Speach in Mexico's Cámara de Diputados

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Abstract

Text as data: speaches in lower chamber of Mexico's federal Congress. Analysis covers three pre-midterm election legislative terms since 2006. Argument, findings.¹

Keywords: Speach, Congress, presidentialism, Mexico

¹Data and supporting materials necessary to reproduce the numerical results in the article are available in the following repository (https://github.com/emagar/legdeb). Supplementary material for this article is available in the appendix in the online edition.

1 Introduction

2 Paste in paper

2.1 Terminology

- A *Legislature* is an elected chamber for a legislative term, called a Congress in the U.S. Concurrent with presidential elections the chamber of deputies renovates in whole, and again at the presidential mid-term. Diputados remain three years in office and were single term-limited up to 2021. The 2021 mid-term election will be the first since 1932 to allow incumbents on the ballot, a major change in Mexican legislative politics. Analysis includes the 60th, 62nd, and 64th Legislatures (the Mexican Congress relies on Roman numerals to distinguish Legislatures since the second half of the Nineteenth century).

- Legislative years break into two *ordinary periods*, one covering the months of September through December, inclusive, another February through April, also inclusive. *Extraordinary periods* may be convened during the recess in order to consider a specific bill. Analysis aggregates each member's speeches in the duration of a given period (merging together all extraordinary periods that year, if any). So members in a legislative year like 2012-13 (that had no extraordinary periods) have two word aggregates in the dataset, one for each ordinary period; in a year like 2013-14 (that did), they have three word aggregates in the data. Periods are the units of observation in the analysis.

- A *plenary session* (or simply a session) is a specific date in the calendar when diputados met. During ordinary periods, sessions are usually held on Tuesdays and Thursdays, and may be scheduled in other weekdays if the Jucopo so decides. Diputados met on forty and thirty-one days in the first and second ordinary periods of 2013-14, respectively, and nine days in extraordinary periods, for a yearly total of eighty session days. (A session in North-American legislative parlance is a Mexican period.)

2.2 The dependent variables

Descriptive statistics

```
* * Descriptives for diputados vs presiding officers * *
> print(summ)
$mean
  Group.1 dv.nword ev.pot.dys dv.nword.by.dy
    dips 7539.702 150.8411
2 pres.off 65215.062
                    188.8500
                                395.45338
$median
  Group.1 dv.nword ev.pot.dys dv.nword.by.dy
                      180
            3677
                                26.37005
2 pres.off
            46710
                       182
                                258.64011
$sd
  Group.1 dv.nword ev.pot.dys dv.nword.by.dy
   dips 14811.50 73.40782
2 pres.off 63481.32
                  79.74391
                           501.86471
$min
  Group.1 dv.nword ev.pot.dys dv.nword.by.dy
1 dips 0 1 0.00000
2 pres.off
             3141
                        32
                               19.64286
$max
  Group.1 dv.nword ev.pot.dys dv.nword.by.dy
     dips 242434 522 1212.170
2 pres.off 320155
                      522
                                3563.254
```

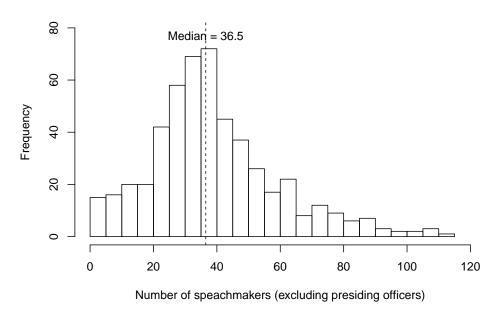
2.3 N speakers

2.4 By legislatura

```
* * Descriptives by legislatura * *
> print(round(summ, 0))
 Group.1 nword.day.50% nword.day.90% nword.day.max nspeakers.day.50%
    60 652 1215 15932
1
2
     62
                            1254
                611
                                        9765
                                                          38
      64
                 547
                            975
                                         6358
                                                          56
 nspeakers.day.90%
1
2
              61
3
              81
```

Figure 1: Average marginal effects from model 3. Dots report how the probability of an urgent bill changes in response to a unit change in each independent variable, all else at mean values; bars are 95-percent confidence intervals.

