

CAUSAL INFERENCE FOR ASSET PRICING

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- Traditional methods: “everything is connected,” Euler equation tests, factor models, ...

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How to interpret estimates? Implicit assumptions on spillovers?

- Quantitative demand systems
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Which results are robust outside of these models and which are specific to these structures?

How do those approaches account for substitution and spillovers across assets?

- Traditional methods: “everything is connected,” Euler equation tests, factor models, ...

WHY IS IT IMPORTANT?

- Learning from asset quantity data
- Learning from natural experiments
 - Counterpart to growth of micro-empirical methods in macro (e.g. Nakamura Steinsson 2018, Sraer Thesmar 2022)
- Classic models far from the data
- Many important questions are about quantities:
 - Quantitative easing policies (e.g. Haddad Moreira Muir 2025)
 - International capital flows, China and US Treasuries (e.g. Jansen Li Schmid 2025)
 - Rise of passive investing (Haddad, Huebner, Loualiche)

OUR FRAMEWORK

- **Simple portable assumption:** *homogeneous substitution conditional on observables*
 - Diagnostics, empirical design, ...
- **Flexible but parsimonious:** captures the forces of many demand structures, particularly specific to finance
 - **Key missing piece of existing models:** elasticity to price of factors/characteristics = substitution depends on characteristics
- **Easy estimation:** set of IV/diff-in-diff regression
 - “Separable:” map from different types of natural experiment to different counterfactual
 - Precisely define what is a valid instrument
 - Lots of work on finding instruments, not the focus here

The Problem

QUESTION

How do an investor's portfolio decisions respond to prices?

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Elasticity matrix: sensitivity of demand to prices

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Elasticity matrix: sensitivity of demand to prices

- Defined in any theory
 - mean-variance: $D = \frac{1}{\gamma} \Sigma^{-1} (M - P)$, $\mathcal{E} = -\frac{1}{\gamma} \Sigma^{-1}$
- Could be log, levels, shares, changes or not, ...
- Flipside: price impact \mathcal{E}^{-1} , how do shifts in demand affect prices?

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- Could be log, levels, shares, changes or not, ...
- Flipside: price impact \mathcal{E}^{-1} , how do shifts in demand affect prices?
- *How does CalPERS adjust its position in 10-year corporate bonds of Ford and GM when their spread changes?*
- *How does AQR move across value and momentum based on their risk premia?*

⇒ Answer to such questions about **different parts** of \mathcal{E}

AN EXAMPLE: CALPERS AND CORPORATE BONDS

- Prices have moved **and no other news**. CalPERS adjusts its bond portfolio:

	Price change	Change in position
1. 10-yr Ford	+ 5%	sell 200
2. 10-yr GM	+ 2%	sell 100
3. 5-yr First Solar	- 1%	buy 100
⋮	⋮	⋮

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$$\Delta D_1 = \underbrace{\mathcal{E}_{11}\Delta P_1}_{\text{became more expensive}} + \underbrace{\mathcal{E}_{12}\Delta P_2}_{\text{substitutes from GM}} + \underbrace{\sum_{k \geq 3} \mathcal{E}_{1k}\Delta P_k}_{\text{substitutes from First Solar, ...}}$$

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→ **Stuck without additional assumptions**

TWO PATHS

- **Causal inference:** impose elementary restriction keeping as much flexibility on mechanism as possible while letting the data speak
- **Structural approach:** choose a microfoundation and estimate the corresponding model

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TWO PATHS

- **Causal inference:** impose elementary restriction keeping as much flexibility on mechanism as possible while letting the data speak
 - Canonical assumption (SUTVA): when I give you medication, it affects your health but not the control group's health
 - Does not work here: demand for each asset only depends on its own price $D_i(P_i) \Rightarrow$ diagonal \mathcal{E}
- **Structural approach:** choose a microfoundation and estimate the corresponding model

LEARNING FROM STANDARD FINANCE MODELS

- Assume returns follow a factor structure: exposures β_i and idiosyncratic risk $\sigma_{\epsilon,i}^2$
- *Force 1: factor management*, if expected return only depends on exposures β
 - only buy portfolios replicating the factors (mutual fund theorem)
 - choose exposure to the factors based on the expected returns of those factors
- *Force 2: “arbitrage”*: if expected returns deviate from factor pricing
 - buy more of cheap (= high alpha) assets, less of expensive ones
- A demand formula from Koijen Yogo 2019:
 - Factor model for returns: $R = \beta F + \epsilon$
 - Variance covariance matrix: $\Sigma = \beta\beta' + \sigma_\epsilon^2 \mathbf{I}$
 - Demand elasticity follows factor structure (Sherman-Morrison):

$$\mathcal{E} = -\frac{1}{\gamma} \Sigma^{-1} = \underbrace{-\frac{1}{\gamma \sigma_\epsilon^2} \mathbf{I}}_{\text{diagonal}} + \underbrace{c \beta \beta'}_{\text{substitution matrix}} \quad \underbrace{c}_{\text{scalar}} = \frac{1}{\gamma \sigma_\epsilon^2} \frac{1}{\sigma_\epsilon^2 + \beta' \beta}$$

