## Environmental stress drives intermediate plant and pollinator specialisation

E. Fernando Cagua¹ (efc29@uclive.ac.nz)
Audrey Lustig² (audrey.lustig@canterbury.ac.nz)
Jason M. Tylianakis¹ (jason.tylianakis@canterbury.ac.nz)
Daniel B. Stouffer¹ (daniel.stouffer@canterbury.ac.nz)

- <sup>1</sup> Centre for Integrative Ecology, School of Biological Sciences, University of Canterbury, Private
- 2 Bag 4800, Christchurch 8041, New Zealand
- <sup>2</sup> Geospatial Research Institute, University of Canterbury, Private Bag 4800, Christchurch 8041,
- 4 New Zealand
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### 17 Abstract

What determines whether or not a species is a generalist or a specialist? Evidence that the environment can influence species interactions is rapidly accumulating. However, a systematic link between environment and the number of partners a species interacts with has been elusive 20 so far; presumably, because environmental gradients appear to have contrasting effects on species 21 depending on the specific environmental variable. Here, we test whether the stresses imposed 22 by the environment, instead of environmental gradients directly, influence species specialisation 23 using a global dataset of plant-pollinator interactions. We found that environmental stress can significantly affect specialisation, even when accounting for changes in community composition, likely by interacting with species' traits and evolutionary history. Species that have a large number of interactions tended to focus on fewer, presumably higher-quality, interactions under stressful environmental conditions. Contrastingly, the specialists present in multiple locations tended to broaden their niche, presumably engaging in opportunistic interactions to cope with increased environmental stress. Indeed, many apparent specialists effectively behaved as facultative generalists. Overall, many of the species we analysed are not inherently generalist or specialist. Instead, the species' level of specialisation should be considered on a relative scale depending on the environmental conditions at a given location.

### 34 Introduction

Species interactions are known to vary widely across space and time (Laliberté & Tylianakis 2012; Schleuning et al. 2012; Poisot et al. 2015; Trøjelsgaard et al. 2015). There are multiple examples of species that interact with a large number of partners in a particular community or season, but 37 with fewer in another (Olesen et al. 2008, 2011; Dupont et al. 2009; Benadi et al. 2014; Rabeling et al. 2019). Some of this variation can be attributed to environmental drivers (Tylianakis & Morris 2017). However, it remains unknown how generalisable characteristics of the environment, specifically the stress it imposes on species, affects whether two species interact, and, ultimately, the species' specialisation. Understanding how the environment drives the number of partners is crucial because it underpins the species' role in its community and shapes the structure of the network of interactions (Cirtwill et al. 2018). This structure, in turn, determines ecosystem function and stability (Thebault & Fontaine 2010). Species interactions are determined in part by niche processes (the matching of traits) and partly by neutral processes (more abundant species are more likely to encounter each other and, thus, interact) (Vázquez et al. 2009a, b). The environment can influence both of these processes (Godsoe et al. 2017). It is, therefore, not surprising that, despite limitations on the spatial extent or the number of environmental gradients considered, multiple studies have shown how changes to interactions can be related to environmental change (Tylianakis & Morris 2017). Despite widespread evidence that species interactions and their networks can change along environmental gradients, the direction of change can be variable (Devoto et al. 2005; Pires et al. 2016; Baskett & Schemske 2018), such that until now it has been impossible to extrapolate previous findings to new kinds of gradients. Overall, while it looks clear that pairwise interactions respond to environmental drivers, there is high variability in the response (Tylianakis et al. 2008). One possible explanation for the seemingly contradictory evidence is that different bioclimatic factors (like temperature or precipitation) can have contrasting effects on species and their partners, such that we lack a common currency with which to derive general responses to the environment. Here we attempt to overcome these barriers to deriving general rules by reducing multiple factors into a single measure of environmental stress. Previous research suggests that environmental stress may