How many spoke in a plenary session



3 Stuff from urge.tex follows

A Model of Urgency-as-Fast-Track Authority

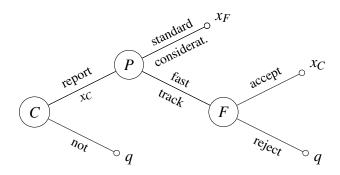
Hypothesis 1: Presidents are more likely to fast-track bills when the committee chair with jurisdiction over the bill belongs to the president's party than otherwise.

Data and Analysis

We collected original data to test the hypothesis, compiling bill histories of every draft law that the executive introduced in Congress between 1998 and 2014:² when each was introduced, in which chamber of the bicameral Congress, the issue it deals with, the status at the time of consultation,

²We scraped the *Cámara de Diputados*' web page (www.camara.cl) in November 2014 to retrieve the record (*boletín*) of every proposal made between 11 March 1998 and 10 March 2014, inclusive. Data and commented code for replication accompany the online appendix. We conducted data analysis with a multiplicity of R's libraries.

Figure 2: The president rules game



and so forth. We also gathered information on the chronological detail of the bill's legislative process in the House: committee referrals and reports, floor discussion and voting, when it was sent to the Senate, and more. Of direct relevance, we coded all bills marked urgent in the *Cámara de Diputados*. Earlier years antedate Internet publication and were dropped, as data completeness in the primary source remains to be verified. The period selected fully covers two Senates, four *Cámaras*, and three presidencies (plus the last two years of an earlier presidency).

Table 1 offers a general summary of bill introductions, passages, and fast-track incidences. The executive sent 1,467 bills to Congress between 1998 and 2014, on average ninety-one yearly. (Members of Congress proposed 79 percent of all bills, not analyzed.) More than one thousand proposals became law during the period, putting the executive's success rate at 72 percent—high by Latin American standards (Morgenstern and Nacif 2002). And 540 bills were on the fast-track during lower chamber consideration, 37 percent of all bills introduced. And 77 percent of those became law, so about 40 percent of bills that became law did it under restrictive procedures.

The variance in urgency incidence across administrations, reported in part B, is considerable. But differences in urgency authority usage do not seem related with specifics traits of the president. For instance, one could argue that presidents with previous legislative experience might be less inclined to interfere with Congressional priorities. Frei (with previous legislative experience) and Bachelet (without) were about on average, Lagos (with experience) quite below, and Piñera (without) quite above.

Table 1: Proposals, legislation, and fast-track authority

Part A. Executive bills

	Bills	frequency
I	introduced	1,467
II	passed	1,059
	as % of introduced	72
III	fast-tracked	540
	as % of introduced	37
ĪV	fast-tracked & passed	415
	as % of fast-tracked	77

Part B. Urgent bills by presidency

President and period	N bills	% fast-tracked
Frei 1998–2000 [†]	128	38
Lagos 2000-2006	544	25
Bachelet 2006-2010	392	39
Piñera 2010–2014	403	50
All 1998–2014	1,467	37

[†] Last third of the six-year term in the analysis only.

Fast-Track predictors

Multivariate analysis of the data is revealing. The unit of analysis is individual executive proposals: the dependent variable *Fast-tracked Bill* equals 1 for proposals marked with 'supreme urgency' while in the Cámara, 0 otherwise. Our main independent variable accounts for preference coincidence between the president and the reporting committee. We include controls for bill features, for timing, and for the strategic environment. Formal variable definitions and descriptive statistics appear in the online appendix.

With respect to our main independent variable, which accounts for preferences, we include *Copartisan Chair* and *Coalition Chair*, which seek to identify committee chairs' preference location vis-à-vis the president, and *Multiple Referrals*, which identifies bills referred to multiple committees. *Co-partisan Chair* equals 1 if the bill was referred to a standing committee presided by a member of the president's party, and equals 0 otherwise; *Coalition Chair* equals 1 for bills referred to committees chaired by members of any party in the presidential coalition, 0 otherwise. These are two different ways in which we measure our key explanatory variable, spatial proximity be-

Table 2: The president's status in Congress and its committees. Percent chairs/seats by party. The president's coalition in 1998–2010 was Concertación; it was Alianza afterwards. Regional includes major-party splinters (from Christian Democrats and UDI). President's status in the Senate slightly and briefly oscillated above and below majority due to vacant seats. Source: prepared with information from www.camara.cl.

