

# Bayesian Modelling of Metabolic Syndrome Risk in AVIS Blood Donors

## An analysis of routine laboratory data

Noemi Bongiorni, Martina Caliandro, Davide Marchesi, Greta Minazzi, Elisa Nordera, Matteo Zanetti

**Tutor:** Simone Colombara

Bayesian Statistics 2025/2026

February 19, 2026



POLITECNICO  
MILANO 1863

- ▶ 100,203 laboratory measurements.
- ▶ 4,300 donors - longitudinal data.
- ▶ 5 target variables: Blood pressure, BMI, Glucose, HDL Cholesterol, Triglycerides.

## Research question:

Is it feasible to implement a **Metabolic Syndrome screening**, based on AVIS laboratory measurements?  
Is this able to identify the predictive biomarkers for high-risk patients and support clinicians in treatment strategies?



# Modeling framework

Hierarchical multivariate Bayesian model:

$$\mathbf{y}_{ij} \mid \boldsymbol{\mu}_{ij}, \Sigma \stackrel{\text{ind}}{\sim} \mathcal{N}_K(\boldsymbol{\mu}_{ij}, \Sigma), \quad i = 1, \dots, I, j = 1, \dots, n_i$$

$$\boldsymbol{\mu}_{ij} = \mathbf{x}_{ij} \boldsymbol{\beta} + \mathbf{b}_i$$

$K$ : number of target variables (5)

$I$ : number of donors (35)

$n_i$ : number of visits per donor

$\boldsymbol{\beta}$ : fixed effects

$b_i$ : donor specific random intercept

Residual covariance: LKJ

$$\Sigma = \mathbf{L}_\Sigma \mathbf{L}_\Sigma^\top, \quad \mathbf{L}_\Sigma = \text{diag}(\boldsymbol{\tau}) \mathbf{L}_\Omega, \quad \mathbf{L}_\Omega \sim \text{LKJ-Cholesky}(4)$$
$$\tau_k \stackrel{\text{iid}}{\sim} \mathcal{N}^+(0, 0.5)$$

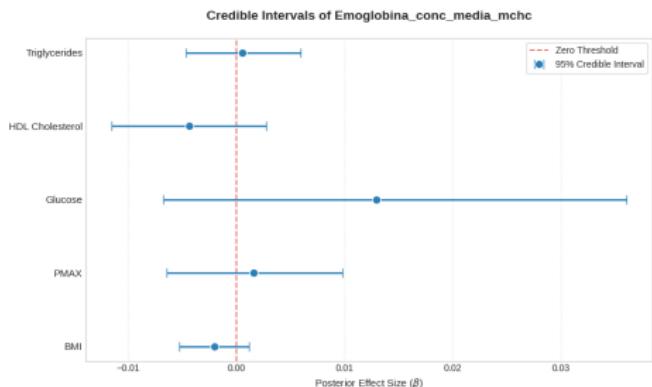
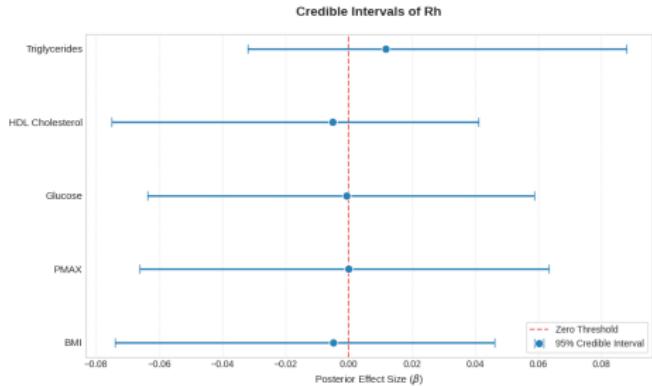
Random effects: Gaussian

$$\mathbf{b}_i = \mathbf{L}_{\Sigma_b} \mathbf{z}_{b_i}, \quad \mathbf{z}_{b_i} \stackrel{\text{iid}}{\sim} \mathcal{N}_K(\mathbf{0}, \mathbf{I}_K)$$
$$\mathbf{L}_{\Sigma_b} = \text{diag}(\boldsymbol{\tau}_b) \mathbf{L}_{\Omega_b}$$
$$\mathbf{L}_{\Omega_b} \sim \text{LKJ-Cholesky}(4)$$
$$\tau_{b_k} \stackrel{\text{iid}}{\sim} \mathcal{N}^+(0, 0.5)$$