affect the number of interaction partners a species has in different ways depending on the species' role in the community (for example its trophic guild) (Tylianakis & Morris 2017). Specifically, we propose two alternative hypotheses of how environmental stress may affect specialisation. First, species under environmental stress might be "pressured" to focus on partners with which they are best adapted to interact. For instance, Hoiss et al. (2012) found increased phylogenetic clustering between plants and pollinators at higher altitudes; while Peralta et al. (2015) found that parasitoids in plantation forest, where environmental stress was higher than in native forests, constrained their host use to those with which phylogenetic matching was strongest. Similarly, Lavandero & Tylianakis (2013) showed that environmental stress due to higher temperature reduced the trophic niche breadth (range of host genotypes used) of parasitoids, suggesting higher specialisation.

Alternatively, it is also possible that when species are under environmental stress, they are forced to be more flexible in their interactions. Higher environmental stress is likely to be reflected in greater energetic or reproductive costs. Therefore, they might not be able to sustain encounter rates with their preferred partners at sufficient levels. In line with this hypothesis, Hoiss *et al.* (2015) found that the specialisation of plant-pollinator networks decreased both with elevation and after extreme drought events. Likewise, Pellissier *et al.* (2010) found a positive relationship between niche breadth and environmental stress: disk- or bowl-shaped blossoms (which allow a large number of potential pollinator species to access pollen and nectar rewards) dominated at high altitude flower communities.

Here, we investigate whether and how environmental stress can systematically affect specialisation.
Our main aim is to test the two hypotheses mentioned above that relate environmental stress with
a species' number of partners and investigate whether this changes systematically across species
or between trophic guilds. We propose that specialist species can become "facultative" generalists
to reduce their vulnerability to the absence of preferred partners (for example, when variations in
climate decouple phenologies; Benadi et al. 2014). In other words, we expect that, as environmental
stress increases, specialists should be more likely to engage with more partners. Species with many
partners, conversely, should have a larger pool of available partners and might, therefore, be more
likely to specialise on the most beneficial partners under environmental stress. Importantly, when
testing these hypotheses, we control for the potential effects of the environment on community

We test these hypotheses using data on plant-pollinator interactions. These interactions provide a particularly interesting system to test these hypotheses because, due to the multiple trade-offs involved in the pollination service, there are multiple intuitive ways in which we could imagine

composition (which has been previously shown to be a determinant factor; Gravel et al. 2018).

species responding to environmental stress given the available partners. We estimate the stress

species might experience a given community by calculating its bioclimatic suitability based on the

97 species' patterns of global occurrence.

### $_{ ext{ iny 98}}$ $\operatorname{Methods}$

We retrieved plant-pollinator networks from the Web of Life database (Fortuna et al. 2014). This database contains datasets originating from 58 studies published in the primary literature between 100 1923 and 2016. Calculating the environmental stress of species in their community and their potential 101 partners required us to reduce both the taxonomic and distributional/locational uncertainty. A 102 critical step towards reducing this uncertainty is to ensure that the names used to identify species 103 are valid and unambiguous, which in turn allows us to obtain further information from biological 104 databases and accurately match species across studies. Therefore, our first step was to ensure 105 consistent spelling and standardisation of species names and synonyms (see Supplementary Methods). 106 The cleaning process resulted on a total of 2,555 plants and 8,406 pollinator species distributed 107 across 73 locations arround the globe (Figure 1 and S1). 108

After matching species across studies as accurately as possible, we performed two more steps. First,
we calculated the environmental stress of species in their communities. Second, we relate each
species' stress in its community with the number of partner species it has as a metric of its level of
specialisation.

### Environmental stressdispla

We calculated the environmental stress of species in their communities. We assume that the stress a species experiences in a particular location is inversely related to the suitability of the average

#### number of species at multiple locations frequency distribution

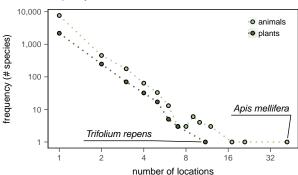


Figure 1: Frequency distribution of the number of locations in which a species is present. The most common pollinator species was Apis mellifera, which was sampled at 42 locations, while the most common plant species was Trifolium repens, which was sampled at 11 locations.

