Hashing and Sketching

Part One

Outline for Today

Hash Functions

Understanding our basic building blocks.

• Frequency Estimation

• Estimating how many times we've seen something.

• Concentration Inequalities

 "Correct on expectation" versus "correct with high probability."

• Probability Amplification

Increasing our confidence in our answers.

Preliminaries: *Hash Functions*

Hashing in Practice

- Hash functions are used extensively in programming and software engineering:
 - They make hash tables possible: think C++ std::hash, Python's __hash__, or Java's Object.hashCode().
 - They're used in cryptography: SHA-256, HMAC, etc.
- Question: When we're in Theoryland, what do we mean when we say "hash function?"

Hashing in Theoryland

- In Theoryland, a hash function is a function from some domain called the *universe* (typically denoted **/) to some codomain.
- The codomain is usually a set of the form

$$[m] = \{0, 1, 2, 3, ..., m - 1\}$$

$$h: \mathcal{U} \to [m]$$

Hashing in Theoryland

- *Intuition*: No matter how clever you are with designing a specific hash function, that hash function isn't random, and so there will be pathological inputs.
 - You can formalize this with the pigeonhole principle.
- *Idea*: Rather than finding the One True Hash Function, we'll assume we have a collection of hash functions to pick from, and we'll choose which one to use randomly.

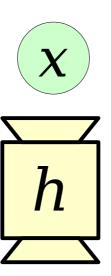
Families of Hash Functions

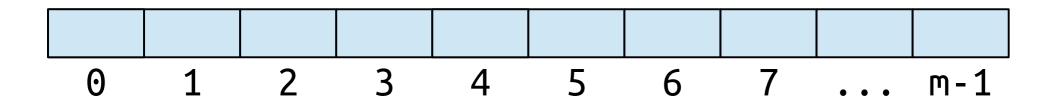
- A *family* of hash functions is a set \mathscr{H} of hash functions with the same domain and codomain.
- We can then introduce randomness into our data structures by sampling a random hash function from \mathcal{H} .
- **Key Point:** The randomness in our data structures almost always derives from the random choice of hash functions, not from the data.

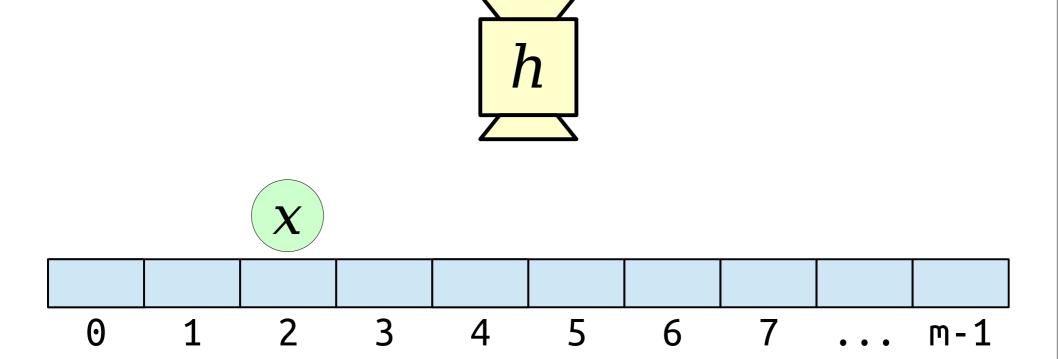
Data is adversarial. Hash function selection is random.

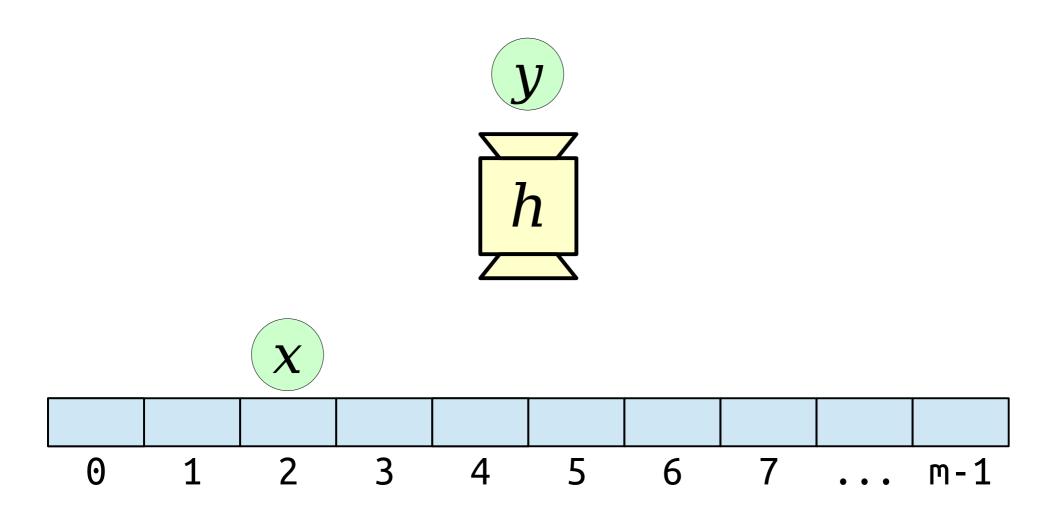
 Question: What makes a family of hash functions \(\mathscr{H}\) a "good" family of hash functions?

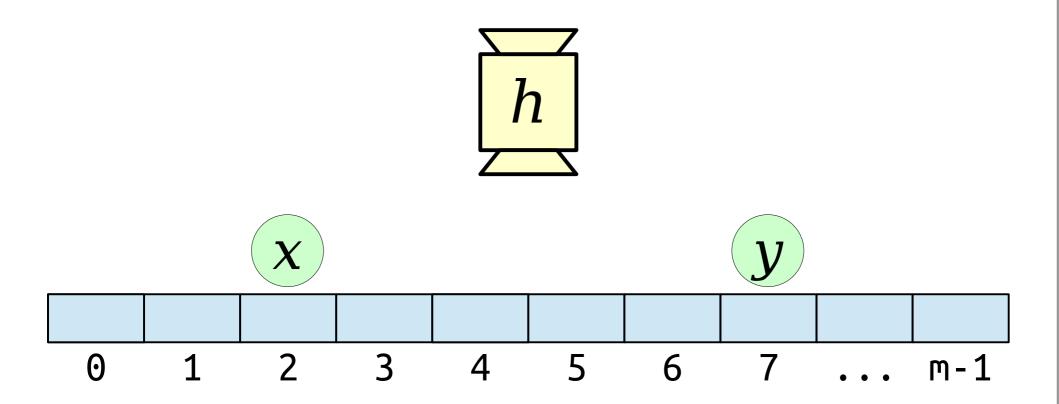


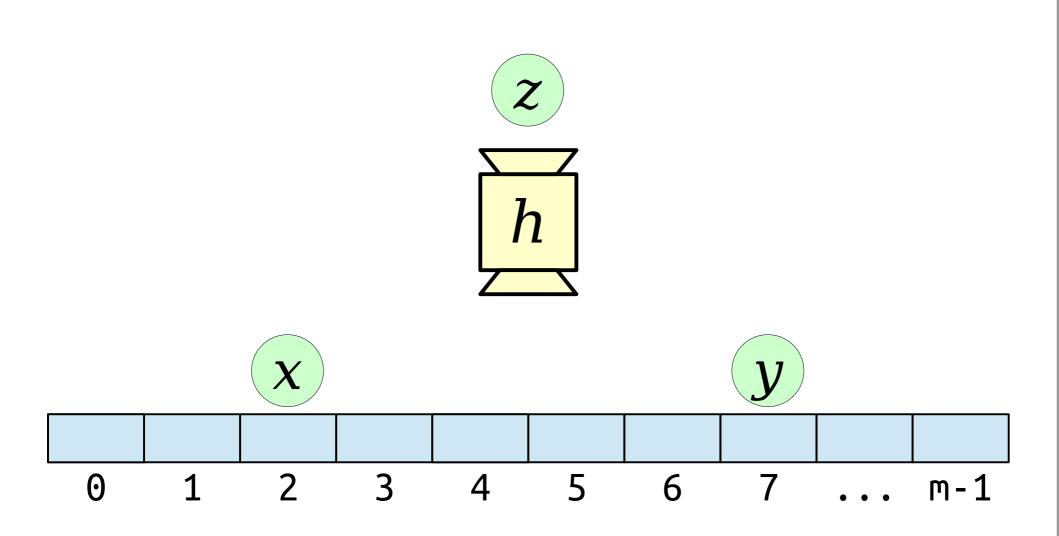






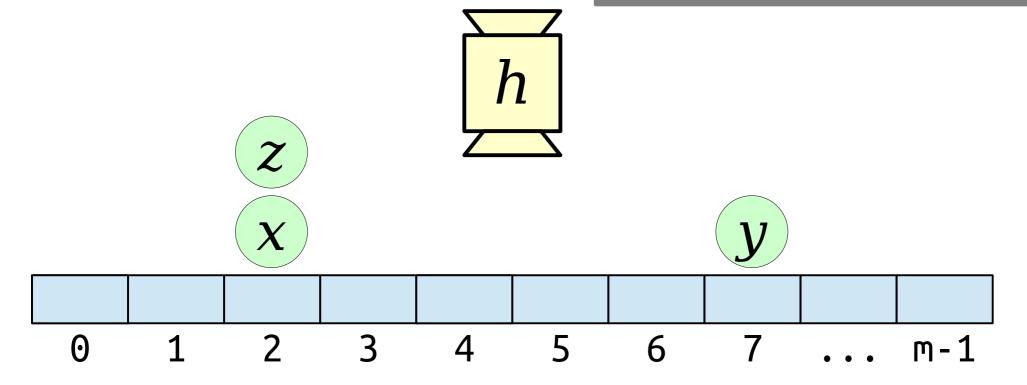




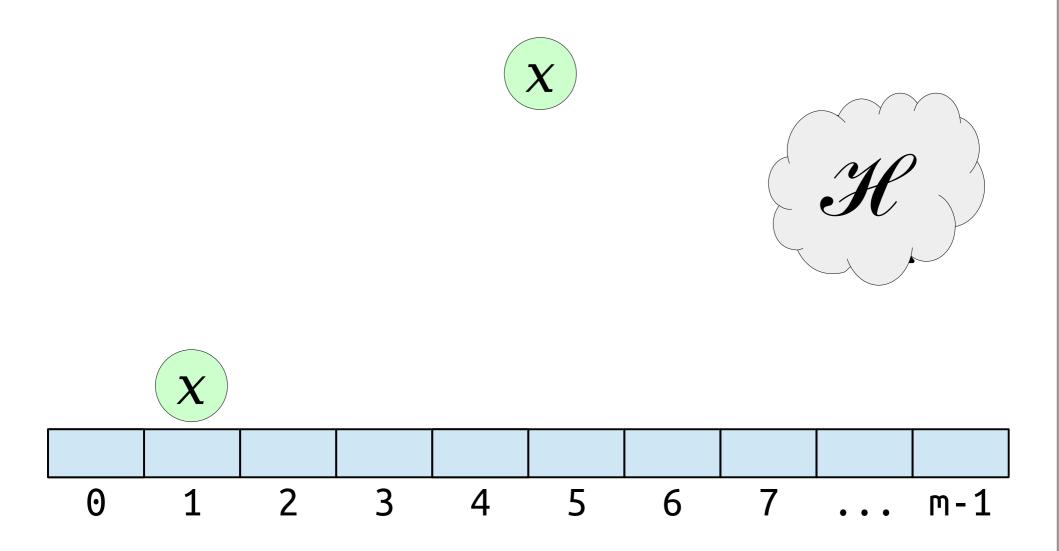


Problem: A hash function that distributes n elements uniformly at random over [m] requires $\Omega(n \log m)$ space in the worst case.

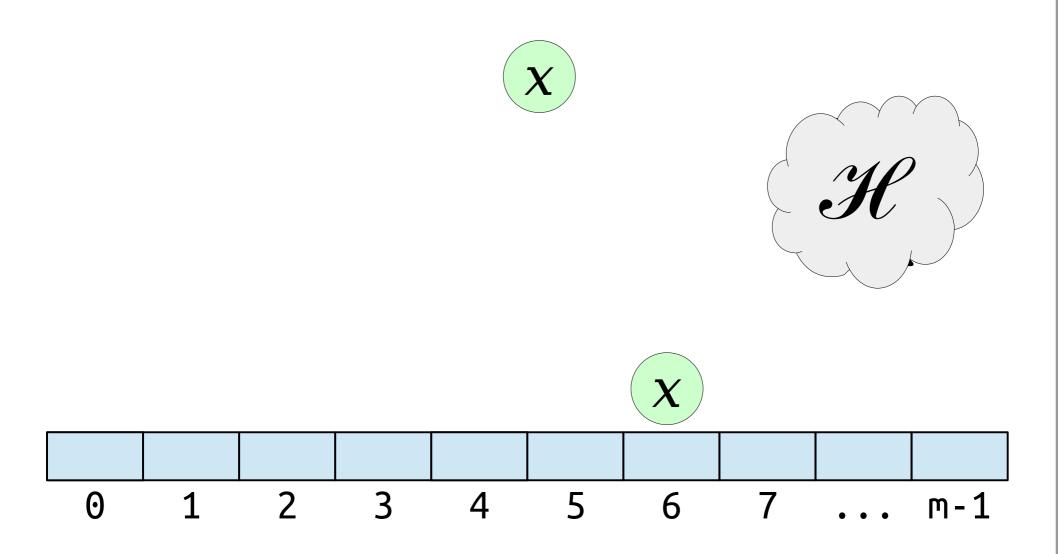
Question: Do we actually need true randomness? Or can we get away with something weaker?



Each element should have an equal probability of being placed in each slot. For any $x \in \mathcal{U}$ and random $h \in \mathcal{H}$, the value of h(x) is uniform over [m].



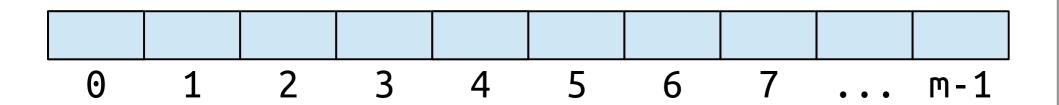
Each element should have an equal probability of being placed in each slot. For any $x \in \mathcal{U}$ and random $h \in \mathcal{H}$, the value of h(x) is uniform over [m].



