

1. a)  $X \sim N(0, \sigma^2)$

$$f_X(x) = \frac{1}{\sqrt{2\pi}\sigma} e^{-x^2/2\sigma^2}$$

$$Y = |X| = \begin{cases} x & \text{if } x \geq 0 \\ -x & \text{if } x < 0 \end{cases} \Rightarrow \begin{matrix} x = y & \text{if } x \geq 0 \\ x = -y & \text{if } x < 0 \end{matrix}$$

$$f_Y(y) = f_X(y) \cdot \left| \frac{d}{dy} y \right|$$

$$= \frac{1}{\sqrt{2\pi}\sigma} e^{-y^2/2\sigma^2} \quad \text{if } x \geq 0 \quad y \geq 0$$

$$f_Y(y) = f_X(-y) \cdot \left| \frac{d}{dy} -y \right|$$

$$= \frac{1}{\sqrt{2\pi}\sigma} e^{-y^2/2\sigma^2} \quad \text{if } x < 0 \quad y \geq 0$$

Thus  $f_Y(y) = \frac{2}{\sqrt{2\pi}\sigma} e^{-y^2/2\sigma^2}$  for  $y \geq 0$

b)  $E(Y) = \int_0^\infty y f_Y(y) dy = \int_0^\infty y \sqrt{\frac{2}{\pi}} \cdot \frac{1}{\sigma} e^{-y^2/2\sigma^2} dy$

let  $-y^2/2\sigma^2 = t$   
 $y dy = \sigma^2 dt$

$$E(Y) = \sqrt{\frac{2}{\pi}} \frac{1}{\sigma} \sigma^2 \int_0^\infty e^{-t} dt$$

$$= \left[ \sqrt{\frac{2}{\pi}} \sigma (-e^{-t}) \right]_0^\infty$$

$$\boxed{= \sqrt{\frac{2}{\pi}} \sigma}$$

$$E(Y^2) = \int_0^\infty y^2 \sqrt{\frac{2}{\pi}} \frac{1}{\sigma} e^{-y^2/2\sigma^2} dy$$

$$= \sqrt{\frac{2}{\pi}} \sigma \int_0^\infty \sqrt{2\sigma^2 t} e^{-t} dt$$

$$= \frac{2}{\sqrt{\pi}} \sigma^2 \int_0^\infty t^{\frac{1}{2}} e^{-t} dt \Rightarrow \text{Gamma dist}$$

$$= \frac{2}{\sqrt{\pi}} \sigma^2 \Gamma\left(\frac{3}{2}\right)$$

$$= \frac{1}{\sqrt{\pi}} \sigma^2 \Gamma\left(\frac{1}{2}\right)$$

$$= \sigma^2$$

$$\begin{aligned} \text{Var}(Y) &= E(Y^2) - E(Y)^2 \\ &= \left(1 - \frac{2}{\pi}\right) \sigma^2 \end{aligned}$$

$$c) L(\sigma^2 | \mathcal{X}) = \prod_{i=1}^N f_Y(y_i) = \left(\frac{2}{\pi}\right)^{\frac{N}{2}} \left(\frac{1}{\sigma^2}\right)^{\frac{N}{2}} e^{-\frac{\sum y_i^2}{2\sigma^2}}$$

$$\log L(\sigma^2 | \mathcal{X}) = \frac{N}{2} \log \frac{2}{\pi} - \frac{N}{2} \log \sigma^2 - \frac{\sum y_i^2}{2\sigma^2}$$

$$\frac{\partial \log L(\sigma^2 | \mathcal{X})}{\partial \sigma^2} = -\frac{N}{2\sigma^2} + \frac{\sum y_i^2}{2(\sigma^2)^2} = 0$$

