

Unit 4 E14.1 problems A, B and C

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Econometrics II, Bachelor degree in Economics

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E14.1 On the text website, www.pearsonglobaleditions.com/Stock_Watson, you will find the data file **USMacro_Quarterly**, which contains quarterly data on several macroeconomic series for the United States; the data are described in the file **USMacro_Description**. The variable $PCEP$ is the price index for personal consumption expenditures from the U.S. National Income and Product Accounts. In this exercise you will construct forecasting models for the rate of inflation, based on $PCEP$. For this analysis, use the sample period 1963:Q1–2012:Q4 (where data before 1963 may be used, as necessary, as initial values for lags in regressions).

- a.** i. Compute the inflation rate, $Infl = 400 \times [\ln(PCEP_t) - \ln(PCEP_{t-1})]$. What are the units of $Infl$? (Is $Infl$ measured in dollars, percentage points, percentage per quarter, percentage per year, or something else? Explain.)
ii. Plot the value of $Infl$ from 1963:Q1 through 2012:Q4. Based on the plot, do you think that $Infl$ has a stochastic trend? Explain.
- b.** i. Compute the first four autocorrelations of $\Delta Infl$.
ii. Plot the value of $\Delta Infl$ from 1963:Q1 through 2012:Q4. The plot should look “choppy” or “jagged.” Explain why this behavior is consistent with the first autocorrelation that you computed in part (i).

- c
- Run an OLS regression of $\Delta Infl_t$ on $\Delta Infl_{t-1}$. Does knowing the change in inflation this quarter help predict the change in inflation next quarter? Explain.
 - Estimate an AR(2) model for $\Delta Infl$. Is the AR(2) model better than an AR(1) model? Explain.
 - Estimate an AR(p) model for $p = 0, \dots, 8$. What lag length is chosen by BIC? What lag length is chosen by AIC?
 - Use the AR(2) model to predict the change in inflation from 2012:Q4 to 2013:Q1—that is, predict the value of $\Delta Infl_{2013:Q1}$.
 - Use the AR(2) model to predict the level of the inflation rate in 2013:Q1—that is, $Infl_{2013:Q1}$.

```
[13]: # TOOLS
import numpy as np
import pandas as pd
import matplotlib.pyplot as plt
from statsmodels.tsa.arima_model import ARIMA
from statsmodels.formula.api import ols
```

```
[12]: # Data without null values and with the obs column as the index
dataExcel = pd.read_csv("us_macro_quarterly.csv", index_col='obs').dropna()
data_converted_index = pd.to_datetime(dataExcel.index)
dataExcel.index = pd.DatetimeIndex(data_converted_index.values, freq = "Q",
    →data_converted_index.inferred_freq)

# New inflation rate variable
dataExcel['INFL'] = 400 * (np.log(dataExcel['PCECTPI']) - np.
    →log(dataExcel['PCECTPI']).shift())

# New change in inflation rate variable
dataExcel['changeINFL'] = dataExcel['INFL'] - dataExcel['INFL'].shift()
dataExcel
```

```
[12]:
```

	GDPC96	JAPAN_IP	PCECTPI	GS10	GS1	TB3MS	\
1959-01-01	2973.782	9.425179	17.137	3.990000	3.503333	2.773333	
1959-04-01	3046.096	10.080844	17.204	4.256667	3.916667	3.000000	
1959-07-01	3040.235	10.681870	17.307	4.503333	4.603333	3.540000	
1959-10-01	3052.194	11.419493	17.401	4.583333	4.916667	4.230000	
1960-01-01	3120.195	12.184435	17.424	4.486667	4.570000	3.873333	
...	
2012-07-01	15533.985	95.792017	106.193	1.643333	0.183333	0.103333	
2012-10-01	15539.628	94.258812	106.622	1.706667	0.173333	0.086667	
2013-01-01	15583.948	94.725440	106.909	1.950000	0.153333	0.086667	

2013-04-01	15679.677	95.992001	106.878	1.996667	0.126667	0.050000
2013-07-01	15839.347	97.558537	107.387	2.710000	0.123333	0.033333

	UNRATE	EXUSUK	CPIAUCSL	INFL	changeINFL
1959-01-01	5.833333	2.809507	28.993333	NaN	NaN
1959-04-01	5.100000	2.814537	29.043333	1.560818	NaN
1959-07-01	5.266667	2.808283	29.193333	2.387652	0.826833
1959-10-01	5.600000	2.802463	29.370000	2.166653	-0.220999
1960-01-01	5.133333	2.802970	29.396667	0.528356	-1.638297
...
2012-07-01	8.033333	1.581367	230.029667	1.672150	0.615436
2012-10-01	7.833333	1.606433	231.277000	1.612670	-0.059480
2013-01-01	7.700000	1.550633	232.102667	1.075254	-0.537416
2013-04-01	7.500000	1.536700	232.086667	-0.116003	-1.191258
2013-07-01	7.233333	1.552300	233.597000	1.900454	2.016457

