

QUANTIFYING THE RESPONSE OF BLAINVILLE'S BEAKED WHALES TO US NAVAL SONAR EXERCISES IN HAWAII

Eiren K. Jacobson^{1*}, E. Elizabeth Henderson², Cornelia S. Oedekoven¹, David
L. Miller¹, Stephanie L. Watwood³, David J. Moretti³, Len Thomas¹

¹ *Centre for Research into Ecological and Environmental Modelling, School of Mathematics and
Statistics, University of St Andrews, St Andrews, Scotland*

² *Naval Information Warfare Center Pacific, San Diego, CA, USA*

³ *Naval Undersea Warfare Center, Newport, RI, USA*

Corresponding author (email: eiren.jacobson@st-andrews.ac.uk)

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Abstract

Behavioral responses of beaked whales (family Ziphiidae) to naval use of mid-frequency active sonar (MFAS) have been quantified for some species and regions using risk functions, which give the probability of a response as a function of covariates such as received level. We develop a novel risk function for Blainville's beaked whales (BBWs) on the US Navy Pacific Missile Range Facility (PMRF) in Hawaii and compare our risk function to one developed for the same species in a different ocean basin. We used passive acoustic data collected at bottom-mounted hydrophones before and during six naval training exercises at PMRF along with modelled sonar received levels to describe the effect of training and MFAS on foraging groups of BBWs. We used a multi-stage generalized additive modelling (GAM) approach to control for the underlying spatial distribution of vocalizations under baseline conditions. At a MFAS received level of 150 dB re 1 μ Pa the probability of detecting groups of BBWs decreased by 78% (95% CI 62%-100%) when compared to periods when general training activity was ongoing and by 92% (95% CI 87%-100%) when compared to baseline conditions. Our results indicate a more pronounced response to naval training and MFAS than has been previously reported. [200/200]

1 Introduction

Beaked whales (family Ziphiidae) are a group of deep-diving cetaceans that rely on sound to forage, navigate, and communicate (Aguilar de Soto et al., 2012; Johnson, Madsen, Zimmer, Aguilar de Soto, and Tyack, 2004; Macleod and D’Amico, 2006). Multiple mass strandings of beaked whales have been associated with high-intensity anthropogenic sound sources. These acute events have motivated research into whether and how beaked whales respond to different types and intensities of anthropogenic noise (Cox et al., 2006).

Anthropogenic sound can disrupt the patterned foraging dive cycles of beaked whales (Falcone et al., 2017), potentially leading to cumulative sublethal impacts resulting from reduced foraging opportunities (L. F. New, Moretti, Hooker, Costa, and Simmons, 2013), or to symptoms similar to decompression sickness that can lead to injury or death [CITE DE QUIROS]. For example, research on Blainville’s beaked whales *Mesoplodon densirostris* on a US Navy range in the Bahamas has shown that animals may stop foraging and/or move away from naval sonar sources (Joyce et al., 2019; Tyack et al., 2011).

Naval sonar can be broadcast from various platforms, including vessels, helicopters, buoys, submarines, and torpedoes (Harris et al., 2019; Navy, 2018). Most research has focused on the impacts of mid-frequency active sonar (MFAS) broadcast from naval vessels. Separately, researchers have shown that, in the absence of MFAS, beaked whales may alter their behavior in response to vessel noise (Aguilar Soto et al., 2006; Pirodda et al., 2012).

The US Navy is interested in quantifying the effects of sonar on beaked whales for the purpose of risk assessments and permitting associated with training activities [CITE EG US DON 2017]. There are different experimental and analytical ways of quantifying responses to sonar. Here, we focus on analyses of data from cabled hydrophone arrays.

For example, (CITE name (date) McCarthy et al., 2011) used data from the cabled hydrophone array at the US Navy’s Atlantic Undersea Test and Evaluation Center (AUTEK) in the Bahamas collected before, during, and after naval training exercises involving MFAS. The authors used separate generalized additive models (GAMs) for each period, and modelled the acoustic detection of groups of Blainville’s beaked whales (group vocal periods; GVPs) as a function of location on the range and time. They found that the number of GVPs was lower during the exercises than before or after.

Building on this work, (CITE name (date) Moretti et al., 2014) used a GAM to model the presence of acoustic detections of groups of Blainville’s beaked whales on the AUTEK range as a smooth function of MFAS received level. They then compared the expected probability of detecting animals when no sonar was present to the expected probability of detecting animals across sonar received levels to estimate the probability of disturbance. They found that the probability of detecting groups of Blainville’s beaked whales was reduced by 50% at 150 dB re 1 μ Pa, which they interpreted as a 50% probability of disturbance.

Our primary objective was to replicate the effort of Moretti et al. [CITE DATE] with the same species on a different US Navy training range in a different oceanic environment. Unlike AUTEK, which occurs in a deep isolated basin surrounded by steep slopes, the Pacific Missile

Range Facility (PMRF) range in Hawaii is located on the side of an ancient volcano, with a steep slope down to the deep ocean floor. Density of Blainville’s beaked whales at PMRF is lower and more variable than at AUTC, so we wanted to explicitly account for differences in underlying beaked whale presence across the range.

An additional objective was to isolate the effect of general training activity from the effect of MFAS, so that beaked whale response to MFAS could be quantified relative to pre-training baseline periods and to periods when general training activities were present on the range.

We used a spatially referenced dataset of Blainville’s beaked whale foraging dives recorded at the PMRF off the island of Kauai, Hawaii (Fig. 1). Acoustic detections of Blainville’s beaked whales were collected via a cabled hydrophone array at PMRF before and during training exercises involving MFAS broadcast from navy ships. Previous work in this region has shown that Blainville’s beaked whales are present year-round at this site, prefer slope habitats, and that acoustic detections decrease during multi-day training events involving MFAS (CITE FIX FOUR AUTHORS? Henderson, Martin, Manzano-Roth, and Matsuyama, 2016; Manzano-Roth, Henderson, Martin, Martin, and Matsuyama, 2016).

A series of three linked models were fitted. The first was fitted to data collected before the training exercises began in order to estimate the underlying spatial distribution of acoustic detections. The expected values from this model were used as offsets in the second model, which was fitted to data collected when training exercises were ongoing by no MFAS was broadcast. The expected values from this second model were then used as offsets in the third model, which was fitted to data collected when training exercises that did use MFAS were ongoing. Uncertainty was propagated through all models using posterior simulation. With this set of model results, we quantified the expected decrease in detection of GVPs across increasing sonar received levels relative to both the pre-training baseline period and the period when training activities were ongoing but no MFAS was present. [813]

2 Methods

2.1 Data Collection and Processing

2.1.1 Acoustic detection of beaked whales

The Pacific Missile Range Facility (PMRF) is an instrumented U.S. Naval range extending 70 km NW of the island of Kauai, Hawaii and encompassing 2,800 km². The range includes a cabled hydrophone array (Fig. 1) with hydrophones at depths ranging from approximately 650 m to 4,700 m. We used data collected before and during six Submarine Command Courses (SCCs) at PMRF. SCCs are training exercises that occur biannually in February and August and typically last 6-7 days. Acoustic recordings were made for a minimum of two days before each SCC as well as during the exercise. During data collection, hydrophones sampled at a rate of 96 kHz, with the high pass filter on each hydrophone set at either 50 Hz,

106 100 Hz, or 10 kHz. Up to 62 of the range hydrophones were recorded simultaneously by the
 107 Naval Information Warfare Center (NIWC).

