Flexibility of interactions depends on habitat suitability

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16 Abstract

Evidence that the environment influences the interaction between species is rapidly accumulating. However, how it happens is currently unclear as environmental gradients appear to have contrasting or non-linear effects on the species' trophic niche breadth depending on the environmental variable. Here, we explore the relationship between the stresses imposed by the environment, instead of environmental gradients directly, and niche breadth using a global dataset of plant-pollinator interactions. We found that environmental stress plays a significant role in determining the number of partners a species interacts with, but this role is highly variable across species. In particular, when faced with environmental stress, species that have a large number of interactions are more likely to focus on a smaller number of, presumably higher-quality, interactions. Contrastingly, the specialists that can cope with increased stress are more likely to broaden their niche and engage in opportunistic interactions, effectively behaving as facultative generalists.

Introduction

Species interactions are known to vary widely across space and time. There are multiple examples of species that interact tightly in an ecological community or a particular season but not in another. It is crucial to understand why and how this turnover of interactions occur because the network of interactions of an ecological community is known to underpin ecosystem function and stability. Despite much progress in the last decades, we are still unable to predict species interactions, partially because we still have not accurately identified the factors that drive species to interact under certain conditions and not under another. There are two main processes involved in whether two species in a community interact or not. The first is an "Eltonian" niche process that is underpinned by the traits of the interacting species. Compared to other kinds of trophic interactions, mutualistic interactions, and plant-pollination interactions, in particular, are relatively generalised. While a large proportion of the interactions appear to be opportunistic, a fraction of interactions shows a strong signature of deep co-evolutionary history (Hutchinson, Cagua, and Stouffer 2017). In contrast to opportunistic interactions with low traitmatching, the more co-evolved interactions with optimally matched traits are more persistent and presumably lead to a more substantial per-capita benefit for the interacting species. The second is a neutral process underpinned by the encounter rate of the interacting species. All else being equal, when the encounter rate of interacting species is high, there is a higher probability to interact than if the rate were lower, for example, due to species rarity. Perhaps, because it operates at scales larger than those of the Eltonian niche concept, the abiotic environment is one of the factors that is commonly overlooked by community ecologists when explaining interaction turnover (but see Gravel et al. 2018). The most evident—and best-studied—way

through which the environment can affect the structure of ecological networks is by way of patterns of community composition. At the heart of biogeography is the idea that the abiotic environment can determine the "Grinellian" niche of a species, this is whether the species can have a positive growth-rate when rare in a particular location. For two species to interact, they first need to co-occur in the same community, which can only happen if the environmental conditions of the community fall within the Grinellian niche of both species. Although less understood, the environment can also affect the interactions themselves, after community composition has been established (Tylianakis and Morris 2017). For instance, some studies show that trophic interactions, like predation (McKinnon et al. 2010; Vucic-Pestic et al. 2011) and herbivory (Baskett and Schemske 2018), can increase with temperature but might decrease with precipitation (Pires et al. 2016), and both plant-plant (Bertness and Callaway 1994) and plant-pollinator (Tur et al. 2016) interactions can switch from facilitative to competitive along an elevation/temperature gradient. Contrastingly, other studies have shown either no effect or nonlinear effects of temperature or precipitation on plant-pollinator interactions (Devoto, Medan, and Montaldo 2005; Gravel et al. 2018). One way to summarise the seemingly contradictory evidence of whether the environment can influence whether two species interact or not is by moving from gradients of temperature or precipitation to gradients of environmental stress. Previous research is so far equivocal and, taken together, suggest that environmental stress can have two contrasting effects on species' Eltonianniche (Tylianakis and Morris 2017). On the one hand, it is possible that when species are under environmental stress, they might be "pressured" to focus on partners with which they are best adapted to interact. For instance, Hoiss et al. (2012) found increased phylogenetic clustering between plants and pollinators at higher altitudes; while Peralta et al. (2015) found that parasitoids in plantation forest, where environmental stress was higher than

in native forests, were constrained to interact with hosts they were best adapted to attack. Similarly, Lavandero and Tylianakis (2013) found that environmental stress due to higher temperature reduced the breadth of the Eltonian niche of parasitoids.

