

Summary of Chapter 6

1 Concepts

- **Confidence Interval Estimate for μ :**

1. σ is known:

Two-sided: A $100(1 - \alpha)\%$ confidence interval (C.I.) on μ is given by

$$\bar{X} - Z_{\alpha/2} \frac{\sigma}{\sqrt{n}} \leq \mu \leq \bar{X} + Z_{\alpha/2} \frac{\sigma}{\sqrt{n}}$$

One-sided: A $100(1 - \alpha)\%$ upper (or lower)-confidence bound for μ is

$$\mu \leq \bar{X} + Z_{\alpha} \frac{\sigma}{\sqrt{n}} \quad \text{or} \quad \bar{X} - Z_{\alpha} \frac{\sigma}{\sqrt{n}} \leq \mu$$

2. σ is unknown:

Two-sided: A $100(1 - \alpha)\%$ C.I. on μ is given by

$$\bar{X} - t_{\alpha/2, n-1} \frac{S}{\sqrt{n}} \leq \mu \leq \bar{X} + t_{\alpha/2, n-1} \frac{S}{\sqrt{n}}$$

One-sided: A $100(1 - \alpha)\%$ upper (or lower)-confidence bound for μ is

$$\mu \leq \bar{X} + t_{\alpha, n-1} \frac{S}{\sqrt{n}} \quad \text{or} \quad \bar{X} - t_{\alpha, n-1} \frac{S}{\sqrt{n}} \leq \mu$$



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- **Confidence Interval Estimate for p :** A $100(1 - \alpha)\%$ C.I. on p is

$$p_s - Z_{\alpha/2} \sqrt{\frac{p_s(1 - p_s)}{n}} \leq p \leq p_s + Z_{\alpha/2} \sqrt{\frac{p_s(1 - p_s)}{n}}$$

where p_s is the sample proportion.

- **Confidence Interval Estimate for Population Total:** Let N be the population size, a $100(1 - \alpha)\%$ C.I. for population total T is

$$N\bar{X} - Nt_{\alpha/2, n-1} \frac{S}{\sqrt{n}} \sqrt{\frac{N-n}{N-1}} \leq T \leq N\bar{X} + Nt_{\alpha/2, n-1} \frac{S}{\sqrt{n}} \sqrt{\frac{N-n}{N-1}}$$

- **Confidence Interval Estimate for Total Difference:** A $100(1 - \alpha)\%$ C.I. for total difference D is

$$N\bar{D} - Nt_{\alpha/2, n-1} \frac{S_D}{\sqrt{n}} \sqrt{\frac{N-n}{N-1}} \leq D \leq N\bar{D} + Nt_{\alpha/2, n-1} \frac{S_D}{\sqrt{n}} \sqrt{\frac{N-n}{N-1}}$$

where $\bar{D} = \frac{1}{n} \sum_{i=1}^n D_i$, $S_D = \sqrt{\frac{1}{n-1} \sum_{i=1}^n (D_i - \bar{D})^2}$.

- **Determining Sample Size:** Let e denote the sampling error,

$$1. \text{ For } \mu: n = \frac{Z^2 \sigma^2}{e^2} \quad 2. \text{ For } p: n = \frac{Z^2 p(1-p)}{e^2}$$

2 Examples

Example 1. The quality control manager at a lightbulb factory needs to estimate the mean life of a large shipment of lightbulbs. The process standard deviation is known be 100 hours. A random sample of 64 lightbulbs indicated a sample mean life of 350 hours.

- Set up a 95% confidence interval estimate of the true population mean life of lightbulbs in this shipment.
- Do you think the manufacturer has the right to state that the lightbulbs last an average of 400 hours?
- Does the population of the lightbulb life have to be normally distributed here? Explain.
- Explain why an observed value of 320 hours for a single lightbulb is not unusual, even though it is outside the confidence interval you calculated.

a. $\alpha = 0.05$, $n = 64$, $\bar{X} = 350$, $\sigma = 100$, and $Z_{\alpha/2} = Z_{0.025} = 1.96$.