108 A beaked whale detector from the Navy Acoustic Range WHale AnaLysis (NARWHAL)
 109 algorithm suite [CITE Martin et al. 2020] was run on the recordings. This detector first
 110 compares signal-to-noise (SNR) thresholds within the expected beaked whale click frequency
 111 range (16 - 44 kHz) versus the bandwidth outside the click in a running 16,384-pt fast Fourier
 112 transform (FFT) spectrogram. The detected clicks were then passed to a 64-pt FFT stage
 113 that measured power, bandwidth, slope, and duration characteristics to classify the clicks to
 114 species. This process was followed by an automated routine in Matlab (CITE Mathworks
 115 2019) to group detections of individual beaked whale echolocation clicks into GVPs [CITE
 116 Henderson et al. 2019]. If a group of whales was detected by more than one hydrophone, the
 117 GVP was assigned to the hydrophone that recorded the most clicks. The data were then
 118 aggregated to indicate the presence or absence of the start of a GVP for each hydrophone
 119 within each half-hour period.

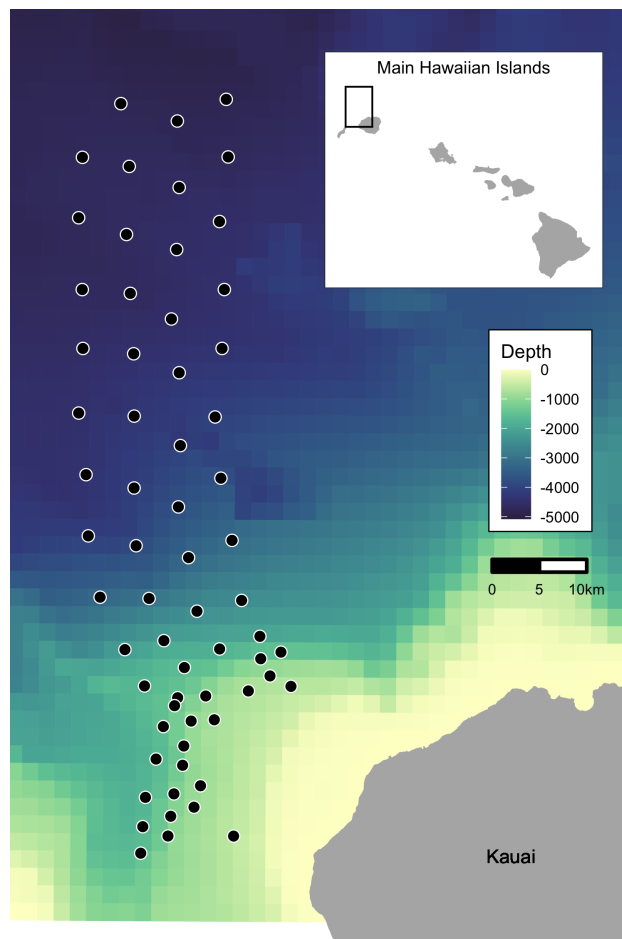


Figure 1: Map of approximate locations of hydrophones (black points) at the Pacific Missile Range Facility near the island of Kauai, Hawaii. Color scale indicates bathymetry. Inset map shows range location relative to the Main Hawaiian Islands.

2.1.2 Modelling received levels of hull-mounted mid-frequency active sonar

LIZ SEE Qs IN BRACKETS. ALSO I REARRANGED SOME SENTENCES SO PLS CHECK.

For security reasons, classified data regarding activity that occurred on the range during each SCC was passed from PMRF to one author with clearance (EEH). These data indicated the locations of the ships during the training periods and the start and stop times of each individual training event. However, no information was provided on the start and stop of sonar use; hence, periods of active sonar were determined from the range hydrophone recordings by running a sonar detector from the NARWHAL algorithm suite tuned to MFAS. The hydrophone recordings cannot be reliably used to determine received level when the received level exceeds XXX dB re. $1 \mu \text{ Pa}$ [LIZ Q: HOW DO WE SAY THIS? I WANT TO EXPLAIN WHY WE DIDN'T USE THE RECORDED LEVELS.]. Therefore, we used propagation modelling to estimate transmission loss and calculate the expected maximum received level of hull-mounted MFAS around the location of each hydrophone. First, the locations of all surface ships were noted for each half-hour period and the closest ship to each hydrophone was determined. MFAS propagation was modelled using the parabolic equation propagation model in the program Peregrine (OASIS; Heaney and Campbell, 2016). Transmission loss was estimated using a 200 Hz band around the center frequency of the sonar (3.5 kHz). The transmission loss was estimated along the radial from the ship to the hydrophone from a distance of 1 km before the hydrophone to 1 km past the hydrophone in 200 m increments and converted to received levels based on the source level of the sonar. [LIZ Q: WAS THE SOURCE LEVEL KNOWN OR ASSUMED? CAN WE SAY WHAT IT WAS?] The maximum modeled received level along that radial was determined for each hydrophone and half-hour period. However, if the distance between the ship and the hydrophone was less than the depth of the water column, the parabolic equation would overestimate transmission loss at that angle. In these cases, a simple sonar equation was used to estimate transmission loss instead. Transmission loss was estimated at depth since Blainville's beaked whales do not begin clicking until they have reached approximately 200-500 m depth of their foraging dive and spend most of their foraging dive at around 1,000-1,500 m [CITE Johnson et al. 2004, 2006, Madsen et al. 2013]. For hydrophones shallower than 1,000 m the received level was estimated at a point 20 m above the sea floor with a +/- 10 m buffer, while for hydrophones deeper than 1,000 m the received level was estimated at a depth of 1,000 m with a +/- 10 m buffer. This process resulted in an estimate of received level for each hydrophone and half-hour period. Uncertainty in the modeled received levels was not considered.

2.2 Spatial Modelling

We used a multi-stage generalized additive modelling (GAM; S. N. Wood, 2017) approach to control for the underlying spatial distribution of Blainville's beaked whales when modelling the effects of training activities and of MFAS. We first used a tessellation to determine the area effectively monitored by each hydrophone. Then, we used pre-activity data to create a spatial model of the probability of GVPs across the range prior to the onset of naval activity.

We used the predicted values from this model as an offset in a model created using data from when naval activity was present on the range, but MFAS was not. We then used the predicted values from this second model as an offset in a third model created using data when naval activity and MFAS were present on the range. Finally, we used posterior simulation to calculate confidence intervals and quantified the change in the probability of detecting GVPs when Naval activity was present and across received levels of MFAS.

