# ANOVA III. Application to Categorical Predictors

Miaoyan Wang

Department of Statistics UW Madison

# Two representations of multiple samples comparison

• Data table may look like:

Factor	Treatment(Trt) 1	Trt 2		Trt k
Observation	<i>y</i> <sub>11</sub>	<i>y</i> <sub>21</sub>		<i>y</i> <sub>k1</sub>
	<i>y</i> <sub>12</sub>	<i>y</i> <sub>22</sub>		<i>y</i> <sub>k2</sub>
	<u>:</u>	:	:	:
	$y_{1n_1}$	$y_{2n_2}$		$y_{kn_k}$
mean	$\bar{y}_1$ .	$\bar{y}_2$ .		$\bar{y}_k$ .

• Or more often,

Trt	originally indexed y	re-indexed $y'$	
1	<i>y</i> 11	$y_1'$	
÷	:	÷	
1	$y_{1n_1}$	$y'_{n_1}$	
:	:		
k	<i>Yk</i> 1	$y'_{n_1++n_{k-1}+1}$	
	• • •		
k	$y_{kn_k}$	$y'_{n_1++n_{k-1}+n_k}$	

### Model Formulation

Consider a linear model with a K-level categorical predictor:

$$Y_n' = \beta_1 \mathbb{1}_{Trt1} + \dots + \beta_k \mathbb{1}_{Trtk} + \varepsilon_n, \quad \text{for } n = 1, \dots, (n_1 + \dots + n_k).$$

Equivalently, let

$$Y_{ij}=\mu_i+\varepsilon_{ij},$$

#### where

- $j = 1, ..., n_i$  indexes sample unit and i = 1, ..., k indexes treatment levels.
- Y<sub>ij</sub> is the response variable of the jth unit receiving the ith level of treatment (TrT).
- ullet  $\mu_i$  is the population mean for the ith treatment level
- $\varepsilon_{ij}$  is a random error for the *j*th sample unit and the *i*th treatment level. We assume  $\varepsilon_{ij} \sim \text{iid } N(0, \sigma^2)$ .
- The model parameters are  $\mu_1, \ldots, \mu_k, \sigma^2$  estimated by  $\bar{Y}_1, \ldots, \bar{Y}_k$ , and MSE, respectively.

## Alternative Model Formulation

Let

$$Y_{ij} = \mu + \alpha_i + \varepsilon_{ij}$$
, with constraints  $\sum_{i=1}^k \alpha_i = 0$ 

#### where

- $j = 1, ..., n_i$  indexes sample unit and i = 1, ..., k indexes treatment levels.
- $\mu$  is the grand population mean  $\mu = \frac{1}{k} \sum_{i=1}^{k} \mu_i$ .
- $\alpha_i = \mu_i \mu$  is the difference between the *i*th treatment mean and the grand mean (note:  $\sum_{i=1}^k \alpha_i = 0$ ).
- The random errors are assumed to follow

$$\varepsilon_{ij} \sim \text{iid } N(0, \sigma^2).$$

• The model parameters are  $\mu, \alpha_1, \ldots, \alpha_k, \sigma^2$  estimated by  $\bar{Y}_{...}$ ,  $\bar{Y}_{1.} - \bar{Y}_{...}, \ldots, \bar{Y}_{k.} - \bar{Y}_{...}$ , and MSE, respectively.



### Test for Overall Trt Effect

A one-way ANOVA analysis begins with testing

$$H_0: \mu_1 = \mu_2 = \cdots = \mu_k$$
 vs  $H_A:$  not all  $\mu_i$ 's are equal.

Or equivalently, test

$$H_0: \alpha_i = 0$$
 for all  $i$  vs  $H_A:$  not all  $\alpha_i$ 's are zero.

- The approach is to partition the total sum of squares, construct an ANOVA table, and perform an F test.
- Notation regarding the means:
  - $ightharpoonup \bar{Y}$ .. is the grand mean,
  - $\bar{Y}_{i}$  is the *i*th group (Trt) mean.

# Partition of Total Sum of Squares

Partition the total sum of squares (SS):

$$SST = SSR + SSE$$
,

#### where

► Total SS: SST =  $\sum_{i=1}^{k} \sum_{j=1}^{n_i} (Y_{ij} - \bar{Y}_{..})^2$  on df = N - 1. If balanced design:  $\sum_{i=1}^{k} \sum_{i=1}^{n} (Y_{ij} - \bar{Y}_{..})^2$  on df = kn - 1.

