

# Relational Frictions Along the Supply Chain: Evidence from Senegalese Traders\*

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## Abstract

Search and trust frictions have historically made it hard for small firms in lower-income countries to buy inputs from foreign markets. The growth in smartphone ownership and social media usage has the potential to alleviate these barriers. Informed by a dynamic model of relational contracting, we run a field experiment leveraging these technological tools to provide exogenous variation in (1) search frictions and (2) trust frictions (adverse selection and moral hazard) in a large international import market. In our search treatment, we connect a randomly selected 80% of 1,862 small garment firms in Senegal to new suppliers in Turkey. We then cross-randomize two trust treatments that provide additional information about the types (adverse selection) and incentives (moral hazard) of these new suppliers. Alleviating search frictions is sufficient to increase access to foreign markets: in all treated groups, firms are 26% more likely to have the varieties a mystery shopper requests and the goods sold are 30% more likely to be high quality. However, the trust treatments are necessary for longer-term impact: using both transaction-level mobile payments data and a follow-up survey, we show that these groups are significantly more likely to develop the connections into relationships that persist beyond the study. These new relationships lead to increases in medium-run profit and sales. Finally, we use the treatment effects to estimate the model and evaluate counterfactuals where we set various combinations of the frictions to zero, finding that the largest gains come from eliminating adverse selection.

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## 1 Introduction

Search and trust frictions make it hard for small firms to access foreign inputs. Consider a clothing wholesaler in Senegal who wants to start selling high quality European-made jeans. First, they must find a supplier, most of whom are in Europe, as well as some way of seeing what this supplier sells. Second, even if they manage to do this and decide to make an order, they face both adverse selection and moral hazard problems as they are typically required to pay before observing the quality. These issues are particularly severe in lower-income settings, whose vast informal sectors complicate information aggregation and lower state capacities make contracts unenforceable. Yet, while there is a long theoretical tradition studying these frictions—dating back to Stigler (1961) on search and Shapiro (1983) on trust—it has proven harder to analyze them empirically as researchers rarely observe variation that is both exogenous and a close map onto the specific theoretical objects.

In recent years, there has been substantial growth in smartphone ownership and social media usage in lower-income countries, which has the potential to fundamentally change how small firms approach buying and selling decisions. Not only are firms increasingly selling online—with e-commerce revenue in Africa estimated to have doubled between 2019 and 2024—but they are primarily doing so through social media rather than traditional platforms.<sup>1</sup> This growth in “social commerce” may reflect the fact that social media could meaningfully alleviate both search and trust frictions in supply chains. For search, the supplier in Europe can send photos of their wares. For trust, domestic firms may be able to much more easily share information and coordinate action to discipline suppliers that cheat—a modern day version of the mechanism in Greif (1993). However, whether it actually does alleviate these frictions is ultimately an empirical question.

In this paper, we provide the first experimental evidence on the extent to which search and trust frictions limit access to foreign input markets and whether these new technologies alleviate them. We designed a field experiment leveraging key features of social media to address various combinations of these frictions in the context of a large international import market. Specifically, we randomly allocated 1,862 small garment firms in Dakar across treatment arms that connected them to new suppliers in Turkey and varied the information available about the types and incentives of these suppliers. We then measured how these interventions affected their access to foreign goods, supplier relationships, and profits and sales, using data from a mystery shopping exercise, real-time transactions from the largest mobile payments provider in Senegal, and a follow-up survey.

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<sup>1</sup>In a survey across six African countries, a large non-profit found that among firms that use some e-commerce, 60% use social media exclusively (GSMA, 2023). Statista estimate that total B2C e-commerce revenue across Africa increased from 18 billion USD in 2019 to 34 billion USD in 2024 (Statista, 2023).

