

X_i is the number of seedlings observed in quadrat # i .

$X_1, \dots, X_n \stackrel{\text{iid}}{\sim} \text{Poisson}(\lambda)$

Goal: a confidence interval for λ based on the large-sample normal approx. to the sampling distribution of

$$\hat{\lambda}^{\text{MLE}} = \bar{X} = \frac{1}{n} \sum_{i=1}^n X_i$$

The normal approx. is,

$$\hat{\lambda}^{\text{MLE}} \sim \text{Normal}\left(\lambda, \underbrace{\frac{1}{J(\hat{\lambda}^{\text{MLE}})}}\right)$$

observed Fisher information
evaluated at $\hat{\lambda}^{\text{MLE}}$ (estimate
based on observed sample data)

$$\underline{J(\lambda^*)} = - \frac{d^2}{d\lambda^2} \log \left\{ \prod_{i=1}^n f_{X_i}(x_i | \lambda) \right\} \Big|_{\lambda=\lambda^*}$$

$$= - \frac{d^2}{d\lambda^2} \sum_{i=1}^n \log \left\{ \frac{e^{-\lambda} \lambda^{x_i}}{x_i!} \right\} \Big|_{\lambda=\lambda^*}$$

$$= - \frac{d^2}{d\lambda^2} \sum_{i=1}^n \left\{ -\lambda + x_i \log(\lambda) - \log(x_i!) \right\} \Big|_{\lambda=\lambda^*}$$

$$= - \frac{d}{d\lambda} \sum_{i=1}^n \left\{ -1 + \frac{x_i}{\lambda} \right\} \Big|_{\lambda=\lambda^*}$$

$$= \sum_{i=1}^n \frac{x_i}{\lambda^2} \Big|_{\lambda=\lambda^*}$$

$$= \frac{1}{(\lambda^*)^2} \cdot \sum_{i=1}^n x_i$$

Approximately, if n large,

$$\hat{\Delta}^{MLE} \sim \text{Normal}\left(\lambda, \frac{1}{\frac{1}{\bar{x}^2} n \cdot \frac{1}{n} \sum_{i=1}^n x_i}\right)$$

$$\hat{\Delta}^{MLE} \sim \text{Normal}\left(\underline{\lambda}, \frac{\bar{x}}{n}\right)$$

Need a pivotal quantity:

- can involve data and parameters
- its distribution can't depend on unknown parameters

Many options; most common:

$$\frac{\hat{\Delta}^{MLE} - \lambda}{\sqrt{\bar{x}/n}} \sim \underline{\text{Normal}(0, 1)}$$

(same thing)

$$\textcircled{1} \quad \frac{\bar{x} - \lambda}{\sqrt{\bar{x}/n}} \sim \text{Normal}(0, 1)$$

This is an approximation

The pivotal quantity in green does not exactly follow a normal distribution

② If n large,

$$P\left(z\left(\frac{\alpha}{2}\right) \leq \frac{\hat{\lambda}^{MLE} - \lambda}{\sqrt{\bar{x}/n}} \leq z\left(1 - \frac{\alpha}{2}\right)\right) \approx 1 - \alpha$$

↑ because based on a large-sample normal approx.

where $z\left(\frac{\alpha}{2}\right)$ and $z\left(1 - \frac{\alpha}{2}\right)$ are the $\frac{\alpha}{2}$ and $1 - \frac{\alpha}{2}$ quantiles of a Normal $(0, 1)$ distribution,

③ Continuing from above:

$$P\left(z\left(\frac{\alpha}{2}\right)\sqrt{\bar{x}} \leq \hat{\lambda}^{MLE} - \lambda \leq z\left(1 - \frac{\alpha}{2}\right)\sqrt{\bar{x}}\right) \approx 1 - \alpha$$

$$\Rightarrow P\left(-\hat{\lambda}^{MLE} + z\left(\frac{\alpha}{2}\right)\sqrt{\bar{x}} \leq -\lambda \leq -\hat{\lambda}^{MLE} + z\left(1 - \frac{\alpha}{2}\right)\sqrt{\bar{x}}\right) \approx 1 - \alpha$$

$$\Rightarrow P\left(\hat{\lambda}^{MLE} - z\left(\frac{\alpha}{2}\right)\sqrt{\bar{x}} \geq \lambda \geq \hat{\lambda}^{MLE} - z\left(1 - \frac{\alpha}{2}\right)\sqrt{\bar{x}}\right) \approx 1 - \alpha$$

$$\Rightarrow P\left(\hat{\lambda}^{MLE} - z\left(1 - \frac{\alpha}{2}\right)\sqrt{\bar{x}} \leq \lambda \leq \hat{\lambda}^{MLE} - z\left(\frac{\alpha}{2}\right)\sqrt{\bar{x}}\right) \approx 1 - \alpha$$

So an approximate $(1 - \alpha) * 100\%$ C.I. for λ is:

$$\left[\hat{\lambda}^{MLE} - z\left(1 - \frac{\alpha}{2}\right)\sqrt{\bar{x}}, \hat{\lambda}^{MLE} - z\left(\frac{\alpha}{2}\right)\sqrt{\bar{x}} \right]$$

This will not have a actual confidence level of $(1 - \alpha) * 100\%$ because it's based on a normal approximation to the distribution of the pivotal quantity, not the actual distribution of

$$\frac{\hat{\lambda}^{MLE} - \lambda}{\sqrt{\bar{x}/n}} !$$

Def.: The nominal confidence level (or the nominal coverage rate) of a confidence interval is the claimed proportion of samples for which the interval contains the actual value of the parameter being estimated.

$$(1 - \alpha) \times 100\%$$

Comment: If the confidence interval was derived based on an approximate sampling distribution, the actual coverage rate (or confidence level) may be different from the nominal coverage rate.