Each element should have an equal probability of being placed in each slot. For any $x \in \mathcal{U}$ and random $h \in \mathcal{H}$, the value of h(x) is uniform over [m].

Find an "obviously bad" family of hash functions that satisfies the distribution property.

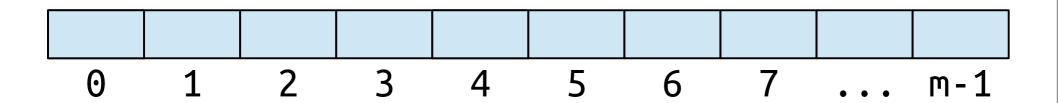
Formulate a hypothesis, but *don't post anything* in chat just yet.



Each element should have an equal probability of being placed in each slot. For any $x \in \mathcal{U}$ and random $h \in \mathcal{H}$, the value of h(x) is uniform over [m].

Find an "obviously bad" family of hash functions that satisfies the distribution property.

Now, private chat me your best guess. Not sure? Just answer "??".



Each element should have an equal probability of being placed in each slot. For any $x \in \mathcal{U}$ and random $h \in \mathcal{H}$, the value of h(x) is uniform over [m].

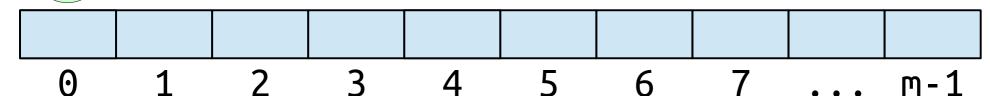
Problem: This rule doesn't guarantee that elements are spread out.



2

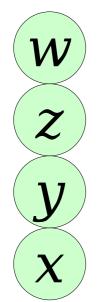
y

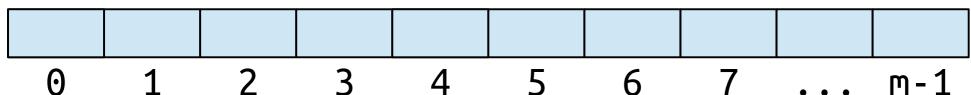
X



Each element should have an equal probability of being placed in each slot. For any $x \in \mathcal{U}$ and random $h \in \mathcal{H}$, the value of h(x) is uniform over [m].

Problem: This rule doesn't guarantee that elements are spread out.

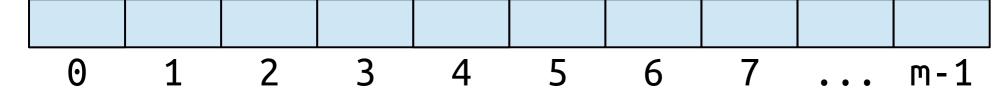




Each element should have an equal probability of being placed in each slot. For any $x \in \mathcal{U}$ and random $h \in \mathcal{H}$, the value of h(x) is uniform over [m].

Problem: This rule doesn't guarantee that elements are spread out.

w z y x

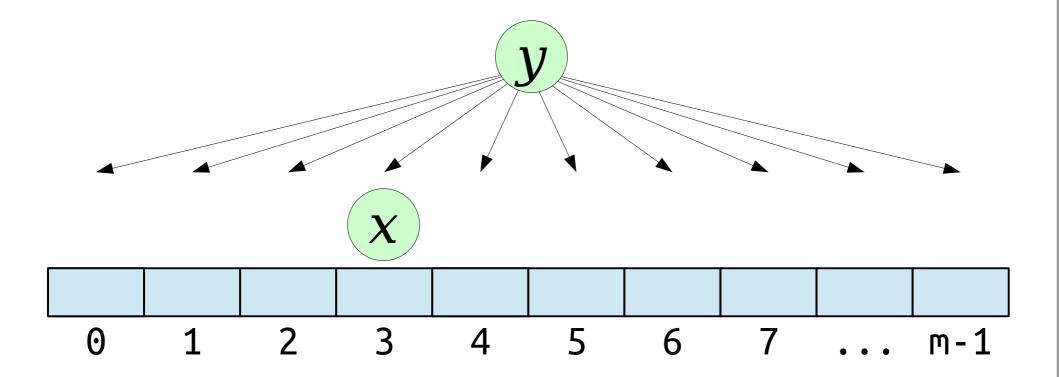


Each element should have an equal probability of being placed in each slot. For any $x \in \mathcal{U}$ and random $h \in \mathcal{H}$, the value of h(x) is uniform over [m].

Independence Property:

Where one element is placed shouldn't impact where a second goes.

For any distinct $x, y \in \mathcal{U}$ and random $h \in \mathcal{H}$, h(x) and h(y) are independent random variables.



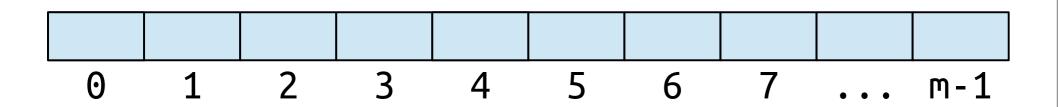
Each element should have an equal probability of being placed in each slot. For any $x \in \mathcal{U}$ and random $h \in \mathcal{H}$, the value of h(x) is uniform over [m].

Independence Property:

Where one element is placed shouldn't impact where a second goes.

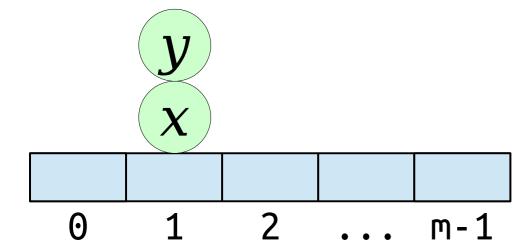
For any distinct $x, y \in \mathcal{U}$ and random $h \in \mathcal{H}$, h(x) and h(y) are independent random variables.

A family of hash functions \mathscr{H} is called **2-independent** (or **pairwise independent**) if it satisfies the distribution and independence properties.



For any distinct $x, y \in \mathcal{U}$ and random $h \in \mathcal{H}$, h(x) and h(y) are independent random variables.

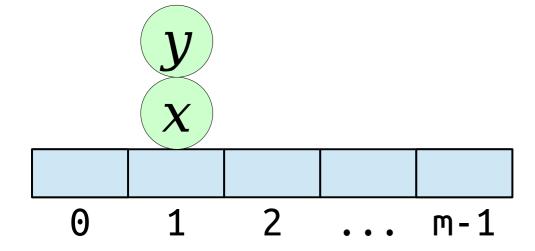
Intuition:



For any distinct $x, y \in \mathcal{U}$ and random $h \in \mathcal{H}$, h(x) and h(y) are independent random variables.

Intuition:

$$\Pr[h(x) = h(y)]$$

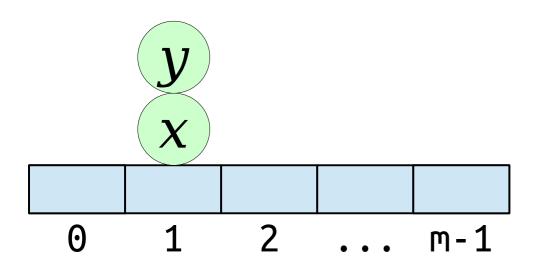


For any distinct $x, y \in \mathcal{U}$ and random $h \in \mathcal{H}$, h(x) and h(y) are independent random variables.

Intuition:

2-independence means any pair of elements is unlikely to collide.

$$\Pr[h(x) = h(y)]$$

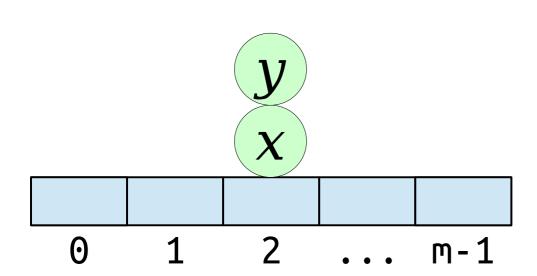


For any distinct $x, y \in \mathcal{U}$ and random $h \in \mathcal{H}$, h(x) and h(y) are independent random variables.

Intuition:

2-independence means any pair of elements is unlikely to collide.

$$\Pr[h(x) = h(y)]$$

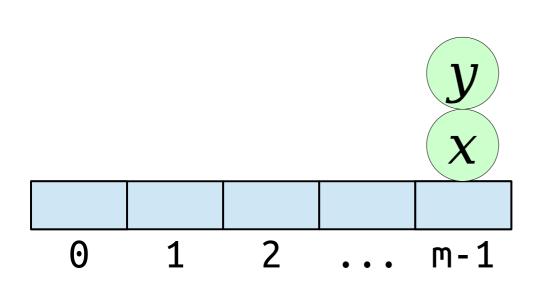


For any distinct $x, y \in \mathcal{U}$ and random $h \in \mathcal{H}$, h(x) and h(y) are independent random variables.

Intuition:

2-independence means any pair of elements is unlikely to collide.

$$\Pr[h(x) = h(y)]$$

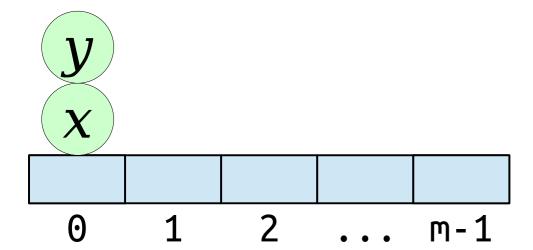


For any distinct $x, y \in \mathcal{U}$ and random $h \in \mathcal{H}$, h(x) and h(y) are independent random variables.

Intuition:

2-independence means any pair of elements is unlikely to collide.

$$\Pr[h(x) = h(y)]$$



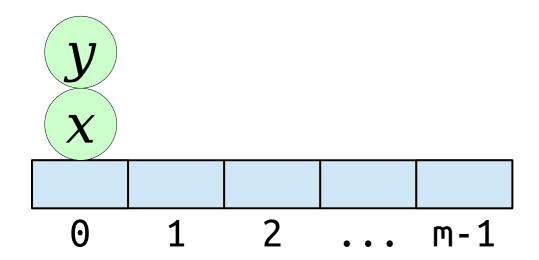
For any distinct $x, y \in \mathcal{U}$ and random $h \in \mathcal{H}$, h(x) and h(y) are independent random variables.

Intuition:

2-independence means any pair of elements is unlikely to collide.

$$\Pr[h(x) = h(y)]$$

$$= \sum_{i=0}^{m-1} \Pr[h(x) = i \land h(y) = i]$$

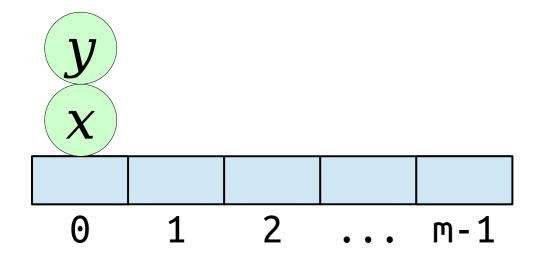


For any distinct $x, y \in \mathcal{U}$ and random $h \in \mathcal{H}$, h(x) and h(y) are independent random variables.

Intuition:

$$\Pr[h(x) = h(y)]$$

$$= \sum_{i=0}^{m-1} \Pr[h(x) = i \land h(y) = i]$$

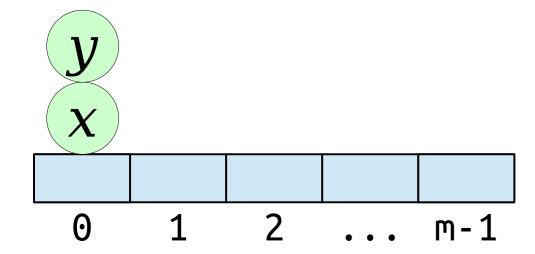


For any distinct $x, y \in \mathcal{U}$ and random $h \in \mathcal{H}$, h(x) and h(y) are independent random variables.

Intuition:

$$\Pr[h(x) = h(y)]$$

$$= \sum_{i=0}^{m-1} \Pr[h(x) = i \land h(y) = i]$$



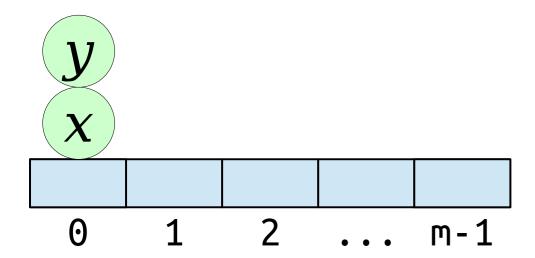
For any distinct $x, y \in \mathcal{U}$ and random $h \in \mathcal{H}$, h(x) and h(y) are independent random variables.

Intuition:

$$\Pr[h(x) = h(y)]$$

$$= \sum_{i=0}^{m-1} \Pr[h(x) = i \land h(y) = i]$$

$$= \sum_{i=0}^{m-1} \Pr[h(x) = i] \cdot \Pr[h(y) = i]$$



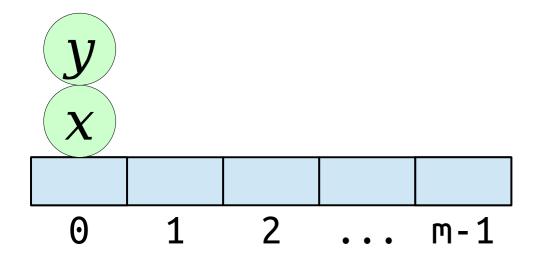
For any distinct $x, y \in \mathcal{U}$ and random $h \in \mathcal{H}$, h(x) and h(y) are independent random variables.

Intuition:

$$\Pr[h(x) = h(y)]$$

$$= \sum_{i=0}^{m-1} \Pr[h(x) = i \land h(y) = i]$$

$$= \sum_{i=0}^{m-1} \Pr[h(x) = i] \cdot \Pr[h(y) = i]$$



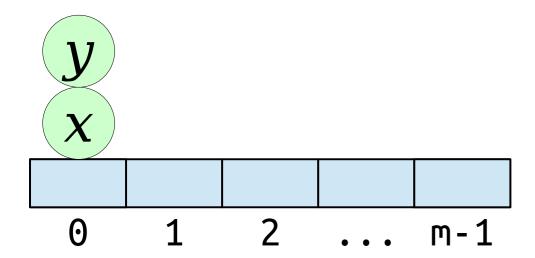
For any distinct $x, y \in \mathcal{U}$ and random $h \in \mathcal{H}$, h(x) and h(y) are independent random variables.

