

$$1. p(a,b) = p(a,b,c=0) + p(a,b,c=1)$$

$$= \begin{cases} 0.336 & a=0 & b=0 \\ 0.264 & a=0 & b=1 \\ 0.256 & a=1 & b=0 \\ 0.144 & a=1 & b=1 \end{cases}$$

$$p(a) = p(a,b=0) + p(a,b=1) = \begin{cases} 0.6 & a=0 \\ 0.4 & a=1 \end{cases}$$

$$p(b) = p(b,a=0) + p(b,a=1) = \begin{cases} 0.592 & b=0 \\ 0.408 & b=1 \end{cases}$$

Since  $p(a=1,b=1) = 0.144$ , and  $p(a=1)p(b=1) = 0.4 \times 0.408 = 0.1632$

Then  $p(a,b) \neq p(a)p(b)$

$$p(c) = \sum p(a,b,c) = \begin{cases} 0.480 & c=0 \\ 0.520 & c=1 \end{cases}$$

$$p(a,b|c) = \frac{p(a,b,c)}{p(c)} = \begin{cases} 0.400 & a=0, b=0, c=0 \\ 0.277 & a=0, b=0, c=1 \\ 0.100 & a=0, b=1, c=0 \\ 0.415 & a=0, b=1, c=1 \\ 0.400 & a=1, b=0, c=0 \\ 0.123 & a=1, b=0, c=1 \\ 0.100 & a=1, b=1, c=0 \\ 0.185 & a=1, b=1, c=1 \end{cases}$$

$$p(a|c) = \frac{p(a,c)}{p(c)} = \begin{cases} 0.500 & a=0, c=0 \\ 0.692 & a=0, c=1 \\ 0.500 & a=1, c=0 \\ 0.308 & a=1, c=1 \end{cases}$$

$$p(b|c) = \frac{p(b,c)}{p(c)} = \begin{cases} 0.800 & b=0 & c=0 \\ 0.400 & b=0 & c=1 \\ 0.200 & b=1 & c=0 \\ 0.600 & b=1 & c=1 \end{cases}$$

Thus,  $p(a=1, b=1 | c=0) = p(a=1 | c=0) p(b=1 | c=0) = 0.1$

$$p(a,b|c) = p(a|c) p(b|c)$$

$$2. \quad p(a) = p(a,b=0) + p(a,b=1) = \begin{cases} 0.6 & a=0 \\ 0.4 & a=1 \end{cases}$$

$$p(b|c) = \frac{p(b,c)}{p(c)} = \begin{cases} 0.8 & b=0 & c=0 \\ 0.4 & b=0 & c=1 \\ 0.2 & b=1 & c=0 \\ 0.6 & b=1 & c=1 \end{cases}$$

$$p(c|a) = \frac{p(a,c)}{p(a)} = \begin{cases} 0.4 & a=0 & c=0 \\ 0.6 & a=0 & c=1 \\ 0.6 & a=1 & c=0 \\ 0.4 & a=1 & c=1 \end{cases}$$

Thus,  $p(a,b,c) = p(a) p(c|a) p(b|c) = \text{the table.}$



$$3.$$

	$x=0$	$x=1$	$x=2$
$y=0$	0.1	0.1	0.2
$y=1$	0.1	0.1	0.1
$y=2$	0.2	0.1	0

4. By Bayes Thm :  $p(\theta|x) \propto p(x|\theta)p(\theta)$

$$\ln p(\theta|x) \propto \ln p(x|\theta) + \ln p(\theta)$$

$$\begin{aligned} \ln p(\theta|x) &\propto \ln \left\{ \sum_z p(x, z|\theta) \right\} + \ln p(\theta) \\ &= \ln \left[ \left[ \sum_z p(x, z|\theta) \right] \cdot p(\theta) \right] \end{aligned}$$

$$\begin{aligned} \text{Then } Q'(\theta, \theta_{old}) &= \sum_z p(z|x, \theta_{old}) \ln [p(x, z|\theta) p(\theta)] \\ &= \sum_z p(z|x, \theta_{old}) (\ln p(x, z|\theta) + \ln p(\theta)) \\ &= \sum_z p(z|x, \theta_{old}) \ln p(x, z|\theta) + \ln p(\theta) \\ &= Q(\theta, \theta_{old}) + \ln p(\theta) \end{aligned}$$

$$5. \frac{\partial \ln p}{\partial \Sigma} = \frac{\partial}{\partial \Sigma} \left\{ \sum_{n=1}^N \ln a_n \right\} = \sum_{n=1}^N \frac{1}{a_n} \frac{\partial a_n}{\partial \Sigma}$$

$$\text{where } a_n = \sum_{k=1}^K \pi_k \mathcal{N}(x_n | \mu_k, \Sigma)$$

$$\text{Since } \frac{\partial \ln \mathcal{N}(x_n | \mu_k, \Sigma)}{\partial \Sigma} = -\frac{1}{2} \Sigma^{-1} + \frac{1}{2} \Sigma^{-1} S_{n,k} \Sigma^{-1}$$