Fixed effects: Bayesian Lasso

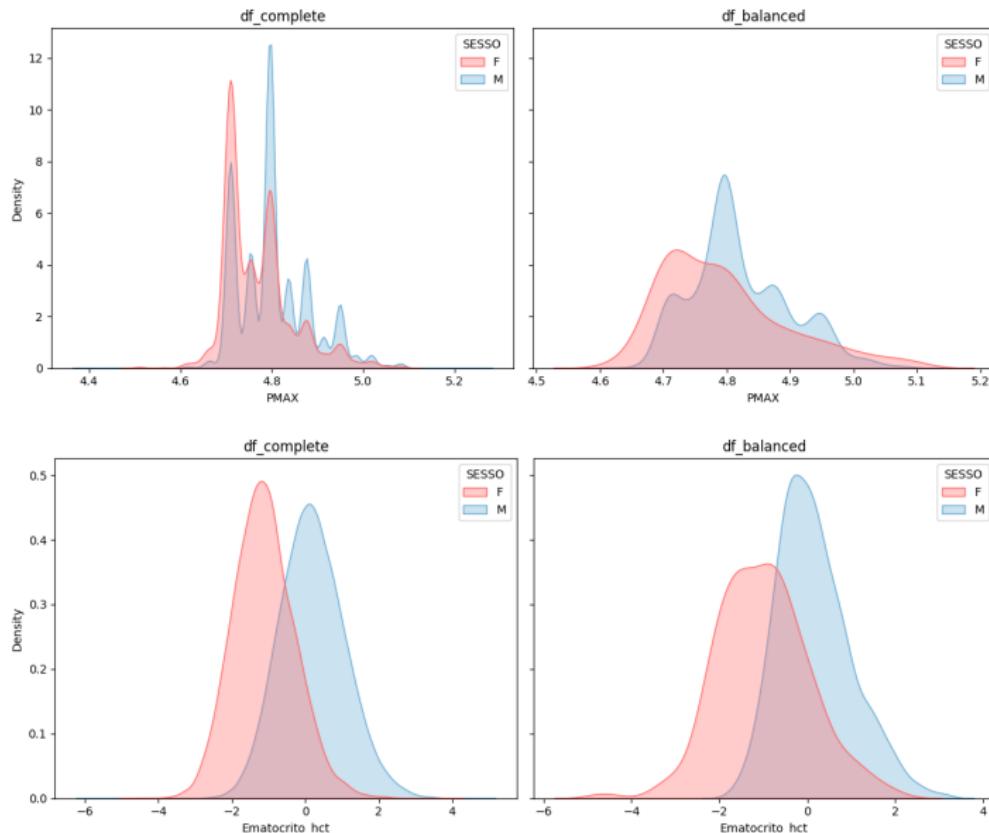
$$\beta_{pk} \mid \tau_{pk} \stackrel{\text{ind}}{\sim} \mathcal{N}(0, \tau_{pk})$$
$$\tau_{pk} \stackrel{\text{iid}}{\sim} \text{Exponential}\left(\frac{\lambda^2}{2}\right)$$
$$\lambda \sim \text{Gamma}(1, 1)$$