Intuition:

$$\Pr[h(x) = h(y)]$$

$$= \sum_{i=0}^{m-1} \Pr[h(x) = i \land h(y) = i]$$

$$= \sum_{i=0}^{m-1} \Pr[h(x) = i] \cdot \Pr[h(y) = i]$$



For any distinct $x, y \in \mathcal{U}$ and random $h \in \mathcal{H}$, h(x) and h(y) are independent random variables.

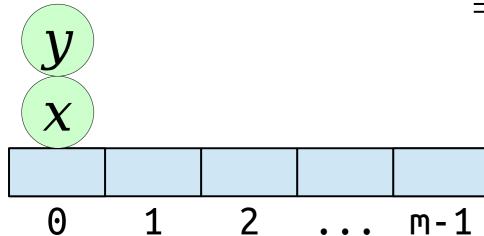
Intuition:

$$\Pr[h(x) = h(y)]$$

$$= \sum_{i=0}^{m-1} \Pr[h(x) = i \land h(y) = i]$$

$$= \sum_{i=0}^{m-1} \Pr[h(x) = i] \cdot \Pr[h(y) = i]$$

$$=\sum_{i=0}^{m-1}\frac{1}{m^2}$$



For any distinct $x, y \in \mathcal{U}$ and random $h \in \mathcal{H}$, h(x) and h(y) are independent random variables.

Intuition:

$$\Pr[h(x) = h(y)]$$
= $\sum_{i=0}^{m-1} \Pr[h(x) = i \land h(y) = i]$
= $\sum_{i=0}^{m-1} \Pr[h(x) = i] \cdot \Pr[h(y) = i]$
= $\sum_{i=0}^{m-1} \frac{1}{m^2}$

For any $x \in \mathcal{U}$ and random $h \in \mathcal{H}$, the value of h(x) is uniform over [m].

For any distinct $x, y \in \mathcal{U}$ and random $h \in \mathcal{H}$, h(x) and h(y) are independent random variables.

Intuition:

2-independence means any pair of elements is unlikely to collide.

$$\Pr[h(x) = h(y)]$$
= $\sum_{i=0}^{m-1} \Pr[h(x) = i \land h(y) = i]$
= $\sum_{i=0}^{m-1} \Pr[h(x) = i] \cdot \Pr[h(y) = i]$
= $\sum_{i=0}^{m-1} \frac{1}{m^2}$

For any $x \in \mathcal{U}$ and random $h \in \mathcal{H}$, the value of h(x) is uniform over [m].

For any distinct $x, y \in \mathcal{U}$ and random $h \in \mathcal{H}$, h(x) and h(y) are independent random variables.

Intuition:

2-independence means any pair of elements is unlikely to collide.

$$\Pr[h(x) = h(y)]$$
= $\sum_{i=0}^{m-1} \Pr[h(x) = i \land h(y) = i]$
= $\sum_{i=0}^{m-1} \Pr[h(x) = i] \cdot \Pr[h(y) = i]$
= $\sum_{i=0}^{m-1} \frac{1}{m^2}$

This is the same as if *h* were a truly random function.

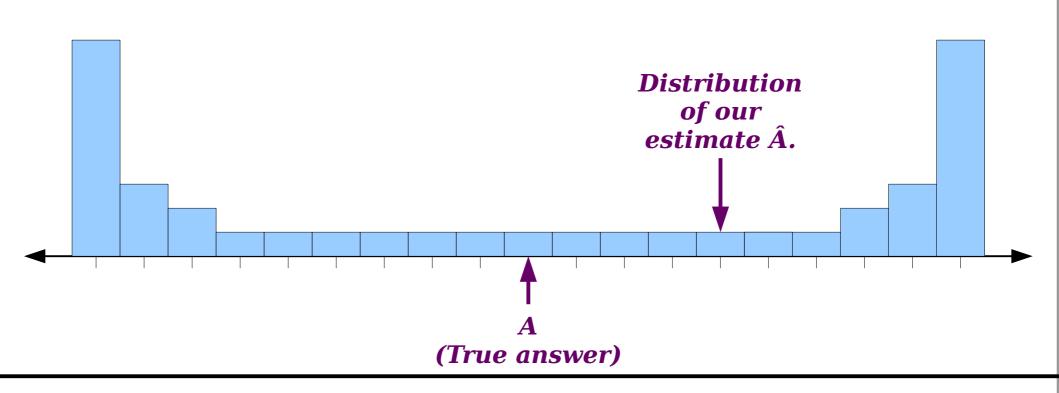
For more on hashing outside of Theoryland, check out *this Stack Exchange post*.

Approximating Quantities

What makes for a good "approximate" solution?

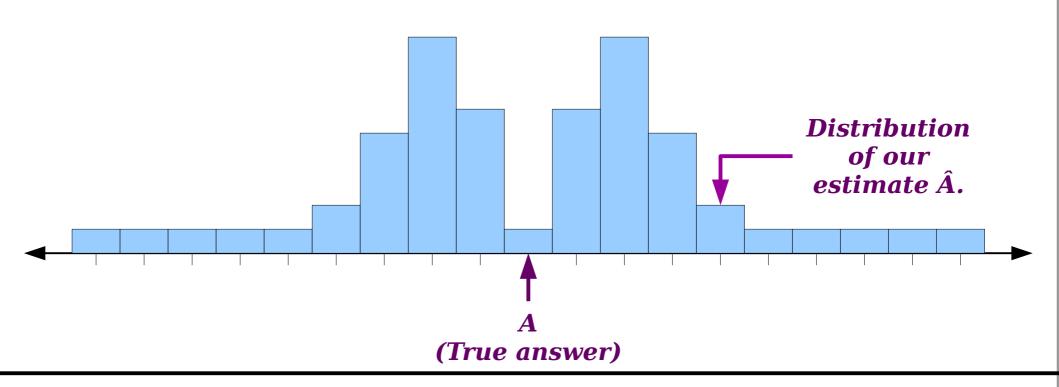
This would not make for a good estimate. However, we have $E[\hat{A}] = A$.

Observation 1: Being correct in expectation isn't sufficient.



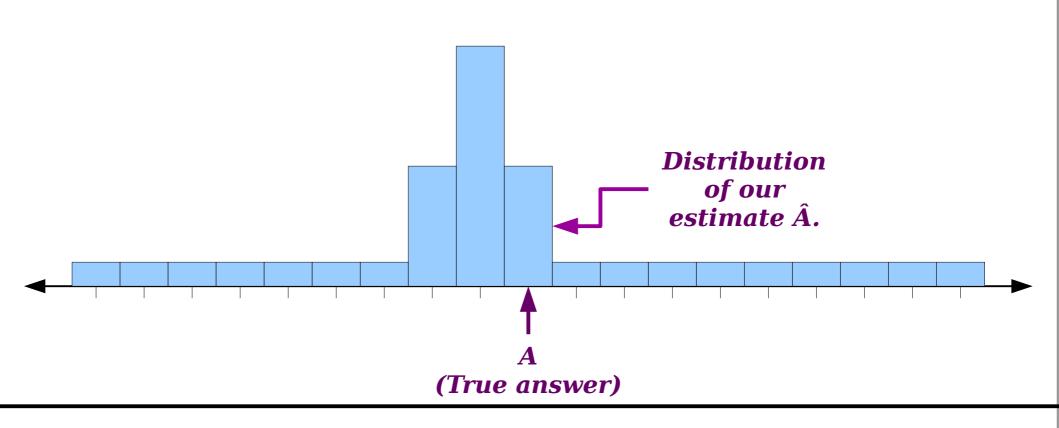
It's unlikely that we'll get the right answer, but we're probably going to be close.

Observation 2: The difference $|\hat{A} - A|$ between our estimate and the truth should ideally be small.

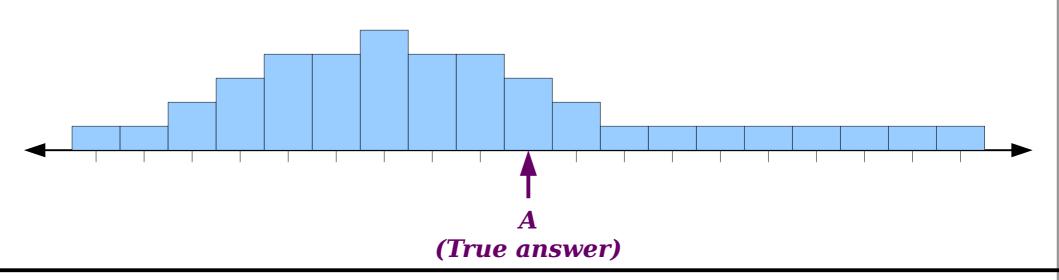


This estimate skews low, but it's very close to the true value.

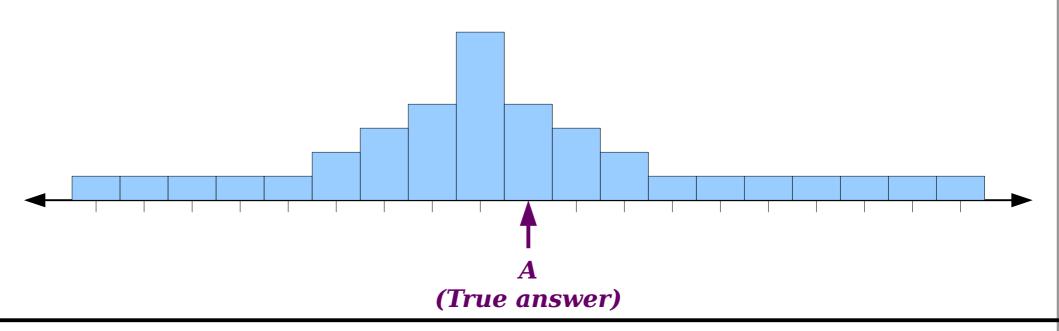
Observation 3: An estimate doesn't have to be unbiased to be useful.

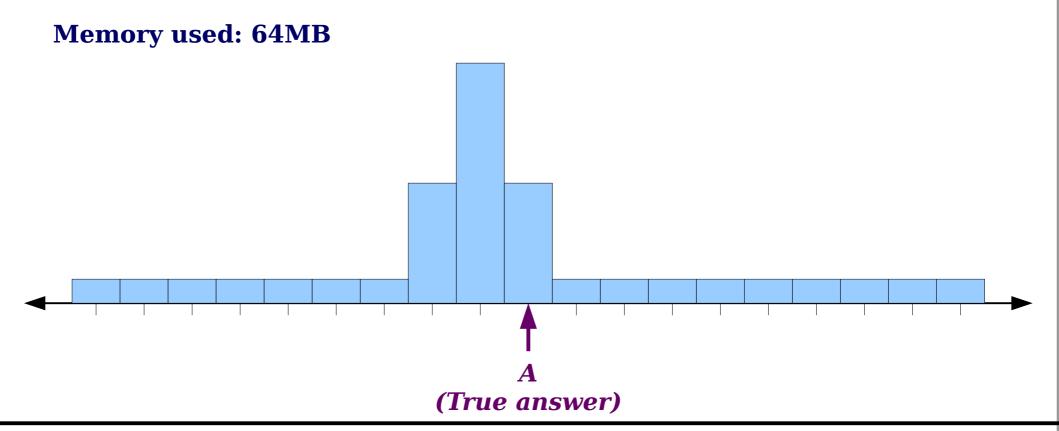


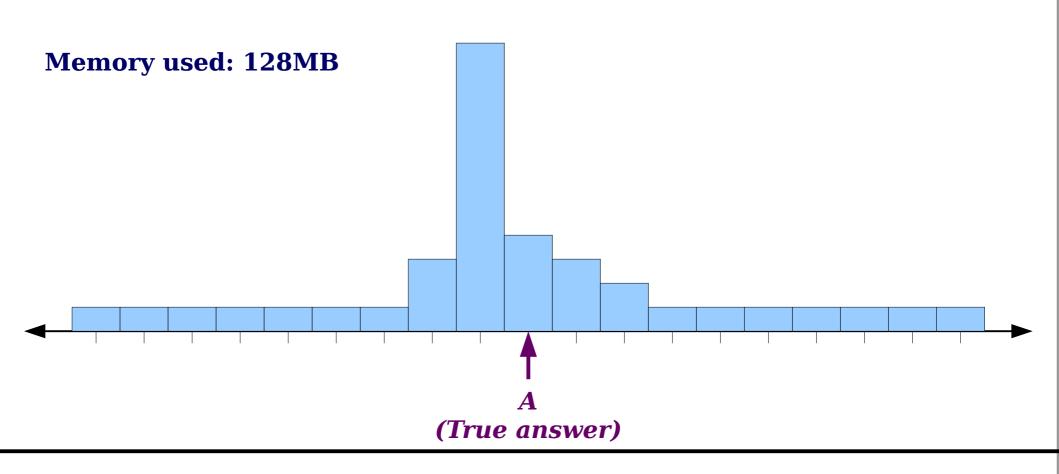
Memory used: 16MB



Memory used: 32MB

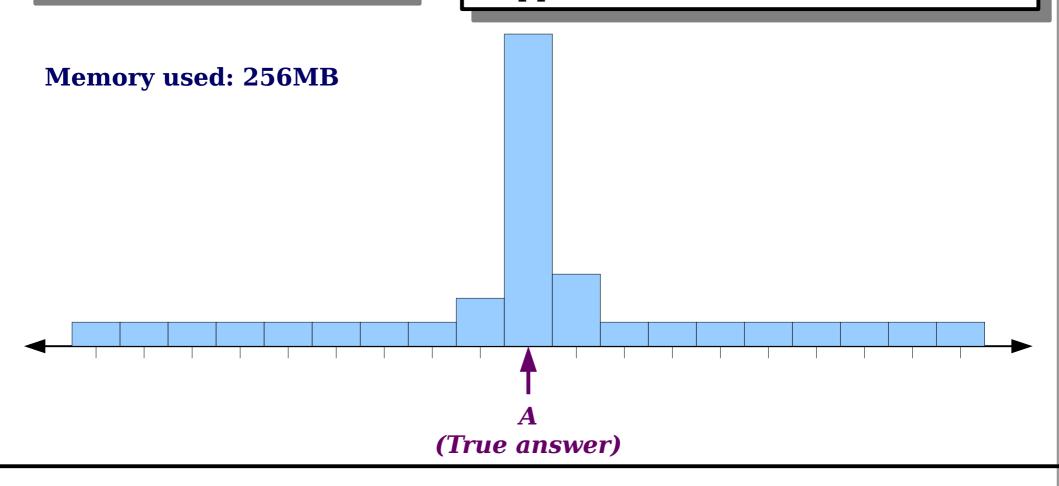






The more resources we allocate, the better our estimate should be.

