1: Hybrid Gibbs Sampler to estimate Poisson Distribution λ (30%)

Derivation:

For these Poisson distributed random variables (r.v.s) (n = 500) with mean parameter λ , unobserved variables are $\lambda, y_1, y_2, \ldots, y_{78}$, in which y's denote the r.v.s which are larger than or equal to five.

Prior:
$$\pi(\lambda) \propto \frac{1}{\lambda}$$
;

$$P(X,Y|\lambda) = \prod_{i=1}^{422} \frac{e^{-\lambda} \lambda^{x_i}}{x_i!} \prod_{j=1}^{78} \frac{e^{-\lambda} \lambda^{y_i}}{y_i!} I(y_i \ge 5);$$

$$P(\lambda|X,Y) \propto P(X,Y|\lambda) P(\lambda) \propto e^{-n\lambda} \lambda^{\sum_i x_i + \sum_j y_j - 1}$$

$$P(y_j|X,\lambda) = \frac{e^{-\lambda} \lambda^{y_i}}{y_i!} I(y_i \ge 5) \propto \frac{\lambda^{y_i}}{y_i!} I(y_i \ge 5)$$

Then we can know that $\lambda | X, Y \propto Gamma(\sum_i x_i + \sum_j y_j, n)$.

The MH-Step to sample 78 unobserved y's is

$$y_{j}^{*} = \begin{cases} y_{j}^{(t)} - 1, & \text{with probability } \frac{1}{3}; \\ y_{j}^{(t)}, & \text{with probability } \frac{1}{3}; \\ y_{j}^{(t)} + 1, & \text{with probability } \frac{1}{3}. \end{cases}$$

$$r = \min \left\{ \frac{[\lambda^{(t+1)}]^{y_{j}^{*}} / y_{j}^{*}!}{[\lambda^{(t+1)}]^{y_{j}^{(t)}} / y_{i}^{(t)}!} I(y_{j}^{*} \geq 5), 1 \right\}$$

where r is the accept-reject ratio.

Result: $\hat{\lambda} = 2.803$.

2: Gibbs Sampler for Clustering (30%)

Derivation:

We use $\{X_{ij}\}_{i=1,2,3,\dots,1000}^{j=1,2,3}$ to denote the datum of *i*-th sample in *j*-th dimension. Then using the same notation in the question, the complete-data likelihood function is:

$$f(X, Z|\Pi, \Theta) = \prod_{i=1}^{1000} \prod_{k=1}^{3} \left[P(Z_i = k|\Pi, \Theta) P(X_{ij}, j = 1, 2, 3|Z_j, \Pi, \Theta) \right]^{I(Z_j = k)}$$

$$= \prod_{i=1}^{1000} \prod_{k=1}^{3} \left[P(Z_i = k|\Pi) \prod_{j=1}^{3} P(X_{ij}|Z_j, \Theta) \right]^{I(Z_j = k)}$$

$$= \prod_{i=1}^{1000} \prod_{k=1}^{3} \left[\pi_k \prod_{j=1}^{3} \binom{10j}{X_{ij}} \theta_{jk}^{X_{ij}} (1 - \theta_{jk})^{10j - X_{ij}} \right]^{I(Z_j = k)}$$

Note that given $Z_i = k$, for each sample i, $X_{ij} \sim Bino(10j, \theta_{jk})$; $P(\theta_{jk}) \propto Beta(a, b)$ (Prior: $P(\theta_{jk}) \propto Beta(1, 1) \propto 1$); and $(\pi_1, \pi_2, \pi_3) \sim Dirichlet(\alpha_1, \alpha_2, \alpha_3)$, we can derive by the following:

$$P(\Pi, \Theta, Z|X) \propto P(\Pi)P(\Theta)P(X, Z|\Pi, \Theta)$$

$$\propto \prod_{k=1}^{3} \pi_{k}^{\alpha_{k}-1} \prod_{j=1}^{3} \prod_{k=1}^{3} \theta_{jk}^{a-1} (1-\theta_{jk})^{b-1} \prod_{i=1}^{1000} \prod_{k=1}^{3} \left[\pi_{k} \prod_{j=1}^{3} \binom{10j}{X_{ij}} \theta_{jk}^{X_{ij}} (1-\theta_{jk})^{10j-X_{ij}} \right]^{I(Z_{j}=k)}$$

