

TAM 598 Lecture 6 :

RANDOM VECTORS AND  
MULTIVARIATE GAUSSIANS

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Announcements:

- HW 1 covers lectures 1-4 ; due on Feb 12 ;  
to be submitted via CANVAS
- HW2 covers lectures 4-8 ; due on Feb 26

**Random Vectors** - take  $N$  random variables  $X_1, X_2, \dots, X_N$  and put them in a vector

$$\underline{X} = (X_1, \dots, X_N)$$

Random vectors are used to model uncertain quantities such as

→ state of a multibody system : vector of coordinates & velocities

→ an unknown function : vector of its function values at  $N$  test points

→ an image : vector of pixel values

**PDF of a Random Vector** is the joint PDF of its components.

$$p(\underline{x}) = p(x_1, x_2, \dots, x_N)$$

① can marginalize  $p(x_i) = \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} p(x_1, x_2, \dots, x_N) dx_2 dx_3 \dots dx_N$

② can integrate out subsets  $p(x_1, x_2) = \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} p(x_1, x_2, \dots, x_N) dx_3 \dots dx_N$

Expectation of a Random Vector is the vector of expectations of each component

$$\mathbb{E}[\underline{x}] = \begin{pmatrix} \mathbb{E}[x_1] \\ \vdots \\ \mathbb{E}[x_n] \end{pmatrix}$$

This expectation is linear:

$$\mathbb{E}[ax + b\underline{y}] = a\mathbb{E}[\underline{x}] + b\mathbb{E}[\underline{y}]$$

### Covariance Matrix of Two Random Vectors

Let  $\underline{x}$  and  $\underline{y}$  be random vectors of dimension N and M respectively. The covariance of  $\underline{x}$  and  $\underline{y}$  is the  $N \times M$  matrix consisting of all covariances between components of  $\underline{x}$  and  $\underline{y}$ , ie

$$C[\underline{x}, \underline{y}] = (C[x_i, y_j])$$

① sometimes  $C[\underline{x}, \underline{y}]$  is called  $\underline{\Sigma}$  ("sigma")

② we can rewrite the covariance matrix as

$$C[\underline{x}, \underline{y}] = E[(\underline{x} - E[\underline{x}])(\underline{y} - E[\underline{y}])^T]$$

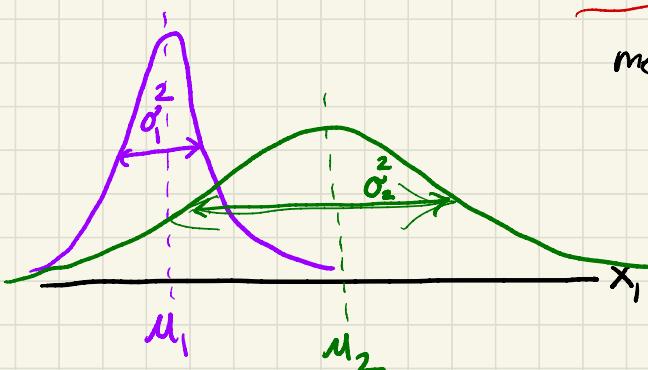
③ the  $N \times N$  matrix  $C[\underline{x}, \underline{x}]$  is the self covariance or just the covariance matrix of  $\underline{x}$ .

diagonal: variance of each component of  $\underline{x}$

off-diagonal: covariance of  $x_i, x_j$  for  $i \neq j$

## Multivariate Gaussian Distributions

recall univariate Gaussian distributions



"drawn from"

$$X \sim N(\mu, \sigma^2)$$

mean      ↑      Variance

PDF:

$$p(x) = \frac{1}{\sqrt{2\pi\sigma^2}} \exp\left\{-\frac{1}{2\sigma^2}(x-\mu)^2\right\}$$

properties of univariate Gaussian distributions :

