

Joint Probability Mass Function

consider two discrete random variables X and Y . The joint probability mass function $f_{XY}(x,y)$ gives us the probability that $X = x$ and $Y = y$.

$$p(x,y) = p(X=x, Y=y) = f_{XY}(x,y) := p(\{w : X(w)=x, Y(w)=y\})$$

properties of $p(x,y)$:

① non-negative $p(x,y) \geq 0$

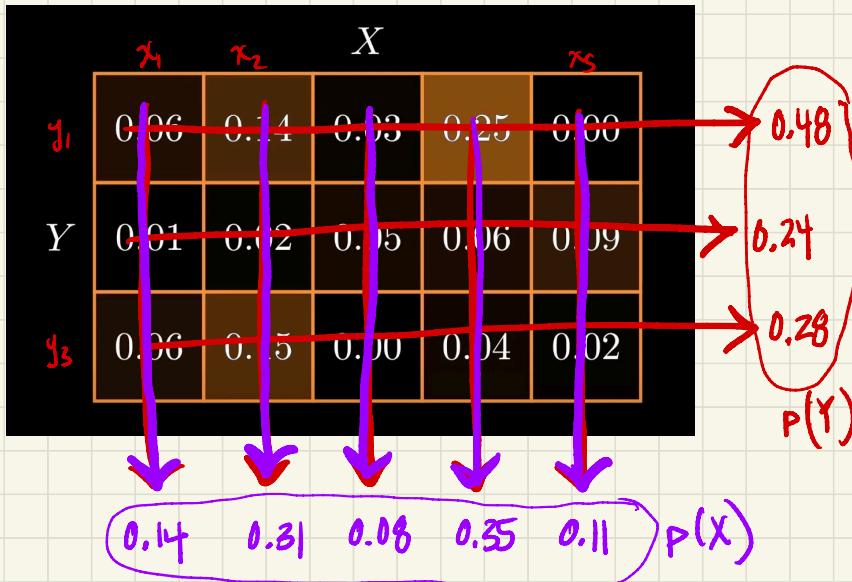
② normalized $\sum_x \sum_y p(x,y) = 1$

③ marginalization: if you marginalize one of the random variables values, you get the other's pmf

$$p(x) = \sum_y p(x,y) \quad p(y) = \sum_x p(x,y)$$

MARGINALIZATION : a marginal distribution contains a subset of variables

e.g) pmf of X, Y



"What distribution would I get for X if I ignored Y ?"

for $p(x, y)$, two marginals are available: $p(x), p(y)$

$$p(x) = \sum_y p(x, y)$$

sum out the other variables

$$p(y) = \sum_x p(x, y)$$

e.g samples from X and Y together
but somebody chose and threw away all the Y s

JOINT PROBABILITY DENSITY FUNCTION

- ④ applies now to two continuous random variables X, Y
- ⑤ $f_{XY}(x,y)$ is the function that yields the probability that (X,Y) belong to any Borel subset A of \mathbb{R}^2

$$P((X,Y) \in A) = \iint_A \underbrace{f_{XY}(x,y)}_{\text{or } p(x,y)} dx dy$$

- ⑥ $f_{XY}(x,y)$ is also non-negative and normalized. And, we can again marginalize out one variable at a time:

$$p(x) = \int_{-\infty}^{\infty} p(x,y) dy \quad p(y) = \int_{-\infty}^{\infty} p(x,y) dx$$

CONDITIONING A RANDOM VARIABLE ON ANOTHER $p(x|y)$

- ① a conditional distribution fixes a subset of variables. "What distribution would I get for X if I set $Y=y$?"
- ② for both discrete and continuous, divide the joint by the marginal

$$P(A|B) = \frac{P(A,B)}{P(B)}$$

A, B are logical sentences.
this is just the product rule from Lect. 2

let $A = (X=x)$
 $B = (Y=y)$

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$$p(x|y) = \frac{p(x,y)}{p(y)}$$

- ③ we can also say

$$p(x|y) = \frac{p(x,y)}{\int_{-\infty}^{\infty} p(x,y) dx'}$$

← knowing the joint pdf, can find the conditional of one given the other

eg) pmf of X, Y

		X		
		x_1	x_2	x_3
y_1		0.06	0.14	0.03
Y	0.01	0.02	0.05	0.06
	0.09			
	0.06	0.15	0.00	0.04

$$P(X | Y = y_2)$$

$\rightarrow 0.24$

$$\times \frac{1}{0.24} \quad \times \frac{1}{0.24} \quad \times \frac{1}{0.24} \quad \times \frac{1}{0.24} \quad \times \frac{1}{0.24}$$

