

1)

(a) (i) The inverse of a symmetric matrix is itself symmetric.
True.

We want to show $A^{-1} = A^{-T}$ given $A = A^T$

Proof: $I = A A^{-1} = A^T A^{-1}$

multiply A^{-T} on both sides (left multiply)

$$A^{-T} = A^{-T} A^T A^{-1}$$

$$A^{-T} = A^{-1} \quad \#$$

(ii) All 2×2 orthogonal matrices have the following form

$$\begin{bmatrix} \cos \theta & -\sin \theta \\ \sin \theta & \cos \theta \end{bmatrix} \text{ or } \begin{bmatrix} \cos \theta & \sin \theta \\ \sin \theta & -\cos \theta \end{bmatrix} \quad \text{True}$$

let $P \in \mathbb{R}^{2 \times 2}$, $P = \begin{bmatrix} a & b \\ c & d \end{bmatrix}$

if P is orthogonal, then $P^T P = I$

$$\Rightarrow \begin{pmatrix} a^2 + c^2 & ab + cd \\ ab + cd & b^2 + d^2 \end{pmatrix} = \begin{pmatrix} 1 & 0 \\ 0 & 1 \end{pmatrix}$$

$$\Rightarrow a^2 + c^2 = 1, \quad b^2 + d^2 = 1, \quad ab + cd = 0$$

Without loss of generality, let $a = \cos \theta$, $c = \sin \theta$ and $b = -c$

Which are $\begin{bmatrix} \cos \theta & -\sin \theta \\ \sin \theta & \cos \theta \end{bmatrix}$ or $\begin{bmatrix} \cos \theta & \sin \theta \\ \sin \theta & -\cos \theta \end{bmatrix}$ $b = \sin \theta$ $d = \cos \theta$ $a = d$
 or $b = c$ $a = -d$

(iii) Assume $A = C C^T$, then we have

$$x^T A x = x^T C C^T x = \|C^T x\|_2^2 \geq 0 \Rightarrow A \text{ is positive semi-definite.}$$

Then $A = U \Sigma U^T$ with $\Sigma_{ii} \geq 0$ which means all eigenvalues of A is non-negative.

Also $\text{tr}(A) = \sum \lambda_i$ λ_i $i=1, 2, \dots$ are the eigenvalues.

$$\text{tr}(A) = -8 - 5 - 2 = -15 \Rightarrow \text{Not all eigenvalues are non-negative.}$$

Therefore, A cannot be written as $A = C C^T$

2. (a) (i) $E_Y[E_X(X|Y)] = E[X]$

Proof: $E_Y[E_X(X|Y)] = \int_{R_Y} E[X|y] P(y) dy = \int_{R_Y} \int_{R_X} x P(x|y) dx P(y) dy$

$$= \int_{R_X} \int_{R_Y} x P(x,y) dy dx$$

$$= \int_{R_X} x P(x) dx = E[X]$$

R_X, R_Y are the region of X and Y respectively.

(ii) $E[I_{X \in C}] = P(X \in C)$

Proof: $E[I_{X \in C}] = \int_{\mathbb{R}} P(x) I_{X \in C} dx = \int_C P(x) dx = P(X \in C)$

(iii) $\text{var}[X] = E_Y[\text{var}_X[X|Y]] + \text{var}_Y[E_X[X|Y]]$

Proof: $\text{var}_X[X|Y] = E_X[(X - E[X|Y])^2 | Y]$

~~$E_Y[\text{var}_X[X|Y]] = E_Y[E_X[(X - E[X|Y])^2 | Y]]$~~
 ~~$= E_Y[E_X[X^2 - 2XE[X|Y] + E[X|Y]^2 | Y]]$~~
 ~~$= E_Y[E_X[X^2 | Y] - 2E[X|Y]E[X|Y] + E[X|Y]^2]$~~
 ~~$= E_Y[E[X^2 | Y] - 2E[X|Y]^2 + E[X|Y]^2]$~~
 ~~$= E_Y[E[X^2 | Y] - E[X|Y]^2]$~~