	1998-2002	2002-06	2006-10	2010-14					
Part A. Committee chairs, Cámara									
President's party	53	35	17	23					
Other coalition party	41	41	83	50					
Opposition	6	24		27					
Total	100	100	100	100					
<i>N</i> standing committees	17	17	18	22					
Part B. Seats, Cámara									
President's coalition	58	53	51	50					
Opposition	42	48	47	48					
Regional			3	2					
Total	100	100	100	100					
Part C. Seats, Senate									
President's coalition	50	50	55	45					
Opposition	50	50	45	55					
Total	100^{\dagger}	100	100	100					
†vacant seats dronned									

†vacant seats dropped

tween the chief executive and the reporting committee, and we expect each to associate positively with the dependent variable. Part A of Table 2 shows that the number of standing committee chairs in hands of the president's party varied in the period, from a high of 53 percent in the 1998–2002 legislature to a low of 17 percent in 2006–10. And the opposition chaired no standing committee in 2006–10, but up to 24 and 27 percent in 2002–06 and 2010–14, respectively.³

We also control for multiple referrals. Nearly one quarter (24 percent) of bills in the period were referred to more than one standing committee. The "other committee" count excludes the Finance Committee, with jurisdiction over any form of new spending (and discussed next—multiple referrals reach 32 percent when the Finance Committee is considered). Also excluded are special

³Largesse towards opposition parties was probably aimed at beefing up the president's legislative support. Unlike the Senate, the coalition remained in control of the Cámara throughout the period. But, by requiring 67, 60, and 57 percent votes of each chamber, respectively, constitutional reform, constitution-interpreting legislation, and organic laws therefore always required support across the aisle.

and bicameral committees. *Multiple Referrals* should capture any effect of agenda control sharing among several committee chairs during the proposal's negotiation—reflecting the need to rein on unruly chairs through a friendlier committee. A single co-partisan or coalition chair among multiple referees suffices for the indicator previously discussed to equal 1.

The variable intended to capture bill-specific features is *Hacienda Referral*, which equals 1 for bills referred to the powerful Finance Committee with special status in the Chilean Congress, 0 otherwise. Unlike other standing committees, the Finance Committee has jurisdiction over *every* bill authorizing spending in any domain. Moreover, the unanimous exception rule discussed earlier is inapplicable to *Hacienda* bills, which must be reported prior to floor consideration.⁴ Committee members, working in tandem with the Finance Ministry, may or may not appropriate funds from the budget in their report (Alemán and Navia 2009). Not unlike the Appropriations and Rules committees in the U.S. House, *Hacienda* has the status of a control committee, a key asset for agenda power (Kiewiet and McCubbins 1991). *Hacienda* referral therefore controls for a subset of generally important proposals, and we expect it to associate positively with urgency authority.

Three controls account for the strategic environment. *Presidential Approval* is the net general population presidential approval rate at bill initiation (i.e., the percentage of respondents who approve of the president's job minus the percentage who disapprove). We have no a priori expectation here. If presidents with better public opinion standing are also more successful in the legislative arena (Alemán and Navia 2009; Bond and Fleisher 1990), they might also need restrictive rules less often, and reliance on the fast-track might therefore drop (in some issue areas, at least). The contrary might hold if popular presidents were more successful in obtaining more likeable reports from the average committee chair, that would then require protection against floor amendments. *Introduced in Senate* equals 1 for bills sent to the Upper Chamber, 0 otherwise. By virtue of being smaller, enjoying longer terms, and not being firmly under the president's coalition control during most of the period, bills sent to the Senate might present systematic differences in

⁴Ley Orgánica del Congreso arts. 17 and 21.

⁵We estimate the models using subsets of bills by issue area in the online appendix. Smaller numbers of bills tend to hamper statistical significance, but the results are in line with those reported.

fast-track usage. And *Senate Majority* equals 1 if the president's coalition controlled half or more of Upper Chamber seats when the bill was initiated, 0 otherwise.⁶ Other things equal, presidents with sufficient partisan legislative resources in both chambers will find it easier to push proposals through Congress, and might be less inclined to use the fast-track prerogative to successfully navigate log-rolls through the plenary session.

