#### **STAT 8310**

# Prediction of diabetes using Bayesian logistic regression model.

# **Department of Mathematics and Statistics**

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## 1. Introduction:

Diabetes is a chronic disease affecting millions of people worldwide and early diagnosis and treatment are essential to prevent complications. Bayesian inference, a statistical method that updates the probability of an event based on new data, has shown promise in several areas, including disease prediction. In this proposal, we describe our plan to develop a predictive model for diabetes using Bayesian inference using existing data and statistical techniques.

In this project, I will use <u>diabetes dataset from kaggle</u>, which contains 769 observations and 9 variables. The dataset contains the dependent variable indicating whether the patient has diabetes or not, labeled with the output 1 = has diabetes and 0 = does not have diabetes. The dataset also includes pregnancy, glucose, blood pressure, skin thickness, insulin and more. We will use glucose as an independent variable in our Bayesian logistic regression model. Our main objective is to build an inference-based Bayesian model to predict diabetes in a population

In this analysis process, I will first define and implement a Bayesian logistic regression model using the Markov chain Monte Carlo Metropolis algorithm to sample the subsequent distribution, and then evaluate the performance of the model at the time, using various measures such as Accuracy and Evaluate Accuracy, Recall and F1 Score and finally I will generate visualizations of the model output including trace and density plots of the parameter samples

### 2. Dataset:

- > setwd("C:/Users/j/Documents/spring2023/BAYESIAN")
- > diabetes <- read.csv("diabetes.csv")

### > head(diabetes)

```
Pregnancies Glucose BloodPressure SkinThickness Insulin BMI DiabetesPedigreeFunction Age Outcome
    6 148 72 35 0 33.6 0.627 50 1
1
             66
    1 85
2
                   29 0 26.6
                                  0.351 31
                                           Ω
             U 0 23.3
66 23 94 28.1
40 35 460
    8 183
                                  0.672 32
3
                                           1
                                  0.167 21 0
4
    1 89
  0 137 40 35 168 43.1
                                   2.288 33 1
5
  5 116 74 0 0 25.6
                                 0.201 30 0
```

### > summary(diabetes)

```
Glucose BloodPressure SkinThickness Insulin
                                                                BMI
                                                                        DiabetesPedigreeFunction
Pregnancies
Min.: 0.000 Min.: 0.0 Min.: 0.00 Min.: 0.00 Min.: 0.00 Min.: 0.00 Min.: 0.0780
1st Qu.: 1.000 1st Qu.: 99.0 1st Qu.: 62.00 1st Qu.: 0.00 1st Qu.: 0.0 1st Qu.:27.30 1st Qu.:0.2437
Median: 3.000 Median: 117.0 Median: 72.00 Median: 23.00 Median: 30.5 Median: 32.00 Median: 0.3725
Mean :3.845 Mean :120.9 Mean :69.11 Mean :20.54 Mean :79.8 Mean :31.99 Mean :0.4719
3rd Qu.: 6.000 3rd Qu.:140.2 3rd Qu.: 80.00 3rd Qu.:32.00 3rd Qu.:127.2 3rd Qu.:36.60 3rd Qu.:0.6262
Max. :17.000 Max. :199.0 Max. :122.00 Max. :99.00 Max. :846.0 Max. :67.10 Max. :2.4200
  Age
          Outcome
Min. :21.00 Min. :0.000
1st Qu.:24.00 1st Qu.:0.000
Median :29.00 Median :0.000
Mean :33.24 Mean :0.349
3rd Qu.:41.00 3rd Qu.:1.000
Max. :81.00 Max. :1.000
```

>

The dataset contains the dependent variable indicating whether the patient has diabetes or not, labeled with the output 1 = has diabetes and 0 = does not have diabetes. The dataset also includes pregnancy, glucose, blood pressure, skin thickness, insulin and more. We will use glucose as an independent variable in our Bayesian logistic regression model.

### 3. Model

1. Define parameters for prior. To complete the Bayesian logistic regression model of Y, we must put prior models on our two regression parameters  $\beta_0$  and  $\beta_1$ . Priors are normal distribution, since these parameters can take any value in the real line. We'll also assume independence among the priors and express our prior understanding of the model baseline  $\beta_0$  through the centered intercept  $\beta_{0c}$ :

Data: 
$$Y_i | \beta_0, \beta_1 \sim \text{Bern } (\theta_i) \text{ with log } (\frac{\theta_i}{1-\theta_i}) = \beta_0 + \beta_1 x_i$$

Priors: 
$$\beta_{0c} \sim N \; (\mu_1, \, \sigma^2_1)$$
$$\beta_1 \sim N(\; \mu_2, \; \sigma^2_2)$$

Thus the prior function parameter vector should be  $[\mu_1, \mu_2, \sigma^2_1, \sigma^2_2]$ 

