

Math 50 Final – Practice

Instructor: Ethan Levien

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Name: _____

Section: _____

Instructions

- You have 3 hours to complete the exam.
- You may have a one page (single-sided) “cheat sheet” which must be turned in with the exam, but no electronics (including calculators).
- Each problem is worth 5 points.
- Write your solutions in the boxes.
- Don’t cheat.

Exercise 1 (Converting code to math): Consider the following code

```
x1 = np.random.normal(0,1,100)
x2 = 2*x1 + np.random.normal(0,1,100)
y = x1 + x2 + np.random.normal(0,1,100)
b1 = np.cov(y,x1)/np.var(x1)
b2 = np.cov(y,x2)/np.var(x2)
```

What are b1 and b2 (approximately)?

Solution:

Exercise 2 (Joint distribution): Consider the model of a time series X_1, X_2, X_3, \dots :

$$X_{i+1}|X_i \sim \text{Normal}(1 + X_i/2, 1/3)$$

- (a) Write a Python function `generatesim(L)` to generates a simulation of L steps of this time series starting with $X_i = 0$. The function should return a length L numpy array.

- (b) After many steps, the process reaches a steady state where $E[X_{i+1}] = E[X_i]$. What is the distribution of X_i in steady-state?

Exercise 3 (Regression model comparison): (X, Y) data is fit to a single-predictor regression model in statsmodels using OLS, yielding the following output:

	coef	std err	t	P> t	[0.025	0.975]
const	1.0685	0.186	5.756	0.000	0.700	1.437
x1	1.9622	0.177	11.062	0.000	1.610	2.314

A second predictor X_2 is then included in the model, which yields the following output:

	coef	std err	t	P> t	[0.025	0.975]
const	1.0001	0.011	92.003	0.000	0.979	1.022
x1	0.9810	0.012	82.470	0.000	0.957	1.005
x2	2.0122	0.012	168.877	0.000	1.989	2.036

If X_2 and X_1 were fit to a linear regression model with X_2 as the response variable, what would be the regression slope?

Solution:

Exercise 4 (Sample distribution): Consider the model

$$Y_1 \sim \text{Normal}(\beta_1 X_1 + \beta_2 X_2, \sigma_\epsilon^2)$$

After fitting the model, we find $\hat{\beta}_1 = 100$, $\hat{\beta}_2 = -101$, $\hat{\sigma}_\epsilon^2 = 1/4$. The model is then fit to a different data set and it is found that $\hat{\beta}_1 = -100$, $\hat{\beta}_2 = 100.4$.

- (a) Is $\text{cov}(X_1, X_2)$ likely to be positive or negative for the fitted data?

Solution:

- (b) Based on this information is it possible that R^2 is very close to 1?

Solution:

Exercise 5 (Bernoulli regression model): Consider the two predictor linear regression model with an interaction term:

$$Y = \beta_1 X_1 + \beta_2 X_2 + J_{1,2} X_1 X_2 + \epsilon$$

The following plot shows data generated from such a model.

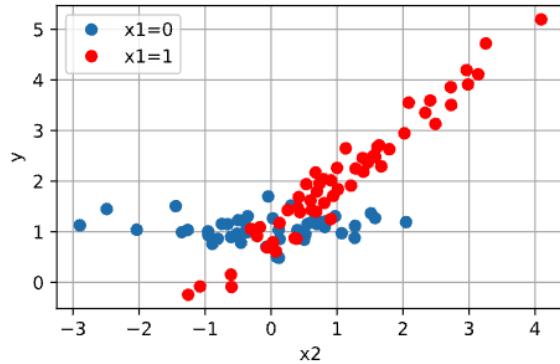


Figure 1:

What are the values of β_1 , β_2 and $J_{1,2}$?

Solution:

Exercise 6 (Orthogonality): Consider the features $\phi_1(x) = x^3$ and $\phi_2(x) = x^2$.

- (a) Are these orthogonal with respect to $X \sim \text{Uniform}(-1, 1)$? (in the sense that $E[\phi_1(X)\phi_2(X)] = 0$)

- (b) What is an example of a distribution for X such that ϕ_1 and ϕ_2 are not orthogonal?

Exercise 7 (Bayesian posterior): Consider a the Bayesian model σ_ϵ and β_1 , but unknown intercept with Normal priors:

$$\beta_0 \sim \text{Normal}(0, \tau_0^2) \quad (1)$$

$$Y|X, \beta \sim \text{Normal}(\beta_0 + \beta_1 X, \sigma_\epsilon^2) \quad (2)$$

Calculate the posterior of β_0 .

Exercise 8 (Missing data): You are given data with predictors X_1, X_2 . You want to fit a linear regression model

$$Y = \beta_0 + \beta_1 X_1 + \beta_2 X_2 + \epsilon$$

but some of the X_2 values have been corrupted and are not reliable. One idea to handle this is called *imputation* and involves generating the missing values X_2 .

