Environmental stress affects niche breadth in plant-pollinator communities

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Abstract

What determines whether or not a species is a generalist or a specialist?

Evidence that the environment can influence species interactions is rapidly

accumulating. However, a systematic link between environment and the num-

ber of partners a species interacts with—the species degree—has been elusive

so far; presumably, because environmental gradients appear to have con-

trasting effects on species depending on the environmental variable. Here, we $+e_5+$ $+e_7 +e_8+$ $+e_8+$ $+e_8+$

environment, instead of environmental gradients directly, and degree using

a global dataset of plant-pollinator interactions. We found that the environ-25

ment can play a significant effect on species degree, even when accounting

for community composition, likely by an interaction interact with species

traits and evolutionary history Specifically, species are more likely to attain

intermediate levels of partners when environmental stress increases Species

that have a large number of interactions are more likely to focus on a

smaller number of, presumably higher-quality, interactions. Contrastingly,

the specialists that can cope with increased stress are more likely to broaden

their niche, presumably engaging in opportunistic interactions Many of the

species we analysed are not inherently generalist or specialist, many as some

specialists can effectively behave as facultative generalists. Species' degree of

generalisation should be measured on a relative scale depending on where

they are found and the environmental conditions at that location

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38 Introduction

stability.

Species interactions are known to vary widely across space and time. There
are multiple examples of species that interact with a large number of partners
in a particular community or season, but with fewer in another. Some of
this variation can be attributed to environmental drivers. However, how
exactly the environment, specifically the stress it imposes on species, affects
whether two species interact or not, and ultimately the number of partners a
species has is still unknown. Understanding how the number of partners—the
species degree—is driven by the environment is crucial because it underpins
the species role in its community and shapes the structure of the network
of interactions. This structure, in turn, determines ecosystem function and

where are the rets?

Species interactions are determined in part by niche processes (the matching of traits) and partly by neutral processes (more abundant species are more likely to encounter each other and, thus, interact). The environment can influence both of these processes. It is, therefore, not surprising that, despite limitations on the spatial extent or the number of environmental gradients considered, multiple studies have been able to show how changes to interactions can be related to environmental change (Tylianakis and Morris 2017). For instance, some studies suggest that the strength of some trophic interactions, like predation (McKinnon et al. 2010; Vucic-Pestic et al. 2011) and herbivory (Baskett and Schemske 2018), can increase with temperature but might decrease with precipitation (Pires et al. 2016). Some other studies, however, have shown either no effect (on average) or non-linear effects of temperature or precipitation on plant-pollinator interactions (Devoto, Medan, and Montaldo 2005; Gravel et al. 2018). Overall, while it looks clear that pairwise interactions respond to environmental drivers, there is high variability in the response (Tylianakis et al. 2008).

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One possible explanation for the seemingly contradictory evidence is that each species can have multiple partners. Each of these partners, as well as the interactions with them, can be simultaneously affected by the environmental conditions. Therefore environmental stress may affect the number of partners in different ways depending on its role in the community (for example its trophic guild) or even the species itself. Previous research suggests that there might be two alternative hypotheses of how environmental stress may affect species degree Tylianakis and Morris 2017). On the one hand, it is possible that when species are under environmental stress, they might be "pressured" to focus on partners with which they are best adapted to interact. For instance, Hoiss et al. (2012) found increased phylogenetic clustering between plants and pollinators at higher altitudes; while Peralta et al. (2015) found that parasitoids in plantation forest, where environmental stress was higher than in native forests, were constrained to interact with hosts they were best adapted to attack. Similarly, Lavandero and Tylianakis (2013) found

the Eltonian niche of parasitoids.

that environmental stress due to higher temperature reduced the breadth of

83 On the other hand, it is also possible that when species are under environ-

mental stress, they are forced to be more flexible in their interactions as

85 higher environmental stress is likely to be reflected in greater energetic or

86 reproductive costs. Therefore they might not be able to sustain encounter

87 rates with their preferred partners at sufficient levels. In line with this

hypothesis, Hoiss, Krauss, and Steffan-Dewenter (2015) found that the spe-

89 chalisation of plant-pollinator networks decreased both with elevation and

after extreme drought events. Likewise, Pellissier et al. (2010) found a

positive relationship between niche breadth and environmental stress: disk-

or bowl-shaped blossoms (which allow a large number of potential pollinator

species to access pollen and nectar rewards) dominated at high altitude

94 flower communities.

