1: The Bisection Method (25%)

For function $f(x) = x^3 + 6x^2 + \pi x - 12$, the derivative is $f'(x) = 3x^2 + 12x + \pi$. The we can calculate that zeros of the derivative are $\frac{-12 - \sqrt{12(12 - \pi)}}{6}$ and $\frac{-12 + \sqrt{12(12 - \pi)}}{6}$.

 $f(\frac{-12-\sqrt{12(12-\pi)}}{6}) = 7.864841$ and $f(\frac{-12+\sqrt{12(12-\pi)}}{6}) = -12.43121$ Hence, the function f has totally 3 zeros.

Algorithm: Bisection Method in the R file.

Result: Zeros: -4.837944, -2.259727, and 1.097664.

2: Poisson Regression - Newton's Method (25%)

(1) Since $y_i \sim Poisson(\lambda_i)$ and $log(\lambda_i) = \alpha + \beta x_i + \gamma x_i^2$, we can get the Likelihhod function:

$$L(\alpha, \beta, \gamma | \mathbf{x}, \mathbf{y}) = \prod_{i=1}^{n} \frac{\lambda_i^{y_i} e^{-\lambda_i}}{y_i!} = \prod_{i=1}^{n} \frac{e^{(\alpha + \beta x_i + \gamma x_i^2) y_i} e^{-e^{\alpha + \beta x_i + \gamma x_i^2}}}{y_i!}$$

(2) The log-Likelihood function is

$$l(\alpha, \beta, \gamma | \mathbf{x}, \mathbf{y}) = \sum_{i=1}^{n} (\alpha + \beta x_i + \gamma x_i^2) y_i - e^{\alpha + \beta x_i + \gamma x_i^2} - \log y_i!$$

Then we have:

$$\frac{\partial l(\alpha, \beta, \gamma | \mathbf{x}, \mathbf{y})}{\partial \alpha} \Big|_{\hat{\alpha}} = \sum_{i=1}^{n} [y_i - e^{\alpha + \beta x_i + \gamma x_i^2}] \Big|_{\hat{\alpha}} = 0$$

$$\frac{\partial l(\alpha, \beta, \gamma | \mathbf{x}, \mathbf{y})}{\partial \beta} \Big|_{\hat{\beta}} = \sum_{i=1}^{n} [x_i y_i - x_i e^{\alpha + \beta x_i + \gamma x_i^2}] \Big|_{\hat{\beta}} = 0$$

$$\frac{\partial l(\alpha, \beta, \gamma | \mathbf{x}, \mathbf{y})}{\partial \gamma} \Big|_{\hat{\gamma}} = \sum_{i=1}^{n} [x_i^2 y_i - x_i^2 e^{\alpha + \beta x_i + \gamma x_i^2}] \Big|_{\hat{\gamma}} = 0$$

Let
$$\mathbf{F}(\mathbf{x}) = \begin{pmatrix} \frac{\partial l(\alpha, \beta, \gamma | \mathbf{x}, \mathbf{y})}{\partial \alpha} \\ \frac{\partial l(\alpha, \beta, \gamma | \mathbf{x}, \mathbf{y})}{\partial \beta} \\ \frac{\partial l(\alpha, \beta, \gamma | \mathbf{x}, \mathbf{y})}{\partial \gamma} \end{pmatrix}$$
, then $\mathbf{F}'(\mathbf{x}) = \begin{pmatrix} \frac{\partial^2 l(\alpha, \beta, \gamma | \mathbf{x}, \mathbf{y})}{\partial \alpha^2} & \frac{\partial^2 l(\alpha, \beta, \gamma | \mathbf{x}, \mathbf{y})}{\partial \beta \partial \alpha} & \frac{\partial^2 l(\alpha, \beta, \gamma | \mathbf{x}, \mathbf{y})}{\partial \gamma \partial \alpha} \\ \frac{\partial^2 l(\alpha, \beta, \gamma | \mathbf{x}, \mathbf{y})}{\partial \alpha \partial \beta} & \frac{\partial^2 l(\alpha, \beta, \gamma | \mathbf{x}, \mathbf{y})}{\partial \beta^2} & \frac{\partial^2 l(\alpha, \beta, \gamma | \mathbf{x}, \mathbf{y})}{\partial \gamma \partial \beta} \\ \frac{\partial^2 l(\alpha, \beta, \gamma | \mathbf{x}, \mathbf{y})}{\partial \alpha \partial \gamma} & \frac{\partial^2 l(\alpha, \beta, \gamma | \mathbf{x}, \mathbf{y})}{\partial \beta \partial \gamma} & \frac{\partial^2 l(\alpha, \beta, \gamma | \mathbf{x}, \mathbf{y})}{\partial \gamma^2} \end{pmatrix}$, in