2.3.1 Determining hydrophone effort

For security reasons, randomly jittered locations and depths of hydrophones at PMRF were used. We projected the coordinates of each hydrophone into Universal Transverse Mercator Zone 4. Because the beaked whale detection algorithm assigned GVPs to the hydrophone that recorded the most echolocation clicks, and because the spatial separation of the hydrophones was not uniform, effort was not the same for all hydrophones. This means that some hydrophones may have detected more GVPs because they were further away from other hydrophones, not because they were located in higher-density areas. We account for this by determining the area effectively monitored by each hydrophone. To do this, we used a Voronoi tessellation implemented in the R (R Core Team, 2018) package `deldir` (Turner, 2019) to define a tile for each hydrophone that contained all points on the range that were closest to that hydrophone. The area of each tile corresponded to the effective area monitored. We assumed that beaked whale groups occur within the tessellation tile of the hydrophone to which the GVP is assigned. For hydrophones on the outside of the range, i.e., not surrounded by other hydrophones, we used a cutoff radius of 6,500 m to bound the tessellation tile. This distance is based on the maximum detection distance of individual Blainville’s beaked whale clicks at a U.S. Naval range in the Bahamas (T. A. Marques, Thomas, Ward, DiMarzio, and Tyack, 2009). Different combinations of hydrophones were used during different SCCs, so separate tessellations were created for each SCC.

2.3.2 M1: Modelling the pre-activity probability of dive detection

We used data collected prior to SCCs, when no naval ships were present on the range and no other naval activity was known to occur, to model the spatial distribution of GVP detections across the range. Because of the way that GVPs were assigned to hydrophones (see Section 2.1.1) the data were not continuous in space. To account for this, we used a Markov random field (MRF) to model the spatial distribution of GVP detections. Markov random fields (Rue and Held, 2005) model correlation in space between discrete spatial units (henceforth, “tiles”). The correlation between two tiles is dictated by distance, as measured by the number of other tiles one needs pass through to travel between two tiles (“hops”); correlation is strongest between a tile and its direct neighbors (those tiles it shares a border with) and decreases with additional hops. This is appropriate for our data as we did not know where in each tile a given GVP occurred, but we assumed that it did occur in that tile.

We modelled the probability of a GVP at tile i as a Bernoulli trial: $\text{GVP}_i \sim \text{Bin}(1, \mu_{\text{M1},i})$. The linear predictor for the model was modelling on the logit scale:

$$\text{logit}(\mu_{\mathbf{M1},i}) = \beta_{\mathbf{M1},0} + f(\mathbf{MRF}_i) + f(\text{Depth}_i) + \log_e A_i, \quad (\mathbf{M1})$$

where $\beta_{\mathbf{M1},0}$ is an intercept, $f(\mathbf{MRF}_i)$ denotes the Markov random field used to smooth space, $f(\text{Depth}_i)$ is a smooth of depth (using a thin plate spline; Simon N Wood (2003)) and $\log_e A_i$ is an offset for the area (in km^2) of each tile, A_i . The offset term accounts changes in probabilities of GVP detection due to the differing area monitored by each hydrophone. Because the hydrophone tessellation changed between SCCs (as there were different sets of hydrophones recorded during each SSC), separate MRFs were used for each SCC, but a single smoothing parameter was estimated across all MRFs. This allows for different spatial smooths for each SCC, but constrains the smooths to have the same amount of wiggleness. The smooth of depth was shared across SCCs.

2.3.3 M2: Modelling the effect of Naval activity

For a few days prior to the onset of hull-mounted MFA sonar used during SCCs, other naval training activities occurred at PMRF. Various vessels were present on the range during this period and other noise sources, including torpedoes and submarines, may have been present. We used data collected when training activity was present on the range, but hull-mounted MFAS was not used, to model the effect of general naval activity on beaked whale GVPs. Initially, we tried to use low-frequency noise levels in the 10-999 Hz range measured on range hydrophones as a covariate in this model, but found that the measured noise levels were not consistent with known locations of naval training activities (see Appendix X for details).

We used the predicted baseline probability of a GVP detection from M1 as an offset to control for the underlying spatial distribution of GVPs. The model for the data when ships were present was intercept-only, with an offset derived from M1. We again modelled GVP presence at tile i as $\text{GVP}_i \sim \text{Bin}(1, \mu_{\mathbf{M2},i})$, with a linear predictor on the logit scale:

$$\text{logit}(\mu_{\mathbf{M2},i}) = \beta_{\mathbf{M2},0} + \log_e \xi_{\mathbf{M1},i}, \quad (\mathbf{M2})$$

where $\beta_{\mathbf{M2},0}$ is an intercept and $\xi_{\mathbf{M1},i}$ is the prediction (on the logit scale) for tile i using model M1, included as an offset term.

2.3.4 M3: Modelling the effect of hull-mounted MFA sonar

We used data collected when hull-mounted MFAS was present on the range to model the effect of sonar on beaked whales. The probability of a GVP when sonar was present was modeled as a function of the maximum received level (modeled at each hydrophone for each half-hour period; see section 2.2). We assumed that as the maximum received level increased, the probability of dives decreased and modeled this using a monotonically decreasing smooth so that the relationship held for all possible realizations of the smooth (Pya and Wood, 2015). To ensure that the model predictions were the same at a maximum received level of 0 dB

and when ships were not present, we did not include an intercept. GVP presence at tile i was modelled as a Bernoulli trial $\text{DivePresent}_i \sim \text{Bin}(1, \mu_{\text{M3},i})$ where the linear predictor on the logit scale was:

$$\text{logit}(\mu_{\text{M3},i}) = f(\text{MaxRL}_i) + \log_e \xi_{\text{M2},i}, \quad (\text{M3})$$

where $f(\text{MaxRL}_i)$ was modeled as a monotonic decreasing smooth, $\xi_{\text{M2},i}$ denotes the prediction (on the logit scale) for tile i when Naval training activities were present on the range using model M2.

2.3.5 Uncertainty propagation

We used posterior simulation (sometimes referred to as a parametric bootstrap; Wood, Li, Shaddick, and Augustin (2017)) to propagate uncertainty through M1, M2, and M3. This consisted of sampling from the posterior distribution of the parameters for each model in turn, calculating predictions using these parameters and then refitting the subsequent model with updated offsets. Following this procedure through from M1 to M2 to M3 incorporated uncertainty from each model in the final results.

The prediction grid contained all possible combinations of covariates within the realized covariate space; i.e., each hydrophone for each SCC with associated location, hydrophone depth, and area of the tessellation tile, presence/absence of naval activity, and, if naval activity was present, then either sonar absence or sonar received level between 35 and 190 dB in intervals of 5 dB.

Based on the resulting final posterior distribution of results (for model M3) we can use appropriate quantiles to obtain average predictions and intervals. Mathematical details of the procedure are given in Appendix B.