$$\frac{\sum y_i^2}{2(\sigma^2)^2} = \frac{N}{2\sigma^2}$$

$$MLE: \hat{\sigma}^2 = \frac{\sum y_i^2}{N}$$

2. a) Since  $P(A, B) = P(A)P(B)$ , with unobserved check ~~point~~  $C$ ,

Then  $P(A|B) = P(A)$ ,  $P(B|A) = P(B)$

Then  $P(A|B, \phi) = P(A|\phi)$

Thus  $A \perp\!\!\!\perp B | \phi$

$$b) P(A, B, C, D) = P(A)P(B)P(C|A, B)P(D|C)$$

$$P(A, B, C|D) = \frac{P(A)P(B)P(C|A, B)P(D|C)}{P(D)}$$

$$P(A, B|D) = \frac{P(A)P(B) \cdot \sum_C P(C|A, B)P(D|C)}{P(D)}$$

$$\text{Then } P(A|B, D) = \frac{P(A) \cdot \sum_C P(C|A, B)P(D|C)}{P(D)} \neq P(A|D)$$

Thus  $A \nperp\!\!\!\perp B | D$

$$\begin{aligned}
c) P(C=1 | d=2, a=0, b=1) &= \frac{P(a, b, C, d)}{P(a, b, C=0, d) + P(a, b, C=1, d)} \\
&= \frac{P(a=0) P(b=1) P(C=1 | a=0, b=1) P(d=2 | C=1)}{P(a=0) P(b=1) P(C=1 | a=0, b=1) P(d=2 | C=1) + P(a=0) P(b=1) P(C=0 | a=0, b=1) P(d=2 | C=0)} \\
&= \frac{0.5 \times 0.5 \times 0.5 \times 0.00443}{0.5 \times 0.5 \times 0.5 \times (0.00443 + 0.05399)} \\
&= \frac{0.00443}{0.00443 + 0.05399} \\
&= 0.0758
\end{aligned}$$

$$\begin{aligned}
d) P(d=2) &= \sum_{a, b, c} P(a, b, C, d) \\
&= 0.5 \times 0.5 \times [(0.1 + 0.1 + 0.3 + 0.5) \times 0.05399 + (0.9 + 0.9 + 0.7 + 0.5) \times 0.00443] \\
&= 0.01682
\end{aligned}$$

$$\begin{aligned}
P(C=1) &= \sum_{a, b} P(a, b, C) = 0.5 \times 0.5 \times (0.9 + 0.9 + 0.7 + 0.5) \\
&= 0.75
\end{aligned}$$

$$\begin{aligned}
P(C=1 | d=2) &= \frac{P(d=2 | C=1) P(C=1)}{P(d=2)} \\
&= \frac{0.00443 \times 0.75}{0.01682} = 0.0197
\end{aligned}$$

3, a)

$$\parallel t - Xw \parallel^2$$

$$L(\lambda_1, \lambda_2, w) = (t - Xw)^T (t - Xw) + \lambda_1 \|w\|_1 + \lambda_2 \|w\|_2^2$$

$$P(w) = \text{Lasso}(w) = \parallel t - X\hat{w} \parallel^2 + \lambda_1 \|\hat{w}\|_1$$

$$P_2(w) = \text{Ridge}(w) = \parallel t - Xw \parallel^2 + \lambda_2 \|w\|_2^2$$

$$P_3(x) = \text{Linear} = \parallel t - Xw \parallel^2$$

b) ① let  $t = 0$ ,  $w_1^{(0)} = 0.2$ ,  $w_2^{(0)} = 1$

$$x^{(1)} = \arg \min \parallel t - Xw^{(1)} \parallel_1$$

$$③ \quad w_1^{(1)} = \arg \min [ \parallel t - Xw^{(1)} \parallel_2^2 + \lambda_1 \|w^{(1)}\|_1 ]$$

$$w_2^{(1)} = \arg \min [ \parallel t - Xw^{(1)} \parallel_2^2 + \lambda_2 \|w^{(1)}\|_2^2 ]$$

