

Fiscal Policy and Inequality

12. Fixed Effects Regression

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Empirical Application: Estimating Incidence

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- ▶ How do changes in the cigarette tax affect prices?
 - ▶ Does the burden fall on cigarette companies or smokers?
 - ▶ Why?
 - ▶ What are the welfare implications of this policy?

Evans, Ringel & Stech (1999)

- ▶ Cigarettes taxed at both federal and state level in US
 - ▶ Total revenue \simeq \$35 billion per year, similar to estate taxation
- ▶ Variation in excise tax among states within US
 - ▶ from \$0.30 per pack in Virginia to \$4.35 in New York

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- ▶ Variation in excise tax among states within US
 - ▶ from \$0.30 per pack in Virginia to \$4.35 in New York
- ▶ Since 1975, more than 200 changes in state taxes
- ▶ Exploit these state-level changes using simple **diff-in-diff** research design

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- ▶ D = “First difference” estimator
- ▶ Identification assumption (D): absent the tax change, cigarette prices in state A would not have changed between period 0 and 1.

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- ▶ ID assumption is likely violated:
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- ▶ Identification assumption (DD): “**parallel trends**”
 - ▶ Absent tax change, trend in prices would have been the same in states A and B.
 - ▶ Allows for time-invariant differences between the two groups.

Diff-in-Diff Regression

- ▶ Can estimate the diff-in-diff effect using

$$P_{jt} = \alpha + \gamma \text{Treat}_{jt} + \lambda \text{After}_{jt} + \rho \text{Treat} * \text{After}_{jt} + \varepsilon_{jt}$$

where *Treat* is a dummy for being the reform state, and *After* is a dummy for years after the reform.

Diff-in-Diff Regression

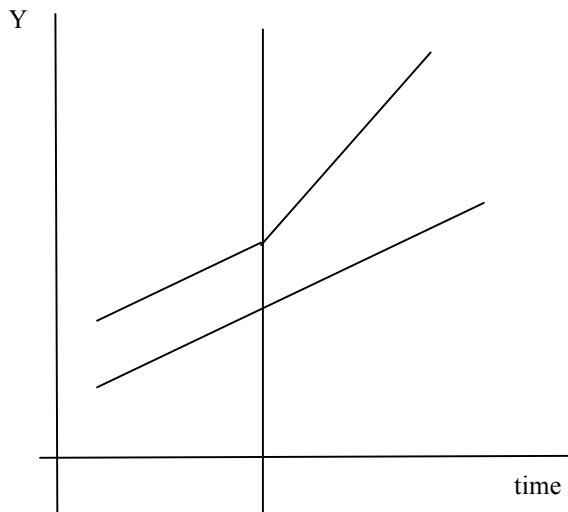
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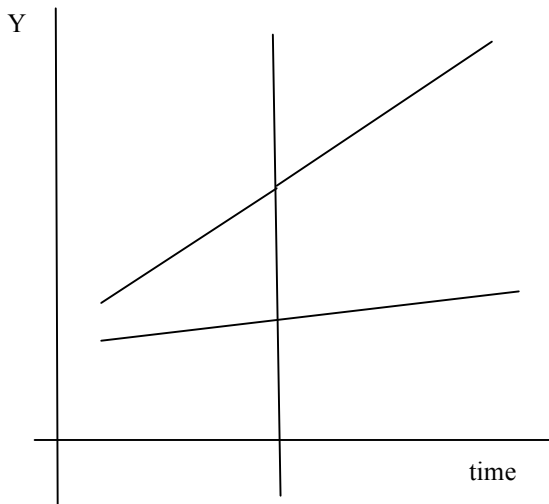
- ▶ Interpreting coefficients:
 - ▶ α , average in non-treated group, pre-treatment
 - ▶ γ , difference between treated and non-treated in pre-treatment period
 - ▶ λ , change in the control group after reform
 - ▶ ρ , the diff-in-diff treatment effect estimate (change in treatment group, relative to change in control group).