The group of variables accounting for time-related effects control for different aspects of the congressional cycle. *Year Remaining* (and its squared value to capture non-linearity, if any) measures the percentage of the legislative year remaining at bill initiation. Chilean legislative years start after the (meridional) summer break. So the variable adopts value 100 for proposals introduced on March 1 (the first day of the legislative year), and value 0 for proposals introduced the last day of February. It should control for stationarity in the data (the online appendix elaborates the temporal dimension in the use of urgencies). And *Relax Deadline* equals 1 for bills initiated in July 2010 or later, 0 otherwise. Congress doubled deadlines to consider and vote urgent bills five months into the 2010–14 legislature (see online appendix). Any systematic shift in urgency usage attributable to this reform should be reflected in this coefficient.

Model Specification and Results

Given that we pool observations from four elected legislatures, with important differences in the types and the volume of proposals considered (Alemán and Navia 2009), heterogeneity might interfere. We fit two additional model specifications for robustness verification. One includes fixed legislature effects—i.e., three dummies for bills initiated in the 2002–06, 2006–10, and 2010–14 periods, respectively; the excluded 1998–2002 dummy is the baseline. Another adds further flexibility by also estimating separate errors for bills initiated in each legislature (a so-called mixed effects model, Gelman and Hill 2007, 262, 302). We rely on a generalized linear model for mixed effects estimation, and logit for the rest. We normalized continuous variables *Presidential Ap-*

 $^{^6}$ Parties in the presidential and opposition coalitions were tied throughout most of the 1998–2006 Senate (majority briefly oscillating back and forth in the first years due to member indictments, impeachments, and deaths in both coalitions). Ties are coded as *Senate Majority* = 1.

proval and Year Remaining to speed the GLM's convergence. Normalized measures were used throughout for model comparability.

Table 3 reports results. We find support for our main hypothesis. In line with expectations, both variables measuring the proximity of committee chairs to the president increase the probability of fast-track. *Co-partisan Chair* has a positive coefficient in model 1, as hypothesized, and the effect achieves conventional statistical significance (parentheses in the table report p-values). The evidence is much stronger for the variable's other specification: the coefficient for *Coalition Chair* in models 2–4 is also positive, triples its size, and achieves p-values at .005 or below. In other words, if a bill is referred to a committee where the president finds support (either because the chair belongs to her party or to her coalition), the report is more likely to be fast-tracked for floor consideration. This matches previous scholarship, i.e., that the coalition is as good a predictor of presidential support in Congress—or better, in our case—as the party. The finding is robust across model specifications. In general, all model coefficients remain pretty much unchanged in size and significance when fixed and mixed effects are included on the right side (*Senate Majority* and *Relax Deadline* must be dropped due to colinearity with legislature dummies).

Figure 1 reports changes in the average predicted probability that a bill is fast-tracked associated with unit changes in model 3's explanatory variables (all other regressors at their mean value). This is a convenient way to gauge logit regression coefficients, by translating them into interpretable quantities. The report from a committee with a coalition chair experiences a 0.17 hike (0.06 standard error) in the likelihood of receiving a closed rule compared to a report by an opposition-chaired committee. The effect is as big as the average marginal effects of *Hacienda Referral* (0.18), which capures mostly high-significance draft laws, and that of *Multiple Referrals* (0.16), which we view as an indicator of issue complexity. We therefore find no statistical evidence

⁷As suggested in http://stackoverflow.com/questions/23478792/warning-messages-when-trying-to-run-glmer-in-r and https://rstudio-pubs-static.s3.amazonaws.com/33653_57fc7b8e5d484c909b615d8633c01d51.html. Normalization re-scales and centers the measures in order to improve parameter identification.

⁸The regression model performs satisfactorily. A likelihood-ratio test of overall fit rejects the hypothesis, at below the .001 level, that an intercept-only fit is as good as our models. Predictors across model specifications correctly classify 67–68 percent of the observations.

Table 3: Executive bill fast-track predictors. Standard errors in parentheses. Model 3 includes fixed Legislatura effects (not reported). Model 4 estimates separate error terms by Legislatura. Method of estimation: model 4 with generalized linear model, others with logit (fitted with R base's glm and library lme4, Bates et al. 2015).