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Here, we investigate whether, and how environmental stress can systematically affect species degree. Our main aim is to test the two compet hypotheses that relate environmental stress and species degree and investigate whether this changes across species or between trophic guilds. We propose that specialist species can become "facultative" generalists to reduce their vulnerability to the absence of preferred partners (for example, when 100 variations in climate decouple phenologies; Benadi et al. 2014). We therefore 101 also expect that, as environmental stress increases species with a relatively 102 small number of partners are more likely to engage with more partners and 103 broaden their trophic niche. Species with a large number of partners, on 104 the other hand, should have a larger pool of available partners and might, therefore, be more likely to narrow their niche under environmental stress 106 by focusing on the most beneficial partners. Importantly, when testing 107 these hypotheses, we control for the potential effects of the environment in 108 community composition and the size of the species fundamental niche, both 109 from an Eltonian (interactions) and Grinellian (environment) perspective. 110

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We test these hypotheses using data on plant-pollinator interactions. We use the species' patterns of occurrence to estimate the environmental suitability in their communities as an indirect measure of the environmental stress they might experience. Condensing the environmental variation over multiple factors (like temperature and precipitation) into a single metric is crucial to generalise our findings at a global scale.

$_{^{17}}$ Methods

We retrieved plant-pollinator networks from the Web of Life database (Fortuna, Ortega, and Bascompte 2014). This database contains datasets originating from 57 studies published in the primary literature between 1923 and 2016. Calculating the environmental stress of species in their community

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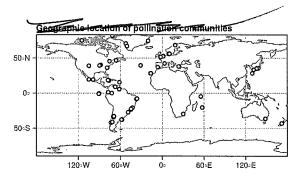
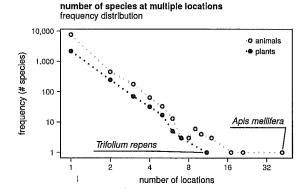


Figure 1: Worldwide distribution of pollination communities included in this study



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Figure 2: Frequency distribution of the number of locations in which a species is present. The most common pollinator species was Apis mellifera, which was sampled on 42 locations, while the most common plant species was Trifolium repens, which was sampled on 11 locations.

and their Eltonian niche breadth required us to reduce both the taxonomic 122 and distributional/locational uncertainty. A critical step towards reducing 123 this uncertainty is to ensure that the names used to identify species are valid 124 and unambiguous, which in turn allow us to obtain further information from 125 biological databases and accurately match species across studies. Therefore, 126 our first step was to ensure consistent spelling and standardisation of species names synonyms (See Supplementary Methods)/ This step ensured that 128 the matching of species across studies was as accurate as possible. 129 cleaning process resulted pn a total of 2,555 plants and 8,406 pollinator species distributed across 73 locations around the globe (Figure 1, 2). 131

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second, we calculated the suitability of the environment for a species in a

Why not just use suitability? is it necessary to make it a proxy?

particular community as a proxy of environmental stress. We assume that the

environmental stress a species experience in a particular location is inversely

related to the suitability of the average environmental conditions in that

place. Our third and final step was to relate the environmental suitability

to the relative number of partners a species has in a community, as a proxy

138 for Eltionian niche breadth. To explore this relationship within and across

species, we used a multilevel Bayesian model in which we controlled for the

potential effects of the environment on co-occurrence.

141 Species suitability

Our next step was to determine the habitat suitability of the species as 142 a proxy of the environmental stress they experience in their community. 143 As we aim to compare the trophic niche for different suitability levels, we only do this for species that were present in at least two communities. To 145 calculate the suitability of a species in a particular location, we used a 146 niche-factor analysis (Hirzel et al. 2002; Broennimann et al. 2012). This 147 approach is based on the probability density function of species distribution 148 in an environmental variable space. In a nutshell, habitats (characterised 149 by a collection of environmental variables) in which the species occurs most 150 often are deemed to be suitable for the species than habitats in which the 151 species has never been observed. This approach to estimating the habitat 152 suitability requires two critical pieces of information. First, we require 153 information about the occurrences of the species of interest. Second, we 154 require information about the environmental conditions for all the locations 155 in which the species occurs. 156

We retrieved 38.1 million occurrences from the Global Biodiversity Information Facility (GBIF; https://www.gbif.org). Issues with data quality are a central issue hampering the use of publicly available species occurrence GBIF data in ecology and biogeography (Jetz et al. 2019). We therefore