① closed under linear transformation

$$X \sim N(\mu, \sigma^2) \quad Y = aX + b$$

corrected typo in class lecture

$a\mu + b$

$$Y \sim N(a\mu + b, a^2\sigma^2)$$

② adding / subtracting Gaussian RVs

$$X \sim N(\mu_x, \sigma_x^2)$$

$$Y \sim N(\mu_y, \sigma_y^2)$$

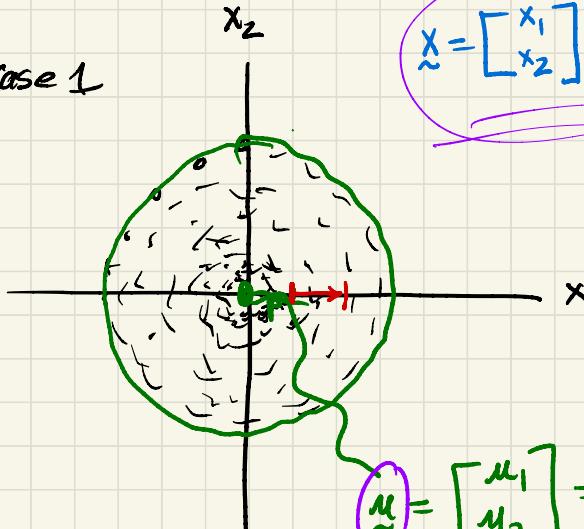
$$X \pm Y \sim N(\mu_x \pm \mu_y, \sigma_x^2 + \sigma_y^2)$$

$N \sim (\mu, \sigma^2)$

## Bivariate Gaussian Distributions

we construct a two-component random vector

case 1



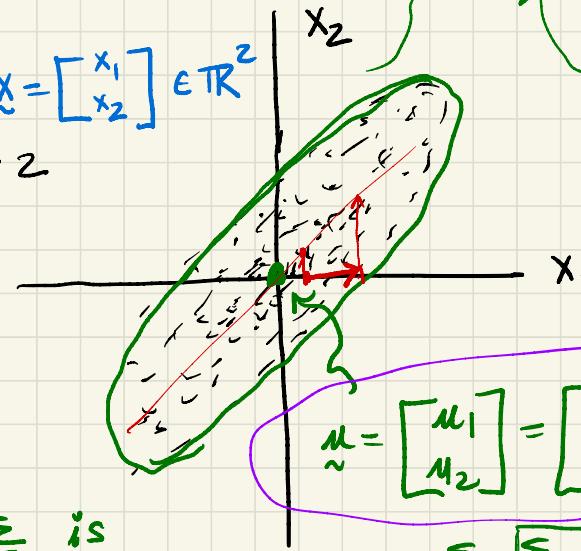
$$\mu = \begin{bmatrix} \mu_1 \\ \mu_2 \end{bmatrix} = \begin{bmatrix} 0 \\ 0 \end{bmatrix}$$

$$\Sigma = \begin{bmatrix} \Sigma_{11} & \Sigma_{12} \\ \Sigma_{21} & \Sigma_{22} \end{bmatrix} = \begin{bmatrix} 5 & 0 \\ 0 & 5 \end{bmatrix}$$

- two random variables from which

$N \sim (\mu, \Sigma)$

case 2



$\Sigma$  is  
"positive  
definite"

$$\Sigma = \begin{bmatrix} \Sigma_{11} & \Sigma_{12} \\ \Sigma_{21} & \Sigma_{22} \end{bmatrix}$$

$$= \begin{bmatrix} 5 & 2.5 \\ 2.5 & 5 \end{bmatrix}$$

recall how we defined covariance of two random variables :

$$\text{cov}(x_1, x_2) = \mathbb{E}((x_1 - \mathbb{E}[x_1])(x_2 - \mathbb{E}[x_2]))$$

① if  $\mathbb{E}[x_1] = 0$ ,  $\mathbb{E}[x_2] = 0$  then

$$\text{cov}(x_1, x_2) = \mathbb{E}[x_1 x_2] \quad \Sigma_{12}$$

②  $\text{cov}(x_1, x_1) = \mathbb{E}((x_1 - \mathbb{E}[x_1])(x_1 - \mathbb{E}[x_1]))$

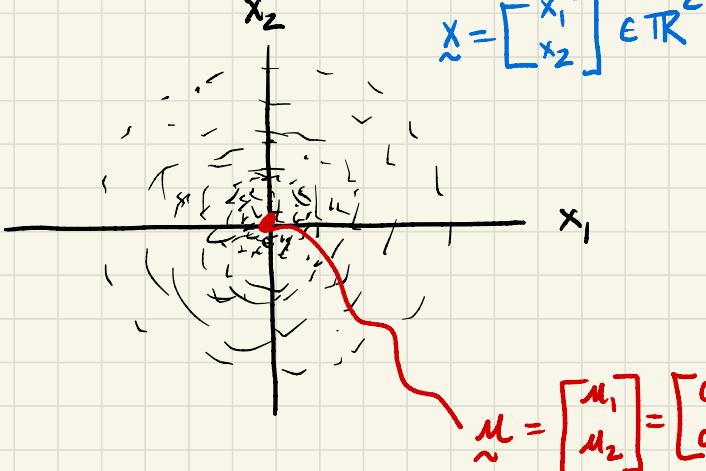
$$= \mathbb{E}[x_1^2] = \mathbb{V}(x_1)$$