2.3.6 Quantifying the change in probability of GVPs

Finally, we calculated the expected change in $\mathbb{P}(\text{GVP})$ relative to either the distribution of GVPs when no general Naval training activity was present and no MFA sonar was present ($\Delta_{\text{M3}':\text{M1}'}$), or relative to the distribution of GVPs when general Naval training activity was present but no MFA sonar was present ($\Delta_{\text{M3}':\text{M2}'}$).

Using the N_b posterior samples from the model, we calculated the expected $\mathbb{P}(\text{GVP})$ under each set of covariates as

$$\mathbb{P}(\text{GVP}) = \text{logit}^{-1}(\mu_{\mathbf{M}'}), \quad (1)$$

for each $\text{M1}'$, $\text{M2}'$, and $\text{M3}'$. Then, we calculated the change in $\mathbb{P}(\text{GVP})$ for each set of covariates $\text{M3}'$ and $\text{M1}'$ ($\Delta_{\text{M3}':\text{M1}'}$) and between $\text{M3}'$ and $\text{M2}'$ ($\Delta_{\text{M3}':\text{M2}'}$) for each realization of the posterior simulation.

$$\Delta_{M3':M1'} = \frac{\mathbb{P}(\text{GVP})_{M3'} - \mathbb{P}(\text{GVP})_{M1'}}{\mathbb{P}(\text{GVP})_{M1'}} \quad (2)$$

$$\Delta_{M3':M2'} = \frac{\mathbb{P}(\text{GVP})_{M3'} - \mathbb{P}(\text{GVP})_{M2'}}{\mathbb{P}(\text{GVP})_{M2'}} \quad (3)$$

For each received level we calculated the 2.5th, 50th, and 97.5th quantiles of $\Delta_{M3':M1'}$ and $\Delta_{M3':M2'}$ to create 95% CIs of change in $\mathbb{P}(\text{GVP})$ across possible received levels. We consider that the probability of disturbance is equal to 1 wherever the 95% CI does not include 0, and 0 otherwise.

2.3.7 Implementation

NOTE: this section may be moved to an appendix if we have space constraints, but I do prefer to acknowledge the contributions of software developers in the main text if possible.

Statistical analyses presented in this manuscript were conducted in R (v. 3.5.2; R Core Team, 2018). Data import and manipulation was accomplished using the packages `dplyr` (Wickham, François, Henry, and Müller, 2020), `lubridate` (Grolemund and Wickham, 2011), `readr` (Wickham, Hester, and François, 2018), `readxl` (Wickham and Bryan, 2019), and `tidyr` (Wickham and Henry, 2020). `mgcv` (S. N. Wood, 2017) and `scam` (Pya and Wood, 2015) were used to formulate and fit models. Map creation was facilitated by the `fields` (Douglas Nychka, Reinhard Furrer, John Paige, and Stephan Sain, 2017), `ggsn`, `marmap`, `rgdal`, `sf`, and `sp` packages. All graphics were produced using `ggplot2`, with color palettes from the `viridis` and `cmocean` packages. The manuscript was written in `rmarkdown` (Xie, Allaire, and Grolemund, 2018).

3 Results

3.1 Results of Data Collection and Processing

Data were collected before and during six SCCs: two each in 2013, 2014, and 2017 (Table 1). The number of hydrophones for which recordings were available varied from 49 to 61. A total of 190928 30-min observations were made.

Table 1: No. of hydrophones used and number of observations made (no. 30-min periods) for each SCC before the exercise began, when Naval activity was present, and when Naval activity and MFA sonar were present.

SCC	HPs	Pre-Activity	Nav. Activity	MFA Sonar
Feb13	61	114	193	124

SCC	HPs	Pre-Activity	Nav. Activity	MFA Sonar
Aug13	61	209	115	97
Feb14	60	513	111	129
Aug14	61	263	120	128
Feb17	59	450	97	108
Aug17	49	270	106	113

The exact timing of activities during these exercises varied (Fig. 2). For most SCCs, pre-activity data were available immediately preceding the onset of Naval training activity; however, in February 2013 the only available pre-activity data were collected almost a month prior to the onset of Naval training activity. In some SCCs, weekends or other breaks in training resulted in a break in training activity on the range during the days preceding MFA sonar use. MFA sonar was used for 3-4 days during each training event.

Across all SCCs, hydrophones, and conditions, a total of 2312 GVPs were identified. The average $\mathbb{P}(\text{GVP})$ in the dataset was therefore 1%. The spatial distribution of GVPs differed during the pre-activity phases of SCCs (Fig. SX; top panel).

Modelled maximum received levels ranged from 38 to 186 dB re. $1 \mu \text{ Pa}$, with a median value when MFA sonar was present of 147 dB re. $1 \mu \text{ Pa}$. The intensity and spatial distribution of MFA received levels varied across the range and across SCCs (Fig. SX).

Based on the observed data, $\mathbb{P}(\text{GVP})$ changed by -57% when Naval activity was present compared to when Naval activity was absent, by -47% when Naval activity and MFA sonar were present compared to when only ships were present, and by -77% when Naval activity and MFA sonar were present compared to when neither ships nor sonar were present (Fig. SX).

3.2 Results of spatial modelling

We created separate tessellations for each SCC (Fig. SX). In August 2017, data were available from fewer hydrophones, and so in some cases the tessellated tiles, with bounding radius of 6,500 m, did not completely cover the range. Hydrophone depths varied from 648 to 4716 m.

M1 fitted a spatial model of $\mathbb{P}(\text{GVP})$ to data collected prior to the onset of Naval training activity. This model used a MRF smooth to account for the spatial structure of the range and a spline on depth, with an offset for the log of the area effectively monitored by each hydrophone. Both the MRF and spline on depth were significant (p-value $< 2\text{E-}16$). The model explained 14.1% of deviance in the dataset, and visual inspection of observed versus predicted values indicated a good fit to the data (Fig. SX). The model **M1** predicted highest $\mathbb{P}(\text{GVP})$ at hydrophone depths between 1,500 and 2,000 m (Fig. SX).