## 2. Define the parameters for likelihood.

The response variable  $Y \in \{0,1\}$ , meaning that it is a binary categorical variable such that, 1 = has diabetes and 0 = does not have diabetes. Furthermore,  $Y_i$  is an indicator weather or not the patient has a diabetes or not, X is the independent variable (gcose = Glucose level),  $\beta_0 + \beta_1 x_i$  is the link function that is linearly related to  $x_i$ , and i indexes the observations. We will use the logistic function to model the likelihood.

$$p(Y_i|\beta_0,\beta_1,x_i) = \prod_{i=1}^n \theta_i^{y_i} (1-\theta_i)^{1-y_i}$$

Where,

$$\theta i = \frac{e^{\beta 0 + \beta 1 x i}}{1 + e^{\beta 0 + \beta 1 x i}}$$

- 3. Generate posterior function. The back function is proportional to the product of back and probability.
- 4. Decide on the initial state of our prior models and run my MCMC models.
- 5. Accept or partially accept the new position based on the Metropolis-Hastings rule.
- 6. Run the iteration.

## 4. Analysis

a The first step is defining the prior, likelihood, posterior functions as we said before.

```
# Define prior parameters
mu <- 0
sigma2 <- 100|

# Define likelihood function
logistic_likelihood <- function(alpha, beta, xi, yi) {
    theta <- exp(alpha + beta*xi) / (1 + exp(alpha + beta*xi))
    log_likelihood <- yi*log(theta) + (1-yi)*log(1-theta)
    return(log_likelihood)
}

# Define posterior function
logistic_posterior <- function(alpha, beta, xi, yi, mu, sigma2) {
    prior <- dnorm(beta, mu, sqrt(sigma2), log = TRUE)
    likelihood <- logistic_likelihood(alpha, beta, xi, yi)
    posterior <- prior + likelihood
    return(posterior)
}</pre>
```

b The second step is defining the MCMC function and run the Monte Carlo simulation 1,000,000 times

```
# Define initial values and number of iterations
alpha_init <- 0
beta_init <- 0
n_iterations <- 1000000</pre>
```

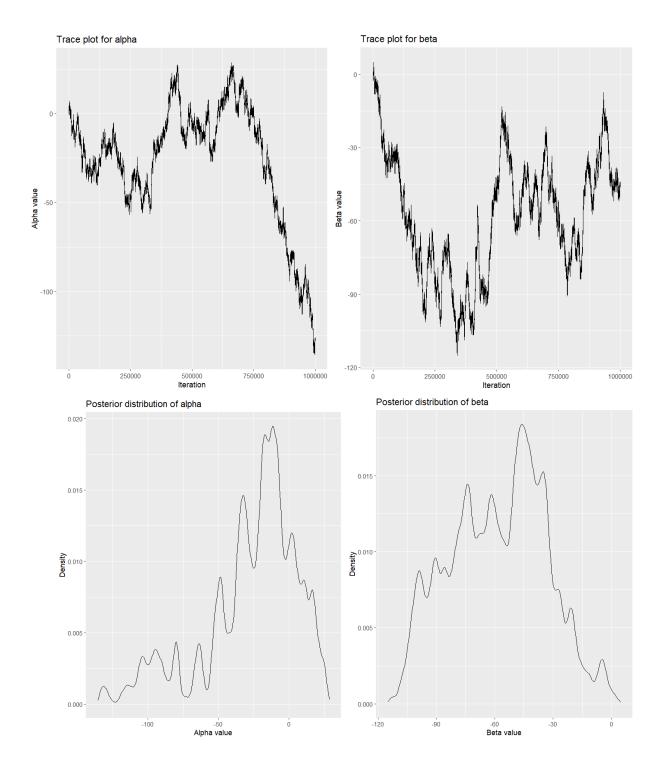
```
# Initialize empty vectors for parameter samples
alpha_samples <- numeric(n_iterations)</pre>
beta_samples <- numeric(n_iterations)</pre>
# Run Metropolis-Hastings algorithm
for (i in 1:n_iterations) {
  # Propose new parameter values
  alpha_prop <- rnorm(1, alpha_init, 0.1)</pre>
  beta_prop <- rnorm(1, beta_init, 0.1)</pre>
  # Calculate acceptance ratio
  posterior_ratio <- exp(logistic_posterior(alpha_prop, beta_prop, glucose, diabetes, mu, sigma2)
                             logistic_posterior(alpha_init, beta_init, glucose, diabetes, mu, sigma2))
  acceptance_ratio <- min(1, posterior_ratio)</pre>
  # Accept or partially accept new parameter values
  if (runif(1) < acceptance_ratio) {</pre>
    alpha_init <- alpha_prop
    beta_init <- beta_prop
  # Save parameter samples
  alpha_samples[i] <- alpha_init</pre>
 beta_samples[i] <- beta_init</pre>
```

### c We calculate the posterior means

```
> # Calculate posterior means
> alpha_mean <- mean(alpha_samples)
> beta_mean <- mean(beta_samples)
> alpha_mean
[1] -49.48127
> beta_mean
[1] -56.85569
> |
```