For π :

$$f(\Pi|\Theta,Z) \propto \prod_{k=1}^{3} \pi_k^{\alpha_k - 1} \prod_{i=1}^{1000} \prod_{k=1}^{3} \pi_k^{I(Z_i = k)}$$

$$\propto \prod_{k=1}^{3} \pi_k^{\alpha_k - 1} \prod_{k=1}^{3} [\pi_k]^{\sum_{i=1}^{1000} I(Z_i = k)}$$

$$\propto \prod_{k=1}^{3} \pi_k^{\alpha_k + \sum_{i=1}^{1000} I(Z_i = k) - 1}$$

Then it follows that:

$$\Pi|\Theta, Z \sim Dirichlet(\alpha_1 + \sum_{i=1}^{1000} I(Z_i = 1), \alpha_2 + \sum_{i=1}^{1000} I(Z_i = 2), \alpha_3 + \sum_{i=1}^{1000} I(Z_i = 3))$$

For θ :

$$\theta_{jk}|-\propto \prod_{j=1}^{3} \prod_{k=1}^{3} \theta_{jk}^{a-1} (1-\theta_{jk})^{b-1} \prod_{i=1}^{1000} \prod_{k=1}^{3} \left[\prod_{j=1}^{3} \binom{10j}{X_{ij}} \theta_{jk}^{X_{ij}} (1-\theta_{jk})^{10j-X_{ij}}\right]^{I(Z_{j}=k)}$$

$$\propto \theta_{jk}^{a-1} (1-\theta_{jk})^{b-1} \prod_{i=1}^{1000} \left[\theta_{jk}^{X_{ij}} (1-\theta_{jk})^{10j-X_{ij}}\right]^{I(Z_{j}=k)}$$

$$\propto \theta_{jk}^{\sum_{i=1}^{1000} X_{ij} I(Z_{i}=k) + a-1} (1-\theta_{jk})^{\sum_{i=1}^{1000} (10j-X_{ij}) I(Z_{i}=k) + b-1}$$

Then it follows that:

$$\theta_{jk}|-\sim Beta(a+\sum_{i=1}^{1000}X_{ij}I(Z_i=k),b+\sum_{i=1}^{1000}(10j-X_{ij})I(Z_i=k))$$

For Z:

$$\begin{split} Z_{i}|-&\propto\prod_{i=1}^{1000}\prod_{k=1}^{3}\pi_{k}^{I(Z_{i}=k)}\prod_{i=1}^{1000}\prod_{k=1}^{3}\prod_{j=1}^{3}\binom{10j}{X_{ij}}^{I(Z_{i}=k)}\theta_{jk}^{X_{ij}I(Z_{i}=k)}[(1-\theta_{jk})^{10j-X_{ij}}]^{I(Z_{i}=k)}\\ &\propto\prod_{k=1}^{3}\pi_{k}^{I(Z_{i}=k)}\prod_{k=1}^{3}\prod_{j=1}^{3}\binom{10j}{X_{ij}}^{I(Z_{i}=k)}\theta_{jk}^{X_{ij}I(Z_{i}=k)}[(1-\theta_{jk})^{10j-X_{ij}}]^{I(Z_{i}=k)}\\ &\propto\prod_{k=1}^{3}\left[\pi_{k}\left[\prod_{j=1}^{3}\binom{10j}{X_{ij}}\theta_{jk}^{X_{ij}}(1-\theta_{jk})^{10j-X_{ij}}\right]^{I(Z_{i}=k)}\right] \end{split}$$

Then for Gibbs Sampler algorithm:

Given $\Pi^{(t)}, \Theta^{(t)}, Z^{(t)}$, we update the parameters by the following:

$$\begin{split} \pi_1^{(t+1)}, \pi_2^{(t+1)}, \pi_3^{(t+1)}| - &\sim Dirichlet(\alpha_1 + \sum_{i=1}^{1000} I(Z_i^{(t)} = 1), \alpha_2 + \sum_{i=1}^{1000} I(Z_i^{(t)} = 2), \alpha_3 + \sum_{i=1}^{1000} I(Z_i^{(t)} = 3)) \\ \theta_{jk}^{(t+1)}| - &\sim Beta(a + \sum_{i=1}^{1000} X_{ij} I(Z_i^{(t)} = k), b + \sum_{i=1}^{1000} (10j - X_{ij}) I(Z_i^{(t)} = k)) \\ P(Z_i^{(t+1)} = k| -) = \frac{\pi_k^{(t+1)} \prod_{j=1}^3 \binom{10j}{X_{ij}} (\theta_{jk}^{(t+1)})^{X_{ij}} (1 - \theta_{jk}^{(t+1)})^{10j - X_{ij}}}{\sum_{l=1}^3 \pi_l^{(t+1)} \prod_{j=1}^3 \binom{10j}{X_{ij}} (\theta_{jl}^{(t+1)})^{X_{ij}} (1 - \theta_{jl}^{(t+1)})^{10j - X_{ij}}} \end{split}$$

Result:

Estimation of Π :

 $\hat{\pi}_1 = 0.499, \ \hat{\pi}_2 = 0.298, \ \text{and} \ \hat{\pi}_3 = 0.203.$

Estimation of Θ :

 $\theta_{11} = 0.807$, $\theta_{12} = 0.490$, $\theta_{13} = 0.196$, $\theta_{21} = 0.206$, $\theta_{22} = 0.804$, $\theta_{23} = 0.515$, $\theta_{31} = 0.502$, $\theta_{32} = 0.196$, and $\theta_{33} = 0.797$.

Estimation of Z:

Samples of cluster 1:

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> which(estimated_z==1)
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[55] 96 99 104 106 108 112 113 115 118 120 122 123 128 129 130 131 135 137 140 141 143 146 147 150 152 153 154
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[109] 212 213 215 216 221 224 225 226 232 233 236 237 238 239 240 241 242 243 245 251 252 253 258 260 262 263 267
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[163] 320 321 323 324 325 326 328 329 330 331 334 335 336 337 339 344 346 347 350 351 352 353 357 358 359 365 366
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Figure 1: Samples in cluster 1

Samples of cluster 2:

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> which(estimated_z==2)
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 [295]
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Figure 2: Samples in cluster 2

Samples of cluster 3:

```
> which(estimated_z==3)
[1] 1 9 22 30 35 41 53 59 62 64 77 83 85 86 87 95 101 102 103 110 111 117 125 132 138 149 158
[28] 167 168 177 178 179 190 193 201 207 208 211 214 218 222 228 244 246 248 250 255 256 257 259 261 265 266 287
[55] 288 297 299 303 306 314 327 333 338 340 341 343 355 356 364 377 380 384 385 386 390 392 398 399 400 403 408
[82] 410 421 426 442 452 458 459 464 465 466 467 468 469 470 476 477 480 482 484 493 503 505 508 509 510 514 518
[109] 519 522 524 533 538 544 547 551 556 559 565 566 569 570 572 574 576 582 583 589 603 607 611 616 621 627 628
[136] 630 631 635 636 644 645 648 653 654 655 662 671 678 685 691 694 698 702 722 723 729 731 732 736 740 747 751
[163] 756 765 767 768 769 772 776 778 781 789 794 801 804 807 813 819 830 847 851 858 859 860 877 891 899 900 901
[190] 909 916 923 927 938 950 956 958 959 971 980 983 986 991
```

Figure 3: Samples in cluster 3

Traceplots:

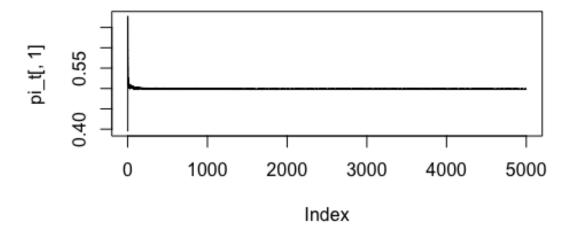


Figure 4: Traceplot of π_1

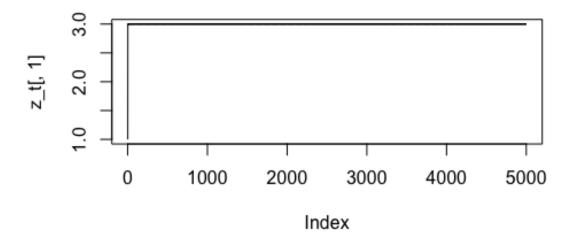


Figure 5: Traceplot of Z_1

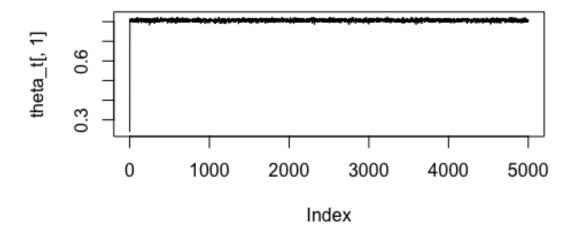


Figure 6: Traceplot of θ_{11}

3: Hybrid Gibbs Sampler (40%)