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- **What next?**
 - KY 2019: add additional restrictions, get to logit or similar forms
 - This paper: generalize, keeping only the basic structure of diagonal + substitution driven by observables

A PRACTICAL EXAMPLE: SIMULATION

- **Simulation:** Start from factor model demand, increase equally the supply of all assets \rightarrow equilibrium is that price of high beta assets drops more

- Mean-variance model: $\log D_t = \gamma^{-1} (\beta\beta' + \sigma_\epsilon^2 \mathbf{I})^{-1} (\mu_t - \log P_t)$
- Effect of supply shock δ :

$$\Delta \log P_t = \gamma (\beta\beta' + \sigma_\epsilon^2 \mathbf{I}) \delta$$

- Simple one factor model (CAPM)

$$(\Delta \log P_t)_i = \gamma \sigma_\epsilon^2 \delta_i + \gamma \beta_i \sum_k \beta_k \delta_k$$

- Uniform shock $\delta = \delta \mathbf{1}$ has heterogeneous effects:

$$\frac{(\Delta \log P_t)_i}{\delta} = \gamma \sigma_\epsilon^2 + \gamma \beta_i \sum_k \beta_k$$

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A PRACTICAL EXAMPLE: ESTIMATION

Can we back out from prices the structure of the supply shock?

- KY 2019: demand and predict what the change in supply was: *erroneously* predict larger increase of supply for high beta assets
 - Adding a macro elasticity (Gabaix Koijen 2025) does not fix it
- Logit estimation (as in KY 2019)

$$\log D_{i,t} = b_0 + \hat{\mathcal{E}} \log P_{i,t} + \theta_t \beta_i + e_{it}$$

- Econometrician recovers demand shocks that are linear in the price change.
- Misses the differential preferences for assets with different β

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 - Adding a macro elasticity (Gabaix Koijen 2025) does not fix it
- Logit estimation with Macro elasticity (as in Gabaix-Koijen)
 - Separate logit (cross-section) from aggregate (time-series)
 - Construct aggregate and idiosyncratic variables:

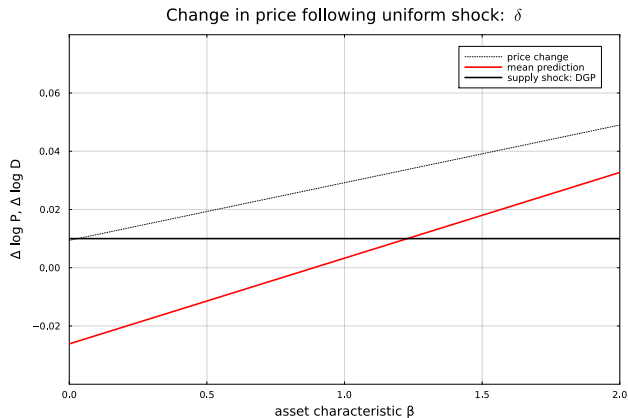
$$D_{t,agg} = \sum_i D_{i,t}, \quad D_{i,t}^{idio} = D_{i,t} - D_{t,agg}.$$

- Fixes the mean of the estimated supply shock
- Still misses the differential preferences for assets with different β

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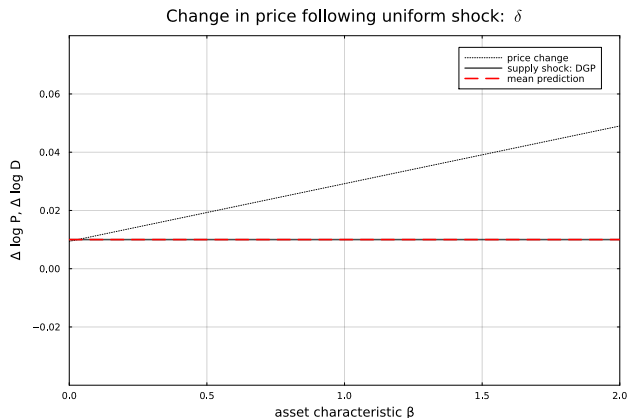
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- KY 2019: demand and predict what the change in supply was: *erroneously* predict larger increase of supply for high beta assets
 - Adding a macro elasticity (Gabaix Koijen 2025) does not fix it
- Our methodology: account for factor substitution:



Framework

HOMOGENEOUS SUBSTITUTION CONDITIONAL ON OBSERVABLES

A simple assumption:

- **Homogeneous substitution conditional on observables**
 - CalPERS substitutes across bonds based on their observables (e.g. duration, greenness) only

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■ Compare bonds with same observables: Ford vs. GM

- E.g.: CalPERS adjusts Ford and GM equally in response to price of First Solar $\varepsilon_{13} = \varepsilon_{23}$

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$$\text{Diff-in-diff:} \quad \Delta D_1 - \Delta D_2 = \hat{\mathcal{E}}(\Delta P_1 - \Delta P_2) \text{ if same relative elasticity}$$

■ Compare bonds with same observables: Ford vs. GM

- E.g.: CalPERS adjusts Ford and GM equally in response to price of First Solar $\mathcal{E}_{13} = \mathcal{E}_{23}$

→ *comparing assets with same observables differences out substitution*

FORMAL SETUP

■ Homogeneous substitution conditional on observables X

$$\boxed{\mathcal{E}_{il} = \mathcal{E}_{jl} \text{ if } X_i = X_j} \quad \text{for all } i, j \in \mathcal{S}, \text{ and } l \neq i, j,$$

- If price of 3rd asset move, response of demand for 2 assets with same observables is the same
- Parametrize linearly: $\mathcal{E}_{il} = \mathcal{E}_{\text{cross}}(X_i, X_l) = X_i' \mathcal{E}_X X_l$

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■ Decomposition of demand elasticity:

$$\begin{aligned} \mathcal{E} &= \text{relative elasticity} + \text{substitution} \\ &= \text{diagonal matrix} + X \underbrace{\mathcal{E}_X}_{K \times K} X' \end{aligned}$$

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- Assume constant relative elasticity $\hat{\mathcal{E}}$ for simplicity, relax in the paper

QUESTIONS REVISITED

$$\begin{aligned}\mathcal{E} &= \text{relative elasticity} + \text{substitution} \\ &= \hat{\mathcal{E}}I + X\mathcal{E}_X X'\end{aligned}$$

Different questions are about different parts of \mathcal{E}

- *How does CalPERS adjust its position in 10-year corporate bonds of Ford and GM when their spread changes?*
- *How does AQR move across factors based on factor risk premia?*

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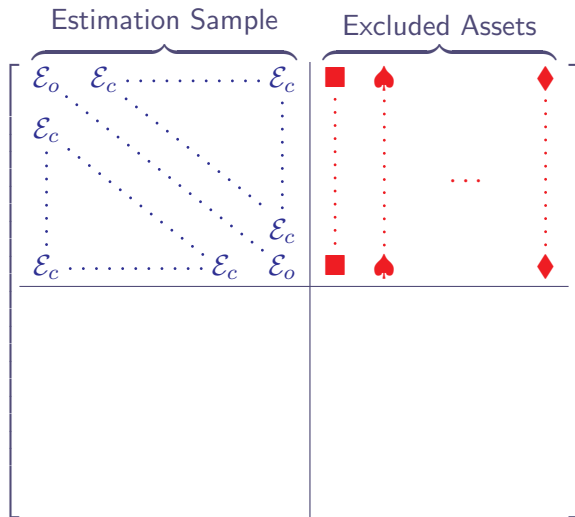
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Different questions are about different parts of \mathcal{E}

- *How does CalPERS adjust its position in 10-year corporate bonds of Ford and GM when their spread changes?*
 - Asset-specific behavior characterized by the relative elasticity $\hat{\mathcal{E}}$
- *How does AQR move across factors based on factor risk premia?*
 - Question about substitution characterized by \mathcal{E}_X