Observation 4: A good approximation should be tunable.



We have two user-provided values

$$\varepsilon \in (0, 1]$$
 $\delta \in (0, 1]$

where ϵ represents **accuracy** and δ represents **confidence**.

Goal: Make an estimator \hat{A} for some quantity A where

With probability at least
$$1 - \delta$$
, $|\hat{A} - A| \le \varepsilon \cdot size(input)$

for some measure of the size of the input.

We have two user-provided values

$$\varepsilon \in (0, 1]$$
 $\delta \in (0, 1]$

where ϵ represents **accuracy** and δ represents **confidence**.

Goal: Make an estimator \hat{A} for some quantity A where

With probability at least
$$1 - \delta$$
, $|\hat{A} - A| \le \epsilon \cdot size(input)$

for some measure of the size of the input.

We have two user-provided values

$$\varepsilon \in (0, 1]$$
 $\delta \in (0, 1]$

where ϵ represents **accuracy** and δ represents **confidence**.

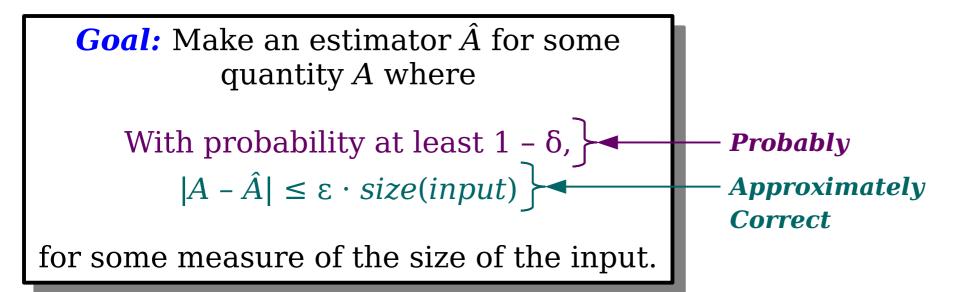
Goal: Make an estimator \hat{A} for some quantity A where

With probability at least $1 - \delta$, $|\hat{A} - A| \leq \epsilon \cdot size(input)$ Probably
Approximately
Correct

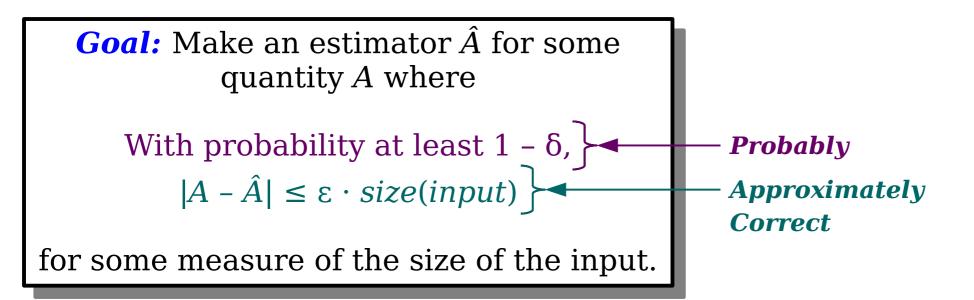
for some measure of the size of the input.

Goal: Make an estimator \hat{A} for some quantity A where

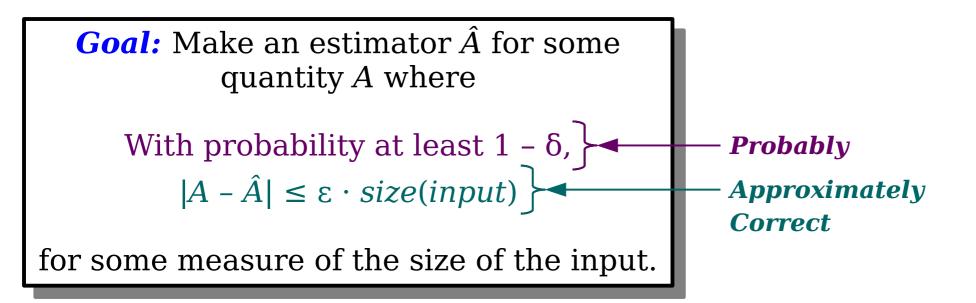
With probability at least $1 - \delta$, $|A - \hat{A}| \leq \epsilon \cdot size(input)$ Approximately Correct for some measure of the size of the input.



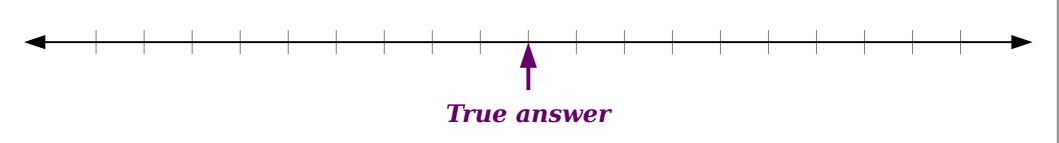
$$\delta = \frac{1}{2}$$
 $\epsilon \text{ small}$

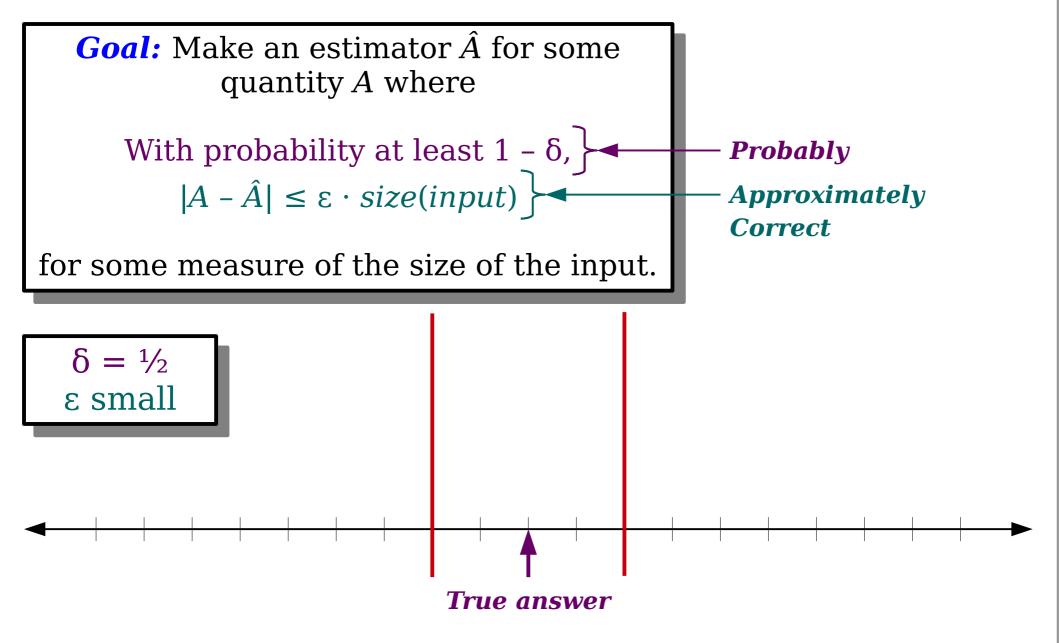


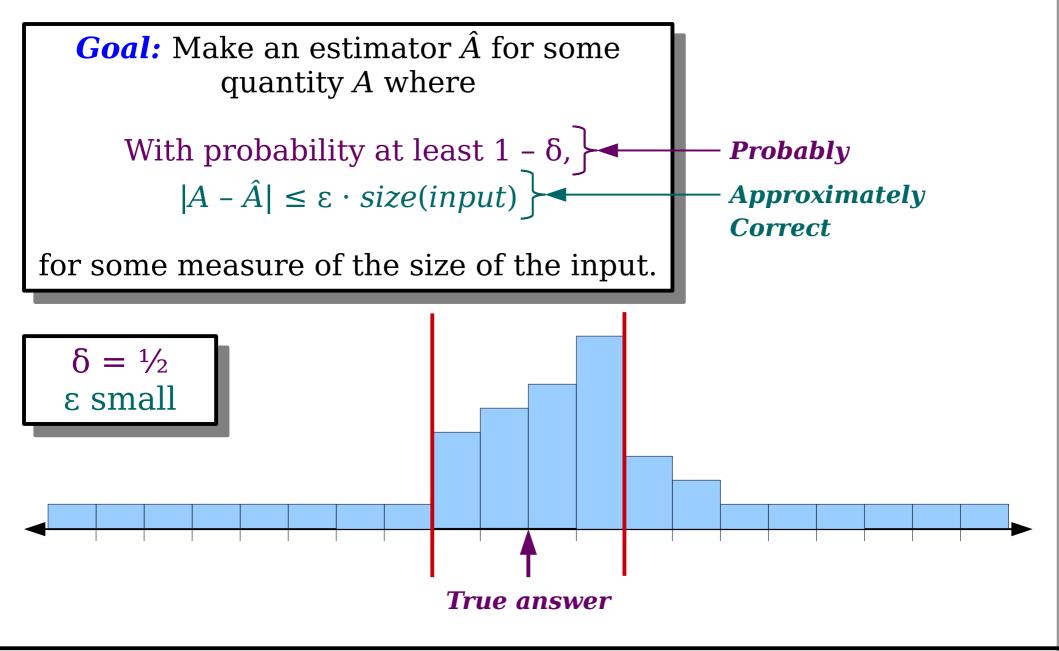
$$\delta = \frac{1}{2}$$
 $\epsilon \text{ small}$

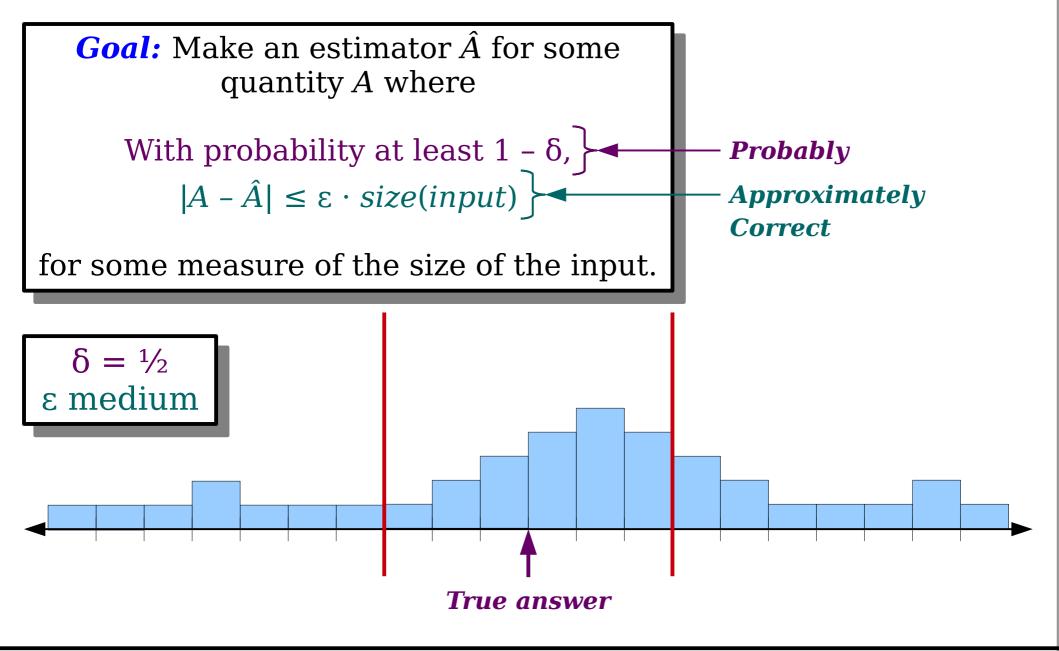


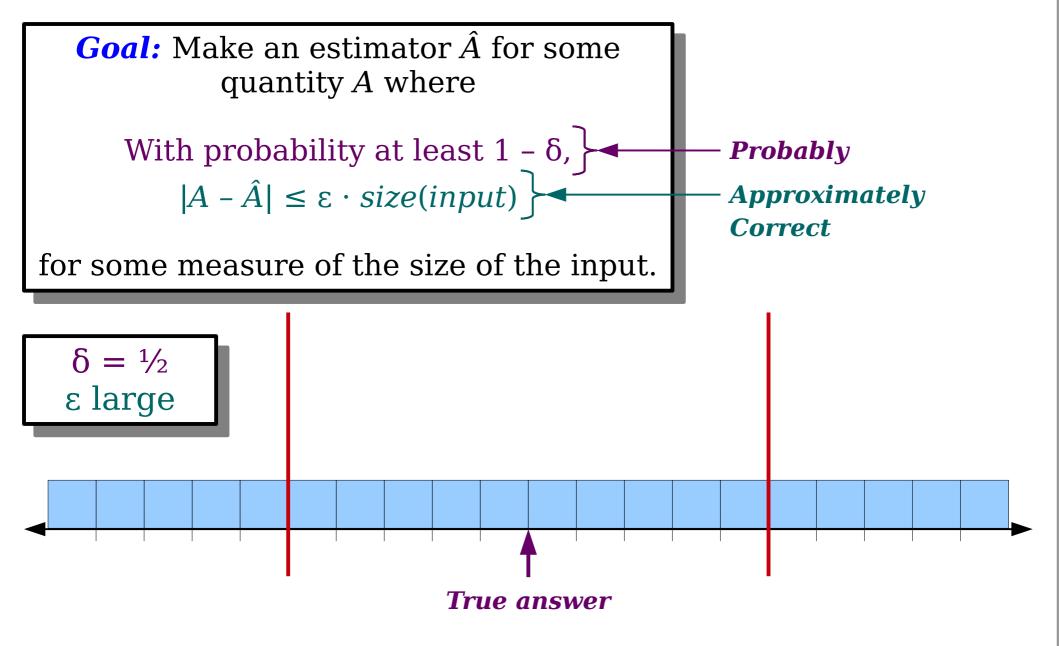
$$\delta = \frac{1}{2}$$
 $\epsilon \text{ small}$

