1: Hybrid Gibbs Sampler to estimate Poisson Distribution λ (30%)

Derivation:

For these Poisson distributed random variables (r.v.s) (n = 500) with mean parameter λ , unobserved variables are $\lambda, y_1, y_2, \ldots, y_{78}$, in which y's denote the r.v.s which are larger than or equal to five.

Prior:
$$\pi(\lambda) \propto \frac{1}{\lambda}$$
;

$$P(X,Y|\lambda) = \prod_{i=1}^{422} \frac{e^{-\lambda} \lambda^{x_i}}{x_i!} \prod_{j=1}^{78} \frac{e^{-\lambda} \lambda^{y_i}}{y_i!} I(y_i \ge 5);$$

$$P(\lambda|X,Y) \propto P(X,Y|\lambda) P(\lambda) \propto e^{-n\lambda} \lambda^{\sum_i x_i + \sum_j y_j - 1}$$

$$P(y_j|X,\lambda) = \frac{e^{-\lambda} \lambda^{y_i}}{y_i!} I(y_i \ge 5) \propto \frac{\lambda^{y_i}}{y_i!} I(y_i \ge 5)$$

Then we can know that $\lambda | X, Y \propto Gamma(\sum_i x_i + \sum_j y_j, n)$.

The MH-Step to sample 78 unobserved y's is

$$y_{j}^{*} = \begin{cases} y_{j}^{(t)} - 1, & \text{with probability } \frac{1}{3}; \\ y_{j}^{(t)}, & \text{with probability } \frac{1}{3}; \\ y_{j}^{(t)} + 1, & \text{with probability } \frac{1}{3}. \end{cases}$$

$$r = \min \left\{ \frac{[\lambda^{(t+1)}]^{y_{j}^{*}} / y_{j}^{*}!}{[\lambda^{(t+1)}]^{y_{j}^{(t)}} / y_{i}^{(t)}!} I(y_{j}^{*} \geq 5), 1 \right\}$$

where r is the accept-reject ratio.

Result: $\hat{\lambda} = 1.674$.

2: Gibbs Sampler for Clustering (30%)

Derivation:

We use $\{X_{ij}\}_{i=1,2,3,\dots,1000}^{j=1,2,3}$ to denote the datum of *i*-th sample in *j*-th dimension. Then using the same notation in the question, the complete-data likelihood function is:

$$f(X, Z|\Pi, \Theta) = \prod_{i=1}^{1000} \prod_{k=1}^{3} \left[P(Z_i = k|\Pi, \Theta) P(X_{ij}, j = 1, 2, 3|Z_j, \Pi, \Theta) \right]^{I(Z_j = k)}$$

$$= \prod_{i=1}^{1000} \prod_{k=1}^{3} \left[P(Z_i = k|\Pi) \prod_{j=1}^{3} P(X_{ij}|Z_j, \Theta) \right]^{I(Z_j = k)}$$

$$= \prod_{i=1}^{1000} \prod_{k=1}^{3} \left[\pi_k \prod_{j=1}^{3} \binom{10j}{X_{ij}} \theta_{jk}^{X_{ij}} (1 - \theta_{jk})^{10j - X_{ij}} \right]^{I(Z_j = k)}$$

Note that given $Z_i = k$, for each sample i, $X_{ij} \sim Bino(10j, \theta_{jk})$; $P(\theta_{jk}) \propto 1$; and $(\pi_1, \pi_2, \pi_3) \sim Dirichlet(\alpha_1, \alpha_2, \alpha_3)$, we can derive by the following:

$$\begin{split} P(\Pi,\Theta,Z|X) &\propto P(\Pi)P(\Theta)P(X,Z|\Pi,\Theta) \\ &\propto \prod_{k=1}^{3} \pi_{k}^{\alpha_{k}-1} \prod_{i=1}^{1000} \prod_{k=1}^{3} \left[\pi_{k} \prod_{j=1}^{3} \binom{10j}{X_{ij}} \theta_{jk}^{X_{ij}} (1-\theta_{jk})^{10j-X_{ij}}\right]^{I(Z_{j}=k)} \\ &\propto \prod_{k=1}^{3} \pi_{k}^{\alpha_{k}-1} \prod_{i=1}^{1000} \prod_{k=1}^{3} \pi_{k}^{I(Z_{i}=k)} \prod_{i=1}^{1000} \prod_{k=1}^{3} \prod_{j=1}^{3} \binom{10j}{X_{ij}}^{I(Z_{i}=k)} \theta_{jk}^{X_{ij}I(Z_{i}=k)} [(1-\theta_{jk})^{10j-X_{ij}}]^{I(Z_{i}=k)} \end{split}$$