# Variable selection



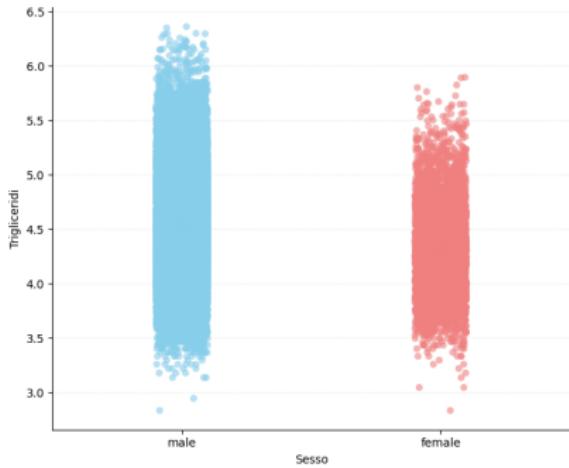
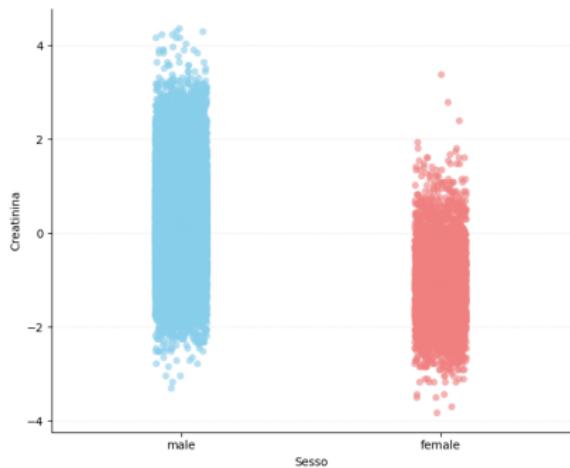
**Non-informative:**  
Basophili\_perc  
Creatinina  
Emoglobina  
Ferritina  
Ferro\_totale  
Data  
Data\_nascita  
Rh  
ABO  
SESSO .

# Covariate Sesso



# Covariate Sesso

Creatinina and Trigliceridi values of male and females





# DPMM for Random Effects

Idea:

$$\mathbf{b}_i \mid \boldsymbol{\theta}_i \stackrel{iid}{\sim} k(\cdot; \boldsymbol{\theta}_i)$$

$$\boldsymbol{\theta}_i \mid P \stackrel{iid}{\sim} P$$

$$P \sim DP(\alpha, P_0)$$

$k(\cdot; \boldsymbol{\theta}_i)$  is a multivariate Gaussian density

$i = 1, \dots, I$  (35): donor index

Actually:

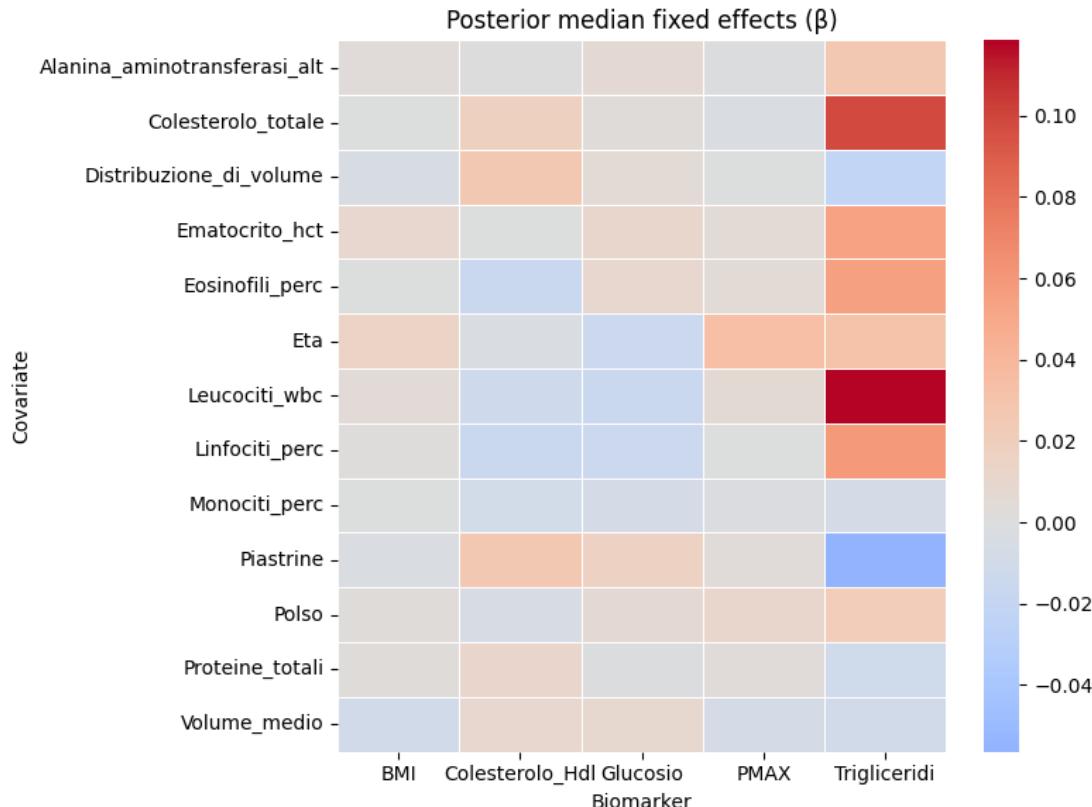
$$P = \sum_{m=1}^M \pi_m \delta_{\boldsymbol{\theta}_m^*}, \quad \pi_1, \dots, \pi_M : \text{stick-breaking weights}$$