M2 used the predicted values from **M1** as an offset and fitted a model of to data when Naval activity was ongoing, as indicated by the presence of Naval ships on the range. This model was intercept-only, and $\mathbb{P}(\text{GVP})$ when Naval training was ongoing was significantly different

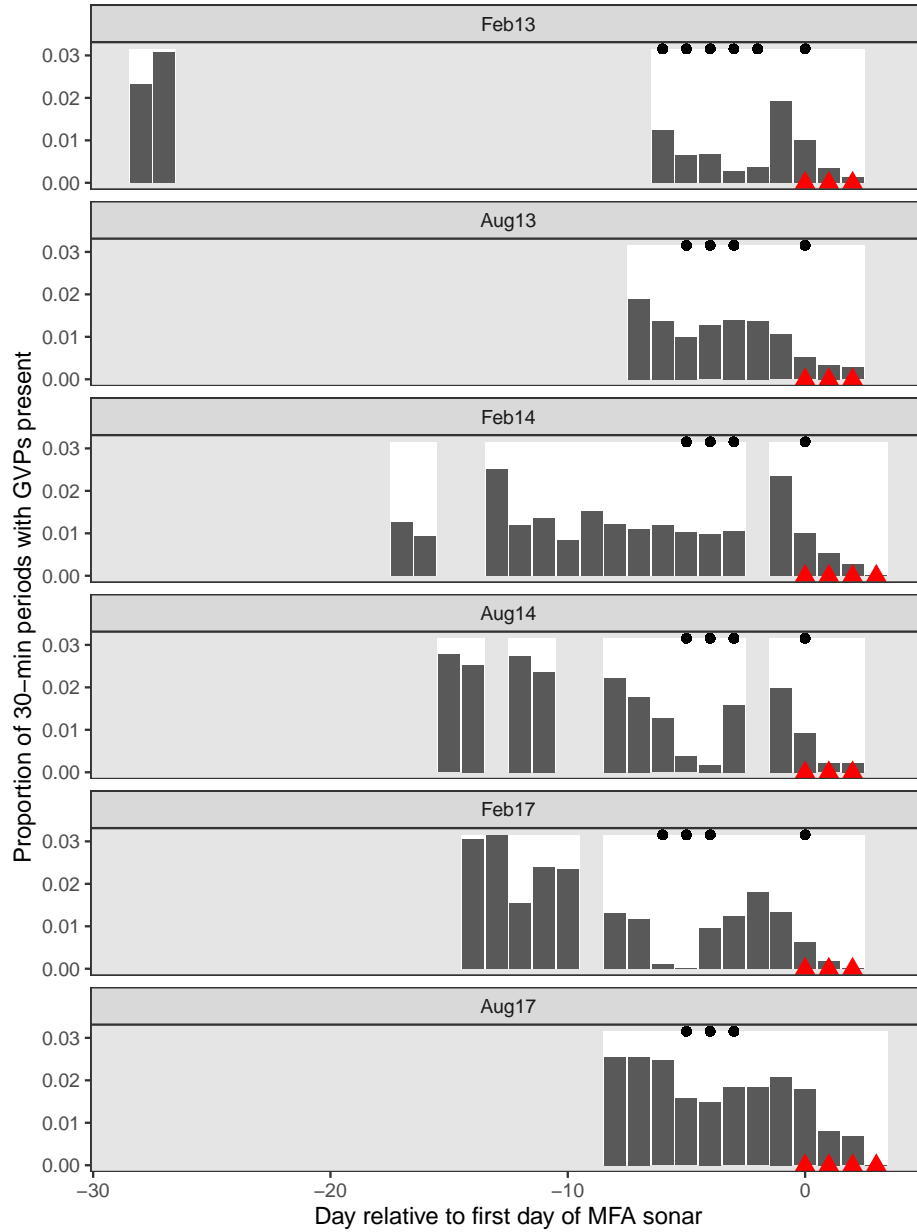


Figure 2: Timeseries of six recorded Naval training activities at PMRF. The timeseries are aligned relative to the first day that MFA sonar (red triangles) was used in each exercise (x-axis). Days with white background indicate days for which recordings and data were available. Gray bars indicate the proportion of 30-min periods on each day, across all hydrophones, when GVPs were detected. Black dots indicate days when Naval activity was present on the range.

from the baseline period (p-value < 2E-16). The expected $\mathbb{P}(\text{GVP})$ decreased by a median of 64% (95% CI 59% - 68%) when Naval training activity was present compared to when it was absent.

M3 used the predicted values from M2 as an offset and fitted a model to data when Naval activity and MFA sonar were present. This model used a monotonically decreasing spline on modelled MFA sonar received level (Fig. SX) and did not include an intercept term. The smooth on MFA sonar received level had significant explanatory power (p-value = 6.74E-10) and the model explained 12.4% of deviance in the data.

For MFA sonar received levels above 100 dB, change in $\mathbb{P}(\text{GVP})$ was calculated relative to the pre-activity baseline period ($\Delta_{M3':M1'}$) and to the period when Naval activity was present on the range ($\Delta_{M3':M2'}$; Fig. 4 & Fig. 5). At a received level of 150 dB, $\Delta_{M3':M1'}$ was -92% (95% CI -100% - -87%) and $\Delta_{M3':M2'}$ was -78% (95% CI -100% - -62%).

4 Discussion

We used a series of three linked models to quantify the response of Blainville's beaked whales to naval training exercises involving MFAS: the first model was fitted to pre-exercise baseline data, the second was fitted to data collected when naval training exercises were ongoing but no MFAS was present, and the third model was fitted to data collected during naval training exercises that used MFAS. We found that the probability of acoustic detections of Blainville's beaked whales decreased when both naval training exercises and naval training exercises using MFAS were present.

In comparison to the risk function developed by Moretti et al. 2014 for Blainville's beaked whales at AUTECH, our risk function predicts a more intense response to naval sonar. This may be because Moretti et al. were not able to explicitly account for the effects of naval training activities that did not include MFAS. Their baseline period consisted of 19 hours of data before the onset of MFAS; as at PMRF, it is likely that training activities during this period included sound sources other than MFAS. Therefore, their risk function is probably more analogous to our expected change in the probability of a detection when MFAS is present relative to when naval training activity was present (Fig. 4).

Additionally, we used spatially-explicit methods that account for the spatial confounding of animal distribution and naval training activity. The data used in this study are from an undesigned experiment, where the spatial intensity of the treatments (naval activity and MFAS) were not applied randomly with respect to either the study area or Blainville's beaked whale presence. We did not want the spatial distribution of training exercises and MFAS to influence our understanding of the baseline spatial distribution of Blainville's beaked whales. Due to the spatial confounding of animal distribution and naval training activity at PMRF, fitting a single model to all of the data would lead to underestimates of the impact of sonar, since changes in distribution due to MFAS could be explained as spatial changes by the MRF (Appendix X). Our three-stage modelling approach addresses this issue while propagating uncertainty between the models.

