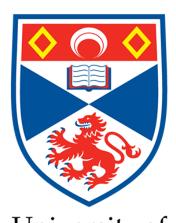
MT5751: Distance Sampling Project. Bowhead Whales

Group K

2023 - 02 - 14

Student numbers xxx, 180015716, 220013309





University of St Andrews

Abstract

Introduction

Methods

Data Collection To observe bowhead whales, visual aerial line-transect surveys were conducted in coast land offshore areas in West Greenland between $65^{\circ}400$ N and $75^{\circ}300$ N, largely covering the area of the local spring aggregation in Disko Bay. The study region on 242,650km² was divided into 16 strata based on prior knowledge of anticipated densities of bowhead whales. Between 24th March and 14th April 2012 7,836.5km of total east-west oriented transect lines were observed, with the targeted altitude and speed being 213m and 167 km/h respectively. Surveying was only carried out if Beaufort Sea States Code was 2 or less, i.e. only when the sea was calm. Inclinometers were used to measure the angle of declination for each observed group of bowhead whales (Suunto). The angles were then converted to perpendicular distances (x) using the following equation from Buckland et al. (2001):

$$x = v * tan(90 - \psi),$$

where v is the altitude of the airplane, ψ is the declination angle. Forward distance (y) to each sighting was calculated based on time of first sighting, time when passing abeam and speed of aircraft. For a more extensive description of the data collection refer to Rekdal et al. (2015).

Distance Sampling To estimate the abundance of bowhead whales, a sample representative portion of the area, i.e. the plot distance sampling, is used. Distance sampling is used in ecological surveys to account for imperfect detections and is incredibly useful considering limited resources. The bowhead whales further away are less likely to be detected, thus distance is employed as a detection bias. The detection function describes how detectability varies with distance, it includes the proportion of animals seen at a given distance and a probability of seeing an animal at a given distance. To employ the distance sampling, first of the assumptions need to be met: - the transect lines were placed at random and the bowhead whales were not affected by their placement; - animals on the transect line, in this case from distance from the plane of XX were seen with probability p = 1; - perpendicular distances from line to bowhead whales are uniform random variables $\pi(x) = 1/w$;; - distances were measured without error and to the initial location.

To then define the detection function g(x), with the assumptions stated above fulfilled, the key functions are fitted to the model: - Half-normal $g(x) = \exp(-x^2/2\sigma^2)$, - Hazard-rate $g(x) = \exp(-(x/\sigma)^{-b})$,

where σ is a scale parameter ("standard deviation") and b controls the width of the shoulder.

Data Analysis

Due to an obscured view close to the transect line, detections were left-truncated at 100m Rekdal et al. (2015).

8 models were fit to the distance data using all combinations of hazard-rate and half-normal detection functions with cosine, hermite polynomial and simple polynomial adjustments.

Model selection using Akaike information criterion (AIC) found that the best model was a half-normal with no adjustments (see figure 1). AICc was also applied due to the ratio of model parameters to observation being less than 1:40 and found the same result (Takezawa, 2014). This model had an good fit; Cramer-von Mises test (T = 0.073, p = 0.732), bootstrap Kolmogorov-Smirnov test for goodness-of-fit (D = 0.073, p = 1). Additionally a QQ-plot comparing empirical and fitted CDF had points lying approximately on the 1:1 line (see figure 1).

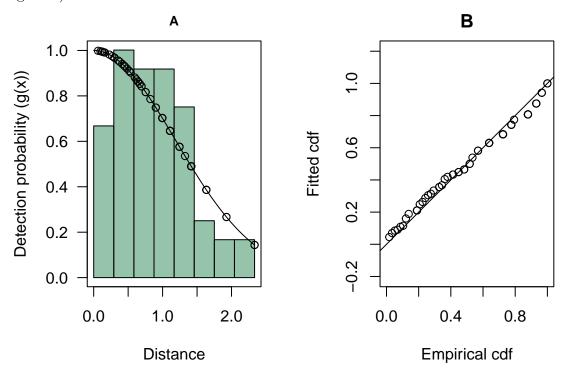


Figure 1: Half–Normal No Adjustment Model Fit. A is the model fit over the distribution of distances. B is the QQ–plot of the expected vs observed CDF