	DV: Bill on fast-track (1) or not (0)				
	(1)	(2)	(3)	(4)	
Co-partisan Chair	.289** (.024)				
Coalition Chair		.825*** (.005)	.874*** (<.001)	.847*** (<.001)	
Multiple Referrals	.772*** (<.001)	.795*** (<.001)	.808*** (<.001)	.809*** (.004)	
Hacienda Referral	1.002*** (<.001)	.940*** (<.001)	.917*** (<.001)	.923*** (<.001)	
Pres. Approval	078 (.286)	096 (.187)	.029 (.710)	044 (.567)	
Introduced in Senate	716^{***} (<.001)	698*** (<.001)	744*** (<.001)	730*** (<.001)	
Senate Majority	251 (.214)	319 (.110)			
Year Remaining	.072 (.223)	.065 (.268)	.053 (.370)	.053 (.368)	
(Year Remaining) ²	224^{***} (<.001)	238*** (<.001)	255*** (<.001)	251*** (<.001)	
Relax Deadline	.479* (.057)	.394 (.104)			
Intercept	-1.046*** (<.001)	-1.589*** (<.001)	-1.933*** (<.001)	$-1.719^{***} $ (<.001)	
Effects Observations	none 1,467	none 1,467	fixed 1,467	mixed 1,467	
Log <i>L</i> % correct	-864 67	-862 68	$-852 \\ 68$	-859 68	

^{*}p<.1; **p<.05; ***p<.01 (p-values in parentheses)

to reject our Hypothesis 1. The results also confirm hypothesis 2, showing that a bill reported by a generally less friendly committe (chaired by the opposition), has a higher probability of receiving an open rule on the floor, thereby allowing the floor majority to bring back the bill to the median through floor amendments. Thus, presidents use open rules to control bills coming from preference distant committee chairs.

The substantial effects of *Hacienda Referral* and *Multiple Referrals* deserve comment. They suggest, first, that when spending gets in the way, restrictive rules are the norm in Chile. Recall that *Multiple Referrals* exclude the Finance Committee, so there is an independent effect of bills with jurisdictional overlaps worth investigating further, and which must be associated, in part at least, to influencing the report through a friendlier committee. Furthermore, note that the Finance Committee was always chaired by a coalition member but, with the exception of the 1998–2000 period, never by a co-partisan of the president. This may explain the milder effect of the partisan specification of our key variable in model 1.

Another effect worth highlighting is *Introd. in Senate*. Bills successfully passing the Upper Chamber first, where the opposition was systematically larger and at times in control, were much less likely to get urgent status (the average marginal effect is -0.15 and significant). This suggests that agreements and compromises reached in the Senate ignited less, not more, protection from floor amendments in the *Cámara's* plenary, most likely as a consequence of the greater preference divergence between the President and the opposition-led Senate. Analysis of inter-chamber negotiation and the reliance on urgency in the Upper Chamber are interesting venues for future research.

Finally, there are time trends in fast-track authority that simulations reveal neatly. Figure ?? portrays the predicted probability that a bill enters the fast-track throughout the legislative year. Regressors in model 3 are held constant to simulate a bill sent to the *Cámara* in the 2006–10 Legislature that was referred to a single committee, excluding *Hacienda*. *Presidential Approval*

⁹We are grateful for this insight from an anonymous referee. According to Soto Velasco (2015, 118), multiple-referees may act sequentially or in tandem, as decided by the Cámara's presiding officer. When sequential, a divergent secondary committee's report is treated as an amendment to the primary committee's—an urgency overrides the secondary (and subsequent) report(s). We unfortunately lack information on this important aspect of multiple referrals.

(insignificant across models) is set to the mean for President Bachelet's first term, coinciding in full with the 2006–10 Legislature. The inverted-U shape shows how fast-track probability, predicted at 0.17 for coalition-chaired committees at the start, and 0.08 for the rest, becomes much likelier in the first half of the legislative year. By the second quarter (June–August), the probability is at its peak, about 0.32 percent and 0.17, respectively. It then experiences a sharp drop, ending the austral Summer break at 0.13 for coalition-chaired committees, and 0.05 for others. And while 95-percent confidence bands overlap, they barely do so at the middle of the legislative year, lending confidence that we are picking up a signal and not just random noise.

Discussion

This paper has argued that beyond the effect of urgency authority on the timing and deadlines for bill consideration, its procedural effects conceal a much more significant effect, by allowing presidents to shield desired policy from amendments on the floor. Furthermore, the president's decision not to use this tool is also important. An open rule means that bills can receive amendments on the floor that may move it closer to the president's preferences. This paper portrays urgency authority, found in seven Latin American constitutions, as equivalent to the fast-track authority that United States presidents enjoy periodically. Doing so shows how proposals that presidents qualify as urgent are considered under restrictive rules by the chamber's plenary: they must be voted up-or-down, without amendments.

