

TAMS65 - Lecture 9

Random vector - Normal random vector

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Syllabus

- ▶ Examiner, Teaching assistants, Textbook, Course Literature and Course plan are the same as VT1.
- ▶ 4 Distance lectures (on Lisam) - Two new topics.
- ▶ 2 Distance lessons - Microsoft Teams.
- ▶ 1 Distance seminar (on Lisam) - Same way as lectures.
- ▶ Distance Project (on Lisam : Lisam - Course documents - 7 Project) - **Instructions**.

Content

- ▶ Syllabus
- ▶ Repetitions
- ▶ Random vector
- ▶ Standard normal random vector
- ▶ General normal random vector
- ▶ Example 3
- ▶ Example 4



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Syllabus

- ▶ Distance quiz will be given on May 6, 2020.
 - ▶ which is based on the project and 4 lectures.
 - ▶ which contains five questions with multiple choice options.
 - ▶ Details will be given in the last Lecture.

Pass the course =
=Pass the written Exam (VT1) + Pass the project(VT2) + Pass the quiz(VT2)

Repetitions

If X, Y are independent (oberoende), then $E(XY) = E(X)E(Y)$.

Covariance (kovarians) of X and Y is defined as

$$\begin{aligned}\text{cov}(X, Y) &= c(X, Y) = \sigma_{X,Y} = E[(X - E(X))(Y - E(Y))] \\ &= E(XY) - E(X)E(Y)\end{aligned}$$

Remark 1:

- 1 Covariance measures how much two random variables change together.
- 2 $\text{cov}(X, Y) = \text{cov}(Y, X)$
- 3 $\text{cov}(X, X) = V(X)$
- 4 $\text{cov}(a, X) = 0$ for any constant a .

Example 1

Example 1, If $X_1 \sim N(0, 1)$, $X_2 \sim N(1, 2)$ and $X_3 \sim N(1, 3)$ are independent, calculate

$$\text{cov}(2X_1 - X_2 + 4X_3, 3X_1 + X_2).$$

$$\begin{aligned}\text{cov}(2X_1 - X_2 + 4X_3, 3X_1 + X_2) \\ &= 6\text{cov}(X_1, X_1) + 2\text{cov}(X_1, X_2) - 3\text{cov}(X_2, X_1) - \text{cov}(X_2, X_2) \\ &\quad + 12\text{cov}(X_3, X_1) + 4\text{cov}(X_3, X_2) \\ &= 6V(X_1) - V(X_2) = 2.\end{aligned}$$

Repetitions

5 If X, Y are independent, then $\text{cov}(X, Y) = 0$

- ▶ Independent \Rightarrow covariance is 0, but covariance is 0 \nRightarrow independent.
- ▶ Covariance is NOT 0 \Rightarrow NOT independent.

6 $V(aX + bY) = a^2V(X) + 2ab\text{cov}(X, Y) + b^2V(Y)$.

- ▶ If X, Y are independent, for any constants a, b , then $V(aX + bY) = a^2V(X) + b^2V(Y)$.

7 If X, Y, W, V are random variables, a, b, c, d, e are constants, then

$$\begin{aligned}\text{cov}(aX + bY + c, dW + eV) \\ = ad\text{cov}(X, W) + ae\text{cov}(X, V) + bd\text{cov}(Y, W) + be\text{cov}(Y, V)\end{aligned}$$

Repetitions

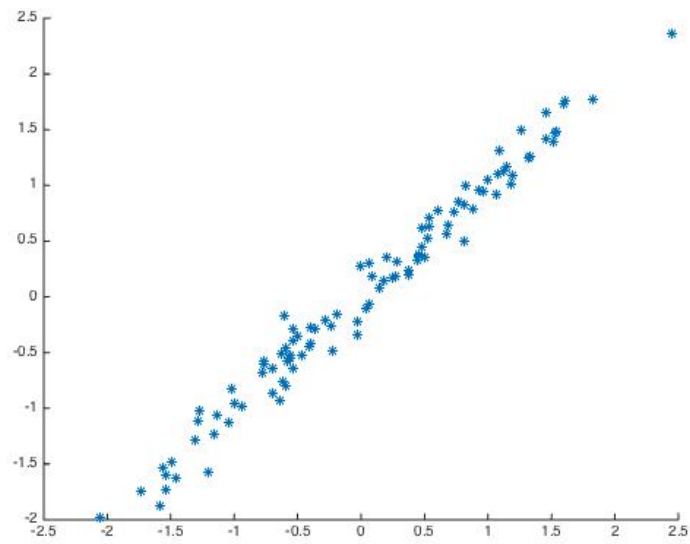
Correlation/Correlation coefficient (korrelations koefficient) of X and Y is defined as

$$\rho_{X,Y} = \frac{\text{cov}(X, Y)}{\sqrt{V(X)V(Y)}}$$

Remark 2:

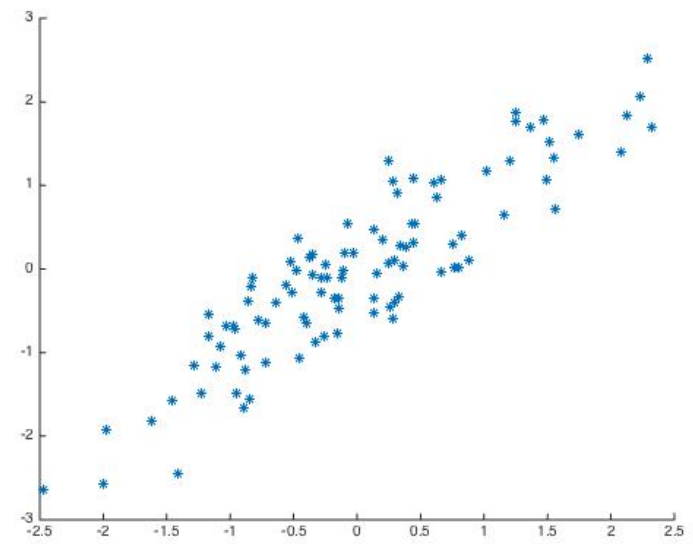
- 1 Correlation measures how much linear relation between X and Y .
 - ▶ If $|\rho_{X,Y}| \approx 1$, then X and Y are full linear.
 - ▶ If $|\rho_{X,Y}| \approx 0$, then X and Y are non-linear.