$$E[\text{var}(X|Y)] = E[E[(X - E[X|Y])^2 | Y]]$$

$$= E[(X - E[X|Y])^2] \quad (1)$$

$$\text{var}(E_X[X|Y]) = E[(E[X|Y] - E(E[X|Y]))^2]$$

$$= E[(E[X|Y] - E[X])^2] \quad (2)$$

$$\text{var}(X) = E[(X - E[X])^2] = E[(X - E[X|Y] + E[X|Y] - E[X])^2]$$

$$= E[(X - E[X|Y])^2] + E[(E[X|Y] - E[X])^2]$$

$$+ 2E[(X - E[X|Y])(E[X|Y] - E[X])] \quad (3)$$

$$\begin{aligned}
& E[(X - E[X|Y]) \cdot (E[X|Y] - E[X])] \\
&= E[X \cdot E[X|Y] - X \cdot E[X] - E[X|Y] \cdot E[X] + E[X] \cdot E[X|Y]] \\
&= E[X \cdot E[X|Y]] - E^2[X] - E[E^2[X|Y]] + E[X] \cdot E(E[X|Y]) \\
&= E\{E[X \cdot E[X|Y] | Y]\} - E[E^2[X|Y]] - E^2[X] + E^2[X] \\
&= E\{E[X|Y] \cdot E[X|Y]\} - E[E^2[X|Y]] = 0 \quad (4)
\end{aligned}$$

Combining (1) (2) (3) (4), we get

$$\text{var}(X) = E[\text{var}(X|Y)] + \text{var}(E[X|Y])$$

$$\begin{aligned}
\text{(iv)} \quad E[XY] &= \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} xy P_{XY}(x, y) dx dy \quad \text{if } X, Y \\
&= \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} xy P_X(x) P_Y(y) dx dy \quad \text{by independence} \\
&= \int_{-\infty}^{\infty} x P_X(x) dx \cdot \int_{-\infty}^{\infty} y P_Y(y) dy \\
&= E[X] \cdot E[Y]
\end{aligned}$$

$$\text{(v)} \quad E[X] = 0 \cdot P(X=0) + 1 \cdot P(X=1) = P(X=1)$$

$$E[Y] = P(Y=1)$$

$$E[XY] = \sum_{x=0}^1 \sum_{y=0}^1 xy P(X=x, Y=y)$$

$$= P(X=1, Y=1)$$

$$E[XY] = E[X] \cdot E[Y]$$

$$\Rightarrow P(X=1, Y=1) = P(X=1) \cdot P(Y=1)$$

$\Rightarrow X, Y$ are independent.

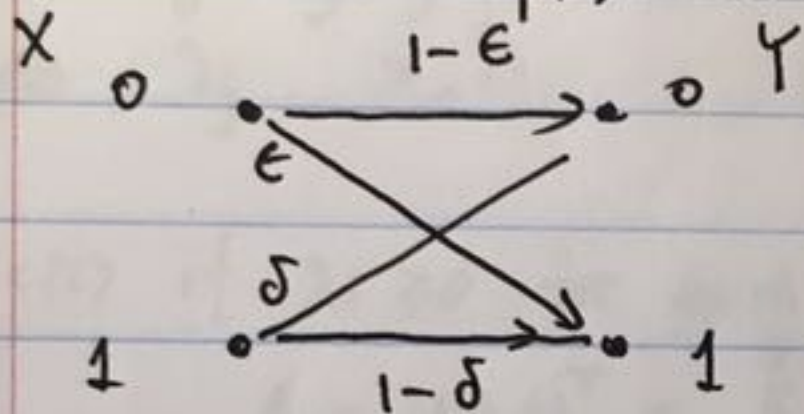
(b) (i) $P(H=h, D=d) \leq P(H=h)$

$$P(H=h) = \sum_x P(H=h, D=x) > P(H=h, D=d)$$

(ii) $P(H=h | D=d)$ depends $P(H=h)$

$$P(H=h) = \sum_k P(H=h | D=k) P(D=k)$$

For example, consider Binary Channel Model in Communication system



$$P(Y=0 | X=0) = 1-\epsilon$$

$$P(Y=1 | X=0) = \epsilon$$

$$P(Y=0 | X=1) = \delta$$

$$P(Y=1 | X=1) = 1-\delta$$

$$P(X=0) = \frac{1}{2}$$

$$P(X=1) = \frac{1}{2}$$

Thus
$$P(Y=0) = P(Y=0 | X=0) P(X=0) + P(Y=0 | X=1) P(X=1)$$

$$= (1-\epsilon) \times \frac{1}{2} + \delta \times \frac{1}{2} = \frac{1}{2} (1-\epsilon + \delta)$$

$$P(Y=0 | X=1) = \delta$$

The relationship between $P(Y=0)$ and $P(Y=0 | X=1)$ depends on the choice of ϵ, δ .