③ another way to think of  $\text{cov}(x_1, x_2)$  : a measure of similarity of  $x_1$  and  $x_2$ . It is a lot like a dot product

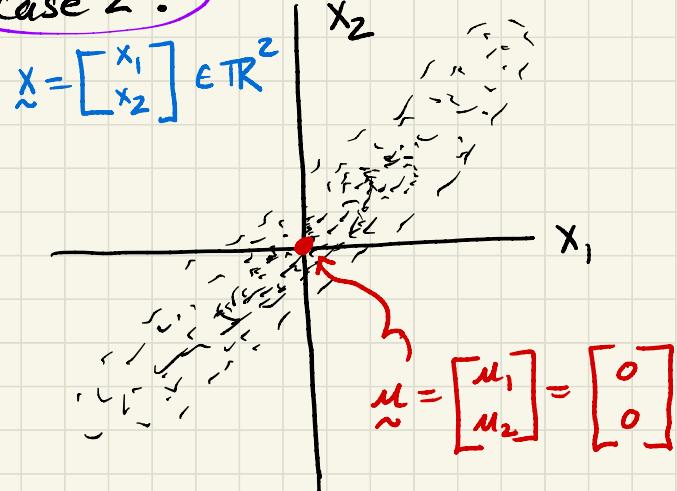
$$\begin{bmatrix} 1 & 0 \end{bmatrix} \begin{bmatrix} 1 \\ 0 \end{bmatrix} = 1$$

$$\begin{bmatrix} 1 & 0 \end{bmatrix} \begin{bmatrix} 0 \\ 1 \end{bmatrix} = 0$$

Case 1:



Case 2:

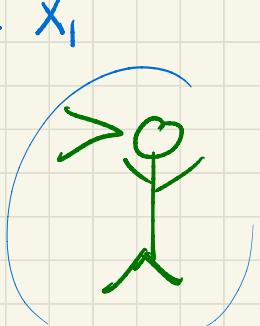
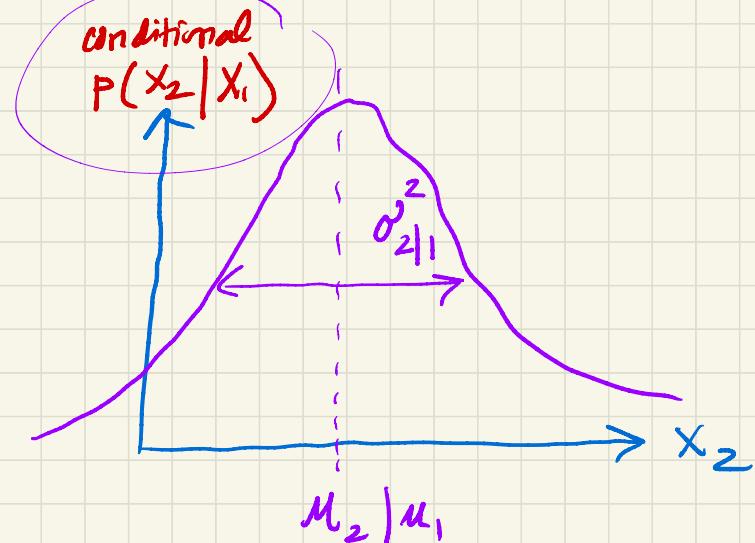
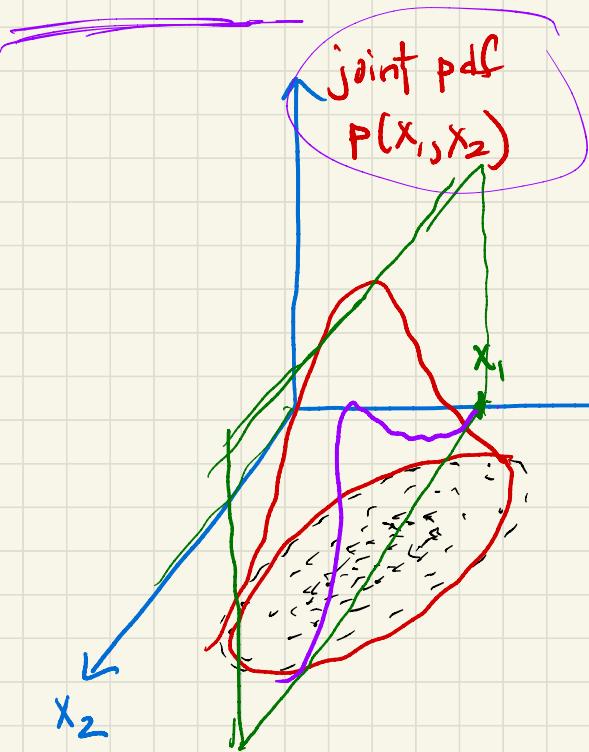