correlation $\rho = 0.99$



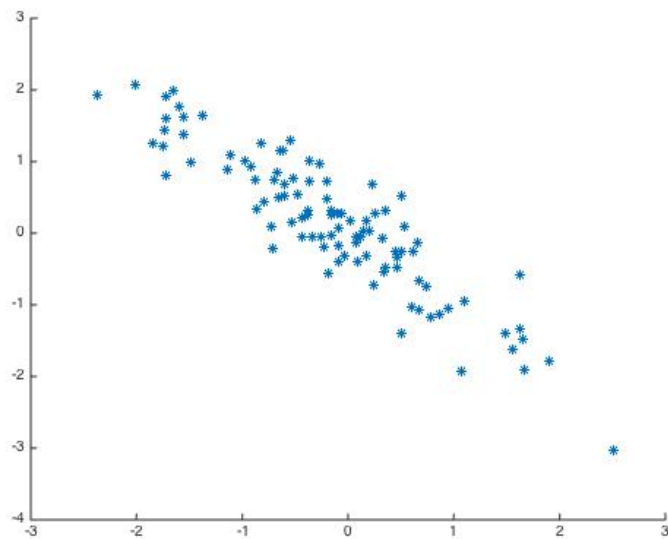
$\hat{\rho} = 0.9886$

correlation $\rho = 0.9$



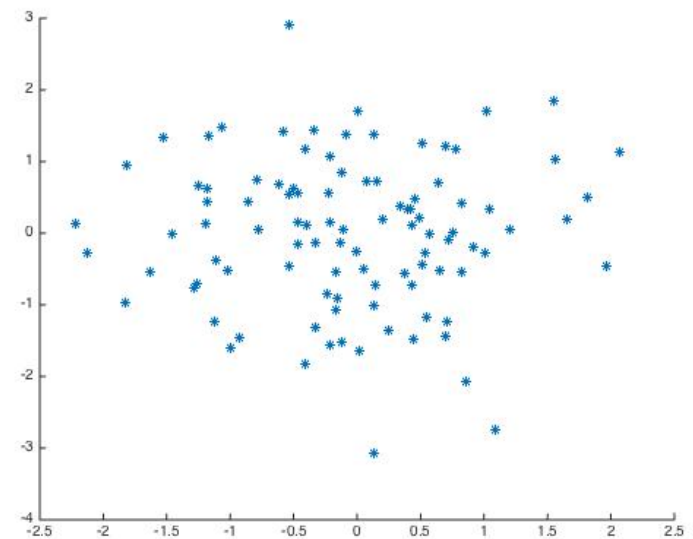
$\hat{\rho} = 0.8983$

correlation $\rho = -0.9$



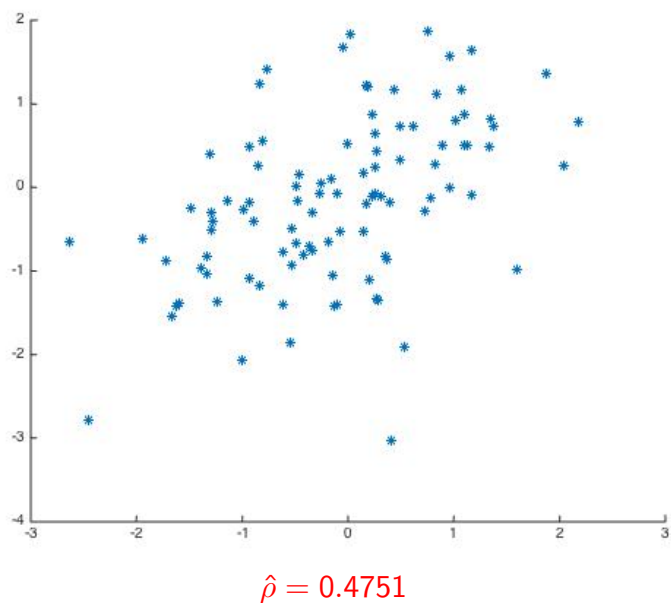
$\hat{\rho} = -0.9075$

correlation $\rho = 0$

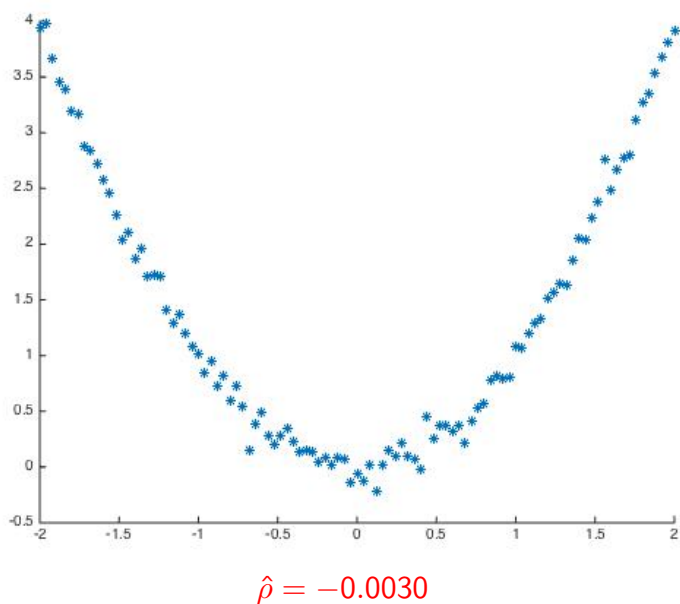


$\hat{\rho} = -0.0041$

correlation $\rho = 0.5$



correlation $\rho = 0$, non linear but quadratic relation.



Repetitions

Remark 2:

$$-1 \leq \rho_{X,Y} \leq 1$$

- ▶ Proof: Cauchy-Schwarz inequality.
- ▶ $|\rho_{X,Y}| = 1$ if and only if X and Y are full linear.

- ▶ If X and Y are independent, then $\rho_{X,Y} = 0$.

Definition X and Y is called **uncorrelated**(okorrelerade) if $\rho_{X,Y} = 0$.

Note:

- ▶ Independent \Rightarrow uncorrelated, but uncorrelated \nRightarrow independent.
- ▶ $|\rho_{X,Y}| \approx 0 \Rightarrow X$ and Y are non-linear, but it does NOT imply that they don't have other relation.

Random vector

Definition: A **random vector**(stokastisk vektor) is defined as

$$\mathbf{X} = \begin{pmatrix} X_1 \\ X_2 \\ \vdots \\ X_n \end{pmatrix}_{n \times 1},$$

where the components(komponenterna) $X_i, i = 1, \dots, n$ are one dimensional random variables.

