MATH 392 Problem Set 5

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6.39

 $\{\hat{\beta}_n\}$ is consistent if and only if $\lim_{n\to\infty} P(|\hat{\beta}_n - \beta| < \varepsilon) = 1 \ \forall \ \varepsilon > 0.$

It has been shown before that \forall n ϵ N, $E[\hat{\beta}_n] = \frac{n}{n+1}\beta$. Thus,

$$lim_{n\to\infty}P(|\hat{\beta}_n - \beta| < \varepsilon) = 1 \Leftrightarrow lim_{n\to\infty}P(|E[\hat{\beta}_n] - \beta| < \varepsilon) = 1$$
$$\Leftrightarrow lim_{n\to\infty}|E[\hat{\beta}_n] - \beta| < \varepsilon$$
$$\Leftrightarrow lim_{n\to\infty}|\frac{n}{n+1}\beta - \beta| < \varepsilon$$
$$\Leftrightarrow lim_{n\to\infty}|(\frac{n}{n+1} - 1)\beta| < \varepsilon$$

It suffices to show that $\lim_{n\to\infty}\left|\frac{n}{n+1}-1\right|<\varepsilon$:

Let $\varepsilon > 0$. Then $\exists m \in \mathbb{N}$ st $\forall n \geq m$,

$$\frac{1}{n}<\varepsilon\Rightarrow\frac{1}{n+1}<\varepsilon\Rightarrow|\frac{-1}{n+1}|<\varepsilon\Rightarrow|\frac{n-n-1}{n+1}|<\varepsilon\Rightarrow|\frac{n}{n+1}-\frac{n+1}{n+1}|<\varepsilon\Rightarrow|\frac{n}{n+1}-1|<\varepsilon\Rightarrow|\frac{n}{n+1}-1|<\varepsilon\Rightarrow|\frac{n}{n+1}-1|<\varepsilon\Rightarrow|\frac{n}{n+1}-1|<\varepsilon\Rightarrow|\frac{n}{n+1}-1|<\varepsilon\Rightarrow|\frac{n}{n+1}-1|<\varepsilon\Rightarrow|\frac{n}{n+1}-1|<\varepsilon\Rightarrow|\frac{n}{n+1}-1|<\varepsilon\Rightarrow|\frac{n}{n+1}-1|<\varepsilon\Rightarrow|\frac{n}{n+1}-1|<\varepsilon\Rightarrow|\frac{n}{n+1}-1|<\varepsilon\Rightarrow|\frac{n}{n+1}-1|<\varepsilon\Rightarrow|\frac{n}{n+1}-1|<\varepsilon\Rightarrow|\frac{n}{n+1}-1|<\varepsilon\Rightarrow|\frac{n}{n+1}-1|<\varepsilon\Rightarrow|\frac{n}{n+1}-1|<\varepsilon\Rightarrow|\frac{n}{n+1}-1|<\varepsilon\Rightarrow|\frac{n}{n+1}-1|<\varepsilon\Rightarrow|\frac{n}{n+1}-1|<\varepsilon\Rightarrow|\frac{n}{n+1}-1|<\varepsilon\Rightarrow|\frac{n}{n+1}-1|<\varepsilon\Rightarrow|\frac{n}{n+1}-1|<\varepsilon\Rightarrow|\frac{n}{n+1}-1|<\varepsilon\Rightarrow|\frac{n}{n+1}-1|<\varepsilon\Rightarrow|\frac{n}{n+1}-1|<\varepsilon\Rightarrow|\frac{n}{n+1}-1|<\varepsilon\Rightarrow|\frac{n}{n+1}-1|<\varepsilon\Rightarrow|\frac{n}{n+1}-1|<\varepsilon\Rightarrow|\frac{n}{n+1}-1|<\varepsilon\Rightarrow|\frac{n}{n+1}-1|<\varepsilon\Rightarrow|\frac{n}{n+1}-1|<\varepsilon\Rightarrow|\frac{n}{n+1}-1|<\varepsilon\Rightarrow|\frac{n}{n+1}-1|<\varepsilon\Rightarrow|\frac{n}{n+1}-1|<\varepsilon\Rightarrow|\frac{n}{n+1}-1|<\varepsilon\Rightarrow|\frac{n}{n+1}-1|<\varepsilon\Rightarrow|\frac{n}{n+1}-1|<\varepsilon\Rightarrow|\frac{n}{n+1}-1|<\varepsilon\Rightarrow|\frac{n}{n+1}-1|<\varepsilon\Rightarrow|\frac{n}{n+1}-1|<\varepsilon\Rightarrow|\frac{n}{n+1}-1|<\varepsilon\Rightarrow|\frac{n}{n+1}-1|<\varepsilon\Rightarrow|\frac{n}{n+1}-1|<\varepsilon\Rightarrow|\frac{n}{n+1}-1|<\varepsilon\Rightarrow|\frac{n}{n+1}-1|<\varepsilon\Rightarrow|\frac{n}{n+1}-1|<\varepsilon\Rightarrow|\frac{n}{n+1}-1|<\varepsilon\Rightarrow|\frac{n}{n+1}-1|<\varepsilon\Rightarrow|\frac{n}{n+1}-1|<\varepsilon\Rightarrow|\frac{n}{n+1}-1|<\varepsilon\Rightarrow|\frac{n}{n+1}-1|<\varepsilon\Rightarrow|\frac{n}{n+1}-1|<\varepsilon\Rightarrow|\frac{n}{n+1}-1|<\varepsilon\Rightarrow|\frac{n}{n+1}-1|<\varepsilon\Rightarrow|\frac{n}{n+1}-1|<\varepsilon\Rightarrow|\frac{n}{n+1}-1|<\varepsilon\Rightarrow|\frac{n}{n+1}-1|<\varepsilon\Rightarrow|\frac{n}{n+1}-1|<\varepsilon\Rightarrow|\frac{n}{n+1}-1|<\varepsilon\Rightarrow|\frac{n}{n+1}-1|<\varepsilon\Rightarrow|\frac{n}{n+1}-1|<\varepsilon\Rightarrow|\frac{n}{n+1}-1|<\varepsilon\Rightarrow|\frac{n}{n+1}-1|<\varepsilon\Rightarrow|\frac{n}{n+1}-1|<\varepsilon\Rightarrow|\frac{n}{n+1}-1|<\varepsilon\Rightarrow|\frac{n}{n+1}-1|<\varepsilon\Rightarrow|\frac{n}{n+1}-1|<\varepsilon\Rightarrow|\frac{n}{n+1}-1|<\varepsilon\Rightarrow|\frac{n}{n+1}-1|<\varepsilon\Rightarrow|\frac{n}{n+1}-1|<\varepsilon\Rightarrow|\frac{n}{n+1}-1|<\varepsilon\Rightarrow|\frac{n}{n+1}-1|<\varepsilon\Rightarrow|\frac{n}{n+1}-1|<\varepsilon\Rightarrow|\frac{n}{n+1}-1|<\varepsilon\Rightarrow|\frac{n}{n+1}-1|<\varepsilon\Rightarrow|\frac{n}{n+1}-1|<\varepsilon\Rightarrow|\frac{n}{n+1}-1|<\varepsilon\Rightarrow|\frac{n}{n+1}-1|<\varepsilon\Rightarrow|\frac{n}{n+1}-1|<\varepsilon\Rightarrow|\frac{n}{n+1}-1|<\varepsilon\Rightarrow|\frac{n}{n+1}-1|<\varepsilon\Rightarrow|\frac{n}{n+1}-1|<\varepsilon\Rightarrow|\frac{n}{n+1}-1|<\varepsilon\Rightarrow|\frac{n}{n+1}-1|<\varepsilon\Rightarrow|\frac{n}{n+1}-1|<\varepsilon\Rightarrow|\frac{n}{n+1}-1|<\varepsilon\Rightarrow|\frac{n}{n+1}-1|<\varepsilon\Rightarrow|\frac{n}{n+1}-1|<\varepsilon\Rightarrow|\frac{n}{n+1}-1|<\varepsilon\Rightarrow|\frac{n}{n+1}-1|<\varepsilon\Rightarrow|\frac{n}{n+1}-1|<\varepsilon\Rightarrow|\frac{n}{n+1}-1|<\varepsilon\Rightarrow|\frac{n}{n+1}-1|<\varepsilon\Rightarrow|\frac{n}{n+1}-1|<\varepsilon\Rightarrow|\frac{n}{n+1}-1|<\varepsilon\Rightarrow|\frac{n}{n+1}-1|<\varepsilon\Rightarrow|\frac{n}{n+1}-1|<\varepsilon\Rightarrow|\frac{n}{n+1}-1|<\varepsilon\Rightarrow|\frac{n}{n+1}-1|<\varepsilon\Rightarrow|\frac{n}{n+1}-1|<\varepsilon\Rightarrow|\frac{n}{n+1}-1|<\varepsilon\Rightarrow|\frac{n}{n+1}-1|<\varepsilon\Rightarrow|\frac{n}{n+1}-1|<\varepsilon\Rightarrow|\frac{n}{n+1}-1|<\varepsilon\Rightarrow|\frac{n}{n+1}-1|<\varepsilon\Rightarrow|\frac{n}{n+1}-1|<\varepsilon\Rightarrow|\frac{n}{n+1}-1|<\varepsilon\Rightarrow|\frac{n}{n+1}-1|<\varepsilon\Rightarrow|\frac{n}{n+1}-1|<\varepsilon\Rightarrow|\frac{n}{n+1}-1|<\varepsilon\Rightarrow|\frac{n}{n+1}-1|<\varepsilon\Rightarrow|\frac{n}{n+1}-1|<\varepsilon\Rightarrow|\frac{n}{n+1}-1|<\varepsilon\Rightarrow|\frac{n}{n+1}-1|<\varepsilon\Rightarrow|\frac{n}{n+1}-1|<\varepsilon\Rightarrow|\frac{n}{n+1}-1|<\varepsilon\Rightarrow|\frac{n}{n+1}-1|<\varepsilon\Rightarrow|\frac{n}{n+1}-1|<\varepsilon\Rightarrow|\frac{n}{n+1}-1|<\varepsilon\Rightarrow|\frac{n}{n+1}-1|<\varepsilon\Rightarrow|\frac{n}{n+1}-1|<\varepsilon\Rightarrow|\frac{n}{n+1}-1|<\varepsilon\Rightarrow|\frac{n}{n+1}-$$