$$\begin{bmatrix} x_1 \\ x_2 \end{bmatrix} \sim N \left( \begin{bmatrix} \mu_1 \\ \mu_2 \end{bmatrix}, \begin{bmatrix} \Sigma_{11} & \Sigma_{12} \\ \Sigma_{21} & \Sigma_{22} \end{bmatrix} \right)$$

↑  
covariance matrix

$$\Sigma_{ij} = \text{cov}(x_i, x_j)$$

$$\begin{bmatrix} x_1 \\ x_2 \end{bmatrix} \sim N \left( \begin{bmatrix} \mu_1 \\ \mu_2 \end{bmatrix}, \begin{bmatrix} \Sigma_{11} & \Sigma_{12} \\ \Sigma_{21} & \Sigma_{22} \end{bmatrix} \right)$$



# MULTIVARIATE GAUSSIAN THEOREM

text: Kevin Murphy Ch 4  
Bishop Ch 2.3

(allows us to reason about the marginals & conditionals of a multivariate Gaussian.)

Suppose  $\underline{\underline{X}} = (\underline{x}_1, \underline{x}_2)$  is jointly Gaussian with parameters

$$\underline{\underline{\mu}} = \begin{bmatrix} \underline{\mu}_1 \\ \underline{\mu}_2 \end{bmatrix} \quad \underline{\Sigma} = \begin{pmatrix} \underline{\Sigma}_{11} & \underline{\Sigma}_{12} \\ \underline{\Sigma}_{21} & \underline{\Sigma}_{22} \end{pmatrix} \quad \underline{\underline{\Lambda}} = \underline{\Sigma}^{-1} = \begin{pmatrix} \underline{\Lambda}_{11} & \underline{\Lambda}_{12} \\ \underline{\Lambda}_{21} & \underline{\Lambda}_{22} \end{pmatrix}$$

Then:

the marginals are

$$p(\underline{x}_1) = N(\underline{x}_1 | \underline{\mu}_1, \underline{\Sigma}_{11})$$

$$p(\underline{x}_2) = N(\underline{x}_2 | \underline{\mu}_2, \underline{\Sigma}_{22})$$

the conditionals are

$$p(\underline{x}_1 | \underline{x}_2) = N(\underline{x}_1 | \underline{\mu}_{1|2}, \underline{\Sigma}_{1|2})$$

where

$$\begin{aligned} \underline{\mu}_{1|2} &= \underline{\mu}_1 + \underline{\Sigma}_{12} \underline{\Sigma}_{22}^{-1} (\underline{x}_2 - \underline{\mu}_2) \\ &= \underline{\mu}_1 - \underline{\Lambda}_{11}^{-1} \underline{\Lambda}_{12} (\underline{x}_2 - \underline{\mu}_2) \end{aligned}$$

$$\begin{aligned} \underline{\Sigma}_{1|2} &= \underline{\Sigma}_{11} - \underline{\Sigma}_{12} \underline{\Sigma}_{22}^{-1} \underline{\Sigma}_{21} \\ &= \underline{\Lambda}_{11}^{-1} \end{aligned}$$