A 95% C.I. on μ is

$$\bar{X} \pm Z_{\alpha/2} \frac{\sigma}{\sqrt{n}} = 350 \pm (1.96) \frac{100}{\sqrt{64}} = 350 \pm 24.5$$

$$\text{or} \quad 325.5 \leq \mu \leq 374.5$$

We are 95% confident that the mean life of lightbulbs in this shipment is between 325.5 and 374.5 hours.



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b. No. The manufacturer cannot support a claim that the bulbs last an average of 400 hours. Based on the data from the sample, a mean of 400 hours does not fall in the 95% confidence interval. In fact, a mean of 400 hours would present a distance of 4 standard deviations above the sample mean of 350 hours ($\sigma_{\bar{X}} = 100/\sqrt{64} = 12.5$).

c. No. Since σ is known and $n = 64$, from the **Central Limit Theorem**, we may assume that the sampling distribution of \bar{X} is approximately normal.

d. An individual value of 320 is only 0.30 standard deviations ($0.3\sigma = 0.3 \times 100 = 30$) below the sample mean of 350. The confidence interval represents bounds on the estimate of the mean of a sample of 64, not an individual value.

Example 2. The compressive strength of concrete is being tested by a civil engineer. He tests 12 specimens and obtains the following data.

2216 2237 2249 2204 2225 2301 2281 2263 2318 2255 2275 2295

- Using one of graphical statistical tools check the assumption that compressive strength is normally distributed. Include a graphical display in your answer.
- Construct 95% two-sided confidence interval on the mean strength.
- Construct a 95% low-confidence bound on the mean strength. Compare this bound with the lower bound of the two-sided confidence interval and discuss why they are different.

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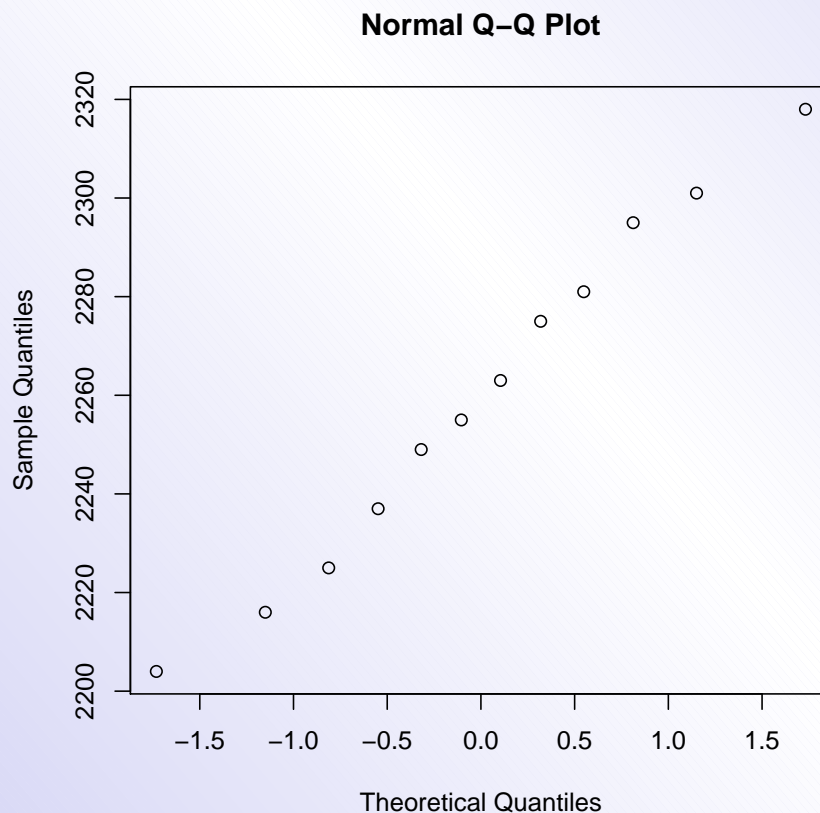
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a. The data appear to be normally distributed based on examination of the normal probability plot below.