0.05	0.10	0.21	0.25	0.39
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$$\rightarrow P(X | Y = y_2)$$

$$P(x_1 | y_2) \dots$$

EXPECTATION OF TWO RANDOM VARIABLES

let X, Y be two random variables. Let $Z = g(X, Y)$ be a third random variable. Then the expectation of Z is

$$E[Z] = E[g(X, Y)] = \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} g(x, y) p(x, y) dx dy$$

if $Z = X + Y$, $E[X+Y] = E[X] + E[Y]$

INDEPENDENCE, DEPENDENCE, COVARIANCE, AND CORRELATION

- ① Two random variables X, Y are **independent** if they don't influence each other. Knowing $X = x$ conveys no info about Y

Joint probabilities
are interesting
only when they
capture
dependence
between
variables

- ② Independence means that the joint distribution factorizes

$$P(X, Y) = P(X) P(Y)$$

we write

$$X \perp Y$$

This also implies that the two conditionals are equal to their marginals:

$$\begin{aligned} P(X|Y) &= \frac{P(X,Y)}{P(Y)} \\ &= \frac{P(X)P(Y)}{P(Y)} = P(X) \end{aligned}$$

$$P(X|Y) = P(X) \quad \text{and} \quad P(Y|X) = P(Y)$$

- ③ Independence also tells us that the expectation of X, Y also factorizes

$$E[XY] = E[X]E[Y]$$

Example: $p(x, Y)$ pmf

if independent, only need 8 variables
to obtain the full matrix

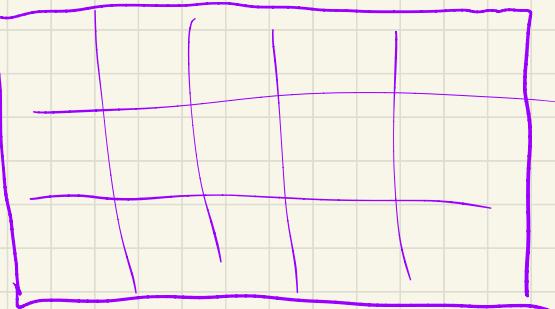
		X				
		0.03	0.03	0.03	0.05	0.07
Y		0.04	0.05	0.04	0.07	0.10
		0.07	0.08	0.06	0.12	0.16

3x5 matrix, Elements sum to 1

=

0.21
0.29
0.50

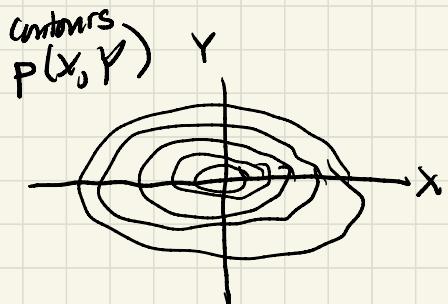
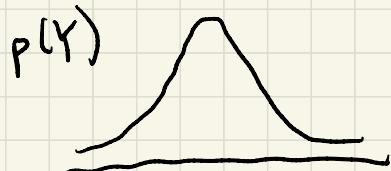
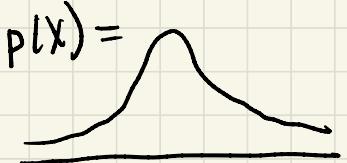
0.14	0.16	0.13	0.24	0.33
$p(x)$				



this is a "rank 1" matrix

$p(Y)$
marginal

Example: two independent continuous RVs



conditional independence: R.V.s X and Y are conditionally independent given Z if

$$P(X, Y | Z = z) = P(X | Z = z) P(Y | Z = z)$$

for all $Z = z$.

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comes up a lot  
when reasoning  
about  
latent  
structure  
in data

example: customers buy umbrellas at a store

$B_i$  = event that customer  $i$  buys an umbrella  $i=1, 2, \dots$

$R$  = event that it is raining today

say  $P(B_i|R) = 0.7$

$$P(B_i|\neg R) = 0.1$$

$$P(R) = 0.25$$

(marginalize)

$$\begin{aligned} P(B_i) &= P(B_i|R)P(R) + P(B_i|\neg R)P(\neg R) \\ &= (0.7)(0.25) + (0.1)(0.75) \\ &= 0.25 \end{aligned}$$

consider joint probability

$$\begin{aligned} P(B_1, B_2) &= P(B_1, B_2, R) + P(B_1, B_2, \neg R) \\ &= P(B_1, B_2|R)P(R) + P(B_1, B_2|\neg R)P(\neg R) \\ &= P(B_1|R)P(B_2|R)P(R) + P(B_1|\neg R)P(B_2|\neg R)P(\neg R) \\ &= (0.7)(0.7)(0.25) + (0.1)(0.1)(0.75) = 0.13 \end{aligned}$$