$$S_{n,k} = (x_n - \mu_k)(x_n - \mu_k)^T$$

$$\text{we can obtain: } \frac{\partial a_n}{\partial \Sigma} = \frac{\partial}{\partial \Sigma} \left\{ \sum_{k=1}^K \pi_k \mathcal{N}(x_n | \mu_k, \Sigma) \right\}$$

$$\begin{aligned} &= \sum_{k=1}^K \frac{\partial}{\partial \Sigma} \pi_k \mathcal{N}(x_n | \mu_k, \Sigma) \\ &= \sum_{k=1}^K \pi_k \frac{\partial}{\partial \Sigma} \ln \mathcal{N}(x_n | \mu_k, \Sigma) \\ &= \sum_{k=1}^K \pi_k \cdot \exp[\ln \mathcal{N}(x_n | \mu_k, \Sigma)] \cdot \frac{\partial}{\partial \Sigma} [\ln \mathcal{N}(x_n | \mu_k, \Sigma)] \end{aligned}$$

$$= \sum_{k=1}^K \pi_k \mathcal{N}(x_n | \mu_k, \Sigma) \cdot \left( -\frac{1}{2} \Sigma^{-1} + \frac{1}{2} \Sigma^{-1} S_{nk} \Sigma^{-1} \right)$$

$$\begin{aligned} \frac{\partial \ln p}{\partial \Sigma} &= \frac{N}{\sum_{n=1}^N} \frac{1}{a_n} \frac{\partial a_n}{\partial \Sigma} \\ &= \frac{N}{\sum_{n=1}^N} \frac{\sum_{k=1}^K \pi_k \mathcal{N}(x_n | \mu_k, \Sigma) \cdot \left( -\frac{1}{2} \Sigma^{-1} + \frac{1}{2} \Sigma^{-1} S_{nk} \Sigma^{-1} \right)}{\sum_{k=1}^K \pi_k \mathcal{N}(x_n | \mu_k, \Sigma)} \\ &= \sum_{n=1}^N \sum_{k=1}^K \mathcal{N}(z_{nk}) \cdot \left( -\frac{1}{2} \Sigma^{-1} + \frac{1}{2} \Sigma^{-1} S_{nk} \Sigma^{-1} \right) \end{aligned}$$

$$\text{Set } = -\frac{1}{2} \left( \sum_{n=1}^N \sum_{k=1}^K \mathcal{N}(z_{nk}) \Sigma^{-1} + \frac{1}{2} \Sigma^{-1} \left[ \sum_{n=1}^N \sum_{k=1}^K \mathcal{N}(z_{nk}) S_{nk} \right] \Sigma^{-1} \right) = 0$$

$$\text{Then } \Sigma = \frac{\sum_{n=1}^N \sum_{k=1}^K \mathcal{N}(z_{nk}) S_{nk}}{\sum_{n=1}^N \sum_{k=1}^K \mathcal{N}(z_{nk})}$$

6. assume  $k$  is fixed for Beta prior  $\mu_k$

$$p(\mu_k | a_k, b_k) = \frac{\Gamma(a_k + b_k)}{\Gamma(a_k) \Gamma(b_k)} \mu_k^{a_k-1} (1-\mu_k)^{b_k-1}$$

$$\text{Contribution to } p(\theta) = \sum_{k=1}^K \sum_{i=1}^D (a_i - 1) \ln \mu_{ki} + (b_i - 1) \ln (1 - \mu_{ki})$$

$$p(\pi | \alpha) = \frac{\Gamma(\alpha_0)}{\prod_{k=1}^K \Gamma(\alpha_k)} \prod_{k=1}^K \pi_k^{\alpha_k-1}$$

Then

$$\text{Contribution to Dirichlet prior } \ln p(\theta) = \sum_{k=1}^K (a_k - 1) \ln \pi_k$$

$$\begin{aligned} Q'(\theta, \theta_{\text{old}}) &= E_Z(\ln p) + \sum_{k=1}^K \sum_{i=1}^D [(a_i - 1) \ln \mu_{ki} + (b_i - 1) \ln (1 - \mu_{ki})] \\ &\quad + \sum_{k=1}^K (a_k - 1) \ln \pi_k \end{aligned}$$

$$\begin{aligned}\frac{\partial Q'}{\partial \mu_{ki}} &= \frac{\partial E_z(\ln p)}{\partial \mu_{ki}} + \frac{a_i - 1}{\mu_{ki}} - \frac{b_i - 1}{1 - \mu_{ki}} \\ &= \frac{N_k \bar{x}_{ki} + a_i - 1}{\mu_{ki}} - \frac{N_k - N_{k\cdot} \bar{x}_{ki} + b_i - 1}{1 - \mu_{ki}}\end{aligned}$$

$$\sum_{n=1}^N x_{ni} \cdot r(z_{nk}) = N_k \left[ \frac{1}{N_k} \sum_{n=1}^N x_{ni} \cdot r(n_k) \right] = N_k \cdot \bar{x}_{ki}$$

$$\mu_{ki} = \frac{N_k \bar{x}_{ki} + a_i - 1}{N_k + a_i - 1 + b_i - 1}$$

$$L \propto E_z + \sum_{k=1}^K (a_k - 1) \ln \pi_k + \lambda \left( \sum_{k=1}^K \pi_k - 1 \right)$$

$$\frac{\partial L}{\partial \pi_k} = \sum_{n=1}^N \frac{N(z_{nk})}{\pi_k} + \frac{a_k - 1}{\pi_k} + \lambda = 0$$

$$\lambda = - \sum_{k=1}^K N_k - \sum_{k=1}^K (a_k - 1) = -N - a_0 + K$$

$$\pi_k = \frac{\sum_{n=1}^N N(z_{nk}) + a_k - 1}{- \lambda} = \frac{N_k + a_k - 1}{N + a_0 - K}$$