Results

Clusters

Individuals ### Discussion

References

Appendix

R Analysis Code

```
##### SETTING UP #####

# Set up packages
```

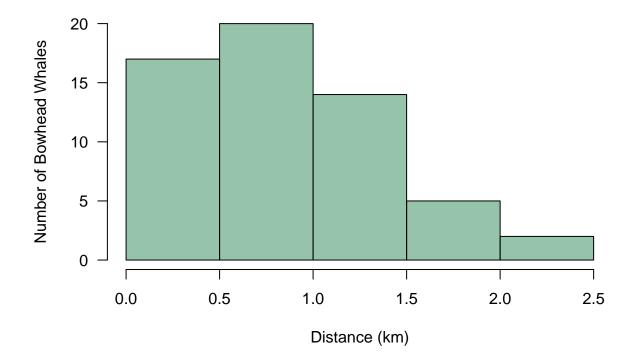
```
library(Distance)
library(tidyverse)
library(remotes)

# Getting Package Statsecol from github which contains the data
remotes::install_github("https://github.com/chrissuthy/statsecol")

# Loading in the data package
library(statsecol)
```

```
84 obs. of 9 variables:
## 'data.frame':
                      2 2 2 2 2 2 2 2 2 2 . . .
   $ Region.Label: int
   $ Area
                      : int
##
   $ Sample.Label: int
                      3 3 4 4 4 4 4 4 4 4 ...
                : num 98.8 98.8 181.4 181.4 181.4 ...
##
   $ Effort
   $ distance
                : num 0.996 1.108 0.128 0.357 0.301 ...
##
                      0.55 0.015 0.023 0.0025 0.57 0.51 0.34 0.091 0.029 0.021 ...
##
   $ у
                : num
##
   $ size
                : num 1 1 1 1 1 2 2 1 2 2 ...
   $ bf
                      0 0 0 0 0 0 0 0 0 0 ...
##
                : int
                      1 2 3 4 5 6 7 8 9 10 ...
##
   $ object
                : int
```