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🚮 followed a series of filters and geographic heuristics to correct or remove erroneous and imprecise referencing records (See supplementary methods; Zizka et al. 2019) which allowed us to identify and remove 7.5 million prob-163 lematic occurrences from further analysis. We integrated the occurrences 164 from our plant-pollinator communities to the cleaned occurrences retrieved 165 from GBIF. We retrieved environmental data from WorldClim V2.0, which includes 19 167 bioclimatic variables commonly used in species distribution modelling (Fick 168 and Hijmans 2017). We then complemented data obtained from WorldClim 169 with data from Environ (Title and Bemmels 2017), which includes 16 extra 170 bioclimatic and two topographic variables. The additional set of variables 171 from Envirem are relevant to ecological or physiological processes and 172 spen have the potential to improve our suitability estimation (Title and 173 Bemmels 2018). We obtained all environmental data as rasters composed 174 by cells of 2.5 arc-minutes. We chose this resolution because it provides 175 a reasonable match to the locational accuracy of the species occurrences found in GBIF, particularly those that come preserved specimens in museum 177 bad grammer collections. 178 After obtaining information about species occurrence and the environment, 179 we then merged these two datasets such that a vector with details of our 180 37 bioclimatic and topographic variables characterised the location of each 181 occurrence. Sets of occurrence data tend to be spatially aggregated due to 182 sample bias (tendency to collect close to cities, certain countries) 183 autocorrelation arises in ecological data because geographically clumped 184 records tend to be more similar in physical characteristics and/or species 185 abundances/than pairs of locations that are farther apart. To account 186 for such spatial dependency in occurrence data, if a species had more than 187 one occurrence percents within one of the cells of the bioclimatic raster, we 188 only included one MM occurrence records We did this to avoid giving

more weight to areas with a high number of occurrences, a common scenario 190 in occurrence records collected opportunistically as the ones we use here. In this step we removed 85.4% of the occurrences which resulted in a total of 192 4.5 million occurrences used in our niche analysis. 193 A common issue of terrestrial bioclimatic datasets is that the boundaries of the cells with information do not precisely match the landmass boundaries. 195 The result of this missmatch is that not all environmental variables was not 196 available for 3,273 of the raster cells with occurrences (0.8% of the total). 197 As expected, the vast majority of these problematic cells were close to the 198 shore. To address this issue, we calculated the average value of environmental 199 variables within an 5km buffer of the centre of the cell where the variable 200 was missing and used it to approximate the value of the variable in that cell. Using this procedure, we were able to fill environmental variables for 202 89.3% of the cells were missing. To fill the remaining 350 cells, we 203 repeated the aforementioned procedure but instead using a 10km buffer. We 204 removed from further analysis occurrences located within the 135 cells for 205 which we were unable to fill environmental variables. $(\ \ \ \ \)_o$ 206 Next, we calculate the probability density function of the species distribution in environmental space. To determine the environmental space, we use 208 the first two components from a principal component analysis of the 37 209 bioclimatic variables associated with the species occurrences. Specifically 210 we use the dudi.pca function from the R package ade4 1.7.13 (Dray and 211 Dufour 2007) and center and scale all bioclimatic variables to have a mean 212 of 0 and a unit standard deviation. We then determine the position of 213 species occurrences in the environmental space and estimate their bivariate 214 probability density function. We use a kernel method to estimate this density 215 and normalise it such that it ranges between zero and one. Specifically, to 216 calculate the probability density function we use ecospat.grid.clim.dyn from the R package ecospat 3.0 (Broennimann, Di Cola, and Guisan 2018)

with a grid resolution of 200. We then determine the location in the environ-219 mental space of the plant-pollinator communities using the function suprow 220 from ade4. The normalised density at that particular location corresponds 221 to our suitability metric, which we calculate using the R package raster 222 2.8.19 (Hijmans 2019). We use the kernel density method in the niche-factor 223 analysis (Broennimann et al. 2012) rather than the distance from the mode 224 (Hirzel et al. 2002), as it has been proposed earlier, as it has been shown to 225 reduce the procedure's sensitivity to sampling effort and the resolution of 226 the environmental space. There are so many scenningly tech steps here I think you should summ. We used a sensitivity analysis to determine the minimum number of occur-227 228 rences that are necessary to have robust environmental suitability values 229 in our communities. For that we used the species with most occurrences 230 available, Archilochus colubris, and calculated the mean absolute error of the suitability values obtained with one thousand subsamples from the 74,791 232 occurrences available from GBIF. 233