On the other hand, it is also possible that when species are under environmental stress, they are forced to be more flexible in their interactions
as higher environmental stress is likely to be reflected in greater energetic
or reproductive costs and therefore they might not be able to sustain encounter rates with their preferred partners at sufficient levels. In line with
this hypothesis, Hoiss, Krauss, and Steffan-Dewenter (2015) found that the
specialisation of plant-pollinator networks decreased both with elevation
and after extreme drought events. Likewise, Pellissier et al. (2010) found a
positive relationship between niche breadth and environmental stress: diskor bowl-shaped blossoms (which allow a large number of potential pollinator
species to access pollen and nectar rewards) dominated at high altitude
flower communities.

We analyse data on plant-pollinator interactions globally and calculate the environmental suitability of all species present at two or more communities 105 to. Our main aim is to leverage that information to test the two competing 106 hypotheses that relate environmental stress and Eltionian niche breadth 107 and whether there is a pattern within and across species. Specialist species 108 can become "facultative" generalists to reduce their vulnerability to the 109 absence of preferred partners (for example, when variations in climate de-110 couple phenologies; Benadi et al. 2014). We therefore also expect that 111 as environmental stress increases species with a relatively small number of 112 partners are more likely to broaden their Eltonian niche. Species with a 113 large number of partners, on the other hand, should have a larger pool of 114 available partners and might, therefore, be more likely to narrow their niche 115 under environmental stress. Importantly, when testing these hypotheses, we

control for the potential effects of the environment in community composition and the size of the species fundamental niche, both from an Eltonian (interactions) and Grinellian (environment) perspective.

$_{20}$ Methods

We retrieved plant-pollinator networks from the Web of Life database (Fortuna, Ortega, and Bascompte 2014). This database contains datasets origi-122 nating from 57 studies published in the primary literature between 106 and 123 2016 (Kaiser-Bunbury et al. 2010, 2014; Bartomeus, Vilà, and Santamaría 2008; Arroyo, Primack, and Armesto 1982; Hatterslev-Smith 1985; Barrett 125 and Helenurm 1987; Dicks, Corbet, and Pywell 2002; Dupont, Hansen, and 126 Olesen 2003; Elberling and Olesen 1999; Herrera 1988; Hocking 1968; Inouye 127 and Pyke 1988; Kato et al. 1990; Medan et al. 2002; Memmott 1999; Motten 1986; Olesen, Eskildsen, and Venkatasamy 2002; Ollerton 2003; Petanidou 129 and Vokou 1993; McMullen 106AD; Mosquin 1967; Clements and Long 130 1923; Kevan 1970; Small 1976; Montero 2005; Stald, Valido, and Olesen 2003; Ingversen 2006; Bundgaard 2003; Bek 2006; KATo 2000; Dupont and Olesen 2009; Lundgren and Olesen 2005; Percival 1974; Philipp et al. 2006; Primack 1983; Ramirez 1989; Ramirez and Brito 1992; Schemske et al. 1978; Smith-Ramírez et al. 2005; Vázquez 2002; Robertson 1929; Ferreira 135 Canela 2006; Kohler 2011; Lasprilla and others 2003; Sabatino 2010; Lara 136 2006; Las-Casas, Azevedo Júnior, and Dias Filho 2012; del Coro Arizmendi and Ornelas 1990; Abreu and Vieira 2004; Vizentin-Bugoni et al. 2016; 138 Bezerra, Machado, and Mello 2009; INouE et al. 1990; KATo, Matsumoto, 139 and Kato 1993; Yamazaki and Kato 2003; Gutierrez Z, Rojas-Nossa, and 140 Stiles 2004; Kato and Miura 1993; Kakutani et al. 1990). Calculating the environmental stress of species in their community and their Eltonian niche breadth requires that the scientific names used to identify them are valid and unambiguous. This is because valid scientific names are necessary to

both accurately match species across studies and to obtain other information from biological databases. However, a significant proportion of the data obtained contain spelling mistakes, taxonomic synonyms, or non-standard abbreviations. Therefore, our first step was to validate and correct scientific names across studies to ensure that we matching of species across studies was as accurate as possible.