## hw3

Enbo Tian

2022/2/28

### Graphical Model 1

a)

```
rm(list = ls())
X <- c("cold", "hot", "mild")
# day 0
day0 = replicate(5, sample(X, size = 1, prob = c(1/3, 1/3, 1/3)))
# function of day
dayk <- function(day){
  dayk <- rep(0, 5)
  for (i in 1:5){
    if (day[i] == "cold"){
      dayk[i] = sample(c("cold", "hot", "mild"), size = 1, prob = c(1/2, 1/4,
1/4))
    }
    else if (day[i] == "hot"){
      dayk[i] = sample(c("cold", "hot", "mild"), size = 1, prob = c(1/3, 1/3,
1/3))
    }
    else if (day[i] == "mild"){
      dayk[i] = sample(c("cold", "hot", "mild"), size = 1, prob = c(1/4, 1/4,
1/2))
    }
  }
  dayk
}
# day 1:5
day1 <- dayk(day0)
day2 <- dayk(day1)
day3 <- dayk(day2)
day4 <- dayk(day3)

day <- data.frame(day0, day1, day2, day3, day4)
day

##   day0 day1 day2 day3 day4
## 1 mild mild  hot  hot mild
## 2 mild  hot cold  hot  hot
## 3 cold cold mild mild cold
```

```
## 4 hot cold hot mild mild
## 5 mild cold hot mild mild
```

b)

```
# P(day0)
p0 <- c(1/3,1/3,1/3)
# p (k given k-1)
pgiven <- matrix(c(1/2,1/4,1/4,1/3,1/3,1/3,1/4,1/4,1/2),ncol=3)
# marginal prob
p1 <- pgiven%%p0
p2 <- pgiven%%p1
p3 <- pgiven%%p2
margp <- data.frame(p0,p1,p2,p3)
margp

##           p0           p1           p2           p3
## 1 0.3333333 0.3611111 0.3634259 0.3636188
## 2 0.3333333 0.2777778 0.2731481 0.2727623
## 3 0.3333333 0.3611111 0.3634259 0.3636188
```

c)

```
# 3/2hot
p3g2 <- c(1/3,1/3,1/3)
p3g2

## [1] 0.3333333 0.3333333 0.3333333

#  $p(1/2="hot") = p(2/1)*p(1)/p(2 = "hot")$ 
p1 <- c(p1)
p2 <- c(p2)
p3 <- c(p3)
p1g2 <- pgiven * p1 / p2
p1g2 <- c(p1g2[,2])
p1g2

## [1] 0.3312102 0.3389831 0.3312102

#  $p(0/1) = p(1/0)*p(0)/p(1/2="hot")$ 
p0g1 <- pgiven * p0 / p1g2
p0g1

##           [,1]      [,2]      [,3]
## [1,] 0.5032051 0.3354701 0.2516026
## [2,] 0.2458333 0.3277778 0.2458333
## [3,] 0.2516026 0.3354701 0.5032051
```

d)

```
# give day2 is hot
day2 = "hot"
# get most probable day1
if (max(p1g2) == p1g2[1]){
```

```

    day1 = "cold"
    i = 1
  }else if(max(p1g2) == p1g2[2]){
    day1 = "hot"
    i = 2
  }else if(max(p1g2) == p1g2[3]){
    day1 = "mild"
    i = 3
  }
  # get most probable day0
  new_p0g1 <- p0g1[,i]
  if (max(new_p0g1) == new_p0g1[1]){
    day0 = "cold"
  }else if(max(new_p0g1) == new_p0g1[2]){
    day0 = "hot"
  }else if(max(new_p0g1) == new_p0g1[3]){
    day0 = "mild"
  }
  # Same prob for day3 given hot of day2
  c(day0,day1,day2)

## [1] "cold" "hot" "hot"

```

the most probable report for day 0 to 2 are “cold” “hot” “hot”, and we have the same probability for day3.

## Graphical Model 2

a)

```

mi <- c(-2, 2 ,0)
# height function
height <- function(statep){
  m <- replicate(5,sample(mi,size = 1,prob = statep))
  y <- rep(0,5)
  for (i in 1:5){
    y[i] <- rnorm(1,m,1)
  }
  y
}
# get height
y0 <- height(p0)
y1 <- height(p1)
y2 <- height(p2)
y3 <- height(p3)
datay <- data.frame(y0,y1,y2,y3)
datay