So if Multivariate Gaussians on  $\mathbb{R}^n$   $\stackrel{\Sigma \text{ pos def}}{\equiv}$ ,  $\stackrel{\Sigma^{-1} \text{ also pos def.}}{\equiv}$

$$\underline{x} = \begin{bmatrix} x_1 \\ x_2 \\ x_3 \\ \vdots \\ x_n \end{bmatrix} \sim N\left(\begin{bmatrix} \underline{\mu}_1 \\ \underline{\mu}_2 \\ \underline{\mu}_3 \\ \vdots \\ \underline{\mu}_N \end{bmatrix}, \begin{bmatrix} \Sigma_{11} & \Sigma_{12} & \Sigma_{13} & \cdots & \Sigma_{1n} \\ \Sigma_{21} & \Sigma_{22} & \Sigma_{23} & \cdots & \Sigma_{2n} \\ \Sigma_{31} & \Sigma_{32} & \Sigma_{33} & \cdots & \Sigma_{3n} \\ \vdots & \vdots & \vdots & \ddots & \vdots \\ \Sigma_{n1} & \Sigma_{n2} & \Sigma_{n3} & \cdots & \Sigma_{nn} \end{bmatrix}\right)$$

mean  
 $\underline{\mu} \in \mathbb{R}^n$

covariance  $\Sigma \stackrel{n}{=} \mathbb{R}^{n \times n}$   
 positive definite

joint PDF  $p(\underline{x}) = \frac{1}{\sqrt{(2\pi)^n \det(\Sigma)}} \exp \left\{ -\frac{1}{2} (\underline{x} - \underline{\mu})^\top \Sigma^{-1} (\underline{x} - \underline{\mu}) \right\}$

① closure under linear transf:

$$\underline{x} \sim N(\underline{\mu}, \Sigma) \quad \underline{A}\underline{x} + \underline{b} \sim N(\underline{A}\underline{\mu} + \underline{b}), \quad \underline{\Sigma} \stackrel{A}{=} \underline{\Sigma} \underline{A}^\top$$

So if Multivariate Gaussians on  $\mathbb{R}^n$

$$\tilde{x} = \begin{bmatrix} x_1 \\ x_2 \\ x_3 \\ \vdots \\ x_n \end{bmatrix} \sim N\left(\begin{bmatrix} \mu_1 \\ \mu_2 \\ \mu_3 \\ \vdots \\ \mu_N \end{bmatrix}, \begin{bmatrix} \Sigma_{11} & \Sigma_{12} & \Sigma_{13} & \cdots & \Sigma_{1n} \\ \Sigma_{21} & \Sigma_{22} & \Sigma_{23} & \cdots & \Sigma_{2n} \\ \Sigma_{31} & \Sigma_{32} & \Sigma_{33} & \cdots & \Sigma_{3n} \\ \vdots & \vdots & \vdots & \ddots & \vdots \\ \Sigma_{n1} & \Sigma_{n2} & \Sigma_{n3} & \cdots & \Sigma_{nn} \end{bmatrix}\right)$$

mean  
 $\tilde{\mu} \in \mathbb{R}^n$

covariance  
 $\tilde{\Sigma} \in \mathbb{R}^n \times \mathbb{R}^n$  pos. def.

joint PDF  $p(\tilde{x}) = \frac{1}{\sqrt{(2\pi)^n \det(\tilde{\Sigma})}} \exp\left\{-\frac{1}{2} (\tilde{x} - \tilde{\mu})^T \tilde{\Sigma}^{-1} (\tilde{x} - \tilde{\mu})\right\}$

① addition & subtraction :  $\tilde{x} \sim N(\tilde{\mu}_x, \tilde{\Sigma}_x)$   $\tilde{y} \sim N(\tilde{\mu}_y, \tilde{\Sigma}_y)$

$$\tilde{x} + \tilde{y} \sim N(\tilde{\mu}_x + \tilde{\mu}_y, \tilde{\Sigma}_x + \tilde{\Sigma}_y)$$

So if Multivariate Gaussians on  $\mathbb{R}^n$

MVGT

$$\tilde{x} = \begin{bmatrix} x_1 \\ x_2 \\ x_3 \\ \vdots \\ x_n \end{bmatrix} \sim N\left(\begin{bmatrix} \mu_1 \\ \mu_2 \\ \mu_3 \\ \vdots \\ \mu_N \end{bmatrix}, \begin{bmatrix} \Sigma_{11} & \Sigma_{12} & \Sigma_{13} & \cdots & \Sigma_{1n} \\ \Sigma_{21} & \Sigma_{22} & \Sigma_{23} & \cdots & \Sigma_{2n} \\ \Sigma_{31} & \Sigma_{32} & \Sigma_{33} & \cdots & \Sigma_{3n} \\ \vdots & \vdots & \vdots & \ddots & \vdots \\ \vdots & \vdots & \vdots & \ddots & \vdots \\ \Sigma_{n1} & \Sigma_{n2} & \Sigma_{n3} & \cdots & \Sigma_{nn} \end{bmatrix}\right)$$