Second, we calculated the suitability of the environment for a species in a 151 particular community as a proxy of environmental stress. We assume that the 152 environmental stress a species experience in a particular location is inversely 153 related to the suitability of the average environmental conditions in that 154 place. Our third and final step was to relate the environmental suitability 155 to the relative number of partners a species has in a community, as a proxy 156 for Eltionian niche breadth. To explore this relationship within and across species, we used a multilevel Bayesian model in which we controlled for the 158 potential effects of the environment on co-occurrence. 159

160 Plant-pollinator interactions

All together, interaction data from the included studies included 11,231 unique organism names. From these 1,166 were present in more than one study. From the total number of organisms, 159 were identified to the subspecies or variety level, 6,759 to the species level, 1,755 to the genus level, whereas the remaining 2,558 were unidentified. As the species level was the most common taxonomic rank available in our interaction datasets, in all further analysis, we grouped together subspecies or varieties within the same species.

We were able to confirm the validity of 5,263 of the scientific names used to identify organisms (roughly 76%). We assessed the validity of a name by querying the Global Names Resolver database (https://resolver.globalnames.

org) which includes data from 98 taxonomic sources. We accessed this database using the function gnr_resolve from the R package taxize 0.9.6 (Chamberlain and Szocs 2013; S. Chamberlain, Szoecs, et al. 2019).

From the remaining 1,655 names we were unable to validate, we were able to 175 identify and correct 726 that contained spelling mistakes. These spelling mis-176 takes were corrected automatically by fuzzy matching the canonical names in our data sources with those in the Global Names Resolver database. However, 178 on rare occasions, the fuzzy matching algorithm can suggest a scientific name 179 that has a similar spelling, but that corresponds to an organism in a different 180 taxonomic group, often a separate kingdom. To address this potential prob-181 lem, we checked the taxonomic hierarchy of suggested names and confirmed 182 that it matched our expected taxon. We retrieved all taxonomic hierarchies 183 from the National Center for Biotechnology Information taxonomic database (https://www.ncbi.nlm.nih.gov/taxonomy). 185

As species names are constantly changing, we subsequently checked for possible synonyms of the canonical names in our data sources. Using data from the Integrated Taxonomic Information System database (http://www. itis.gov), we found synonyms and alternative names for 611 species.

Finding these alternative names was required for two main reasons. First, because we wanted to be able to identify the cases in which the same species might have been recorded with different names in various data sources. This can occur not only when the canonical name has been changed but also when there are widely used orthographic variants. Second, because retrieving occurrence data is often only possible using the latest accepted/valid name for a particular species.

All together, from the 1,655 names we were unable to validate, it was not possible to automatically correct or find synonyms 332 of them. We then manually consulted multiple online databases, chiefly Wikispecies (https://species.wikimedia.org/), and looked for canonical names that both,

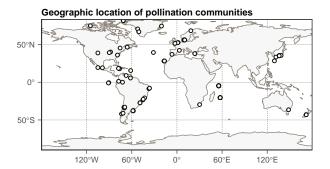


Figure 1: Worldwide distribution of pollination communities included in this study

resembled the unvalidated names and matched the geographic and taxonomic expectations. In this fashion, we were able to further correct 25 names. Most 202 manual corrections were made on names that have been abbreviated or had 203 more than two spelling mistakes. A complete list of manual name corrections 204 can bee seen in Table S1. 205 This cleaning process allowed us to match further 270 names across data 206 sources and, by doing so, identify another 72 species that were present in 207 more than one study. The process also allowed us to identify problematic data 208 sources in which some names were included as both plants and pollinators. 209 These data sources were removed from further analysis. In seven of our data 210 sources interaction data was recorded at multiple points in time. When this was the case, we combined interaction data into one single interaction

The cleaning process resulted on a total of 2,555 plants and 8,406 pollinator species distributed across 73 locations arround the globe (Figure 1, 2).

Species suitability

network.

213

Our next step was to determine the habitat suitability of the species as a proxy of the environmental stress they experience in their community. As we aim to compare the trophic niche for different suitability levels, we only do this for species that were present in at least two communities. To calculate

number of species at multiple locations frequency distribution

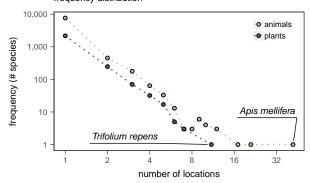


Figure 2: Frequency distribution of the number of locations in which a species is present. The most common pollinator species was *Apis mellifera*, which was sampled on 42 locations, while the most common plant species was *Trifolium repens*, which was sampled on 11 locations.

the suitability of a species in a particular location, we used a (Grinellian) niche-factor analysis (Hirzel et al. 2002; Broennimann et al. 2012). This 222 approach is based on the probability density function of species distribution 223 in an environmental variable space. In a nutshell, habitats (characterised by a collection of environmental variables) in which the species occurs most 225 often are deemed to be suitable for the species than habitats in which the 226 species has never been observed. This approach to estimating the habitat suitability requires two critical pieces of information. First, we require 228 information about the occurrences of the species of interest. Second, we 220 require information about the environmental conditions for all the locations 230 in which the species occurs.

