# MATH 392 Problem Set 2

# EJ Arce

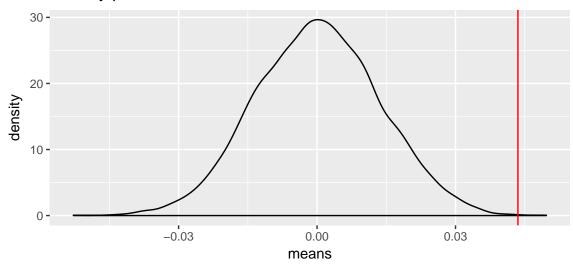
2 February 2018

#### 3.4

 $\mathbf{a}$ 

```
FlightDelays <- FlightDelays %>%
  mutate(Delayed20 = ifelse(Delay > 20, 1, 0))
# Calculating observed difference in the two groups' mean proportions
xobs <- mean(FlightDelays$Delayed20[FlightDelays$Carrier == "UA"] -</pre>
               mean(FlightDelays$Delayed20[FlightDelays$Carrier=="AA"]))
# Running simulation for hypothesis testing
nsim <- 10000
means <- rep(NA, nsim)
for(i in 1:nsim){
 perm <- sample(FlightDelays$Carrier, replace=F)</pre>
 means[i] <- mean(FlightDelays$Delayed20[perm=="UA"] -</pre>
                     mean(FlightDelays$Delayed20[perm=="AA"]))
}
simdf <- data.frame(means)</pre>
# Plotting the simulated null distribution
ggplot(simdf, aes(x = means)) +
  geom_density() +
  geom_vline(xintercept = xobs, col = "red") +
  ggtitle("Density plot of observed mean differences from 10000 simulations")
```

# Density plot of observed mean differences from 10000 simulations



The red vertical line indicates the difference in proportions observed in the actual dataset. The p-value for a two-tailed test is calculated below.

```
(sum(means > xobs) + 1)/(length(means) + 1) * 2
## [1] 0.0009999
```

We are testing to see if the variance in flight delays for United Airlines is greater than the variance for American Airlines. Thus, we are conducting a one-tailed significance test. Speicically,

```
H_0: \rho_{UA} \leq \rho_{AA}
```

```
H_A: \rho_{UA} > \rho_{AA}
```

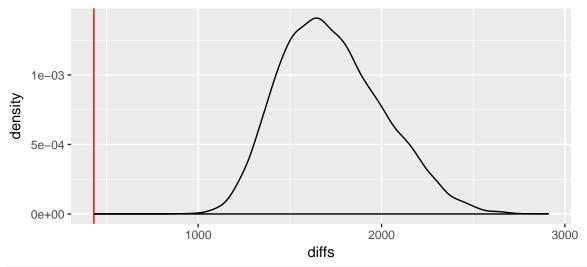
```
# Variance in flight delay lengths for each carrier
varUA <- var(FlightDelays$Delay[FlightDelays$Carrier == "UA"])
varAA <- var(FlightDelays$Delay[FlightDelays$Carrier == "AA"])
varUA
## [1] 2038</pre>
```

#### ## [1] 1606

b

varUA and varAA indicate the variances of United Airlines' and American Airlines' delay times, respectively. A simulation just like the last problem will be used to test if the difference in these variances is statistically significant.

# Density plot of observed variance differences from 10000 simulatic



```
(sum(diffs > obs.diff) + 1)/(length(diffs) + 1)
```

## [1] 1

# 3.16

 $\mathbf{a}$ 

# table(GSS2002\$Gender,GSS2002\$Pres00)

b

```
gender <- GSS2002$Gender
pres <- GSS2002$Pres00
chisq.test(gender,pres)</pre>
```

```
##
## Pearson's Chi-squared test
##
## data: gender and pres
## X-squared = 33, df = 4, p-value = 1e-06
```

 $\mathbf{c}$ 

```
# Chi-squared test using permutations
x2.obs <- chisq.test(gender,pres)$statistic
nsim <- 10000
x2.stats <- rep(NA,nsim)
for(i in 1:nsim){
   perm <- xtabs(~sample(Gender, replace=F) + Pres00, data = GSS2002)
   x2.stats[i] <- chisq.test(perm)$statistic
}
(sum(x2.stats > x2.obs) + 1)/(nsim+1)
```

#### ## [1] 9.999e-05

None of the simulated  $\chi^2$  values were greater than our observed  $\chi^2_{obs}$  value of 33.29, resulting in our very low p-value.