Red circles highlight the first element of each vector ( $x_1, \mu_1, \Sigma_{11}$ )

marginals are easy

$$p(x_1) = \int \dots \int p(x_1, \dots, x_n) dx_2 \dots d x_n = N(\mu_1, \Sigma_{11})$$

So if Multivariate Gaussians on  $\mathbb{R}^n$

$$x = \begin{bmatrix} x_1 \\ x_2 \\ \vdots \\ x_n \end{bmatrix} \sim N\left(\mu_A, \Sigma_A\right)$$

$\mu_A$

$\Sigma_A$

$$\Sigma_A = \begin{bmatrix} \Sigma_{11} & \Sigma_{12} & \Sigma_{13} & \cdots & \Sigma_{1n} \\ \Sigma_{21} & \Sigma_{22} & \Sigma_{23} & \cdots & \Sigma_{2n} \\ \Sigma_{31} & \Sigma_{32} & \Sigma_{33} & \cdots & \Sigma_{3n} \\ \vdots & \vdots & \vdots & \ddots & \vdots \\ \Sigma_{n1} & \Sigma_{n2} & \Sigma_{n3} & \cdots & \Sigma_{nn} \end{bmatrix}$$

marginals are easy : of sub-vector  $(x_1, x_2)$

$$p(x_1, x_2) = \int \dots \int p(x_1, x_2, \dots, x_n) dx_3 \dots dx_n = N(\mu_A, \Sigma_A)$$

So if

$$x = \begin{bmatrix} x_1 \\ x_2 \\ x_3 \\ \vdots \\ x_n \end{bmatrix} \sim N\left( \begin{bmatrix} \mu_A \\ \mu_B \end{bmatrix}, \Sigma \right)$$

where  $\Sigma$  is a block-diagonal matrix:

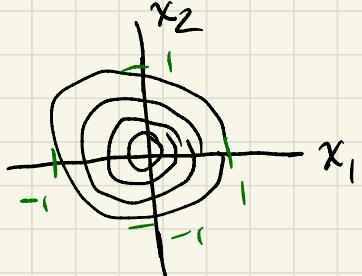
$$\Sigma = \begin{bmatrix} \Sigma_{AA} & & & & & \\ & \ddots & & & & \\ & & \Sigma_{AB} & & & \\ & & & \Sigma_{BA} & & \\ & & & & \Sigma_{BB} & \\ & & & & & \ddots \end{bmatrix}$$

with  $\Sigma_{AA}$  and  $\Sigma_{BB}$  being symmetric matrices.

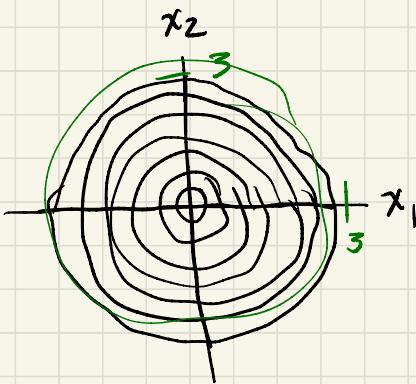
Conditionals are also not so bad

$$p(x_1, x_2 | \underbrace{x_3, \dots, x_N}_y) = N\left(\mu_A + \sum_{AB} \sum_B^{-1} (y - \mu_B), \sum_A - \sum_{BA} \sum_B^{-1} \sum_{AB}\right)$$

