

1. Review some important theoretical concepts

- Law of iterated expectations (LIE)

$$E(Y) = E_X(E[Y|X])$$

- Populations, parameters, samples, statistics... and their distributions!

$$Y \sim f(y; \theta) \quad \theta = (\mu_Y, \sigma_Y^2) \quad \{Y_1, Y_2, \dots, Y_N\} \quad T(Y_1, \dots, Y_N)$$

- Estimators (aka statistics) vs estimates

$$T(Y_1, \dots, Y_N) \quad T(y_1, \dots, y_N)$$

- Properties: finite sample (bias, variance, efficiency) vs “large” sample or asymptotic (consistency, asymptotic variance)

$$T(Y_1, \dots, Y_N) \text{ is unbiased for } \theta \text{ whenever } E[T(Y_1, \dots, Y_N)] = \theta$$

$$T(Y_1, \dots, Y_N) \text{ is consistent for } \theta \text{ whenever } \text{plim}_{N \rightarrow \infty} T(Y_1, \dots, Y_N) = \theta$$

$$\text{Note that plim is equivalent to the regular limits: } \lim_{N \rightarrow \infty} \text{Bias}[T(Y_1, \dots, Y_N)] = 0$$

$$\lim_{N \rightarrow \infty} \text{Var}[T(Y_1, \dots, Y_N)] = 0$$

- Law of large numbers (LLN)

If $\{Y_1, \dots, Y_N\}$ is a random sample from population $Y \sim f(y; \mu_Y, \sigma_Y^2)$, then

$$\text{plim}_{N \rightarrow \infty} \bar{Y} = \mu_Y$$

In other words, the mean of a random sample $\bar{Y} = \frac{1}{N} \sum_{i=1}^N Y_i$ is always consistent

for the population mean $\mu_Y := E[Y]$ of any random variable Y !

- Central limit theorem (CLT)

If $\{Y_1, \dots, Y_N\}$ is a random sample from population $Y \sim f(y; \mu_Y, \sigma_Y^2)$ with finite variance $\sigma_Y^2 < \infty$, then the “standardized” random variable

$$Z := \frac{\bar{Y} - \mu_Y}{\sigma_Y}$$

has an asymptotic distribution that is standard normal: $Z \sim N(0,1)$ as $N \rightarrow \infty$.

- Regression as conditional expectation

- Prove that the method of moments estimator $\hat{\sigma}_{MOM}^2 := \frac{1}{N} \sum_{i=1}^N (X_i - \bar{X})^2$ is consistent for population variance parameter $\sigma_X^2 := E[(X - \mu_X)^2]$ of random variable $X \sim f_X(\mu_X, \sigma_X^2)$.

$$\begin{aligned} \text{plim}_{N \rightarrow \infty} \hat{\sigma}_{MOM}^2 &= \text{plim}_{N \rightarrow \infty} \frac{1}{N} \sum_{i=1}^N (X_i - \bar{X})^2 = \text{plim}_{N \rightarrow \infty} \frac{1}{N} \sum_{i=1}^N (X_i^2 - 2X_i\bar{X} + \bar{X}^2) \\ &= \text{plim}_{N \rightarrow \infty} \left[\frac{1}{N} \sum_{i=1}^N X_i^2 - \bar{X}^2 \right] = \text{plim}_{N \rightarrow \infty} \left[\frac{1}{N} \sum_{i=1}^N X_i^2 \right] - \text{plim}_{N \rightarrow \infty} \bar{X}^2 \\ &= E[X^2] - \left[\text{plim}_{N \rightarrow \infty} \bar{X} \right]^2 = E[X^2] - E[X]^2 \\ &= E[(X - \mu_X)^2] = \sigma_X^2 \end{aligned}$$

- Let $\{X_1, \dots, X_N\}$ denote a random sample of size N from $X \sim f_X(\mu_X, \sigma_X^2)$. Consider the following candidate estimators of the population mean:

$$\hat{\mu}_1 := \bar{X} + \frac{1}{N}$$

$$\hat{\mu}_2 := 0.9 \cdot \bar{X}$$

$$\hat{\mu}_3 := \frac{X_1 + X_N}{2}$$

Which estimators are unbiased for μ_X ? Which are consistent for μ_X ?