[219 rows x 11 columns]

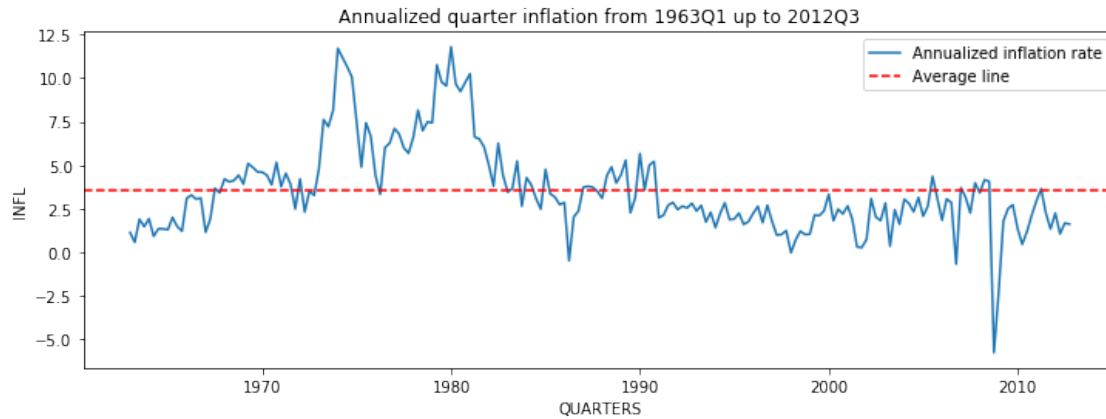
a. i.

The new variable **INFL** measures the approximate annualized percentage change of the **PCECTPI**. It is measured in percentage per year because the approximate percentage change of the quarter was annualized when multiplying it by four.

a. ii.

```
[11]: # Plot the infl from the first quarter of 1963 up to the fourth quarter of 2012
plt.plot(dataExcel.loc['1963-01-01': '2012-10-01', 'INFL'], label='Annualized_
    ↳inflation rate')

# INFL mean and plot information
plt.axhline(y=np.mean(dataExcel['INFL'].loc['1963-01-01': '2012-10-01']),
    ↳linestyle='--', color='r', label='Average line')
plt.title('Annualized quarter inflation from 1963Q1 up to 2012Q3')
plt.ylabel('INFL')
plt.xlabel('QUARTERS')
plt.rcParams["figure.figsize"] = [12,4]
plt.legend()
plt.show()
```



INFL does look to have a non-stationary or a weak stationary stochastic trend, because the values seem to be inside of a range of random numbers that are indexed by time but they do not seem to have the same autocovariance and same mean for all the **quarters** as the *red dotted line* intuitively shows.

b. i.

```
[4]: # Compute correlation coefficient from 1 to 4
dataset = dataExcel['changeINFL'].loc['1963-01-01': '2012-10-01']
acf(dataset, fft=True, nlags=4)
```

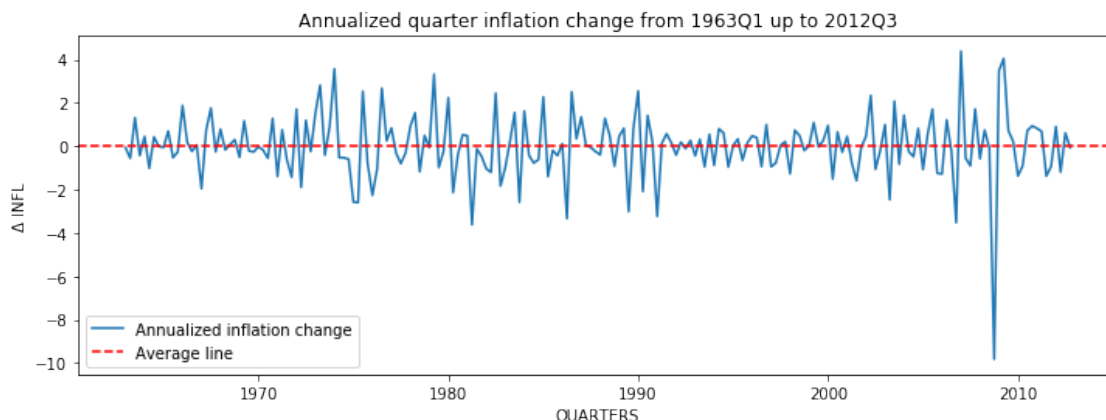
```
[4]: array([ 1.          , -0.25309496, -0.18285883,  0.13264088, -0.09238671])
```