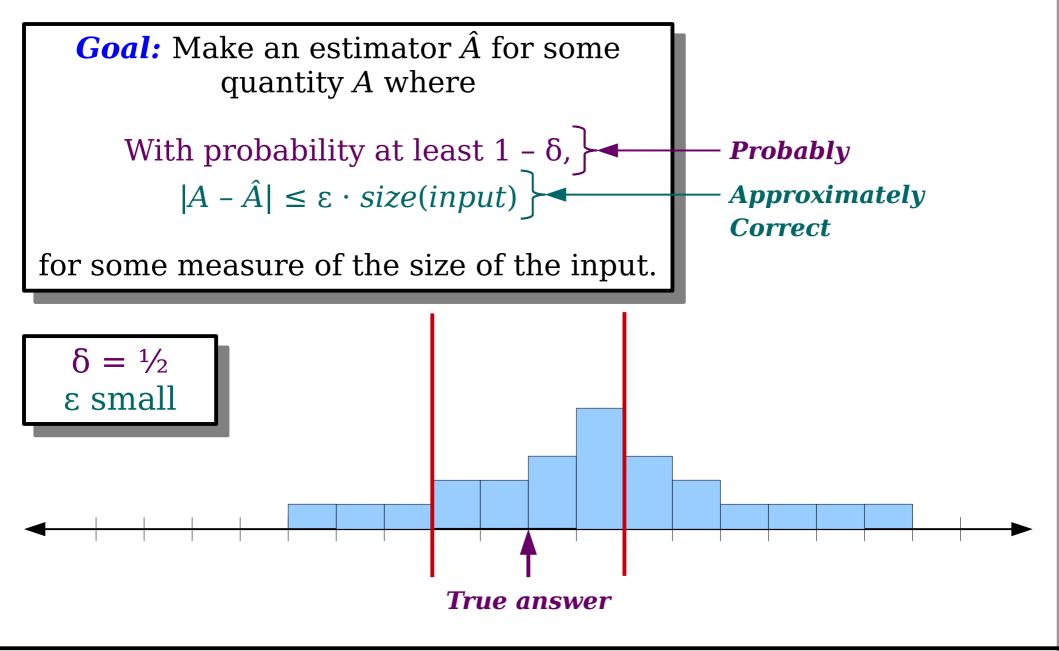


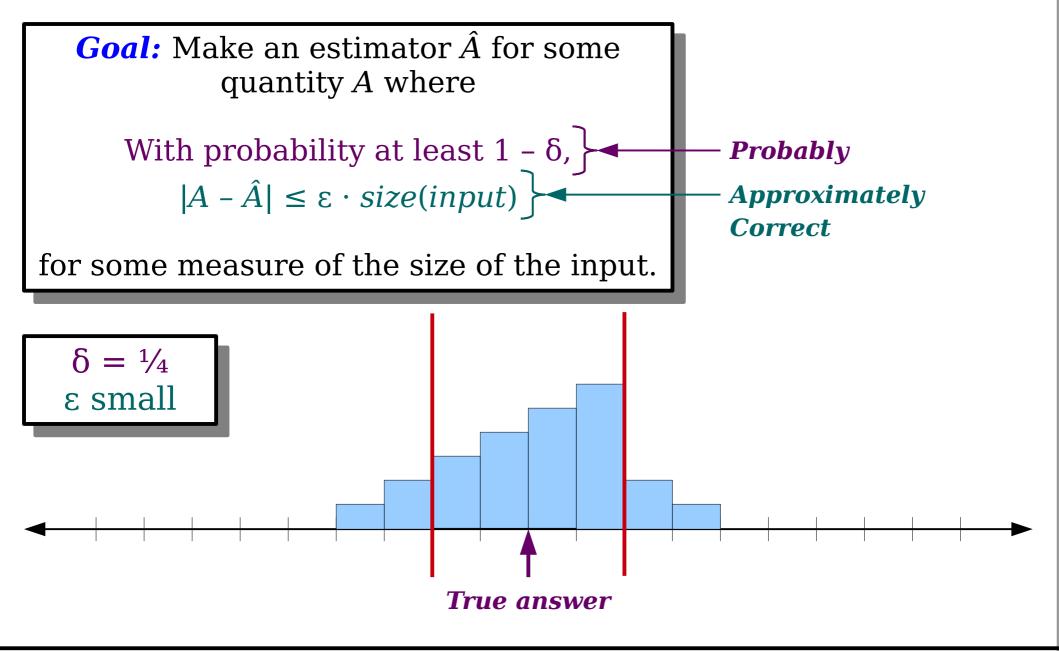


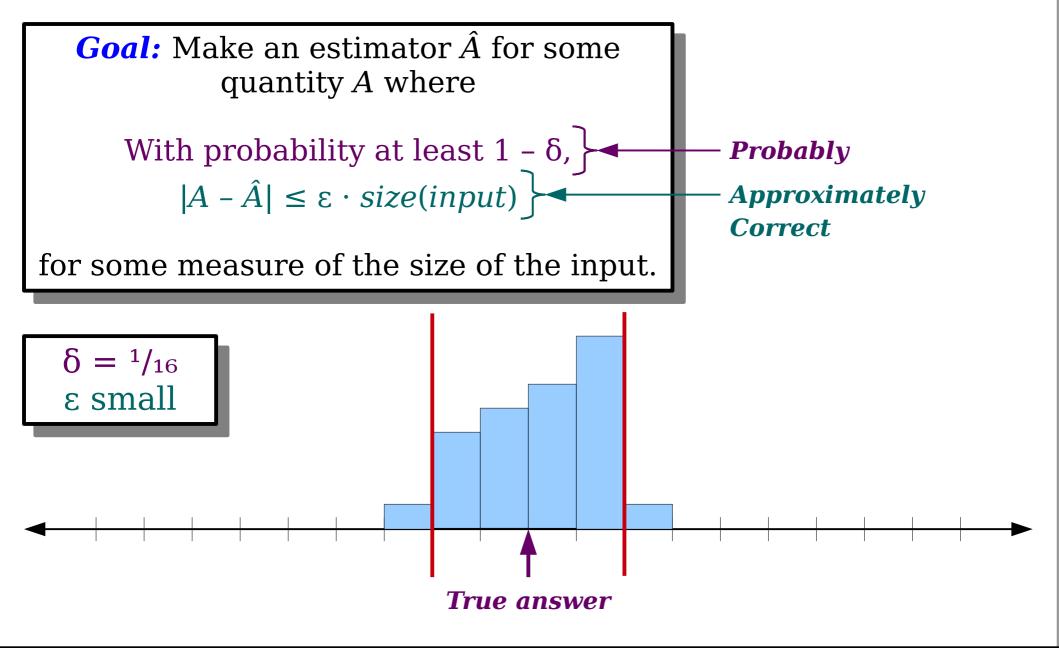












Frequency Estimation

Frequency Estimators

- A *frequency estimator* is a data structure supporting the following operations:
 - *increment*(*x*), which increments the number of times that *x* has been seen, and
 - estimate(x), which returns an estimate of the frequency of x.
- Using BSTs, we can solve this in space $\Theta(n)$ with worst-case $O(\log n)$ costs on the operations.
- Using hash tables, we can solve this in space $\Theta(n)$ with expected O(1) costs on the operations.

Frequency Estimators

- Frequency estimation has many applications:
 - Search engines: Finding frequent search queries.
 - Network routing: Finding common source and destination addresses.
- In these applications, $\Theta(n)$ memory can be impractical.
- *Goal:* Get *approximate* answers to these queries in sublinear space.

The Count-Min Sketch

How to Build an Estimator

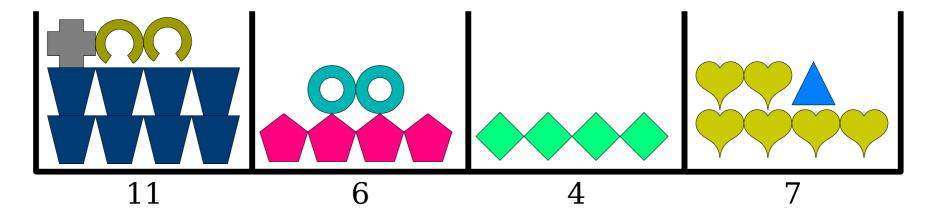
- 1. Design a simple data structure that, intuitively, gives you a good estimate.
- 2. Use a *sum of indicator variables* and *linearity of expectation* to prove that, on expectation, the data structure is pretty close to correct.
- 3. Use a *concentration inequality* to show that, with decent probability, the data structure's output is close to its expectation.
- 4. Run multiple copies of the data structure in parallel to amplify the success probability.

How to Build an Estimator

- 1. Design a simple data structure that, intuitively, gives you a good estimate.
- 2. Use a *sum of indicator variables* and *linearity of expectation* to prove that, on expectation, the data structure is pretty close to correct.
- 3. Use a *concentration inequality* to show that, with decent probability, the data structure's output is close to its expectation.
- 4. Run multiple copies of the data structure in parallel to amplify the success probability.

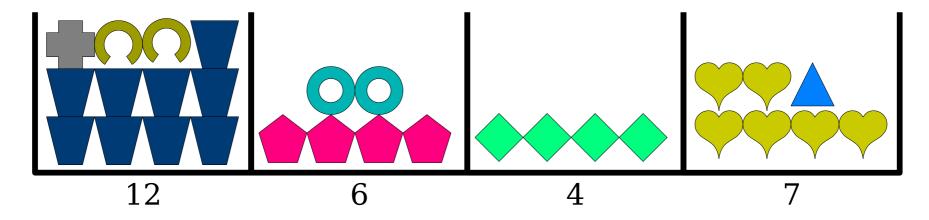
Revisiting the Exact Solution

- In the exact solution to the frequency estimation problem, we maintained a single counter for each distinct element. This is too space-inefficient.
- *Idea*: Store a fixed number of counters and assign a counter to each $x_i \in \mathcal{U}$. Multiple x_i 's might be assigned to the same counter.
- To *increment*(x), increment the counter for x.
- To *estimate*(x), read the value of the counter for x.



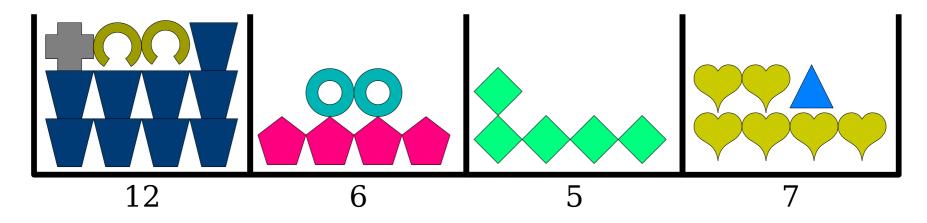
Revisiting the Exact Solution

- In the exact solution to the frequency estimation problem, we maintained a single counter for each distinct element. This is too space-inefficient.
- *Idea*: Store a fixed number of counters and assign a counter to each $x_i \in \mathcal{U}$. Multiple x_i 's might be assigned to the same counter.
- To *increment*(x), increment the counter for x.
- To *estimate*(x), read the value of the counter for x.



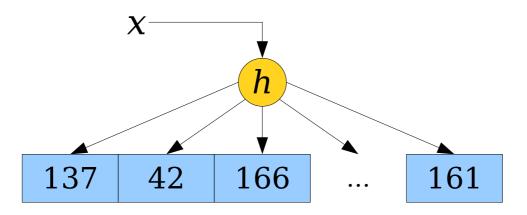
Revisiting the Exact Solution

- In the exact solution to the frequency estimation problem, we maintained a single counter for each distinct element. This is too space-inefficient.
- *Idea*: Store a fixed number of counters and assign a counter to each $x_i \in \mathcal{U}$. Multiple x_i 's might be assigned to the same counter.
- To *increment*(x), increment the counter for x.
- To *estimate*(x), read the value of the counter for x.



Our Initial Structure

- We can model "assigning each x_i to a counter" by using hash functions.
- Choose, from a family of 2-independent hash functions \mathcal{H} , a uniformly-random hash function $h: \mathcal{U} \to [w]$.
- Create an array count of w counters, each initially zero.
 - We'll choose w later on.
- To *increment*(x), increment count[h(x)].
- To **estimate**(x), return **count**[h(x)].



Analyzing our Structure

For each $x_i \in \mathcal{U}$, let \mathbf{a}_i denote the number of times we've seen x_i .

Similarly, let \hat{a}_i denote our estimated value of the frequency of x_i .

Goal: Bound the probability that the error $(\hat{a}_i - a_i)$ is too high.

Idea: Think of our element frequencies a_1 , a_2 , a_3 , ... as a vector

$$a = [a_1, a_2, a_3, ...].$$

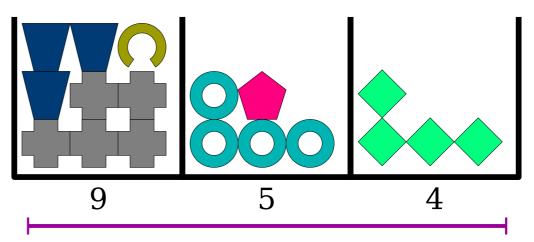
The total number of objects is the sum of the vector entries.

This is called the L_1 norm of a, and is denoted $||a||_1$:

$$\|\boldsymbol{a}\|_1 = \sum_i |\boldsymbol{a}_i|$$

There are $\|a\|_1$ total elements distributed across w buckets. We're using a 2-independent hash family.

Reasonable guess: each bin has $\|a\|_1$ / w elements in it, so $\hat{a}_i - a_i \le \|a\|_1$ / w



Number of buckets: w

Question: Intuitively, what should we expect our approximation error to be?

How to Build an Estimator

- 1. Design a simple data structure that, intuitively, gives you a good estimate.
- 2. Use a *sum of indicator variables* and *linearity of expectation* to prove that, on expectation, the data structure is pretty close to correct.
- 3. Use a *concentration inequality* to show that, with decent probability, the data structure's output is close to its expectation.
- 4. Run multiple copies of the data structure in parallel to amplify the success probability.

How to Build an Estimator

- 1. Design a simple data structure that, intuitively, gives you a good estimate.
- 2. Use a *sum of indicator variables* and *linearity of expectation* to prove that, on expectation, the data structure is pretty close to correct.
- 3. Use a *concentration inequality* to show that, with decent probability, the data structure's output is close to its expectation.
- 4. Run multiple copies of the data structure in parallel to amplify the success probability.

