MATH 392 Problem Set 6

EJ Arce 19 March 2018

```
# Load data

x <- c(47, 126, 285, 318, 142, 55, 231,

102, 164, 85, 242, 62, 289, 290)
```

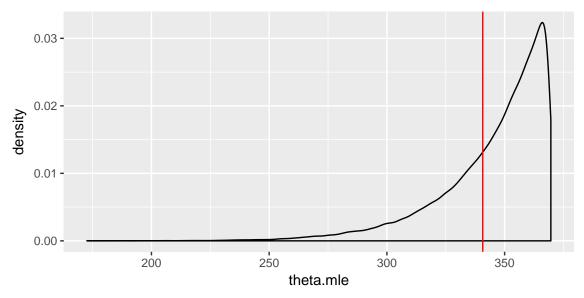
1.1

$$X \sim Unif(0, \theta)$$
 $H_0: \theta = 345$ $H_A: \theta > 345$

$$\hat{\theta}_{MLE,corr} = \frac{n+1}{n} X_{max}$$

```
# Calculate test statistic
xmax <- max(x)
n <- length(x)
theta.test.mle <- (n+1)*xmax/n

# Simulate sampling distribution of thetas from null distribution
nsim <- 10^5
max.null <- 345
theta.mle <- rep(NA,nsim)
for(i in 1:nsim){
    sample <- runif(n,0,max.null)
    theta.mle[i] <- (n+1)*max(sample)/n
}
null.samp.mle <- data.frame(theta.mle)
ggplot(null.samp.mle, aes(x=theta.mle)) +
    geom_density() +
    geom_vline(xintercept = theta.test.mle, col = "red")</pre>
```



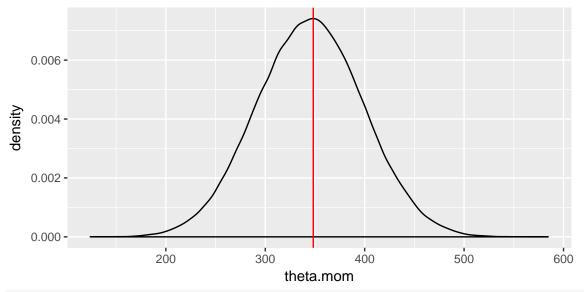
Now, calculate the proportion of simulations that produced estimates more extreme than the test statistic: (sum(null.samp.mle\$theta.mle > theta.test.mle) + 1)/(length(null.samp.mle\$theta.mle) + 1)

[1] 0.6810232

Using a reasonable α level, the simulation results are not statistically significant enough to reject the null hypothesis.

1.2

```
theta.test.mom <- 2*mean(x)
nsim <- 10^5
max.null <- 345
theta.mom <- rep(NA,nsim)
for(i in 1:nsim){
   sample <- runif(n,0,max.null)
    theta.mom[i] <- 2*mean(sample)
}
null.samp.mom <- data.frame(theta.mom)
ggplot(null.samp.mom, aes(x=theta.mom)) +
   geom_density() +
   geom_vline(xintercept = theta.test.mom, col = "red")</pre>
```



(sum(null.samp.mom\$theta.mom > theta.test.mom) + 1)/(length(null.samp.mom\$theta.mom) + 1)

[1] 0.4769752

Again, we cannot reject the null hypothesis.

1.3

The two types of error committed when concluding hypothesis tests are

- Type I: rejecting H_0 when it is actually true
- Type II:retaining H_0 when H_A is actually true

In the case of the German tank problem, committing a type I error would mean that the Western allies believe that there are more German tanks than there actually are. A type II error would mean that the Western allies believe the Germans have 345 tanks, when in fact they have more. Thus, a type II error is more consequential than a type I error.

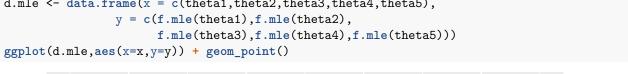
1.4

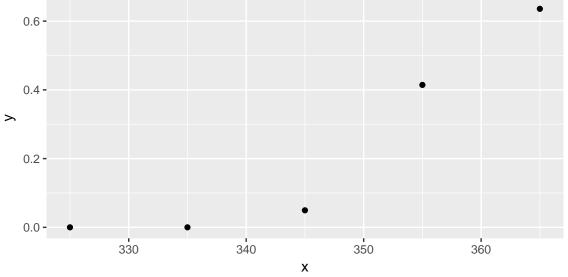
Suppose $\alpha = .05$.

```
n <- 17
nsim <- 10^5
theta1 <- 325
theta2 <- 335
theta3 <- 345
theta5 <- 365

# Power curve for theta parameters
f.mle <- function(theta.a){
    nsim <- 10^5
    theta.0 <- 345
    thetas.0 <- rep(NA,nsim)
    thetas.a <- rep(NA,nsim)
    for(i in 1:nsim){</pre>
```

```
sample.0 <- runif(n,0,theta.0)</pre>
    sample.a <- runif(n,0,theta.a)</pre>
    thetas.0[i] <- (n+1)*(max(sample.0))/n
    thetas.a[i] <- (n+1)*(max(sample.a))/n
  sampling.0 <- data.frame(thetas.0)</pre>
  sampling.a <- data.frame(thetas.a)</pre>
  alpha.crit <- quantile(sampling.0$thetas.0,.95)</pre>
  power <- sum(sampling.a$thetas.a > alpha.crit)/length(sampling.a$thetas.a)
  power
d.mle <- data.frame(x = c(theta1,theta2,theta3,theta4,theta5),</pre>
                 y = c(f.mle(theta1),f.mle(theta2),
                       f.mle(theta3),f.mle(theta4),f.mle(theta5)))
```





Now do the same thing for the method of moments estimator:

```
f.mom <- function(theta.a){</pre>
  nsim <- 10<sup>5</sup>
  theta.0 <- 345
  thetas.0 <- rep(NA,nsim)
  thetas.a <- rep(NA,nsim)</pre>
  for(i in 1:nsim){
    sample.0 <- runif(n,0,theta.0)</pre>
    sample.a <- runif(n,0,theta.a)</pre>
    thetas.0[i] <- 2*mean(sample.0)</pre>
    thetas.a[i] <- 2*mean(sample.a)</pre>
  sampling.0 <- data.frame(thetas.0)</pre>
  sampling.a <- data.frame(thetas.a)</pre>
  alpha.crit <- quantile(sampling.0$thetas.0,.95)</pre>
  power <- sum(sampling.a$thetas.a > alpha.crit)/length(sampling.a$thetas.a)
  power
}
```

As expected for both tests, as the true alternative hypothesis increases, the power increases. For true alternative hypotheses with $\theta > 345$ (which is what we would get when we do reject H_0), the power is greater using the MLE statistic. Thus, the MLE test would be a better option.