We retrieved occurrences from the Global Biodiversity Information Facility (GBIF; https://www.gbif.org) using the R package rgbif 0.9.6 (Chamberlain and Boettiger 2017; S. Chamberlain, Barve, et al. 2019). Specifically, for each species, we only requested occurrences for which the coordinates of the observation were available and that had no known geospatial issue in the GBIF database. Roughly, we downloaded 38.1 million occurrences for the 986 species we were interested on. This occurrences, however, contain observations of mixed quality. Therefore, we followed (ref?) and applied a

series of simple filters and geographic heuristics to remove those of lower quality. Specifically, we removed all occurrences with (i) a coordinate 241 uncertainty larger than 100km; (ii) those recorded prior to 1945 (ref? or 242 why?); (iii) those in which the number of individuals registered was either zero or more than 100 as such large counts are often erroneous (ref? or why?); and (iv) those occurrences in which the "basis of record" was not a 245 human observation or a preserved specimen (ref? why?). We then used the R package CoordinateCleaner 0.9.6 (Zizka et al. 2019) and land mass and country data from Natural Earth (https://www.naturalearthdata.com) with 248 a 1:10,000,000 scale to further identify and remove problematic occurrences. 249 We removed occurrences for which their coordinates (v) fell outside the 250 borders of the country where they were recorded; (vi) those around a country 251 capital or the centroid of the country and province centroids; (vii) those 252 around a biodiversity institution; and (viii) those located within oceans. 253 Thorugh this cleaning process, we removed with 7.2 million occurrences 254 distributed accross 916 species. Finally, we integrated the occurrences from 255 our plant-pollinator communities to the cleaned occurrences retrieved from 256 GBIF.

We retrieved environmental data from WorldClim V2.0, which includes 19 bioclimatic variables commonly used in species distribution modelling (Fick 259 and Hijmans 2017). We then complemented data obtained from WorldClim 260 with data from Environ (Title and Bemmels 2017), which includes 16 extra 261 bioclimatic and two topographic variables. The additional set of variables from Envirem are relevant to ecological or physiological processes and as 263 such, have the potential to improve our suitability estimation (Title and 264 Bemmels 2018). We obtained all environmental data as rasters composed by cells of 2.5 arc-minutes. We chose this resolution because it provides 266 a reasonable match to the locational accuracy of the species occurrences 267 found in GBIF, particularly those that come preserved specimens in museum

269 collections.

After obtaining information about species occurrence and the environment, 270 we then merged these two datasets such that the location of each occurrence was characterised by a vector with details of our 37 bioclimatic and 272 topographic variables. If a species had more than one occurrence records 273 within one of the cells of the bioclimatic raster, we only included one of the occurrence records. We did this to avoid giving more weight to areas 275 with a high number of occurrences, a common scenario in occurrence records collected opportunistically as the ones we use here. In this step we removed 85.2% of the occurrences which resulted in a total of 4.6 million occurrences 278 used in our niche analysis. 279 A common issue of terrestrial bioclimatic datasets is that the boundaries of 280 the cells with information do not precisely match the landmass boundaries. 281 The result of this missmatch is that not all environmental variables was not 282 available for 3,315 of the raster cells with occurrences (0.8% of the total). 283 As expected, the vast majority of these problematic cells were close to the 284 shore. To address this issue, we calculated the average value of environmental 285 variables within an 5km buffer of the centre of the cell where the variable 286 was missing and used it to approximate the value of the variable in that 287 cell. Using this procedure, we were able to fill environmental variables for 288 89.4% of the cells were they were missing. To fill the remaining 350 cells, we 289 repeated the aforementioned procedure but instead using a 10km buffer. We 290 removed from further analysis occurrences located within the 135 cells for 291 which we were unable to fill environmental variables. Next, we calculate the probability density function of the species distribution 293 in environmental space. To determine the environmental space, we use the first two components from a principal component analysis of the 37 295 bioclimatic variables associated with the species occurrences. Specifically 296