$$\begin{aligned} E[\hat{\mu}_1] &:= E\left[\bar{X} + \frac{1}{N}\right] = E\left[\frac{1}{N} \sum_{i=1}^N X_i\right] + E\left[\frac{1}{N}\right] = \frac{1}{N} E\left[\sum_{i=1}^N X_i\right] + \frac{1}{N} \\ &= \frac{1}{N} \sum_{i=1}^N E[X_i] + \frac{1}{N} = \frac{1}{N} \sum_{i=1}^N \mu_X + \frac{1}{N} = \frac{1}{N} N \mu_X + \frac{1}{N} = \mu_X + \frac{1}{N} \neq \mu_X \rightarrow \text{biased!} \end{aligned}$$

$$E[\hat{\mu}_2] := E[0.9 \cdot \bar{X}] = 0.9 \cdot E[\bar{X}] = 0.9 \cdot \mu_X \neq \mu_X \rightarrow \text{biased!}$$

$$E[\hat{\mu}_3] := E\left[\frac{X_1 + X_N}{2}\right] = \frac{1}{2} \cdot E[X_1 + X_N] = \frac{1}{2} (\mu_X + \mu_X) = \mu_X \rightarrow \text{unbiased!}$$

$$\text{plim}_{N \rightarrow \infty} \hat{\mu}_1 = \text{plim}_{N \rightarrow \infty} \left[\bar{X} + \frac{1}{N} \right] = \text{plim}_{N \rightarrow \infty} [\bar{X}] + \text{plim}_{N \rightarrow \infty} \left[\frac{1}{N} \right] = \mu_X \quad \rightarrow \text{consistent!}$$

$$\text{plim}_{N \rightarrow \infty} \hat{\mu}_2 = \text{plim}_{N \rightarrow \infty} [0.9 \cdot \bar{X}] = 0.9 \cdot \text{plim}_{N \rightarrow \infty} [\bar{X}] = 0.9 \cdot \mu_X \neq \mu_X \quad \rightarrow \text{not consistent!}$$

$$\text{Since } \text{Var}(\hat{\mu}_3) = \frac{1}{2} \sigma_X^2, \text{ we have } \lim_{N \rightarrow \infty} \frac{1}{2} \sigma_X^2 \neq 0 \quad \rightarrow \text{not consistent!}$$

4. Let $Y = \beta_0 + \beta_1 X + e$ denote a linear population regression function. Prove that whenever $\text{Cov}(X, e) = 0$ we can write the values of $\{\beta_0, \beta_1\}$ in terms of $E(X)$, $E(Y)$, $\text{Var}(Y)$, and $\text{Cov}(X, Y)$. What is the economic meaning behind the assertion that the value of the parameter $\text{Cov}(X, e)$ must be $= 0$ in the population? (Bonus: if it fails, why would an infinite sample, or the whole population, be useless for establishing causality?)

$$\text{Cov}(X, e) = 0$$

$$\text{Cov}(X, Y - \beta_0 - \beta_1 X) = 0$$

$$\text{Cov}(X, Y) - \text{Cov}(X, \beta_0) - \text{Cov}(X, \beta_1 X) = 0$$

$$\text{Cov}(X, Y) - 0 - \beta_1 \text{Cov}(X, X) = 0$$

$$\text{Cov}(X, Y) - \beta_1 \text{Var}(X) = 0$$

$$\beta_1 = \frac{\text{Cov}(X, Y)}{\text{Var}(X)} = \frac{E[XY] - E[X]E[Y]}{E[X^2] - E[X]^2}$$

$$\Rightarrow \hat{\beta}_1^{MOM} := \frac{\widehat{\text{Cov}(X, Y)}}{\widehat{\text{Var}(X)}} = \hat{\beta}_1^{OLS}$$

$$E[Y] = E[\beta_0 + \beta_1 X + e]$$

$$E[Y] = \beta_0 + \beta_1 E[X] + E[e]$$

$$E[Y] = \beta_0 + \beta_1 E[X] + 0$$

$$\beta_0 = E[Y] - \beta_1 E[X]$$

$$\Rightarrow \hat{\beta}_0^{MOM} := \widehat{E[Y]} - \hat{\beta}_1^{MOM} \cdot \widehat{E[X]} = \bar{Y} - \hat{\beta}_1^{MOM} \bar{X} = \hat{\beta}_1^{OLS}$$