The autocorrelation with 1, 2, 3 and 4 lags are - 0.25309496, - 0.18285883, 0.13264088 and - 0.09238671 respectively.

b. ii.

```
[5]: # Plot the infl from the first quarter of 1963 up to the fourth quarter of 2012
plt.plot(dataExcel['changeINFL'].loc['1963-01-01': '2012-10-01'],
        label='Annualized inflation change')

# INFL mean and plot information
plt.title('Annualized quarter inflation change from 1963Q1 up to 2012Q3')
plt.axhline(y=0, linestyle='--', color='r', label='Average line')
plt.ylabel('Δ INFL')
plt.xlabel('QUARTERS')
plt.rcParams["figure.figsize"] = [12,4]
plt.legend()
plt.show()
```



When there is a negative correlation we expect to see a pattern of values crossing the *red dotted line* too often. This pattern is caused because the previous value seems to be related with an opposite sign in the next value. As we see in the plot, the values cross the *red dotted line* too often and this is consistent with the fact that in the previous exercise, the first autocorrelation was of **-0.25309756405695916**, a negative correlation.

c. i.

```
[6]: # Computing the model with ols
modelA = ols('changeINFL ~ changeINFL.shift()', data=dataExcel.loc['1963-01-01':
↪ '2012-10-01']).fit(cov_type='HC1')
print(modelA.summary())
```

```

                                OLS Regression Results
=====
Dep. Variable:                  changeINFL    R-squared:                0.064
Model:                            OLS        Adj. R-squared:            0.059
Method:                 Least Squares    F-statistic:                 12.67
Date:                Fri, 04 Dec 2020    Prob (F-statistic):          0.000466
Time:                  11:28:35        Log-Likelihood:              -356.48
No. Observations:                199        AIC:                        717.0
Df Residuals:                    197        BIC:                        723.5
Df Model:                        1
Covariance Type:                HC1
=====
=====
                                coef    std err          z      P>|z|      [0.025
-----
0.975]
-----
Intercept                0.0030    0.103      0.029      0.977      -0.200
0.206
```

```

changeINFL.shift()    -0.2531      0.071      -3.560      0.000      -0.392
-0.114
=====
Omnibus:                90.025      Durbin-Watson:                2.133
Prob(Omnibus):          0.000      Jarque-Bera (JB):            970.079
Skew:                  -1.389      Prob(JB):                    2.24e-211
Kurtosis:              13.453      Cond. No.                    1.50
=====

```

Warnings:

```
[1] Standard Errors are heteroscedasticity robust (HC1)
```

It seems to be that the next change in inflation is approximately - **0.2531** times the current change in inflation.

It is important to notice that is reasonable to be sceptic for two reasons, the obvious one is that, even though the correlation coefficient is significantly different from zero, the R squared of this model is quite small and therefore this model apparently just explains a small proportion of the variance in the sampled future changes. The second reason is that for seasonal patterns the OLS method does not assume normality and we would have wider confidence intervals and therefore lower precision (heteroskedasticity and autocorrelation (HC1) robust standard errors can improve it but it is still a problem. So, we could argue that the current inflation change helps to predict the change in inflation next quarter, but it is not appropriate to use OLS to compute the autoregressive models since we might get slightly less precise estimates.

c. ii.

```

[7]: # Make AR(1) model
modelB = ARIMA(dataExcel['changeINFL'].loc['1963-01-01': '2012-10-01'], (1, 0, 0, 0)).fit()
print(modelB.summary())

# Make AR(2) model
modelC = ARIMA(dataExcel['changeINFL'].loc['1963-01-01': '2012-10-01'], (2, 0, 0, 0)).fit()
print(modelC.summary())

```

ARMA Model Results

```

=====
Dep. Variable:          changeINFL      No. Observations:          200
Model:                  ARMA(1, 0)      Log Likelihood             -357.800
Method:                 css-mle         S.D. of innovations        1.448
Date:                   Fri, 04 Dec 2020 AIC                               721.600
Time:                   11:28:35        BIC                               731.495
Sample:                 01-01-1963      HQIC                          725.604
                        - 10-01-2012
=====
=====