we use the dudi.pca function from the R package ade4 1.7.13 (Dray and

Dufour 2007) and center and scale all bioclimatic variables to have a mean 298 of 0 and a unit standard deviation. We then determine the position of 299 species occurrences in the environmental space and estimate their bivariate 300 probability density function. We use a kernel method to estimate this density 301 and normalise it such that it ranges between zero and one. Specifically, to calculate the probability density function we use ecospat.grid.clim.dyn 303 from the R package ecospat 3.0 (Broennimann, Di Cola, and Guisan 2018) 304 with a grid resolution of 200. We then determine the location in the environmental space of the plant-pollinator communities using the function suprow 306 from ade4. The normalised density at that particular location corresponds 307 to our suitability metric, which we calculate using the R package raster 2.8.19 (Hijmans 2019). We use the kernel density method in the niche-factor 309 analysis (Broennimann et al. 2012) rather than the distance from the mode 310 (Hirzel et al. 2002), as it has been proposed earlier, as it has been shown to 311 reduce the procedure's sensitivity to sampling effort and the resolution of 312 the environmental space. 313

We used a sensitivity analysis to determine the minimum number of occurrences that are necessary to have robust environmental suitability values
in our communities. For that we used the species with most occurrences
available, *Archilochus colubris*, and calculated the mean absolute error of the
suitability values obtained with one thousand subsamples from the 74,791
occurrences available from GBIF.

Data analysis

We then used a set of bayesian multilevel models to evaluate the impact of environmental suitability on the size of the realised Eltonian trophic niche. We measure the size of the Eltonian niche using the normalised degree of species, this is, the number of species it interacts with given the number of species in the opposite guild (Martín González, Dalsgaard, and Olesen

2010). The normalised degree was modelled using a logit link function, and 326 a binomial distribution in which the number of species interacts with is the 327 number of successes, and the number of species in the opposite guild is the 328 number of trials. We are aware that wether species interacts or not is not a 329 Bernoulli process as species interactions are not strictly independent from each other. However, a binomial distribution allows us to account for the 331 differences in species richness across communities indirectly. Importantly, 332 however, results are qualitatively similar when we model species degree directly using a Poisson distribution and a logarithmic link function (Supp 334 info). 335

We evaluate four models to assess the relative importance of suitability. A first model, our baseline model, included five population-level predictors 337 and two grouping levels, species and the community. The population-level predictors in the baseline model, commonly called fixed effects, were the 330 habitat suitability, the species guild (plant or a pollinator), a metric of the 340 species generalism, and its number of known possible partners. We included the overall level of generalism as one would expect species with a large 342 number of interactions globally to be more likely to interact with a large 343 number of species locally. Concretely, we estimated an overall species level of generalism as the total number of interacting species across communities in our dataset. We also included the number of known possible partners as a predictor in our models as it allows us to control for the environmental 347 effects on species co-occurrence. We calculate this metric by determining the number of partners with which the species is known to interact in 349 any other community. Controlling for the number of potential partners 350 makes our model a particularly stringent test of our environmental stress hypotheses. Often the potential and the actual number of partners is the 352 same or very close to each other, especially for rare species present only in a 353 few communities.

As we were particularly interested in understanding whether the effect of habitat suitability is conditional on the species guild (plant or pollinator), we, 356 therefore, included guild and its interaction with suitability in the model. We 357 allowed the intercept of degree and slope of the suitability-degree relationship 358 to vary among species, this allowed us to investigate two questions. First, it allows us to inspect the extent to which suitability is a population or 360 a group level effect. Second, by investigating the correlation between the 361 intercept and the slope as a model parameter, it allowed us to inspect the extent by which species with a small or large number of interactions respond 363 to increasing levels of environmental stress. To account for unmeasured 364 differences between communities, like sampling effort, sampling method, or diversity, we calculated an intercept for each community in our study. To 366 facilitate model interpretation and convergence, we scaled all continuous 367 variables to have a mean of zero and a unit standard deviation. 368

