MATH 392 Problem Set 2

EJ Arce

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3.4

 \mathbf{a}

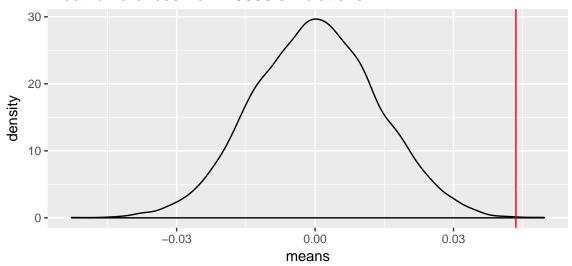
We are testing to see if the distributions of the proportion of flights delayed by more than 20 minutes, θ , differs by airline:

$$H_0: \theta_{UA} - \theta_{AA} = 0$$

$$H_A: \theta_{UA} - \theta_{AA} \neq 0$$

```
FlightDelays <- FlightDelays %>%
  mutate(Delayed20 = ifelse(Delay > 20, 1, 0))
# Calculating observed difference in the two groups' mean proportions
xobs <- mean(FlightDelays$Delayed20[FlightDelays$Carrier == "UA"] -</pre>
               mean(FlightDelays$Delayed20[FlightDelays$Carrier=="AA"]))
# Running simulation for hypothesis testing
nsim <- 10000
means <- rep(NA, nsim)
for(i in 1:nsim){
  perm <- sample(FlightDelays$Carrier, replace=F)</pre>
  means[i] <- mean(FlightDelays$Delayed20[perm=="UA"]) -</pre>
                      mean(FlightDelays$Delayed20[perm=="AA"])
simdf <- data.frame(means)</pre>
# Plotting the simulated null distribution
ggplot(simdf, aes(x = means)) +
  geom_density() +
  geom_vline(xintercept = xobs, col = "red") +
  ggtitle("Mean differences from 10000 simulations")
```

Mean differences from 10000 simulations



The red vertical line indicates the difference in proportions observed in the actual dataset. The p-value for a two-tailed test is calculated below.

```
(sum(means > xobs) + 1)/(length(means) + 1) * 2
```

[1] 0.0014

b

We are testing to see if the variance in flight delays for United Airlines is greater than the variance for American Airlines. Thus, we are conducting a one-tailed significance test. Specicially,

 $H_0: \rho_{UA} \leq \rho_{AA}$

 $H_A: \rho_{UA} > \rho_{AA}$

```
# Variance in flight delay lengths for each carrier
varUA <- var(FlightDelays$Delay[FlightDelays$Carrier == "UA"])
varAA <- var(FlightDelays$Delay[FlightDelays$Carrier == "AA"])
varUA</pre>
```

[1] 2038 varAA

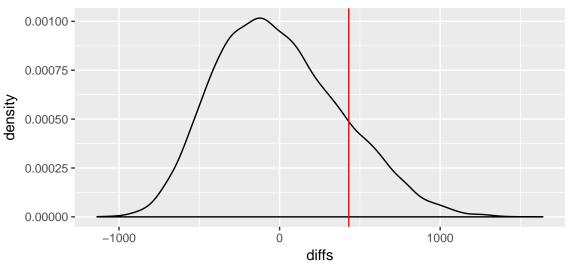
[1] 1606

varUA and varAA indicate the variances of United Airlines' and American Airlines' delay times, respectively. A simulation just like the last problem will be used to test if the difference in these variances is statistically significant.

```
# Calculate test statistic
obs.diff <- varUA-varAA

# Run simulation
nsim <- 10000
diffs <- rep(NA, nsim)</pre>
```

Variance differences from 10000 simulations



```
(sum(diffs > obs.diff) + 1)/(length(diffs) + 1)
```

[1] 0.1451

The density plot and calculated p-value show that the observed variance for United Airlines delays is not significantly greater than observed variances for American Airlines.

3.16

a

table(GSS2002\$Gender,GSS2002\$Pres00)

b

```
gender <- GSS2002$Gender
pres <- GSS2002$Pres00
chisq.test(gender,pres)
##
    Pearson's Chi-squared test
##
## data: gender and pres
## X-squared = 33, df = 4, p-value = 1e-06
\mathbf{c}
# Remove NAs
GSS2002 <- GSS2002 %>%
 filter(!is.na(Gender),
         !is.na(Pres00))
# Chi-squared test using permutations
x2.obs <- chisq.test(gender,pres)$statistic</pre>
nsim <- 10000
x2.stats <- rep(NA,nsim)
for(i in 1:nsim){
  perm <- xtabs(~sample(Gender, replace=F) + Pres00, data = GSS2002)</pre>
  x2.stats[i] <- chisq.test(perm)$statistic</pre>
}
(sum(x2.stats > x2.obs) + 1)/(nsim+1)
```

[1] 9.999e-05

None of the simulated χ^2 values were greater than our observed χ^2_{obs} value of 33.29, resulting in our very low p-value.