which

$$\frac{\partial^{2}l(\alpha, \beta, \gamma | \mathbf{x}, \mathbf{y})}{\partial \alpha^{2}} = \sum_{i=1}^{n} -e^{\alpha + \beta x_{i} + \gamma x_{i}^{2}}$$

$$\frac{\partial^{2}l(\alpha, \beta, \gamma | \mathbf{x}, \mathbf{y})}{\partial \alpha \partial \beta} = \frac{\partial^{2}l(\alpha, \beta, \gamma | \mathbf{x}, \mathbf{y})}{\partial \beta \partial \alpha} = \sum_{i=1}^{n} -x_{i}e^{\alpha + \beta x_{i} + \gamma x_{i}^{2}}$$

$$\frac{\partial^{2}l(\alpha, \beta, \gamma | \mathbf{x}, \mathbf{y})}{\partial \alpha \partial \gamma} = \frac{\partial^{2}l(\alpha, \beta, \gamma | \mathbf{x}, \mathbf{y})}{\partial \gamma \partial \alpha} = \sum_{i=1}^{n} -x_{i}^{2}e^{\alpha + \beta x_{i} + \gamma x_{i}^{2}}$$

$$\frac{\partial^{2}l(\alpha, \beta, \gamma | \mathbf{x}, \mathbf{y})}{\partial \beta^{2}} = \sum_{i=1}^{n} -x_{i}^{2}e^{\alpha + \beta x_{i} + \gamma x_{i}^{2}}$$

$$\frac{\partial^{2}l(\alpha, \beta, \gamma | \mathbf{x}, \mathbf{y})}{\partial \beta \partial \gamma} = \frac{\partial^{2}l(\alpha, \beta, \gamma | \mathbf{x}, \mathbf{y})}{\partial \gamma \partial \beta} = \sum_{i=1}^{n} -x_{i}^{3}e^{\alpha + \beta x_{i} + \gamma x_{i}^{2}}$$

$$\frac{\partial^{2}l(\alpha, \beta, \gamma | \mathbf{x}, \mathbf{y})}{\partial \gamma^{2}} = \sum_{i=1}^{n} -x_{i}^{4}e^{\alpha + \beta x_{i} + \gamma x_{i}^{2}}$$

Therefore, by Newton's Methhod, given initial guess $\alpha^{(0)}$, $\beta^{(0)}$, and $\gamma^{(0)}$, for each iteration:

$$\begin{pmatrix} \alpha^{(n)} \\ \beta^{(n)} \\ \gamma^{(n)} \end{pmatrix} = \begin{pmatrix} \alpha^{(n-1)} \\ \beta^{(n-1)} \\ \gamma^{(n-1)} \end{pmatrix} - \mathbf{F}'[(\mathbf{x})]^{-1} \mathbf{F}(\mathbf{x}).$$

Hence, for the algorithm: $\mathbf{x}^{(\mathbf{n})} = \begin{pmatrix} \alpha^{(n)} \\ \beta^{(n)} \\ \gamma^{(n)} \end{pmatrix}$,

STEP 1: Solve $\mathbf{F}'(\mathbf{x}^{(n)}) \triangle \mathbf{x}^{(n)} = -\mathbf{F}(\mathbf{x}^{(n)});$

STEP 2: Update by $\mathbf{x^{(n+1)}} = \mathbf{x^{(n)}} + \triangle \mathbf{x^{(n)}}$

(3)

Algorithm: Newton's Method code in the R file.