```



```
##          y0          y1          y2          y3
## 1 2.511762 -1.536019  0.03102007 -0.88238828
## 2 2.133691 -2.829667 -0.71771823 -0.03544213
## 3 2.609272 -2.645629 -0.45645577  1.62213386
## 4 1.782848 -1.309067 -0.26663822 -0.52945762
## 5 1.400379 -2.158521 -0.42731199  0.83383134
```

b)

```
m <- c(2,0,-2,-2)
y0 <- replicate(5,rnorm(1,m[1],1))
y1 <- replicate(5,rnorm(1,m[2],1))
y2 <- replicate(5,rnorm(1,m[3],1))
y3 <- replicate(5,rnorm(1,m[4],1))
data2y<- data.frame(y0,y1,y2,y3)
data2y

##          y0          y1          y2          y3
## 1 1.6746606 -0.7783034 -1.6931646 -1.7411468
## 2 1.3366571 -0.6038038  0.1484034 -3.7888110
## 3 0.9875342  1.2656318 -2.1430355 -1.2360920
## 4 1.7316862 -1.1488038 -1.0097901 -0.8994421
## 5 2.4540002 -0.9465946 -1.6004229 -1.1189140
```

c)

```
Y0 =0.7
Y1 =1.5
Y2 =-1.8
Y3 =-1

#p(y0/x0)
d0 <- rep(0,3)
d0[1] <- dnorm(Y0,-2,1)
d0[2] <- dnorm(Y0,0,1)
d0[3] <- dnorm(Y0,2,1)
#p(y1/x1)
d1 <- rep(0,3)
d1[1] <- dnorm(Y1,-2,1)
d1[2] <- dnorm(Y1,0,1)
d1[3] <- dnorm(Y1,2,1)
#p(y2/x2)
d2 <- rep(0,3)
d2[1] <- dnorm(Y2,-2,1)
d2[2] <- dnorm(Y2,0,1)
d2[3] <- dnorm(Y2,2,1)
#p(y3/x3)
d3 <- rep(0,3)
d3[1] <- dnorm(Y3,-2,1)
d3[2] <- dnorm(Y3,0,1)
d3[3] <- dnorm(Y3,2,1)
# p(x0)*p(y0/x0)*p(x1/x0)*p(y1/x1)*p(x2/x1)*p(y2/x2)*p(x3/x2)*p(y3/x3)
```

```

mp <- p0*d0*pgiven*d1*pgiven*d2*pgiven*d3
sum(mp[,1])

## [1] 4.060199e-06

sum(mp[,2])

## [1] 9.549811e-06

sum(mp[,3])

## [1] 4.031672e-06

```

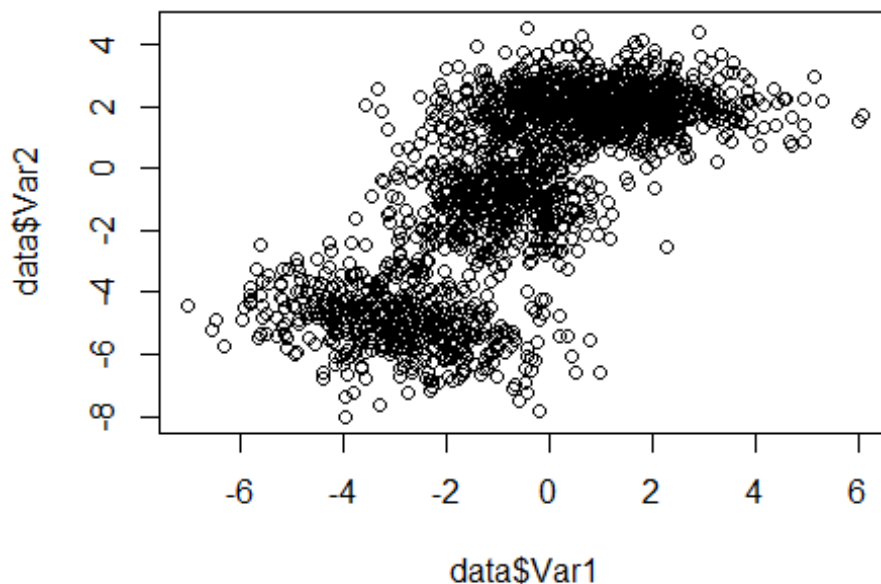
## Gaussian Mixture Model

a)

```

rm(list = ls())
library("readxl")
data <- read_excel("gmm_data.xlsx")
plot(data$Var1, data$Var2)

```



of clusters is 3.