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environmental conditions in that place. As we aim to compare specialisation levels for different levels of environmental stress, we only calculate bioclimatic suitability for species that were present in at 117 least two communities. To calculate the bioclimatic suitability of a species in a particular location, 118 we used a niche-factor analysis (Hirzel et al. 2002; Broennimann et al. 2012). This approach is 119 based on the probability density function of species distribution in an environmental variable space. 120 Habitats are characterised by a collection of environmental variables. In a nutshell, those habitats in which the species occurs more often are deemed to be more suitable for the species than habitats 122 in which the species has never been observed. As bioclimatic suitability is calculated in a scale from 123 zero to one following the niche-factor analysis, for simplicity, we define environmental stress as one 124 minus suitability. 125

The niche factor analysis requires two critical pieces of information. First, it requires information about the occurrences of the species of interest. Second, the method requires information about the environmental conditions for all the locations in which the species occurs. We retrieved 38.1 million occurrences of our focal species from the Global Biodiversity Information Facility (GBIF.org 2019a, b, c, d, e, f, g, h, i, j). Issues with data quality are a central issue hampering the use of publicly available species occurrence GBIF data in ecology and biogeography (Jetz et al. 2019). We, therefore, followed a series of filters and geographic heuristics to correct or remove erroneous and imprecise referencing records (see supplementary methods; Zizka et al. 2019) which allowed us to identify and remove 7.5 million potentially problematic occurrences from further analysis. We

integrated the occurrence records from our plant-pollinator communities to the cleaned occurrences retrieved from GBIF.

We retrieved environmental data from WorldClim V2.0, which includes 19 bioclimatic variables 137 commonly used in species distribution modelling (Fick & Hijmans 2017). We then complemented 138 data obtained from WorldClim with data from Environ (Title & Bemmels 2017), which includes 139 16 extra bioclimatic and two topographic variables. The additional set of variables from Environ 140 are relevant to ecological or physiological processes and thus have the potential to improve our suitability estimation (Title & Bemmels 2018). We obtained all environmental data as rasters 142 composed by cells of 2.5 arc-minutes. We chose this resolution because it provides a reasonable 143 match to the locational accuracy of the species occurrences found in GBIF, particularly those that originate from preserved specimens in museum collections. 145

After obtaining information about species occurrence and the environment, we then merged these two datasets such that a vector with details of our 37 bioclimatic and topographic variables characterised 147 the location of each occurrence. Sets of occurrence data tend to be spatially aggregated due to 148 sample bias (tendency to collect close to cities, certain countries). Moreover, spatial autocorrelation arises in ecological data because geographically clumped records tend to be more similar in physical 150 characteristics and/or species abundances than do pairs of locations that are farther apart. To 151 account for such spatial dependency in occurrence data, we only included one occurrence record if a species had more than one within a cell of the bioclimatic raster. We did this to avoid giving 153 more weight to areas with a high number of occurrences, a common scenario in occurrence records 154 collected opportunistically as the ones we use here. In this step we removed 85.4% of the occurrences 155 which resulted in a total of 4.5 million occurrences used in our niche analysis. 156

A common issue of terrestrial bioclimatic datasets is that the boundaries of the cells with information do not precisely match the landmass boundaries. The result of this missmatch is that not all environmental variables were available for 3,273 of the raster cells with occurrences (0.8% of the total). As expected, the vast majority of these problematic cells were close to the shore. To address this issue, we calculated the average value of environmental variables within a 5km buffer of the centre of the cell where the variable was missing and used it to approximate the value of the variable in that cell. Using this procedure, we were able to fill environmental variables for 89.3% of the cells

where they were missing. To fill the remaining 350 cells, we repeated the aforementioned procedure but instead using a 10km buffer. We removed from further analysis occurrences located within the 135 cells for which we were unable to fill environmental variables (0.03% of the total).