Result: $\alpha = 1.503533$, $\beta = 1.052351$, and $\gamma = 1.957396$.

3: Logistic Regression - Newton's Method (20%)

(1) Since $y_i \sim Bernoulli(p_i)$ and $logit(p_i) = \alpha + \beta x_i$, we can know that $f(y_i, p_i) = p_i^{y_i} (1 - p_i)^{1 - y_i}$, and $p_i = \frac{e^{\alpha + \beta x_i}}{1 + e^{\alpha + \beta x_i}}$. Then we can get the following Likelihood function:

$$L(\alpha, \beta | \mathbf{x}, \mathbf{y}) = \prod_{i=1}^{n} p_i^{y_i} (1 - p_i)^{1 - y_i}$$
$$= \prod_{i=1}^{n} \left(\frac{e^{\alpha + \beta x_i}}{1 + e^{\alpha + \beta x_i}}\right)^{y_i} \left(1 - \frac{e^{\alpha + \beta x_i}}{1 + e^{\alpha + \beta x_i}}\right)^{1 - y_i}$$

(2) The log-Likelihood function is

$$l(\alpha, \beta | \mathbf{x}, \mathbf{y}) = \sum_{i=1}^{n} y_i (\alpha + \beta x_i - \log(1 + e^{\alpha + \beta x_i})) + (1 - y_i) \log(\frac{1}{1 + e^{\alpha + \beta x_i}})$$
$$= \sum_{i=1}^{n} \alpha x_i + \beta x_i y_i - \log(1 + e^{\alpha + \beta x_i})$$

Then we have:

$$\frac{\partial l(\alpha, \beta | \mathbf{x}, \mathbf{y})}{\partial \alpha} \Big|_{\hat{\alpha}} = \sum_{i=1}^{n} \left[y_i - \frac{e^{\alpha + \beta x_i}}{1 + e^{\alpha + \beta x_i}} \right] \Big|_{\hat{\alpha}} = 0$$
$$\frac{\partial l(\alpha, \beta | \mathbf{x}, \mathbf{y})}{\partial \beta} \Big|_{\hat{\beta}} = \sum_{i=1}^{n} \left[x_i y_i - \frac{x_i e^{\alpha + \beta x_i}}{1 + e^{\alpha + \beta x_i}} \right] \Big|_{\hat{\beta}} = 0$$

Let
$$\mathbf{F}(\mathbf{x}) = \begin{pmatrix} \frac{\partial l(\alpha, \beta | \mathbf{x}, \mathbf{y})}{\partial \alpha} \\ \frac{\partial l(\alpha, \beta | \mathbf{x}, \mathbf{y})}{\partial \beta} \end{pmatrix}$$
, then $\mathbf{F}'(\mathbf{x}) = \begin{pmatrix} \frac{\partial^2 l(\alpha, \beta | \mathbf{x}, \mathbf{y})}{\partial \alpha^2} & \frac{\partial^2 l(\alpha, \beta | \mathbf{x}, \mathbf{y})}{\partial \beta \partial \alpha} \\ \frac{\partial^2 l(\alpha, \beta | \mathbf{x}, \mathbf{y})}{\partial \alpha \partial \beta} & \frac{\partial^2 l(\alpha, \beta | \mathbf{x}, \mathbf{y})}{\partial \beta^2} \end{pmatrix}$, in which
$$\frac{\partial^2 l(\alpha, \beta | \mathbf{x}, \mathbf{y})}{\partial \alpha^2} = -\frac{e^{\alpha + \beta x_i}}{(1 + e^{\alpha + \beta x_i})^2}$$
$$\frac{\partial^2 l(\alpha, \beta | \mathbf{x}, \mathbf{y})}{\partial \beta \partial \alpha} = \frac{\partial^2 l(\alpha, \beta | \mathbf{x}, \mathbf{y})}{\partial \alpha \partial \beta} = -\frac{x_i e^{\alpha + \beta x_i}}{(1 + e^{\alpha + \beta x_i})^2}$$
$$\frac{\partial^2 l(\alpha, \beta | \mathbf{x}, \mathbf{y})}{\partial \beta^2} = -\frac{x_i^2 e^{\alpha + \beta x_i}}{(1 + e^{\alpha + \beta x_i})^2}$$