#### 3.22

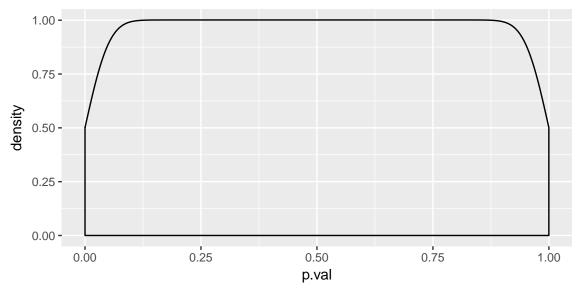
```
q \leftarrow c(.2,.4,.6,.8)
d <- data.frame("quantile" = q)</pre>
obs.stats <-c(12.57,16.87,20.73,24.66)
d<-cbind(d,obs.stats)</pre>
norm.stats <- rep(NA,4)
for(i in 1:4){
  norm.stats[i] <- qnorm(q[i],22,7)</pre>
d <- cbind(d,norm.stats)</pre>
    quantile obs.stats norm.stats
## 1
         0.2
                  12.57
                             16.11
## 2
          0.4
                  16.87
                               20.23
## 3
                   20.73
          0.6
                               23.77
## 4
          0.8
                   24.66
                               27.89
# Calculate observed chi-squared value
((3.54^2)/16.11) + ((3.36^2)/20.23) + ((3.04^2)/23.77) + ((3.23^2)/27.89)
## [1] 2.099
```

#### 3.31

# **Empirical Solution**

```
nsim <- 10000
t.obs <- rnorm(nsim,0,1)
ts <- data.frame(t.obs)
ts <- ts %>%
    arrange(t.obs) %>%
    mutate(t.obs = abs(t.obs))
p.val <- rep(NA,nsim)
for(i in 1:nsim){
    p.val[i] <- (sum(ts$t.obs>ts$t.obs[i])+1)/(length(ts$t.obs)+1)
}
```

ts <- cbind(ts, p.val)
ggplot(ts, aes(x=p.val)) +
 geom\_density()</pre>



As we'd expect, the simulated density plot follows a uniform distribution.

# **Analytical Solution**

Consider a test statistic  $t = T(x_1, ..., x_n)$ . Its corresponding p-value is calculated by solving

$$p = Pr(T(X) > t|H_0).$$

This makes p a random variable as well, since its calculated probability depends on the random variable T(X). Thus the p-value follows some probability distribution P(T). Since the p-values are drawn from the distribution of T(X), then the p-values have a one-to-one correspondence to each observed test statistic t. Thus,

$$Pr(P(T) \ge p) = Pr(T(X) \ge t) = p.$$

This shows that P(T) follows a uniform distribution, where Pr(p) = 1/n.

#### 3.32

Let Z denote the standard normal random variable. Then  $Z \sim N(0,1)$ . Suppose  $X = Z^2$ . Show that  $X \sim \chi^2_{df=1}$ .

The pdf of Z is already known to be

$$f_Z(z) = \frac{1}{\sqrt{2\pi}} e^{-\frac{x^2}{2}}$$

$$F_Z(z) = \frac{1}{2}(1 + erf(\frac{x}{\sqrt{2}}))$$

Using the cdf method,

$$F_X(x) = P(X \le x) = P(Z^2 \le x) = P(Z \le \sqrt{x}) = F_Z(\sqrt{x})$$

$$F_Z(\sqrt{x}) = \int_{-\sqrt{x}}^{\sqrt{x}} \frac{1}{\sqrt{2\pi}} e^{\frac{-x^2}{2}} dx$$

$$F_Z(\sqrt{x}) = 2 \int_0^{\sqrt{x}} \frac{1}{\sqrt{2\pi}} e^{\frac{-x^2}{2}} dx$$

$$F_Z(\sqrt{x}) = 2\frac{1}{2}erf(\frac{\sqrt{x}}{\sqrt{2}})$$

$$F_Z(\sqrt{x}) = erf(\frac{\sqrt{x}}{\sqrt{2}})$$

$$f_X(x) = \frac{\partial}{\partial x} F_Z(\sqrt{x}) = f_Z(\sqrt{x}) \frac{1}{2} x^{-\frac{1}{2}}$$

$$f_X(x) = \frac{1}{\sqrt{2\pi}} e^{-\frac{x}{2}} \frac{1}{2} x^{-\frac{1}{2}}$$

$$f_X(x) = \frac{x^{-\frac{1}{2}}e^{-\frac{x}{2}}}{2\sqrt{2\pi}}$$

Notice that

$$\Gamma(1/2) = \int_0^\infty t^{-1/2} e^{-t} dt = \sqrt{2\pi},$$

so we get

$$f_X(x) = \frac{x^{-\frac{1}{2}}e^{-\frac{x}{2}}}{2\Gamma(1/2)}$$

Thus,

$$f_X(x) \sim \chi^2_{df=1}$$