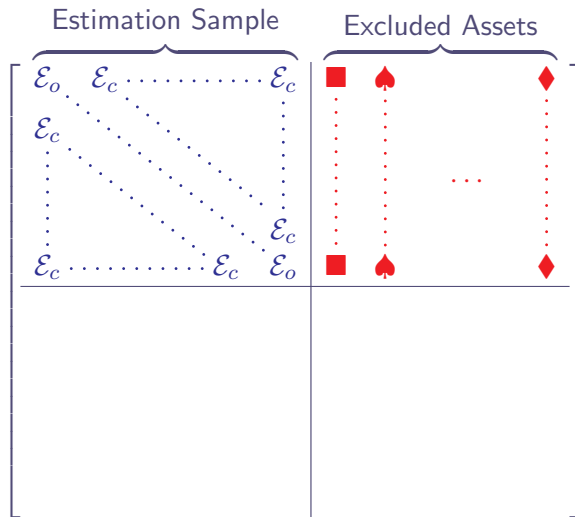
LOCAL EXPERIMENTS

- With few close assets: ignore observables & assume full homogeneity
 - Same own- and cross-price elasticity for every pair of assets in \mathcal{S}
 - Demand for all assets in \mathcal{S} responds in same way to price of 5-year First Solar bond (outside \mathcal{S})



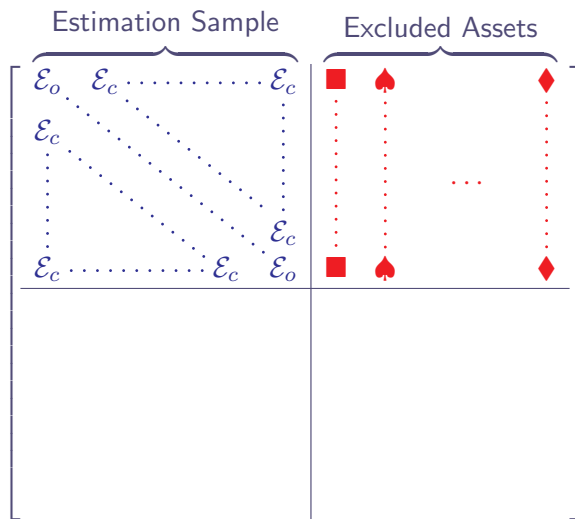
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- Diagnostic: balance between treated (high Z_i) and control (low Z_i) on covariance with broad factors



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- *Risk based motives:* care about portfolio-level factor exposure, so X_i are factor loadings or characteristics that proxy for them
 - Markowitz: $D = \frac{1}{\gamma} \Sigma^{-1} (\mu - P) \Rightarrow \mathcal{E} = -\frac{1}{\gamma} \Sigma^{-1}$
 - If Σ has factor structure: idio risk drives $\hat{\mathcal{E}}$, factor risk drives \mathcal{E}_X

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- *Risk based motives:* care about portfolio-level factor exposure, so X_i are factor loadings or characteristics that proxy for them
- *Non-risk motives:* X_i is asset weight in this objective

$$\max_D \quad D'(\mu - P) - \frac{\gamma}{2} D' \Sigma D - \frac{\kappa}{2} \left(D' X^{(1)} \right)^2$$

such that $D' X^{(2)} \leq \Theta$

- Binding constraints (leverage), regulatory score (capital ratio), or stakeholders pressure (greenness)

CROSS-SECTIONAL IDENTIFICATION

- **Data-Generating-Process:** Elasticity matrix \mathcal{E} + *homogeneous substitution conditional on observable X*

$$\Delta \mathbf{D} = \mathcal{E} \Delta \mathbf{P} + \epsilon$$

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- Demand shift ϵ correlated with prices: Ford is more expensive because the new F150 is amazing, change in CalPERS financial health, ...
- **Proposition 1** Under our assumption, and the *usual exclusion and relevance restrictions*, the IV estimator identifies the **relative elasticity** $\hat{\mathcal{E}} = \mathcal{E}_{ii} - \mathcal{E}_{ji}$ for $X_i = X_j$

$$\Delta D_i = \hat{\mathcal{E}} \Delta P_i + \theta' X_i + e_i$$

$$\Delta P_i = \lambda Z_i + \eta' X_i + u_i$$

with Z_i instrument for prices ($Z_i \perp \epsilon_i | X_i$)

- E.g.: Fed buys some bonds but not others

ABSORBING SUBSTITUTION

- Key step: coefficient on observables θ absorbs substitution from other assets

$$\begin{aligned}\Delta D_i &= \varepsilon_{ii} \Delta P_i + \sum_{j \neq i} X_i' \varepsilon_X X_j \Delta P_j + \epsilon_i \\ &= (\varepsilon_{ii} - X_i' \varepsilon_X X_i) \Delta P_i + \sum_j X_i' \varepsilon_X X_j \Delta P_j + \epsilon_i \\ &= \underbrace{(\varepsilon_{ii} - X_i' \varepsilon_X X_i)}_{\text{relative elasticity}} \Delta P_i + X_i' \underbrace{\sum_j \varepsilon_X X_j \Delta P_j}_{\text{constant across assets, absorbed in } \theta} + \epsilon_i\end{aligned}$$