We compared this baseline model with three alternative models in which we remove one predictor at a time. To quantify the difference between models, in terms of their expected out-of-sample performance, we use the 371 Wanatabe-Akaike information criterion (WAIC). All models were fitted under 372 a bayesian framework using the R package brms 2.8.0 (Bürkner 2017, 2018) 373 as an interface for Stan (Carpenter et al. 2017). For each model, we used 374 four Markov chains of 4,000 iterations each; half of the iterations were used 375 for warmup. We used weakly informative priors for all model parameters. 376 Specifically we used normal priors of mean zero and standard deviation ten for the population-level effects and the intercepts, a half-Cauchy prior with 378 a location of zero and a scale of two for the standard deviations, and, when 379 applicable, an LKJ-correlation prior with parameter $\zeta = 1$ for the correlation matrix between group-level parameters.

error of environmental suitability of communities for a species present in two plant–pollinator communities

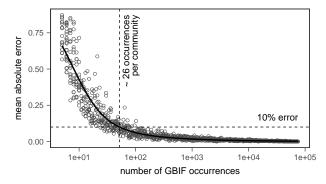


Figure 3: Sensitivity analysis of environmental suitability error. The number of independent occurrences retrieved from GBIF is inversely related to the error of environmental suitability for our plant-pollinator networks. The sensitivity analysis was performed by subsampling occurrences of *Archilochus colubris* the species in our dataset with the largest number of occurrences in GBIF, which was recorded in two of our communities.

Results

After performing our sensitivity analysis, we found that we need roughly 26, 18 independent occurrences for each community for which we calculated a 384 suitability value in order to obtain a mean absolute error below 0.1 (Fig. 3). 385 We therefore removed from further analyses species for which we did not 386 have enough occurrences to obtain robust estimates. When inspecting the 387 suitability values of the analysed species we found that some communitues 388 have larger overal suitability values. However, most communities included 389 species for which habitat suitability was low and species for which it was 390 high (Fig. 4). 391 We found that our models performed relatively well. The bayesian R-squared for our baseline model was 0.91, which indicates our models were able to 393 capture a large proportion of the variability on the data. Overall we found 394 only weak evidence of a relationship between the environmental niche size 395 and the normalised degree (Figure 5a). This relationship was stronger for 396 plants than for animals. Similarly environmental suitability does not show a consistent pattern across species. Indeed, when looking at the population

median suitability per community

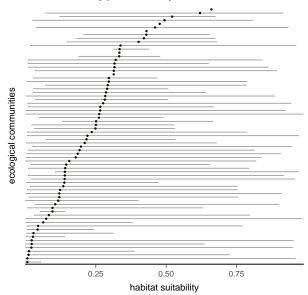


Figure 4: Median habitat suitability of communities in our dataset. Each row represents a different community and horizontal lines represent span the 2.5 and 97.5 quantiles.

level effects, suitability virtually no relationship with normalised degree, 399 neither for plants or pollinators (Figure 5b). However, suitability is still an 400 important predictor as the WAIC difference between our baseline model and 401 that that did not include suitability was 574 \pm 101. Both among plants and pollinators for some species there is a strong negative relationship between 403 suitability and normalised degree and some others with a strong positive 404 relationship (Figure 6a). As expected, we found a positive relationship 405 between a species generalism and normalised degree 5c). However our model 406 comparisons based on WAIC shows that including generalism in the model 407 adds relatively little information once we have accounted for the other 408 variables in the model. Finally, we found a strong and positive relationship between the number of possible interactions and the number of realised interactions in the community. The group level variation among communities was larger than that among 412 species. The standard deviation (in the parameters scale) of the community 413 intercepts was 1.02 [0.85, 1.23] while the standard deviation of the species