Diff-in-diff: Parallel trends assumption

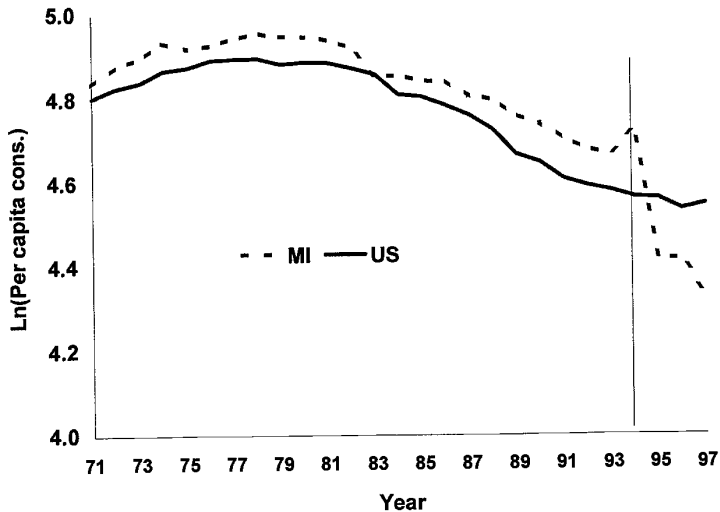


Diff-in-diff: Parallel trends assumption

Things don't always work as we wished they did...

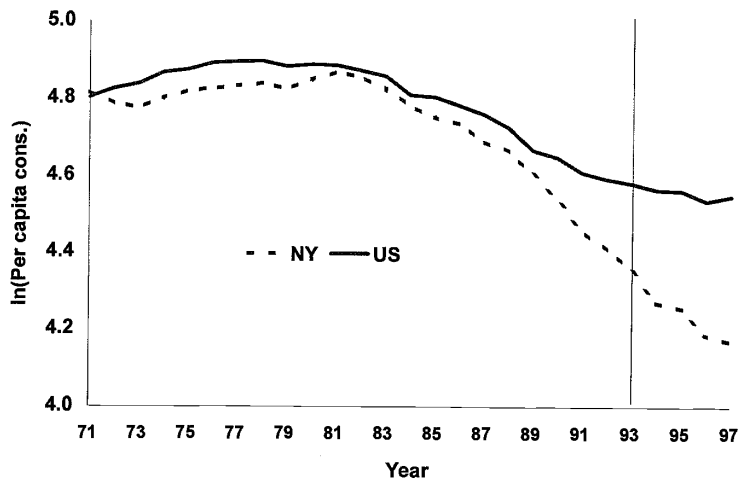


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- ▶ We can use “placebo” DD to test parallel trends assumption
 - ▶ Replicate DD estimate at other points in time when there was *no tax change*
- ▶ If DD in other periods is not zero, then $DD_{t=1}$ likely biased
 - ▶ Useful to plot long time series of outcomes for treatment and control
 - ▶ Pattern should be parallel lines, with sharp change just after reform

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- ▶ Problem: what if states with higher taxes T_{jt} also have more anti-tobacco campaigns?
 - ▶ $\text{Cov}(T_{jt}, \varepsilon_{jt}) \neq 0$ and $\hat{\beta}^{OLS}$ is **biased**.

Fixed-Effects Estimation

- ▶ Fixed-effects estimation:

$$P_{jt} = \alpha + \beta T_{jt} + \delta_j + \gamma_t + \varepsilon_{jt}$$

- ▶ δ_j are state fixed effects
 - ▶ a dummy variable equaling one for state j 's observations, and zero otherwise
- ▶ γ_t are year fixed effects
 - ▶ a dummy variable equaling one for year t 's observations, and zero otherwise
- ▶ Fixed-effects estimation generalises *DD* to $J > 2$ groups and $S > 2$ periods
- ▶ Requires panel (longitudinal) data

Fixed-Effects Estimation (FE)

- ▶ ID assumption (FE): absent the tax change, the trend in the outcome (price) would have been the same in treatment and control groups
 - ▶ In other words, there are no time-varying omitted variables.
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 - ▶ Equivalent to “parallel trends”
- ▶ Implicit assumption in FE: treatment effect is additive and constant (in this case, across states)
- ▶ FE obtains identification from within-state variation over time
 - ▶ nationwide factors, such as federal tax changes, do not affect FE estimates

Threats to validity for FE and DiD

- ▶ If the groups are different in levels, maybe they evolve differently?
- ▶ Why did the treatment group adopt the policy, and not the control group?
- ▶ Policies are usually implemented in bundles (the timing of the treatment may not be by chance) → the outcome variable may be affected by these other policies
- ▶ The treatment should not affect the control group
- ▶ The composition of the treatment and control groups should not change as a result of treatment

Usual checks

- ▶ The two groups evolved similarly in the past (although this is not a guarantee for validity)
- ▶ The timing of the adoption of the policy was as good as random
- ▶ No other policies were adopted at the same time
- ▶ Verify that there is no reason to believe that the control group might be affected
- ▶ Add group-specific trends in outcome variable as additional regressor.