We first provide novel descriptive evidence from our baseline survey showing that firms are using social media extensively in ways consistent with it alleviating search and trust frictions. For search, we document that 86% of firms are in WhatsApp groups managed by suppliers (both domestic and international), with the median firm in four. These groups help firms see what suppliers abroad sell, but it is unclear how much of the friction it alleviates as firms must still find the suppliers—and thus join the groups—in the first place. For trust, one-quarter of firms are in WhatsApp groups with other firms for the purpose of sharing business information, and two-thirds of firms have recommended or warned against particular suppliers to other firms in the past year. However, these networks are often highly fragmented, so it is unclear how relevant the information that they provide is to the average firm. While we focus on WhatsApp, other apps, such as TikTok, Instagram, and Facebook, are also popular, each used by around a third of firms. This descriptive evidence motivates an experiment that leverages social media to create exogenous variation in search and trust frictions in international trade.

To make precise what we mean by search and trust frictions, to understand how the frictions interact, and to discuss what types of variation might separate them, we develop a model of relational contracting featuring sequential search for suppliers and both adverse selection and moral hazard, which we refer to as trust frictions. A firm can either buy inputs from a local supplier without frictions, or can pay a fixed search cost to match with a random foreign supplier. Foreign suppliers may be thought of as selling newer varieties, higher quality varieties, or the same varieties at a lower price. Foreign suppliers must take a costly but unobservable action to ensure that the goods are high quality. Bad-type suppliers will never do this, while good-type suppliers will only do so if the future value of the relationship exceeds the current-period cost. We characterize the optimal contract and derive an equation showing that adverse selection and moral hazard create wedges between marginal revenue and marginal cost that distort quantity downwards. Moreover, these trust frictions lower the initial value of a relationship and thus the return to searching.

The model highlights the types of variation that the ideal experiment aiming to isolate the three frictions would generate. To isolate search frictions, treatments should either create matches or lower the cost of finding a random new supplier. To isolate adverse selection, treatments should either directly give information about a particular supplier or improve the ability to learn this information over time, but should not affect suppliers' incentives. To isolate moral hazard, treatments should strengthen the incentives of the supplier or firms' perceptions of these incentives, but should not provide other information.

The experiment comprises three treatments. Each treatment is designed to be of real-world interest in its own right, while also mapping onto one of the three specific frictions (search,

adverse selection, and moral hazard). In the search treatment, we add treated firms to the supplier WhatsApp groups of three different suppliers in Turkey. We inform firms that the suppliers were recruited by a local team in Turkey and export to Senegal, but do not provide any further details.<sup>2</sup> Thus, the treatment creates new matches.

We then cross-randomise the adverse selection and moral hazard treatments among the firms in the search treatment. In the adverse selection treatment, we add treated firms to a fourth WhatsApp group containing other firms matched with the same suppliers, the purpose of which is for them to privately share information about whether these suppliers are good or bad. Importantly, we seed these groups with initial information: treated firms receive a recommendation for one of the suppliers, based on mystery orders that we commissioned prior to the study. Thus, the treatment directly provides (positive) information about suppliers' types and improves learning.

In the moral hazard treatment, we inform firms that we will ask them to rate the study suppliers, and that any supplier receiving consistently negative feedback will be removed from the study, thereby losing access to 150-200 potential clients. We emphasize that we have communicated this information to the suppliers and that they therefore have strong incentives to exert effort. Thus, the treatment shifts firms' perceptions about the incentives of the suppliers.

Altogether, we have five equally sized groups: Pure Control, Search Only, Search + Adverse Selection, Search + Moral Hazard, and Search + Adverse Selection + Moral Hazard.

Our primary outcome is a revealed preference measure of access to foreign goods. We designed a mystery shopping exercise in which trained surveyors, acting as real customers, attempt to buy goods from all firms. We then measured the type and quality of the goods that they sold to us. This has two advantages: it captures real behaviour (as opposed to self-reported outcomes), and it allows us to separately measure a horizontal dimension (access to a wide set of differentiated varieties) and a vertical dimension (access to high quality varieties). On the horizontal dimension, each good that we attempt to buy is defined by 5 criteria, such as colour and sleeve style, and our outcome is an indicator for whether the firm has a good matching at least 3 criteria. On the vertical dimension, the outcome is an index that aggregates three measures: two based on a detailed quality scorecard that we designed together with hired experts, and one based on whether the good was made in Turkey (a strong signal of quality in this setting). We pre-specified these outcomes and the regression specification that we use throughout the paper in our Pre-Analysis Plan (PAP).