Explain how you could implement imputation assuming you are given a dataframe with rows Y, X_1, X_2, C where $C = 1$ for the rows with corrupted data and $C = 0$ otherwise. What assumptions are being made for the procedure to be unbiased?

Solution:

Solutions

1. b_1 is the regression coefficient with only X_1 as a predictor and the regression coefficient $\beta_{1,2}$ of X_1 on X_2 is given to be 2 as well as $\beta_1 = \beta_2 = 1$. Thus we have $b_1 = 1+2 = 3$. The regression coefficient of X_2 on X_1 is $\beta_{2,1} = \text{cov}(X_1, X_2)/\text{var}(X_2) = \text{var}(X_1)\beta_{1,2}/(2^2 + 1) = 1 \times 2/5 = 2/5$. Hence $b_2 = 1 + 2/5 = 7/5$.

2. (a) One valid implementation:

```
def generatesim(L):
    X = np.zeros(L)
    for i in range(1,L):
        X[i] = 1.0 + 0.5*X[i-1] + np.random.normal(0.0, np.sqrt(1/3))
    return X
```

(b) In steady-state we have the marginals are equal so $\mu_X = E[X_{i+1}] = 1 + E[X_i]/2 = 1 + \mu_X/2$. Thus $\mu_X = 2$. Similarly $\sigma_X^2 = \sigma_X^2/4 + 1/3$. Thus $\sigma_X^2 = 4/9$ and $X_i \sim \text{Normal}(2, 4/9)$.

3. We are given the single predictor regression coefficient $\hat{\beta}'_1 \approx 2$ and the two predictor model coefficients $\hat{\beta}_1 \approx 1$, $\hat{\beta}_2 \approx 2$. Using $\hat{\beta}'_1 = \hat{\beta}_1 + \hat{\beta}_2\beta_{1,2}$ we deduce the coefficient of X_1 with X_2 as the response variable, denoted $\beta_{1,2}$ here, is $\beta_{1,2} \approx 1/2$.
4. (a) This suggests the $\hat{\beta}_1$ and $\hat{\beta}_2$ are negatively correlated and therefore X_1 and X_2 are positively correlated, hence $\text{cov}(X_1, X_2) > 0$. (b) Yes, because even though we know $\hat{\sigma}_\epsilon^2$ we aren't given σ_Y^2 , which could be quite large especially with such large values of $\hat{\beta}_i$.
5. β_2 is the slope of X_2 vs. Y when $X_1 = 0$, which is approximately 0 in the plot. $\beta_2 + J_{1,2}$ is the slope when $X_1 = 1$, which appears to be 1. β_1 is the expected difference in Y between $X_1 = 0$ and $X_1 = 1$ groups when $X_2 = 0$, which appears to be 0. In summary, $\beta_1 = \beta_2 = 0$ and $J_{1,2} = 1$.

6. (a) With $X \sim \text{Unif}(-1, 1)$,

$$E[\phi_1(X)\phi_2(X)] = E[X^5] = \int_{-1}^1 x^5 \cdot \frac{1}{2} dx = 0, \quad (3)$$

so they are orthogonal. (b) Any non-symmetric distribution makes $E[X^5] \neq 0$. For example, if $X \sim \text{Unif}(0, 1)$,

$$E[X^5] = \frac{1}{6} \neq 0, \quad (4)$$

so they are not orthogonal.

7. The (frequentist) estimator $\hat{\beta}_0 = \bar{Y} - \beta_1 \bar{X}$. This is Normal when conditioned on β_0 : $\hat{\beta}_0 | \beta_0 \sim \text{Normal}(\beta_0, \hat{\sigma}_\epsilon^2/N)$. If we invert this regression model we get

$$\beta_0 | \hat{\beta}_0 \sim \text{Normal}\left(\beta_0 \frac{\tau^2}{\tau^2 + \hat{\sigma}_\epsilon^2/N}, \frac{\tau^2 \hat{\sigma}_\epsilon^2/N}{\tau^2 + \hat{\sigma}_\epsilon^2/N}\right) \quad (5)$$

8. Using rows with $C = 0$, fit

$$X_2 = \alpha_0 + \alpha_1 X_1 + \zeta, \quad (6)$$

and estimate $\hat{\sigma}_\zeta^2 \approx \text{var}(\zeta)$. For each $C = 1$ row,

$$\tilde{X}_2 = \hat{\alpha}_0 + \hat{\alpha}_1 X_1 + \tilde{\zeta}, \quad \tilde{\zeta} \sim \text{Normal}(0, \hat{\sigma}_\zeta^2). \quad (7)$$

Then we fit

$$Y = \beta_0 + \beta_1 X_1 + \beta_2 \tilde{X}_2 + \epsilon \quad (8)$$

on the completed data. We have assumed C is independent of X_1 and X_2 .