

Logistic Regression on the Presence of Heart Disease

Edward Yu

February 09, 2018

Contents

1	Growing Problem of Heart Disease	2
2	Data Set Variables and Descriptions	2
3	Data Wrangling	3
4	Exploratory Analysis	3
4.1	Dealing with Missing Values: Making the Cut	4
4.2	Imputation of Continuous and Categorical Missing Values	6
4.3	Correlation Plots	7
5	Model Selection	8
5.1	LASSO method	8
5.2	Forward & Backwards Stepwise Regression	12
5.3	Model Conclusion	13

Abstract

This project will create a logistic model regressing potential risk factors of heart disease to predict whether an individual has the presence or absence of heart disease. The results of this project can be used in a few ways:

1. To investigate which risk factors are under more or less important in determining an individual's heart condition.
2. This model can be built into pre-screening applications in which users can test for their risk of heart disease, supplementing an individual's regiment of physician check-ups and overall health and body consciousness.
3. Aid physicians more mathematically and precisely define risk percentages for their patients, particularly for patients who have mixed heart condition signals.

1 Growing Problem of Heart Disease

Heart and cardiovascular disease is the leading cause of death worldwide, accounting for approximately 17.3 million deaths per year, and the pangs of this disease are only expected to worsen. By 2030, that number is expected to grow to more than 23.6 million. This accounts for nearly one out of every three deaths. To put a dollar sign to this problem, direct and indirect costs of cardiovascular diseases add up to more than \$320.1 billion in the U.S. in medical costs and lost productivity¹.

As we explore more about the risk factors of heart disease and find innovative ways to detect its presence, we can begin to make smarter decisions about our health.

2 Data Set Variables and Descriptions

This project uses the UCI machine learning repository's [Heart Disease Data Set](#). The directory contains 4 databases concerning heart disease diagnosis from 4 locations:

- Cleveland Clinic Foundation (cleveland.data)
- Hungarian Institute of Cardiology, Budapest (hungarian.data)
- V.A. Medical Center, Long Beach, CA (long-beach-va.data)
- University Hospital, Zurich, Switzerland (switzerland.data)

Each database contained the same data formatting with a total of 14 attributes. Below are the data variables and descriptions:

Variable	Definition
age	Age in Years
sex	Sex of Patient (1:male; 0:female)
cp	Chest pain type (1=typical angina; 2=atypical angina; 3=non-anginal pain; 4=asymptomatic)
trestbps	Resting blood pressure in mm Hg
chol	Serum cholesterol in mg/dl
fbs	Fasting Blood sugar 120 mg/dl (1=true; 0=false)
restecg	Resting electrocardiographic results (0=normal; 1=having ST-T wave abnormality; 2= showing probable or definite left ventricular hypertrophy by Estes' criteria)
thalach	Maximum heart rate achieved
exang	Exercise induced angina (1=yes; 2=no)
oldpeak	ST depression induced by exercise relative to rest
slope	Slope of the peak exercise ST segment (1=upsloping; 2=flat; 3=downsloping)
ca	Number of major vessels
thal	Thallium heart scan, indicates areas of heart that are not getting adequate supply of blood (3=normal; 6=fixed defect; 7=reversible defect)
num	Diagnosis of heart disease angiographic diseases status (0= <50% diameter narrowing; 1= <50% diameter narrowing)

¹[Heart Disease Stats](#)

3 Data Wrangling

Since the data spread across 4 locations contained the same variable structure, the primary goal was to append the rows of each data set into one master data set. Because each observation in the four data sets represented a unique, independent observation the data was already in TIDY structure. Before appending the data observations together variables within each data set was changed to the same variable type. The most important step of this process was converting numerically coded values into proper factors. For example, variable 'cp', chest pain type, was originally coded with a numeric value from 1-4, each associated with a type of chest pain. These values were recoded to character variables describing the chest pain and then factorized.