- **Relative elasticity:** difference between own-price and cross-price elasticity for assets with same observables
 - How does the relative demand for Ford and GM respond to their relative price?
 - Useful to answer relative Qs and construct relative counterfactuals
 - In large cross-sections with substantial idiosyncratic risk \approx own-price elasticity
 - What GE theorists call the Morishima elasticity Gabaix Koijen 2025 the micro-elasticity

FAQ

- 1 “I exclude some assets from my sample because I don’t have data on them. Is this a problem?” **No**, as long as you assume the **same structure** applies to **excluded assets**. Conditional on X_i , the omitted asset affects others symmetrically, which differences out.

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- 2 *“What if I have a repeated cross-section?”* Control for observables X_i **interacted with time fixed effects**.
- 3 *“What do I estimate in a price impact regression of exogenous demand shocks on prices?”* Under **our assumption**, the identified coefficient is the **relative multiplier** $-1/\hat{\epsilon}$.
- 4 *“How about equilibrium spillovers?”* If the Fed buys a bond but not a similar one, the **price of the similar one also moves**. If LOOP holds, this may threaten instrument relevance (testable), but **not the exclusion restriction**.

FAQ

- 1 “I exclude some assets from my sample because I don’t have data on them. Is this a problem?” **No**, as long as you assume the **same structure** applies to **excluded assets**. Conditional on X_i , the omitted asset affects others symmetrically, which differences out.
- 2 “What if I have a repeated cross-section?” Control for observables X_i **interacted with time fixed effects**.
- 3 “What do I estimate in a price impact regression of exogenous demand shocks on prices?” Under **our assumption**, the identified coefficient is the **relative multiplier** $-1/\hat{\mathcal{E}}$.
- 4 “How about equilibrium spillovers?” If the Fed buys a bond but not a similar one, the **price of the similar one also moves**. If LOOP holds, this may threaten instrument relevance (testable), but **not the exclusion restriction**.
- 5 “Can I recover the own-price elasticity from my cross-sectional regression?” In general, **no** because the **own-price elasticity** combines **both** the relative elasticity and substitution. Cross-sectional regressions only identify part of \mathcal{E} .

SUBSTITUTION AND ITS ESTIMATION

Estimating **substitution** \mathcal{E}_X crucial for many questions:

- How does CalPERS adjust its portfolio when the price of all bonds drops?
- Will CalPERS maintain its green tilt if green bonds become very expensive relative to brown bonds?

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Proposition 2 Impossible to identify **substitution** with the cross-section alone

$$\Delta D_i = \hat{\mathcal{E}} \Delta P_i + X_i' \underbrace{\sum_j \mathcal{E}_X X_j \Delta P_j}_{\text{BOTH absorbed in } \theta} + \epsilon_i$$

- Coefficient on X_i measures both **substitution** and **shift in demand for observable**
 - Does CalPERS reduce its green tilt because of **expensive green bonds** or **weaker environmental priorities**?
 - This is a **missing coefficients** problem

DEMAND–PRICE DECOMPOSITION

Classic strategy: construct portfolios sorted on observables, and measure their price and demand (= portfolio tilt)

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- X_i (which is normalized) for example captures greenness

$$\Delta D_{agg} = \frac{1}{N} \sum_i \Delta D_i,$$

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- *Separate the response of demand to prices into three univariate components:*

Relative:

$$\Delta D_{idio,i} = \hat{\epsilon} \Delta P_{idio,i}$$

Meso:

$$\Delta D_X = \tilde{\epsilon}_{agg} \Delta P_{agg} + \tilde{\epsilon}_X \Delta P_X$$

Macro:

$$\Delta D_{agg} = \bar{\epsilon}_{agg} \Delta P_{agg} + \bar{\epsilon}_X \Delta P_X$$

ESTIMATING SUBSTITUTION WITH THE TIME SERIES

- **Proposition 3** Regressing portfolio tilts on portfolio prices with **time series instruments** identifies **substitution** \mathcal{E}_X

$$\begin{aligned}\Delta D_{X,t} &= \tilde{\mathcal{E}}_{agg} \Delta P_{agg,t} + \tilde{\mathcal{E}}_X \Delta P_{X,t} + \epsilon_{X,t} \\ \Delta D_{agg,t} &= \bar{\mathcal{E}}_{agg} \Delta P_{agg,t} + \bar{\mathcal{E}}_X \Delta P_{X,t} + \epsilon_{agg,t}\end{aligned}$$

- Effectively only K assets = portfolios
- E.g. Fed does more or less QE and operation twist over time

SUMMARY

Homogeneous substitution conditional on observables X :

$$\begin{aligned}\mathcal{E} &= \text{relative elasticity} + \text{substitution} \\ &= \hat{\mathcal{E}}I + X\mathcal{E}_X X'\end{aligned}$$

Consistent with many motives: risk, constraints, non-pecuniary preferences, irrational, ...