Next, we calculated the probability density function of the species distribution in environmental space. 167 To determine the environmental space, we used the first two components from a principal component 168 analysis of the 37 bioclimatic variables associated with the species occurrences. Specifically we used 169 the dudi.pca function from the R package ade4 1.7.13 (Dray & Dufour 2007) and centered and scaled all bioclimatic variables to have a mean of zero and a unit variance. We then determined the 171 position of species occurrences in the environmental space and estimated their bivariate probability 172 density function. We used a kernel method to estimate this density and normalised it such that 173 it ranges between zero and one. We used the kernel density method in the niche-factor analysis 174 (Broennimann et al. 2012) rather than the distance from the mode (Hirzel et al. 2002) (as has been 175 proposed earlier) because it has been shown to reduce the procedure's sensitivity to sampling effort 176 and the resolution of the environmental space. Specifically, to calculate the probability density 177 function we used ecospat.grid.clim.dyn from the R package ecospat 3.0 (Broennimann et al. 178 2018) with a grid resolution of 200. We then determined the location in the environmental space of 179 the plant-pollinator communities using the function suprow from ade4. The normalised density at 180 that particular location (which we calculated using the R package raster 2.8.19: Hijmans 2019) 181 corresponds with the bioclimatic suitability. The result of all these steps is the generation of a value 182 of environmental stress, which corresponds to one minus the bioclimatic suitability for a species of a 183 particular location. 184

We used a sensitivity analysis to determine the minimum number of occurrences that are necessary to have robust environmental stress estimations. For that we used the species with most occurrences available, *Archilochus colubris*, and calculated the mean absolute error of the bioclimatic suitability values obtained with one thousand subsamples from the 74,791 occurrences available from GBIF.

This analysis indicated that, for a species, we need roughly 18 independent occurrences for each community for which we aim to estimate the environmental stress. This is the number of occurrences necessary to maintain the mean absolute error of bioclimatic suitability below 0.1 (Figure 2). We therefore removed from further analyses 286 species for which we did not have enough occurrences

# error of bioclimatic suitability for a species present in two plant–pollinator communities 0.8 0.8 0.6 0.7 0.9 0.9 0.9 0.9 0.9 10% error 10% error

Figure 2: Sensitivity analysis of environmental stress error. The number of independent occurrences retrieved from GBIF is inversely related to the error of bioclimatic suitability for our plant-pollinator networks. The sensitivity analysis was performed by subsampling occurrences of *Archilochus colubris*, the species in our dataset with the largest number of occurrences in GBIF, which was recorded in two of our communities.

to obtain robust estimates.

### Data analysis

We then used a set of Bayesian multilevel models to evaluate the impact of environmental stress on 195 species specialisation. Specifically, we use the normalised degree of species as our response variable; 196 that is, the number of species it interacts with given the number of species in the opposite guild 197 (Martín González et al. 2010). The normalised degree was modelled using a logit link function and 198 a binomial distribution in which the number of partner species a focal species interacts with is the 199 number of successes, and the number of species in the opposite guild is the number of trials. We 200 are aware that whether species interact or not is not a Bernoulli process as species interactions are 201 not strictly independent from each other. However, the use of a binomial distribution allows us to 202 account for the differences in species richness across communities indirectly. Importantly, results 203 are qualitatively similar when we model species degree directly using a Poisson distribution and a 204 logarithmic link function. 205

<sup>206</sup> We evaluated four models to assess the relative importance of stress for each focal species.