$$4. a) \quad p(x) = \sum_{k=1}^K \pi_k p(x|k)$$

$$p(x_b, x_a) = \sum_{k=1}^K \pi_k p(x_b, x_a | k)$$

$$= \sum_{k=1}^K \pi_k p(x_b | x_a, k) p(x_a | k)$$

$$p(x_a) = \int_{x_b} \sum_k \pi_k p(x_b, x_a | k) dx_b$$

$$= \sum_k \pi_k \int_{x_b} p(x_b, x_a | k) dx_b$$

$$= \sum_k \pi_k p(x_a | k)$$

$$p(x_b | x_a) = \frac{p(x_b, x_a)}{p(x_a)}$$

$$= \frac{\sum_{k=1}^K \pi_k p(x_b, x_a | k)}{\sum_k \pi_k p(x_a | k)}$$

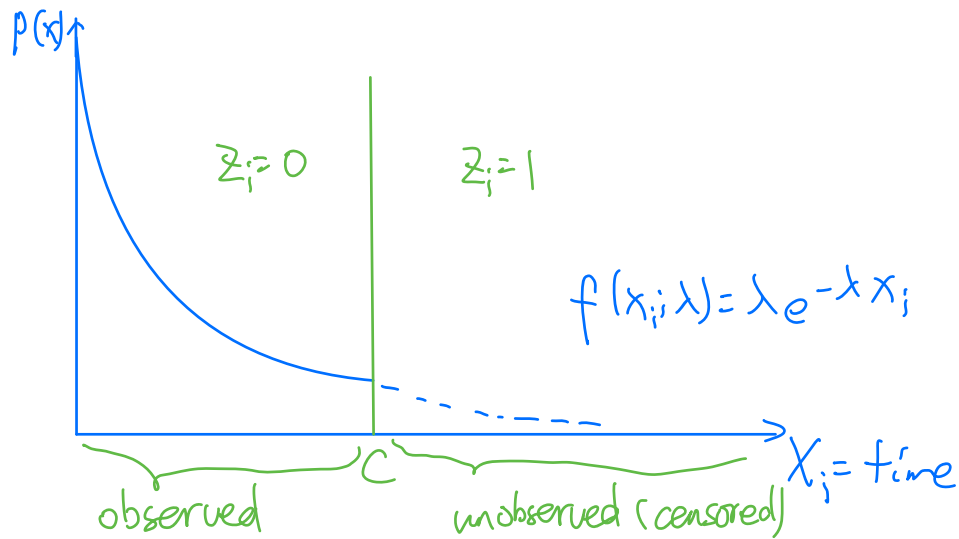
$$= \frac{\sum_{k=1}^K \pi_k p(x_b | x_a, k) p(x_a | k)}{\sum_{k=1}^K \pi_k p(x_a | k)}$$

$$= \sum_{k=1}^K \frac{\pi_k p(x_a | k)}{\sum_{k=1}^K \pi_k p(x_a | k)} p(x_b | x_a, k)$$

$$b) \text{ mixing coef: } \frac{\pi_k p(x_a | k)}{\sum_{k=1}^K \pi_k p(x_a | k)}$$

$$\text{component densities: } p(x_b | x_a, k)$$

5. a) Since  $X_1, \dots, X_n \sim \exp(\lambda)$



$$\begin{aligned}
 \text{b) } P(X=x | X \geq c) &= \frac{P(X=x, X \geq c)}{P(X \geq c)} \\
 &= \frac{f(x, \lambda)}{1 - F_X(x \geq c)} \\
 &= \frac{\lambda e^{-\lambda x}}{1 - (1 - e^{-\lambda c})} \\
 &= \lambda e^{-\lambda(x-c)}
 \end{aligned}$$

$$C) L(x_i, y_i, z_i, i=1..n; \lambda) = \prod_{i=1}^n p(x_i)^{1-z_i} p(y_i)^{z_i}$$

$$p(Y=y) = \lambda e^{-\lambda(y-c)}$$

$$E[\log L(\cdot)] = E\left[\sum_{i=1}^n (1-z_i) \log p(x_i) + z_i \log p(y_i)\right]$$

$$= \sum_{i=1}^n (1-z_i) (\log \lambda - \lambda x_i) + z_i E(p(Y=y_i)) (\log \lambda - \lambda y_i)$$