While all seven constitutions remain silent about the procedural implications of urgency authority, and we have only looked for these restrictive procedures in statutes and chamber rules in the case of Chile, there are good reasons to expect that other countries have included similar provisions. After all, the rationale of the urgency authority is expediting the legislative process, and restrictive rules are a natural choice to speed urgent bills' consideration before an explicit deadline expires.

The classification of constitutional urgency authority into the plenary arrest (Brazil), automatic

adoption (Ecuador, Paraguay, Uruguay), and indeterminate (Chile, Colombia, Mexico) variants, which we presented in Section 2, naturally brings two pending issues to the fore. One is cross-national validation of the presence of restrictive procedures at the level of statutes or chamber rules—especially where urgency effects are seemingly indeterminate, as in Chile. The task ahead is to identify procedures that restrict the choices available to legislators, such as closed rules do in Chile. The other is the need to model the bargaining logic of plenary-arrest and automatic-adoption variants in search of similarities and differences with the urgency-as-fast-track. We expect these efforts to generalize our argument beyond Chile, our case study.

Game-theoretic treatment of urgency-as-fast-track authority shows that preference overlap between the president and the reporting committee is the mechanism driving the choice to put bills on the fast-track. The reverse also applies: bills referred to opposition-chaired committees, who might report them to the floor with fundamental changes, are less likely to be fast tracked. The president would prefer them to go to the floor with an open rule, especially if on the floor a (friendlier) majority can restore the original intent of the president and her party. Systematic analysis of proposals in the Chilean lower chamber in recent years yields evidence that, other things constant, bills reported by committees chaired by members of the president's coalition/party are about twice as likely to be fast-tracked than the rest. To the extent that parties and coalitions indicate preference overlap (as is accepted in Chile), the evidence supports our main theoretical result.

The paper's results and findings are of natural interest to students of comparative political institutions, especially those interested in legislative procedure and separation of powers. They also shed light on the field of American Politics. In a provocative book, Howell and Moe (2016) make an argument in favor of giving U.S. presidents permanent fast-track authority not limited to trade agreements. In order to have a coherent and effective government, they argue in favor of constitutional reform to put the executive at the center of the legislative process: "presidents should be granted enhanced agenda-setting powers to propose bills to Congress, which Congress should then be required to vote on without amendment, on a strictly majoritarian basis, within a fixed period of time" (Howell and Moe 2016, 145). Failure to vote in due time would turn those

bills into law. Equating urgency and fast-track authorities shows this to be the automatic adoption variant. Reform would therefore make U.S. presidents similar to those in Ecuador, Paraguay, and Uruguay in this respect.

Our results also speak to students of executive decrees and unilateral policymaking more generally. As discussed at the start of the paper, urgency authority and decree authority may overlap, as in fact they do in Brazil, Chile, and Colombia. We have established that, while urgency authority increases the president's leverage over the legislative process, an agreement is needed between the president and her congressional allies without which this prerogative is rendered powerless. Congress retains authority to reject urgent proposals, and cannot therefore be made worse-off than the status quo. Executive decrees, on the other hand, imply an abdication of decision rights on the part of legislators to participate in the process, and force legislators to respond to the president's enacted decree instead once it has already altered the status quo.

Yet for this precise reason executive decrees are not a true alternative in certain cases (Palanza 2019). Figueiredo and Limongi (2000, 164) report that urgent bills in Brazil are quite rare since the provisional decree "is a much more efficient way of speeding up and approving legislation." More research should be done to establish whether effectively the urgency prerogative is used, for instance, on issues that are excluded from executive decree authority in the Brazilian constitution. Previous findings indicate that reliance on decrees effectively diminishes when policy deals with these issues—although it does not disappear. It is undeniable that Chile, where executive decrees are seldom used (only two decrees have been enacted since 1990), resorts to urgency authority expansively. Future research, and hopefully comparative work, will help tackle this issue.

Chilean legislative scholars will find interest in investigating a missing piece in our argument. To be clear, the research advanced in this paper sheds light on the factors that affect the likelihood of reliance on urgency authority, and we provide an argument explaining the mechanism that leads to this usage. But we do not provide verification that closed rules do in fact dispel amendments—a key and untested premise of our approach. Systematic contrasting of bills that presidents drafted in the period, the various amendments that were introduced, the committee reports, and the text

ultimately adopted, would enable analysts to demonstrate whether or not urgencies in fact dodge amendments. This is a difficult but promising venue for future research.