the number

b)

```

library(fMultivar)

## 载入需要的程辑包: timeDate

```

```

## 载入需要的程辑包: timeSeries

## 载入需要的程辑包: fBasics

msnFit(data)

##
## Title:
## Skew Normal Parameter Estimation
##
## Call:
## msnFit(x = data)
##
## Model:
## Skew Normal Distribution
##
## Estimated Parameter(s):
## $beta
##           Var1      Var2
## [1,] 1.455257 3.191867
##
## $Omega
##           Var1      Var2
## Var1  8.146794 11.70847
## Var2 11.708474 22.46951
##
## $alpha
##           Var1      Var2
## -0.8598872 -10.2055688
##
##
## Description:
## Mon Feb 28 19:07:39 2022 by user: 11193

library(ellipse)

##
## 载入程辑包: 'ellipse'

## The following object is masked from 'package:graphics':
##
##      pairs

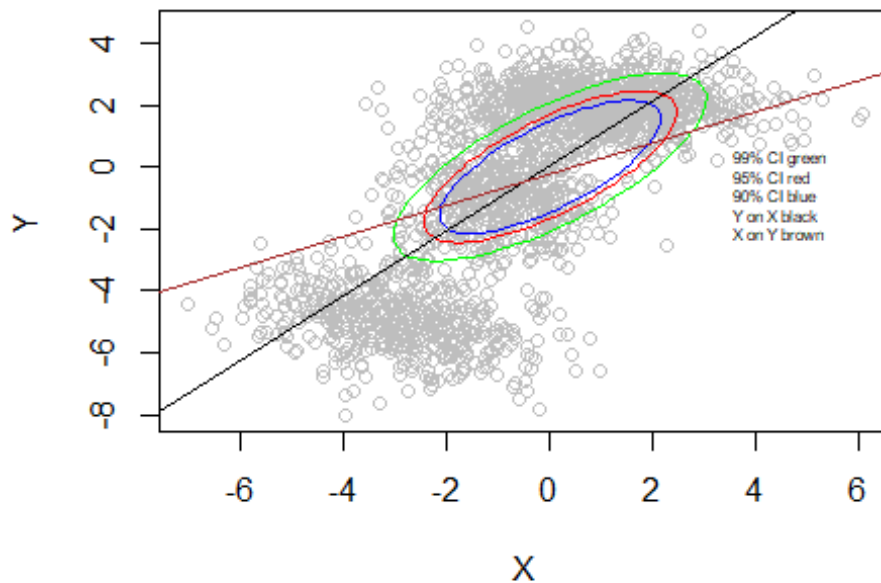
rho = cor(data)
y_on_x <- lm(data$Var2 ~ data$Var1)
x_on_y <- lm(data$Var1 ~ data$Var2)
plot_legend <- c("99% CI green", "95% CI red", "90% CI blue",
                 "Y on X black", "X on Y brown")
plot(data, xlab = "X", ylab = "Y", col = "grey")
lines(ellipse(rho), col="red")
lines(ellipse(rho, level = .99), col="green")

```

```

lines(ellipse(rho, level = .90), col="blue")
abline(y_on_x)
abline(x_on_y, col="brown")
legend(3,1,legend=plot_legend,cex = .5, bty = "n")

```



## c)

```

library(MGMM)
d <- as.matrix(data)
K2GMM <- FitGMM(d,k=2)

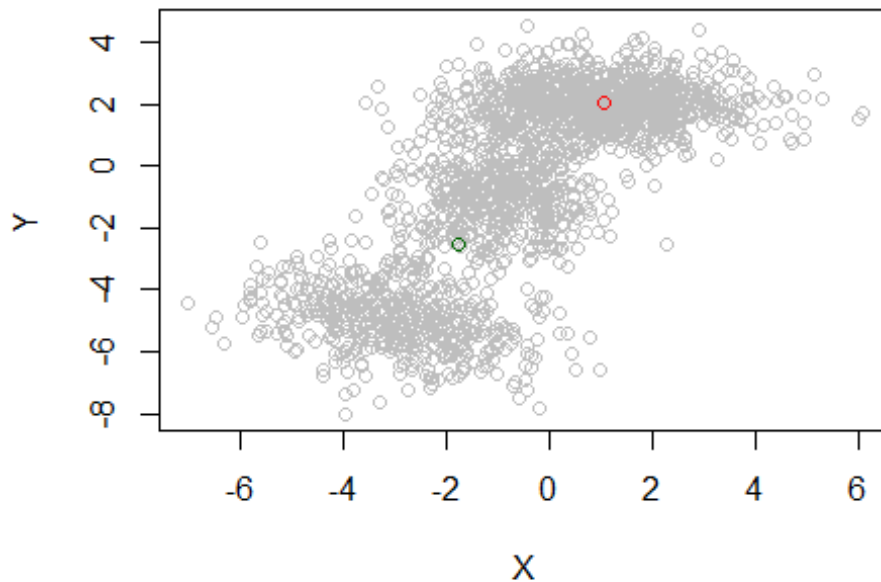
## Objective increment: 10.8
## Objective increment: 0.793
## Objective increment: 0.21
## Objective increment: 0.184
## Objective increment: 0.218
## Objective increment: 0.27
## Objective increment: 0.336
## Objective increment: 0.418
## Objective increment: 0.519
## Objective increment: 0.644
## Objective increment: 0.799
## Objective increment: 0.99
## Objective increment: 1.23
## Objective increment: 1.52
## Objective increment: 1.89
## Objective increment: 2.35
## Objective increment: 2.93
## Objective increment: 3.67