A random vector \mathbf{X} has a **mean vector (väntevärdesvektor)** which is defined as

$$\mu_{\mathbf{X}} = \begin{pmatrix} E(X_1) \\ E(X_2) \\ \vdots \\ E(X_n) \end{pmatrix}_{n \times 1}.$$

A random vector \mathbf{X} has also a **covariance matrix (kovarians-matris)** which is defined as

$$\mathbf{C}_{\mathbf{X}} = \begin{pmatrix} \text{cov}(X_1, X_1) & \text{cov}(X_1, X_2) & \dots & \text{cov}(X_1, X_n) \\ \text{cov}(X_2, X_1) & \text{cov}(X_2, X_2) & \dots & \text{cov}(X_2, X_n) \\ \vdots & \vdots & \ddots & \vdots \\ \text{cov}(X_n, X_1) & \text{cov}(X_n, X_2) & \dots & \text{cov}(X_n, X_n) \end{pmatrix}_{n \times n},$$

where the components $\text{cov}(X_i, X_j) = E[(X_i - E(X_i))(X_j - E(X_j))]$ for $i, j = 1, 2, \dots, n$.

Remark 3:

- 1 $\text{cov}(X_i, X_i) = V(X_i)$, $i = 1, \dots, n$. Note: On main diagonal, there are $V(X_i)$.
- 2 If $n = 1$, then $\mathbf{C}_{\mathbf{X}} = V(X_1)$.
- 3 $\mathbf{C}_{\mathbf{X}}$ is symmetric since $\text{cov}(X_i, X_j) = \text{cov}(X_j, X_i)$. Thus, we can see

$$\mathbf{C}_{\mathbf{X}} = \mathbf{C}'_{\mathbf{X}} = \mathbf{C}_{\mathbf{X}}^T \quad \text{transpose(transponat) of } \mathbf{C}_{\mathbf{X}}$$

Question: Can a random vector \mathbf{X} have a variance, if $n > 1$? NO!

Theorem 1: Let \mathbf{X} be a n dimensional random vector with mean vector $\mu_{\mathbf{X}}$ and covariance matrix $\mathbf{C}_{\mathbf{X}}$. Define a new random vector

$$\mathbf{Y} = \mathbf{A}\mathbf{X} + \mathbf{b},$$

where \mathbf{A} is a constant matrix and \mathbf{b} is a constant vector.

Then \mathbf{Y} has the following mean vector and covariance matrix

$$\mu_{\mathbf{Y}} = \mathbf{A}\mu_{\mathbf{X}} + \mathbf{b},$$

$$\mathbf{C}_{\mathbf{Y}} = \mathbf{A}\mathbf{C}_{\mathbf{X}}\mathbf{A}'.$$

Recall: $E(aX + b) = aE(X) + b$, $V(aX + b) = a^2V(X)$.

Example 2

Example 2, A random vector $\begin{pmatrix} X \\ Y \end{pmatrix}$ has covariance matrix $\begin{pmatrix} 2 & 3 \\ 3 & 6 \end{pmatrix}$.

- (1) Find $V(X)$ and $V(Y)$.
- (2) Are X and Y independent?

- (1) $V(X) = 2$, $V(Y) = 6$.
- (2) X and Y are not independent since $\text{cov}(X_1, X_2) = 3 \neq 0$.

Standard normal random vector

- 2 The standard normal random vector is denoted by

$$\mathbf{X} \sim N(\boldsymbol{\mu}_X, \mathbf{C}_X) \text{ or } \mathbf{X} \sim N(\vec{0}, \mathbf{I}_{n \times n}).$$

- 3 The density function(täthesfunktion) is

$$f_{\mathbf{X}}(\mathbf{x}) = f_{\mathbf{X}}(x_1, \dots, x_n) = f_{X_1}(x_1) \dots f_{X_n}(x_n) = \frac{1}{(\sqrt{2\pi})^n} e^{-\frac{1}{2}\mathbf{x}'\mathbf{x}},$$

$$\text{where } \mathbf{x} = \begin{pmatrix} x_1 \\ x_2 \\ \vdots \\ x_n \end{pmatrix}.$$

Standard normal random vector

Definition: A random vector $\mathbf{X} = \begin{pmatrix} X_1 \\ X_2 \\ \vdots \\ X_n \end{pmatrix}_{n \times 1}$ is called **standard normal random vector** if its components satisfy

$$X_i \sim N(0, 1), i = 1, \dots, n \text{ and } X_1, \dots, X_n \text{ are independent.}$$

Remark 4:

$$\mu_{\mathbf{X}} = \begin{pmatrix} 0 \\ 0 \\ \vdots \\ 0 \end{pmatrix}_{n \times 1} = \vec{0}, \mathbf{C}_{\mathbf{X}} = \begin{pmatrix} 1 & 0 & \dots & 0 \\ 0 & 1 & \dots & 0 \\ \vdots & \vdots & \ddots & \vdots \\ 0 & 0 & \dots & 1 \end{pmatrix}_{n \times n} = \mathbf{I}_{n \times n}.$$

General Normal random vector

Recall: If $X \sim N(0, 1)$, then $Y = \sigma X + \mu \sim N(\mu, \sigma)$.