- Let's look at $\hat{a}_i = \text{count}[h(x_i)]$ for some choice of x_i .
- For each element x_j :
 - If $h(x_i) = h(x_j)$, then x_j contributes a_j to count $[h(x_i)]$.
 - If $h(x_i) \neq h(x_j)$, then x_j contributes 0 to **count**[$h(x_i)$].

- Let's look at $\hat{a}_i = \text{count}[h(x_i)]$ for some choice of x_i .
- For each element x_j :
 - If $h(x_i) = h(x_j)$, then x_j contributes a_j to count $[h(x_i)]$.
 - If $h(x_i) \neq h(x_j)$, then x_j contributes 0 to **count**[$h(x_i)$].
- To pin this down precisely, let's define a set of random variables $X_1, X_2, ...,$ as follows:

$$X_{j} = \begin{cases} 1 & \text{if } h(x_{i}) = h(x_{j}) \\ 0 & \text{otherwise} \end{cases}$$

- Let's look at $\hat{a}_i = \text{count}[h(x_i)]$ for some choice of x_i .
- For each element x_j :
 - If $h(x_i) = h(x_j)$, then x_j contributes a_j to count $[h(x_i)]$.
 - If $h(x_i) \neq h(x_j)$, then x_j contributes 0 to **count**[$h(x_i)$].
- To pin this down precisely, let's define a set of random variables $X_1, X_2, ...,$ as follows:

$$X_{j} = \begin{cases} 1 & \text{if } h(x_{i}) = h(x_{j}) \\ 0 & \text{otherwise} \end{cases}$$

Each of these variables is called an *indicator* random variable, since it "indicates" whether some event occurs.

- Let's look at $\hat{a}_i = \text{count}[h(x_i)]$ for some choice of x_i .
- For each element x_j :
 - If $h(x_i) = h(x_j)$, then x_j contributes a_j to count $[h(x_i)]$.
 - If $h(x_i) \neq h(x_j)$, then x_j contributes 0 to count[$h(x_i)$].
- To pin this down precisely, let's define a set of random variables $X_1, X_2, ...,$ as follows:

$$X_{j} = \begin{cases} 1 & \text{if } h(x_{i}) = h(x_{j}) \\ 0 & \text{otherwise} \end{cases}$$

• The value of $\hat{a}_i - a_i$ is then given by

$$\hat{\boldsymbol{a}}_i - \boldsymbol{a}_i = \sum_{j \neq i} \boldsymbol{a}_j X_j$$

$$\mathbf{E}[\hat{\boldsymbol{a}}_i - \boldsymbol{a}_i] = \mathbf{E}[\sum_{j \neq i} \boldsymbol{a}_j X_j]$$

$$\begin{split} \mathbf{E}[\hat{\boldsymbol{a}}_i - \boldsymbol{a}_i] &= \mathbf{E}[\sum_{j \neq i} \boldsymbol{a}_j X_j] \\ &= \sum_{j \neq i} \mathbf{E}[\boldsymbol{a}_j X_j] \end{split}$$

$$E[\hat{\boldsymbol{a}}_i - \boldsymbol{a}_i] = E[\sum_{j \neq i} \boldsymbol{a}_j X_j]$$
$$= \sum_{j \neq i} E[\boldsymbol{a}_j X_j]$$

This follows from *linearity*of expectation. We'll use
this property extensively
over the next few days.

$$\begin{split} \mathbf{E}[\hat{\boldsymbol{a}}_i - \boldsymbol{a}_i] &= \mathbf{E}[\sum_{j \neq i} \boldsymbol{a}_j X_j] \\ &= \sum_{j \neq i} \mathbf{E}[\boldsymbol{a}_j X_j] \\ &= \sum_{j \neq i} \boldsymbol{a}_j \mathbf{E}[X_j] \end{split}$$

$$E[\hat{\boldsymbol{a}}_i - \boldsymbol{a}_i] = E[\sum_{j \neq i} \boldsymbol{a}_j X_j]$$

$$= \sum_{j \neq i} E[\boldsymbol{a}_j X_j]$$

$$= \sum_{j \neq i} \boldsymbol{a}_j E[X_j]$$

The values of a_j are not random. The randomness comes from our choice of hash function.

$$\begin{split} \mathbf{E}[\hat{\boldsymbol{a}}_i - \boldsymbol{a}_i] &= \mathbf{E}[\sum_{j \neq i} \boldsymbol{a}_j X_j] \\ &= \sum_{j \neq i} \mathbf{E}[\boldsymbol{a}_j X_j] \\ &= \sum_{j \neq i} \boldsymbol{a}_j \mathbf{E}[X_j] \end{split}$$

$$\begin{split} \mathbf{E}[\hat{\boldsymbol{a}}_i - \boldsymbol{a}_i] &= \mathbf{E}[\sum_{j \neq i} \boldsymbol{a}_j X_j] \\ &= \sum_{j \neq i} \mathbf{E}[\boldsymbol{a}_j X_j] \\ &= \sum_{j \neq i} \boldsymbol{a}_j \mathbf{E}[X_j] \end{split}$$

$$\begin{split} \mathbf{E}[\hat{\boldsymbol{a}}_{i} - \boldsymbol{a}_{i}] &= \mathbf{E}[\sum_{j \neq i} \boldsymbol{a}_{j} X_{j}] \\ &= \sum_{j \neq i} \mathbf{E}[\boldsymbol{a}_{j} X_{j}] \\ &= \sum_{j \neq i} \boldsymbol{a}_{j} \mathbf{E}[X_{j}] \end{split}$$

$$E[X_i] =$$

$$X_{j} = \begin{cases} 1 & \text{if } h(x_{i}) = h(x_{j}) \\ 0 & \text{otherwise} \end{cases}$$

$$\begin{split} \mathbf{E}[\hat{\boldsymbol{a}}_{i} - \boldsymbol{a}_{i}] &= \mathbf{E}[\sum_{j \neq i} \boldsymbol{a}_{j} X_{j}] \\ &= \sum_{j \neq i} \mathbf{E}[\boldsymbol{a}_{j} X_{j}] \\ &= \sum_{j \neq i} \boldsymbol{a}_{j} \mathbf{E}[X_{j}] \end{split}$$

$$E[X_j] = 1 \cdot Pr[h(x_i) = h(x_j)] + 0 \cdot Pr[h(x_i) \neq h(x_j)]$$

$$X_{j} = \begin{cases} 1 & \text{if } h(x_{i}) = h(x_{j}) \\ 0 & \text{otherwise} \end{cases}$$

$$\begin{split} \mathbf{E}[\hat{\boldsymbol{a}}_i - \boldsymbol{a}_i] &= \mathbf{E}[\sum_{j \neq i} \boldsymbol{a}_j X_j] \\ &= \sum_{j \neq i} \mathbf{E}[\boldsymbol{a}_j X_j] \\ &= \sum_{j \neq i} \boldsymbol{a}_j \mathbf{E}[X_j] \end{split}$$

$$E[X_j] = 1 \cdot Pr[h(x_i) = h(x_j)] + 0 \cdot Pr[h(x_i) \neq h(x_j)]$$

$$\begin{split} \mathbf{E}[\hat{\boldsymbol{a}}_{i} - \boldsymbol{a}_{i}] &= \mathbf{E}[\sum_{j \neq i} \boldsymbol{a}_{j} X_{j}] \\ &= \sum_{j \neq i} \mathbf{E}[\boldsymbol{a}_{j} X_{j}] \\ &= \sum_{j \neq i} \boldsymbol{a}_{j} \mathbf{E}[X_{j}] \end{split}$$

$$\begin{split} \mathbf{E}[X_j] &= 1 \cdot \Pr[h(x_i) = h(x_j)] + 0 \cdot \Pr[h(x_i) \neq h(x_j)] \\ &= \Pr[h(x_i) = h(x_j)] \end{split}$$

$$\begin{split} \mathbf{E}[\hat{\boldsymbol{a}}_{i} - \boldsymbol{a}_{i}] &= \mathbf{E}[\sum_{j \neq i} \boldsymbol{a}_{j} \boldsymbol{X}_{j}] \\ &= \sum_{j \neq i} \mathbf{E}[\boldsymbol{a}_{j} \boldsymbol{X}_{j}] \\ &= \sum_{j \neq i} \boldsymbol{a}_{j} \mathbf{E}[\boldsymbol{X}_{j}] \end{split}$$

$$\begin{split} \mathbf{E}[X_j] &= \mathbf{1} \cdot \Pr[h(x_i) = h(x_j)] + \mathbf{0} \cdot \Pr[h(x_i) \neq h(x_j)] \\ &= \Pr[h(x_i) = h(x_j)] \end{split}$$

If X is an indicator variable for some event \mathcal{E} , then $\mathbf{E}[X] = \mathbf{Pr}[\mathcal{E}]$. This is really useful when using linearity of expectation!

$$\begin{split} \mathbf{E}[\hat{\boldsymbol{a}}_{i} - \boldsymbol{a}_{i}] &= \mathbf{E}[\sum_{j \neq i} \boldsymbol{a}_{j} X_{j}] \\ &= \sum_{j \neq i} \mathbf{E}[\boldsymbol{a}_{j} X_{j}] \\ &= \sum_{j \neq i} \boldsymbol{a}_{j} \mathbf{E}[X_{j}] \end{split}$$

$$\begin{split} \mathbf{E}[X_j] &= 1 \cdot \Pr[h(x_i) = h(x_j)] + 0 \cdot \Pr[h(x_i) \neq h(x_j)] \\ &= \Pr[h(x_i) = h(x_j)] \end{split}$$

$$\begin{split} \mathbf{E}[\hat{\boldsymbol{a}}_i - \boldsymbol{a}_i] &= \mathbf{E}[\sum_{j \neq i} \boldsymbol{a}_j X_j] \\ &= \sum_{j \neq i} \mathbf{E}[\boldsymbol{a}_j X_j] \\ &= \sum_{j \neq i} \boldsymbol{a}_j \mathbf{E}[X_j] \end{split}$$

$$\begin{split} \mathbf{E}[X_j] &= 1 \cdot \Pr[h(x_i) = h(x_j)] + 0 \cdot \Pr[h(x_i) \neq h(x_j)] \\ &= \Pr[h(x_i) = h(x_j)] \end{split}$$

Hey, we saw this earlier!

$$\begin{split} \mathbf{E}[\hat{\boldsymbol{a}}_i - \boldsymbol{a}_i] &= \mathbf{E}[\sum_{j \neq i} \boldsymbol{a}_j X_j] \\ &= \sum_{j \neq i} \mathbf{E}[\boldsymbol{a}_j X_j] \\ &= \sum_{j \neq i} \boldsymbol{a}_j \mathbf{E}[X_j] \end{split}$$

$$\begin{split} \mathbf{E}[X_j] &= 1 \cdot \Pr[h(x_i) = h(x_j)] + 0 \cdot \Pr[h(x_i) \neq h(x_j)] \\ &= \Pr[h(x_i) = h(x_j)] \\ &= \frac{1}{\mathbf{Hev. we sa}} \end{split}$$

Hey, we saw this earlier!

$$\begin{split} \mathbf{E}[\hat{\boldsymbol{a}}_{i} - \boldsymbol{a}_{i}] &= \mathbf{E}[\sum_{j \neq i} \boldsymbol{a}_{j} X_{j}] \\ &= \sum_{j \neq i} \mathbf{E}[\boldsymbol{a}_{j} X_{j}] \\ &= \sum_{j \neq i} \boldsymbol{a}_{j} \mathbf{E}[X_{j}] \\ &= \sum_{j \neq i} \frac{\boldsymbol{a}_{j}}{w} \end{split}$$

$$\begin{split} \mathbf{E}[X_j] &= 1 \cdot \Pr[h(x_i) = h(x_j)] + 0 \cdot \Pr[h(x_i) \neq h(x_j)] \\ &= \Pr[h(x_i) = h(x_j)] \\ &= \frac{1}{\mathbf{Hev. we sa}} \end{split}$$

Hey, we saw this earlier!

$$\begin{split} \mathbf{E}[\hat{\boldsymbol{a}}_{i} - \boldsymbol{a}_{i}] &= \mathbf{E}[\sum_{j \neq i} \boldsymbol{a}_{j} X_{j}] \\ &= \sum_{j \neq i} \mathbf{E}[\boldsymbol{a}_{j} X_{j}] \\ &= \sum_{j \neq i} \boldsymbol{a}_{j} \mathbf{E}[X_{j}] \\ &= \sum_{j \neq i} \frac{\boldsymbol{a}_{j}}{w} \end{split}$$

$$\begin{split} \mathbf{E}[X_j] &= \mathbf{1} \cdot \Pr[h(x_i) = h(x_j)] + \mathbf{0} \cdot \Pr[h(x_i) \neq h(x_j)] \\ &= \Pr[h(x_i) = h(x_j)] \\ &= \frac{1}{w} \end{split}$$

$$\begin{split} \mathbf{E}[\hat{\boldsymbol{a}}_{i} - \boldsymbol{a}_{i}] &= \mathbf{E}[\sum_{j \neq i} \boldsymbol{a}_{j} X_{j}] \\ &= \sum_{j \neq i} \mathbf{E}[\boldsymbol{a}_{j} X_{j}] \\ &= \sum_{j \neq i} \boldsymbol{a}_{j} \mathbf{E}[X_{j}] \\ &= \sum_{j \neq i} \frac{\boldsymbol{a}_{j}}{w} \\ &\leq \frac{\|\boldsymbol{a}\|_{1}}{w} \end{split}$$

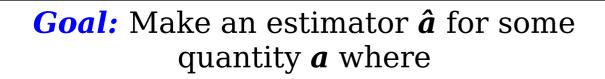
$$\begin{split} \mathbf{E}[X_j] &= \mathbf{1} \cdot \Pr[h(x_i) = h(x_j)] + \mathbf{0} \cdot \Pr[h(x_i) \neq h(x_j)] \\ &= \Pr[h(x_i) = h(x_j)] \\ &= \frac{1}{w} \end{split}$$

How to Build an Estimator

- 1. Design a simple data structure that, intuitively, gives you a good estimate.
- 2. Use a *sum of indicator variables* and *linearity of expectation* to prove that, on expectation, the data structure is pretty close to correct.
- 3. Use a *concentration inequality* to show that, with decent probability, the data structure's output is close to its expectation.
- 4. Run multiple copies of the data structure in parallel to amplify the success probability.

