\beta}_1, \hat{\beta}_2, \hat{\beta}_3, \hat{\beta}_4, \hat{\beta}_5, \hat{\beta}_6, \hat{\beta}_7, \hat{\beta}_8)^T then the predicted value Y_0 = X_0^T \beta = 2.389053 the standard error of the predicted value se_{Y_0} = \sqrt{1 + X_0^T (X^T X)^{-1} X_0} \hat{\sigma}
```

:. The 95%  $CI \ = \ (Y_0 - t_{(0.025,88)} \ se_{Y_0} \ , \ Y_0 + t_{(0.025,88)} \ se_{Y_0})$ 

The result is shown as below:

```
## fit lwr upr
## 1 2.389053 0.9646584 3.813447
```

iii.

Now, the observation has been changed into  $X_1 = (1.44692, 3.62301, 20, 0.3001, 0, -0.79851, 7, 15)^T$ , and then do the same calculation as above to attain the 95% CI.

The result is been shown as below:

## fit lwr upr ## 1 3.272726 1.538744 5.006707

因為 age 變數數據全部都落在  $41\sim79$  的區間之中,所以 age=20 為外插 (extrapolation) 的資料,在根據此筆數據做估計時,誤差範圍 (即信賴區間寬度) 就會變得較寬。

#### b.

Construct the reduced model:

$$lqsa = \beta_0 + \beta_1 \ lcavol + \beta_2 \ lweight + \beta_5 \ svi + \epsilon$$

```
g3 = lm(lpsa ~ lcavol+lweight+svi, data = data2)
```

i.

Use the same method to calculate the prediction and the confidence interval.

```
## fit lwr upr
## 1 2.372534 0.9383436 3.806724
```

計算出來的估計值與 part a 差異不大,信賴區間的寬度也只有稍微寬於 part a 所計算出的寬度。

我會更傾向於選擇 part b 的模型,因為使用較少的變數,即代表在時間和金錢上的成本花費較少,而且在此題的情況下,計算出的預測值和信賴區間都沒有太大的差異。

ii.

$$\begin{cases} H_0: reduced\ model\ fits\ better\\ H_1: full\ model\ fits\ better \end{cases} \Longleftrightarrow \begin{cases} H_0: \beta_3 = \beta_4 = \beta_6 = \beta_7 = \beta_8\\ H_1: at\ least\ one\ \beta_i\ \neq\ 0\ ,\ i=3,4,6,7,8 \end{cases}$$

```
anova(g3, g2)
```

```
## Analysis of Variance Table
##
## Model 1: lpsa ~ lcavol + lweight + svi
## Model 2: lpsa ~ lcavol + lweight + age + lbph + svi + lcp + gleason +
## pgg45
## Res.Df RSS Df Sum of Sq F Pr(>F)
## 1 93 47.785
## 2 88 44.163 5 3.6218 1.4434 0.2167
```

```
\because p-value \ = \ 0.2167 \ > \ 0.05 \ \Rightarrow \ fail \ to \ reject \ H_0
```

 $<sup>\</sup>therefore$  We do not have enough evidence to show that the full model fits better than reduced model. The reduced model is prefered.