The dependent variable 'num' is an integer valued from presence (values 1-4) from 0 (no presence). Because we are simply attempting to distinguish from a set of features whether or not a patient has the presence of heart disease, values >0 were coded as 1, thus creating two levels "Presence" and "Absence". From this step, the dependent variable is primed for logistic regression techniques.

A new variable called 'dsetid' was created denoting which data set each observation came from. This was used to detect any non-random missing value distributions from data set to data set. This variable will not be used in the actual modeling and analysis sections.

Lastly, after variables from each data set were coerced to uniform variable types, each regional data set was row binded together and 'heartDT' became the resulting final data set ready for exploratory analysis.

4 Exploratory Analysis

The initial heart data set contains a total of 916 observations and 15 variables. Below is a breakdown of each variable, class, and type:

Variables	Class	Type
age	Numerical	Continuous
sex	Factor	Categorical
cp	Factor	Categorical
trestbps	Numerical	Continuous
chol	Numerical	Continuous
fbs	Factor	Categorical
restecg	Factor	Categorical
thalach	Numerical	Continuous
exang	Factor	Categorical
oldpeak	Numerical	Continuous
slope	Factor	Categorical
ca	Factor	Categorical
thal	Factor	Categorical
num	Factor	Categorical

4.1 Dealing with Missing Values: Making the Cut

A total of 12.74% of the data was missing. Below is a visual graph of the missingness of each variable, we see the following breakout:

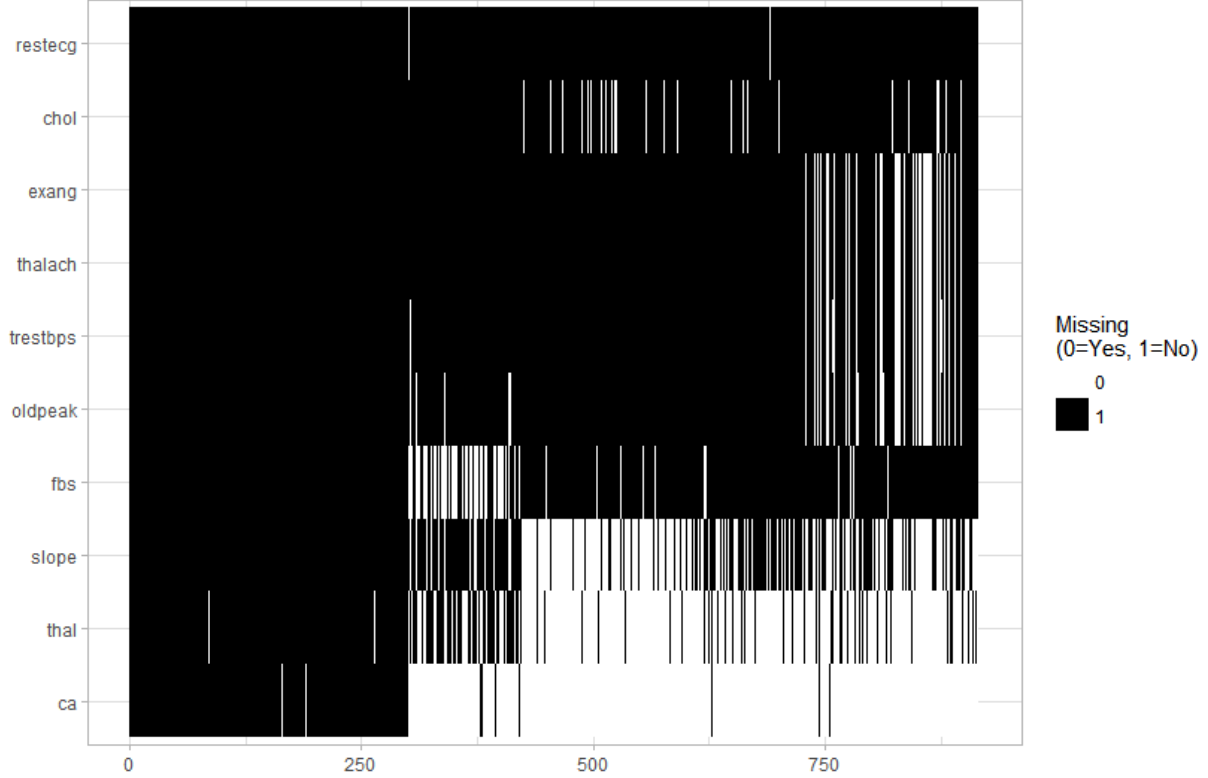


Figure 1: Missing Observations by Variable

The figure above is sorted in increasing number of missing observations from top to bottom with restecg having the least number of missing variables and ca having the greatest. Variables with no missing values were omitted from the figure. The variables restecg, chol, exang, thalach, trestbps, oldpeak, and fbs seem to be missing at random. Although it is impossible to verify statistically, it is substantively reasonable to claim these variables do not have any significant patterns of missingness. Some argument can be made that the variables exang, thalach, trestbps, and old peak have similar lines appearing for the same observations because of the straight white lines that cross through all three variables. These observations are most likely the result of nonresponse from the same individual patients. However, because these observations constitute only a minute percentage of the whole data set the missingness will not affect future analysis significantly and will be dealt with by imputation.

Variables slope, thal, and ca have 33.62%, 66.37%, and 52.72% missingness respectively, which is significantly larger than other variables as denoted by the large white

chunks in Figure 1. Because of the particular method of binding rows used:

```
heartdata <- -bind_rows(cleveland, switzerland, hungarian, longbeach)
```

We know that the first observations within the *heartdata* set come from the cleveland data set with 302 observations. The observations of switzerland, hungarian, and long beach data sets follow thereafter. This indicates that the majority of missingness from *slope*, *thal*, and *ca* are likely to be from the data set regions of switzerland, hungarian, and long beach. This is further confirmed in the table below:

Missing Variable			
Data Set Region	<i>slope</i>	<i>ca</i>	<i>thal</i>
Cleveland	0.0%	0.4%	.2%
Switzerland	1.8%	12.7%	5.5%
Hungarian	20.6%	31.7%	28.9%
Long Beach	11.1%	21.5%	18.0%

From the table above, we can conclude that there is a systemic presence of missingness due to regional data collection. Since *slope*, *thal*, and *ca* have a significant percentage of data missing and is systemic, these variables have been removed from the data set and will be left out in further analysis. Although it may be alluring to impute these variables, the amount of missing data is not trivial and if done so will bias the analysis.

4.2 Imputation of Continuous and Categorical Missing Values

Missing values for the remaining variables were split into numerical and categorical classifications. Numerical variables were imputed using the predicted mean method. Categorical variables were imputed using Bayesian polytomous regression since the factored are unordered and contain more than two levels.

The distribution of original and imputed numeric data is displayed in the density graph below:

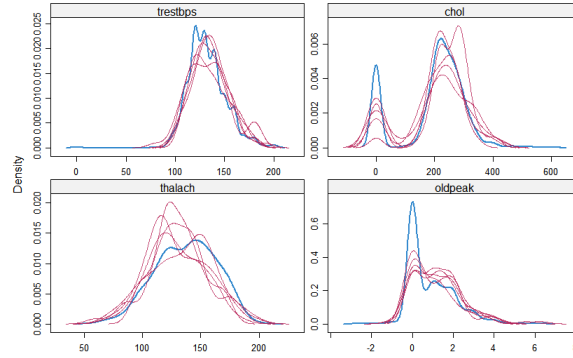


Figure 2: Density graph of imputed numeric data

The blue lines represent the density of the original data while the magenta line represents the imputed data distribution. We see that `trestbps` and `chol` have very similar lines densities. `thalach`'s imputed distribution shows a higher peak distribution in the low 100's as compared to the original data. `oldpeak` original has a higher density centered around 0. For the most part the matching shape of the blue and magenta lines indicate that the imputed values are plausible values with similar distributions as the original data. Another graph, the strip plot, shares a similar story:

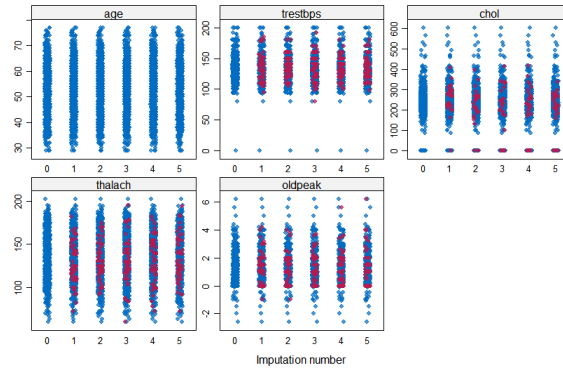


Figure 3: Density graph of imputed numeric data

Age is also displayed within the strip plot, but because age did not contain any missing observations all dots were blue.

A strip plot of categorical data imputation was also created:

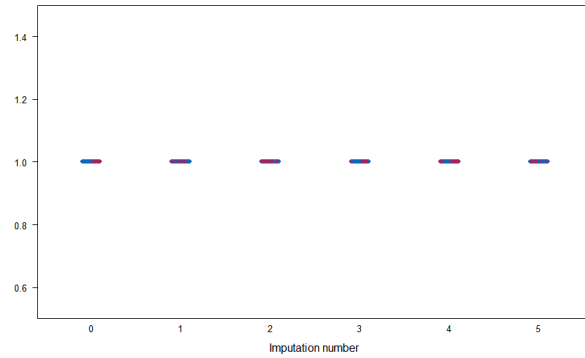


Figure 4: Density graph of imputed numeric data

Blue and magenta dots overlap on each other indicating the imputed values are close to the real values. Both numeric and categorical imputed values replaced the original missing values within the 'heartDT' set, thus finalizing the 'heartDT' data set for model selection and logistic regression analysis.

4.3 Correlation Plots

Numerical variables were separated from categorical variables from the 'heartDT' data set. The corplot function was used to create the following correlation plot:

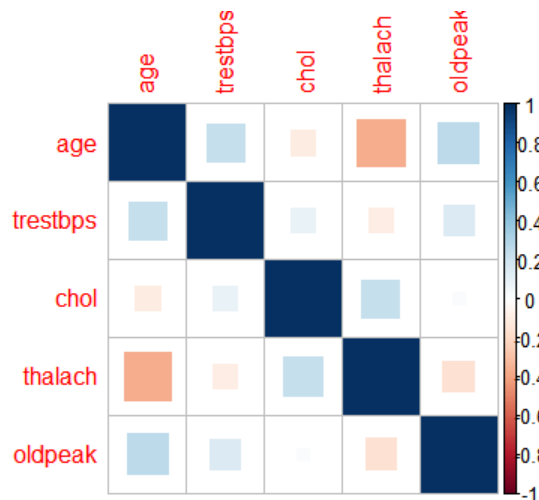


Figure 5: Correlation Plot of Numeric Variables

From the correlation plot we can see there is no strong correlation between any of the numeric variables. Age did have a weak correlation with trestbps and oldpeak, and chol with thalach had a similar correlation of 0.4.

5 Model Selection

Starting with 916 total observations in the 'heartDT' data set the master data set was split into training and testing data sets. 75% of the data was allocated to the training and 25% to the testing. All model selections were trained on the training set and then tested using the testing set. The featured model selection methods tested in this project will be forward stepwise, backward stepwise, and L1 regularization (LASSO). The forward and backwards stepwise's performance will be tested on AIC criteria. The strengths and weaknesses of each model selection method will then be discussed and compared.