For π :

$$f(\Pi|\Theta, Z) \propto \prod_{k=1}^{3} \pi_k^{\alpha_k - 1} \prod_{i=1}^{1000} \prod_{k=1}^{3} \pi_k^{I(Z_i = k)}$$

$$\propto \prod_{k=1}^{3} \pi_k^{\alpha_k - 1} \prod_{k=1}^{3} [\pi_k]^{\sum_{i=1}^{1000} I(Z_i = k)}$$

$$\propto \prod_{k=1}^{3} \pi_k^{\alpha_k + \sum_{i=1}^{1000} I(Z_i = k) - 1}$$

Then it follows that:

$$\Pi|\Theta, Z \sim Dirichlet(\alpha_1 + \sum_{i=1}^{1000} I(Z_i = 1), \alpha_2 + \sum_{i=1}^{1000} I(Z_i = 2), \alpha_3 + \sum_{i=1}^{1000} I(Z_i = 3))$$

For θ :

$$\theta_{jk}|-\propto \prod_{i=1}^{1000} \prod_{k=1}^{3} \left[\prod_{j=1}^{3} \binom{10j}{X_{ij}} \theta_{jk}^{X_{ij}} (1-\theta_{jk})^{10j-X_{ij}}\right]^{I(Z_{j}=k)}$$

$$\propto \prod_{i=1}^{1000} \left[\theta_{jk}^{X_{ij}} (1-\theta_{jk})^{10j-X_{ij}}\right]^{I(Z_{j}=k)}$$

$$\propto \theta_{jk}^{\sum_{i=1}^{1000} X_{ij} I(Z_{i}=k)} (1-\theta_{jk})^{\sum_{i=1}^{1000} (10j-X_{ij}) I(Z_{i}=k)}$$

Then it follows that:

$$\theta_{jk}|-\sim Beta(1+\sum_{i=1}^{1000}X_{ij}I(Z_i=k),1+\sum_{i=1}^{1000}(10j-X_{ij})I(Z_i=k))$$

For Z:

$$Z_{i}|-\propto \prod_{i=1}^{1000} \prod_{k=1}^{3} \pi_{k}^{I(Z_{i}=k)} \prod_{i=1}^{1000} \prod_{k=1}^{3} \prod_{j=1}^{3} \binom{10j}{X_{ij}}^{I(Z_{i}=k)} \theta_{jk}^{X_{ij}I(Z_{i}=k)} [(1-\theta_{jk})^{10j-X_{ij}}]^{I(Z_{i}=k)}$$

$$\propto \prod_{k=1}^{3} \pi_{k}^{I(Z_{i}=k)} \prod_{k=1}^{3} \prod_{j=1}^{3} \binom{10j}{X_{ij}}^{I(Z_{i}=k)} \theta_{jk}^{X_{ij}I(Z_{i}=k)} [(1-\theta_{jk})^{10j-X_{ij}}]^{I(Z_{i}=k)}$$

$$\propto \prod_{k=1}^{3} \left[\pi_{k} \left[\prod_{j=1}^{3} \binom{10j}{X_{ij}} \theta_{jk}^{X_{ij}} (1-\theta_{jk})^{10j-X_{ij}} \right]^{I(Z_{i}=k)} \right]^{I(Z_{i}=k)}$$

Then for Gibbs Sampler algorithm:

Given $\Pi^{(t)}, \Theta^{(t)}, Z^{(t)}$, we update the parameters by the following:

$$\pi_{1}^{(t+1)}, \pi_{2}^{(t+1)}, \pi_{3}^{(t+1)}| - \sim Dirichlet(\alpha_{1} + \sum_{i=1}^{1000} I(Z_{i}^{(t)} = 1), \alpha_{2} + \sum_{i=1}^{1000} I(Z_{i}^{(t)} = 2), \alpha_{3} + \sum_{i=1}^{1000} I(Z_{i}^{(t)} = 3))$$

$$\theta_{jk}^{(t+1)}| - \sim Beta(1 + \sum_{i=1}^{1000} X_{ij}I(Z_{i}^{(t)} = k), 1 + \sum_{i=1}^{1000} (10j - X_{ij})I(Z_{i}^{(t)} = k))$$

