

TAM 598 Lecture 14 :

Bayesian Linear Regression

Announcements:

- Hw 4 covers lectures 13-16; due on Mar 31 (Mon)
- 1.5 recorded class over spring break
- No class Monday Mar 24th and Mon Mar 31

Least Squares Linear Regression - "traditional"

Can we interpret it from a more "modern" Bayesian or probabilistic way?

observations $\underline{x}_{1:n}$, outputs $y_{1:n}$

generalized linear model w/ m basis functions:

model the measurement process using a likelihood function

assume that outcomes of single measurements are Gaussian, with mean $\underline{w}^T \underline{\phi}(\underline{x})$ and noise variance σ^2

notation: $N(y|\mu, \sigma^2)$ is a pdf

$$p(y) = (2\pi\sigma^2)^{-1/2} \exp\left\{-\frac{(y-\mu)^2}{2\sigma^2}\right\}$$

for independent measurements, the likelihood of the data factorizes

$$P(\underline{y}_{1:n} | \underline{x}_{1:n}, \underline{w}) =$$

I How do we find the parameters? Maximum likelihood for weights \underline{w} , and σ^2

$$\max_{\underline{w}, \sigma^2}$$

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So: maximizing likelihood wrt \underline{w} is equivalent to minimizing sum of squares! And our weights should therefore satisfy

AND also we can estimate σ^2 as well via max likelihood

Now: we can incorporate uncertainty σ^2 when making predictions

point predictive distribution : $P(y | \underline{x}, \underline{w}, \sigma^2) = N(y | \underline{w}^\top \phi(\underline{x}), \sigma^2)$

the measured output y is normal distributed around the model (least squares) prediction, with variance σ^2

II How do we find the parameters? Maximum a posteriori estimates

here: maximize the log prob of the posterior rather than the likelihood. Helps avoid overfitting

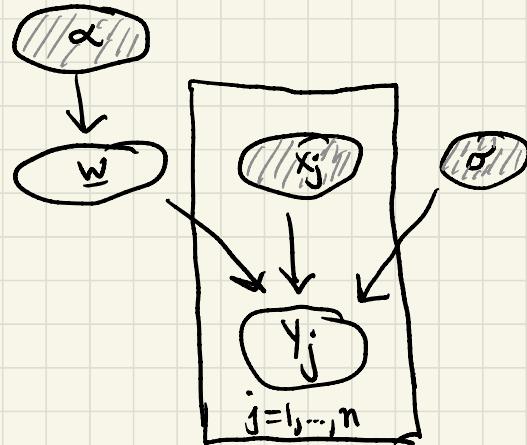
again: $y(x; w) = \underline{w}^T \underline{\phi}(x)$

$$y_{1:n} \mid x_{1:n}, w, \sigma^2 \sim N(\underline{w}^T \underline{\phi}(x), \sigma^2)$$

now: uncertainty in model parameters described by a prior $w \sim p(w)$

e.g.) gaussian prior
on weights

graphically



posterior from Bayes' Rule

$$p(\underline{w} | \underline{x}_{1:n}, \underline{y}_{1:n}, \sigma, \alpha) =$$

$$\frac{P(\underline{y}_{1:n} | \underline{x}_{1:n}, \underline{w}, \sigma, \alpha) P(\underline{w} | \alpha)}{\int P(\underline{y}_{1:n} | \underline{x}_{1:n}, \underline{w}', \sigma, \alpha) P(\underline{w}' | \alpha) d\underline{w}'}$$

our state of knowledge about \underline{w} , after we see the data

Now assuming σ and α are known.

a point estimate of \underline{w} is

$$\underline{w}_{MPE} = \operatorname{argmax}_{\underline{w}} p(\underline{y}_{1:n} | \underline{x}_{1:n}, \underline{w}, \sigma) p(\underline{w} | \alpha)$$

For gaussian likelihood and weight prior :

$$\log p(\underline{w} | \underline{x}_{1:n}, \underline{y}_{1:n}, \sigma, \alpha) =$$

maximizing :

III How do we find the parameters? Bayesian Linear Regression

similar to II :

$$\left[\begin{array}{l} \text{again: } y(x; \underline{w}) = \underline{w}^T \underline{\phi}(x) \\ Y_{1:n} \mid X_{1:n}, \underline{w}, \sigma^2 \sim N(\underline{w}^T \underline{\phi}(x), \sigma^2) \\ \text{uncertainty in model parameters described by a prior } \underline{w} \sim p(\underline{w}) \end{array} \right]$$

but no point estimate. Keep the posterior distribution and work with it directly. Why? Can now quantify the epistemic uncertainty arising from limited # of observations.

$$\begin{matrix} \text{posterior} & p(\underline{w} \mid X_{1:n}, Y_{1:n}, \sigma, \alpha) = \\ \text{is gaussian} & \end{matrix}$$

posterior predictive distribution: what can we say about y at some new \underline{x} after seeing the data?

$$P(y | \underline{x}, \underline{x}_{1:n}, \underline{y}_{1:n}, \sigma^2, \alpha) =$$

for all gaussian priors, this is analytically available

$$P(y | \underline{x}, \underline{x}_{1:n}, \underline{y}_{1:n}, \sigma^2, \alpha) = N(y | m(\underline{x}), s^2(\underline{x}))$$

$$\text{where } m(\underline{x}) =$$

$$\hat{s}^2(\underline{x}) =$$