5.1 LASSO method

In the world of statistics, LASSO and its related techniques such as Ridge regression and Elastic net are fairly newer methods compared to stepwise regression that have enhanced our ability for prediction accuracy and interpretability of our statistical models. The LASSO method is a way of both variable selection and regularization and offers a parsimonious model by constraining the magnitude of the coefficients, thus making the model simpler and more elegant.

The 'glmnet' function was used to complete the LASSO regression. After splitting the original 'heartDT', 687 observations was appropriated to the 'heartTrain' data set. The following code was run to fit the first LASSO regression:

$$fit = glmnet(x = X_train, y = Y_train, family = "binomial")$$

where X_train is a matrix of independent variables and Y_train is an object containing the 'num' regressed variable. The plot below shows the fit:

Each curve corresponds to a variable within the model, including each categorical factor with a total of 13 variables. The path of the curve shows its coefficient against the L_1norm as λ varies. The axis above indicates the number of non-zero coefficients at the current λ . The λ controls trade off between fitting the training data well (minimizing deviance) and keeping the coefficients small. It effectively controls keeping hypothesis relatively simple to avoid overfitting.

When λ is small the LASSO regularization term, L_1norm , the LASSO solution is similar to OLS solution retaining all coefficients. In contrast, a large λ value will penalize coefficients highly, increasing the regularization reducing the parameters to zero ending up with an underfit model. Controlling for the λ value is essential for managing the complexity of the model.

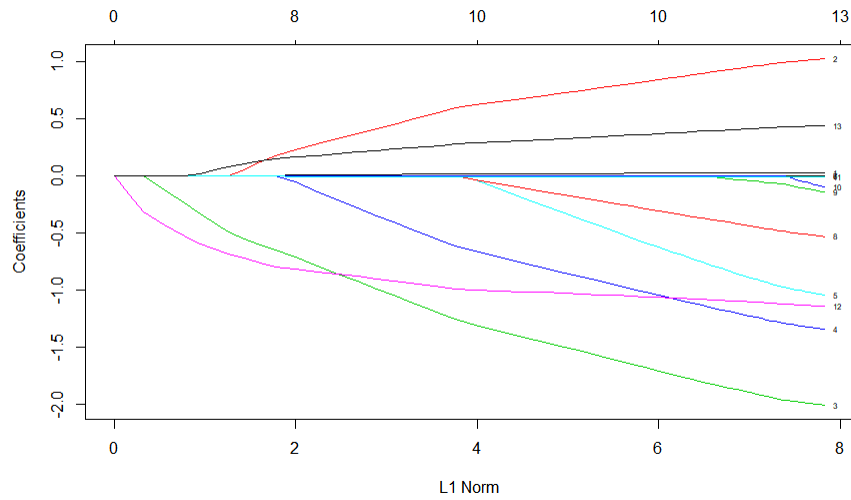


Figure 6: L1 Norm vs. Coefficients from LASSO regression

Cross-validation is used to select the most optimal λ . In choosing a correct value we will consider the end goal of prediction accuracy. The function `cv.glmnet` conveniently provides us a way of cross-validation:

```
cvfit = cv.glmnet(x = X_train, y = Y_train, family = "binomial", type.measure = "class")
```

Below is a graph of the resulting cross-validation error curve:

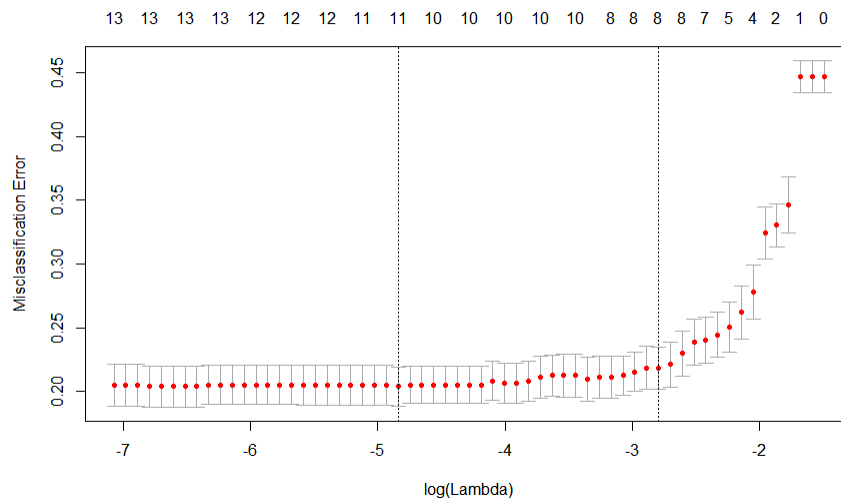


Figure 7: Cross-Validation Curve vs. Lambda

The red dotted line represents the cross-validation curve and the error bars represent the the upper and lower standard deviation. A similar trend is seen here with larger λ values leading to an increasing effect of regularization and reducing more coefficients to 0. The two black dotted lines represent the λ value with the minimum CV error and the λ value one standard error away (also representing the slightly more regularized model). We select the λ value with the least CV error because the cross-validation error estimates prediction error, so by choosing the λ with the minimum CV error we are also achieving minimal prediction error.

Now, we test predictions for both $\lambda_{min} = 0.00789$ and $\lambda_{1se} = 0.0611$ using the 'heartTest' data set:

```
cv_pred_min <- predict(cvfit, newx = X_test, type = "class", s = cvfit$lambda.min)
```

```
cv_pred_1se <- predict(cvfit, newx = X_test, type = "class", s = cvfit$lambda.1se)
```

Analyzing the prediction accuracy with the confusion matrix:

<pre>> confusionMatrix(cv_pred_min, heartTest\$num) Confusion Matrix and Statistics</pre> <table border="1"> <thead> <tr> <th></th> <th colspan="2">Reference</th> </tr> <tr> <th>Prediction</th> <th>Absence</th> <th>Presence</th> </tr> </thead> <tbody> <tr> <th>Absence</th> <td>84</td> <td>13</td> </tr> <tr> <th>Presence</th> <td>18</td> <td>114</td> </tr> </tbody> </table> <pre> Accuracy : 0.8646 95% CI : (0.8134, 0.9061) No Information Rate : 0.5546 P-Value [Acc > NIR] : <2e-16 Kappa : 0.7247 Mcnemar's Test P-Value : 0.4725 Sensitivity : 0.8235 Specificity : 0.8976 Pos Pred Value : 0.8660 Neg Pred Value : 0.8636 Prevalence : 0.4454 Detection Rate : 0.3668 Detection Prevalence : 0.4236 Balanced Accuracy : 0.8606 'Positive' class : Absence </pre>		Reference		Prediction	Absence	Presence	Absence	84	13	Presence	18	114	<pre>> confusionMatrix(cv_pred_1se, heartTest\$num) Confusion Matrix and Statistics</pre> <table border="1"> <thead> <tr> <th></th> <th colspan="2">Reference</th> </tr> <tr> <th>Prediction</th> <th>Absence</th> <th>Presence</th> </tr> </thead> <tbody> <tr> <th>Absence</th> <td>83</td> <td>12</td> </tr> <tr> <th>Presence</th> <td>19</td> <td>115</td> </tr> </tbody> </table> <pre> Accuracy : 0.8646 95% CI : (0.8134, 0.9061) No Information Rate : 0.5546 P-Value [Acc > NIR] : <2e-16 Kappa : 0.7241 Mcnemar's Test P-Value : 0.2812 Sensitivity : 0.8137 Specificity : 0.9055 Pos Pred Value : 0.8737 Neg Pred Value : 0.8582 Prevalence : 0.4454 Detection Rate : 0.3624 Detection Prevalence : 0.4148 Balanced Accuracy : 0.8596 'Positive' class : Absence </pre>		Reference		Prediction	Absence	Presence	Absence	83	12	Presence	19	115
	Reference																								
Prediction	Absence	Presence																							
Absence	84	13																							
Presence	18	114																							
	Reference																								
Prediction	Absence	Presence																							
Absence	83	12																							
Presence	19	115																							

Figure 8: Left: λ_{min} , Right: λ_{1se}

The accuracy of both models are essentially the same with a small difference of λ_{min} having a higher sensitivity and lower specificity than λ_{1se} .