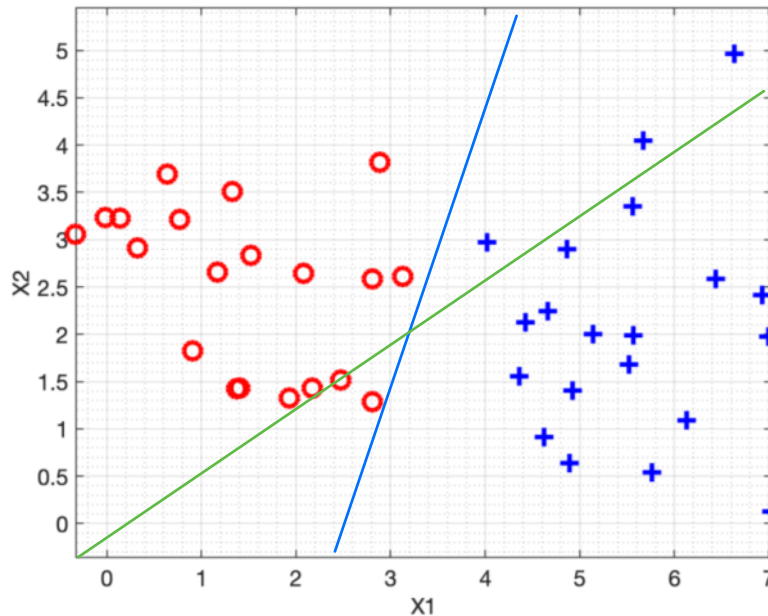
$$= \sum_{i=1}^n (1-z_i) (\log \lambda - \lambda x_i) + z_i (\log \lambda - \lambda (c + \frac{1}{\lambda^{old}}))$$

$$Q(\lambda, \lambda^{old}) = n \log \lambda - \lambda \sum (1-z_i) x_i + z_i (c + \frac{1}{\lambda^{old}})$$

$$\frac{\partial Q}{\partial \lambda} = \frac{n}{\lambda} - \sum_{i=1}^n (1-z_i) x_i + z_i (c + \frac{1}{\lambda^{old}}) = 0$$

$$\hat{\lambda} = \frac{n}{\sum_{i=1}^n (1-z_i) x_i + z_i (c + \frac{1}{\lambda^{old}})}$$

6.



$$a) \quad x_2 = -\frac{x_1}{x_2} \cdot x_1 - \frac{1}{x_2}$$

$$x_2 = -\frac{x_1^2}{x_2} - \frac{1}{x_2}$$

$$x_1^2 + x_2^2 = 1$$

The decision boundary would be a line, start at  $x_1$ -axis,  $w$  would be about 0.5 for each.

b) Since there is no constant term in this model, the decision boundary would become a line start from (0,0) and  $w$  may not balance, may be  $(0.2, 0.8) - (0.3 - 0.7)$  about