234 Data analysis

We then used a set of bayesian multilevel models to evaluate the impact of 235 environmental suitability on the number of partners a species has. Specifically, 236 we use the normalised degree of species as our response variable; this is, the 237 number of species it interacts with green the number of species in the opposite guild (Martín González, Dalsgaard, and Olesen 2010). The normalised degree 239 was modelled using a logit link function, and a binomial distribution in which the number of species interacts with is the number of successes, and the 240 number of species in the opposite guild is the number of trials. We are 242 aware that whether species interact or not is not a Bernoulli process as species interactions are not strictly independent from each other. However, 244 a binomial distribution allows us to account for the differences in species richness across communities indirectly. Importantly, however, results are qualitatively similar when we model species degree directly using a Poisson

We evaluate four models to assess the relative importance of suitability. A first model, our baseline model, included three population-level predictors 250 and two grouping levels, species and the community. The population-level 251 predictors in the baseline model, commonly called fixed effects, were the 252 habitat suitability, the species guild (plant or a pollinator), and its number of known possible partners. We included the number of known possible partners 254 as a predictor in our models as it allows us to control for the environmental 255 effects on species co-occurrence. We calculate this metric by determining the number of partners with which the species is known to interact in any other community. Controlling for the number of potential partners makes our 258 model a particularly stringent test of our environmental stress hypotheses 259 because this variable could explain a large proportion of variance. Often the potential and the actual number of partners is the same or very close to each 261 other, especially for rare species present only in a few communities. As we were interested in understanding whether the effect of habitat suitability is

distribution and a logarithmic link function.

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conditional on the species guild (plant or pollinator), we Markforth included both guild and its interaction with suitability in the model. Is this the only introduce?

We allowed the intercept of degree and slope of the suitability-degree relationship to vary among species. This approach allowed us to investigate two questions. First, it allows us to inspect the extent to which suitability is a population or a group evel effect. Second, by investigating the correlation between the intercept and the slope as a model parameter, it allowed us to inspect the extent by which species with a small or large number of interactions respond to increasing levels of environmental stress. To account for unmeasured differences between communities, like sampling effort, sampling method, or diversity, we also calculated an intercept for each community in our study. To fadilitate model interpretation and convergence, we scaled all

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continuous variables to have a mean of zero and a unit standard deviation.

We compared this baseline model with three alternative models in which 277 we remove one predictor at a time. To quantify the difference between 278 models, in terms of their expected out-of-sample performance, we use the 279 Wanatabe-Akaike information criterion (WAIC). All models were fitted under 280 a Sayesian framework using the R package brms 2.8.0 (Bürkner 2017, 2018) 281 as an interface for Stan (Carpenter et al. 2017). For each model, we used 282 four Markov chains of 4,000 iterations each; we used half of the iterations 283 for warmup. We used weakly informative priors for all model parameters. 284 Specifically we used normal priors of mean zero and standard deviation ten 285 for the population-level effects and the intercepts, a half-Cauchy prior with a location of zero and a scale of two for the standard deviations, and, when applicable, an LKJ-correlation prior with parameter $\zeta=1$ for the correlation 288 matrix between group-level parameters.

Results

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After performing our sensitivity analysis, we found that we need roughly 26,
18 independent occurrences for each community for which we calculated a
suitability value in order to obtain a mean absolute error below 0.1 (Fig. S1).
We therefore removed from further analyses species for which we did not
have enough occurrences to obtain robust estimates. When inspecting the
suitability values of the analysed species, we found that most communities
included species for which habitat suitability was low and species for which

Our models performed relatively well. The Payesian R-squared for our baseline model was 0.91, which indicates our models were able to capture a large proportion of the variability on the data. Overall, we found that environmental suitability does not show a consistent pattern across species.

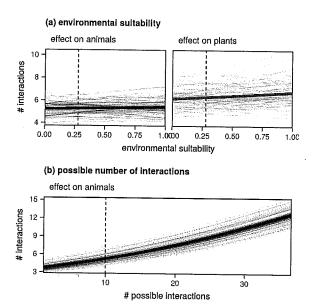


Figure 3: Conditional effects of predictors in our baseline model. The predicted values of the number of interacting species are based on a hypothetical community with 76 plants and 33 pollinators. These values correspond to the median number of species in each guild, respectively. In each panel, we condition on the mean values of all other predictors in the model. We indicate mean values for each predictor with a vertical dashed line. For model fitting, we scaled all predictors to have a mean of zero and unit variance. However, except for environmental niche size, here we show the unscaled predictors to facilitate interpretation. To illustrate the uncertainty around the fitted estimates, we plot the fits of 100 independent draws from the posterior distribution. The thick lines indicate the mean values of the response distribution. As there was no interaction between guild and generality or the number of possible interactions for these two predictors, we only show conditional the conditional effect of pollinators.