Identification:

- **Relative elasticity**: compare similar assets = cross-sectional IV controlling for X
- **Substitution**: demand for portfolios based on X = time-series portfolio level instruments

WHAT ABOUT LOGIT?

- Koijen Yogo (2019), Koijen Richmond Yogo (2024):
 - \exists factor models where volatility and expected returns depend non-linearly of prices which yield asset demand in the logit form
 - Logit has non-zero substitution and can be inferred from the cross-section alone

WHAT ABOUT LOGIT?

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 - \exists factor models where volatility and expected returns depend non-linearly of prices which yield asset demand in the logit form
 - Logit has non-zero substitution and can be inferred from the cross-section alone
- Logit satisfies our assumption, and its parameter can be robustly interpreted as **relative elasticity**
- Logit strongly restricts **substitution**: **an arbitrary factor model is not equivalent to logit**
 - Logit: when the price of any bond \uparrow , CalPERS replaces it proportionally to its existing portfolio
 - Factor model: CalPERS replaces it disproportionately with bonds loading on similar factors

WHAT ABOUT LOGIT?

- Koijen Yogo (2019), Koijen Richmond Yogo (2024):
 - \exists factor models where volatility and expected returns depend non-linearly of prices which yield asset demand in the logit form
 - Logit has non-zero substitution and can be inferred from the cross-section alone
- Logit demand is a very special case of factor demand:
 - Single factor elasticity: substitution based on factor loadings (characteristics β)

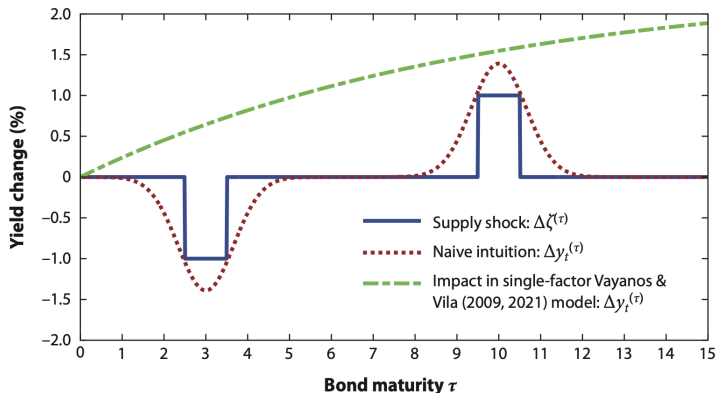
$$\gamma^{-1} \text{diag}(\sigma_{\epsilon}^{-2}) + c\beta\beta'$$

- Logit elasticity: substitution based on shares (ω):

$$\alpha \text{diag}(\omega) + \alpha\omega\omega'$$

GROUP-BASED SUBSTITUTION VS FACTOR MODELS

- Nested logit (Fang 2023, Koijen Yogo 2024): symmetric groups based on values of observables → can use the cross-section of groups to estimate substitution
 - Predict strong local effect and diffuse effect across all other groups
 - Sharply different from factor model with exposure depending on observable (see Cochrane 2008, Vayanos Vila 2021)



EXAMPLE: CORPORATE BOND RELATIVE MULTIPLIER

- U.S. investment-grade corporate bonds (following Chaudhary Fu Li, 2024)
- Steps to go through when conducting causal inference in asset pricing:
 - 1 choose a source of variation
 - 2 assess exogeneity
 - 3 assess assumptions A1 and A2 and select observables + units
 - 4 implement the regression analysis
- Step 1: flow-induced demand shock Z_{it} : fund flow in mutual funds \times portfolio composition (Coval Stafford 2007, Lou 2012)
- Step 2: assess exogeneity, i.e., $Z_{it} \perp \epsilon_{it} | X_{it}$
 - example threat to identification: another investor type (e.g., insurance companies) buying the same bonds held by mutual funds that receive a lot of inflows

STEP 3: DIAGNOSTIC FOR HOMOGENEOUS SUBSTITUTION – BALANCE ON COVARIANCES

Do treated & control bonds comove the same way with broad portfolios?