Last, we emphasize that our investigation contributes to the literature on restrictive procedures. To the general class of procedures as instruments of political control, we add the urgency authority. One difference sets fast-track mechanisms apart from standard closed rules. Standard restrictive rules give agenda power to legislators—it may be the plenary (McCubbins, Noll and Weingast 1987), standing committees (Weingast 1992), the bicameral conference (Shepsle and Weingast 1987), or the majority party (Cox and McCubbins 1997). Urgency puts the executive in control of protecting vulnerable legislative agreements. The unique role and function to determine whether or not a major bill will be considered in the U.S. House, and which amendments and motions will be allowed, is firmly in control of the legislative leadership, acting as sole gatekeepers to the plenary (Cox 2006). As with France's *vote bloqué* (Huber 1996), fast-track authority lets the executive branch pull some of the gatekeeping levers, earning a priceless legislative prerogative worthy of further research.

Acknowledgements

The author is grateful to xxx for research assistance. Eric Magar received financial support from the Asociación Mexicana de Cultura A.C. and CONACYT's Sistema Nacional de Investigadores. The authors are responsible for mistakes and shortcomings in the study.

References

Alemán, Eduardo and Patricio Navia. 2009. "Institutions and the Legislative Success of 'Strong' Presidents: An Analysis of Government Bills in Chile." *Journal of Legislative Studies* 15(4):401–19.

Bates, Douglas, Martin Maechler, Benjamin M. Bolker and Steven Walker. 2015. "Fitting Linear

- Mixed-Effects Models using lme4.". ArXiv e-print; in press, *Journal of Statistical Software*. **URL:** http://arxiv.org/abs/1406.5823
- Bond, Jon R. and Richard Fleisher. 1990. *The President in the Legislative Arena*. Chicago: Chicago University Press.
- Cox, Gary W. 2006. The Organization of Democratic Legislatures. In *The Oxford Handbook of Political Economy*, ed. Barry R. Weingast and Donald A. Wittman. New York: Oxford University Press pp. 141–61.
- Cox, Gary W. and Mathew D. McCubbins. 1997. "Toward a theory of legislative rules changes: Assessing Schickler and Rich's evidence." *American Journal of Political Science* 41(4):1376–86.
- Figueiredo, Argelina Cheibub and Fernando Limongi. 2000. "Presidential Power, Legislative Organization, and Party Behavior in Brazil." *Comparative Politics* 32(1):151–70.
- Gelman, Andrew and Jennifer Hill. 2007. Data Analysis Using Regression and Multi-level/Hierarchical Models. Cambridge University Press.
- Howell, William G. and Terry M. Moe. 2016. *Relic: How our Constitution Undermines Effective Government—and Why We Need a More Powerful Presidency*. New York: Basic Books.
- Huber, John D. 1996. *Rationalizing Parliament: Legislative Institutions and Party Politics in France*. New York: Cambridge University Press.
- Kiewiet, D. Roderick and Mathew D. McCubbins. 1991. *The Logic of Delegation: Congressional Parties and the Appropriations Process*. Chicago: University of Chicago Press.
- McCubbins, Mathew D., Roger G. Noll and Barry R. Weingast. 1987. "Administrative Procedures as Instruments of Political Control." *Journal of Law, Economics, & Organization* 3(1):243–77.
- Morgenstern, Scott and Benito Nacif. 2002. *Legislative Politics in Latin America*. New York: Cambridge University Press.

- Palanza, Valeria. 2019. Checking Presidential Power: Executive Decrees and the Legislative Process in New Democracies. New York: Cambridge University Press.
- Shepsle, Kenneth A. and Barry R. Weingast. 1987. "The Institutional Foundations of Committee Power." *American Political Science Review* 81(1):85–104.
- Soto Velasco, Sebastián. 2015. *Congreso Nacional y Proceso Legislativo: Teoría y Práctica*. Santiago: Thomson Reuters.
- Weingast, Barry R. 1992. Fighting Fire with Fire: Amending Activity and Institutional Change in the Postreform Congress. In *The Postreform Congress*, ed. Roger H. Davidson. St. Martin's Press.