Therefore, by Newton's Methhod, given initial guess $\alpha^{(0)}$ and $\beta^{(0)}$, for each iteration:

$$\begin{pmatrix} \alpha^{(n)} \\ \beta^{(n)} \end{pmatrix} = \begin{pmatrix} \alpha^{(n-1)} \\ \beta^{(n-1)} \end{pmatrix} - \mathbf{F}'[(\mathbf{x})]^{-1}\mathbf{F}(\mathbf{x}).$$

Hence, for the algorithm: $\mathbf{x}^{(\mathbf{n})} = \begin{pmatrix} \alpha^{(n)} \\ \beta^{(n)} \end{pmatrix}$,

STEP 1: Solve $\mathbf{F}'(\mathbf{x}^{(n)}) \triangle \mathbf{x}^{(n)} = -\mathbf{F}(\mathbf{x}^{(n)});$

STEP 2: Update by $\mathbf{x}^{(\mathbf{n}+\mathbf{1})} = \mathbf{x}^{(\mathbf{n})} + \triangle \mathbf{x}^{(\mathbf{n})}$.

(3)

Algorithm: Newton's Method code in the R file.

Result: $\alpha = 1.564284$ and $\beta = 1.771093$.

4: EM Algorithm (30%)

(1) Observed data: Y_i for i = 1, 2, ..., 8000; Missing data: Z_i for i = 1, 2, ..., 8000., where $Z_i = 1, 2, or 3$ for low, middle, and high income respectively.

Since $Y_i|(Z_i = k) \sim N(\mu_k, \sigma_k^2)$, with proportion π_k ($\pi_3 = 1 - (\pi_1 + \pi_2)$), we can formulate the **complete** – **data Likelihhod function** as:

$$L(\pi_1, \pi_2, \mu_1, \mu_2, \mu_3, \sigma_1, \sigma_2, \sigma_3 | \mathbf{Y}, \mathbf{Z})$$

$$= \prod_{i=1}^{n} \left[\pi_1 \frac{1}{\sqrt{2\pi}\sigma_1} e^{-\frac{(y_i - \mu_1)^2}{2\sigma_1^2}} \right]^{I(Z_i = 1)} \left[\pi_2 \frac{1}{\sqrt{2\pi}\sigma_2} e^{-\frac{(y_i - \mu_2)^2}{2\sigma_2^2}} \right]^{I(Z_i = 2)} \left[(1 - \pi_1 - \pi_2) \frac{1}{\sqrt{2\pi}\sigma_3} e^{-\frac{(y_i - \mu_3)^2}{2\sigma_3^2}} \right]^{I(Z_i = 3)}$$

The observed-data Likelihhod function is:

$$L(\pi_{1}, \pi_{2}, \mu_{1}, \mu_{2}, \mu_{3}, \sigma_{1}, \sigma_{2}, \sigma_{3} | \mathbf{Y})$$

$$= \prod_{i=1}^{n} \left[\left[\pi_{1} \frac{1}{\sqrt{2\pi}\sigma_{1}} e^{-\frac{(y_{i} - \mu_{1})^{2}}{2\sigma_{1}^{2}}} \right] + \left[\pi_{2} \frac{1}{\sqrt{2\pi}\sigma_{2}} e^{-\frac{(y_{i} - \mu_{2})^{2}}{2\sigma_{2}^{2}}} \right] + \left[(1 - \pi_{1} - \pi_{2}) \frac{1}{\sqrt{2\pi}\sigma_{3}} e^{-\frac{(y_{i} - \mu_{3})^{2}}{2\sigma_{3}^{2}}} \right] \right]$$