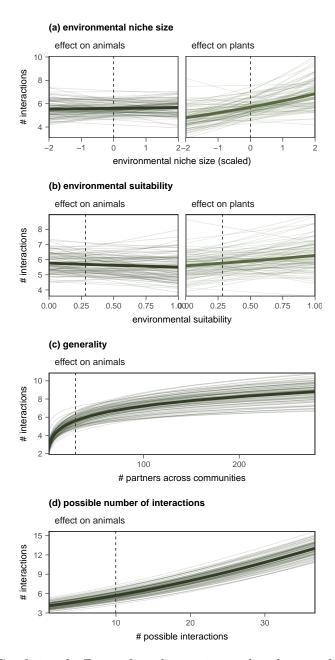


Figure 5: Conditional effects of predictors in our baseline model. The predicted values of the number of interacting species are based on a hypothetical community with 76 plants and 33 pollinators. These values correspond to the median number of species in each guild respectively. In each panel we condition on the mean values of all other predictors in the model, mean values for each predictor are indicated with a vertical dashed line. For model fitting, we scaled all predictors to have mean of zero and unit variance. However, with the exception of environmental niche size, here we show the unscaled predictors to facilitate interpretation. To illustrate the uncertainty around the fitted estimates we plot the fits of 100 independent draws from the posterior distribution. The thick lines indicate the mean values of the response distribution. As there was no interaction between guild and generality or the number of possible interactions, for these two predictors we only show conditional the conditional effect of pollinators.

Table 1: Comparison in out of sample predictive power of the baseline model (bold) and their alternatives. We rank models by their expected log predictive density based on their leave-one-out cross-validation information criterion (LOO). The standard error of the LOO difference provides rough guidance to the uncertainty of the model ranking. We also show the Wanatabe-Akaike information criterion (WAIC) of each model for comparison.

predictors	WAIC	SE
Suit. + Gen. + Pot. + Env. Suit. + Pot. + Env. Suit. + Gen. + Pot. Suit. + Gen. + Env. Gen. + Pot. + Env. unknown	6,167 6,157 6,170 7,737 6,740 6,163	135 138 135 281 183 136

intercept was 0.5 [0.44, 0.57], and that of the species' suitability slope was 0.34 [0.28, 0.4] (95% credible intervals shown within square brakets). Interestingly, the slope the relationship between suitability and normalised degree was negatively correlated with the species' intercept in the model (Figure 6b). The mean correlation coefficient was -0.51 [-0.69, -0.3].

420 Discussion

We found that environmental stress, which we measured using habitat suitability, was an important predictor for the number of partners a species interacts with. However, this relationship was highly variable among species. So much so that, at the population level, there is no discernible pattern between stress and the number of partners. Interestingly, within species, there was a significant correlation between the number of partners a species has and the stress-partners relationship itself. Specifically, species with a large number of partners in low-stress communities were more likely to have a negative relationship and hence reduce the number of partners as stress increases. Contrastingly, species in our datasets with a small number of partners in low-stress communities were more likely to have a larger number of partners in low-stress communities were more likely to have a larger number of partners in more stressful communities.

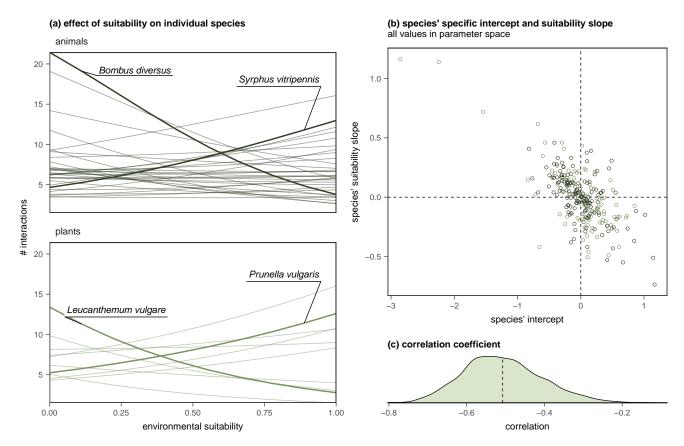


Figure 6: Species level effects of suitability. (a) Conditional effect of suitability for individual species. To facilitate visualization we show only species for which there is suitability information in at least six communities (10 plants and 33 pollinators). As in the previous figure, fitted values assume a hypothetical community of median size. In each panel we highlight two species for which the relationship between environental suitability and normalised degree was particularly strong. (b) The correlation between the species' intercept and the species' slope of suitability was negatively correlated.