```

```
## Objective increment: 4.6
## Objective increment: 5.77
## Objective increment: 7.21
## Objective increment: 8.89
## Objective increment: 10.6
## Objective increment: 12
## Objective increment: 12.2
## Objective increment: 10.8
## Objective increment: 8.22
## Objective increment: 5.53
## Objective increment: 3.51
## Objective increment: 2.22
## Objective increment: 1.42
## Objective increment: 0.928
## Objective increment: 0.614
## Objective increment: 0.411
## Objective increment: 0.276
## Objective increment: 0.187
## Objective increment: 0.127
## Objective increment: 0.0864
## Objective increment: 0.059
## Objective increment: 0.0403
## Objective increment: 0.0276
## Objective increment: 0.0189
## Objective increment: 0.0129
## Objective increment: 0.00888
## Objective increment: 0.0061
## Objective increment: 0.00419
## Objective increment: 0.00288
## Objective increment: 0.00198
## Objective increment: 0.00136
## Objective increment: 0.000935
## Objective increment: 0.000643
## Objective increment: 0.000442
## Objective increment: 0.000304
## Objective increment: 0.000209
## Objective increment: 0.000144
## Objective increment: 9.91e-05
## Objective increment: 6.82e-05
## Objective increment: 4.69e-05
## Objective increment: 3.23e-05
## Objective increment: 2.22e-05
## Objective increment: 1.53e-05
## Objective increment: 1.05e-05
## Objective increment: 7.25e-06
## Objective increment: 4.99e-06
## Objective increment: 3.43e-06
## Objective increment: 2.36e-06
## Objective increment: 1.63e-06
## Objective increment: 1.12e-06
```

```
## Objective increment: 7.71e-07
## 68 update(s) performed before reaching tolerance limit.

plot(data, xlab = "X", ylab = "Y", col = "grey")
points(K2GMM@Means[[1]][1], K2GMM@Means[[1]][2], col = "red")
points(K2GMM@Means[[2]][1], K2GMM@Means[[2]][2], col = "dark green")
```



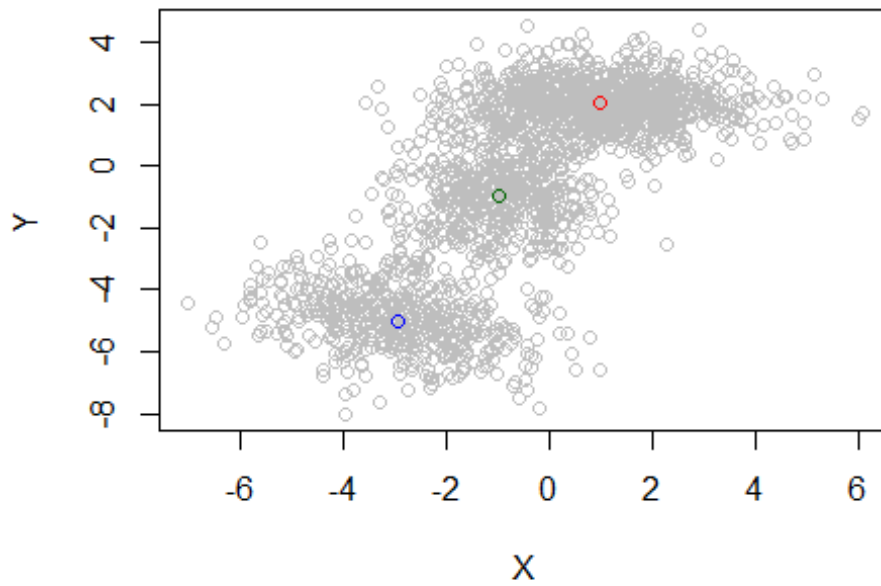
## d)

```
K3GMM <- FitGMM(d, k=3)

## Objective increment: 17.1
## Objective increment: 3.64
## Objective increment: 1.16
## Objective increment: 0.414
## Objective increment: 0.163
## Objective increment: 0.0716
## Objective increment: 0.0356
## Objective increment: 0.02
## Objective increment: 0.0126
## Objective increment: 0.00851
## Objective increment: 0.00604
## Objective increment: 0.0044
## Objective increment: 0.00325
## Objective increment: 0.00241
## Objective increment: 0.0018
## Objective increment: 0.00134
## Objective increment: 0.000998
## Objective increment: 0.000743
```

```
## Objective increment: 0.000553
## Objective increment: 0.000412
## Objective increment: 0.000306
## Objective increment: 0.000228
## Objective increment: 0.000169
## Objective increment: 0.000126
## Objective increment: 9.32e-05
## Objective increment: 6.91e-05
## Objective increment: 5.13e-05
## Objective increment: 3.8e-05
## Objective increment: 2.82e-05
## Objective increment: 2.09e-05
## Objective increment: 1.55e-05
## Objective increment: 1.15e-05
## Objective increment: 8.52e-06
## Objective increment: 6.32e-06
## Objective increment: 4.68e-06
## Objective increment: 3.47e-06
## Objective increment: 2.57e-06
## Objective increment: 1.9e-06
## Objective increment: 1.41e-06
## Objective increment: 1.05e-06
## Objective increment: 7.74e-07
## 40 update(s) performed before reaching tolerance limit.

plot(data, xlab = "X", ylab = "Y", col = "grey")
points(K3GMM@Means[[1]][1], K3GMM@Means[[1]][2], col = "red")
points(K3GMM@Means[[2]][1], K3GMM@Means[[2]][2], col = "dark green")
points(K3GMM@Means[[3]][1], K3GMM@Means[[3]][2], col = "blue")
```