(2) The complete-data log-Likelihhod function is:

$$l(\pi_1, \pi_2, \mu_1, \mu_2, \mu_3, \sigma_1, \sigma_2, \sigma_3 | \mathbf{Y}, \mathbf{Z}) = \sum_{i=1}^n \left[I(Z_i = 1) \left[\log \pi_1 - \log \sqrt{2\pi} - \frac{1}{2} \log \sigma_1^2 - \frac{(y_i - \mu_1)^2}{2\sigma_1^2} \right] + I(Z_i = 2) \left[\log \pi_2 - \log \sqrt{2\pi} - \frac{1}{2} \log \sigma_2^2 - \frac{(y_i - \mu_2)^2}{2\sigma_2^2} \right] + I(Z_i = 3) \left[\log(1 - \pi_1 - \pi_2) - \log\sqrt{2\pi} - \frac{1}{2} \log\sigma_3^2 - \frac{(y_i - \mu_3)^2}{2\sigma_3^2} \right] \right]$$

Then given initial guess $\pi_1^{(0)}, \pi_2^{(0)}, \mu_1^{(0)}, \mu_2^{(0)}, \mu_3^{(0)}, \sigma_1^{(0)}, \sigma_2^{(0)}, \sigma_3^{(0)}$, we define a Q function by $Q(\Theta; \Theta^{(t)}) = E_{\mathbf{\Theta}^{(t)}} (l(\mathbf{\Theta}|\mathbf{Y}, \mathbf{Z})|\mathbf{Y})$:

$$Q(\Theta; \Theta^{(t)}) = E_{\pi_1^{(t)}, \pi_2^{(t)}, \mu_1^{(t)}, \mu_2^{(t)}, \mu_3^{(t)}, \sigma_1^{2(t)}, \sigma_2^{2(t)}, \sigma_3^{2(t)}} \left(l(\pi_1, \pi_2, \mu_1, \mu_2, \mu_3, \sigma_1, \sigma_2, \sigma_3 | \mathbf{Y}, \mathbf{Z}) | \mathbf{Y} \right)$$

$$= \sum_{i=1}^{n} \left[I(Z_i^{(t)} = 1) \left[\log \pi_1 - \log \sqrt{2\pi} - \frac{1}{2} \log \sigma_1^2 - \frac{(y_i - \mu_1)^2}{2\sigma_1^2} \right] + I(Z_i^{(t)} = 2) \left[\log \pi_2 - \log \sqrt{2\pi} - \frac{1}{2} \log \sigma_2^2 - \frac{(y_i - \mu_2)^2}{2\sigma_2^2} \right] + I(Z_i^{(t)} = 3) \left[\log(1 - \pi_1 - \pi_2) - \log \sqrt{2\pi} - \frac{1}{2} \log \sigma_3^2 - \frac{(y_i - \mu_3)^2}{2\sigma_3^2} \right] \right]$$