The final features the LASSO model selects are:

```
#finding actual coefficients for lasso model lse
fitlse<- glmnet(x=X_train, y=Y_train, family = "binomial", alpha=1, lambda=cv$lambda.min)
fitlse$beta[,1]
      age      sexMale cpAtypical Angina cpNon-Anginal Pain  cpTypical Angina
0.008505194  0.367322413 -0.925752036 -0.282797951  0.000000000
trestbps      chol      fbs2      restecg2      restecg3
0.000000000 -0.001198026  0.000000000  0.000000000  0.000000000
thalach      exang2      oldpeak
-0.006350008 -0.880883107  0.206453483
#finding actual coefficients for lasso model min
fitmin<-glmnet(x=X_train, y=Y_train, family = "binomial", alpha=1, lambda=cv$lambda.1se)
fitmin$beta[,1]
      age      sexMale cpAtypical Angina cpNon-Anginal Pain  cpTypical Angina
0.007063912  0.323097582 -0.859330080 -0.212982965  0.000000000
trestbps      chol      fbs2      restecg2      restecg3
0.000000000 -0.001021031  0.000000000  0.000000000  0.000000000
thalach      exang2      oldpeak
-0.006140946 -0.861199350  0.192817722
```

Figure 9: Lasso Model Coefficients; Top: λ_{lse} , Bottom: λ_{min}

Plugging in the coefficients into the logistic regression equation we have:

$$\log\left(\frac{p}{1-p}\right) = \beta_0 + \beta_1 * age + \beta_2 * sexMale + \beta_3 * cpAtypical Angina + \dots + \beta_{13} * oldpeak$$

where each β is represented by the coefficients listed in Figure 9. In interpreting the model, for example, holding all other variables at a fixed value, the odds of the presence of heart disease if a patient is male rather than female is $e^{0.323} = 1.381$. In other words, we can say that the odds for males having heart disease are 38% higher than the odds for females.

5.2 Forward & Backwards Stepwise Regression

The stepwise regression method is an older approach of choosing predictive variables in an automated fashion. Within each step of the process a variable is considered for addition or subtraction from the overall model starting from either a null or full model. The forward model begins by starting with a null model, then slowly steps upwards by adding the most significant model. The backward model begins with the full model and cuts out the least significant model.

A null model was first created regressing 'num' 1. Afterwards the 'step' function from the leap package was used to perform the forward stepwise regression. The final model contained the following variables: cp, exang, chol, oldpeak, sex, age, thalach, and fbs. After each step the AIC criteria decreased to a final value of 614.4.

Likewise, a full model containing all variables was first created regressing 'num' .. Afterwards, the 'step' function was used to perform the backward stepwise regression. The final contained the following variables: age, sex, cp, chol, fbs, thalach, exang, and oldpeak. The AIC criteria also was 614.4 for the final model as expected. The coefficients for the forwards and backwards stepwise models are shown below:

For the logistic regression equation we have:

$$\log\left(\frac{p}{1-p}\right) = \beta_0 + \beta_1 * age + \beta_2 * sexMale + \beta_3 * cpAtypicalAngina + \dots + \beta_{10} * oldpeak$$

The coefficients for the forwards and backwards stepwise models are shown below:

Coefficients:				
	Estimate	Std. Error	z value	Pr(> z)
(Intercept)	1.441405	1.157776	1.245	0.213140
age	0.027290	0.012332	2.213	0.026905 *
sexMale	1.043895	0.265229	3.936	8.29e-05 ***
cpAtypical Angina	-2.048267	0.318472	-6.432	1.26e-10 ***
cpNon-Anginal Pain	-1.359940	0.249053	-5.460	4.75e-08 ***
cpTypical Angina	-1.080908	0.422742	-2.557	0.010561 *
chol	-0.003773	0.001027	-3.674	0.000239 ***
fbs2	-0.539679	0.291270	-1.853	0.063904 .
thalach	-0.008642	0.004513	-1.915	0.055534 .
exang2	-1.135706	0.238350	-4.765	1.89e-06 ***
oldpeak	0.443262	0.104712	4.233	2.30e-05 ***

Figure 10: Coefficients for Stepwise Models

The confusion model is shown below:

```
> confusionMatrix(fwdpred, heartTest$num) > confusionMatrix(bwdpred, heartTest$num)
Confusion Matrix and Statistics          Confusion Matrix and Statistics
```

Reference			Reference		
Prediction	Absence	Presence	Prediction	Absence	Presence
Absence	80	11	Absence	82	12
Presence	22	116	Presence	20	115

```

      Accuracy : 0.8559
      95% CI : (0.8036, 0.8987)
    No Information Rate : 0.5546
    P-Value [Acc > NIR] : < 2e-16

      Kappa : 0.7052
  Mcnemar's Test P-Value : 0.08172

    Sensitivity : 0.7843
    Specificity : 0.9134
    Pos Pred Value : 0.8791
    Neg Pred Value : 0.8406
    Prevalence : 0.4454
    Detection Rate : 0.3493
    Detection Prevalence : 0.3974
    Balanced Accuracy : 0.8488

    'Positive' class : Absence

```

Reference			Reference		
Prediction	Absence	Presence	Prediction	Absence	Presence
Absence	82	12	Absence	82	12
Presence	20	115	Presence	20	115

```

      Accuracy : 0.8603
      95% CI : (0.8085, 0.9024)
    No Information Rate : 0.5546
    P-Value [Acc > NIR] : <2e-16

      Kappa : 0.715
  Mcnemar's Test P-Value : 0.2159

    Sensitivity : 0.8039
    Specificity : 0.9055
    Pos Pred value : 0.8723
    Neg Pred value : 0.8519
    Prevalence : 0.4454
    Detection Rate : 0.3581
    Detection Prevalence : 0.4105
    Balanced Accuracy : 0.8547

    'Positive' class : Absence

```

Figure 11: Left: Forward Step, Right: Backward Step

Both the forward and backward stepwise models produced the same list of variables and had a final AIC value of 614.4. The accuracy of the backwards regression model was slightly higher than the forward model. The sensitivity of the forward model had a lower sensitivity but a greater specificity.

5.3 Model Conclusion

The primary point of model comparison between LASSO and Stepwise regression methods is prediction accuracy. On testing with 'heartTest' data set, the LASSO method performed slightly better than the Stepwise regression methods. Considering the subset of features within this data set was fairly small with a total of 13 variables, the LASSO regularization method did not have as much of a chance to shine. The LASSO method finds its strength where $p \gg n$. p , representing the number of features within a model and n , representing the number of observations. With a large number of features, the LASSO method does a great job in coercing coefficients to 0, reducing the amount of overfitting. One inherent weakness of L1 regularization is that it tends to do poorly on variable sets with high levels of multicollinearity by often selecting a group of highly correlated variables. However, as we have seen there is not a high level of correlation amongst the numeric variables of this particular data set, thus is not a significant issue here.

Overall, regarding the condition of heart disease, the most significant contributor seems to be sexMale having the largest coefficient of 0.323 within the LASSO model. Other indicators such as oldpeak, representing ST depression is often a sign of myocardial ischemia, which is the lack of blood supply to the heart. In a cardiac stress test, a

change of at least 2 mm signifies reversible ischemia. Any mm drop greater than 2 from from vertical distance between the patient's trace and isoelectric line is indicative of irreversible ischemia ².

²Ischemia Site