

1. Since we have  $E[f(z_n)] = \sum_{l=1}^L w_n^{(l)} f(z_n^{(l)})$

let  $\{z_n^{(l)}\}$  is a set of  $p(z_n | x_{n-1})$

$$w_n^{(l)} = \frac{p(x_n | z_n^{(l)})}{\sum_{m=1}^L p(x_n | z_n^{(m)})}$$

2.  $E(\hat{f}) = E\left[\frac{1}{L} \sum_{l=1}^L f(z^{(l)})\right]$

$$= \frac{1}{L} \sum_{l=1}^L E[f(z^{(l)})]$$

$$= E[f]$$

$$\text{var}[\hat{f}] = E[\hat{f}^2] - E[\hat{f}]^2 = E\left[\left(\frac{1}{L} \sum_{l=1}^L f(z^{(l)})\right)^2\right] - E[f]^2$$

$$= \frac{1}{L^2} E\left[\left(\sum f(z^{(l)})\right)^2\right] - E[f]^2$$

$$= \frac{1}{L^2} E\left[\sum_{l=1}^L f^2(z^{(l)}) + \sum_{i,j=1, i \neq j}^L f(z^{(i)}) f(z^{(j)})\right] - E[f]^2$$

$$= \frac{1}{L^2} E\left[\sum_{l=1}^L f^2(z^{(l)})\right] + \frac{L^2 - L}{L^2} E[f]^2 - E[f]^2$$

$$= \frac{1}{L^2} \sum_{l=1}^L E[f^2(z^{(l)})] - \frac{1}{L} E[f]^2$$

$$= \frac{1}{L^2} \cdot L \cdot E(f^2) - \frac{1}{L} E(f)^2$$

$$= \frac{1}{L} E(f^2) - \frac{1}{L} E(f)^2 = \frac{1}{L} E[(f - E(f))^2]$$

3. Since we have  $q(z) = p(y) \cdot \left| \frac{dy}{dz} \right|$

$$y = \arctan \frac{z-c}{b} \Rightarrow \frac{dy}{dz} = \frac{1}{b} \cdot \frac{1}{1 + [(z-c)/b]^2}$$

$$q(z) = 1 \cdot \frac{1}{b} \cdot \frac{1}{1 + [(z-c)/b]^2} \text{ is a pdf.}$$

$$\text{By (11.16)} \quad k = \frac{1}{b}.$$

## Application Questions

Since there's some compelling was too slow, I can only do the screen shot for the problems, but not knit it to pdf.

```
1 ---
2 title: "cs539 hw4 Application Questions"
3 author: "Enbo Tian"
4 date: '2022-03-25'
5 output: pdf_document
6 ---
7 # Viterbi Decoding Algorithm
8 ## a)
9 ```{r}
10 library(HMM)
11 # Viterbi algorithm
12 Viterbi=function(v,a,b,initial_distribution) {
13
14   T = length(v)
15   M = nrow(a)
16   prev = matrix(0, T-1, M)
17   omega = matrix(0, M, T)
18
19   omega[, 1] = log(initial_distribution * b[, v[1]])
20   for(t in 2:T){
21     for(s in 1:M) {
22       probs = omega[, t - 1] + log(a[, s]) + log(b[s, v[t]])
23       prev[t - 1, s] = which.max(probs)
24       omega[s, t] = max(probs)
25     }
26   }
27
28   S = rep(0, T)
29   last_state=which.max(omega[,ncol(omega)])
30   S[1]=last_state
31
32   j=2
33   for(i in (T-1):1){
34     S[j]=prev[i,last_state]
35     last_state=prev[i,last_state]
36     j=j+1
37   }
38
39   S[which(S==1)]='A'
40   S[which(S==2)]='B'
41 }
```

```

38
39     S[which(S==1)]='A'
40     S[which(S==2)]='B'
41
42     S=rev(S)
43
44     return(S)
45
46 }
47
48 #repeat the result presented in the pdf file
49
50 hmm = initHMM(c("H","L"), c("A","C","G","T"),startProbs = c(0.5,0.5),
51 transProbs=matrix(c(0.5,0.5,0.4,0.6),2),
52 emissionProbs=matrix(c(0.2,0.3,0.3,0.2,0.3,0.2,0.2,0.3),2))
53
54 observations = c("G","G","C","A","C","T","G","A","A")
55 viterbi = viterbi(hmm,observations)
56 viterbi