Indeed, when looking at the population level effects, suitability has virtually no relationship with the normalised degree white for plants or pollinators (Figure 8b) However, suitability is an important predictor as the WAIC difference between our baseline model and that that did not include suitability was 581 ± 104 (Table 1). This apparent discrepancy can be explained by the variability of the suitability-degree relationship across species.

suitability and normalised degree for some others with a strong positive relationship (Figure 4a). Interestingly, the slope of this relationship correlates

Table 1: Comparison in out of sample predictive power of the baseline model (bold) and their alternatives. We rank models by their expected log predictive density based on their Wanatabe-Akaike information criterion (WAIC).

predictors	WAIC	SE
Suit. * Guild + Pot.	6,149	140
Suit. * Guild	7,652	273
Guild + Pot.	6,730	190
Suit. + Pot.	6,151	139

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negatively with the species' intercept in the model (Figure 4b). The mean correlation coefficient was -0.43 [-0.62, -0.22]. In other words, the slope of the suitability-degree relationship was more likely to be positive for species with a smaller number of partners in lower-suitability and negative for species with a larger number of partners in lower-suitability conditions.

As expected, we found a strong and positive relationship between the number of possible interactions and the number of realised interactions in the community. There was a large difference on WAIC between the model that included this predictor and that that excluded it. This result indicates that the availability of potential partners accounts for a large proportion of the variability on species degree.

variability on species degree.

Importantly, our findings of the suitability-degree relationship were qualitatively unchanged whether we included this variable or not.

The group-level variation among communities was larger than that among species which further indicates the importance of the local context when determining species degree. Specifically, the standard deviation (in the parameters scale) of the community intercepts was 1.03 [0.86, 1.23] while the standard deviation of the species intercept was 0.54 [0.48, 0.61], and that of the species' suitability slope was 0.35 [0.29, 0.41] (95% credible intervals shown within square brakets).

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are there species? (a) effect of suitability on individual species (b) species' specific intercept and suitability slope all values in parameter space Bombus diversus Syrphus vitripennis

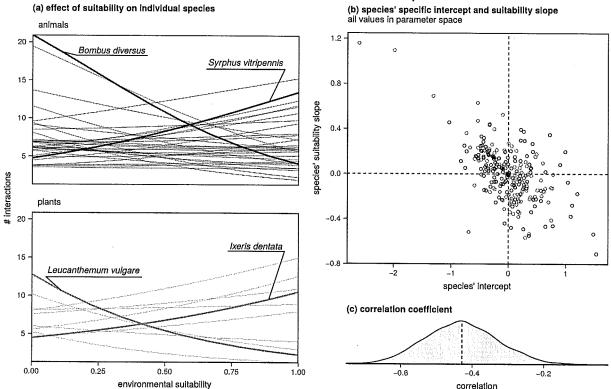


Figure 4: Species-level effects of suitability. (a) Conditional effect of suitabilk ity for individual species. To facilitate visualization, we show only species for which there is suitability information in at least six communities (10 plants and 33 pollinators). As in the previous figure, fitted values assume a hypothetical community of median size. In each panel, we highlight two species for which the relationship between environmental suitability and the normalised degree was particularly strong. (b) The correlation between the species' intercept and the species' slope of suitability was negative favymanesi;

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332 Discussion

We set out to explore whether and how environmental stress can systematically affect species degree. After accounting for the pool of potential 334 partners, we found that wow environmental conditions contribute to deter-335 mining whether a species is a generalist or a specialist in their community. 336 However, we also found that the particular effect of the environment is strongly dependent on the species. We/proposed two alternative hypotheses 338 of how environmental stress may affect the degree, and we found evidence 339 for both of them. Species with a large number of partners in low-stress 340 communities were more likely to have a negative relationship and hence 341 reduce the number of partners as stress increases. Contrastingly, species in 342 our datasets with a small number of partners in low-stress communities were 343 more likely to have a larger number of partners in more stressful communities. In summary, environmental stress pushes species that are flexible enough 345 to change their interaction partners towards intermediate levels of degree, a 346 so-called "regression towards the mean".

Our results suggest that changes in community composition are indeed 348 the primary channel through which the environment determines changes 349 interaction probability. However, they also show that, for a large number 350 of species, the environment may also play a substantial role in determining 351 their realised (Eltonian) niche. While previous research has recognised that 352 environmental factors may help explain the changes in network structure 353 along environmental gradients that cannot be explained by community composition (Tylianakis, Tscharntke, and Lewis 2007), how these two factors 355 were linked had been elusive so far (Gravel et al. 2018). We believe that part 356 of this difficulty could have arisen because species, and ultimately network 357 structure, can respond in multiple and contrasting, ways depending on the 358 particular bioclimatic variable examined (e.g. temperature or precipitation). 359

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Using stress to summarise the effect on species of multiple environmental

gradients allowed us to detect a clear signal of the environment in species' interaction patterns.

