

# Image Denoising

## Course 1 assignments

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Ex 4.1  $x \sim P(\lambda)$

$$E[x] = \sum_{n=0}^{\infty} n \frac{\lambda^n e^{-\lambda}}{n!} = e^{-\lambda} \sum_{n=1}^{\infty} \frac{\lambda^n}{(n-1)!} = e^{-\lambda} \lambda \sum_{n=1}^{\infty} \frac{\lambda^{n-1}}{(n-1)!}$$

Taylor series of  $e^x$

$$= e^{-\lambda} \lambda e^{\lambda} = \lambda$$

$$\sigma^2(x) = \sum_{n=0}^{\infty} n^2 \frac{\lambda^n e^{-\lambda}}{n!} - \lambda^2 = e^{-\lambda} \lambda \sum_{n=1}^{\infty} n \frac{\lambda^{n-1}}{(n-1)!} - \lambda^2$$

$$= e^{-\lambda} \lambda \left[ \sum_{n=1}^{\infty} \frac{(n-1) \lambda^{n-1}}{(n-1)!} + \sum_{n=1}^{\infty} \frac{\lambda^{n-1}}{(n-1)!} \right] - \lambda^2$$

$$= e^{-\lambda} \lambda \left[ \sum_{n=0}^{\infty} \frac{n \lambda^n}{n!} + \sum_{n=1}^{\infty} \frac{\lambda^{n-1}}{(n-1)!} \right] - \lambda^2$$

Previous result!

$$= e^{-\lambda} \lambda [e^{\lambda} \lambda + e^{\lambda}] - \lambda^2$$

$$= \lambda$$

Ex 4.2

$x_1, \dots, x_n \stackrel{iid}{\sim} P(\lambda)$

$y = \sum x_i \sim P(\sum \lambda_i)$

Let's prove that if  $\left. \begin{matrix} X \sim P(\lambda_x) \\ Y \sim P(\lambda_y) \end{matrix} \right\} \Rightarrow X+Y \sim P(\lambda_x + \lambda_y)$

If this holds, then  $\sum_n x_n \sim P(\sum \lambda_n)$

$$\left[ \underbrace{(x_1 + x_2 + \dots + x_{n-1})}_{P} + \underbrace{x_n}_{P} \right]_P$$

$$\begin{aligned} P(X+Y=k) &= \sum_{t=0}^k P(X=t)P(Y=k-t) = \sum_{t=0}^k \frac{\lambda_x^t e^{-\lambda_x}}{t!} \frac{\lambda_y^{k-t} e^{-\lambda_y}}{(k-t)!} \\ &= \sum_{t=0}^k \frac{\lambda_x^t \lambda_y^{k-t} e^{-\lambda_x - \lambda_y}}{t! (k-t)!} = \sum_{t=0}^k \frac{\lambda_x^t \lambda_y^{k-t} k! e^{-\lambda_x - \lambda_y}}{t! (k-t)! k!} \\ &= \sum_{t=0}^k \frac{\lambda_x^t \lambda_y^{k-t} \binom{k}{t} e^{-\lambda_x - \lambda_y}}{k!} = \frac{e^{-(\lambda_x + \lambda_y)} (\lambda_x + \lambda_y)^k}{k!} \end{aligned}$$

Binomial Identity

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### Ex 4.3

We have  $\tilde{u} \approx u + g(u)n$   $n \sim \mathcal{N}(0,1)$

We want  $f$  smooth, such that  $f(\tilde{u})$  gets uniform std independent of  $u$ .

$$\begin{aligned} f(\tilde{u}) &\approx f(u) + f'(u)(\tilde{u}-u) \\ &\approx f(u) + f'(u)g(u)n \end{aligned}$$

we want  $\frac{f'(u)}{g(u)} = c$ ; in our case  $g(u) = \sqrt{u}$

$$\Rightarrow f(u) = \int_0^u \frac{c dt}{\sqrt{t}} = 2c + \frac{1}{2} \Big|_0^u = 2c\sqrt{u}$$