$$P(Z_{i}^{(t+1)} = k| -) = \frac{\pi_{k}^{(t+1)} \prod_{j=1}^{3} \binom{10j}{X_{ij}} (\theta_{jk}^{(t+1)})^{X_{ij}} (1 - \theta_{jk}^{(t+1)})^{10j - X_{ij}}}{\sum_{l=1}^{3} \pi_{l}^{(t+1)} \prod_{j=1}^{3} \binom{10j}{X_{ij}} (\theta_{jl}^{(t+1)})^{X_{ij}} (1 - \theta_{jl}^{(t+1)})^{10j - X_{ij}}}$$

Assignment3

Result:

3: Hybrid Gibbs Sampler (40%)

Note that for $i = 1, 2, \mathbf{Y}_i = (y_{i1}, y_{i2}, y_{i3}, y_{i4}) \sim multinomial(100, p_1, p_2, p_3, p_4)$, then:

Prior : $\pi(\mathbf{p}) \propto Dirichlet(\alpha_1, \alpha_2, \alpha_3, \alpha_4) \propto Dirichlet(2, 2, 2, 2) \propto p_1 p_2 p_3 p_4$;

$$P(y_{i1}, y_{i2}, y_{i3}, y_{i4} | p_1, p_2, p_3, p_4) = \frac{100!}{y_{i1}! y_{i2}! y_{i3}! y_{i4}!} p_1^{y_{i1}} p_2^{y_{i2}} p_3^{y_{i3}} p_4^{y_{i4}}$$

 $P(p_1, p_2, p_3, p_4|y_{i1}, y_{i2}, y_{i3}, y_{i4}) \propto P(y_{i1}, y_{i2}, y_{i3}, y_{i4}|p_1, p_2, p_3, p_4) f(p_1, p_2, p_3, p_4|\alpha_1, \alpha_2, \alpha_3, \alpha_4)$

$$\propto \frac{100!}{y_{i1}! y_{i2}! y_{i3}! y_{i4}!} p_1^{y_{i1}+\alpha_1-1} p_2^{y_{i2}+\alpha_2-1} p_3^{y_{i3}+\alpha_3-1} p_4^{y_{i4}+\alpha_4-1}$$

 $p_1, p_2, p_3, p_4 | y_{i1}, y_{i2}, y_{i3}, y_{i4} \sim Dirichlet(y_{i1} + \alpha_1, y_{i2} + \alpha_2, y_{i3} + \alpha_3, y_{i4} + \alpha_4)$

$$P(y_{i2}|\mathbf{Y},\mathbf{P}) \propto \frac{p_2^{y_{i2}}}{y_{i2}!}$$

Then MH-Step to update y_{i2} (for i = 1, 2) is

$$\begin{split} y_{i2}^{(t+1)} &= \begin{cases} y_{i2}^{(t)} + 1, & \text{with probability } \frac{1}{2} \\ y_{i2}^{(t)} - 1, & \text{with probability } \frac{1}{2} \end{cases} \\ y_{i2}^{(t+1)} &= y_{i2}^{(t)} + 1, & \text{if } y_{i2}^{(t)} = 15 \\ y_{i2}^{(t+1)} &= y_{i2}^{(t)} - 1, & \text{if } y_{i2}^{(t)} = 32 \end{cases} \\ r &= \min \Big\{ \frac{p_{2}^{y_{i2}^{(t+1)}} / y_{i2}^{(t+1)}!}{p_{2}^{y_{i2}^{(t)}} / y_{i2}^{(t)}!}, 1 \Big\} \end{split}$$

where r is the accept-reject ratio. Then for another two unobserved variables:

$$y_{11}^{(t+1)} = 100 - y_{13} - y_{14} - y_{12}^{(t+1)} = 100 - 22 - 31 - y_{12}^{(t+1)}$$
$$y_{24}^{(t+1)} = 100 - y_{21} - y_{23} - y_{22}^{(t+1)} = 100 - 28 - 26 - y_{22}^{(t+1)}$$

Result:

For Stage 1: $p_1 = 0.27$, $p_2 = 0.199$, $p_3 = 0.22$, and $p_4 = 0.311$.

For Stage 2: $p_1 = 0.281$, $p_2 = 0.2$, $p_3 = 0.259$, and $p_4 = 0.259$.