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<sup>2</sup>We chose Turkey as the exporter country because it is the second largest source of ready-to-wear garments in Senegal (after China), and in this setting Turkish-made garments command a large quality premium, which is well-suited for studying trust frictions because firms are worried about suppliers cheating on quality.

We find that the treatments have a large and significant effect on access to foreign goods, on both dimensions. First, we pool the four treated groups together (all of whom received the search treatment). On the horizontal dimension, treated firms are 9.5 percentage points ( $p < 0.01$ ) more likely to have a suitable good (a 26.6% increase). On the vertical dimension, conditional on having a good, the index increases by 0.412 standard deviations ( $p < 0.01$ ). The effect on the price is positive, but small, insignificant, and precisely estimated, so the horizontal and vertical gains do not come at the cost of a large price increase.

Second, when we disaggregate across the four treated groups, we find that the coefficients for the trust-treated groups are not significantly larger than the coefficients for Search Only. This does not necessarily mean that trust frictions do not exist: these are small orders (around 20-40 USD) and may be below the threshold at which trust frictions have bite. We thus conclude that (1) firms face constraints in accessing foreign goods, (2) alleviating search frictions improves this access, (3) social media can be an effective tool to do so.

While alleviating search frictions by connecting firms to new suppliers via social media improves access to foreign goods, the extent to which firms are able to realise this benefit depends on whether these connections develop into lasting relationships. We measure this using data from two sources: (1) a follow-up survey that we conducted after 3 months, and (2) real-time, transaction-level administrative data from the largest mobile money provider in Senegal, tracking a large share of transactions for up to 9 months after the study started.

From the survey data, we find that (pooled) treatment increases the likelihood of having a regular supplier in Turkey by 3.7 percentage points ( $p = 0.083$ ), a 22.2% increase relative to control. Here, disaggregating the treatments matters: the effect comes primarily from the groups with the trust treatments, and is largest in the Search + Adverse Selection + Moral Hazard group (the group with both trust treatments) at 7.5 percentage points ( $p < 0.01$ ). We find no effect on the total number of suppliers, suggesting that firms have substituted away from a local wholesaler and towards importing directly. From the mobile money data, we find that all treated firms order similar amounts from study suppliers in the first few months. However, this changes in the medium- to long-run: firms in the trust treatments order 311.4% ( $p = 0.040$ ) more than Search Only in the 6 months after the study finishes. When we disaggregate this effect into the separate trust treatments, the coefficients are all positive and fairly similar. We thus conclude that the trust treatments meaningfully increased the share of these connections that developed into relationships.

To understand whether alleviating these frictions ultimately flows through to producer surplus, we collect standard summary measures of monthly profit and sales in our follow-up survey. Pooling the treatments together, we find increases of 82.4 USD ( $p = 0.026$ ) in

profits and 245.2 USD ( $p = 0.050$ ) in sales. These are 43.8% and 40.2% increases relative to control. The coefficient in the Search + Adverse Selection + Moral Hazard is, again, much larger than the others. The implied increases are very large, and the magnitudes reduce by around half (but remain significant) when we winsorize at the 1% level. When we look at distributional treatment effects, we find that these average results come primarily from the upper tail of the profit and sales distributions: we see large and significant increases starting at around the 75th percentile for the Search + Adverse Selection + Moral Hazard group. Overall, we conclude that firms are able to realise meaningful gains from accessing a new foreign supplier by using social media to overcome search and trust frictions.