Definition:

If there exist constant matrix \mathbf{A} and constant vector \mathbf{b} , such that $\mathbf{Y} = \mathbf{A}\mathbf{X} + \mathbf{b}$ where \mathbf{X} is a standard normal random vector, then \mathbf{Y} is called (general) **normal random vector**.

Note: Normal random vector is also called **Multivariate random vector**.

Remark 5:

- 1 $\mu_{\mathbf{Y}} = \mathbf{A}\mu_{\mathbf{X}} + \mathbf{b} = \mathbf{b}$, and $\mathbf{C}_{\mathbf{Y}} = \mathbf{A}\mathbf{C}_{\mathbf{X}}\mathbf{A}' = \mathbf{A}\mathbf{A}'$ (by **Theorem 1.**)
- 2 The normal random vector is denoted by

$$\mathbf{Y} \sim N(\mu_{\mathbf{Y}}, \mathbf{C}_{\mathbf{Y}}) \text{ or } \mathbf{Y} \sim N(\mathbf{b}, \mathbf{A}\mathbf{A}').$$

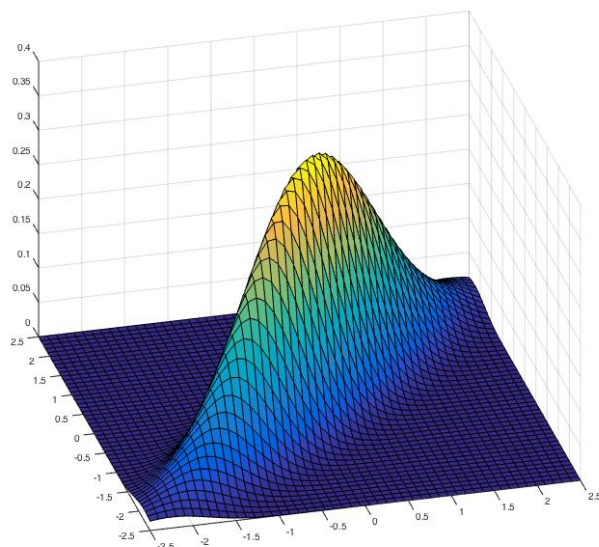
3 The density function(täthesfunktion) is

$$f_{\mathbf{Y}}(\mathbf{y}) = f_{\mathbf{Y}}(y_1, \dots, y_n) = \frac{1}{(\sqrt{2\pi})^n \sqrt{\det \mathbf{C}_{\mathbf{Y}}}} e^{-\frac{1}{2}[(\mathbf{y} - \boldsymbol{\mu}_{\mathbf{Y}})' \mathbf{C}_{\mathbf{Y}}^{-1} (\mathbf{y} - \boldsymbol{\mu}_{\mathbf{Y}})]},$$

where $\mathbf{y} = \begin{pmatrix} y_1 \\ y_2 \\ \vdots \\ y_n \end{pmatrix},$

$\det \mathbf{C}_{\mathbf{Y}}$ is determinant of covariance matrix $\mathbf{C}_{\mathbf{Y}}$
and $\mathbf{C}_{\mathbf{Y}}^{-1}$ is the inverse matrix of $\mathbf{C}_{\mathbf{Y}}$.

For example, $\begin{pmatrix} X \\ Y \end{pmatrix} \sim N(\vec{0}, \begin{pmatrix} 1 & 0.9 \\ 0.9 & 1 \end{pmatrix})$



Theorem 2 $\mathbf{X} = \begin{pmatrix} X_1 \\ \vdots \\ X_n \end{pmatrix}_{n \times 1}$ is a normal random vector if
and only if $c_1 X_1 + \dots + c_n X_n$ is normally distributed for any **not all zero** constants c_1, \dots, c_n .

Theorem 3 If $\begin{pmatrix} X \\ Y \end{pmatrix}$ is a normal random vector, then we get

- X, Y are independent $\Leftrightarrow \text{cov}(X, Y) = 0$.
- X, Y are independent $\Leftrightarrow X, Y$ are uncorrelated ($\rho_{X,Y} = 0$).