```

```
[1] "H" "H" "H" "L" "L" "L" "L" "L" "L"
```

```

57
58 ## b)
59 ```{r}
60 #the observed sequence of: AGTCGTA
61 observations = c("A","G","T","C","G","T","A")
62 viterbi = viterbi(hmm,observations)
63 viterbi
64 ```

```

```
[1] "L" "L" "L" "H" "H" "L" "L"
```

```
66 ▾ # Bayesian Filtering
```

```
67 ▾ ## a)
```

```
68 ▾ ```{r}
```

```
69 # create X and Y
```

```
70 x0 <- rnorm(1,0,1)
```

```
71 x = x0
```

```
72 X <- rep(0,100)
```

```
73 Y <- rep(0,100)
```

```
74 ▾ for(i in 1:100){
```

```
75   X[i] <- rnorm(1,0.99*x+0.1,0.1)
```

```
76   x <- X[i]
```

```
77   Y[i] <- rnorm(1,-2*x+1,0.4)
```

```
78 ▴ }
```

```
79 X <- c(x0,X)
```

```
80 head(X)
```

```
81 head(Y)
```

```
82 ▴
```

```
[1] 1.082742 1.113829 1.268545 1.397311 1.623790 1.670091
```

```
[1] -1.596619 -1.715995 -1.250474 -2.215364 -2.754681 -2.480468
```

```
83
```

```
84 ▾ ## b)
```

```
85 ▾ ```{r}
```

```
86 pygx <- pnorm(Y,-2*X[-1],0.4)
```

```
87 pxgx1 <- pnorm(X[-1],0.99*X[1:100]+0.1,0.1)
```

```
88 pxkgy1k1 <- pygx*pxgx1
```

```
89 z <- pygx[100]*pxkgy1k1[100]
```

```
90 pxkgy1k <- 1/z * pygx *pxkgy1k1
```

```
91 head(pxkgy1k)
```

```
92 ▴
```

```
[1] 0.2707539 0.7765434 0.7172082 0.9888070 0.3311852 0.6504082
```

```
93
```

```

94 ▾ ## c)
95 ▾ ```{r}
96 post <- qnorm(pxkgy1k,0.99*X[1:101]+0.1,0.1)
97 post = post[-100]
98 mean(post)
99 conf <- post/X[-1]
100 alpha <- 1-conf
101 length(alpha[abs(alpha)<0.05])
102 ▴

[1] 2.51459
[1] 1.0160477 0.9991755 0.9986455 0.9975084 1.0124402 0.9994709
[1] -0.0160477017 0.0008244506 0.0013545006 0.0024916007 -0.0124402129
0.0005290819
[1] 95

103
104 ▾ ## d)
105 ▾ ```{r}
106 library("readxl")
107 data<-read_excel("filter_problem.xlsx", col_names = c("index","Xk","Yk"))
108 Y <- data$Yk[-1]
109 X <- data$Xk
110 pygx <- pnorm(Y,-2*X[-1],0.4)
111 pxgx1 <- pnorm(X[-1],0.99*X[1:100]+0.1,0.1)
112 pxkgy1k1 <- pygx*pxgx1
113 z <- pygx[100]*pxkgy1k1[100]
114 pxkgy1k <- 1/z * pygx *pxkgy1k1
115 head(pxkgy1k)
116 post <- qnorm(pxkgy1k,0.99*X[1:101]+0.1,0.1)
117 post = post[-100]
118 mean(post)
119 conf <- post/X[-1]
120 alpha <- 1-conf
121 length(alpha[abs(alpha)<0.05])
122 ▴

[1] 58.38498 73.13255 74.34234 73.62419 69.11370 67.46193

```

Alpha = [1] 96

```

124 ## e)
125 ```{r}
126 x0 <- rnorm(1,0,1)
127 x = x0
128 X <- rep(0,100)
129 for(i in 1:100){
130   X[i] <- rnorm(1,x,0.2)
131   x <- X[i]
132 }
133 X <- c(x0,X)
134 pygx <- pnorm(Y,-2*X[-1],0.4)
135 pxgx1 <- pnorm(X[-1],X[1:100],0.2)
136 pxkgy1k1 <- pygx*pxgx1
137 z <- pygx[100]*pxkgy1k1[100]
138 pxkgy1k <- 1/z * pygx *pxkgy1k1
139 head(pxkgy1k)
140 post <- qnorm(pxkgy1k,X[1:100],0.2)
141 post = post[-100]
142 mean(post)
143 conf <- post/X[-1]
144 alpha <- 1-conf
145 length(alpha[abs(alpha)<0.05])
146 ```