Note that for $i = 1, 2, \mathbf{Y}_i = (y_{i1}, y_{i2}, y_{i3}, y_{i4}) \sim multinomial(100, p_1, p_2, p_3, p_4)$, then:

$$\begin{aligned} \text{Prior}: \pi(\mathbf{p}) &\propto Dirichlet(\alpha_{1}, \alpha_{2}, \alpha_{3}, \alpha_{4}) \propto Dirichlet(2, 2, 2, 2) \propto p_{1}p_{2}p_{3}p_{4}; \\ P(y_{i1}, y_{i2}, y_{i3}, y_{i4} | p_{1}, p_{2}, p_{3}, p_{4}) &= \frac{100!}{y_{i1}! y_{i2}! y_{i3}! y_{i4}!} p_{1}^{y_{i1}} p_{2}^{y_{i2}} p_{3}^{y_{i3}} p_{4}^{y_{i4}} \\ P(p_{1}, p_{2}, p_{3}, p_{4} | y_{i1}, y_{i2}, y_{i3}, y_{i4}) &\propto P(y_{i1}, y_{i2}, y_{i3}, y_{i4} | p_{1}, p_{2}, p_{3}, p_{4}) f(p_{1}, p_{2}, p_{3}, p_{4} | \alpha_{1}, \alpha_{2}, \alpha_{3}, \alpha_{4}) \\ &\propto \frac{100!}{y_{i1}! y_{i2}! y_{i3}! y_{i4}!} p_{1}^{y_{i1} + \alpha_{1} - 1} p_{2}^{y_{i2} + \alpha_{2} - 1} p_{3}^{y_{i3} + \alpha_{3} - 1} p_{4}^{y_{i4} + \alpha_{4} - 1} \\ p_{1}, p_{2}, p_{3}, p_{4} | y_{i1}, y_{i2}, y_{i3}, y_{i4} &\sim Dirichlet(y_{i1} + \alpha_{1}, y_{i2} + \alpha_{2}, y_{i3} + \alpha_{3}, y_{i4} + \alpha_{4}) \\ &P(y_{i2} | \mathbf{Y}, \mathbf{P}) &\propto \frac{p_{2}^{y_{i2}}}{y_{i2}!} \end{aligned}$$

Then MH-Step to update y_{i2} (for i = 1, 2) is

$$\begin{split} y_{i2}^{(t+1)} &= \begin{cases} y_{i2}^{(t)} + 1, & \text{with probability } \frac{1}{2} \\ y_{i2}^{(t)} - 1, & \text{with probability } \frac{1}{2} \end{cases} \\ y_{i2}^{(t+1)} &= y_{i2}^{(t)} + 1, & \text{if } y_{i2}^{(t)} = 15 \\ y_{i2}^{(t+1)} &= y_{i2}^{(t)} - 1, & \text{if } y_{i2}^{(t)} = 32 \end{cases} \\ r &= \min \Big\{ \frac{p_{2}^{y_{i2}^{(t+1)}} / y_{i2}^{(t+1)}!}{p_{2}^{y_{i2}^{(t)}} / y_{i2}^{(t)}!}, 1 \Big\} \end{split}$$

where r is the accept-reject ratio. Then for another two unobserved variables:

$$y_{11}^{(t+1)} = 100 - y_{13} - y_{14} - y_{12}^{(t+1)} = 100 - 22 - 31 - y_{12}^{(t+1)}$$
$$y_{24}^{(t+1)} = 100 - y_{21} - y_{23} - y_{22}^{(t+1)} = 100 - 28 - 26 - y_{22}^{(t+1)}$$

Result:

$$p_1 = 0.301$$
, $p_2 = 0.149$, $p_3 = 0.239$, and $p_4 = 0.311$.