Finally, we use the reduced-form treatment effects to estimate the parameters of the model and the distribution of match-specific productivity between firms and foreign suppliers. We find that the cost of finding a given foreign supplier is relatively low, but the distribution of match-specific productivity implies that firms would have to pay the search cost many times to find a good match. We consider a first counterfactual where we shut down the ability to ever purchase directly from foreign suppliers—for example, if these firms could not use technological tools to interact with foreign suppliers—and find that this would decrease profits by around one-quarter. We then consider a second counterfactual where we progressively turn off the various frictions. The largest gains come from eliminating adverse selection. Since we estimate the share of bad types to be in the order of two-thirds, removing this friction causes many more firms to engage in search. The reason that eliminating moral hazard has a smaller effect is twofold: (1) the severity of adverse selection means firms are already choosing low quantity, (2) the total surplus in most relationships is sufficient to incentivize the supplier without substantial distortions.

Overall, our results show that both search and trust frictions meaningfully limit the ability of small firms to buy inputs from foreign markets, and that social media can indeed be used to partially alleviate them. This provides new evidence on the nature of information frictions in international trade, and also suggests that the rapidly evolving digital landscape may change how firms find, learn about, and develop relationships with suppliers.

This paper builds on several literatures. First, while there is a substantial theoretical tradition studying search and trust frictions, making empirical progress has proved more challenging. This is largely due to challenges in (1) observing buyer-seller relationships, and (2) obtaining exogenous variation that isolates theoretical forces. Recently, a growing literature has begun to overcome some of these challenges, primarily on the data side (e.g., Antras and Foley (2015), Steinwender (2018), Macchiavello and Morjaria (2015, 2021), Startz (2024)). We contribute in two ways. First, this paper is the first experiment systematically testing theories of search and trust frictions in buyer-seller relationships. This allows

us to overcome both challenges: we observe buyer-seller relationships through survey and mobile money data, and we create variation that is both exogenous and specifically designed to capture theoretical moments. Second, our descriptive evidence and experimental results highlight the role that new technologies can play in alleviating these frictions.

Second, we contribute to the literature on how e-commerce improves consumer and producer surplus by alleviating information frictions.<sup>3</sup> The bulk of this literature focuses on formal B2B or B2C e-commerce platforms, typically in higher-income settings. Our contribution is to bring both descriptive and experimental evidence that shows the large and distinct role that digital platforms—and social media in particular—are now playing in lower-income settings. In particular, we show that firms use social media extensively for buying and selling decisions, and that rather than alleviating the frictions by providing centralised repositories of information, social media enables and facilitates bilateral relational contracting.

Third, we contribute to a literature in international trade that emphasises how networks have historically played a large role in solving information frictions and enforcing contracts. For example, Greif (1993) highlights how 11th-century Maghribi traders were able to sustain a multilateral punishment system for overseas agents, and that informal information flows through social networks played a key role. The classic article testing these theories empirically is Rauch (1999), which shows that common language and proximity play a larger role in explaining trade in differentiated goods than homogeneous goods, consistent with a large role of shared ties for the types of goods subject to search and trust frictions. While our study is not about networks *per se*, one of the main channels through which social media may alleviate trust frictions is exactly the mechanism in Greif (1993). Our study thus highlights how social media facilitates a modern manifestation of this idea.

The rest of the paper is organised as follows. Section 2 presents the setting and descriptive evidence on how firms use social media in supply chains. Section 3 describes the model. Section 4 describes the experimental design. Section 5 describes the data and methods. Section 6 presents the results. Section 7 presents the model estimation. Section 8 concludes.

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<sup>3</sup>See Ellison and Ellison (2009), Dinerstein, Einav, Levin, and Sundaresan (2018), Barach, Golden, and Horton (2020), Couture et al. (2021), Bai et al. (2023), or Goldfarb and Tucker (2019) for a recent review. The closest paper to ours is Alhorr (2024), who provides Facebook pages to female business owners in Jordan and finds positive effects on business outcomes, particularly for women who are less able to leave the house. Our paper differs as our focus is on how these technologies alleviate search and trust frictions in international trade.

