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**1: Inverse method for Poisson Distribution (25%)**


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For discrete Poisson Distribution ( $\lambda = 5$ ),

the p.m.f is  $P(x|\lambda) = e^{-\lambda} \frac{\lambda^x}{x!}$  and the c.d.f is  $F(x|\lambda) = \sum_{t \leq x} e^{-\lambda} \frac{\lambda^t}{t!}$ .

**Algorithm:** Inverse method for the Poisson Distribution:

To generate  $X \sim F(x)$ :

STEP 1: Generate  $U \sim \text{unif}[0, 1]$ ;

STEP 2: Transform  $X = F^{-1}(U)$ : if  $F(x|\lambda) < U \leq F(x+1|\lambda)$ , let  $X = x+1$ .

**Plot :**

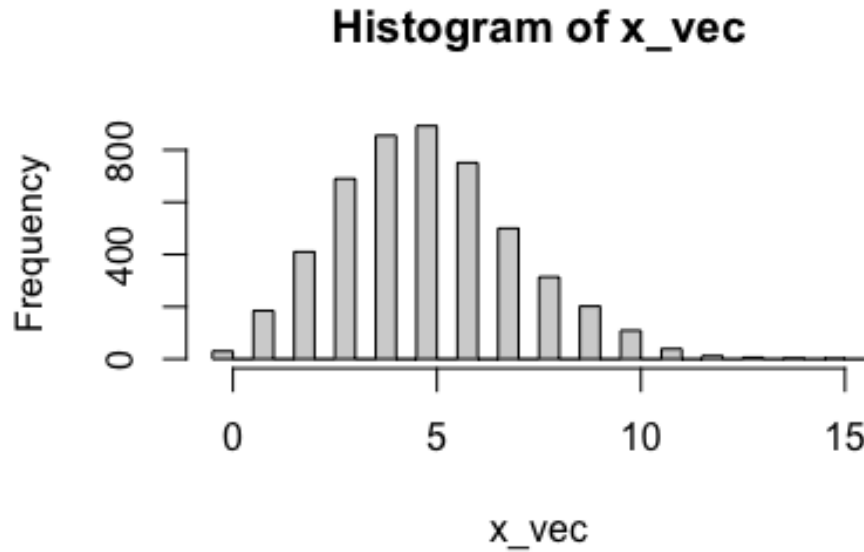


Figure 1: Histogram of 5000 samples

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**2: Accept-Reject method for truncated Gamma Distribution (25%)**


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For  $X \sim \text{Gamma}(\frac{1}{2}, 1)I(x \geq 5)$ ,  $f(x) = \frac{x^{-\frac{1}{2}} e^{-x} I(x \geq 5)}{\int_5^{+\infty} y^{-\frac{1}{2}} e^{-y} dy}$ .

We can define a shifted exponential distribution  $g(x) = e^{-(x-5)} I(x \geq 5)$  and want to find a constant  $M$  such that  $f(x) < Mg(x)$  for any  $x$ .

Then  $M = \sup \frac{f(x)}{g(x)} = \sup \frac{\frac{x^{-\frac{1}{2}} e^{-x} I(x \geq 5)}{\int_5^{+\infty} y^{-\frac{1}{2}} e^{-y} dy}}{e^{-(x-5)} I(x \geq 5)} = \frac{5^{-\frac{1}{2}} e^{-5}}{\int_5^{+\infty} y^{-\frac{1}{2}} e^{-y} dy}$ .

**Algorithm:** Accept-Reject method for truncated Gamma Distribution:

To generate  $X \sim F(x) = \text{c.d.f of } f(x)$ :

STEP 1: Generate  $Y \sim g(y)$ ;

STEP 2: Generate  $U \sim \text{unif}[0, 1]$ ;

STEP 3: Accept  $X = Y$  if  $U \leq \frac{f(Y)}{Mg(Y)}$ .

**Proof :**

From the choice of constant  $M$ , we can know that  $Mg(x) \geq f(x)$ . The goal of this method is to generate  $X \sim F(x) = \text{c.d.f of } f(x)$ .

For the generating algorithm:

$$\begin{aligned}
 P(X \leq x) &= P(Y \leq x | Y \text{ is accepted}) \\
 &= P(Y \leq x | U \leq \frac{f(Y)}{Mg(Y)}) \\
 &= \frac{P(Y \leq x, U \leq \frac{f(Y)}{Mg(Y)})}{P(U \leq \frac{f(Y)}{Mg(Y)})} \\
 &= \frac{\int_{-\infty}^x g(y) \int_0^{\frac{f(y)}{Mg(y)}} 1 dudy}{\int_{-\infty}^{+\infty} g(y) \int_0^{\frac{f(y)}{Mg(y)}} 1 dudy} \\
 &= \frac{\int_{-\infty}^x g(y) \frac{f(y)}{Mg(y)} dy}{\int_{-\infty}^{+\infty} g(y) \frac{f(y)}{Mg(y)} dy} \\
 &= \frac{\int_{-\infty}^x f(y) dy}{\int_{-\infty}^{+\infty} f(y) dy} \\
 &= \frac{\int_{-\infty}^x f(y) dy}{1} \\
 &= F(x)
 \end{aligned}$$

Therefore, this AR method works.

**Comparison :**

Theoretical acceptance probability:

$$\begin{aligned}
 P(U \leq \frac{f(Y)}{Mg(Y)}) &= \int_{-\infty}^{+\infty} g(y) \int_0^{\frac{f(y)}{Mg(y)}} 1 dudy \\
 &= \frac{1}{M} \int_{-\infty}^{+\infty} f(y) dy \\
 &= \frac{1}{M}
 \end{aligned}$$

After computation, this acceptance probability is 0.9207851.