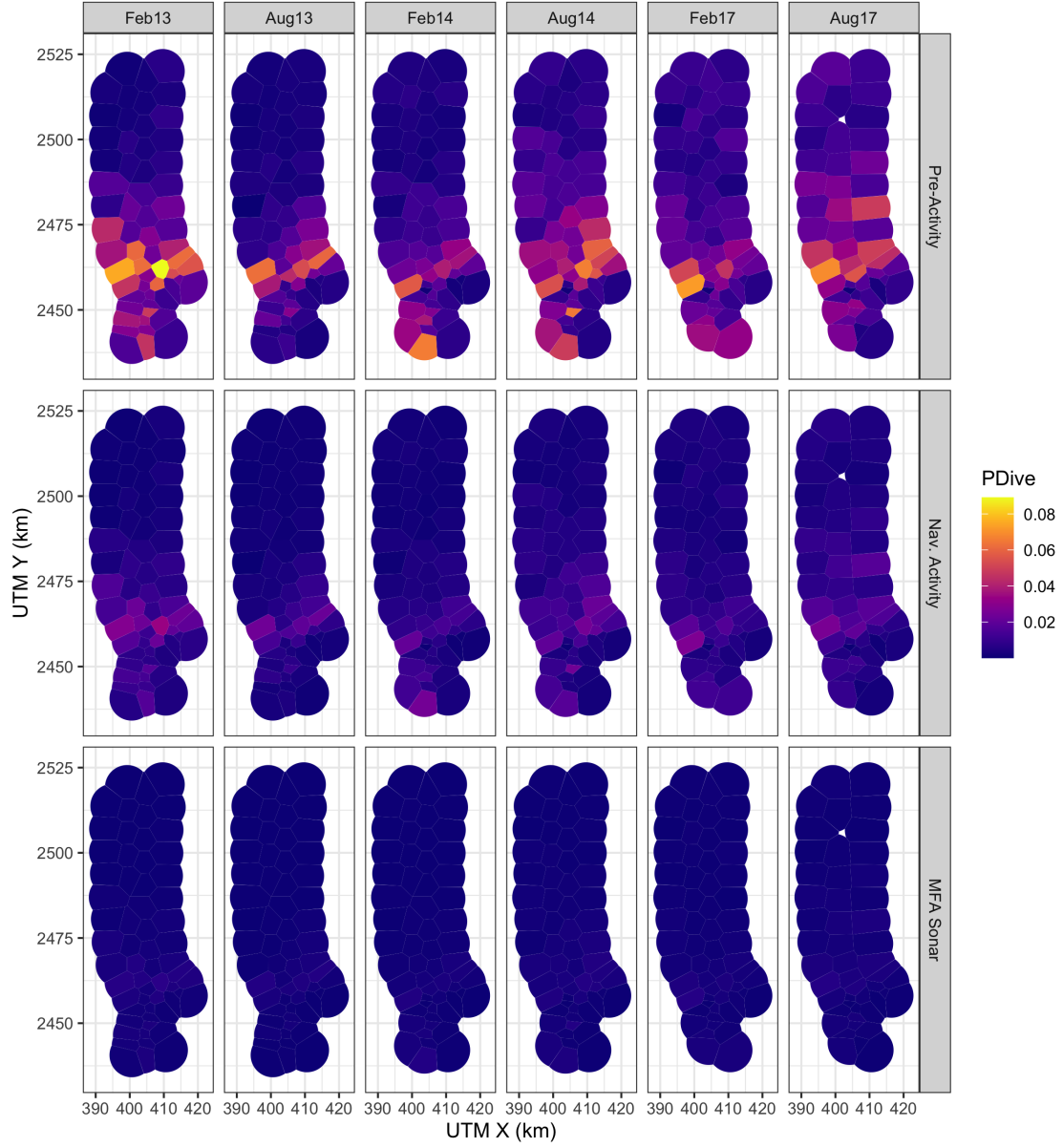


Figure 3: Map of expected probability of diving (color scale) at each hydrophone during each SCC (columns) prior to the onset of Naval training activity, during Naval training activity when no MFA sonar was present, and during Naval training activity when MFA sonar was present at a level of 150 dB re. 1 uPa rms (rows).

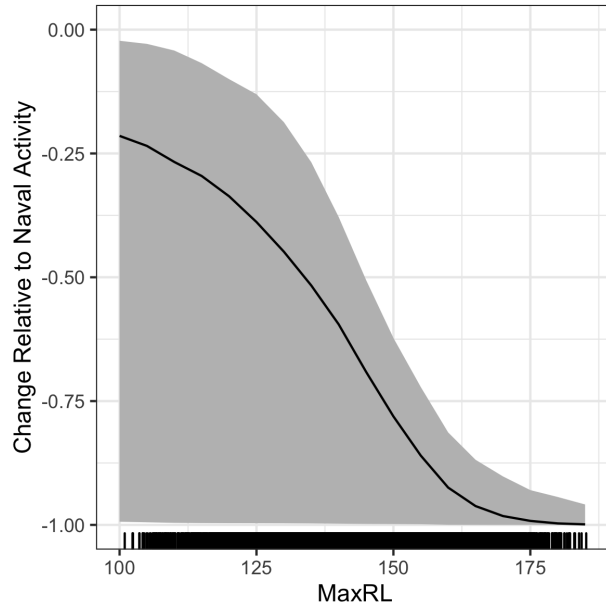


Figure 4: Median (black line) and 95% CIs (gray shading) expected change in the probability of detecting a group vocal period (y-axis) with increasing MFA sonar received level (x-axis) relative to when Naval training activity but no MFA sonar is present on the range.

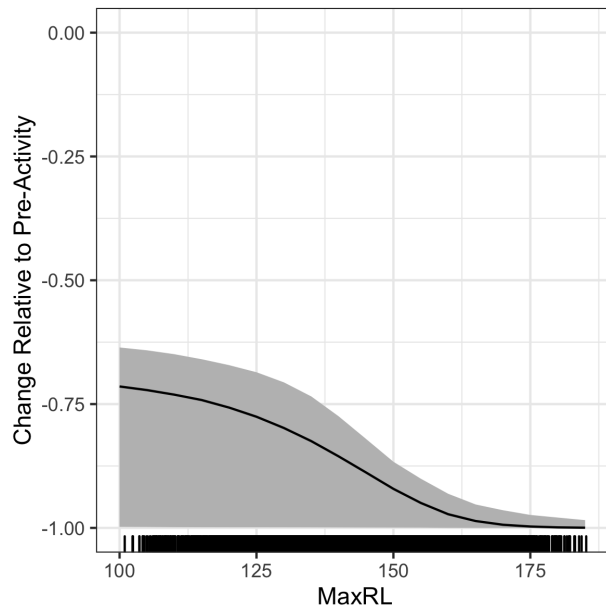


Figure 5: Median (black line) and 95% CIs (gray shading) expected change in the probability of detecting a group vocal period (y-axis) with increasing MFA sonar received level (x-axis) relative to when neither Naval training activity nor MFA sonar is present on the range.

The analytical approach outlined in this article could be applied to other species, regions, and types of disturbance where experimental design is not possible. The use of Markov random fields for the spatial term is useful for cases where exact distance data is not available, avoiding the use of continuous smoothers when true location data is not available. Shape-constrained smoothing is also well-suited to the kind of data we model here – ensuring that values can only stay constant or decrease over time (or any other covariate). Finally, the use of a multi-stage posterior sampling scheme extends to any situation where multiple models are fitted and the results of one part feed into another. Simulation-based approaches such as these bypass the need to derive (often complex) expressions (or shortcut them by assuming independence). We provide example code as Supplementary Material so that other researchers may apply and/or expand on these methods.