## e)

```
# Sigma_i = dx * d covariance matrix
si <- diff(data$Var1) %%% diff(rho)
si[1:20,]
```

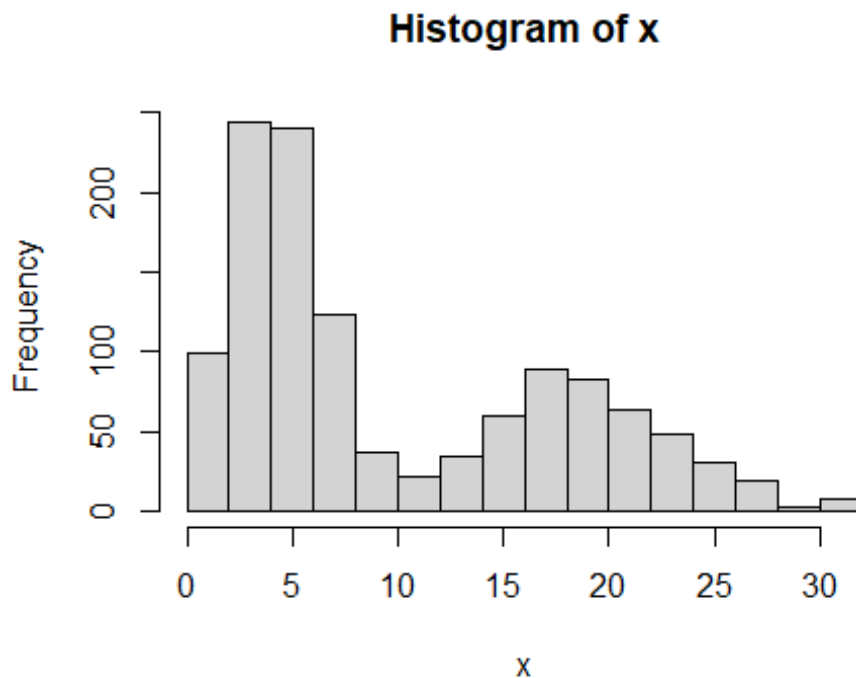
```
##           Var1      Var2
## [1,]  0.2714496 -0.2714496
## [2,]  0.7881399 -0.7881399
## [3,]  0.7121274 -0.7121274
## [4,] -0.8031689  0.8031689
## [5,] -0.6439405  0.6439405
## [6,] -0.3233667  0.3233667
## [7,]  0.8456699 -0.8456699
## [8,] -1.2999374  1.2999374
## [9,]  1.1007313 -1.1007313
## [10,] 0.4108798 -0.4108798
## [11,] 0.1088834 -0.1088834
## [12,] -0.4082431  0.4082431
## [13,] 0.5513425 -0.5513425
## [14,] -0.1791793  0.1791793
## [15,] 0.5668767 -0.5668767
## [16,] -1.3834752  1.3834752
## [17,] 0.2978884 -0.2978884
## [18,] -0.1558065  0.1558065
## [19,] 1.0285411 -1.0285411
## [20,] -0.3159542  0.3159542
```



## Poisson Mixture Model

1)

```
rm(list = ls())  
  
library("readxl")  
data <- read_excel("poisson_data.xlsx")  
x <- data$X  
hist(x)
```

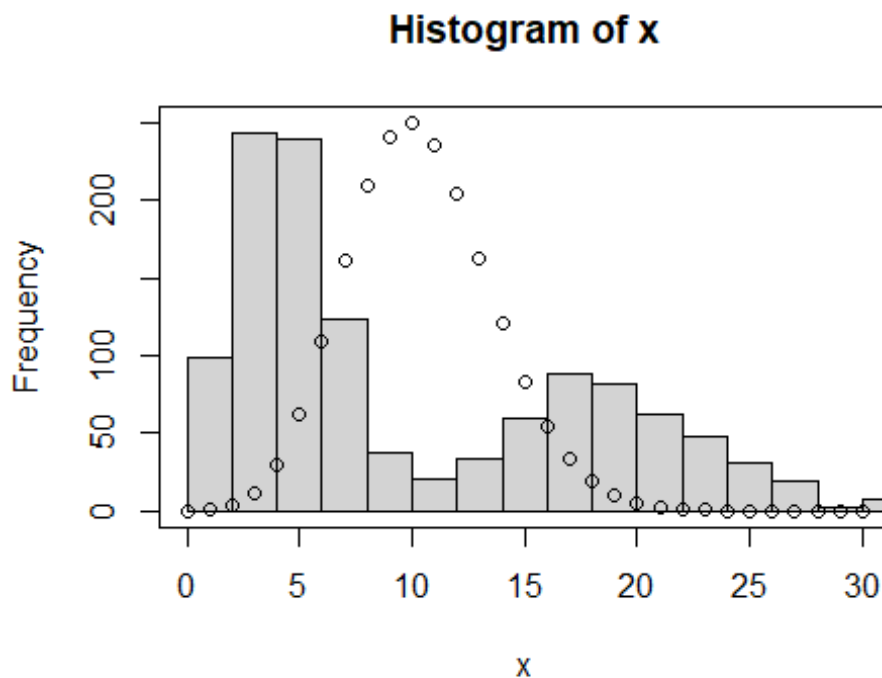


From the

plot it may fit two distribution from 1-10 and 10-30

b)

```
library(MASS)  
lambda <- fitdistr(x, densfun="Poisson")  
lambda  
  
##      lambda  
## 10.37833333  
## ( 0.09299791)  
  
a <- 0:30  
b <- dpois(a, lambda$estimate)  
hist(x, xlim=c(0,30), ylim=c(0,250))  
par(new=TRUE)  
plot(a, b, yaxt="n", xaxt="n", xlab="", ylab="")
```



a simple

poisson distribution is not a good fit.