The actual acceptance rate is 0.918, which is a little bit lower than the theoretical value.

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**3: Importance Sampling for Estimation (25%)**


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(1)

Using 5000 samples from Q2 ( $l$  = length of samples obtained in Q2), the Monte Carlo estimate is

$$\begin{aligned}
 \int_5^{+\infty} \cos(x) x^{-\frac{1}{2}} e^{-x} dx &= \int_{-\infty}^{+\infty} \cos(x) \frac{x^{-\frac{1}{2}} e^{-x} I(x \geq 5)}{\int_5^{+\infty} y^{-\frac{1}{2}} e^{-y} dy} dx \times \int_5^{+\infty} y^{-\frac{1}{2}} e^{-y} dy \\
 &= \frac{\int_5^{+\infty} y^{-\frac{1}{2}} e^{-y} dy}{l} \sum_{i=1}^l \cos(x_i) \\
 &= 0.001708
 \end{aligned}$$

(2)

Using the same notations in Q2, we define  $h(x) = \cos(x)$ .

$$\begin{aligned}
 \int_5^{+\infty} \cos(x) x^{-\frac{1}{2}} e^{-x} dx &= \int_{-\infty}^{+\infty} \cos(x) \frac{x^{-\frac{1}{2}} e^{-x} I(x \geq 5)}{\int_5^{+\infty} y^{-\frac{1}{2}} e^{-y} dy} dx \times \int_5^{+\infty} y^{-\frac{1}{2}} e^{-y} dy \\
 &= \int_5^{+\infty} y^{-\frac{1}{2}} e^{-y} dy \int_{-\infty}^{+\infty} \frac{h(x)f(x)}{g(x)} dx
 \end{aligned}$$

Note that  $\frac{f(x)}{g(x)} \leq M < \infty$ , and  $E_g h^2(x) = \int_{-\infty}^{+\infty} g(x) h^2(x) dx \leq \int_{-\infty}^{+\infty} g(x) dx = 1 < \infty$ , we can use the importance sampling as follows:

**Algorithm:**

STEP 1: Generate  $n = 5000$  samples from  $g(x)$ ;

STEP 2: Compute the Monte Carlo estimate:

$$\int_{-\infty}^{+\infty} \frac{h(x)f(x)}{g(x)} dx = \frac{\sum_{i=1}^n \frac{h(x_i)f(x_i)}{g(x_i)}}{n} = \frac{e^{-5}}{n \int_5^{+\infty} y^{-\frac{1}{2}} e^{-y} dy} \sum_{i=1}^n \cos(x_i) x_i^{-\frac{1}{2}}$$

Therefore,

$$\begin{aligned}
 \int_5^{+\infty} \cos(x) x^{-\frac{1}{2}} e^{-x} dx &= \int_5^{+\infty} y^{-\frac{1}{2}} e^{-y} dy \int_{-\infty}^{+\infty} \frac{h(x)f(x)}{g(x)} dx \\
 &= \int_5^{+\infty} y^{-\frac{1}{2}} e^{-y} dy \frac{e^{-5}}{n \int_5^{+\infty} y^{-\frac{1}{2}} e^{-y} dy} \sum_{i=1}^n \cos(x_i) x_i^{-\frac{1}{2}} \\
 &= \frac{e^{-5}}{n} \sum_{i=1}^n \cos(x_i) x_i^{-\frac{1}{2}} \\
 &= 0.00174
 \end{aligned}$$

#### 4: Stratified Sampling (25%)

(1)

Randomly draw 100 samples from the data set, for each subpopulation:

Standard deviation for age interval 1 is 73.9823615320924;

Standard deviation for age interval 2 is 111.975202857371;

Standard deviation for age interval 3 is 204.753943423479.

(2)

The target of this sampling is to estimate the mean salary of this country.

Using Stratified Sampling, there're 3 strata in total, which are indicated by 1, 2, and 3 respectively. Define  $n_i$  = the sample number of the  $i$  strata,  $x_{ij}$  = the sampled  $j$ -th individual in strata  $i$ ,  $\mu_i$  = the  $i$  strata's proportion in population, and  $S_i$  = the strata  $i$ .

The target is

$$E(\vec{x}) = \sum_{i=1}^3 p(x_i \in S_i) E(x|x \in S_i) = \sum_{i=1}^3 \mu_i \frac{1}{n_i} \sum_{j=1}^{n_i} x_{ij}.$$

It is supposed that the  $n_i$ 's can minimize  $Var(E(\vec{x}))$ .

$$Var(E(\vec{x})) = \sum_{i=1}^3 \mu_i^2 Var\left(\frac{1}{n_i} \sum_{j=1}^{n_i} x_{ij}\right) = \sum_{i=1}^3 \mu_i^2 \frac{1}{n_i} Var(x|x \in S_i)$$

Denote  $Var(x|x \in S_i)$  by  $V_i$ , which can be estimated by variance we obtained in the last step.

Then the minimization problem can be modeled as:

$$\begin{aligned} & \text{minimize} && \sum_{i=1}^3 \frac{\mu_i^2 V_i}{n_i} \\ & \text{subject to} && \sum_{i=1}^3 n_i = 1000 \end{aligned}$$

By Lagrange Multiplier,  $n_i = \frac{1000\mu_i\sqrt{V_i}}{\sum_{k=1}^3 \mu_k\sqrt{V_k}}$ . We use standard deviations estimated in part a to compute variance here.

After computation,

$$n_1 = 68$$

$$n_2 = 307$$

$$n_3 = 625$$

(3)

By randomly drawing  $n_i$  samples in strata  $i$  for each  $i$ , the estimated population mean salary based on these 1000 samples is 4607.275, which is a little bit higher than the true mean salary 4307.189 but is very close. The Stratified Sampling can somehow reflect the real mean salary.