Let's Try It

```
# set index to group and time categories.  
df.set_index(['state', 'year'], inplace=True)  
  
# fit fixed effects model  
from linearmodels.panel.PanelOLS import from_formula  
model = from_formula('log_gdp ~ log_deathtax + EntityEffects + TimeEffects',  
                    data=df)  
results = model.fit()  
results
```

What about all the other regression outputs?

- ▶ Most are not that useful.
- ▶ R^2 tells you how much of the variance in the outcome is explained by the right-hand-side variables.
 - ▶ can be used to decide between two models, e.g., whether to take the log of your outcome variable.
 - ▶ can be arbitrarily increased by adding more variables; this is not a reason to add more variables.

Effect of cig taxes on cig prices (Evans et al 1999)

- ▶ Main regression model:

$$P_{jt} = \alpha + \beta T_{jt} + \delta_j + \gamma_t + \varepsilon_{jt}$$

- ▶ P_{jt} = average retail price per pack in state j , year t (in \$ cents)
- ▶ T_{jt} = total per pack tax (state+federal) in state j , year t (in \$ cents)
- ▶ δ_j = state fixed effects. Control for any time-invariant differences in prices across states (eg, age distribution).
- ▶ γ_t = year fixed effects. Control for shocks to prices that are common to all states but may vary across years (eg, federal tax change).

Evans, Ringel & Stech (1999): Results

TABLE 2
OLS Estimates, Retail Price Model: Tobacco Institute Data

| Independent variable | Average state retail price, 1985–1996 | | Net retail price in Tennessee, 1970–1994 | |
|------------------------------|---------------------------------------|----------------|--|----------------|
| | Nominal (1) | Real (2) | Nominal (3) | Real (4) |
| Nominal/real tax | 1.01 (0.04) | 0.92 (0.04) | | |
| Nominal/real wholesale price | | | 1.07 (0.02) | 0.86 (0.04) |
| R^2 | 0.972 | 0.933 | 0.989 | 0.963 |
| Observations | 612 | 612 | 25 | 25 |

Standard errors in parentheses. Real prices in 1997 cents/pack. Models in columns (1) and (2) control for state effects.

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- ▶ 100% pass-through implies $\varepsilon_D = 0$ or $\varepsilon_S = \infty$ at state level
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 - ▶ Why? Try using a graph to understand this
- ▶ Why not $\varepsilon_D = 0$?
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- ▶ Consider $\varepsilon_S = \infty$:
 - ▶ Tobacco companies can easily send inventories from high-tax to low-tax states, affecting prices by reducing supply
 - ▶ Pass-through would be lower at national level

Evans, Ringel & Stech (1999): Demand Elasticity

- ▶ To estimate ε_D , they use this model:

$$\ln(Q_{jt}) = \beta T_{jt} + X_{jt}\alpha + \mu_{1j} + \mu_{2j} \cdot \text{Time}_t + \nu_t + \varepsilon_{jt}$$

- ▶ $\ln(Q_{jt})$ = log per capita consumption, state j , year t
- ▶ T_{jt} = state+federal tax per pack, in \$ cents
- ▶ μ_{1j} = state fixed effects
- ▶ $\mu_{2j} \cdot \text{Time}_t$ = state-specific time trends
- ▶ ν_t = year fixed effects

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- Since $\varepsilon_S \approx \infty$, then $\Delta P \approx \Delta T$
- Hence:

$$\begin{aligned}\hat{\varepsilon}^D &= \frac{\Delta Q}{\Delta T} \frac{\bar{P}}{\bar{Q}} \\ \Rightarrow \hat{\varepsilon}^D &= \hat{\beta} \cdot \bar{P}\end{aligned}$$