(iii) $P(H=h | D=d) \geq P(D=d | H=h) P(H=h)$

According Bayes Rule,

$$P(H=h | D=d) = \frac{P(D=d | H=h) P(H=h)}{P(D=d)}$$

Since $P(D=d) \leq 1$, thus

$$P(H=h | D=d) \geq P(D=d | H=h) P(H=h)$$

3) (a) Let the j th eigenvector u_j , then

(i) $u_j^T A u_j \geq 0$ since A is PSD.

$$\Leftrightarrow u_j^T U \Lambda U^T u_j \geq 0$$

$$\Leftrightarrow u_j^T \left(\sum_{i=1}^d \lambda_i u_i u_i^T \right) u_j \geq 0$$

$$\Leftrightarrow u_j^T \left(\sum_{i=1}^d \lambda_i u_i u_i^T u_j \right) \geq 0$$

$$\Leftrightarrow u_j^T \lambda_j u_j \geq 0 \quad \text{since } U \text{ is orthogonal}$$

$$\Leftrightarrow \lambda_j \geq 0 \quad \text{for all } j = 1, 2, \dots, d$$

(ii) if $\lambda_i \geq 0$ for each i .

$$A = U \Lambda U^T = \sum_{i=1}^d \lambda_i u_i u_i^T$$

let $x \in \mathbb{R}^d$

$$x^T A x = x^T \sum_{i=1}^d \lambda_i u_i u_i^T x$$

$$= \sum_{i=1}^d \lambda_i \|u_i^T x\|_2^2$$

$$\geq 0 \quad \text{for all } x$$

Thus A is PSD.

Combine (i) and (ii), we have A is PSD iff $\lambda_i \geq 0$ for each i .

(b) (i) If A is PD. By the same arguments before, we have

$$u_j^T A u_j > 0$$

u_j is the j th eigenvector

$$\Leftrightarrow u_j^T \sum_{i=1}^d \lambda_i u_i u_i^T u_j > 0$$

$$\Leftrightarrow u_j^T \lambda_j u_j > 0$$

$$\Leftrightarrow \lambda_j > 0 \quad \text{for all } j = 1, 2, \dots, d$$

(ii) If $\lambda_i > 0$ for each i .

let $x \in \mathbb{R}^d$,

$$\begin{aligned} x^T A x &= x^T \sum_{i=1}^d \lambda_i u_i u_i^T x \\ &= \sum_{i=1}^d \lambda_i \|u_i^T x\|_2^2 \end{aligned}$$

> 0 for all $x \neq 0$

Thus A is PD.

Combine (i) and (ii) A is PD iff $\lambda_i > 0$ for each i .

$$\begin{aligned}
 4) \quad (a) \quad f(tx + (1-t)y) &= a^T(tx + (1-t)y) + b \\
 &= t \cdot a^T x + (1-t) \cdot a^T y + t \cdot b + (1-t) \cdot b \\
 &= t(a^T x + b) + (1-t)(a^T y + b) \\
 &= t f(x) + (1-t) f(y) \quad \textcircled{1}
 \end{aligned}$$

Therefore $f(x) = a^T x + b$ is convex.

By the same reasoning, we can also have ~~$f(x)$~~

$$-f(tx + (1-t)y) = -t f(x) + (1-t) f(y) \quad \textcircled{2}.$$

implies $-f(x)$ is convex, thus $f(x)$ is concave.

From $\textcircled{1} \textcircled{2}$, the equality holds, so $f(x)$ is not strictly convex.