How to Build an Estimator

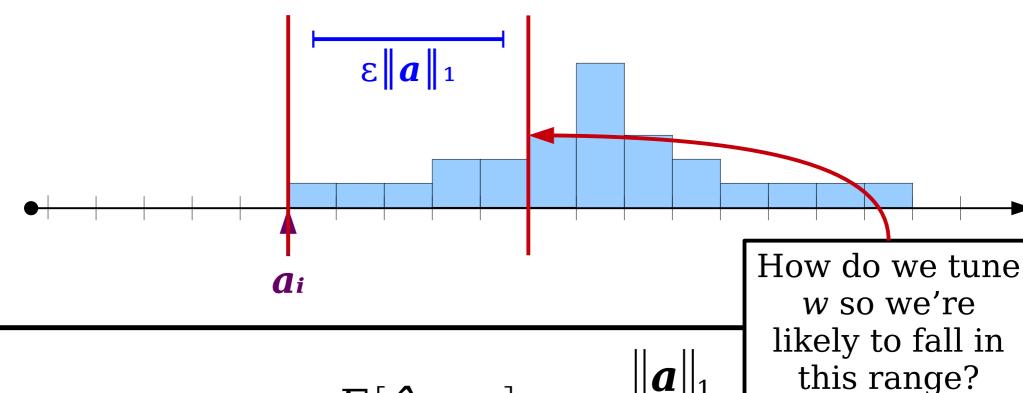
- 1. Design a simple data structure that, intuitively, gives you a good estimate.
- 2. Use a *sum of indicator variables* and *linearity of expectation* to prove that, on expectation, the data structure is pretty close to correct.
- 3. Use a *concentration inequality* to show that, with decent probability, the data structure's output is close to its expectation.
- 4. Run multiple copies of the data structure in parallel to amplify the success probability.



With probability at least
$$1 - \delta$$
, $|\hat{a} - a| \le \varepsilon \cdot size(input)$

for some measure of the size of the input.

- Probably - Approximately Correct



$$E[\hat{\boldsymbol{a}}_i - \boldsymbol{a}_i] \leq \frac{\|\boldsymbol{a}\|_1}{w}$$

$$\Pr\left[\hat{\boldsymbol{a}}_{i}-\boldsymbol{a}_{i} > \varepsilon \|\boldsymbol{a}\|_{1}\right]$$

$$\Pr\left[\hat{\boldsymbol{a}}_{i}-\boldsymbol{a}_{i}\right]>\varepsilon\|\boldsymbol{a}\|_{1}$$

$$\Pr\left[\hat{\boldsymbol{a}}_{i}-\boldsymbol{a}_{i}\right]>\varepsilon\|\boldsymbol{a}\|_{1}$$

We don't know the exact distribution of this random variable.

However, we have a **one-sided error**: our estimate can never be lower than the true value. This means that $\hat{a}_i - a_i \ge 0$.

Markov's inequality says that if *X* is a nonnegative random variable, then

$$\Pr[X \geq c] \leq \frac{E[X]}{c}$$

$$\Pr\left[\hat{\boldsymbol{a}}_{i} - \boldsymbol{a}_{i}\right] > \varepsilon \|\boldsymbol{a}\|_{1}$$

$$\leq \frac{\operatorname{E}\left[\hat{\boldsymbol{a}}_{i} - \boldsymbol{a}_{i}\right]}{\varepsilon \|\boldsymbol{a}\|_{1}}$$

We don't know the exact distribution of this random variable.

However, we have a *one-sided error*: our estimate can never be lower than the true value. This means that $\hat{a}_i - a_i \ge 0$.

Markov's inequality says that if *X* is a nonnegative random variable, then

$$\Pr[X \geq c] \leq \frac{E[X]}{c}.$$

$$\Pr\left[\hat{\boldsymbol{a}}_{i} - \boldsymbol{a}_{i}\right] > \varepsilon \|\boldsymbol{a}\|_{1}$$

$$\leq \frac{E\left[\hat{\boldsymbol{a}}_{i} - \boldsymbol{a}_{i}\right]}{\varepsilon \|\boldsymbol{a}\|_{1}}$$

$$\Pr\left[\hat{\boldsymbol{a}}_{i} - \boldsymbol{a}_{i} > \varepsilon \|\boldsymbol{a}\|_{1}\right]$$

$$\leq \frac{\operatorname{E}\left[\hat{\boldsymbol{a}}_{i} - \boldsymbol{a}_{i}\right]}{\varepsilon \|\boldsymbol{a}\|_{1}}$$

$$\Pr\left[\hat{\boldsymbol{a}}_{i} - \boldsymbol{a}_{i} > \varepsilon \|\boldsymbol{a}\|_{1}\right]$$

$$\leq \frac{\operatorname{E}\left[\hat{\boldsymbol{a}}_{i} - \boldsymbol{a}_{i}\right]}{\varepsilon \|\boldsymbol{a}\|_{1}}$$

$$E[\hat{\boldsymbol{a}}_i - \boldsymbol{a}_i] \leq \frac{\|\boldsymbol{a}\|_1}{w}$$

$$\Pr\left[\hat{\boldsymbol{a}}_{i}-\boldsymbol{a}_{i} > \varepsilon \|\boldsymbol{a}\|_{1}\right]$$

$$\leq \frac{\mathbb{E}\left[\hat{\boldsymbol{a}}_{i}-\boldsymbol{a}_{i}\right]}{\varepsilon\left\|\boldsymbol{a}\right\|_{1}}$$

$$\leq \frac{\|\boldsymbol{a}\|_1}{w} \cdot \frac{1}{\varepsilon \|\boldsymbol{a}\|_1}$$

$$E[\hat{\boldsymbol{a}}_i - \boldsymbol{a}_i] \leq \frac{\|\boldsymbol{a}\|_1}{w}$$

$$\Pr\left[\hat{\boldsymbol{a}}_{i}-\boldsymbol{a}_{i} > \varepsilon \|\boldsymbol{a}\|_{1}\right]$$

$$\leq \frac{\mathbb{E}\left[\hat{\boldsymbol{a}}_{i} - \boldsymbol{a}_{i}\right]}{\varepsilon \|\boldsymbol{a}\|_{1}}$$

$$\leq \frac{\|\boldsymbol{a}\|_1}{w} \cdot \frac{1}{\varepsilon \|\boldsymbol{a}\|_1}$$

$$\Pr\left[\hat{\boldsymbol{a}}_{i}-\boldsymbol{a}_{i} > \varepsilon \|\boldsymbol{a}\|_{1}\right]$$

$$\leq \frac{\mathbb{E}\left[\hat{\boldsymbol{a}}_{i} - \boldsymbol{a}_{i}\right]}{\varepsilon \|\boldsymbol{a}\|_{1}}$$

$$\leq \frac{\|\boldsymbol{a}\|_1}{w} \cdot \frac{1}{\varepsilon \|\boldsymbol{a}\|_1}$$

$$= \frac{1}{\varepsilon w}$$

With probability at least
$$1 - \delta$$
, $|\hat{a} - a| \le \varepsilon \cdot size(input)$

for some measure of input size.

$$\Pr[\hat{\boldsymbol{a}}_i - \boldsymbol{a}_i > \varepsilon \|\boldsymbol{a}\|_1] \leq \frac{1}{\varepsilon w}$$

Probably
Approximately

Approximately Correct

With probability at least
$$1 - \delta$$
, $|\hat{a} - a| \le \varepsilon \cdot size(input)$

for some measure of input size.

$$\Pr[\hat{\boldsymbol{a}}_i - \boldsymbol{a}_i > \varepsilon \|\boldsymbol{a}\|_1] \leq \frac{1}{\varepsilon w}$$

Initial Idea:

Pick
$$w = \varepsilon^{-1} \cdot \delta^{-1}$$
. Then

$$\Pr[\hat{\boldsymbol{a}}_i - \boldsymbol{a}_i > \varepsilon \|\boldsymbol{a}\|_1] \leq \delta$$

Probably

Approximately Correct

With probability at least
$$1 - \delta$$
, $|\hat{a} - a| \le \varepsilon \cdot size(input)$

for some measure of input size.

- Probably - Approximately

Correct.

$$\Pr[\hat{\boldsymbol{a}}_i - \boldsymbol{a}_i > \varepsilon \|\boldsymbol{a}\|_1] \leq \frac{1}{\varepsilon w}$$

Initial Idea:

Pick $w = \varepsilon^{-1} \cdot \delta^{-1}$. Then

$$\Pr[\hat{\boldsymbol{a}}_i - \boldsymbol{a}_i > \varepsilon \|\boldsymbol{a}\|_1] \leq \delta$$

Suppose we're counting 1,000 distinct items.

With probability at least
$$1 - \delta$$
, $|\hat{a} - a| \le \varepsilon \cdot size(input)$

for some measure of input size.

- Probably - Approximately

Correct

$$\Pr[\hat{\boldsymbol{a}}_i - \boldsymbol{a}_i > \varepsilon \|\boldsymbol{a}\|_1] \leq \frac{1}{\varepsilon w}$$

Initial Idea:

Pick $w = \varepsilon^{-1} \cdot \delta^{-1}$. Then

$$\Pr[\hat{\boldsymbol{a}}_i - \boldsymbol{a}_i > \varepsilon \|\boldsymbol{a}\|_1] \leq \delta$$

Suppose we're counting 1,000 distinct items.

If we want our estimate to be within $\varepsilon \| \boldsymbol{a} \|_1$ of the true value with 99.9% probability, how much memory do we need?

With probability at least
$$1 - \delta$$
, $|\hat{a} - a| \le \varepsilon \cdot size(input)$

for some measure of input size.

- Probably - Approximately

Correct

$$\Pr[\hat{\boldsymbol{a}}_i - \boldsymbol{a}_i > \varepsilon \|\boldsymbol{a}\|_1] \leq \frac{1}{\varepsilon w}$$

Initial Idea:

Pick $w = \varepsilon^{-1} \cdot \delta^{-1}$. Then

$$\Pr[\hat{\boldsymbol{a}}_i - \boldsymbol{a}_i > \varepsilon \|\boldsymbol{a}\|_1] \leq \delta$$

Suppose we're counting 1,000 distinct items.

If we want our estimate to be within $\varepsilon \| \boldsymbol{a} \|_1$ of the true value with 99.9% probability, how much memory do we need?

Answer: $1,000 \cdot \varepsilon^{-1}$.

With probability at least
$$1 - \delta$$
, $|\hat{a} - a| \le \varepsilon \cdot size(input)$

for some measure of input size.

$$\Pr[\hat{\boldsymbol{a}}_i - \boldsymbol{a}_i > \varepsilon \|\boldsymbol{a}\|_1] \leq \frac{1}{\varepsilon w}$$

Initial Idea:

Pick $w = \varepsilon^{-1} \cdot \delta^{-1}$. Then

$$\Pr[\hat{\boldsymbol{a}}_i - \boldsymbol{a}_i > \varepsilon \|\boldsymbol{a}\|_1] \leq \delta$$

Suppose we're counting 1,000 distinct items.

If we want our estimate to be within $\varepsilon \| \boldsymbol{a} \|_1$ of the true value with 99.9% probability, how much memory do we need?

Answer: $1,000 \cdot \varepsilon^{-1}$.

Can we do better?

How to Build an Estimator

- 1. Design a simple data structure that, intuitively, gives you a good estimate.
- 2. Use a *sum of indicator variables* and *linearity of expectation* to prove that, on expectation, the data structure is pretty close to correct.
- 3. Use a *concentration inequality* to show that the data structure's output is close to its expectation.
- 4. Run multiple copies of the data structure in parallel to amplify the success probability.

How to Build an Estimator

- 1. Design a simple data structure that, intuitively, gives you a good estimate.
- 2. Use a *sum of indicator variables* and *linearity of expectation* to prove that, on expectation, the data structure is pretty close to correct.
- 3. Use a *concentration inequality* to show that the data structure's output is close to its expectation.
- 4. Run multiple copies of the data structure in parallel to amplify the success probability.

With probability at least
$$1 - \delta$$
, $|\hat{a} - a| \le \varepsilon \cdot size(input)$

for some measure of input size.

$$\Pr[\hat{\boldsymbol{a}}_i - \boldsymbol{a}_i > \varepsilon \|\boldsymbol{a}\|_1] \leq \frac{1}{\varepsilon w}$$

Revised Idea: Pick

$$w = e \cdot \varepsilon^{-1}$$
. Then

$$\Pr[\hat{\boldsymbol{a}}_i - \boldsymbol{a}_i > \varepsilon \|\boldsymbol{a}\|_1] < e^{-1}$$

We could choose $w = k \cdot \varepsilon^{-1}$ for any constant k to get a failure probability of at most k^{-1} . The choice of e is (mostly) arbitrary.

With probability at least
$$1 - \delta$$
, $|\hat{a} - a| \le \varepsilon \cdot size(input)$

for some measure of input size.

$$\Pr[\hat{\boldsymbol{a}}_i - \boldsymbol{a}_i > \varepsilon \|\boldsymbol{a}\|_1] \leq \frac{1}{\varepsilon w}$$

Revised Idea: Pick

$$w = e \cdot \varepsilon^{-1}$$
. Then

$$\Pr[\hat{\boldsymbol{a}}_i - \boldsymbol{a}_i > \varepsilon \|\boldsymbol{a}\|_1] < e^{-1}$$

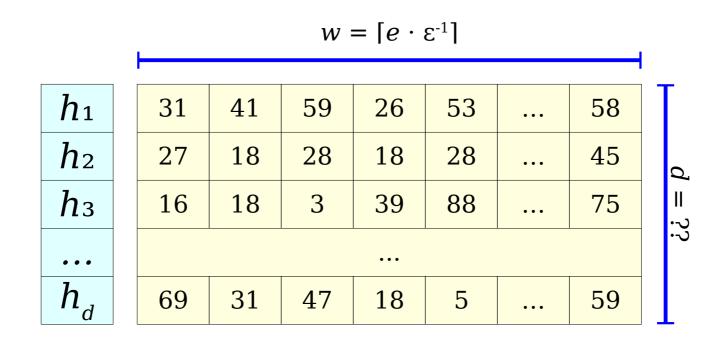
Probably

Approximately Correct

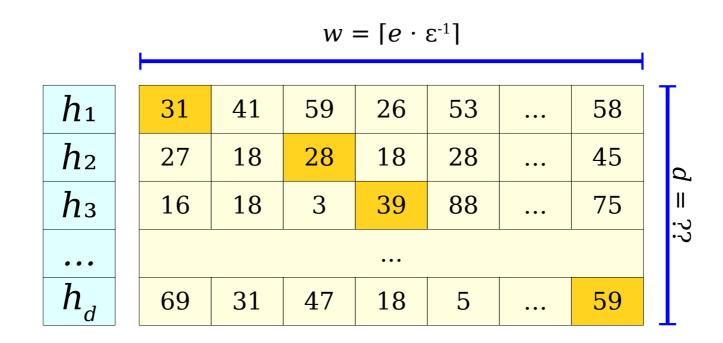
This simple data structure, by itself, is likely to be wrong.