$$c) \quad p(D|M_1) = \int p(D|w_1, M_1) p(w_1|M_1) dw_2$$

$$p(D|M_2) = \int p(D|w_1, M_1) p(w_1|M_1) dw_2$$



# midterm

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2022/3/4

## problem 3

c

```
x <- rnorm(100, 0,1)
e <- rnorm(100,0,0.1)

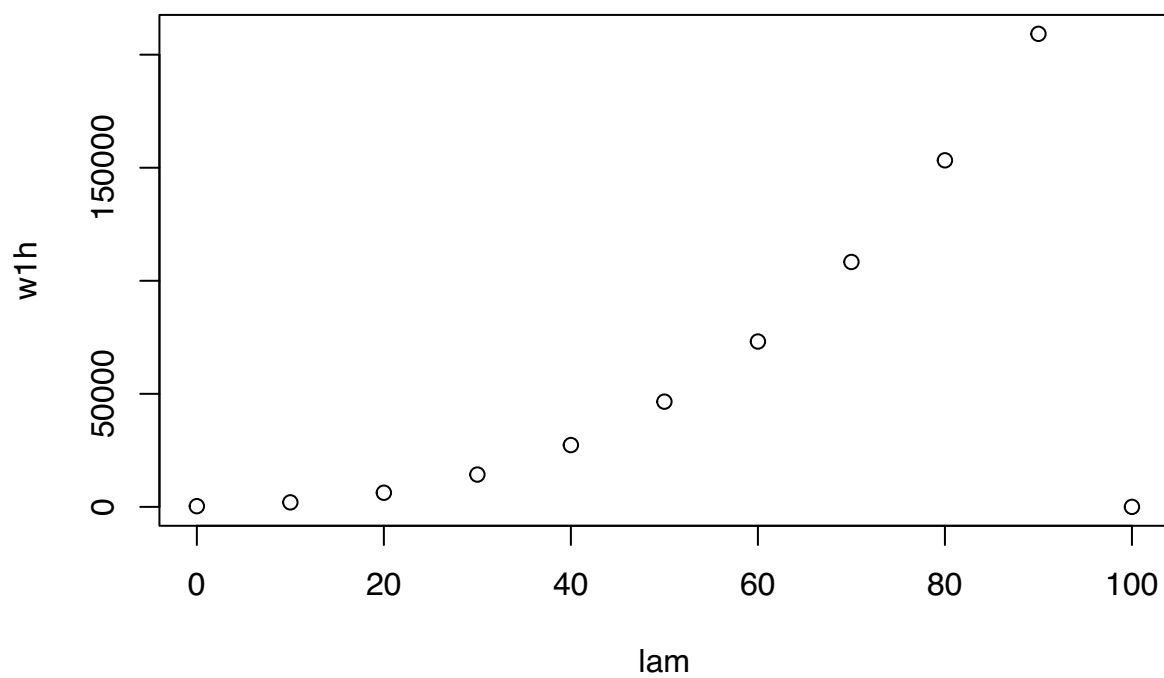
ti <- 1 + 0.2*x -1*x^2 +e

w1l <- function(lam1,l){
  w1l <- e%*%e+lam1*sqrt(sum(w1*w1))
  w1l
}

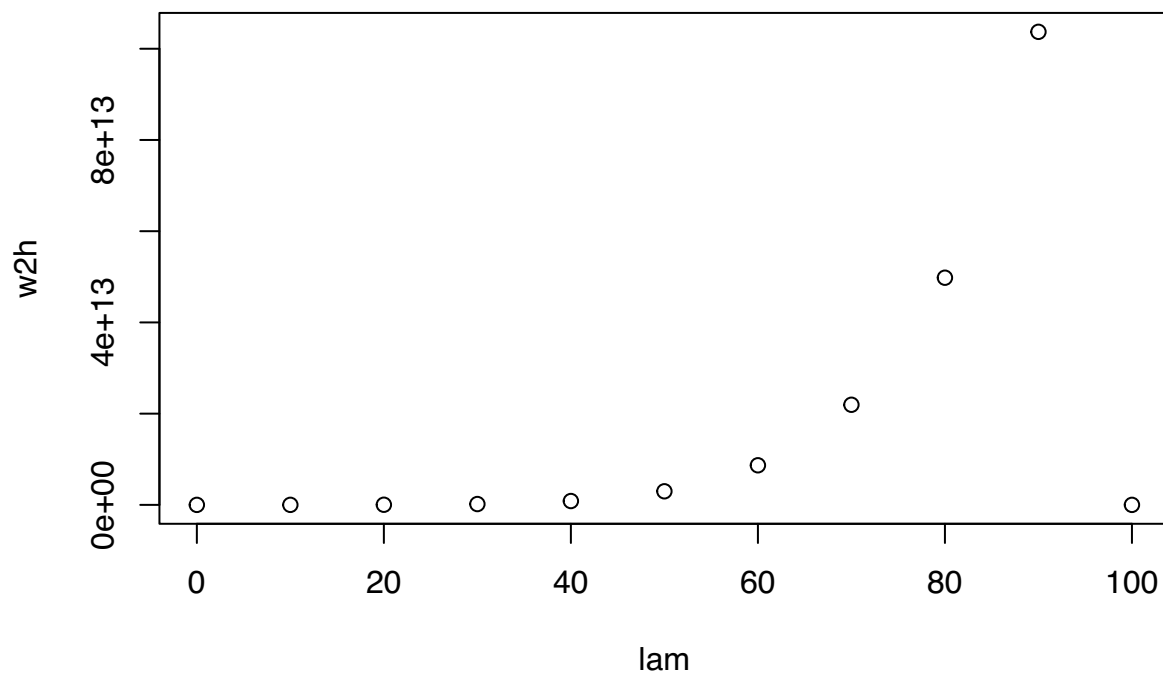
w12 <- function(lam2,l){
  w2l <- e%*%e+lam2*(sum(w2*w2))
  w2l
}

w1h <- rep(0,11)
w2h <- rep(0,11)
h=0
lam = c(0,10,20,30,40,50,60,70,80,90,100)
for( i in lam){
  w1 = 0.2
  w2 = 1
  for(k in 1: 3){
    w1 <- w1l(lam1 = i,l=k)
    w2 <- w12(lam2 = i,l=k)
  }
  w1h[h] = w1
  w2h[h] = w2
  h=h+1
}

plot(lam,w1h)
```



```
plot(lam,w2h)
```