(b) Suppose there exists more than one global minimizers x^*, y^* where $x^* \neq y^*$
we have $f(x^*) = f(y^*)$

Since $f(x)$ is strictly convex on $\text{dom}(f)$, we have

$$f(tx^* + (1-t)y^*) < t f(x^*) + (1-t) f(y^*)$$

$$\Leftrightarrow f(t(x^* - y^*) + y^*) < f(y^*) \quad (\text{or } f(x^*))$$

\Rightarrow there exists $t(x^* - y^*) + y^* \in \text{dom}(f)$ such that

$$f(t(x^* - y^*) + y^*) < f(y^*)$$

which contradicts that y^* is the global minimizer.

Therefore f has at most one global minimizer.

(c) from (a), we have $f(x) = f(y) + \nabla f^T(y)(x-y) + \frac{1}{2}(x-y)^T \nabla^2 f(y+t(x-y))(x-y)$
 let $y = x^*$ and $t=0$, we get

$$f(x) = f(x^*) + \nabla f^T(x^*)(x-x^*) + \frac{1}{2}(x-x^*)^T \nabla^2 f(x^*)(x-x^*)$$

$$\Rightarrow f(x) - f(x^*) = \frac{1}{2}(x-x^*)^T \nabla^2 f(x^*)(x-x^*) \quad \text{since } \nabla f(x^*) = 0$$

$$\Rightarrow (x-x^*)^T \nabla^2 f(x^*)(x-x^*) \geq 0 \quad \text{since } x^* \text{ is local minimum.}$$

$$\Rightarrow \nabla^2 f(x^*) \succeq 0$$

(d) From (a), we get

$$f(x) = f(y) + \nabla f^T(y)(x-y) + \frac{1}{2}(x-y)^T \nabla^2 f(y+t(x-y))(x-y) \quad \text{for all } x, y \in \mathbb{R}^d$$

① Since f is convex, we have

$$f(x) \geq f(y) + \nabla f^T(y)(x-y) \quad \text{1st order condition.}$$

$$\text{Therefore, } \frac{1}{2}(x-y)^T \nabla^2 f(y+t(x-y))(x-y)$$

$$= \frac{1}{2}(x-y)^T \nabla^2 f(t x + (1-t)y)(x-y) \geq 0$$

let $t=1$

$$\Rightarrow \nabla^2 f(x) \succeq 0 \quad \text{for all } x \in \mathbb{R}^d$$

② If $\nabla^2 f(x) \succeq 0$ for all $x \in \mathbb{R}^d$, then

$$a^T \nabla^2 f(x) a \geq 0$$

$$\Rightarrow a^T \nabla^2 f(y+t(x-y)) a \geq 0 \quad \text{since } y+t(x-y) \in \mathbb{R}^d$$

$$\Rightarrow (x-y)^T \nabla^2 f(y+t(x-y))(x-y) \geq 0$$

$$\Rightarrow f(x) \geq f(y) + \nabla f^T(y)(x-y)$$

Therefore f is convex.

Combining ① and ② if f is twice differentiable, then f is convex iff $\nabla^2 f(x)$ is PSD. $\forall x \in \mathbb{R}^d$

$$(e) \quad \frac{\partial^2 f}{\partial x_\ell \partial x_k}(x) = \frac{\partial}{\partial x_\ell} \left[\frac{\partial f(x)}{\partial x_k} \right] = \frac{\partial}{\partial x_\ell} \frac{\partial}{\partial x_k} \left[\frac{1}{2} x^T A x + b^T x + c \right]$$

$$= \frac{\partial}{\partial x_\ell} \frac{\partial}{\partial x_k} \left[\frac{1}{2} \sum_{i=1}^d x_i \sum_{j=1}^d A_{ij} x_j + \sum_{i=1}^d b_i x_i + c \right]$$

$$= \frac{\partial}{\partial x_\ell} \left[\sum_{j=1}^d A_{kj} x_j + b_k \right]$$