Example 3

Example 3, Suppose that \mathbf{X} is a normal random vector with

$$\mathbf{X} = \begin{pmatrix} X_1 \\ X_2 \\ X_3 \end{pmatrix} \sim N \left(\begin{pmatrix} 1 \\ 1 \\ 2 \end{pmatrix}, \begin{pmatrix} 1 & 0 & 2 \\ 0 & 1 & 3 \\ 2 & 3 & 4 \end{pmatrix} \right).$$

Define $Y = (X_1 + X_2 + 2X_3)/2$.

- (1) Find the distributions of X_1, X_2 . Are they independent?
- (2) Find $P(Y \leq 6)$.

(1), By Theorem 1, $X_1 \sim N(1, 1)$ and $X_2 \sim N(1, 1)$.
 $\text{cov}(X_1, X_2) = 0$ and X_1, X_2 are components of a normal random vector which imply that they are independent. (By Theorem 3)
(2) By Theorem 2, Y is normally distributed since it is a linear combination of components of a normal random vector.

Example 3

$$Y = \begin{pmatrix} \frac{1}{2} & \frac{1}{2} & 1 \end{pmatrix} \begin{pmatrix} X_1 \\ X_2 \\ X_3 \end{pmatrix} = \mathbf{A}\mathbf{X}$$

By Theorem 1,

$$\mu_Y = \mathbf{A}\mu_X = \begin{pmatrix} \frac{1}{2} & \frac{1}{2} & 1 \end{pmatrix} \begin{pmatrix} 1 \\ 1 \\ 2 \end{pmatrix} = 3,$$

$$\mathbf{C}_Y = \mathbf{A}\mathbf{C}_X\mathbf{A}' = \begin{pmatrix} \frac{1}{2} & \frac{1}{2} & 1 \end{pmatrix} \begin{pmatrix} 1 & 0 & 2 \\ 0 & 1 & 3 \\ 2 & 3 & 4 \end{pmatrix} \begin{pmatrix} \frac{1}{2} & \frac{1}{2} & 1 \end{pmatrix}' = \frac{19}{2}.$$

Thus $Y \sim N(3, \sqrt{\frac{19}{2}})$.

$$P(Y \leq 6) = \Phi(0.97) = 0.834.$$

Example 4

By Theorem 1,

$$\mu_Y = \mathbf{A}\mu_X = \vec{0},$$

$$\mathbf{C}_Y = \mathbf{A}\mathbf{C}_X\mathbf{A}' = \begin{pmatrix} 2 & 0 \\ 0 & 2 \end{pmatrix} = 2\mathbf{I}_{2 \times 2}.$$

$$\det \mathbf{C}_Y = 4, \mathbf{C}_Y^{-1} = \frac{1}{2}\mathbf{I}_{2 \times 2} \text{ and } \mathbf{y} - \mu_Y = \begin{pmatrix} y_1 \\ y_2 \end{pmatrix}.$$

Thus

$$\begin{aligned} f_Y(y_1, y_2) &= \frac{1}{(\sqrt{2\pi})^2 \sqrt{\det \mathbf{C}_Y}} e^{-\frac{1}{2}[(\mathbf{y} - \mu_Y)' \mathbf{C}_Y^{-1}(\mathbf{y} - \mu_Y)]} \\ &= \frac{1}{4\pi} e^{-\frac{y_1^2 + y_2^2}{4}}. \end{aligned}$$

Example 4

Random variables X_1, X_2 are independent and $X_i \sim N(0, 1), i=1, 2$.

$$Y_1 = X_1 - X_2,$$

$$Y_2 = X_1 + X_2.$$

Find the density function for $\mathbf{Y} = \begin{pmatrix} Y_1 \\ Y_2 \end{pmatrix}$.

$$\mathbf{Y} = \begin{pmatrix} 1 & -1 \\ 1 & 1 \end{pmatrix} \begin{pmatrix} X_1 \\ X_2 \end{pmatrix} = \mathbf{A}\mathbf{X}, \quad \mathbf{X} \sim N(\vec{0}, \mathbf{I}_{2 \times 2})$$

Practice after the lecture:

Exercises:

(I) PS-15, PS-16, PS-20, PS-21.

(II) PS-23, PS-17, PS-18, PS-22.

Thank you!

<http://courses.mai.liu.se/GU/TAMS65/>