d)

w1,2 is increasing when lambda1,2 increasing

## problem 6

d)

```
library("readxl")
train<-read_excel("training.xlsx",col_names = c("x1","x2","group"))
fit1 <- glm(group~ x1+x2,data=train)
summary(fit1)
```

```
##
## Call:
## glm(formula = group ~ x1 + x2, data = train)
##
## Deviance Residuals:
##      Min       1Q   Median       3Q      Max
## -0.76184  -0.15292   0.00143   0.17241   0.63974
##
## Coefficients:
```

```
##           Estimate Std. Error t value Pr(>|t|)
## (Intercept)  2.194053   0.043108  50.897  <2e-16 ***
## x1          -0.214846   0.007725 -27.813  <2e-16 ***
## x2           0.017735   0.014143   1.254    0.211
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## (Dispersion parameter for gaussian family taken to be 0.06226525)
##
## Null deviance: 66.667  on 299  degrees of freedom
## Residual deviance: 18.493  on 297  degrees of freedom
## AIC: 23.442
##
## Number of Fisher Scoring iterations: 2
```

```
fit2 <- glm(group~ x1+x2-1,data = train)
summary(fit2)
```

```
##
## Call:
## glm(formula = group ~ x1 + x2 - 1, data = train)
##
## Deviance Residuals:
##      Min       1Q   Median       3Q      Max
## -1.6618  -0.2639   0.1928   0.7601   1.9616
##
## Coefficients:
##      Estimate Std. Error t value Pr(>|t|)
## x1 -0.03808    0.02148  -1.773  0.0772 .
## x2  0.60288    0.02564  23.514  <2e-16 ***
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## (Dispersion parameter for gaussian family taken to be 0.6033199)
##
## Null deviance: 900.00  on 300  degrees of freedom
## Residual deviance: 179.79  on 298  degrees of freedom
## AIC: 703.76
##
## Number of Fisher Scoring iterations: 2
```

e)

from the coding in d), AIC for model 1 is 23.442, AIC for model 2 is 703.76

f)

```
trdata1 <- predict(fit1, newdata = train, type = "response")
trdata2 <- predict(fit2, newdata = train, type = "response")
test<-read_excel("test.xlsx",col_names = c("x1","x2","group"))
```

```
testdata1 <- predict(fit1, newdata = test, type = "response" )
testdata2 <- predict(fit2, newdata = test, type = "response")
## training model 1
glm.pred1=rep(1,300)
glm.pred1[trdata1 >1.5]=2
table(glm.pred1,factor(train$group))
```

```
##
## glm.pred1    1    2
##           1  90   5
##           2  10 195
```

```
## training model 2
glm.pred2=rep(1,300)
glm.pred2[trdata2 >1.5]=2
table(glm.pred2,factor(train$group))
```

```
##
## glm.pred2    1    2
##           1  62 103
##           2  38  97
```

```
## test model 1
glm.test1 = rep(1,40)
glm.test1[testdata1 >1.5]=2
table(glm.test1,factor(test$group))
```

```
##
## glm.test1    1    2
##           1  19   1
##           2   1  19
```

```
## test model 2
glm.test2=rep(1,40)
glm.test2[testdata2 >1.5]=2
table(glm.test2,factor(test$group))
```

```
##
## glm.test2    1    2
##           1  12  12
##           2   8   8
```