$\boldsymbol{\theta}_m^* = (\boldsymbol{\mu}_m, \boldsymbol{\Sigma}_m), \quad m = 1, \dots, M$  (20): cluster index

$\boldsymbol{\mu}_m \sim \mathcal{N}_K(\mathbf{0}, 2^2 \mathbf{I}), \quad \boldsymbol{\Sigma}_m$  with LKJ prior

$K = 5$ : number of targets

# Fixed effects



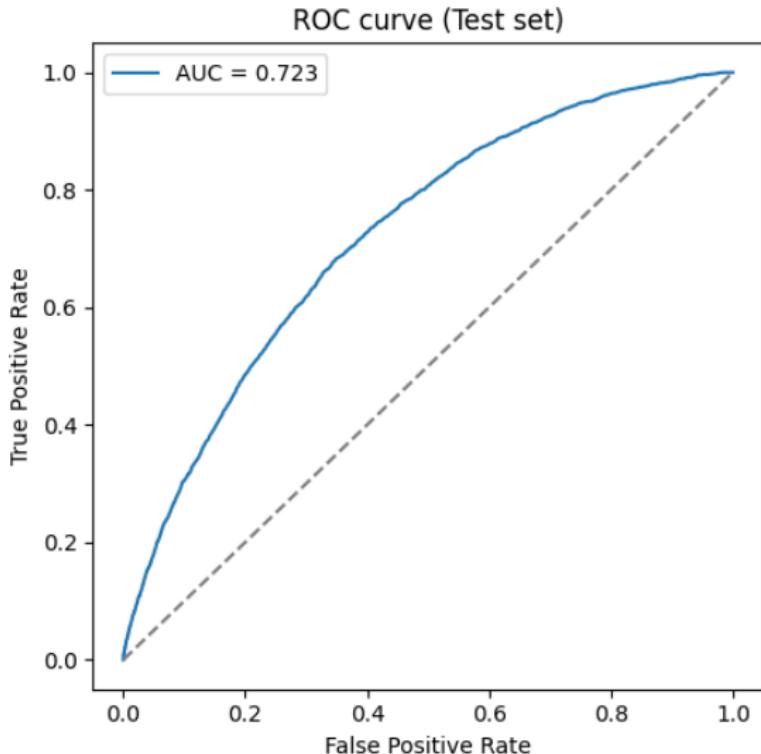
## Cluster-level mean random effects

$$\mu_m = \begin{array}{c|ccccc} & \text{PMAX} & \text{Glucosio} & \text{Trigliceridi} & \text{HDL} & \text{BMI} \\ \text{Cluster 1} & 4.7929 & 4.5373 & 4.6824 & 3.9904 & 3.2908 \end{array}$$