- Between-Trt SS:  $SSE = \sum_{i=1}^{k} \sum_{j=1}^{n_i} (\bar{Y}_{i.} - \bar{Y}_{..})^2 \text{ on df} = k - 1.$ If balanced design:  $n \sum_{i=1}^{k} (\bar{Y}_{i.} - \bar{Y}_{..})^2 \text{ on df} = k - 1.$
- SS for error:  $SSE = \sum_{i=1}^{k} \sum_{j=1}^{n_i} (Y_{ij} - \bar{Y}_{i\cdot})^2 \text{ on } df = N - k.$ If balanced design:  $\sum_{i=1}^{k} \sum_{i=1}^{n} (Y_{ij} - \bar{Y}_{i\cdot})^2 \text{ on } df = k(n-1).$
- Recall that MSE is an estimate of the error variance  $\sigma^2$ .



### ANOVA Table

SS: Sum of squares

MS: Mean sum of squares

Within-group SS  $S_{Within}$ , Between-group SS  $S_{Between}$ 

	SS	df	MS	E(MS)
Between	$S_{Between} = \sum_{j} n_j (y_{j.} - y_{})^2$	K – 1	$S_{Between}/(K-1)$	$\sigma^2 + (K-1)^{-1} \sum_{j=1}^k n_j t_j^2$
Treatment SS (Model)	,			
Within Treatment SS (Error)	$S_{Within} = \sum_{j} \sum_{k=1}^{n_j} (y_{jk} - y_{j.})^2$	n – K	$S_{Within}/(n-K)$	$\sigma^2$
Total SS	$S_{Total} = \sum_{j} \sum_{k=1}^{n_j} (y_{jk} - y_{})^2$	n-1	$S_{Total}/(n-1)$	

where  $t_j = \mu_j - \mu$  and  $\mu = n^{-1} \sum_i n_i \mu_i$ . Q: simplify in a balanced design  $n_1 = ... = n_K$ ?

• In a general one-way ANOVA, under the  $H_0$ :  $\alpha_i = 0$  for all i,

$$F = \frac{MSR}{MSE} \sim F_{k-1,N-k}.$$

# Example

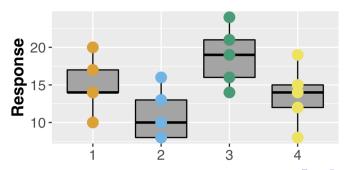
```
library(ggplot2)
y <- c(14, 20, 10, 14, 17, 13, 8, 10, 16, 8, 16, 14, 24, 21, 19, 8, 14, 19, 12, 15)
trt <- factor(rep(1:4, each = 5))
cbind(y, trt)
```

```
y trt
## [1,] 14 1
## [2,] 20
## [3,] 10
## [4,] 14
## [5,] 17
## [6,] 13
## [7,] 8
## [8,] 10
## [9,] 16
## [10,] 8
## [11,] 16
## [12,] 14
## [13,] 24
## [14,] 21
## [15,] 19
## [16,] 8
## [17,] 14
## [18.] 19
## [19,] 12
## [20,] 15
```

## Example

```
ggnlt(data.frame(response=y, trt = trt), aes(x = trt, y = response)) +
geom_boxplot(aes(colour=factor(trt)), fill="#A4A4A4", color="black")+
geom_point(aes(color = factor(trt)), size = 4)+
scale_colour_manual(values = c("#E569F00", "#56B4E9", "#009E73", "#F0E442"))+
scale_fill_manual(values = c("#E569F00", "#56B4E9", "#009E73", "#F0E442"))+
labs(colour = "Treatments")+xlab("")+ylab("Response")+
theme(axis.text=element_text(size=12),
axis.title=element_text(size=14, face="bold"),
legend.text=element_text(size=14, face="bold"),
legend.title=element_text(size=14, face="bold"),
legend.title=element_text(size=14, face="bold"),
legend.title=element_text(size=14, face="bold"),
legend.title=element_text(size=14, face="bold"),
legend.position = "top")
```





## Example

```
anova(lm(y~trt))
## Analysis of Variance Table
##
## Response: y
##
            Df Sum Sq Mean Sq F value Pr(>F)
             3 158.8 52.933 3.6506 0.03533 *
## trt
## Residuals 16 232.0 14.500
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
Marginal evidence against the null hypothesis.
Note that
1-pf(3.6506,3,16)
## [1] 0.03532877
pf(3.6506,3,16,lower.tail = FALSE)
## [1] 0.03532877
```