c)

```
library(mixtools)

## mixtools package, version 1.2.0, Released 2020-02-05
## This package is based upon work supported by the National Science Fo
undation under Grant No. SES-0518772.

##
## 载入程辑包: 'mixtools'

## The following object is masked from 'package:ellipse':
##
##      ellipse

mixture <- normalmixEM(x,lambda = 0.092997,k=2)

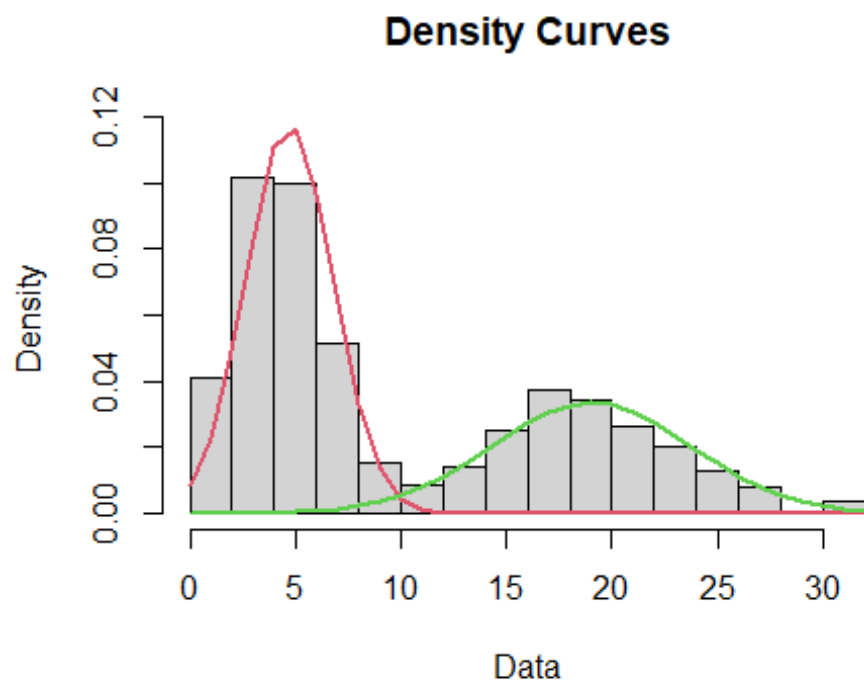
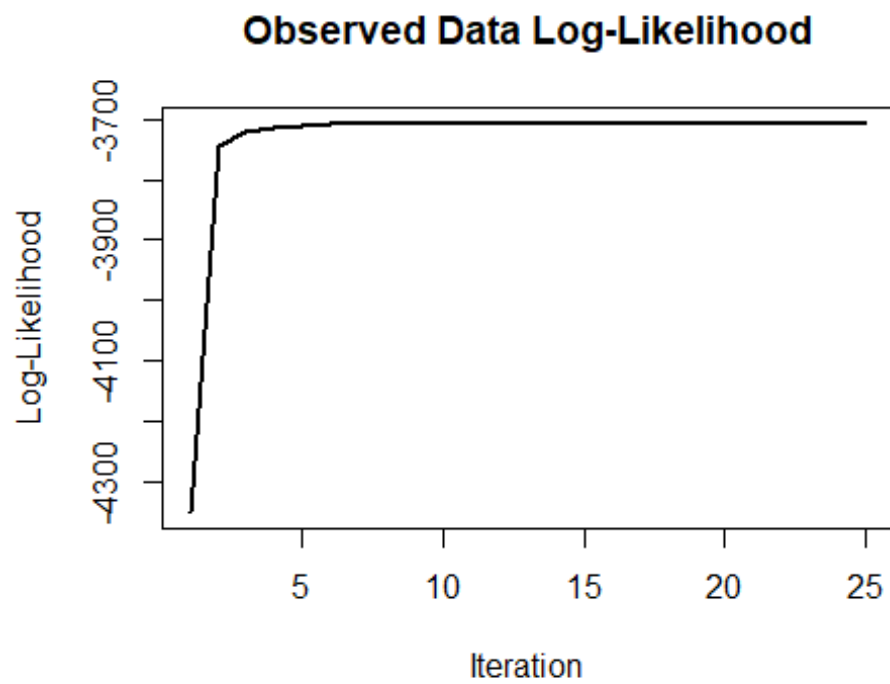
## number of iterations= 24

summary(mixture)

## summary of normalmixEM object:
##      comp 1      comp 2
## lambda 0.604878 0.395122
## mu      4.722448 19.036712
## sigma   2.052662 4.708305
## loglik at estimate: -3707.207
```

d)

```
plot(mixture, density=TRUE)
```



e)

the single poisson distribution can not fit the data very well, the mixture model give two poisson distribution and separate the data into two modle, which fit the data better.