Although both niche and neutral processes are relevant $\overset{\cite{N}}{\cite{M}}$ determining 363 species interactions, our model suggests that niche processes may be the predominant mechanism through which the environment systematically affect 365 species degree. First, it is unlikely that environmental suitability correlates 366 to local species abundances (Pearce and Ferrier 2001; Sagarin, Gaines, and 367 Gaylord 2006). Second, even if there is a relationship between suitability and abundances, a particular environmental gradient could have a positive effect 369 on the abundance of some species and a negative effect on others. Indeed, 370 that within a community there is a wide range of suitability values, 371 even for the relatively limited number of species we were able to include in 372 our analysis. 373

really "show"

Recent research suggests that species are continuously changing their interaction partners wherever environmental conditions change in space or time 375 (Raimundo, Guimarães, and Evans 2018). So far it appears that this rewiring 376 is primarily driven by generalist species (Ponisio, Gaiarsa, and Kremen 2017; 377 Burkle, Marlin, and Knight 2013), presumably because generalist species 378 are less sensitive to trait matching of their interaction partners (CaraDonna 379 et al. 2017). Our results add two important nuances to these findings. 380 First, because generalists seem to focus on a smaller number of partners as 381 environmental conditions deteriorate, we show that trait matching might still 382 play a role in determining the interactions of generalist species. Second, we 383 demonstrate that rewiring is not exclusive to generalists. At least a fraction of the species that appear to be specialist in their communities might be as 385 flexible, if not more, than generalist species, effectively behaving as faculta-386 tive generalists in the face of environmental change. These flexible specialists 387 might therefore have a more significant role in network persistence than previously expected.

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While our model detected how the environment affects the interactions of two 390 types of species, generalists and flexible specialists, there is a third group that 391 remained invisible for our model but has important implications for network 392 persistence and stability. Species that are able to vary their interaction 393 partners flexibly and their role in the network are more likely to persist in 394 their community as environmental conditions vary (Gaiarsa, Kremen, and 305 Ponisio 2019). We propose this third group of is composed of specialists that are constrained to interact with partners of high trait-matching and 397 therefore were not likely to be found in more than one community. If 398 species that are not flexible are unlikely to persist over temporal or spatial 399 environmental gradients, we can expect specialised communities that are highly constrained by trait-matching (like some plant-hummingbird networks; 401 Vizentin-Bugoni, Maruyama, and Sazima 2014; Maruyama et al. 2014) to 402 be far more vulnerable to increased climate change-induced environmental 403 stress and habitat degradation than communities where role and interaction 404 flexibility are more prevalent. 405

Similarly, if the patterns we see in our models have also played a role during the evolutionary history of pollination communities, our results also help 407 explain why only a small fraction of plant-pollinator interactions show a 408 strong signature of deep co-evolutionary history (Hutchinson, Cagua, and 409 Stouffer 2017). The increases on the stress that species are predicted to 410 experience due to rapid environmental change might further erode the co-411 evolutionary history of specialist species. Communities as a whole might be 412 in a trajectory of even more diffuse co-evolution. For specialist species, at least, the longer-term benefits of being able to interact with multiple partners 414 might be more important than the shorter-term benefits of interacting with 415 partners of high trait matching. 416

The structural implications of the "regression towards the mean" that environmental stress promotes are less clear. However, it is plausible to expect that nestedness, and therefore network stability, might be reduced in the
face of rapid environmental change. Determining exactly how the changes
in degree caused by environmental stress reflect on systematic changes in
network structure would be an interesting avenue of research. Answering
this question would require expanding our suitability analysis to all species
in the community and compare the degree distribution of networks along a
gradient of stress for the community as a whole.

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In conclusion, we show that the environment can affect the realised niche of plants and pollinators in systematic ways beyond community composition.
On the one hand, species that are inflexible with their interaction partners are unlikely to persist under more stressful environmental conditions. On the other, species that are flexible with their interactions partners experience a regression towards intermediate levels of degree, where generalist species tend to interact with fewer, presumably more effective, partners and specialist tend to interact with more partners.