$B_1$  and  $B_2$  are not indie  
but they are  
cond. indie given the  
rainy condition that its  
raining

marginalize

sum rule

$B_1|R$   
and  $B_2|R$   
are indep  
(cond indep)

Note  $P(B_1, B_2) \neq P(B_1)P(B_2) = (0.25)(0.25) = 0.0625$

Look at some statistical relationships related to dependence, independence

the COVARIANCE OPERATOR measures linear dependence between two R.V.s

$$C(X, Y) = E[(X - E[X])(Y - E[Y])]$$

$$= E[XY - XE[Y] - YE[X] + E[X]E[Y]]$$

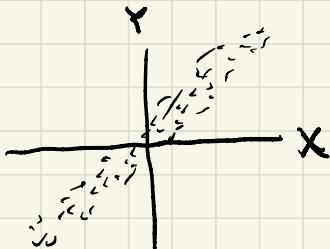
$$= E[XY] - E[X]E[Y]$$

do two RV's deviate from their means in a coincident way

both  $>$  mean  
or less mean  
at same time,  
 $Cov > 0$ .

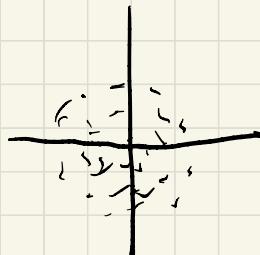
one  $>$ , one  $<$   
then neg

example:



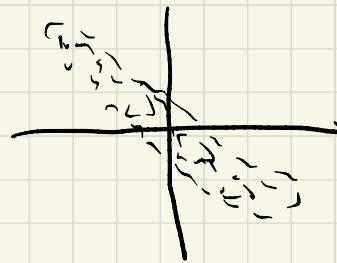
$$C(X, Y) = 2$$

X, Y positively correlated



$$C(X, Y) = 0$$

no linear dependence

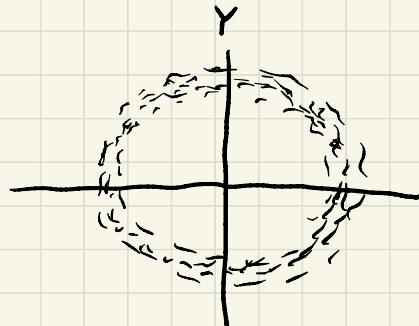


$$C(X, Y) = -3$$

X, Y negatively correlated

be aware: covariance only captures linear dependence. Just because  $C(X, Y) = 0$  does not mean that  $X$  and  $Y$  are independent

eg



- ① Knowing  $X$  gives info about  $Y$ , and v.v.  $\Rightarrow$  not independent
- ② but  $C(X, Y) = 0$  because the dependence is not linear.

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properties of covariance:

① symmetric $C(X, Y) = C(Y, X)$

② linear $C(ax + bY, Z) = aC(X, Z) + bC(Y, Z)$

③ relationship to variance

$$\mathbb{V}(X+Y) = \mathbb{V}(X) + \mathbb{V}(Y) + 2C(X, Y)$$

① relationship to variance

$$\mathbb{V}[x+y] = \mathbb{V}[x] + \mathbb{V}[y] + 2C(x,y)$$

$$\begin{aligned} \text{pf: } \mathbb{V}[x+y] &= \mathbb{E}[(x+y) - \mathbb{E}(x+y)]^2 \\ &= \mathbb{E}[(x+y - \mathbb{E}[x] - \mathbb{E}[y])^2] \\ &= \mathbb{E}\left[(x - \mathbb{E}[x])^2 + (y - \mathbb{E}[y])^2\right. \\ &\quad \left.+ 2\mathbb{E}[(x - \mathbb{E}[x])(y - \mathbb{E}[y])]\right] \\ &= \mathbb{V}[x] + \mathbb{V}[y] + 2C(x,y) \end{aligned}$$

interpret
 $C(x,y) \rightarrow 0$:
independence

The PEARSON CORRELATION COEFFICIENT:
(height - m, ft : weight)

- ① covariance is sensitive to units, and it's generally unknown what range it can take
- ② correlation coefficient addresses this problem by normalizing the covariance by each of the std deviations

$$\text{corr}(x, y) = \rho_{x,y} = \frac{C(x, y)}{\sqrt{V[x] V[y]}} \in [-1, 1]$$

i.e.) define new RVs $X' = X / \sqrt{V[X]}$ and $Y' = Y / \sqrt{V[Y]}$
has $V[X'] \in [-1, 1] \dots$

$$\text{then } \rho_{x,y} = C(X', Y')$$

so correlation is scale free.