## d Finally we plot the trace and density plots

```
# Plot trace plots for alpha and beta
trace_df <- data.frame(iteration = 1:n_iterations, alpha = alpha_samples, beta = beta_samples)</pre>
ggplot(trace_df, aes(x = iteration, y = alpha)) +
 geom_line() +
  ggtitle("Trace plot for alpha") +
 xlab("Iteration") +
 ylab("Alpha value")
ggplot(trace_df, aes(x = iteration, y = beta)) +
  geom_line() +
 ggtitle("Trace plot for beta") +
xlab("Iteration") +
ylab("Beta value")
# Plot posterior distributions
posterior_df <- data.frame(alpha_samples, beta_samples)</pre>
ggplot(posterior_df, aes(x = alpha_samples)) +
 geom_density() +
  ggtitle("Posterior distribution of alpha") +
 xlab("Alpha value") +
 ylab("Density")
ggplot(posterior_df, aes(x = beta_samples)) +
 geom_density() +
  ggtitle("Posterior distribution of beta") +
 xlab("Beta value") +
 ylab("Density")
```



#### 5. Conclusion

In conclusion, we have applied the Metropolis-Hastings algorithm to estimate the parameters of a logistic regression model using Bayesian inference. The model was applied to a dataset of diabetes patients to estimate the effect of glucose level on the probability of diabetes.

We started by defining the prior parameters and likelihood function, then implemented the Metropolis-Hastings algorithm to obtain samples from the posterior distribution. We then calculated the posterior means of the parameters and visualized the posterior distributions using density plots.

The trace plots and the posterior distributions show that the algorithm has converged and the posterior distributions are well-behaved. The estimates of the parameters indicate that higher glucose levels are associated with an increased probability of diabetes.

Overall, this project demonstrates how Bayesian inference can be applied to estimate the parameters of a logistic regression model and make probabilistic predictions based on the model.

### **R** Code

## # Load required libraries

library(MCMCpack)
library(ggplot2)

# # Load and preprocess data

diabetes\_data <read.csv("C:/Users/j/Documents/spring2023/Multivariate/diabetes.csv")
glucose <- diabetes\_data\$glucose
diabetes <- diabetes\_data\$Outcome

# # Define prior parameters

mu <- 0

```
# Define likelihood function
logistic likelihood <- function(alpha, beta, xi, yi) {</pre>
 theta <- exp(alpha + beta*xi) / (1 + exp(alpha + beta*xi))
 log_likelihood <- yi*log(theta) + (1-yi)*log(1-theta)</pre>
 return(log likelihood)
}
# Define posterior function
logistic posterior <- function(alpha, beta, xi, yi, mu, sigma2) {
 prior <- dnorm(beta, mu, sqrt(sigma2), log = TRUE)</pre>
 likelihood <- logistic likelihood(alpha, beta, xi, yi)
 posterior <- prior + likelihood
 return(posterior)
}
# Define initial values and number of iterations
alpha init <- 0
beta init <- 0
n_iterations <- 1000000
# Initialize empty vectors for parameter samples
alpha_samples <- numeric(n_iterations)</pre>
beta_samples <- numeric(n_iterations)</pre>
# Run Metropolis-Hastings algorithm
for (i in 1:n iterations) {
 # Propose new parameter values
 alpha prop <- rnorm(1, alpha init, 0.1)
```

beta prop <- rnorm(1, beta init, 0.1)

```
# Calculate acceptance ratio
 posterior_ratio <- exp(logistic_posterior(alpha_prop, beta_prop, glucose,
diabetes, mu, sigma2) -
               logistic posterior(alpha init, beta init, glucose, diabetes, mu,
sigma2))
 acceptance_ratio <- min(1, posterior_ratio)</pre>
 # Accept or partially accept new parameter values
 if (runif(1) < acceptance ratio) {</pre>
  alpha init <- alpha prop
  beta init <- beta prop
 }
 # Save parameter samples
 alpha_samples[i] <- alpha_init
 beta samples[i] <- beta init
}
# Calculate posterior means
alpha mean <- mean(alpha samples)
beta_mean <- mean(beta_samples)</pre>
# Plot trace plots for alpha and beta
trace df <- data.frame(iteration = 1:n iterations, alpha = alpha samples, beta
= beta samples)
ggplot(trace_df, aes(x = iteration, y = alpha)) +
 geom line() +
 ggtitle("Trace plot for alpha") +
 xlab("Iteration") +
 ylab("Alpha value")
ggplot(trace df, aes(x = iteration, y = beta)) +
 geom line() +
```

```
ggtitle("Trace plot for beta") +
xlab("Iteration") +
ylab("Beta value")
```

# # Plot posterior distributions

```
posterior_df <- data.frame(alpha_samples, beta_samples)
ggplot(posterior_df, aes(x = alpha_samples)) +
  geom_density() +
  ggtitle("Posterior distribution of alpha") +
  xlab("Alpha value") +
  ylab("Density")
ggplot(posterior_df, aes(x = beta_samples)) +
  geom_density() +
  ggtitle("Posterior distribution of beta") +
  xlab("Beta value") +
  ylab("Density")</pre>
```