### Ex 4.5

$$D_{inf} = \arg \min_D \mathbb{E} \{ \|U - D\tilde{U}\|^2 \}$$

$$\arg \min_D \mathbb{E} \left\{ \left\| U - \sum_{i=1}^M \alpha(i) \langle \tilde{U}, b_i \rangle b_i \right\|^2 \right\}$$

$$\arg \min_D \mathbb{E} \left\{ \left\| \sum_{i=1}^M \langle U, b_i \rangle b_i - \sum_{i=1}^M \alpha(i) \langle \tilde{U}, b_i \rangle b_i \right\|^2 \right\}$$

$$\arg \min_D \mathbb{E} \left\{ \left\| \sum_{i=1}^M \langle U, b_i \rangle b_i - \sum_{i=1}^M \alpha(i) \langle U + N, b_i \rangle b_i \right\|^2 \right\}$$

$$\arg \min_D \mathbb{E} \left\{ \left\| \sum_i \langle U - \alpha(i)U, b_i \rangle b_i - \sum_i \alpha(i) \langle N, b_i \rangle b_i \right\|^2 \right\}$$

$$\arg \min_D \mathbb{E} \left\{ \left\| \sum_i (1-\alpha(i)) \langle U, b_i \rangle b_i - \sum_i \alpha(i) \langle N, b_i \rangle b_i \right\|^2 \right\}$$

$$\arg \min_D \mathbb{E} \left\{ \left\| \sum_i (1-\alpha(i)) \langle U, b_i \rangle b_i \right\|^2 \right\} + \mathbb{E} \left\{ \left\| \sum_i \alpha(i) \langle N, b_i \rangle b_i \right\|^2 \right\}$$

$$\arg \min_D \left\langle \sum_i (1-\alpha(i)) \langle U, b_i \rangle b_i, \sum_j (1-\alpha(j)) \langle U, b_j \rangle b_j \right\rangle$$

$$+ \mathbb{E} \left\{ \left\langle \sum_i \alpha(i) \langle N, b_i \rangle b_i, \sum_j \alpha(j) \langle N, b_j \rangle b_j \right\rangle \right\}$$

$$\arg \min_D \sum_i (1-\alpha(i))^2 \langle U, b_i \rangle^2 + \mathbb{E} \left\{ \sum_i \alpha(i)^2 \langle N, b_i \rangle^2 \right\}$$

$$\arg \min_D \sum_i (1-\alpha(i))^2 \langle U, b_i \rangle^2 + \sum_i \alpha(i)^2 \sigma^2$$

The cross terms between  $U$  and  $N$  are 0 in expectation since  $N$  is indep of  $U$  and  $\mathbb{E}(N)=0$

$U$  is not random

$b_i$  is an orthonormal basis

Noise remains uncorrelated in the new basis



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Ex 4.5 The objective is convex in  $\alpha(i)$ , so:

Deriving wrt  $\alpha(i)$  and equal to 0

$$-2(1 - \alpha(i)) \langle U, b_i \rangle^2 + 2\alpha(i) \sigma^2 = 0$$

$$\alpha(i) [2\sigma^2 + 2\langle U, b_i \rangle^2] = 2\langle U, b_i \rangle^2$$

$$\boxed{\alpha(i) = \frac{\langle U, b_i \rangle^2}{\sigma^2 + \langle U, b_i \rangle^2}}$$

Evaluating the MSE on  $\alpha(i)^*$ :

$$\sum_i \alpha_i^2 [\sigma^2 + \langle U, b_i \rangle^2] - 2\alpha_i \langle U, b_i \rangle^2 + \langle U, b_i \rangle^2$$

$$\sum_i \frac{\langle U, b_i \rangle^4 (\sigma^2 + \langle U, b_i \rangle^2)}{(\sigma^2 + \langle U, b_i \rangle^2)^2} - \frac{2\langle U, b_i \rangle^4}{\sigma^2 + \langle U, b_i \rangle^2} + \langle U, b_i \rangle^2$$

$$\sum_i \frac{\langle U, b_i \rangle^4 (-1)}{\sigma^2 + \langle U, b_i \rangle^2} + \langle U, b_i \rangle^2$$

$$\boxed{\sum_i \frac{\langle U, b_i \rangle^2 \sigma^2}{\sigma^2 + \langle U, b_i \rangle^2}} \quad \text{MSE}$$