Thus $\lim_{n\to\infty} \left|\frac{n}{n+1}-1\right| < \varepsilon$, so $\lim_{n\to\infty} P(|\hat{\beta}_n-\beta| < \varepsilon) = 1$. Thus $\{\hat{\beta}_n\}$ is consistent.

6.40

 $\{\hat{\sigma}_n^2\} \text{ is consistent if and only if } \lim_{n\to\infty} P(|\hat{\sigma}_n^2-\sigma^2|<\varepsilon)=1 \ \forall \ \varepsilon>0.$

Again, it suffices to show that $\lim_{n\to\infty} P(|E[\hat{\sigma}_n^2] - \sigma^2| < \varepsilon) = 1$. From Problem Set 4,

$$E[\hat{\sigma}_n^2] - \sigma^2 = \frac{n-1}{n}\sigma^2 - \sigma^2 = \frac{-1}{n}\sigma^2$$

Let $\varepsilon > 0$. Then $\exists m \in \mathbb{N}$ st $\forall n \geq m$, $\frac{1}{n} < \varepsilon$. Thus, $\forall n \geq m$,

$$\left|\frac{1}{n}\right|<\varepsilon\Rightarrow\left|\frac{1}{n}\sigma^2\right|<\varepsilon\Rightarrow\left|(\frac{n-1}{n}-\frac{n}{n})\sigma^2\right|<\varepsilon\Rightarrow\left|\frac{n-1}{n}\sigma^2-\sigma^2\right|<\varepsilon\Rightarrow\left|E[\hat{\sigma}_n^2]-\sigma^2\right|<\varepsilon.$$

Thus, $\lim_{n\to\infty} P(|\hat{\sigma}_n^2 - \sigma^2| < \varepsilon) = 1 \forall \varepsilon > 0.$

7.3

 \mathbf{a}

$$X \sim N(185, 50^2), n = 100, \bar{X} = 210$$

$$P(q_{.05} < \frac{\bar{X} - \mu}{\sigma / \sqrt{n}} < q_{.95}) = 1 - \alpha = .9$$

Solve for μ :

$$P(\bar{X} - q_{.95} \frac{\sigma}{\sqrt{n}} < \mu < \bar{X} - q_{.05} \frac{\sigma}{\sqrt{n}}) = .9$$

$$(210 - 1.645 \frac{50}{\sqrt{100}} < \mu < 210 + .1.645 \frac{50}{\sqrt{100}}) = (205.2 < \mu < 209.8)$$

Thus the 90% confidence interval is (201.8, 218.8)

b

Solve for n:

$$\bar{X} - LB_{.05} = \bar{X} - (\bar{X} - 1.96(\frac{50}{\sqrt{n}})) \le 10$$

$$1.96\frac{50}{10} \le \sqrt{n}$$

$$n > 96.04$$

A sample size of at least 97 is needed.

 \mathbf{c}

Like b,

$$2.576\frac{50}{10} \leq \sqrt{n}$$

$$n \ge 15.9$$

A sample size of at least 166 is needed

7.8

For n = 30, the code generating the simulation is below, along with the proportion of times the confidence interval missed μ . Compared to the next three numbers for n=60, n=100, and n=250, the frequency of missing μ converges to the thoretical frequency of .05.

```
# Run simulation
low <- 0
high <- 0
n <- 30
q <- qt(.975, n-1)
N <- 10^5
for(i in 1:N){
    x <- rgamma(n,shape = 5, rate = 2)
    xbar <- mean(x)
    s <- sd(x)</pre>
```

```
L <- xbar - q*s/sqrt(n)
  U <- xbar + q*s/sqrt(n)
  if(U<5/2){
  low \leftarrow low + 1
  }
 if(5/2<L){
  high <- high + 1
}
(high+low)/N
## [1] 0.05396
## [1] 0.05194
```

[1] 0.05169 ## [1] 0.05009

7.34