$$= A_{k\ell} = A_{\ell k} \quad \text{since } A \text{ is symmetric.}$$

$$\text{Therefore } \nabla_x^2 f(x) = A$$

According to (d), we have

f is convex if and only if A is positive semi-definite on \mathbb{R}^d

Also, by the same argument followed by (d), we conclude

f is strictly convex if and only if A is positive definite on \mathbb{R}^d

In [13]:

```
# Load the Libraries
import math
import numpy as np
import matplotlib.pyplot as plt
%matplotlib inline

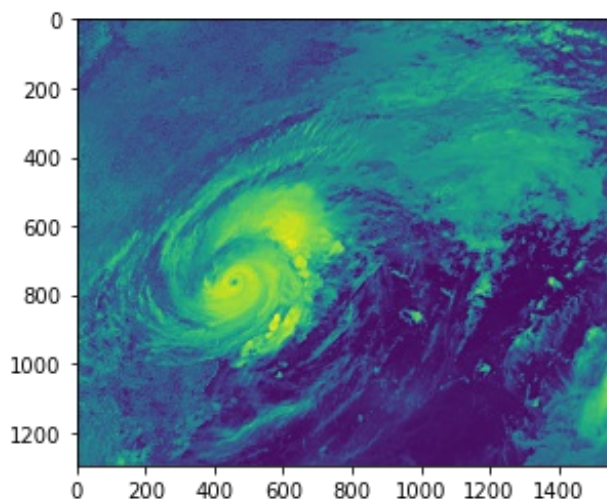
import cv2
```

In [6]:

```
# Load the color image in grayscale
img = cv2.imread('harvey-saturday-goes7am.jpg')
print("The size of the color image is " + str(img.shape))
gray = cv2.cvtColor(img, cv2.COLOR_RGB2GRAY)
print("The size of the grayscale image is " + str(gray.shape))
plt.imshow(gray)
plt.show()
```

The size of the color image is (1296, 1548, 3)

The size of the grayscale image is (1296, 1548)



In [25]:

```
# SVD
U, Sigma, V_trans = np.linalg.svd(gray, full_matrices=1)
print(U.shape)
print(Sigma.shape)
print(V_trans.shape)
```

(1296, 1296)

(1296,)

(1548, 1548)

In [29]:

```
def F_norm(gray):
    """
    This function calculate the F-norm given a input matrix
    """
    tr_AA_transpose = np.trace(np.matmul(gray,gray.transpose()))
    norm = math.sqrt(tr_AA_transpose)
```



```

norm = math.sqrt(U_AA.transpose())
    return norm
print(F_norm(gray))

```

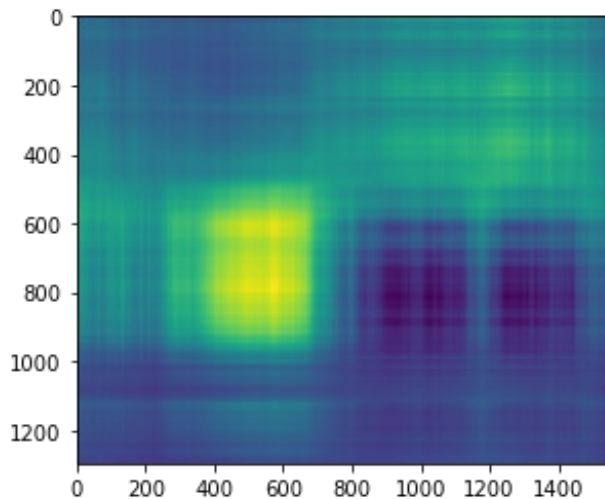
411.85191513455413

In [36]:

```

k = 2
X_bar = 0
for i in range(k):
    X_bar += Sigma[i] * np.outer(U.T[i], V_trans[i])
plt.imshow(X_bar)
plt.show()

```



In [39]:

```

dif = np.subtract(gray,X_bar)
print(F_norm(dif)/F_norm(gray))