Ex 4.6

$$\text{Now } \alpha(i)^* = \begin{cases} 1 & \text{if } \langle U, b_i \rangle^2 \geq c\sigma^2 \\ 0 & \text{otherwise} \end{cases} \quad \text{for some } c > 1$$

$$\text{let's define } S = \{i \mid \langle U, b_i \rangle^2 \geq c\sigma^2\}$$

$$\Rightarrow \text{MSE} = \sum_{i \in S} \sigma^2 + \sum_{i \notin S} \langle U, b_i \rangle^2 \leq \sum_{i \in S} c\sigma^2 + \sum_{i \notin S} \langle U, b_i \rangle^2$$

Using the  
formula from 4.5

$$\text{Then } \sum_{i \in S} c\sigma^2 + \sum_{i \notin S} \langle U, b_i \rangle^2 \text{ is } \sum_i \min(\langle U, b_i \rangle^2, c\sigma^2)$$

$$\text{So, if using } \alpha^*(i) = \begin{cases} 1 & \text{if } \langle U, b_i \rangle^2 \geq c\sigma^2 \\ 0 & \text{otherwise} \end{cases}$$

$$\text{MSE} \leq \sum \min(\langle U, b_i \rangle^2, c\sigma^2) \text{ and the equality holds}$$

$$\text{when } c=1$$



# Ex 4.7 DCT

$$X, Y \in \mathbb{R}^N$$

$$f(X)_k = \alpha_k \sum_{j=0}^{N-1} X_j \cos\left(\pi\left(j + \frac{1}{2}\right) \frac{k}{N}\right) = \langle X, v_k \rangle$$

$$f(Y)_k = \alpha_k \sum_{j=0}^{N-1} Y_j \cos\left(\pi\left(j + \frac{1}{2}\right) \frac{k}{N}\right) = \langle Y, v_k \rangle$$

"Let's call"

$$v_k = 2\alpha_k \begin{pmatrix} \cos\left[\pi\left(0 + \frac{1}{2}\right) \frac{k}{N}\right] \\ \vdots \\ \cos\left[\pi\left(N-1 + \frac{1}{2}\right) \frac{k}{N}\right] \end{pmatrix}$$

We'll prove that <sup>the</sup>  $v_k$  form an orthogonal base of  $\mathbb{R}^N$   
 (= so DCT is an isometry)

$$\langle v_k, v_s \rangle = \sum_{j=0}^{N-1} 4\alpha_k \alpha_s \cos\left[\pi\left(j + \frac{1}{2}\right) \frac{k}{N}\right] \cos\left[\pi\left(j + \frac{1}{2}\right) \frac{s}{N}\right]$$

$$\langle v_k, v_s \rangle = \sum_{j=0}^{N-1} 2\alpha_k \alpha_s \left[ \cos\left[\frac{\pi}{N}\left(j + \frac{1}{2}\right)(k+s)\right] + \cos\left[\frac{\pi}{N}\left(j + \frac{1}{2}\right)(k-s)\right] \right]$$

$$\langle v_k, v_s \rangle = \sum_{j=0}^{N-1} 2\alpha_k \alpha_s \left[ e^{i\left[\frac{\pi}{N}\left(j + \frac{1}{2}\right)(k+s)\right]} + e^{i\left[\frac{\pi}{N}\left(j + \frac{1}{2}\right)(k-s)\right]} + e^{-i\left[\frac{\pi}{N}\left(j + \frac{1}{2}\right)(k-s)\right]} + e^{-i\left[\frac{\pi}{N}\left(j + \frac{1}{2}\right)(k+s)\right]} \right]$$

$$\langle v_k, v_s \rangle = \alpha_k \alpha_s \left[ e^{\frac{i\pi(k+s)}{2N}} \left( \sum_{j=0}^{N-1} e^{\frac{i\pi(k+s)}{N}j} + \sum_{j=0}^{N-1} e^{\frac{-i\pi(k-s)}{N}(j+1)} \right) \right]$$

$$+ e^{\frac{-i\pi(k-s)}{2N}} \left( \sum_{j=0}^{N-1} e^{\frac{i\pi(k-s)}{N}j} + \sum_{j=0}^{N-1} e^{\frac{-i\pi(k+s)}{N}(j+1)} \right) \right]$$