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Environmental stress affects niche breadth in plant-pollinator communities

Supplementary information

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Supplementary methods

Reducing taxonomic uncertainty

Data were obtained from the Web of Life database (Fortuna, Ortega, and Bascompte 2014) which includes data from 57 published studies (Abreu and Vieira 2004; Arroyo, Primack, and Armesto 1982; Barrett and Helenurm 1987; Bartomeus, Vilà, and Santamaría 2008; Bek 2006; Bezerra, Machado, and Mello 2009; Bundgaard 2003; Canela 2006; Clements and Long 1923; del Coro Arizmendi and Ornelas 1990; Dicks, Corbet, and Pywell 2002; Dupont and Olesen 2009; Dupont, Hansen, and Olesen 2003; Elberling and Olesen 1999; Gutierrez, Rojas-Nossa, and Stiles 2004; Hattersley-Smith 1985; Herrera 1988; Hocking 1968; Ingversen 2006; INouE et al. 1990; Inouye and Pyke 1988; Kaiser-Bunbury et al. 2014, 2010; Kakutani et al. 1990; Kato 2000; Kato, Matsumoto, and Kato 1993; Kato and Miura 1996; Kato et al. 1990; Kevan 1970; Kohler 2011; Lara 2006; Las-Casas, Azevedo Júnior, and Dias Filho 2012; Lundgren and Olesen 2005; McMullen 1993; Medan et al. 2002; Memmott 1999; Montero 2005; Mosquin 1967; Motten 1986; Olesen, Eskildsen, and Venkatasamy 2002; Ollerton 2003; Percival 1974; Petanidou and Vokou 1993; Philipp et al. 2006; Primack 1983; Ramirez 1989; Ramirez and Brito 1992; Robertson 1929; Rosero and others 2003; Sabatino 2010; Schemske et al. 1978; Small 1976; Smith-Ramírez et al. 2005; Stald, Valido, and Olesen 2003; Vázquez 2002; Vizentin-Bugoni et al. 2016; Yamazaki and Kato 2003).

Interaction data from the included studies included 11,231 unique organism names. From these 1,166 were present in more than one study. From the total number of organisms, 159 were identified to the subspecies or variety level, 6,759 to the species level, 1,755 to the genus level, whereas the remaining 2,558 were unidentified. As the species level was the most common taxonomic rank available in our interaction datasets, in all further analysis, we grouped together subspecies or varieties within the same species.

We were able to confirm the validity of 5,263 of the scientific names used to identify organisms (roughly 76%). We assessed the validity of a name by querying the Global Names Resolver database (https://resolver.globalnames. org) which includes data from 98 taxonomic sources. We accessed this database using the function gnr_resolve from the R package taxize 0.9.6 (Chamberlain and Szocs 2013; S. Chamberlain, Szoecs, et al. 2019).

From the remaining 1,655 names we were unable to validate, we were able to identify and correct 726 that contained spelling mistakes. These spelling mistakes were corrected automatically by fuzzy matching the canonical names in our data sources with those in the Global Names Resolver database. However, on rare occasions, the fuzzy matching algorithm can suggest a scientific name that has a similar spelling, but that corresponds to an organism in a different taxonomic group, often a separate kingdom. To address this potential problem, we checked the taxonomic hierarchy of suggested names and confirmed that it matched our expected taxon. We retrieved all taxonomic hierarchies from the National Center for Biotechnology Information taxonomic database (https://www.ncbi.nlm.nih.gov/taxonomy).

As species names are constantly changing, we subsequently checked for possible synonyms of the canonical names in our data sources. Using data from the Integrated Taxonomic Information System database (http://www.itis.gov), we found synonyms and alternative names for 611 species.

Finding these alternative names was required for two main reasons. First, because we wanted to be able to identify the cases in which the same species might have been recorded with different names in various data sources. This can occur not only when the canonical name has been changed but also when there are widely used orthographic variants. Second, because retrieving occurrence data is often only possible using the latest accepted/valid name for a particular species.

All together, from the 1,655 names we were unable to validate, it was not possible to automatically correct or find synonyms 332 of them. We then manually consulted multiple online databases, chiefly Wikispecies (https://species.wikimedia.org/), and looked for canonical names that both, resembled the unvalidated names and matched the geographic and taxonomic expectations. In this fashion, we were able to further correct 25 names. Most manual corrections were made on names that have been abbreviated or had more than two spelling mistakes. A complete list of manual name corrections can bee seen in Table S1.