Evans, Ringel & Stech (1999): Demand Elasticity

TABLE 3
*OLS Estimates, Log Per Capita Consumption Model,
Tobacco Institute Data, 1985–1996*

| Independent variable | Coefficients (standard errors) on | | | | | |
|----------------------|-----------------------------------|-------------------|-------------------|-------------------|-------------------|-------------------|
| | Real tax | | | Real price | | |
| | (1) | (2) | (3) | (4) | (5) | (6) |
| Current value | −0.254 (0.037) | −0.165 (0.040) | −0.173 (0.041) | −0.176 (0.027) | −0.176 (0.027) | −0.167 (0.029) |
| 1-year lag | | −0.215 (0.413) | −0.188 (0.047) | | −0.027 (0.032) | −0.031 (0.032) |
| 2-year lag | | | −0.061 (0.045) | | | −0.017 (0.033) |
| Price elasticity | −0.424 (0.062) | −0.635 (0.074) | −0.705 (0.090) | −0.294 (0.045) | −0.337 (0.058) | −0.359 (0.072) |
| R ² | 0.975 | 0.977 | 0.977 | 0.975 | 0.975 | 0.976 |

Standard errors in parentheses. The dependent variable is the log per capita consumption. There are 512 observations in each model. The mean price is \$1.75/pack, and the mean per capita consumption is 105. All models include year effects, state effects, state-specific time trends, and log per capita consumption, plus measures of the fraction of adults in three age, three education, and two race groups.

Evans, Ringel & Stech (1999): Demand Elasticity

- ▶ We have:
 - ▶ $\hat{\beta} = -0.254$
 - ▶ $\bar{P} = \$1.75$
- ▶ Demand model estimate implies: $\varepsilon_D = -0.42$
 - ▶ 10% increase in price induces a 4.2% reduction in consumption
- ▶ This is the **short-run** elasticity. In the paper, they also estimate long-run elasticities.

Practice Exam Question

- Consider the market for diamonds in a closed economy. You estimate the following regression:

$$P_{jt} = \alpha + \beta T_{jt} + \delta_j + \gamma_t + \varepsilon_{jt}$$

where P_{jt} is the consumer price of diamonds in province j and year t , and T_{jt} is the excise tax on diamonds.

1. You obtain a point estimate $\hat{\beta} = 0$. Interpret what this coefficient estimate means. What do you learn about the elasticity of supply and/or demand for diamonds?
2. Given your answer to (a) and economic intuition, who bears the incidence of this tax? Use the tax incidence formula.

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- ▶ Two proposals for assigning to treatment and control:
 1. Flip a coin once: tail, individuals from Zurich are treated, heads, individuals from Zug are treated
 2. Flip coin 2000 times, once for each individual: tails, assigned to treatment; heads, assigned to control

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- ▶ What's wrong? Presence of a common random effect:
 - ▶ In proposal 1, there might be some common shock affecting all individuals in the treatment group or in the control group.
 - ▶ OLS standard errors assume that all observations are independent realizations. Standard errors have to be corrected to account for the presence of a common random effect.

What about fixed effects regressions

- ▶ Consider the case of cigarette taxes. We have 50 states, times 50 years, equals 2500 observations.
 - ▶ if i only included the 10 years before and after the reform, i would have 20 years, or 1000 observations, but its essentially the same information, although standard errors would be bigger.
 - ▶ We need to account for *serial correlation* within state.

Solution: Clustering Standard Errors

- ▶ Cluster standard errors:
 - ▶ statistically acknowledges how many independent sources of information there are in the data.
 - ▶ can cluster by state and by time, to allow for correlation in errors along both dimensions.

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 - ▶ statistically acknowledges how many independent sources of information there are in the data.
 - ▶ can cluster by state and by time, to allow for correlation in errors along both dimensions.
- ▶ In python linearmodels:

```
# fit fixed effects model with clustered errors
results = model.fit(cov_type='clustered',
                    cluster_entity=True,
                    cluster_time=True)
```

How to cluster?

- ▶ In general, cluster at the level of your treatment variation:
 - ▶ state-level reforms → cluster by state
 - ▶ city-level reforms → cluster by city
 - ▶ etc.

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- ▶ In general, cluster at the level of your treatment variation:
 - ▶ state-level reforms → cluster by state
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 - ▶ etc.
- ▶ Clustering is important, and in general will dramatically affect statistical significance of results (Bertrand, Duflo, and Mullainathan QJE 2004).