Randomised Algorithms Winter term 2022/2023, Exercise Sheet No. 5

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Exercise 1.

- (a) We have established in the notes that the probability of not making any mistakes after n-t contractions is at least $\frac{t(t-1)}{n(n-1)}$, in this case, when t=n-1, we would get that this probability is $\frac{n-2}{n}=1-\frac{2}{n}$, which is in fact equal to q_n .
- (b) We denote the probability that S_1 (resp S_2) is correct as P_1 (resp. P_2), we have:

$$P_1 = \mathbb{P}[\{Output \ of \ L5 \ is \ correct\} | \{Contraction \ at \ L4 \ is \ correct\}] \mathbb{P}[\{Contraction \ at \ L4 \ is \ correct\}] = P(n-1)q_n$$

$$P_2 = P(n)$$

The algorithm is successful if either the return at line 7 was a succes (i.e., S_1 is a success) or the return at line 10 was a succes, which yield the following:

$$\begin{split} p(n) &= \mathbb{P}[\{Entered\ L7\}] \mathbb{P}[\{Succes\} | \{entered\ L7\}] + \mathbb{P}[\{Didn't\ Enter\ L7\}] \mathbb{P}[\{Succes\} | \{Didnt\ enter\ L7\}] \\ &= q_n \mathbb{P}[\{Succes\} | \{entered\ L7\}] + (1-q_n) \mathbb{P}[\{Succes\} | \{Didnt\ enter\ L7\}] \\ &= q_n \mathbb{P}[\{S_1\ is\ succesful\}] + (1-q_n) \mathbb{P}[\{The\ best\ of\ S_1,\ S_2\ was\ succesful\}] \\ &= q_n \mathbb{P}[\{S_1\ is\ succesful\}] + (1-q_n)(1-\mathbb{P}[\{S_1\ and\ S_2\ failed\}]) \\ &= q_n P_1 + (1-q_n)[1-(1-P_1)(1-P_2)] \qquad (We\ will\ just\ keep\ re-arranging\ after\ now) \\ &= q_n^2 P(n-1) + (1-q_n)[1-(1-q_nP(n-1))(1-P(n))] \\ &= q_n^2 P(n-1) + (1-q_n)[P(n) + q_nP(n-1) - q_nP(n-1)P(n)] \\ &= q_n^2 P(n-1) + (1-q_n)P(n) - q_n(1-q_n)P(n-1)P(n) \\ &\Longrightarrow \\ q_n P(n) &= q_n P(n-1) - q_n(1-q_n)P(n-1)P(n) \\ &\Longrightarrow \\ P(n) &= P(n-1) - (1-q_n)P(n-1)P(n) \qquad (We\ would\ verify\ line\ 1\ if\ n=2,\ hence\ q_n\ would\ never\ equal\ 0\) \end{split}$$

(c) By dividing the equation that we derived in the last question by P(n)P(n-1), we get the following, for $k \in \{3, ..., n\}$:

$$\frac{1}{P(k-1)} - \frac{1}{P(k)} = -\frac{2}{k}$$

$$\implies$$

$$\sum_{k=3}^{n} \frac{1}{P(k-1)} - \sum_{k=3}^{n} \frac{1}{P(k)} = -\sum_{k=3}^{n} \frac{2}{k}$$

$$\frac{1}{P(2)} - \frac{1}{P(n)} = -\sum_{k=3}^{n} \frac{2}{k}$$

$$\implies$$

$$P(n) = \frac{1}{\sum_{k=3}^{n} \frac{2}{k} + 1}$$

Comparisons of the probability of succes with FastCut:

- GeoContraction has a $\Theta(\frac{1}{\log(n)})$ (Assuming the bounds are exact like suggested in (b))
- FastCut had a $\Omega(\frac{1}{\log(n)})$

Exercise 2.

Let X_t represent the number of walks that are left to reach home at step t. Obviously $X_t \in \{0, \ldots, n\}$, and we are interested in computing $\mathbb{E}[T]$ where $T = \inf\{t \geq 0 | X_t = 0\}$.

We have the following:

$$\begin{split} \mathbb{E}_{(X_t, X_{t+1})}[X_t - X_{t+1} | X_t &= s] = \mathbb{E}_{X_{t+1}}[s - X_{t+1} | X_t = s] \\ &= s - \mathbb{E}[X_{t+1} | X_t = s] \\ &= s - \sum_{X_{t+1}} \mathbb{P}[X_{t+1} | X_t = s] X_{t+1} \end{split}$$

If s = n (we are at the bar), X_{t+1} has two options $\{n, n-1\}$

$$\mathbb{E}[X_t - X_{t+1}|X_t = s] = n - n\mathbb{P}[X_{t+1} = n|X_t = s] - (n-1)\mathbb{P}[X_{t+1} = n - 1|X_t = s]$$

$$= \frac{1}{5}$$

If 0 < s < n, X_{t+1} has two options $\{s - 1, s + 1\}$

$$\mathbb{E}[X_t - X_{t+1}|X_t = s] = s - (s-1)\mathbb{P}[X_{t+1} = s - 1|X_t = s] - (s+1)\mathbb{P}[X_{t+1} = s + 1|X_t = s]$$

$$= \frac{3}{5} + \frac{2}{5}$$

$$= 1$$

Hence, we get, for all $t > 0, s \neq 0$:

$$\frac{1}{5} \le \mathbb{E}[X_t - X_{t+1} | X_t = s] \le 1$$

Which yields the following bounds, assuming $X_0 = n$:

$$n \leq \mathbb{E}[T] \leq 5n$$

Exercise 3.

