MATH 392 Problem Set 4

EJ Arce

17 February 2018

6.2

Analytical Solution

Let $x_1, ..., x_n \sim \text{Poisson}(\lambda)$. Show the maximum likelihood estimate of λ is $\hat{\lambda} = \bar{x}$:

$$L(\lambda|x_1,...,x_n) = P(X_1 = x_1,...,X_n = x_n)$$

$$L(\lambda|x_1,...,x_n) = P(X = x_1)...P(X = x_n)$$

$$L(\lambda|x_1,...,x_n) = \prod_{i=1}^{n} \frac{\lambda^{x_i} e^{-\lambda}}{x_i!}$$

$$L(\lambda|x_1,...,x_n) = \frac{e^{-\lambda n} \lambda^{\sum_{i=1}^{n} x_i}}{\prod_{i=1}^{n} x_i!}$$

Take the natural log of both sides to make the derivation easier:

$$ln(L(\lambda|x_1,...,x_n)) = ln(e^{-n\lambda}) + ln(\lambda^{\sum_{i=1}^n x_i}) - ln(\prod_{i=1}^n x_i!)$$

$$ln(L(\lambda|x_1,...,x_n)) = -n\lambda + ln(\lambda) \sum_{i=1}^{n} x_i - ln(\prod_{i=1}^{n} x_i!)$$

Derive with respect to λ and set equal to 0:

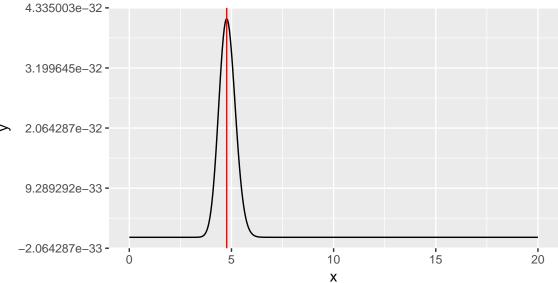
$$\frac{\partial ln(L(\lambda|x_1,...,x_n))}{\partial \lambda} = -n + \frac{\sum_{i=1}^{n} x_i}{\lambda} = 0$$

$$\frac{\partial ln(L(\lambda|x_1,...,x_n))}{\partial \lambda} = \lambda = \frac{\sum_{i=1}^{n} x_i}{n} = \bar{x}$$

Thus $\lambda_{MLE} = \bar{x}$

Empirical Solution

```
# Build function
e \leftarrow exp(1)
pois <- function(lambda, x){</pre>
  n <- length(x)
  sum.x \leftarrow sum(x)
  prod.fact <- prod(factorial(x))</pre>
  e^(-lambda*n)*(lambda^(sum.x))/(prod.fact)
# Create sequence of lambdas for MLE plot
lambdas \leftarrow seq(from = 0, to = 20, by = .05)
\# Draw random sample of x's
x \leftarrow sample(1:10,30,replace=T)
\# Apply pois function to find likelihoods
L_pois <- pois(lambdas,x)</pre>
# Plot
df <- data.frame(x = lambdas,</pre>
                   y = L_pois)
ggplot(df, aes(x=x,y=y)) +
  geom_line() +
  geom_vline(xintercept = mean(x), col = "red")
```



A red line above indicates the mean of the random sample. As expected, the highest likelihood for λ is at the sample mean.

6.8

Analytical Solution

$$f(x;\theta) = \frac{\sqrt{2/\pi}x^2e^{-x^2/2\theta^2}}{\theta^3}$$

$$L(\theta|x_1, ..., x_n) = \prod_{i}^{n} f(x_i; \theta)$$

$$L(\theta|x_1, ..., x_n) = \prod_{i}^{n} \frac{\sqrt{2/\pi} x_i^2 e^{-x_i^2/2\theta^2}}{\theta^3}$$

$$L(\theta|x_1, ..., x_n) = (\frac{\sqrt{2/\pi}}{\theta^3})^n \prod_{i}^{n} x_i^2 e^{-x_i^2/2\theta^2}$$

$$ln(L(\theta|x_1, ..., x_n)) = ln(\frac{\sqrt{2/\pi}}{\theta^3})^n + \sum_{i}^{n} ln(x_i^2) + \sum_{i}^{n} ln(e^{-x_i^2/2\theta^2})$$

$$ln(L(\theta|x_1, ..., x_n)) = ln\sqrt{2/\pi} - ln\theta^3 + \sum_{i}^{n} ln(x_i^2) + \sum_{i}^{n} -x_i^2/2\theta^2$$

Derive with respect to θ and set to 0:

$$\frac{\partial (L(\theta|x_1, ..., x_n))}{\partial \theta} = -\frac{n}{\theta^3} 3\theta^2 + \sum_{i=1}^n x_i^2 / \theta^3 = 0$$

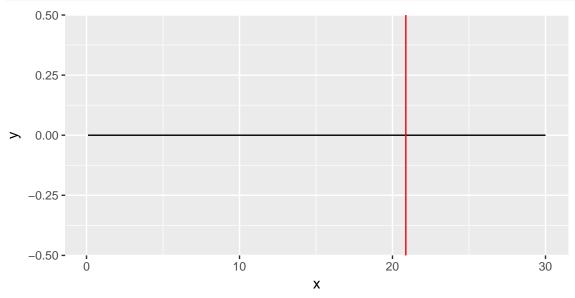
$$\frac{3n}{\theta} = \frac{\sum_{i=1}^n x_i^2}{\theta^3}$$

$$\theta = \frac{\bar{x}\sqrt{n}}{\sqrt{3}}$$

Thus $\theta_{MLE} = \frac{\bar{x}\sqrt{n}}{\sqrt{3}}$.