```

	coef	std err	z	P> z	[0.025
--	------	---------	---	------	--------

0.975]

const	0.0022	0.082	0.027	0.979	-0.158
0.163					
ar.L1.changeINFL	-0.2518	0.068	-3.692	0.000	-0.386
-0.118					
Roots					
=====					
	Real	Imaginary	Modulus	Frequency	

AR.1	-3.9707	+0.0000j	3.9707	0.5000	

ARMA Model Results					
=====					
Dep. Variable:	changeINFL	No. Observations:	200		
Model:	ARMA(2, 0)	Log Likelihood	-350.645		
Method:	css-mle	S.D. of innovations	1.396		
Date:	Fri, 04 Dec 2020	AIC	709.290		
Time:	11:28:35	BIC	722.483		
Sample:	01-01-1963	HQIC	714.629		
	- 10-01-2012				
=====					
=====					
	coef	std err	z	P> z	[0.025
0.975]					

const	0.0023	0.063	0.036	0.971	-0.121
0.125					
ar.L1.changeINFL	-0.3184	0.068	-4.676	0.000	-0.452
-0.185					
ar.L2.changeINFL	-0.2616	0.068	-3.855	0.000	-0.395
-0.129					
Roots					
=====					
	Real	Imaginary	Modulus	Frequency	

AR.1	-0.6085	-1.8579j	1.9550	-0.3004	
AR.2	-0.6085	+1.8579j	1.9550	0.3004	

The coefficient of determination is not a good indicator when comparing different models since it might increase just with the fact of increasing the number of lagged regressors in the model. To avoid that we can use the Akaike Information Criterion (AIC) or the Bayesian Information Criterion (BIC).

To compare both models, I will directly use the BIC and AIC values:

The AR(1) has a AIC of **721.600** and a BIC of **731.495**

The AR(2) has a AIC of **709.290** and a BIC of **722.483**

As we see, the AR(2) seems to be a better model to explain the next quarter change in inflation with the current and last quarter change in inflation since both AIC and BIC in AR(2) have lower values than in AR(1)

c. iii.

```
[15]: # Compute AIC and BID for different models
# I used a function to avoid memory leaks with the assignment of variables
def arima_n(delta):
    aic = []
    bic = []
    for delta in range(1, delta + 1):
        model = ARIMA(dataExcel['changeINFL'].loc['1963-01-01': '2012-10-01'],
            (delta, 0, 0)).fit()
        aic.append(model.aic)
        bic.append(model.bic)
        print(f'AR({delta}) | AIC: {model.aic:0.10f} | BIC: {model.bic:0.10f}')
    print(f'Optimal AIC lenght: {aic.index(min(aic)) + 1}')
    print(f'Optimal BIC length: {bic.index(min(bic)) + 1}')

arima_n(8)
```

```
AR(1) | AIC: 721.5998453243 | BIC: 731.4947974240
AR(2) | AIC: 709.2901820046 | BIC: 722.4834514708
AR(3) | AIC: 711.2728970970 | BIC: 727.7644839297
AR(4) | AIC: 710.8450970553 | BIC: 730.6350012546
AR(5) | AIC: 710.3517631010 | BIC: 733.4399846668
AR(6) | AIC: 710.7607442781 | BIC: 737.1472832105
AR(7) | AIC: 712.5827088755 | BIC: 742.2675651744
AR(8) | AIC: 714.4160715128 | BIC: 747.3992451783
Optimal AIC lenght: 2
Optimal BIC length: 2
```

The optimal AIC length is **2** with a value of **709.2901820046234**

The optimal BIC length is **2** with a value of **722.4834514708156**

c. iv.

```
[9]: # Print the first step out-of-sample forecast for the change in inflation
modelC.forecast(1)
```

```
[9]: (array([-0.13851351]),
      array([1.39621897]),
      array([[-2.87505242,  2.59802539]]))
```

The predicted change in inflation by AR(2) from 2012Q4 to 2013Q1 is - **0.13851351** with a stan-

standard error of **1.39621897** and a 95% confidence interval between - **2.87505242** and **2.59802539**

c. iv.

```
[10]: # Print the first step out-of-sample forecast for the inflation rate
modelD = ARIMA(dataExcel['INFL'].loc['1963-01-01': '2012-10-01'], (2, 0, 0)).
    ↪ fit(trend='nc')
print(modelD.forecast(1))
```

```
(1.5463120230697975, array([1.43504143]), array([[ -1.2663175 ,  4.35894155]]))
```

The predicted inflation by AR(2) for 2013Q1 is **1.87520473** with a standard error of **1.40925276** and a 95% confidence interval between - **0.88687992** and **4.63728937**