$$\langle v_k, v_s \rangle = \alpha_k \alpha_s \left[ e^{\frac{i\pi(k+s)}{2N}} \sum_{j=-N}^{N-1} e^{\frac{i\pi(k+s)}{N}j} + e^{\frac{-i\pi(k-s)}{2N}} \sum_{j=-N}^{N-1} e^{\frac{-i\pi(k-s)}{N}j} \right]$$

$$\langle v_k, v_s \rangle = \alpha_k \alpha_s \left[ e^{\frac{i\pi(k+s)}{2N}} \left( \frac{e^{-i\pi(k+s)} - e^{i\pi(k+s)}}{1 - e^{\frac{i\pi(k+s)}{N}}} \right) + e^{\frac{-i\pi(k-s)}{2N}} \left( \frac{e^{-i\pi(k-s)} - e^{i\pi(k-s)}}{1 - e^{\frac{-i\pi(k-s)}{N}}} \right) \right]$$

If  $k \neq s$ :  $\langle v_k, v_s \rangle = 0$  because the subtracted exponentials are either all 1 or all -1.

If  $k=s=0$

$$\langle v_k, v_s \rangle = \frac{1}{4N} \sum_{j=-N}^{N-1} 2 = \frac{4N}{4N} = 1$$

If  $k=s \neq 0$

$$\langle v_k, v_s \rangle = \frac{1}{2N} \sum_{j=-N}^{N-1} 1 = \frac{2N}{2N} = 1$$



### Ex 4.7 IDCT

The IDCT has the exact same basis as the DCT.  
It's clear that's also an isometry.

#### DCT - IDCT

Let's call  $V$  to the DCT matrix and  $Q$  to the IDCT. Since both matrices are orthogonal

$$V^{-1} = V^T$$

$$Q^{-1} = Q^T$$

$$V_{ij} = \sqrt{\alpha_i} \cos \left[ \pi \left( j + \frac{1}{2} \right) \frac{i}{N} \right]$$

$$Q_{ij} = \sqrt{\beta_j} \cos \left[ \pi \left( i + \frac{1}{2} \right) \frac{j}{N} \right]; \text{ considering } \beta_j = \begin{cases} \sqrt{1/4N} & j=0 \\ \sqrt{1/2N} & \text{otherwise} \end{cases}$$

then  $\beta_i = \alpha_i$

$$\text{Then, } (V^T)_{ij} = V_{ji} = Q_{ij}$$

$$\Rightarrow V^T = V^{-1} = Q$$

and

$$Q^T = Q^{-1} = V$$

$\Rightarrow$  DCT and IDCT are inverse transformations

# Image Denoising

## Course 1 Assignments

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### Ex 4.8

$$\arg \min_{\sum \alpha_k = 1} \sum_k \alpha_k^2 \sigma_k^2$$

$$\mathcal{L}_g = \sum_k \alpha_k^2 \sigma_k^2 + \lambda (\sum \alpha_k - 1)$$

$$\frac{\partial \mathcal{L}_g}{\partial \alpha_k} = 2 \alpha_k \sigma_k^2 + \lambda = 0 \Leftrightarrow \boxed{\lambda = -2 \alpha_k \sigma_k^2} \quad \forall k$$

### Ex 4.9

The optimal  $\alpha_k = \frac{\sigma_k^{-2}}{\sum_k \sigma_k^{-2}} \quad \forall k$

Prove that for Wiener filter

$$\alpha_k = \frac{\|P_{P_k}\|^{-2}}{\sum_j \|P_{P_j}\|^{-2}}$$

Now for each pixel, we have  $K^2$  estimators  $P_k$

$\Rightarrow \sigma_k^2$  is the variance of the  $K$ -th estimator for a pixel

Each estimator of the patch is attenuated by  $f_k(w)$  (on each freq)

The variance is due to the residual noise

On the DCT domain we have

$$\sigma^2 \begin{pmatrix} P_p(0) \\ P_p(1) \\ \vdots \\ P_p(s-1) \end{pmatrix} \underbrace{\quad}_{P_p(k)}$$

which is noise across all frequencies (attenuated)

By Parseval

$$\sigma_k^2 = \|\sigma^2 P_p(k)\|^2 \Rightarrow \boxed{\sigma_k^2 = \sigma^2 \|P_p(k)\|^2}$$