Distances of Bowhead Whales from Line Transects



```
#####
          MODELLING
                         #####
# Due to the long shoulder seen in the histogram, first fit a hazard rate;
# then fit a half-normal
# Fitting hazard-rate detection functions with all adjustment terms
# Hazard-rate with no adjustment
hr <- ds(data = bowhead LT,
         key = "hr",
         adjustment = NULL )
# Hazard-rate with a cosine adjustment
hr cos <- ds(data = bowhead LT,
            key = "hr",
             adjustment = "cos")
# Hazard-rate with a hermite polynomial adjustment
hr_herm <- ds(data = bowhead_LT,</pre>
              key = "hr",
              adjustment = "herm" )
# Hazard-rate with a simple polynomial adjustment
hr poly <- ds(data = bowhead LT,
              key = "hr",
              adjustment = "poly" )
# Fitting half-normal detection functions with all adjustment terms
# Half-normal with no adjustment
hn <- ds(data = bowhead LT,
         key = "hn",
         adjustment = NULL )
# Half-normal with a cosine adjustment
hn_cos <- ds(data = bowhead_LT,</pre>
             key = "hn",
             adjustment = "cos" )
# Half-normal with a hermite polynomial adjustment
hn_herm <- ds(data = bowhead_LT,</pre>
```

```
#####
          MODEL SELECTION
                               #####
# Comparing Models graphically by fit of detection probability over distance distribut
# Half-normal detection function over histogram plot
# Setting up plot size and margins
par(mfrow = c(2, 2),
    mar = c(5, 4, 4, 2) + 0.1)
# Plotting all 4 half-normal models
plot(hn, main = "no adjustment")
plot(hn cos, main = "cosine")
plot(hn_poly, main = "polynomial")
plot(hn_herm, main = "hermite polynomial")
title("Half-Normal Models",
      line = -1,
      # Naming title and setting location
      outer = TRUE)
# Hazard-rate Detection function over histogram plot
par(mfrow = c(2, 2),
    # Setting up plot size and margins
    mar = c(5, 4, 4, 2) + 0.1)
# Plotting all 4 hazard-rate models
plot(hr, main = "no adjustment")
plot(hr cos, main = "cosine")
plot(hr_poly, main = "polynomial")
plot(hr_herm, main = "hermite polynomial")
title("Hazard-Rate Models",
      line = -1,
      # Naming title and setting location
      outer = TRUE)
```

```
# Comparing AIC of all the models
summarize ds models(hn, hn cos, hn herm, hn poly,
                   hr, hr_cos, hr_herm, hr_poly,
                   output = "plain") # Setting Plain instead of equations
# As the dataset is small, checking whether AICc is a more appropriate measure
# To do this, Takezawa (2014) has said AICc should be used when the ratio of
# your parameters to number of data points is less than 1:40
# Print summary of hazard-rate distance model
summary(hr)
 # Parameters = 2 : Observations = 58
 # Therefore ratio is 1:29
      This ratio will be even smaller for the models with adjustments
# Print summary of half-normal distance model
summary(hn)
  # This is the only model that meets the assumptions of AIC however as you
    Can"t compare across model selection parameters we will use AICc
# Using AICc to select models
AICc(hn, hn_cos, hn_herm, hn_poly,
    hr, hr cos, hr herm, hr poly)
# Despite this there is no change in the best model
# The half-normal will be chosen as it has the smallest AIC and AICc
# No adjustment will be chosen as the adjustments don"t model extra variability
# in the data
```

```
##### MODEL FIT #####

# Comparing the detection function to the cramer-von mises test

# Setting plot size
par(mfrow = c(1,2), mar = c(5, 4, 4, 2) + 0.1)

# Plotting the detetion function over the distances distribution
plot(hn,
    which = 2,
    pl.col = adjustcolor("seagreen", 0.5),
    border = NULL,
```

```
ylab = "Detection probability (g(x))",
    xlab = "Distance",
    las = 1,
    main = "Half-normal Model No Adjustments")

# Plotting and running the Cramer-von Mises test and bootstrap Kolmogorov-Smirnov
# test for goodness-of-fit

gof_ds(hn, main = "Expected vs Observed CDF", ks = TRUE)

# The Cramer-von Mises test gives a test statistic of 0.0732325 and a p-value
# of 0.731882
# The Kolomogorov-Smirnov test gives a test statistic stat of 0.0725551 and a
# p-value of 1
# Therefore the model has a good fit as the p values are much more than 0.05
```

```
##### MODEL INFERENCE
                               #####
# Print summary of the model
summary(hn)
# Creating subsets of the column combinations needed for dht() density and
# abundance estimate and variances function.
# Region labels and areas.
region table <- unique(bowhead LT[, c("Region.Label",
                                      "Area")])
# Region and sample labels, effort
sample table <- unique(bowhead LT[, c("Region.Label",</pre>
                                      "Sample.Label",
                                      "Effort")])
# Object, region and sample labels
observation_table <- unique(bowhead_LT[, c("object",
                                           "Region.Label",
                                           "Sample.Label")])
# Finding the abundance and CI
abund bio hn <- dht(model = hn$ddf,
                    region_table,
                    sample table,
                    observation_table)
```

```
# Viewing the abundance and CI and getting the tables for plotting

# Estimated Abundance of individuals
N_ind <- abund_bio_hn$individuals$N

# Estimated Abundance of clusters
N_clu <- abund_bio_hn$clusters$N

# Estimated density of individuals
d_ind <- abund_bio_hn$individuals$D

# Estimated density of clusters
d_clu <- abund_bio_hn$clusters$D

# Average probability of detection for individuals
abund_bio_hn$individuals$average.p

# Average probability of detection for clusters
abund_bio_hn$clusters$average.p</pre>
```

```
# Changing the Region label to a factor
    # This allows it to be plotted in a desired order
fct N ind <- N ind %>%
 mutate(Label = as.factor(Label),
         Label = fct relevel(Label,
                             level = c('2', '3', '9',
                                       '11', '12', '15',
                                       'Total')))
ggplot(fct N ind)+
 geom_point(aes(x = Label, y = Estimate), stat="identity",
             colour = 'red')+
 geom linerange(aes(x = Label,
                     ymin = lcl,
                    ymax = ucl))+
 xlab('Regions')+
 ylab(')
```