where

$$\begin{split} Z_i^{(t)} &= E(Z_i^{(t)}|\Theta^{(t)}) = E(Z_i^{(t)}|\pi_1^{(t)},\pi_2^{(t)},\mu_1^{(t)},\mu_2^{(t)},\mu_3^{(t)},\sigma_1^{2(t)},\sigma_2^{2(t)},\sigma_3^{2(t)}) \\ &= p(Z_i^{(t)} = 1|\Theta^{(t)}) + 2p(Z_i^{(t)} = 2|\Theta^{(t)}) + 3p(Z_i^{(t)} = 3|\Theta^{(t)}) \\ &= 1 + \frac{\pi_2^{(t)} \frac{1}{\sqrt{2\pi}\sigma_2^{(t)}} e^{-\frac{(y_i - \mu_2^{(t)})^2}{2\sigma_2^2(t)}} + 2(1 - \pi_1^{(t)} - \pi_2^{(t)}) \frac{1}{\sqrt{2\pi}\sigma_3^{(t)}} e^{-\frac{(y_i - \mu_3^{(t)})^2}{2\sigma_3^2(t)}} \\ &= 1 + \frac{\pi_2^{(t)} \frac{1}{\sqrt{2\pi}\sigma_2^{(t)}} e^{-\frac{(y_i - \mu_2^{(t)})^2}{2\sigma_2^2(t)}} + 2(1 - \pi_1^{(t)} - \pi_2^{(t)}) \frac{1}{\sqrt{2\pi}\sigma_3^{(t)}} e^{-\frac{(y_i - \mu_3^{(t)})^2}{2\sigma_3^2(t)}} \\ &= \frac{\pi_2^{(t)} \frac{1}{\sqrt{2\pi}\sigma_1^{(t)}} e^{-\frac{(y_i - \mu_2^{(t)})^2}{2\sigma_2^2(t)}} + \pi_2^{(t)} \frac{1}{\sqrt{2\pi}\sigma_2^{(t)}} e^{-\frac{(y_i - \mu_2^{(t)})^2}{2\sigma_2^2(t)}} + (1 - \pi_1^{(t)} - \pi_2^{(t)}) \frac{1}{\sqrt{2\pi}\sigma_3^{(t)}} e^{-\frac{(y_i - \mu_3^{(t)})^2}{2\sigma_3^2(t)}} \end{split}$$

Then we can calculate:

$$\begin{split} \frac{\partial Q}{\partial \pi_1}\bigg|_{\pi_1^{(t+1)},\pi_2^{(t+1)}} &= \sum_{i=1}^n \left[I(Z_i^{(t)}=1)\frac{1}{\pi_1} - I(Z_i^{(t)}=3)\frac{1}{1-\pi_1-\pi_2}\right]\bigg|_{\pi_1^{(t+1)},\pi_2^{(t+1)}} = 0 \\ \frac{\partial Q}{\partial \pi_2}\bigg|_{\pi_1^{(t+1)},\pi_2^{(t+1)}} &= \sum_{i=1}^n \left[I(Z_i^{(t)}=2)\frac{1}{\pi_2} - I(Z_i^{(t)}=3)\frac{1}{1-\pi_1-\pi_2}\right]\bigg|_{\pi_1^{(t+1)},\pi_2^{(t+1)}} = 0 \\ \frac{\partial Q}{\partial \mu_1}\bigg|_{\mu_1^{(t+1)},\sigma_1^{(t+1)}} &= \sum_{i=1}^n \left[I(Z_i^{(t)}=1)\frac{y_i-\mu_1}{\sigma_1^2}\right]\bigg|_{\mu_1^{(t+1)},\sigma_1^{(t+1)}} = 0 \\ \frac{\partial Q}{\partial \mu_2}\bigg|_{\mu_2^{(t+1)},\sigma_2^{(t+1)}} &= \sum_{i=1}^n \left[I(Z_i^{(t)}=2)\frac{y_i-\mu_2}{\sigma_2^2}\right]\bigg|_{\mu_2^{(t+1)},\sigma_2^{(t+1)}} = 0 \\ \frac{\partial Q}{\partial \mu_3}\bigg|_{\mu_3^{(t+1)},\sigma_3^{(t+1)}} &= \sum_{i=1}^n \left[I(Z_i^{(t)}=3)\frac{y_i-\mu_3}{\sigma_3^2}\right]\bigg|_{\mu_3^{(t+1)},\sigma_3^{(t+1)}} = 0 \\ \frac{\partial Q}{\partial \sigma_1^2}\bigg|_{\mu_1^{(t+1)},\sigma_1^{(t+1)}} &= \sum_{i=1}^n \left[I(Z_i^{(t)}=1)\left(-\frac{1}{2}\frac{1}{\sigma_1^2} + \frac{(y_i-\mu_1)^2}{2(\sigma_1^2)^2}\right)\right]\bigg|_{\mu_1^{(t+1)},\sigma_1^{(t+1)}} = 0 \\ \frac{\partial Q}{\partial \sigma_2^2}\bigg|_{\mu_2^{(t+1)},\sigma_2^{(t+1)}} &= \sum_{i=1}^n \left[I(Z_i^{(t)}=2)\left(-\frac{1}{2}\frac{1}{\sigma_2^2} + \frac{(y_i-\mu_2)^2}{2(\sigma_2^2)^2}\right)\right]\bigg|_{\mu_2^{(t+1)},\sigma_2^{(t+1)}} = 0 \\ \frac{\partial Q}{\partial \sigma_3^2}\bigg|_{\mu_3^{(t+1)},\sigma_3^{(t+1)}} &= \sum_{i=1}^n \left[I(Z_i^{(t)}=3)\left(-\frac{1}{2}\frac{1}{\sigma_3^2} + \frac{(y_i-\mu_3)^2}{2(\sigma_3^2)^2}\right)\right]\bigg|_{\mu_3^{(t+1)},\sigma_3^{(t+1)}} = 0 \end{split}$$