Previous research supported the idea that the differences in network structure along an environmental gradient were due solely to changes to community 434 composition. In other words, it showed that the differences between two 435 hypothetical communities with the same species were stochastic for the most 436 part. Here, by abstracting with multiple environmental gradients into a single gradient of environmental stress, we show that this might not be 438 the case. Instead, the environment can also have a direct impact on the 439 number of partners a species interacts with and, therefore, influence network structure. Our results still show that the primary channel through which 441 environment affects interaction probabilities is determining which species 442 are present or absent from a particular community. However, they also show that, for a large number of species, it can have a substantial effect on their realised Eltonian component of the niche. Indeed, as evidenced by the 445 predictive power of competing models, habitat suitability had a substantial 446 contribution to predicting the number of possible partners in the community. Although we cannot tell with certainty how much of the effect of environmental stress can be attributed to either niche or neutral processes, the patterns 440 we observe are likely more related to the niche. On the one hand, it has been 450 shown that environmental models, such as our niche factor analysis, do a 451 poor job predicting species abundances across sites (Pearce and Ferrier 2001; 452 Sagarin, Gaines, and Gaylord 2006). It would be, therefore, somewhat naive 453 to assume that high habitat suitability translates directly to high species 454 abundances. On the other hand, even if species abundances are correlated with suitability, we show that within a community there is a wide range 456 of suitabilities, even for the limited number of species we included in our 457 analyses. Therefore, we might expect little or no overall effect of neutral processes on our response variable, as the abundance of some interaction 450 partners might increase and decrease for others. Furthermore, the fact that the effects of environmental stress do not depend on whether the species is

a plant or a pollinator (for which there can be niche differences based on individual behaviour) emphasises that co-evolutionary history, which in turn 463 shape niche processes driven by, might be driving the patterns we observe. 464 One of our most interesting results was the negative correlation between 465 the species' intercept and the slope of the habitat suitability in our model. 466 Specifically, species with a small intercept, this is, species that were likely to have an above-average number of partners in a community with little 468 environmental stress, were more likely to interact with a smaller number of 469 partners as stress increased. At least for these species, the hypothesis that 470 species focus on partners with which they are best adapted to interact as 471 environmental conditions deteriorate seems to hold weight. Contrastingly, 472 species with a below-average number of partners in a low-stress environment 473 were more likely to interact with a large number of partners as stress increased. At first glance, these species behave as facultative generalists when the 475 environment worsens. This is certainly true for this group of species in our dataset. However, there are some nuances to this observation once we take into account the biases inherent to community-level interaction data. 478 Specifically, we have no information about species for which environmental 470 conditions would, in principle, allow for a positive growth rate but are absent 480 from our datasets. A ramification of this bias is that an undetermined 481 number of species on this second group might also be more likely to focus 482 on partners with higher trait matching as conditions deteriorate. However, 483 we are not able to detect that trend because these species are not recorded in community composition or interaction data. Presumably, because their 485 preferred partners are unavailable and as such, they are absent from the 486 community or too rare to be observed. Note that the correlation between species' intercept and the slope is unlikely to be just a byproduct of our 488 modelling framework, as, in principle, there is nothing preventing species 489 with an above-average number of partners to interact with even more species

491 as environmental stress increases.

These observations have implications for understanding the assembly of 492 ecological networks and how it might be impacted by climate change, which currently threatens to increase the environmental stress of species within its 494 range. Previous research suggests that flexibility of interactions in ecological 495 communities might be a rather ubiquitous phenomenon whenever environmental conditions change in space or time (CaraDonna et al. 2017). This 497 flexibility has been shown to confer robustness and stability to the loss of 498 interaction partners (Kaiser-Bunbury et al. 2010). The current paradigm of 499 interaction turnover suggests first that trait matching is a relatively unim-500 portant driver of interaction turnover for generalist species (CaraDonna et 501 al. 2017). The fact that species with a high number of interactions were 502 more likely to interact with fewer species as environmental stress increased indicates that trait matching might still play a role in determining the in-504 teractions of generalist species. Second, current research suggests that the 505 rewiring of the interaction network is primarily driven by generalist species (Burkle, Marlin, and Knight 2013). We, however, demonstrate that at least 507 a fraction of the species that appear to be specialist in their communities 508 might be more flexible, if not more, than generalist species and, therefore, might have a more significant role in network persistence than previously 510 expected. Contrastingly, specialists that are highly constrained to interact 511 with partners of high trait-matching might be more vulnerable to increased 512 climate change-induced environmental stress and habitat degradation. Taken together, the fact that most species appear to be flexible interaction partners, 514 combined with the fact that those that are not flexible are less likely to per-515 sist, might explain why specialised communities that are highly constrained by trait matching (like some plant-hummingbird networks; Vizentin-Bugoni, Maruyama, and Sazima 2014; Maruyama et al. 2014) might be relatively 519 rare.

In conclusion, we show that...

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