## 2 Setting

### 2.1 Ready-to-Wear Garments in Dakar

Our study focuses on the ready-to-wear garment industry in Dakar, the capital city and economic hub of Senegal. The ready-to-wear garment industry exhibits substantial horizontal and vertical differentiation, making it ideal for our study: horizontal differentiation (a wide range of varieties) is well-suited for studying search frictions, and vertical differentiation (the presence of high and low qualities) is well-suited for studying trust frictions. It is also a large and important industry in its own right: in a consumer survey that we conducted with 400 households in Dakar, ready-to-wear garments represented an average of 6% of total household expenditure.

Within the ready-to-wear garment industry, our study places particular emphasis on goods made in Turkey. We chose Turkish-made goods for two reasons. First, Turkey is the second largest source of ready-to-wear garments in Dakar (after China). Second, Turkish-made goods have a reputation for being higher quality than goods made in China, which is ideal for studying trust frictions. In this setting, highlighting that a good is “Made in Turkey” is a very common way to signal quality. To quantify this, in our consumer survey we showed households an image of a product and randomised whether we said the good was made in Turkey or made in China. We then asked for their willingness to pay. We plot the CDF of willingness to pay in Figure 4, Panel (a). The Turkey CDF is shifted uniformly rightward relative to the China CDF, with an average premium of 34% ( $p < 0.01$ ).

### 2.2 Sample

**Firms in Senegal** The main subjects of the study are 1,862 small firms in the ready-to-wear garments industry in Dakar. These firms are typical of small, informal, owner-operated businesses in many large cities in lower- and middle-income countries. 33% have a physical store in a market, while the remaining 67% operate exclusively online, primarily through social media. Firms with a physical store were recruited through a census in selected markets that sell both high and low quality goods; firms without a physical store were recruited through a combination of advertisements on Facebook and snowball sampling. The firms with a physical store are therefore broadly representative; recent years have seen substantial growth in the number of firms operating online-only businesses and those in our sample are typical of this phenomenon, but as there is no systematic database of such firms we cannot formally assess their representativeness.

At baseline, 91% of firms sell Turkish-made goods, with Turkish-made goods representing 40% of sales for the median firm. 33% of firms sell wholesale, which is an important margin

of heterogeneity that we pre-specified in our PAP. 7% of retailers and 15% of wholesalers have travelled internationally for business at least once in the past 5 years, so the majority of these firms purchase their goods from other firms in Dakar or through e-commerce. 18% of retailers and 28% of wholesalers have at least one regular supplier based in Turkey. As we show in Section 2.3, these firms ubiquitously use social media, and WhatsApp in particular, to receive information from suppliers.

Firms have pessimistic beliefs about unknown foreign suppliers. 60% know multiple other firms that have had bad experiences ordering from a supplier online. To measure firms' priors, we asked them to consider a scenario in which they made an order from 10 unknown foreign suppliers, and to opine on what share of these orders would arrive with the anticipated quality. The median firm's response was that this would happen only 50% of the time.

**Suppliers in Turkey** The study involves connecting firms in Senegal with suppliers in Turkey. We worked with 30 suppliers, all of whom were based in Istanbul and were exporters of ready-to-wear garments to West Africa. We traveled to Turkey and conducted a census of two quarters in Istanbul that are well known for being textile wholesale and export hub for many parts of the world, including to several countries in West Africa. Among the suppliers that met our inclusion criteria, we then conducted a mystery shopping exercise to identify the most active. We focused on suppliers of Senegalese nationality, for a few reasons. First, while many Istanbul-based exporters to West Africa are of Turkish nationality, there is a sizeable Senegalese diaspora operating in this industry, reflecting the evidence in Rauch (2001) and Greif (1993) that cultural ties have historically played important roles in alleviating search and contracting frictions in long-distance trade. These shared ties certainly do not reduce trust frictions to zero, but they do likely facilitate reputation-based mechanisms. Second, they also solve some basic but important frictions such as language barriers and payment technologies. Since we expected our 1,862 firms in Senegal to be highly heterogeneous, introducing these barriers would have substantially complicated the experiment and distracted from the core issues of search and trust. Third, one of the main data sources for our outcomes is transaction-level data from the largest mobile money provider in Senegal. Senegalese suppliers in Turkey typically have accounts and use this system to accept payments, and so focusing on them offers substantial measurement advantages.