Hence, we can conclude that:

$$\pi_1^{(t+1)} = \frac{\sum_{i=1}^n I(Z_i^{(t)} = 1)}{\sum_{i=1}^n I(Z_i^{(t)} = 1) + \sum_{i=1}^n I(Z_i^{(t)} = 2) + \sum_{i=1}^n I(Z_i^{(t)} = 3)}$$
(1)

$$\pi_2^{(t+1)} = \frac{\sum_{i=1}^n I(Z_i^{(t)} = 2)}{\sum_{i=1}^n I(Z_i^{(t)} = 1) + \sum_{i=1}^n I(Z_i^{(t)} = 2) + \sum_{i=1}^n I(Z_i^{(t)} = 3)}$$
(2)

$$\mu_1^{(t+1)} = \frac{\sum_{i=1}^n I(Z_i^{(t)} = 1) y_i}{\sum_{i=1}^n I(Z_i^{(t)} = 1)}$$
(3)

$$\mu_2^{(t+1)} = \frac{\sum_{i=1}^n I(Z_i^{(t)} = 2)y_i}{\sum_{i=1}^n I(Z_i^{(t)} = 2)}$$
(4)

$$\mu_3^{(t+1)} = \frac{\sum_{i=1}^n I(Z_i^{(t)} = 3) y_i}{\sum_{i=1}^n I(Z_i^{(t)} = 3)}$$
(5)

$$\sigma_1^{2(t+1)} = \frac{\sum_{i=1}^n I(Z_i^{(t)} = 1)(y_i - \mu_1^{(t+1)})^2}{\sum_{i=1}^n I(Z_i^{(t)} = 1)}$$
(6)

$$\sigma_2^{2(t+1)} = \frac{\sum_{i=1}^n I(Z_i^{(t)} = 2)(y_i - \mu_2^{(t+1)})^2}{\sum_{i=1}^n I(Z_i^{(t)} = 2)}$$
(7)

$$\sigma_3^{2(t+1)} = \frac{\sum_{i=1}^n I(Z_i^{(t)} = 3)(y_i - \mu_3^{(t+1)})^2}{\sum_{i=1}^n I(Z_i^{(t)} = 3)}$$
(8)

Then for iteration in **EM Algorithm**:

Given initial guess: $\pi_1^{(0)}, \pi_2^{(0)}, \mu_1^{(0)}, \mu_2^{(0)}, \mu_3^{(0)}, \sigma_1^{(0)}, \sigma_2^{(0)}, \sigma_3^{(0)}$, for $t \geq 0$ and $t \in \mathbb{Z}$:

 $\mathbf{E} - \mathbf{step} \text{: Calculate } E(Z_i^{(t)}|\Theta^{(t)}), \text{ where } \Theta^{(t)} = \pi_1^{(t)}, \pi_2^{(t)}, \mu_1^{(t)}, \mu_2^{(t)}, \mu_3^{(t)}, \sigma_1^{2(t)}, \sigma_2^{2(t)}, \sigma_3^{2(t)}.$

 $\mathbf{M} - \mathbf{step}$: Update $\Theta^{(t+1)}$ by equations (1) to (8) that we have obtained before.

(3)

Algorithm: EM Algorithm code in the R file.

Result:

(4)