```

```

[1] 3.752094e+89 7.791352e+88 4.688298e+88 4.846695e+86 2.044845e+79 3.585929e+80
Warning in qnorm(pxkgy1k, X[1:100], 0.2) : 产生了NaNs
[1] NaN
Warning in post/X[-1] :
  longer object length is not a multiple of shorter object length
[1] 33

```

```

148 ## f)
149 Since there is NA value in d and e, y may not follow the distribution which we
150 provided.
151 In addition,c d have a good confidence but e doesn't.
152 # Sequential MCMC
153 ```{r}
154 Y <- data$Yk[-1]
155 X <- data$Xk
156 pygx <- pnorm(Y,-2*X[-1],0.4)
157 pxgx1 <- pnorm(X[-1],0.99*X[1:100]+0.1,0.1)
158 pxkgy1k1 <- pygx*pxgx1
159 z <- pygx[100]*pxkgy1k1[100]
160 pxkgy1k <- 1/z * pygx *pxkgy1k1
161 #a)
162 mean(sample(pxkgy1k,100))
163 #b)
164 mean(sample(pxkgy1k,1000,replace = TRUE))
165 ```

```

```

[1] 31.32629
[1] 32.46146

```

```
# GP and Linear Regression
```

```
##a)
```

```
##{r}
```

```
library(GPfit)
```

```
library(brms)
```

```
data <- read.csv("synchronous_machine.csv")
```

```
### It is too slow for my computer to do the chain with sampling in the model,
```

```
### so I can only pick one parameter for Gp. (nearly 10 min once)
```

```
fitgp <- brm(qIy~gp(PF)+e+dIf+If,chain = 4, data = data)
```

```
summary(fitgp)
```

Warning: Parts of the model have not converged (some Rhats are > 1.05). Be careful when analysing the results! We recommend running more iterations and/or setting stronger priors.

Family: gaussian

Links: mu = identity; sigma = identity

Formula: qIy ~ gp(PF) + e + dIf + If

Data: data (Number of observations: 557)

Draws: 4 chains, each with iter = 2000; warmup = 1000; thin = 1;  
total post-warmup draws = 4000

Gaussian Process Terms:

	Estimate	Est.Error	l-95% CI	u-95% CI	Rhat	Bulk_ESS	Tail_ESS
sdgp(gpPF)	0.73	0.33	0.38	1.73	1.63	8	18
lscale(gpPF)	0.11	0.10	0.01	0.30	2.56	5	19

Population-Level Effects:

	Estimate	Est.Error	l-95% CI	u-95% CI	Rhat	Bulk_ESS	Tail_ESS
Intercept	-1091.42	544.66	-2392.03	-350.81	1.75	6	15
e	-7.07	3.61	-10.51	2.58	1.60	7	13
dIf	-920.50	460.75	-2020.75	-293.37	1.75	6	15
If	927.63	461.14	300.54	2028.76	1.75	6	15

Family Specific Parameters:

	Estimate	Est.Error	l-95% CI	u-95% CI	Rhat	Bulk_ESS	Tail_ESS
sigma	0.45	0.09	0.34	0.57	1.84	6	46

Draws were sampled using sampling(NUTS). For each parameter, Bulk\_ESS and Tail\_ESS are effective sample size measures, and Rhat is the potential scale reduction factor on split chains (at convergence, Rhat = 1).



```
## b)
{r}
fit <- lm(qIy~PF+e+dIf+If,data = data)
summary(fit)
```

```
Call:
lm(formula = qIy ~ PF + e + dIf + If, data = data)

Residuals:
    Min       1Q   Median       3Q      Max
-1.70820 -0.39040 -0.03484  0.31406  2.14232

Coefficients: (2 not defined because of singularities)
              Estimate Std. Error t value Pr(>|t|)
(Intercept)  -7.0398     0.4654  -15.13  <2e-16 ***
PF             10.8102     0.4628   23.36  <2e-16 ***
e              NA         NA      NA      NA
dIf            7.4658     0.2663   28.03  <2e-16 ***
If             NA         NA      NA      NA
---
Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

Residual standard error: 0.5767 on 554 degrees of freedom
Multiple R-squared:  0.5872,    Adjusted R-squared:  0.5857
F-statistic: 394 on 2 and 554 DF,  p-value: < 2.2e-16
```

```
## c)
MSE of gp is much less than linear regression.
```