## 2.3 Social Media and e-Commerce in Supply Chains

In this section, we first describe the main channel that we use for the study and then we present statistics on social media usage and formal e-commerce platform usage.

**Supplier Groups** Many suppliers in our setting—both domestically and internationally—operate WhatsApp groups with their clients to advertise their goods, post prices, and highlight new items in stock. We will regularly refer to these as “supplier groups.” We show examples in Figure 2. A typical group has one supplier and 50-100 clients, most of whom are regular or repeat customers. These are not discussion groups: the purpose is for the supplier to regularly post high quality photos and videos of their goods (typically only the supplier has permission to post). It is of course possible and common for buyers to negotiate with the supplier or to inquire about other goods—they can do so by simply sending a private message. These groups may be usefully thought of as virtual storefronts: clients can see what the supplier is selling, and can talk directly to the supplier about any queries.

These groups play a potentially important role in reducing search frictions and, to some extent, in reducing trust frictions. For search, firms can observe a very large number of goods from all over the world directly on their phone, and can easily negotiate and follow up as needed. Importantly, most firms also use social media extensively to sell to their own customers, and so these groups make it easy for them to forward relevant images to their own clients. For trust, a large group with many clients raises the cost of cheating because cheated buyers can message other members to share information. The group also makes it easier for the supplier to build a brand, further improving reputation-based mechanisms.

**Social Media** In Figure 3, we present statistics from our baseline survey with 1,862 firms in the ready-to-wear garment industry in Dakar. We focus here on sample-wide averages, but we also show in Appendix Figure A1 that the results are almost identical among firms with and without a physical store. In Panel (a), we plot the share of firms that reported using different types of social media to obtain information about suppliers, such as learning about new varieties or price information. WhatsApp is ubiquitous: 92% of firms use WhatsApp Status (a feature not often used in the United States in which content is broadcasted to all contacts for 24 hours), and 86% use supplier groups. TikTok, Instagram, and Facebook are also popular, each used by about a third of firms. Panel (b) shows the distribution of the number of unique supplier WhatsApp groups that firms belong to. Firms are in many supplier groups, with almost half of firms in 5 or more. Since these groups are very active, being in 5 such groups means that firms are observing a lot of information about different suppliers all the time. Importantly, these are not simply groups that

they belong to but ignore: the distribution of the number of groups that they have bought inputs from in the past 12 months is almost identical.

To understand the reason such groups are so widely used, in Panel (c) we present the responses to a question asking what the main search-based advantages of supplier groups are. Firms highlight both how it allows them to see more varieties (both a wider set and higher quality) and how it allows them to compare prices across suppliers. Finally, in Panel (d), we show the location of the suppliers running these groups. The majority (81%) of firms are in a group with a supplier in Senegal, while a large minority are in at least one group with a supplier in a foreign country. 21% are in a group with a supplier in Turkey, 12% are in a group with a supplier in China, and 6% are in a group with a supplier in Dubai.<sup>4</sup> In total, 27% are in at least one supplier WhatsApp group where the supplier is based abroad. Since this focuses only on WhatsApp groups, this is a lower bound on the share of firms using social media more generally for international trade.

Firms are therefore familiar with the concept of using supplier WhatsApp group to transact with foreign suppliers, but, since only 21% are in a group with a supplier in Turkey, our experiment is still able to generate meaningful variation.

**e-Commerce Platforms** Traditional B2B e-commerce platforms, such as Alibaba, have also been shown to alleviate search and trust frictions (see Goldfarb and Tucker, 2019 for a review). Yet, in this setting, they are seldom used: 85% of firms have never purchased from these platforms, and, of those who have, about half have done so only very rarely. This is not because they have not heard of them (88% have). This reflects a broader trend in which large e-commerce companies have had limited success at penetrating African markets. On the surface, it may therefore seem puzzling why social media should be so heavily used in place of formal B2B platforms. In our follow-up survey, we asked firms why they don't use B2B platforms. Aside from the 39% firms who give no particular reason, the two most common answers are that firms find them too complicated to use (40%) and that firms do not trust them (33%).