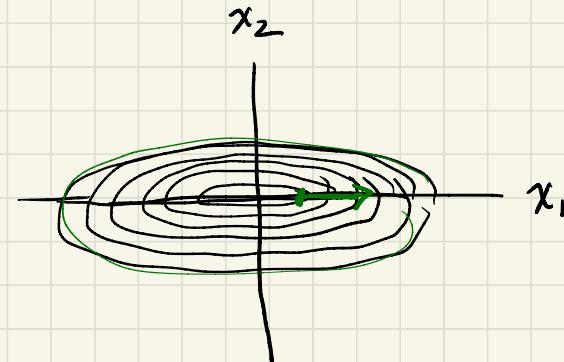
Back to bivariate case



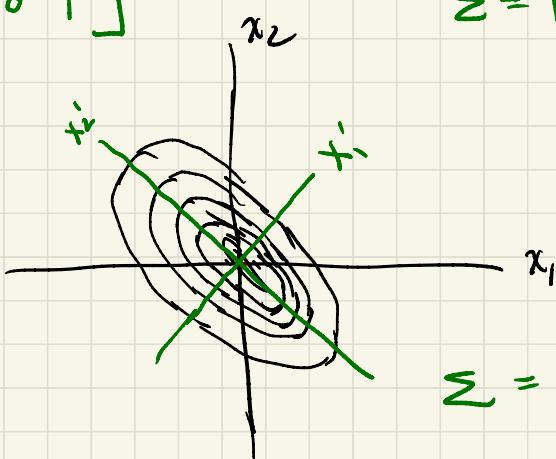
$$\Sigma = \begin{bmatrix} 1 & 0 \\ 0 & 1 \end{bmatrix}$$



$$\Sigma = \begin{bmatrix} 3 & 0 \\ 0 & 3 \end{bmatrix}$$



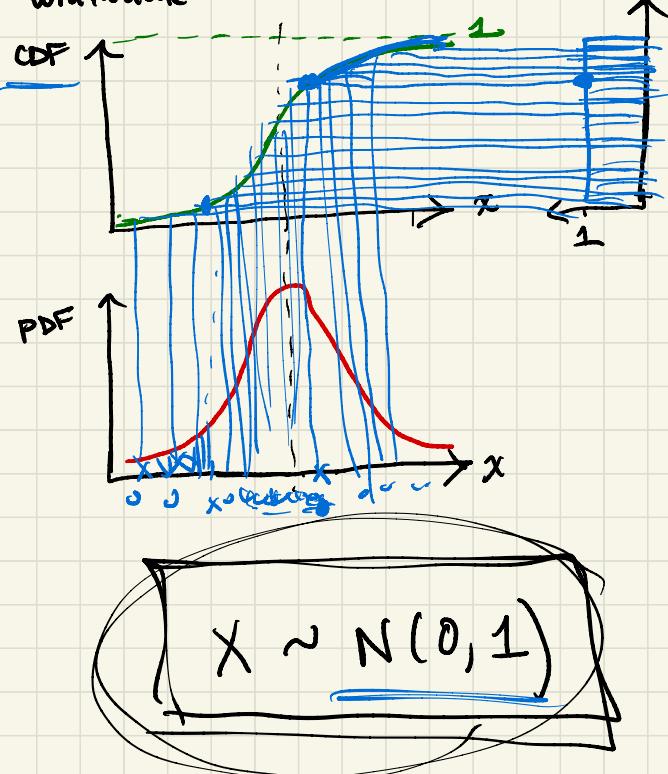
$$\Sigma = \begin{bmatrix} 3 & 0 \\ 0 & 1 \end{bmatrix}$$



$$\Sigma = \begin{bmatrix} 1 & -0.6 \\ -0.6 & 1 \end{bmatrix}$$

## How do we draw samples from an arbitrary multivariate Gaussian?

univariate case:



$$\text{univariate } x' = \underline{N(\mu, \sigma^2)}$$

$$x' = \mu + \sigma X$$

$$= \boxed{\mu + (\sigma N(0, 1))}$$

$\downarrow$  square root of variance

Multivariate

$$X = \begin{bmatrix} x_1 \\ x_2 \end{bmatrix} \sim N \left( \begin{bmatrix} 0 \\ 0 \end{bmatrix}, \begin{bmatrix} 1 & 0 \\ 0 & 1 \end{bmatrix} \right)$$

draw two gaussians,  $N(0, I)$

$$X' \sim N \left( \begin{bmatrix} \mu_1 \\ \mu_2 \end{bmatrix}, \begin{bmatrix} \Sigma_{11} & \Sigma_{12} \\ \Sigma_{21} & \Sigma_{22} \end{bmatrix} \right)$$

$\underline{\mu}$        $\underline{\Sigma}$

$$X' = \mu + L N(0, I)$$

cholesky decomposition  $\Sigma = LL^T$

(17)