A first model, our baseline model, included five variables. The predictors in the baseline model were the environmental stress of the community for each focal species, the species' number of

known possible partners in the community, and both the species guild (plant or a pollinator) and its interaction with environmental stress. We included the number of known possible partners 210 as a predictor in our models as it allows us to control for the effects of the environment on 211 community composition, effectively accounting for species co-occurrence. We calculated this metric 212 by determining the number of partners with which the species is known to interact in any other 213 community and assessing how mant of these potential partners are present in the local community. 214 Controlling for the number of potential partners makes our model a particularly stringent test of our 215 environmental-stress hypotheses because this variable could explain a large proportion of variance. 216 Often, the potential and the actual number of partners is the same or very close to each other, 217 especially for rare species present only in a few communities. 218

We allowed the intercept and slope of the stress-specialisation relationship to vary among species. 219 This approach allowed us to investigate two questions. First, it allows us to inspect the extent to 220 which environmental stress affects species similarly. Second, by investigating the correlation between 221 the intercept and the slope as a model parameter, it allowed us to inspect the extent by which 222 species with a small or large number of partner species respond to increasing levels of environmental 223 stress. To account for unmeasured differences between communities, like sampling effort, sampling 224 method, or diversity, we also allowed the model intercept to be different for each community in our 225 study. To facilitate model interpretation and convergence, we scaled all continuous variables to have 226 a mean of zero and a unit variance. 227

We compared this baseline model with three alternative models in which we removed one predictor 228 at a time. To quantify the difference between models, in terms of their expected out-of-sample 229 performance, we use the Wanatabe-Akaike information criterion (WAIC). All models were fitted 230 under a Bayesian framework using the R package brms 2.8.0 (Bürkner 2017, 2018) as an interface 231 for Stan (Carpenter et al. 2017). For each model, we used four Markov chains of 4,000 iterations 232 each; we used half of the iterations for warmup. We used weakly informative priors for all model 233 parameters. Specifically we used normal priors of mean zero and standard deviation ten for the 234 population-level effects and the intercepts, a half-Cauchy prior with a location of zero and a scale of 235 two for the standard deviations, and, when applicable, an LKJ-correlation prior with parameter  $\zeta = 1$  for the correlation matrix between group-level parameters.

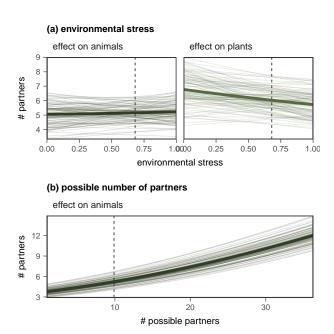


Figure 3: Conditional effects of predictors in our baseline model. The shown values are based on predictions for a hypothetical community with 76 plants and 33 pollinators. These values correspond to the median number of species in each guild across communities. In each panel, we condition on the mean value of the other predictor in the model. We indicate mean values for each predictor with a vertical dashed line. For model fitting, we scaled all predictors to have a mean of zero and unit variance; however, here we show the unscaled predictors to facilitate interpretation. To illustrate the uncertainty around the fitted estimates, we plot the fits of 100 independent draws from the posterior distribution. The thick lines indicate the mean values of the response distribution. As there was no interaction between the guild and the number of possible interactions, we only show the conditional effect of pollinators.

### ${f Results}$

Our models performed relatively well. The Bayesian R<sup>2</sup> for our baseline model was 0.89, which 239 indicates that our models were able to capture a large proportion of the variability on the data. 240 Overall, we found that environmental stress does not have a consistent effect across species. Indeed, 241 when looking at the fixed effects, stress has virtually no relationship with the normalised degree—our 242 metric of specialisation (Figure 3a). However, environmental stress was still an important predictor 243 in our model. The difference in WAIC between our baseline model and the model that did not 244 include environmental stress was  $489 \pm 94$  (Table 1). This apparent discrepancy can be explained 245 by the variability of the specialisation-stress relationship across species. 246

For some species, there is a strong negative relationship between stress and specialisation, while for others, there is a strong positive relationship (Figure 4a). Interestingly, the slope of this

Table 1: Comparison in out of sample predictive power of the baseline model (bold) and their alternatives. We rank models by their expected log predictive density based on their Wanatabe-Akaike information criterion (WAIC).

