Quantitative Methods Human Sciences, 2020–21

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Lecture 2: 22 October 2020

► Recap on probability and counting: the birthday problem.

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- Introduction to conditional probability.

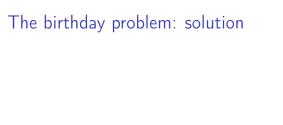
- ► Recap on probability and counting: the birthday problem.
- Introduction to conditional probability.
- Problem sheet 1 (tutorial).

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- Assume each person's birthday is equally likely to be any of the 365 days of the year and assume people's birthdays are independent.
- ▶ What is the probability that at least one pair of people in the group have the same birthday?
- ▶ Hint: Recall that $\mathbb{P}(A) = 1 \mathbb{P}(A^c)$.



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- ▶ $\mathbb{P}(\text{no birthday match}) = \frac{365 \times 364 \times \cdots \times (365 n + 1)}{365^n}$.
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- ▶ \mathbb{P} (no birthday match) = $\frac{365 \times 364 \times \cdots \times (365 n + 1)}{365^n}$.
- ▶ $\mathbb{P}(\text{at least one birthday match}) = 1 \mathbb{P}(\text{no birthday match}).$
- ▶ In this room, $\mathbb{P}(\text{at least one birthday match}) \approx 4\%$.

The birthday problem in R

```
# Create a function
pmatch <- function(n) {
          1 - prod(365:(365 - n + 1)) / (365 ^ n)
}</pre>
```

The birthday problem in R

For loop

for (i in 1:70) {

probs[i] <- pmatch(i)</pre>

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}</pre>
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```
# Alternative method
probs <- sapply(1:70, pmatch)</pre>
```

The birthday problem in **R** (cont.)

```
save <- data.frame("n" = 1:70, "prob" = probs)</pre>
```

The birthday problem in R (cont.)

```
save <- data.frame("n" = 1:70, "prob" = probs)</pre>
```

```
head(save)