R code:

```
> a<-c(2216, 2237, 2249, 2204, 2225, 2301, 2281, 2263, 2318,  
       2255, 2275, 2295)  
> qqnorm(a)
```



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b. $n = 12, \bar{X} = 2259.9, s = 35.6, t_{0.025,11} = 2.201$. A 95% two-sided confidence interval on mean compressive strength is

$$\begin{aligned}\bar{X} - t_{0.025,11} \frac{S}{\sqrt{n}} &\leq \mu \leq \bar{X} + t_{0.025,11} \frac{S}{\sqrt{n}} \\ 2259.2 - 2.201 \left(\frac{35.6}{\sqrt{12}} \right) &\leq \mu \leq 2259.2 + 2.201 \left(\frac{35.6}{\sqrt{12}} \right) \\ 2237.3 &\leq \mu \leq 2282.5\end{aligned}$$

We are 95% confident that the true mean strength is between 2237.3 and 2282.5.

c. A 95% lower-confidence bound on mean strength is

$$\begin{aligned}\bar{X} - t_{0.05,11} \left(\frac{S}{\sqrt{n}} \right) &\leq \mu \\ 2259.9 - 1.796 \left(\frac{35.6}{\sqrt{12}} \right) &\leq \mu \quad \text{or} \quad 2241.4 \leq \mu\end{aligned}$$

We are 95% confident that the true mean strength is not less than 2241.4.

This lower-bound (2241.4) is different from the lower-bound of the two-sided C.I. (2237.3) because the critical values are different in the two cases (in Part b, $t_{0.025,11}$ is used, in Part c, $t_{0.05,11}$ is used).



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Example 3. A study conducted by At-A-Glance communications found that 434 office workers out of a sample of 611 responded to e-mail within an hour or two (D. Haralson and S. Ward, “You Have Mail”, *USA Today*, May 7, 2002, 1A).

- a. Construct a 95% confidence interval for the proportion of workers who responded to e-mail within an hour or two.
- b. Interpret the interval constructed in a.
- c. If you were to conduct a follow-up study to estimate the population proportion of workers who responded to e-mail within an hour or two within ± 0.01 with 99% confidence interval, how many workers would you interview?
- d. In Part c, if the confidence level is reduced to 90%, what is the change in the number of workers you would interview?

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a. $\alpha = 0.05$, $Z_{\alpha/2} = Z_{0.025} = 1.96$, $p_s = 434/611 = 0.7103$, and $n = 611$.
A 95% C.I. on p is

$$\begin{aligned} p_s \pm Z_{\alpha/2} \sqrt{\frac{p_s(1-p_s)}{n}} &= 0.7103 \pm 1.96 \sqrt{\frac{0.7103(1-0.7103)}{611}} \\ &= 0.7103 \pm 0.03597, \quad \text{or} \quad 0.6743 \leq p \leq 0.7463 \end{aligned}$$

b. We are 95% confident that the true proportion of workers who responded to e-mail within an hour or two is between 0.6743 and 0.7463.

c. $\alpha = 0.01$, $Z_{\alpha/2} = Z_{0.005} = 2.5758$,

$$n = \frac{Z_{\alpha/2}^2 p_s(1-p_s)}{e^2} = \frac{2.5758^2 (0.7103)(1-0.7103)}{(0.01)^2} = 13652.58,$$

then we use $n = 13653$.

d $\alpha = 0.1$, $Z_{\alpha/2} = Z_{0.05} = 1.65$,

$$n = \frac{1.65^2 (0.7103)(1-0.7103)}{(0.01)^2} = 5602.20,$$

then we use $n = 5603$.

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