predictors	WAIC	SE
stress x guild + # possible partners	6,592	170
stress + # possible partners	6,595	166
guild + # possible partners	7,081	202
stress x guild	8,041	290

relationship was correlated with the species' specialisation intercept in the model (Figure 4b and c; the mean correlation coefficient was 0.52 [0.33, 0.67]). This correlation indicates that the slope of 250 the stress-specialisation relationship was more likely to be positive for species with many partners 251 under average stress conditions (and negative for species fewer partners). Recall that the model 252 estimates the intercept of the number of partners for each species at the mean value for stress across communities (0.68). Extrapolating to no-stress conditions this means that species that would 254 interact with few partners under no stress are more likely to interact with more partners as stress 255 increases, whereas those that would interact with many partners are more likely to interact fewer. As expected, we found a strong and positive relationship between the number of possible interactions 257 and the number of realised interactions in the community (i.e. species that were capable of being 258 generalists tended to realise this generality at all locations). There was also a large difference 250 of WAIC between the model that included this predictor and that that excluded it. This result 260 indicates that the availability of potential partners—this is, community composition—accounts for a large proportion of the variability in species degree. Importantly, our findings relating to 262 the variability of the stress-specialisation relationship were qualitatively unchanged, whether we 263 accounted for community composition or not. The standard deviation (in the parameters scale) of the community intercepts was 1.02 [0.85, 1.23] 265 which indicates the importance of the local context when determining specialisation. The standard 266 deviation of the species intercept was 0.54 [0.48, 0.61], and that of the species' stress slope was 0.38 267 [0.32, 0.44] (95\% credible intervals shown within square brackets).

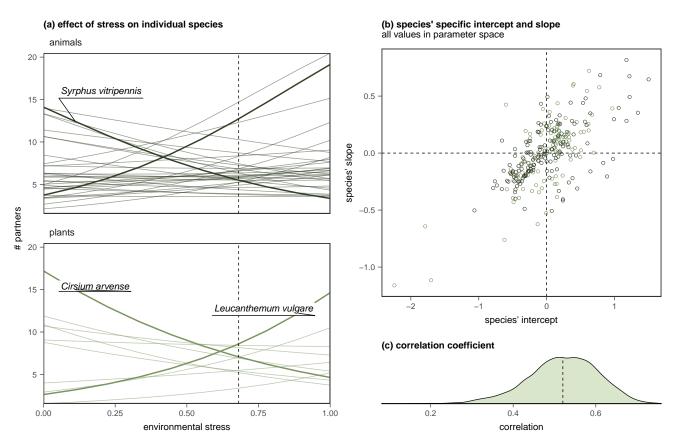


Figure 4: Species-level effects of environmental stress (a) Conditional effect of stress for individual species. Each line corresponds to the median relationship for each species. Although we included in the analysis all species that are present in two or more communities, to facilitate visualisation here, we show only species present in at least six communities (10 plants and 33 pollinators). As in the previous figure, fitted values assume a hypothetical community of median size. In each panel, we highlight two species for which the relationship between environmental suitability and the normalised degree was particularly strong. (b and c) There was a positive correlation between the species' intercept and the species' stress-specialisation slope. The species' intercept can be interpreted as the number of partners a species would have under mean levels of environmental stress.

### 269 Discussion

We set out to explore whether and how environmental stress can systematically affect specialisation. 270 After accounting for the pool of potential partners, we found that environmental conditions contribute 27 to determining whether a species is a generalist or a specialist in their community. We also found that 272 the particular effect of the environment dependends strongly on the species' inherent tendency to 273 be a specialist or generalist. Based on existing literature, we proposed two alternative hypotheses of 274 how environmental stress may affect species' specialisation, and we found evidence for both. Species 275 with many partners in low-stress communities were more likely to have a negative relationship and 276 hence reduce the number of partners as stress increases. Contrastingly, species in our datasets with 277 few partners in low-stress communities were more likely to interact with more partners in more stressful communities. 270 Our results suggest that changes in community composition are indeed the primary channel through 280 which the environment determines changes interaction probability, and this is a common mechanism 281 through which ecological networks respond to environmental change (Tylianakis & Morris 2017). 282 However, they also show that, for a large number of species, the environment may also play a 283 substantial role in determining their level of specialisation. Previous research has recognised that 284 environmental factors may help explain the changes in network structure along environmental 285 gradients that cannot be explained by community composition (Tylianakis et al. 2007). However, how these two factors were linked had been elusive so far (Gravel et al. 2018). We believe that 287 part of this difficulty could have arisen because species, and ultimately network structure, can 288 respond in multiple, and contrasting, ways depending on the particular bioclimatic variable examined 289 (e.g. temperature or precipitation). Using stress to summarise the effect on species of multiple 290 environmental gradients allowed us to detect a clear signal of the environment in species' interaction 291 patterns. 292 Recent research suggests that species are continuously changing their interaction partners wherever 293 environmental conditions change in space or time (Raimundo et al. 2018). So far it appears that 294 this rewiring is primarily driven by generalist species (Burkle et al. 2013; Ponisio et al. 2017), 295 presumably because generalist species are less sensitive to trait matching of their interaction partners 296