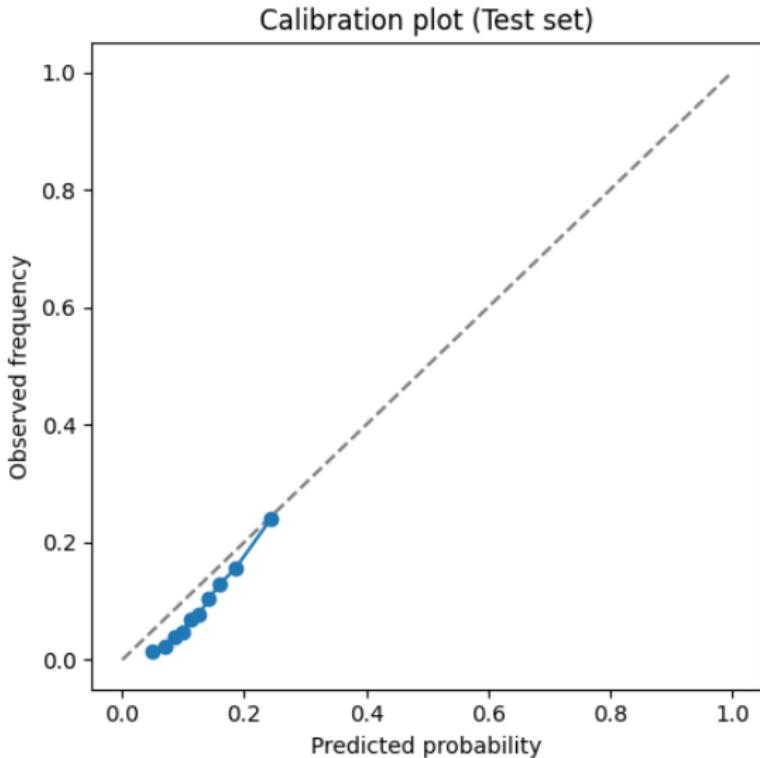
## Covariance matrix of random effects

$$\Sigma_m = \begin{array}{c|ccccc} & P & G & T & H & B \\ \hline P & 0.0020 & 0.0008 & 0.0014 & -0.0010 & 0.0016 \\ G & 0.0008 & 0.0044 & 0.0015 & -0.0030 & 0.0021 \\ T & 0.0014 & 0.0015 & 0.0856 & -0.0279 & 0.0131 \\ H & -0.0010 & -0.0030 & -0.0279 & 0.0303 & -0.0051 \\ B & 0.0016 & 0.0021 & 0.0131 & -0.0051 & 0.0114 \end{array}$$

# Predictive performance: ROC curve



# Predictive performance: calibration plot



# Conclusions

- ▶ Strong predictive performance.
- ▶ Interpretable associations at the population level.
- ▶ No evidence of latent subgroups: population appears continuous.
- ▶ Flexible hierarchical Bayesian model providing an early risk flag and interpretable biomarkers.

# References

- ▶ Park, T., & Casella, G. (2008). *The Bayesian Lasso*. Journal of the American Statistical Association, 103(482), 681–686.  
<https://doi.org/10.1198/016214508000000337>
- ▶ Carvalho, C. M., Polson, N. G., & Scott, J. G. (2010). *The horseshoe estimator for sparse signals*. Biometrika, 97(2), 465–480.  
<https://doi.org/10.1093/biomet/asq017>
- ▶ Piironen, J., & Vehtari, A. (2017). *Sparsity information and regularization in the horseshoe and other shrinkage priors*. Electronic Journal of Statistics, 11(2), 5018–5051. <https://doi.org/10.1214/17-EJS1337SI>
- ▶ Yang, S., Rouder, J. N. Assessing Two Common Priors of Covariance in Hierarchical Designs. *OFS*. [https://doi.org/10.31234/osf.io/jen65\\_v2](https://doi.org/10.31234/osf.io/jen65_v2)
- ▶ Guglielmi, A. (2025). [Bayesian Statistics Slides]. WeBeep, Politecnico di Milano.

**The complete source code and the scripts for this project are available on GitHub at the following repository:**

[https://github.com/elisanordera/  
Bayesian-Statistics-Project.](https://github.com/elisanordera/Bayesian-Statistics-Project)