This cleaning process allowed us to match further 270 names across data

Table S1: Manually corrected canonnical names. More than one correct name have been included when an accepted/valid synonym the cannonical name exists.

incorrect name	corrected name	guild
Acaena pinn	$A caena\ pinnatifida$	plant
Adesmia brachy	$A desmia\ brachy semeon$	plant
Aesculus camea	$Aesculus \ X \ carnea$	plant
Brachyome sinclairii	$Brachyscome\ sinclairii$	plant
Calceolaria arac	Calceolaria arachnoidea	plant
Equium sabulicola	$Echium\ sabulicola$	plant
Euonymus fo rtunei	Euonymus fortunei	plant
Galvezia leucantha pubescen	Galvezia leucantha	plant
Heliconia simulans	$Heliconia\ angusta$	plant
Pitcaimia flammea	Pitcairnia flammea	plant
Psittacanthus flavo viridis	$Psitta can thus \ flavo-viridis$	plant
Rodophiala bifidum	$Rhodophiala\ bifida$	plant
Stachys albi	$Stachys\ albicaulis$	plant
Stenactis annuus	$Erigeron\ annuus$	plant
Thaspium aureum atropurpurem	$Thas pium\ trifoliatum$	plant
Tristhema mauritiana	Tristemma mauritianum	plant
Tropaeolum polyph	Tropaeolum polyphyllum	plant
Tyttnera scabra	$Turnera\ scabra$	plant
	$Turnera\ ulmifolia$	plant
VVedelia biflora	Melanthera biflora	plant
	Wedelia biflora	plant
Cateres pennatus	Kateretes pennatus	pollinator
Eclimus harrisi	Condylostylus crinicauda	pollinator
Ptilandrena g. maculati	$Andrena\ distans$	pollinator
Tapinotaspis caerulea	Chalepogenus caeruleus	pollinator
Tapinotaspis herbsti	$Chale pogenus\ herbsti$	pollinator

sources and, by doing so, identify another 72 species that were present in more than one study. The process also allowed us to identify problematic data sources in which some names were included as both plants and pollinators. These data sources were removed from further analysis. In seven of our data sources interaction data was recorded at multiple points in time. When this was the case, we combined interaction data into one single interaction network.

Reducing location uncertainty

We retrieved occurrences from the Global Biodiversity Information Facility (GBIF; https://www.gbif.org) using the R package rgbif 0.9.6 (Chamberlain and Boettiger 2017; S. Chamberlain, Barve, et al. 2019). Specifically, for each species, we only requested occurrences for which the coordinates of the observation were available and that had no known geospatial issue in the GBIF database. Roughly, we downloaded 38.1 million occurrences for the 986 species we were interested on. This occurrences, however, contain observations of mixed quality. Therefore, we followed Zizka et al. (2019) and

error of environmental suitability of communities for a species present in two plant-poliinator communities

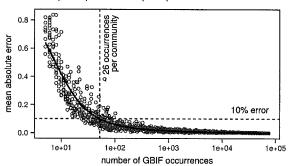


Figure S1: Sensitivity analysis of environmental suitability error. The number of independent occurrences retrieved from GBIF is inversely related to the error of environmental suitability for our plant-pollinator networks. The sensitivity analysis was performed by subsampling occurrences of *Archilochus colubris* the species in our dataset with the largest number of occurrences in GBIF, which was recorded in two of our communities.

applied a series of simple filters and geographic heuristics to remove those of lower quality. Specifically, we removed all occurrences with (i) a coordinate uncertainty larger than 100km; (ii) those recorded prior to 1945 (as records prior to this date have been shown to be often imprecise); (iii) those in which the number of counts in the occurrence was registered was either zero (as that indicates that the species has not been recorded); and (iv) those occurrences in which the "basis of record" was not a human observation or a preserved specimen (as occurrences from unknown and fossil records are known to be highly unreliable). We then used the R package CoordinateCleaner 0.9.6 (Zizka et al. 2019) and land mass and country data from Natural Earth (https://www.naturalearthdata.com) with a 1:10,000,000 scale to further identify and remove problematic occurrences. We removed occurrences for which their coordinates (v) fell outside the borders of the country where they were recorded; (vi) those around a country capital or the centroid of the country and province centroids; (vii) those around a biodiversity institution; and (viii) those located within oceans. Thorugh this cleaning process, we removed with 7.5 million occurrences distributed across 916 species.

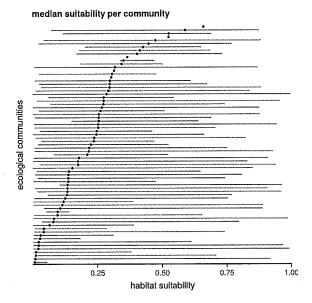


Figure S2: Median habitat suitability of communities in our dataset. Each row represents a different community and horizontal lines represent span the 2.5 and 97.5 quantiles.

Supplementary results

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