(CaraDonna et al. 2017). Our results add two important nuances to these findings. First, because "generalists" seem to focus on fewer partners as environmental conditions deteriorate, we show that trait matching might still play a role in determining the less-flexible core interactions of generalist 299 species. Second, and most importantly, our results suggest that only a small proportion of species 300 are "true generalists" or "true specialists" this is, species that interact with a large or small number 301 of partners regardless of the environmental stress, respectively. This pattern implies that rewiring is 302 not exclusive to species with many partners. Instead, at least a fraction of the species that appear to 303 be specialist in their communities might be as flexible, if not more, than those with a large number 304 of partners, effectively behaving as facultative generalists in the face of environmental change. 305 These flexible "specialists" might therefore have a more significant role in network persistence than 306 previously expected. 307

In our model, we can roughly divide species between true specialists, true generalists, and flexible 308 species. However, there is a fourth group that remained invisible to our model but has important 309 implications for network persistence and stability. Species that can vary their interaction partners 310 flexibly and their role in the network are more likely to persist in their community as environmental 311 conditions vary (Gaiarsa et al. 2019). We propose this fourth group is composed of true specialists 312 that are constrained to interact with partners of high trait-matching and therefore were not likely 313 to be found in more than one community. If species that are not flexible are unlikely to persist over 314 temporal or spatial environmental gradients, we can expect specialised communities that are highly 315 constrained by trait-matching (like some plant-hummingbird networks; Vizentin-Bugoni et al. 2014; 316 Maruyama et al. 2014) to be far more vulnerable to increased climate change-induced environmental 317 stress and habitat degradation than communities where role and interaction flexibility are more 318 prevalent. 319

Similarly, if the patterns we see in our models have also played a role during the evolutionary history of pollination communities, our results also help to explain why only a small fraction of plant-pollinator interactions shows a strong signature of deep co-evolutionary history (Hutchinson *et al.* 2017). The increases in the stress that species are predicted to experience due to rapid environmental change might further erode the co-evolutionary history of specialist species. Communities as a whole might be in a trajectory of even more diffuse co-evolution. For specialists, at least, the

longer-term benefits of being able to interact with multiple partners might be more important than
the shorter-term benefits of interacting with partners of high trait matching.

The structural implications of the "regression towards the mean" that environmental stress promotes are less clear. However, it is plausible to expect that nestedness, and therefore network stability, might be reduced in the face of rapid environmental change. Determining exactly how the changes in degree caused by environmental stress are reflected by systematic changes in network structure would be an interesting avenue of research. Answering this question would require expanding our suitability analysis to all species in the community and comparing the degree distribution of networks along a gradient of stress for the community as a whole.

In conclusion, we show that the environment can affect the specialisation level of plants and pollinators in systematic ways beyond community composition. Species that are inflexible with their interaction partners are unlikely to persist under more stressful environmental conditions. However, we show that many species are flexible in regards to their specialisation levels and therefore are not inherently generalists or specialists. Instead, the species' level of specialisation/generalisation should be considered on a relative scale depending on environmental stress at a given location.

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