- Discuss dose-response and $p(\text{disturbance})$ in context of (Tyack and Thomas, 2019)
- plans to investigate what aspects of general training activity is eliciting a response.

From Liz: Then we can discuss the fact that environment/habitat (e.g. deep basin with shallow slopes all around vs deep open ocean) doesn't seem to play much of a role in Blainville's response, and the response seems to be more of an intrinsic characteristic. Also can mention here the same effort at SCORE with Cuvier's – in light of these results we expect similar results there even though different species but similar habitat to AUTECH. [389]

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Appendix A: Supplementary Tables and Figures

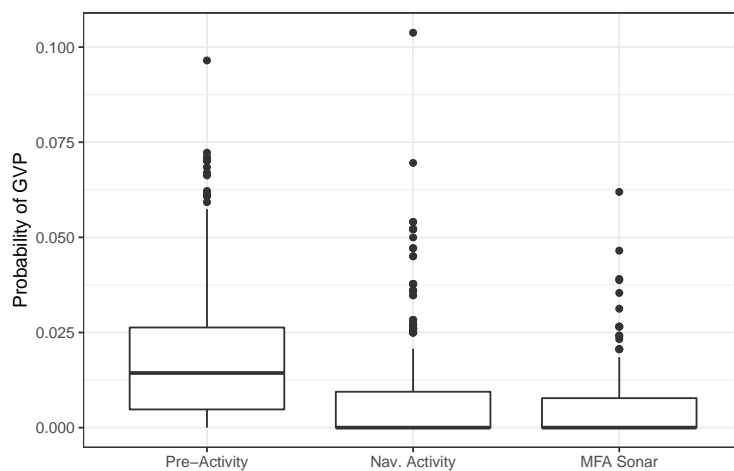
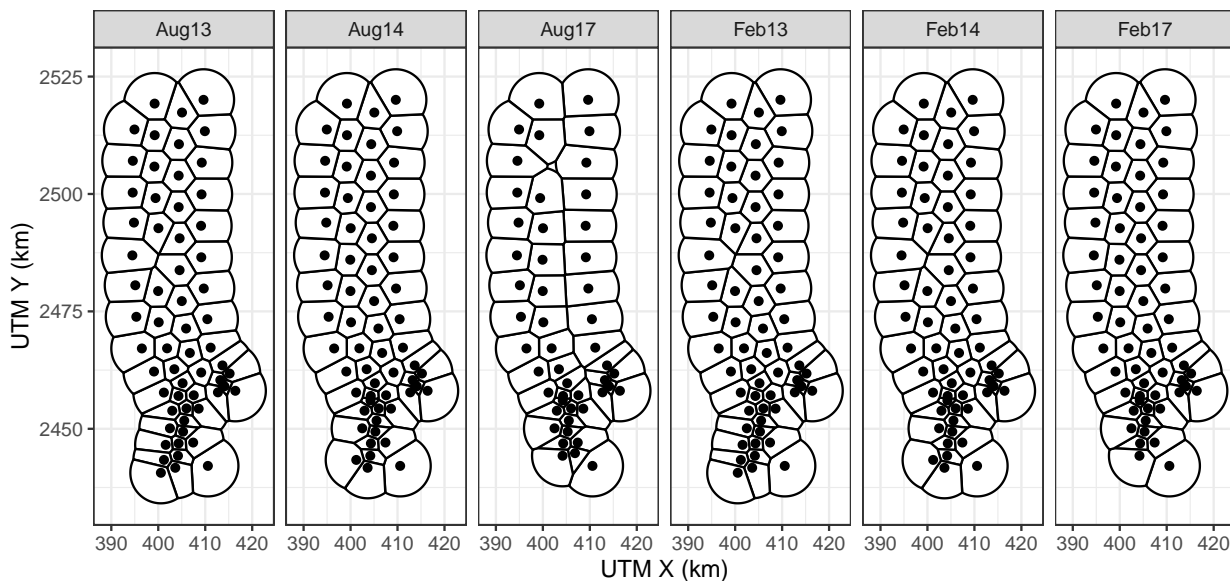


Figure 6: Boxplot of observed probability of diving across all hydrophones and SCCs before, when Naval activity was present, and when MFA sonar was present.