- 1 At each date t , form a long-short portfolio based on treatment status
- 2 Compute the β of the long-short return on broad indices in a window around t (here: 2y)
- 3 β different from zero \Rightarrow substitution likely not homogeneous

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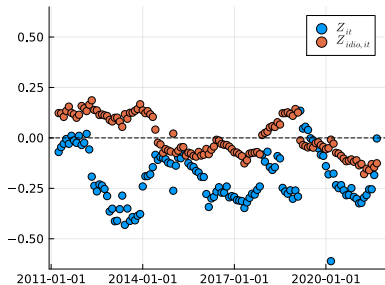
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 \rightarrow natural if investors choose their flows along dimensions like duration and credit risk
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 - $Z_{idio,it}$: residual of instrument regressed on a date fixed effect, **duration** \times date fixed effects and **credit rating** \times date fixed effects

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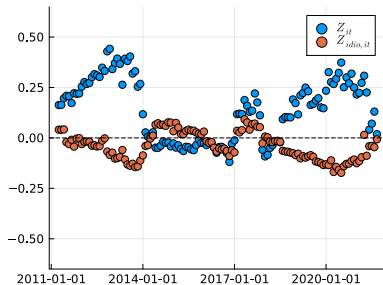
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 - $Z_{idio,it}$: residual of instrument regressed on a date fixed effect, **duration** \times date fixed effects and **credit rating** \times date fixed effects
 - Alternative unit to bond returns: yield changes ▶ A1 yield changes ▶ Multiplier yield changes
 - Similar diagnostic for constant relative elasticity: balance on idiosyncratic volatility
▶ A2 diagnostic

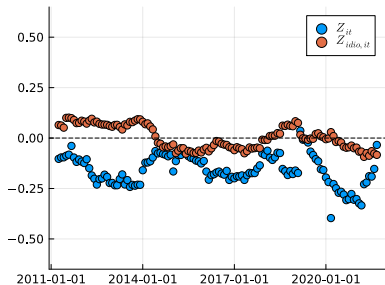
A. Corporate Bond Index



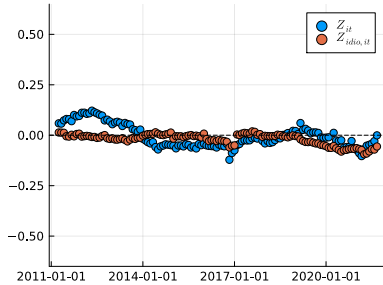
B. High—Low Credit Rating



C. Long—Short Term Bonds



D. Stock Index



STEP 4: IMPLEMENT THE REGRESSION

Relative multiplier $\widehat{\mathcal{M}} \approx 0$

	Return $\Delta P_{it}/P_{i,t-1}$				
	(1)	(2)	(3)	(4)	(5)
<i>Demand shock:</i>					
Z_{it}	1.541*	-0.254	0.019		
	(0.637)	(0.229)	(0.065)		
$Z_{idio,it}$				0.019	0.019
				(0.065)	(0.065)
Date Fixed Effects		Yes	Yes	Yes	Yes
Duration \times Date Fixed Effects			Yes	Yes	
Credit Rating \times Date Fixed Effects			Yes	Yes	
N	646,335	646,335	646,335	646,335	646,335
R^2	0.010	0.415	0.632	0.632	0.415

TAKEAWAYS

- To draw causal inference about demand elasticity, need:
 - A simple assumption: **homogeneous substitution conditional on observables**
 - CalPERS substitutes based on duration and greenness
 - (Standard) source of exogenous variation
 - Fed randomly buys more of some bonds than others, Fed surprisingly engages in QE
- **Relative elasticity** for similar assets: cross-sectional IV
 - Ford vs GM?
- **Substitution** = demand for portfolios: time-series IV
 - Green vs brown? Aggregate price?
- Standard structural models of demand rule out most factor-style substitution

WHY CAUSAL INFERENCE IN ASSET PRICING?