Empirical Solution

```
# Build function
f <- function(theta, x){</pre>
  e \leftarrow exp(1)
  n <- length(x)
  sec.term \langle (x^2) * e^((-x^2)/(2*(theta)^2))
  prod.x <- prod(sec.term)</pre>
  ((sqrt(2/pi)/theta)^n)*prod.x
# Create sequence of thetas for MLE plot
thetas \leftarrow seq(from = .1, to = 30, by = .1)
# Reset seed and draw random sample of x's
set.seed(23)
x <- sample(1:10,30,replace=T)</pre>
# Apply function to find likelihoods
L_theta <- f(thetas,x)</pre>
# Find x-intercept from analytical solution
n <- length(x)
```



6.11

Analytical Solution

As before,

$$L(\lambda) = \prod_{i}^{n} f(x_{i}) \prod_{j}^{m} f(y_{j})$$

$$L(\lambda) = \prod_{i}^{n} \lambda e^{-\lambda x_{i}} \prod_{j}^{m} \lambda e^{-\lambda y_{j}}$$

$$L(\lambda) = \lambda^{n+m} \lambda e^{-\lambda (\sum_{i}^{n} x_{i} + \sum_{j}^{m} y_{j})}$$

$$ln(L(\lambda)) = (n+m)ln(\lambda) - \lambda (\sum_{i}^{n} x_{i} + \sum_{j}^{m} y_{j})$$

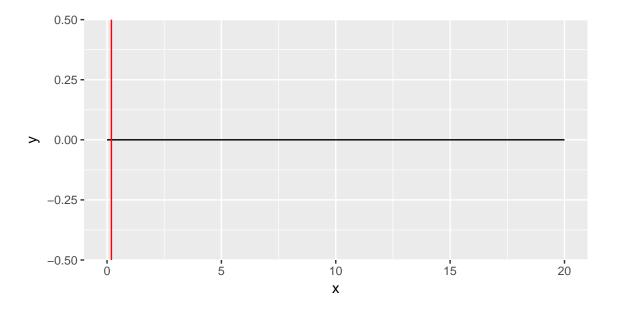
$$\frac{\partial ln(L(\lambda))}{\partial \lambda} = (n+m)/\lambda - (\sum_{i}^{n} x_{i} + \sum_{j}^{m} y_{j}) = 0$$

$$\lambda = \frac{n+m}{\sum_{i}^{n} x_{i} + \sum_{j}^{m} y_{j}}$$

```
Thus, \lambda_{MLE} = \frac{n+m}{\sum_{i}^{n} x_i + \sum_{j}^{m} y_j}.
```

Empirical Solution

```
# Build function
e \leftarrow exp(1)
pois <- function(lambda,x,y){</pre>
  n \leftarrow length(x)
  m <- length(y)
  p.x <- (lambda^n)*prod((e^(-lambda*x)))</pre>
  p.y <- (lambda^m)*prod((e^(-lambda*y)))</pre>
 p.x*p.y
}
# Create sequence of lambdas for MLE plot
lambdas \leftarrow seq(from = 0, to = 20, by = .05)
# Draw random sample of x's
set.seed(11)
x <- sample(1:10,30,replace=T)</pre>
set.seed(35)
y <- sample(1:10,40,replace=T)
# Apply pois function to find likelihoods
L_pois <- pois(lambdas,x,y)</pre>
## Warning in -lambda * x: longer object length is not a multiple of shorter
## object length
## Warning in -lambda * y: longer object length is not a multiple of shorter
## object length
# Calculate analytical solution
obs <- (length(x)+length(y))/(sum(x) + sum(y))
# Plot
df <- data.frame(x = lambdas,</pre>
                  y = L_pois)
ggplot(df, aes(x=x,y=y)) +
  geom_line() +
  geom_vline(xintercept = obs, col = "red")
```



6.13

a

$$L(\alpha|X;\beta) = \alpha\beta x^{\beta-1}e^{-\alpha x^{\beta}}$$

$$ln(L(\alpha|X;\beta)) = ln\alpha + ln\beta + (\beta - 1)lnx + ln(e^{-\alpha x^{\beta}})$$

$$ln(L(\alpha|X;\beta)) = ln\alpha + ln\beta + (\beta - 1)lnx - \alpha x^{\beta}$$

$$\frac{\partial ln(L(\alpha|X;\beta))}{\partial \alpha} = \frac{1}{\alpha} - x^{\beta} = 0$$

$$\alpha = \frac{1}{x^{\beta}}$$

Thus $\alpha_{MLE} = \frac{1}{x^{\beta}}$.

 \mathbf{b}

$$L(\alpha; \beta | X) = \alpha \beta x^{\beta - 1} e^{-\alpha x^{\beta}}$$

Take the natural log of both sides:

$$ln(L(\alpha; \beta|X)) = ln\alpha + ln\beta + (\beta - 1)lnx - \alpha x^{\beta}$$

Differentiate with respect to α and β and set to 0. Solve simultaneously for a and b:

$$\frac{\partial ln(L(\alpha|X))}{\partial \alpha} = 1/\alpha - x^{\beta} = 0$$

and

$$\frac{\partial ln(L(\beta|X))}{\partial \beta} = 1/\beta - + lnx - ax^{\beta} = 0$$