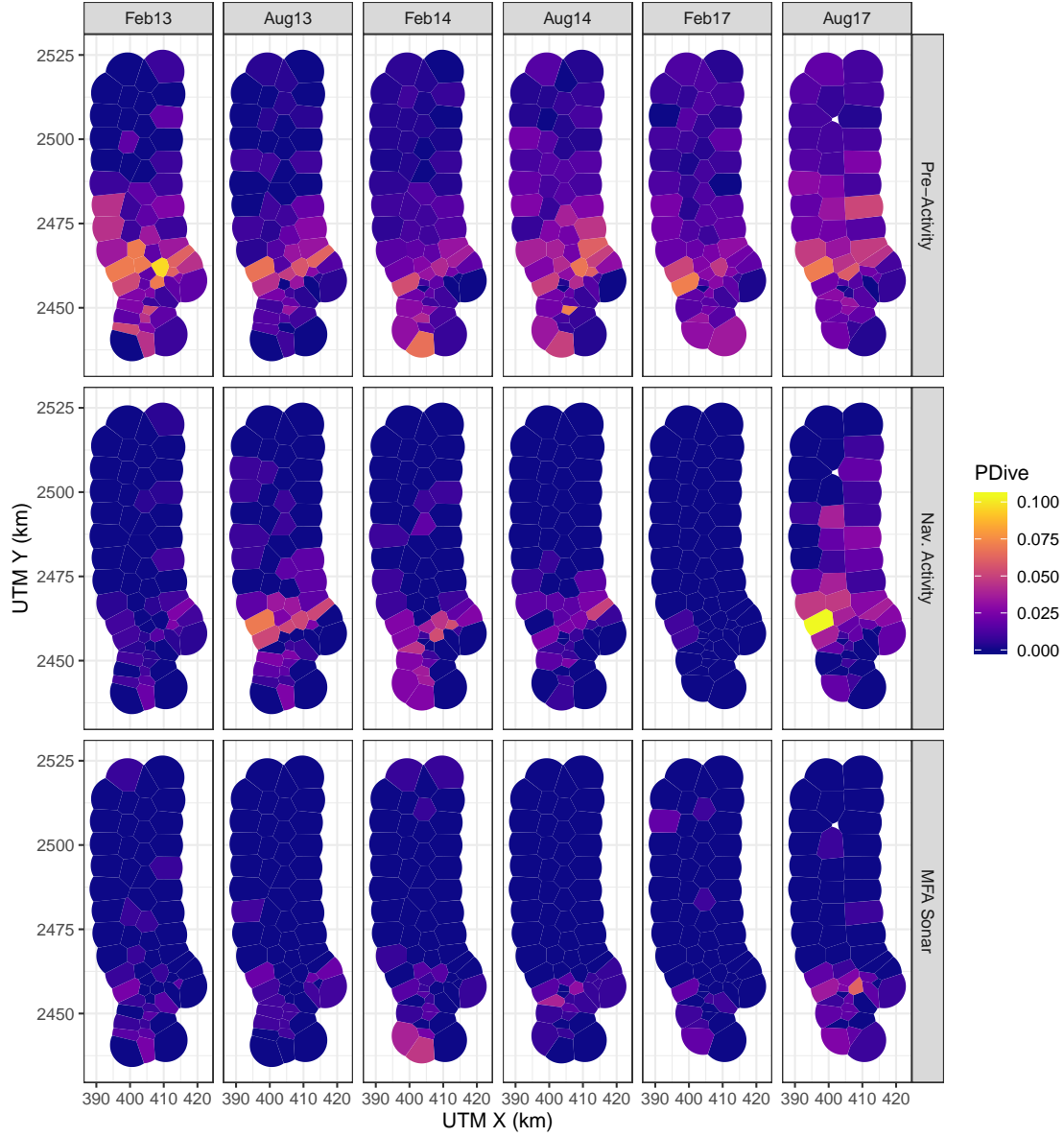


Figure 7: Map of observed probability of diving at each hydrophone before, during Phase A, and during Phase B of each SCC. Note that values of PDive are not corrected for effort (size of the hydrophone tile).

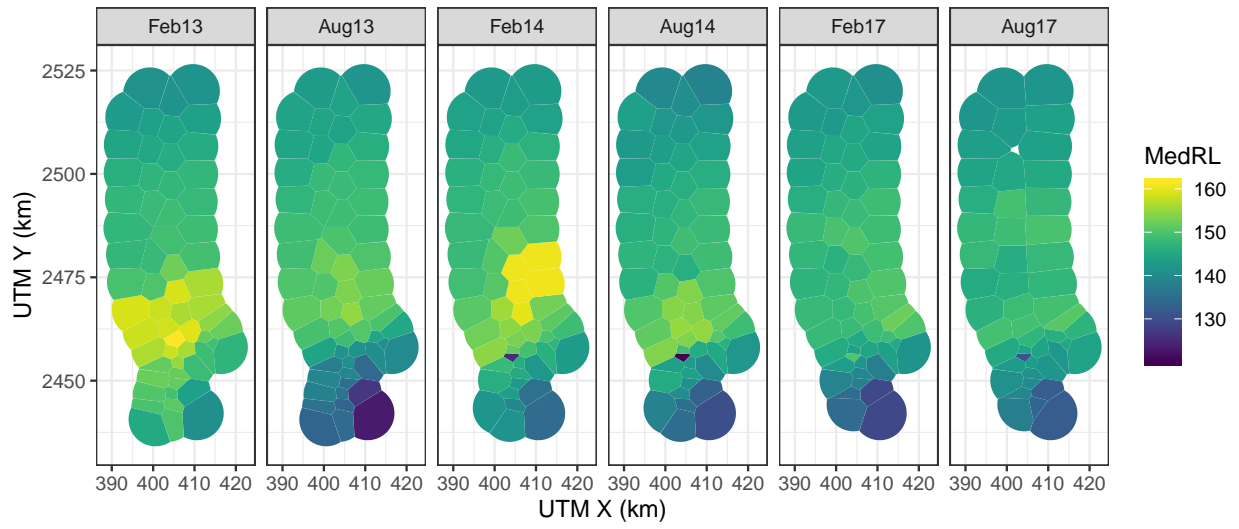


Figure 8: Median received level when MFA sonar was present (color scale) for all hydrophones and SCCs.

Appendix B: Uncertainty estimation details

We used posterior simulation to propagate uncertainty through M1, M2, and M3. Each model was fitted via restricted maximum likelihood (REML), so the results are empirical Bayes estimates. In this case we can generate samples from the (multivariate normal) posterior of the model parameters. After generating a sample, $\beta^* \sim \text{MVN}(\hat{\beta}, \mathbf{V}_\beta)$, we can use the matrix that maps the model parameters to the predictions on the linear predictor scale (often referred to as the \mathbf{L}_p matrix or \mathbf{X}_p matrix; Wood et al. (2017); section 7.2.6), along with the inverse link function to generate predictions for each posterior sample. Here the β for each model includes the coefficients for the smooth terms in the model and fixed effects (e.g., intercept) if present. Predictions, μ^* , can be written as:

$$\mu^* = g^{-1}(\eta^*) = g^{-1}(\mathbf{X}_p \beta^* + \xi),$$

where g is the link function, η^* is the linear predictor and ξ is any offset used by this prediction. By sampling from the posterior of $\hat{\beta}$, and then taking the variance of the resulting predictions we can obtain variance estimates (Wood et al. (2017); section 7.2.6). The prediction grid contained all possible combinations of covariates within the realized covariate space; i.e., each hydrophone for each SCC with associated location, hydrophone depth, and area of the tessellation tile, presence/absence of Naval activity, and, if Naval activity was present, then either sonar absence or sonar received level between 35 and 190 dB in intervals of 5 dB. This procedure was repeated for each model, with refitting to updated offsets from the previous model.

An algorithm for calculating the variance from our multi-stage approach is as follows. First define N_b as the number of samples to make, let $\mathbf{X}_{p,Mj}$ for $j = 1, 2, 3$ be the \mathbf{L}_p matrix that maps coefficients to the predictions for model Mj . For N_b times:

1. Draw a sample from the posterior of M1: $\tilde{\beta}_{M1} \sim \text{MVN}(\hat{\beta}_{M1}, \mathbf{V}_{M1})$.
2. Calculate a new offset for M2, $\tilde{\xi}_{M1} = \mathbf{X}_{p,M1} \tilde{\beta}_{M1} + \log_e \mathbf{A}$.
3. Refit M2 with $\tilde{\xi}_{M1}$ as the offset, to obtain M2'.
4. Draw a sample from the posterior of M2': $\tilde{\beta}_{M2'} \sim \text{MVN}(\hat{\beta}_{M2'}, \mathbf{V}_{M2'})$
5. Calculate a new offset for M3, $\tilde{\xi}_{M2} = \mathbf{X}_{p,M2'} \tilde{\beta}_{M2'} + \tilde{\xi}_{M1}$ (predictions for the sonar data locations for M2').
6. Refit M3 with offset $\tilde{\xi}_{M2}$ to obtain M3'.
7. Predict $\mu_{M1'}$, $\mu_{M2'}$, and $\mu_{M3'}$ over prediction grid and store them.

We can then calculate summary statistics (means and variances) of the N_b values of $\mu_{M1'}$, $\mu_{M2'}$, and $\mu_{M3'}$ we have generated. The empirical variance of the N_b values of $\mu_{M3'}$ will give the uncertainty, incorporating components from all three models. We can take appropriate pointwise quantiles to form confidence bands for the functional relationships between sonar received level and estimated probability of detecting GVPs.