- Causal inference particularly valuable when:
 - existing theories are far from the data
 - it is challenging to understand all sources of variations simultaneously
- First step towards better economic theory

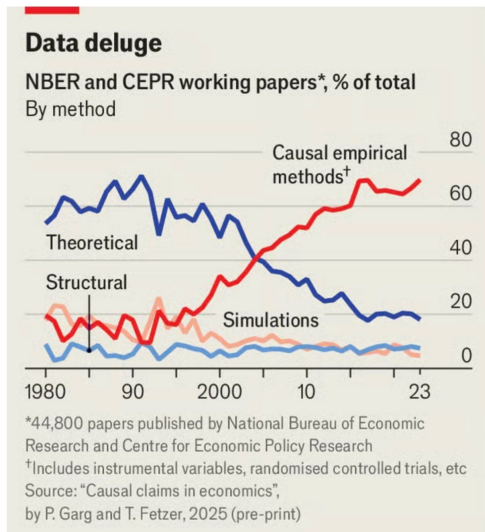
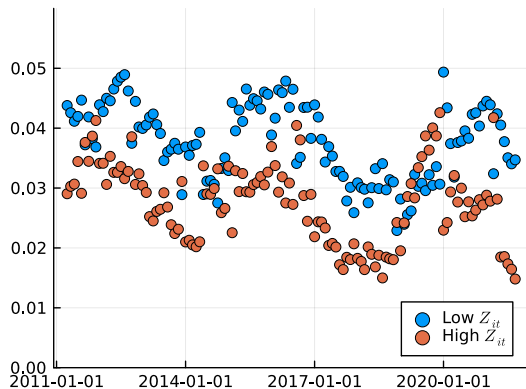


CHART: THE ECONOMIST

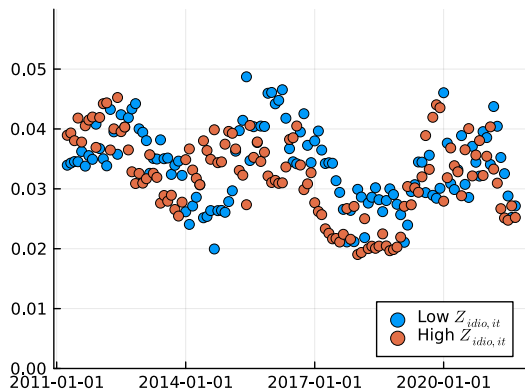
DIAGNOSTIC FOR CONSTANT RELATIVE ELASTICITY

■ Balance on idiosyncratic volatility

A. Idiosyncratic Volatility (Z_{it})

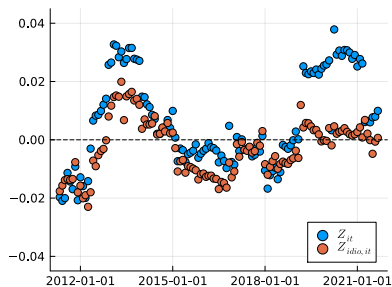


B. Idiosyncratic Volatility ($Z_{idio,it}$)

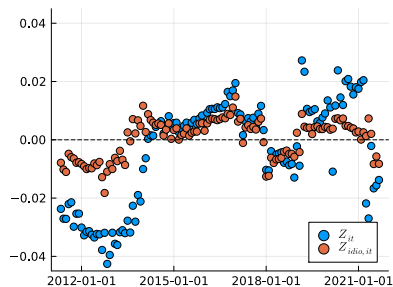


■ Average idiosyncratic volatility among treated versus control bonds

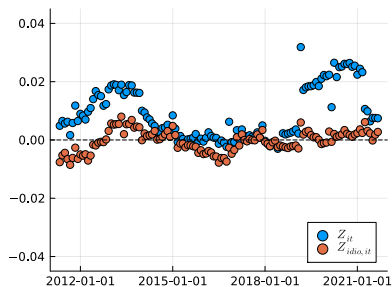
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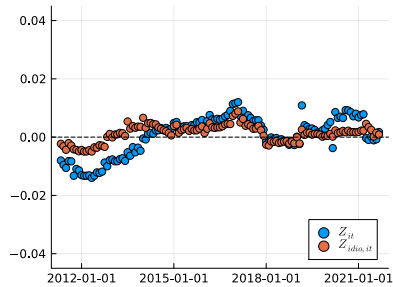
B. High—Low Credit Rating



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Relative multiplier $\widehat{\mathcal{M}} = -0.072$

	Yield change ΔY_{it}				
	(1)	(2)	(3)	(4)	(5)
<i>Demand shock:</i>					
Z_{it}	-0.384*	-0.104*	-0.072**		
	(0.166)	(0.047)	(0.027)		
$Z_{idio,it}$				-0.072**	-0.072**
				(0.027)	(0.027)
Date Fixed Effects		Yes	Yes	Yes	Yes
Duration \times Date Fixed Effects			Yes	Yes	
Credit Rating \times Date Fixed Effects			Yes	Yes	
N	630,255	630,255	630,255	630,255	630,255
R^2	0.004	0.071	0.089	0.089	0.070