## n prob

## 1 1 0.000000000

## 2 2 0.002739726

## 3 3 0.008204166

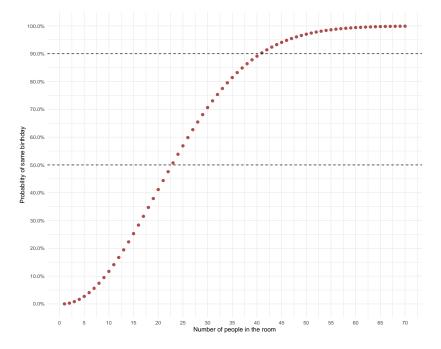
## 4 4 0.016355912

## 5 5 0.027135574

## 6 6 0.040462484
```

The birthday problem in R (cont.)

```
ggplot(save, aes(n, prob)) +
  geom_point( # Modify points
    size = 2.
    colour = "darkred",
    alpha = 0.7) +
 labs( # Axis labels
   x = "Number of people in the room",
    y = "Probability of same birthday") +
  scale_x_continuous( # Modify X-axis
    breaks = seq(0, 70, 5)) +
  scale_y_continuous( # Modify Y-axis
    breaks = seq(0, 1, 0.1),
    label = scales::percent) +
  geom_hline( # When is P(match) > 0.5 or 0.9?
    vintercept = c(0.5, 0.9), linetype = "dashed") +
  theme_minimal() # Remove redundant lines
```



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- Whenever new evidence is observed, we acquire information that may affect out uncertainties.
- Conditional probability allows us to update our beliefs in light of new evidence.
- "Conditioning is the soul of statistics" (Blitzstein and Hwang, 2019: 46).

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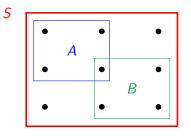
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- ▶ Note that $\mathbb{P}(A \mid A) = 1$.

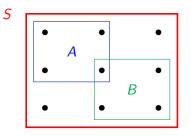
Visualising a conditional probability

Suppose we have a sample space S and two events A and B:



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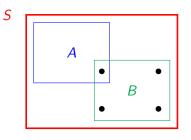
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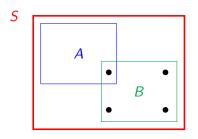
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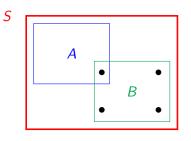
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- ▶ Divide by $\mathbb{P}(B)$, the total mass of the outcomes in B.

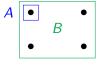
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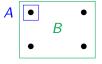
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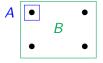
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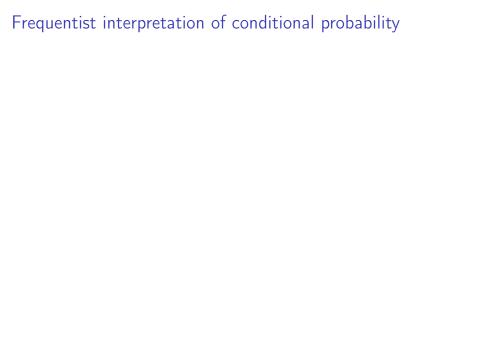
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- Relative measures of uncertainty are redistributed amongst remaining possible outcomes.



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- ► Thus

$$\mathbb{P}(A \mid B) = n_{AB}/n_B = (n_{AB}/n)/(n_B/n) = \mathbb{P}(A \cap B)/\mathbb{P}(B).$$

Joint probability and conditional probability

Theorem

For any events A and B with positive probabilities,

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This can be generalised to the intersection of n events:

$$\mathbb{P}(A_1 \cap \cdots \cap A_n) = \mathbb{P}(A_1)\mathbb{P}(A_2|A_1)\mathbb{P}(A_3|A_1 \cap A_2) \cdots \mathbb{P}(A_n|A_1 \cap \cdots \cap A_n).$$

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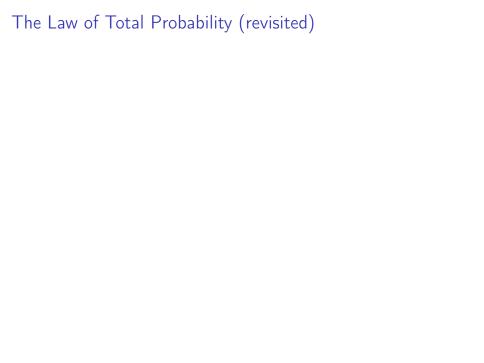
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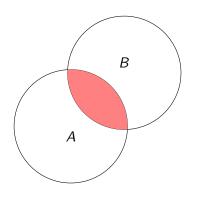
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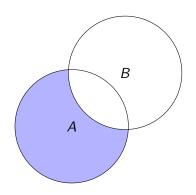
This follows immediately from the previous Theorem (which in turn follows immediately from the definition of conditional probability).



Recall that, for any two events A and B, the Law of Total Probability states that

$$\mathbb{P}(A) = \mathbb{P}(A \cap B) + \mathbb{P}(A \cap B^{c}).$$





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$$\approx 16\%.$$

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Conclusion:

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- All probabilities are conditional probabilities.

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- ► Independence is a symmetric relation: if A is independent of B, B is independent of A.
- ▶ Warning: independence \neq disjointness. In fact, disjoint events can only be independent if $\mathbb{P}(A) = 0$ or $\mathbb{P}(B) = 0$. (Why?)

Conditional independence

Definition

Events A and B are said to be *conditionally independent* given a third event E if

$$\mathbb{P}(A \cap B \mid E) = \mathbb{P}(A \mid E)\mathbb{P}(B \mid E).$$

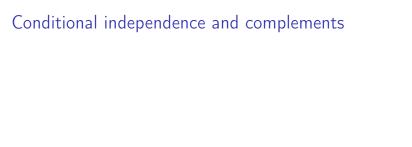
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Health warning: independence does *not* imply conditional independence and vice versa.



Conditional independence and complements

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- Suppose there are two types of teachers: those who give grades that reflect student effort (E), and those who randomly assign grades, regardless of student effort (E^c) .
- ▶ Let W be the event that you work hard and let G be the event that you receive a good grade.
- ▶ Then W and G are conditionally independent given E^c , but they are not conditionally independent given E.

▶ You have one fair coin and one biased coin which lands Heads with probability 3/4.

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- ➤ You have one fair coin and one biased coin which lands Heads with probability 3/4.
- ➤ You pick one of the coins at random, without knowing which one you've chosen, and flip it several times.
- Conditional on choosing either the fair or the biased coin, the coin flips are independent.
- However, the coin flips are not unconditionally independent: without knowing which coin we've chosen, each flip gives us new data from which we can predict outcomes of future flips.

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- ➤ You pick one of the coins at random, without knowing which one you've chosen, and flip it several times.
- Conditional on choosing either the fair or the biased coin, the coin flips are independent.
- However, the coin flips are not unconditionally independent: without knowing which coin we've chosen, each flip gives us new data from which we can predict outcomes of future flips.
- ► (Think about the definition of independence.)

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- ▶ Then A and B are unconditionally independent, with $\mathbb{P}(A) > 0$ and $\mathbb{P}(B) > 0$.
- ► However, given that I receive exactly one call tomorrow (C), A and B are no longer independent:

$$\mathbb{P}(A \mid C) > 0$$
, but $\mathbb{P}(A \mid C \cap B) = 0$.