What happens if we run a bunch of copies of this approach in parallel?

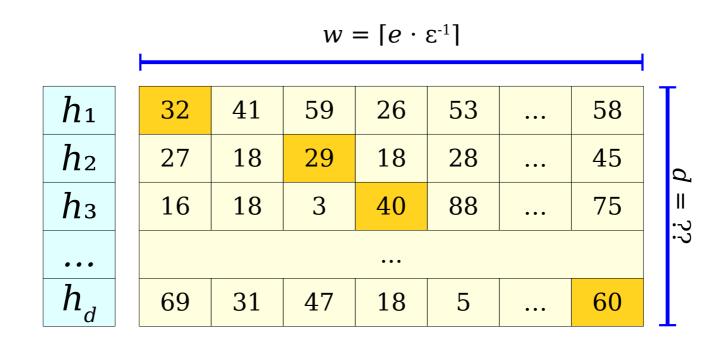
- Let's run d copies of our data structure in parallel with one another.
- Each row has its hash function sampled uniformly at random from our hash family.
- Each time we *increment* an item, we perform the corresponding *increment* operation on each row.



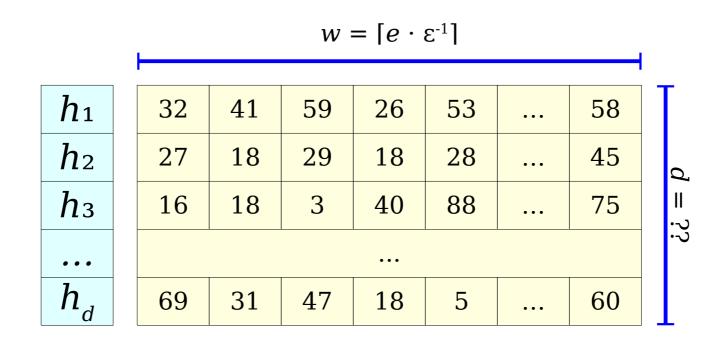
- Let's run *d* copies of our data structure in parallel with one another.
- Each row has its hash function sampled uniformly at random from our hash family.
- Each time we *increment* an item, we perform the corresponding *increment* operation on each row.



- Let's run d copies of our data structure in parallel with one another.
- Each row has its hash function sampled uniformly at random from our hash family.
- Each time we *increment* an item, we perform the corresponding *increment* operation on each row.



- Let's run d copies of our data structure in parallel with one another.
- Each row has its hash function sampled uniformly at random from our hash family.
- Each time we *increment* an item, we perform the corresponding *increment* operation on each row.



- Imagine we call estimate(x) on each of our estimators and get back these estimates.
- We need to give back a single number.

• *Question:* How should we aggregate these numbers into a single estimate?

Formulate a hypothesis, but don't post anything in chat just yet.

Estimator 1:

137

Estimator 2:

271

Estimator 3:

166

Estimator 4:

103

Estimator 5:

- Imagine we call estimate(x) on each of our estimators and get back these estimates.
- We need to give back a single number.

• *Question:* How should we aggregate these numbers into a single estimate?

Now, private chat me your best guess. Not sure? Just answer "??".

Estimator 1:

137

Estimator 2:

271

Estimator 3:

166

Estimator 4:

103

Estimator 5:

- Imagine we call estimate(x) on each of our estimators and get back these estimates.
- We need to give back a single number.
- *Question:* How should we aggregate these numbers into a single estimate?

Estimator 1:

137

Estimator 2:

271

Estimator 3:

166

Estimator 4:

103

Estimator 5:

- Imagine we call estimate(x) on each of our estimators and get back these estimates.
- We need to give back a single number.
- *Question:* How should we aggregate these numbers into a single estimate?

Estimator 1: 137

Estimator 2: 271

Estimator 3: 166

Estimator 4: 103

Estimator 5: 2.61

- Imagine we call estimate(x) on each of our estimators and get back these estimates.
- We need to give back a single number.
- *Question:* How should we aggrinto a single estimate?

Intuition: The smallest estimate returned has the least "noise," and that's the best guess for the frequency.

Estimator 1: **137**

Estimator 2: 271

Estimator 3: 166

Estimator 4: 103

Estimator 5: 2.61

$$\Pr[\min\{\hat{a}_{ij}\} - a_i > \epsilon ||a||_1]$$

$$\Pr[\min \{ \hat{a}_{ij} \} - a_i > \varepsilon ||a||_1]$$

$$\Pr[\min \{ \hat{a}_{ij} \} - a_i > \varepsilon ||a||_1]$$

The only way the minimum estimate is inaccurate is if *every* estimate is inaccurate.

Let $\hat{\boldsymbol{a}}_{ij}$ be the estimate from the jth copy of the data structure.

$$\Pr\left[\min\left\{|\hat{\boldsymbol{a}}_{ij}|\right\} - \boldsymbol{a}_i > \varepsilon ||\boldsymbol{a}||_1\right]$$

=
$$\Pr\left[\bigwedge_{j=1}^{d} \left(\hat{\boldsymbol{a}}_{ij} - \boldsymbol{a}_{i} > \epsilon \|\boldsymbol{a}\|_{1}\right)\right]$$

The only way the minimum estimate is inaccurate is if every estimate is inaccurate.

Let $\hat{\boldsymbol{a}}_{ij}$ be the estimate from the jth copy of the data structure.

$$\Pr\left[\min\left\{|\hat{\boldsymbol{a}}_{ij}|\right\} - \boldsymbol{a}_i > \varepsilon ||\boldsymbol{a}||_1\right]$$

=
$$\Pr\left[\bigwedge_{j=1}^{d} \left(\hat{\boldsymbol{a}}_{ij} - \boldsymbol{a}_{i} > \varepsilon \|\boldsymbol{a}\|_{1}\right)\right]$$

$$\Pr\left[\min\left\{\hat{\boldsymbol{a}}_{ij}\right\} - \boldsymbol{a}_{i} > \varepsilon \|\boldsymbol{a}\|_{1}\right]$$

$$= \Pr\left[\bigwedge_{j=1}^{d} \left(\hat{\boldsymbol{a}}_{ij} - \boldsymbol{a}_{i} > \varepsilon \|\boldsymbol{a}\|_{1}\right)\right]$$

Each copy of the data structure is independent of the others.

Let \hat{a}_{ij} be the estimate from the jth copy of the data structure.

$$\Pr\left[\min\left\{|\hat{\boldsymbol{a}}_{ij}|\right\} - \boldsymbol{a}_i > \varepsilon ||\boldsymbol{a}||_1\right]$$

=
$$\Pr\left[\bigwedge_{j=1}^{d} \left(\hat{\boldsymbol{a}}_{ij} - \boldsymbol{a}_{i} > \varepsilon \|\boldsymbol{a}\|_{1}\right)\right]$$

$$= \prod_{i=1}^{d} \Pr\left[\hat{\boldsymbol{a}}_{ij} - \boldsymbol{a}_{i} > \varepsilon \|\boldsymbol{a}\|_{1}\right]$$

Each copy of the data structure is independent of the others.

Let \hat{a}_{ij} be the estimate from the jth copy of the data structure.

$$\Pr[\min \{ \hat{a}_{ij} \} - a_i > \varepsilon ||a||_1]$$

$$= \Pr\left[\bigwedge_{j=1}^{d} \left(\hat{\boldsymbol{a}}_{ij} - \boldsymbol{a}_{i} > \varepsilon \|\boldsymbol{a}\|_{1}\right)\right]$$

$$= \prod_{i=1}^{d} \Pr\left[\hat{\boldsymbol{a}}_{ij} - \boldsymbol{a}_{i} > \varepsilon \|\boldsymbol{a}\|_{1}\right]$$

$$\Pr[\min \{ \hat{a}_{ij} \} - a_i > \varepsilon ||a||_1]$$

$$= \Pr\left[\bigwedge_{j=1}^{d} \left(\hat{\boldsymbol{a}}_{ij} - \boldsymbol{a}_{i} > \varepsilon \|\boldsymbol{a}\|_{1}\right)\right]$$

$$= \prod_{i=1}^{d} \Pr \left[\hat{\boldsymbol{a}}_{ij} - \boldsymbol{a}_{i} > \varepsilon \|\boldsymbol{a}\|_{1} \right]$$

$$\Pr[\hat{\boldsymbol{a}}_i - \boldsymbol{a}_i \geq \varepsilon \|\boldsymbol{a}\|_1] \leq e^{-1}$$

$$\Pr[\min \{ \hat{a}_{ij} \} - a_i > \epsilon ||a||_1]$$

$$= \Pr\left[\bigwedge_{j=1}^{d} \left(\hat{\boldsymbol{a}}_{ij} - \boldsymbol{a}_{i} > \varepsilon \|\boldsymbol{a}\|_{1}\right)\right]$$

$$= \prod_{i=1}^{d} \Pr\left[\hat{\boldsymbol{a}}_{ij} - \boldsymbol{a}_{i} > \varepsilon \|\boldsymbol{a}\|_{1}\right]$$

$$\leq \prod_{j=1}^{a} e^{-1}$$

$$\Pr[\hat{\boldsymbol{a}}_i - \boldsymbol{a}_i \geq \varepsilon \|\boldsymbol{a}\|_1] \leq e^{-1}$$

$$\Pr[\min\{\hat{a}_{ij}\} - a_i > \epsilon ||a||_1]$$

$$= \Pr\left[\bigwedge_{j=1}^{d} \left(\hat{\boldsymbol{a}}_{ij} - \boldsymbol{a}_{i} > \varepsilon \|\boldsymbol{a}\|_{1}\right)\right]$$

$$= \prod_{i=1}^{d} \Pr\left[\hat{\boldsymbol{a}}_{ij} - \boldsymbol{a}_{i} > \varepsilon \|\boldsymbol{a}\|_{1}\right]$$

$$\leq \prod_{i=1}^{a} e^{-1}$$

$$= e^{-d}$$

With probability at least
$$1 - \delta$$
, $|\hat{a} - a| \le \varepsilon \cdot size(input)$

for some measure of input size.

Probably
Approximately
Correct

$$\Pr[\min\{\hat{\boldsymbol{a}}_{ij}\}-\boldsymbol{a}_i > \varepsilon \|\boldsymbol{a}\|_1] \leq e^{-d}$$

Idea: Choose $d = -\ln \delta$. (Equivalently: $d = \ln \delta^{-1}$.) Then

$$\Pr[\min\{\hat{\boldsymbol{a}}_{ij}\}-\boldsymbol{a}_i > \varepsilon \|\boldsymbol{a}\|_1] \leq \delta$$



 h_1 h_2 h_3 h_d

Sampled uniformly and independently from a 2-independent family of hash functions

h_1	31	41	59	26	53
h_2	27	18	28	18	28
hз	16	18	3	39	88
• • •				•••	
h_d	69	31	47	18	5

```
increment(x):
    for i = 1 ... d:
        count[i][hi(x)]++
```

h_1	31	41	59	26	53	•••	58
h_2	27	18	28	18	28	•••	45
h_3	16	18	3	39	88	•••	75
•••	•••						
h_d	69	31	47	18	5	•••	59

```
increment(x):
    for i = 1 ... d:
        count[i][hi(x)]++
```

h_1	32	41	59	26	53	•••	58
h_2	27	18	28	19	28	•••	45
h_3	16	19	3	39	88	•••	75
•••	•••						
h_d	69	31	47	18	5	•••	60

```
increment(x):
   for i = 1 ... d:
      count[i][hi(x)]++
```

h_1	32	41	59	26	53	•••
h_2	27	18	28	19	28	•••
hз	16	19	3	39	88	•••
•••				•••		
h_d	69	31	47	18	5	•••

```
increment(x):
   for i = 1 ... d:
      count[i][hi(x)]++
```

h_1	3
h_2	2
hз	1
• • •	
h_d	6

```
41
          59
                26
                      53
                                   58
    18
          28
                19
                      28
                                  45
    19
           3
                      88
                                   75
                39
9
    31
          47
                18
                       5
                                  60
```

```
increment(x):
    for i = 1 ... d:
        count[i][hi(x)]++
```

```
estimate(x):
    result = ∞
    for i = 1 ... d:
       result = min(result, count[i][hi(x)])
    return result
```

```
h_1
        32
              41
                                 53
                    59
                          26
                                             58
h_2
        27
              18
                    28
                          19
                                 28
                                             45
h_3
        16
                     3
                                             75
              19
                          39
                                 88
h_d
        69
              31
                    47
                          18
                                 5
                                             60
```

```
increment(x):
    for i = 1 ... d:
        count[i][hi(x)]++
```

```
estimate(x):
    result = ∞
    for i = 1 ... d:
       result = min(result, count[i][hi(x)])
    return result
```

- Update and query times are $\Theta(d)$, which is $\Theta(\log \delta^{-1})$.
- Space usage: $\Theta(\varepsilon^{-1} \cdot \log \delta^{-1})$ counters.
 - This is a *major* improvement over our earlier approach that used $\Theta(\epsilon^{-1} \cdot \delta^{-1})$ counters.
 - This can be *significantly* better than just storing a raw frequency count!
- Provides an estimate to within $\varepsilon \| \boldsymbol{a} \|_1$ with probability at least 1δ .

Major Ideas From Today

- **2-independent hash families** are useful when we want to keep collisions low.
- A "good" approximation of some quantity should have tunable *confidence* and *accuracy* parameters.
- **Sums of indicator variables** are useful for deriving expected values of estimators.
- Concentration inequalities like Markov's inequality are useful for showing estimators don't stay too much from their expected values.
- Good estimators can be built from multiple parallel copies of weaker estimators.

Next Time

Count Sketches

 An alternative frequency estimator with different time/space bounds.

• Cardinality Estimation

• Estimating how many different items you've seen in a data stream.