3.22

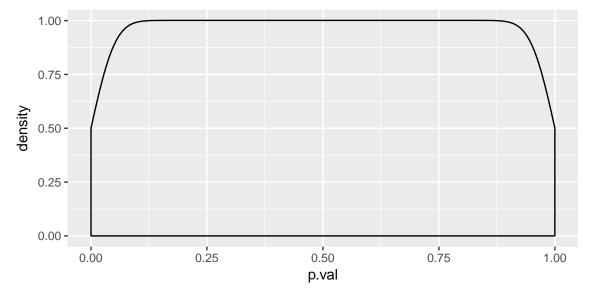
```
q \leftarrow c(.2,.4,.6,.8)
d <- data.frame("quantile" = q)</pre>
obs.stats <-c(12.57,16.87,20.73,24.66)
d<-cbind(d,obs.stats)</pre>
exp.stats <- rep(NA,4)
for(i in 1:4){
  exp.stats[i] \leftarrow qnorm(q[i],22,7)
d <- cbind(d,exp.stats)</pre>
    quantile obs.stats exp.stats
## 1
        0.2
                  12.57
                              16.11
## 2
          0.4
                   16.87
                              20.23
## 3
          0.6
                   20.73
                              23.77
          0.8
                   24.66
                              27.89
# Calculate observed chi-squared value
((3.54^2)/16.11) + ((3.36^2)/20.23) + ((3.04^2)/23.77) + ((3.23^2)/27.89)
```

[1] 2.099

3.31

Empirical Solution

```
nsim <- 10000
t.obs <- rnorm(nsim,0,1)
ts <- data.frame(t.obs)
ts <- ts %>%
    arrange(t.obs) %>%
    mutate(t.obs = abs(t.obs))
p.val <- rep(NA,nsim)
for(i in 1:nsim){
    p.val[i] <- (sum(ts$t.obs>ts$t.obs[i])+1)/(length(ts$t.obs)+1)
}
ts <- cbind(ts, p.val)
ggplot(ts, aes(x=p.val)) +
    geom_density()</pre>
```



As we'd expect, the simulated density plot follows a uniform distribution.

Analytical Solution

Consider a test statistic $t = T(x_1, ..., x_n)$. Its corresponding p-value is calculated by solving

$$p = Pr(T(X) < t|H_0).$$

This makes p a random variable as well, since its calculated probability depends on the random variable T(X). Thus the p-value p follows some probability distribution $p = F_T(T)$. Since the p-values are drawn from the distribution of T(X), then the p-values have a one-to-one correspondence to each observed test statistic t. Thus,

$$F_P(p) = Pr(P \le p) = Pr(T(X) \le t) = Pr(F_T(T) \le p).$$

$$F_P(p) = Pr(F_T^{-1}F_T(T) \le F_T^{-1}(p))$$

$$F_P(p) = Pr(T \le F_T^{-1}(p))$$

$$F_P(p) = F_T(F_T^{-1}(p)) = p$$

This shows that P(T) follows a uniform distribution, where Pr(p) = 1/n.

3.32

Let Z denote the standard normal random variable. Then $Z \sim N(0,1)$. Suppose $X = Z^2$. Show that $X \sim$

The pdf of Z is already known to be

$$f_Z(z) = \frac{1}{\sqrt{2\pi}} e^{-\frac{z^2}{2}}$$

Using the cdf method,

$$\begin{split} F_X(x) &= P(X \le x) = P(Z^2 \le x) = P(-\sqrt{x} \le Z \le \sqrt{x}) = F_Z(\sqrt{x}) - F_Z(-\sqrt{x}) \\ P(-\sqrt{x} \le Z \le \sqrt{x}) &= F_Z(\sqrt{x}) - F_Z(-\sqrt{x}) = \int_{-\sqrt{x}}^{\sqrt{x}} \frac{1}{\sqrt{2\pi}} e^{\frac{-x^2}{2}} dx \\ F_Z(\sqrt{x}) - F_Z(-\sqrt{x}) &= 2 \int_{0}^{\sqrt{x}} \frac{1}{\sqrt{2\pi}} e^{\frac{-x^2}{2}} dx \\ F_Z(\sqrt{x}) - F_Z(-\sqrt{x}) &= 2\frac{1}{2} erf(\frac{\sqrt{x}}{\sqrt{2}}) \\ F_Z(\sqrt{x}) - F_Z(-\sqrt{x}) &= erf(\frac{\sqrt{x}}{\sqrt{2}}) \\ f_X(x) &= \frac{\partial}{\partial x} (F_Z(\sqrt{x}) - F_Z(-\sqrt{x})) = 2f_Z(\sqrt{x}) \frac{1}{2} x^{-\frac{1}{2}} \\ f_X(x) &= \frac{1}{\sqrt{2\pi}} e^{-\frac{x}{2}} x^{-\frac{1}{2}} \\ f_X(x) &= \frac{x^{-\frac{1}{2}} e^{-\frac{x}{2}}}{\sqrt{2\pi}} \\ \Gamma(1/2) &= \int_{0}^{\infty} t^{-1/2} e^{-t} dt = \sqrt{\pi}, \end{split}$$

Notice that

$$\Gamma(1/2) = \int_0^\infty t^{-1/2} e^{-t} dt = \sqrt{\pi},$$

so we get

$$f_X(x) = \frac{x^{-\frac{1}{2}}e^{-\frac{x}{2}}}{\sqrt{2}\Gamma(1/2)}$$

Thus,

$$f_X(x) \sim \chi^2_{df=1}$$