(a) As the state 0 is never reacheable and $X_0 = 1$, T could be formulated as follows:

$$T = \inf\{t > 1 | X_t = -1\}$$

i.e, the first time we reach the state s = -1 (our definition of "success", speaking from a geometric distribution terminology). By noting that each X_t is a Bernoulli experiment with p = 1/2, we conclude that $\mathbb{E}[T] = 1/p = 2$.

(b) The state we are interested to reach in the definition of T is -1, in the theorem the state of interest is s = 0, and X_t only takes positive values in the theorem statement, again, here X_t can take the value -1.

We have for $s \neq 0$: $\mathbb{E}[X_{t+1}|X_t = s] = 0$, and hence:

$$\mathbb{E}[X_t - X_{t+1} | X_t = s] = s - \mathbb{E}[X_{t+1} | X_t = s] = s$$

Which yields, $\mathbb{E}[X_t - X_{t+1} | X_t = s] \leq 1$ for all $s \neq 0$. using the additive drift theorem would lead to conclude that $1 \leq E[T]$

To make it work, we define the random process (Y_t) as $Y_t = X_t + 1$, clearly $Y_t \in \{0, 1, 2\}$, and

$$T = \inf\{t \ge 0 | X_t = -1\}$$

= \inf\{t \ge 0 | X_t + 1 = 0\}
= \inf\{t \ge 0 | Y_t = 0\}

The way we defined $(Y_t)_{t\geq 0}$, and how we formulated T using (Y_t) , matches the theorem's assumptions. Moreover we have: $[s \neq 0 \land \mathbb{P}(Y_t = s) > 0 \implies s = 2]$, and hence we only need to verify the theorem's conditions for s = 2:

$$\begin{split} \mathbb{E}[Y_t - Y_{t+1} | Y_t = 2] &= 2 - \mathbb{E}[Y_{t+1} | Y_t = 2] \qquad ((Y_{t+1} \ can \ only \ move \ to \ 0 \ or \ 2) \\ &= 2 - 2\mathbb{P}[Y_{t+1} = 2 | Y_t = 2] \\ &= 2 - 2\frac{1}{2} \\ &= 1 \end{split}$$

And hence, using the additive drift theorem, we get:

$$\frac{\mathbb{E}[Y_0]}{1} \leq \mathbb{E}[T] \leq \frac{\mathbb{E}[Y_0]}{1} \implies \mathbb{E}[T] = 2, since \ \mathbb{E}[Y_0] = 2$$

Which is the same as the answer that we got in (a).

Exercise 4.

(a) Let's first compute $\mathbb{E}[X_t - X_{t+1}|X_t = s]$ for some $0 < s \le n$:

$$\mathbb{E}[X_t - X_{t+1} | X_t = s] = s - \sum_{i=0}^{s-1} \mathbb{P}[X_{t+1} = i | X_t = s]i$$

$$= s - \frac{1}{s} \sum_{i=0}^{s-1} i$$

$$= \frac{s+1}{2}$$

Hence, we have $\mathbb{E}[X_t - X_{t+1}|X_t = s] \geq \frac{s+1}{2}$, for all $0 < s \leq n$, with $h(s): s \mapsto \frac{s+1}{2}$ being monotonically increasing, we have verified all assumptions of the Variable Drift Theorem, hence, we can conclude the following:

$$\mathbb{E}[T] \le \frac{1}{h(1)} + \int_{1}^{n} \frac{1}{h(x)} dx$$

$$= 1 + \int_{1}^{n} \frac{2}{s+1} dx$$

$$= 1 + 2 \int_{2}^{n+1} (\ln(x))' dx$$

$$= 1 + 2(\ln(n+1) - \ln(2))$$

$$\le 2\ln(n+1) \qquad (1 - 2\ln(2) \approx -0.39)$$

Using the Additive Drift Theorem (since $\mathbb{E}[X_t - X_{t+1} | X_t = s] \ge 1$, for all $0 < s \le n$) would lead to conclude that $\mathbb{E}[T] \le n$, which is less tight than the gap that we obtained using the Variable Drift Theorem.

(b) The function $h(s): s \mapsto \frac{1}{s}$ isn't monotonically increasing, hence we can't use the Variable Drift Theorem.

We have, for $0 < s \le n$, $\mathbb{E}[X_t - X_{t+1}|X_t = s] \ge \frac{1}{n}$, hence, using the Additive Drift Theorem, yields the following: $\mathbb{E}[T] \le n^2$

(c) Blabla

$$\mathbb{E}[T] = \frac{1}{h(1)} + \int_{1}^{n} x^{-\frac{1}{2}} dx$$
$$= 1 + \frac{1}{2} \int_{1}^{n} \sqrt{x} dx$$
$$= \frac{1}{2} (\sqrt{n} + 1)$$

(d) Consider the random process $(X_t)_{t\geq 0}$ on state space $\{0,1,\ldots,n\}$, where knowing that $X_t=s$ for s>0, X_{t+1} is defined as follows:

$$X_{t+1} = \begin{cases} s-1 & \text{with probability } \frac{1}{s} \\ s & \text{with probability } 1 - \frac{1}{s} \end{cases}$$

For s > 0 and $t \ge 1$:

$$\begin{split} \mathbb{E}[X_t - X_{t+1} | X_t &= s] = s - \mathbb{E}[X_{t+1} | X_t = s] \\ &= s - [(s-1)\frac{1}{s} + s(1 - \frac{1}{s})] \\ &= s - [1 - 1/s + s - 1] \\ &= \frac{1}{s} \end{split}$$