```

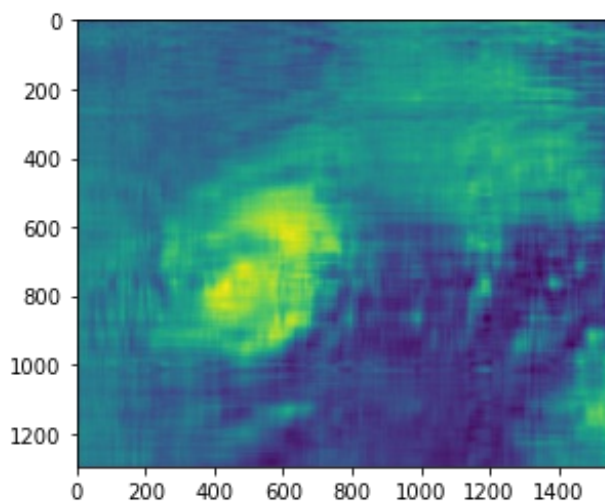
102.4922553316057

In [40]:

```

k = 10
X_bar = 0
for i in range(k):
    X_bar += Sigma[i] * np.outer(U.T[i], V_trans[i])
plt.imshow(X_bar)
plt.show()

```



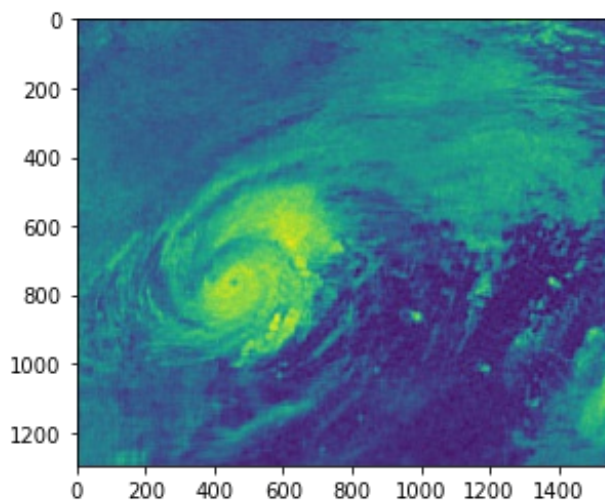
In [41]:

```
dif = np.subtract(gray,X_bar)
print(F_norm(dif)/F_norm(gray))
```

58.93770105042957

In [42]:

```
k = 40
X_bar = 0
for i in range(k):
    X_bar += Sigma[i] * np.outer(U.T[i], V_trans[i])
plt.imshow(X_bar)
plt.show()
```



In [43]:

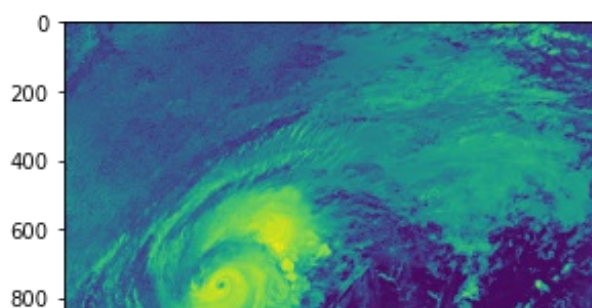
```
dif = np.subtract(gray,X_bar)
print(F_norm(dif)/F_norm(gray))
```

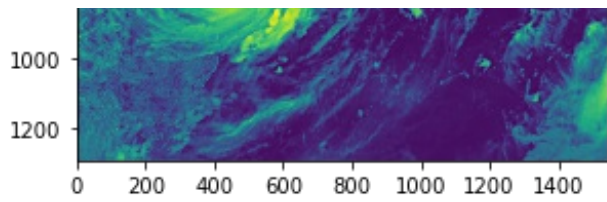
31.300100215056247

In [44]:

```
# this is only for test
k = 1296
X_bar = 0
for i in range(k):
    X_bar += Sigma[i] * np.outer(U.T[i], V_trans[i])
plt.imshow(X_bar)
plt.show()

dif = np.subtract(gray,X_bar)
print(F_norm(dif)/F_norm(gray))
```





(b) According to the fundamental theorem, $(1296 \times k) + (k \times k) + (1548 \times k) = 2844 \times k + k^2$

When $k = 2$, we need $(1296 \times 2) + (2 \times 2) + (1548 \times 2) = 5692$

When $k = 10$, we need $(1296 \times 10) + (10 \times 10) + (1548 \times 10) = 28540$

When $k = 40$, we need $(1296 \times 40) + (40 \times 40) + (1548 \times 40) = 115360$