These statistics confirm the anecdotal observations that led us to run this study: that small firms use social media extensively for their buying and selling activities, and that a sizable share of firms use it as a means of doing international trade.

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<sup>4</sup>The share in China is likely a large underestimate of total social media interactions with China, as WhatsApp is blocked by China's firewall (it is useable with a VPN) and so other social media platforms—such as WeChat—are much more common.

### 3 Theory

In this section, we describe a model of relational contracting featuring adverse selection and moral hazard with on-path learning, embedded within a sequential search framework. The goal is to make precise the role of the frictions and to highlight the types of variation necessary to identify them. In Section 7, we will use the reduced form treatment effects to estimate the parameters governing these frictions and consider counterfactuals in which we set them to zero to estimate the total potential gains. While the forces in the model are all canonical, we are not aware of existing literature that combines them in this way.<sup>5</sup>

#### 3.1 Firms and Suppliers

The model is an infinite horizon repeated game with discrete time, all players have common discount factor  $\delta$ , and all players are risk neutral. There are two sets of players: firms (the principal) and foreign suppliers (the agent). In the stage game, the firm can purchase input  $q$  and re-sell to consumers for revenue  $r(q)$ , where  $r(q)$  is strictly increasing, concave, and has  $\lim_{q \rightarrow \infty} r'(q) = 0$ . The firm has outside option  $\bar{U}$ . The firm is not matched with a foreign supplier by default, and matching will be governed by a search process. We first discuss the principal-agent problem that arises conditional on matching, and then describe the search process and the determination of  $\bar{U}$ .

#### 3.2 Relational Contracting with Foreign Suppliers

Goods sold by foreign suppliers can be high or low quality. For simplicity, we normalise the value of low quality goods to zero, and let the value of  $q_t$  high quality goods to the firm be  $r(q_t)$ , as before. Faced with an order for  $q_t$ , foreign suppliers can choose to take a costly but unobservable action  $a_t \in \{0, 1\}$  that influences the probability that the goods are high quality. If they choose  $a_t = 1$ , then the goods are high quality with probability 1 and the supplier pays cost  $cq_t$ , where  $c > 0$  is a constant marginal cost. If they choose  $a_t = 0$ , they instead pay a lower cost  $(1 - \xi)cq_t$  with  $\xi \in (0, 1)$  but the goods are only high quality with probability  $\lambda \in (0, 1)$ . Avoiding the action and cutting a share  $\xi$  of the cost may be interpreted as purchasing the goods from a cheaper manufacturer that only delivers with

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<sup>5</sup>Our model differs from much of the relational contracts literature as the solution is non-stationary, due to both on-path learning and limited liability. Among prior work in this literature that studies non-stationary equilibria (Hörner (2002), Halac (2012), Yang (2013), Fong and Li (2017)), the principal-agent component of our model combines the on-path learning from Yang (2013) with the explicit treatment of limited liability from Fong and Li (2017). However, unlike both of these papers, we endogenise quantity, which complicates incentive design as the principal can choose not only the terms of the contract but also the stakes of the contract in each period. Martimort et al. (2017) take a mechanism design approach to analyse the endogenous quantity issue, but they focus on separating equilibria (ruled out under our version of the enforcement constraint and limited liability) and thus do not feature on-path learning.

probability  $\lambda$ , or as shipping the goods with a cheaper exporting service where the goods only arrive with probability  $\lambda$ .

If the firm orders from the supplier, they pay a transfer  $\tau_t$ , determined endogenously. This transfer must be paid before receiving the good.<sup>6</sup> Foreign suppliers have outside option 0, reflecting the idea that relationships are separable for the supplier.

**Adverse Selection:** There are two types of foreign suppliers: good and bad. A supplier's type is fixed over time, known to the supplier, and unobservable to firms. The only difference between the two types is that bad types will never choose  $a_t = 1$  (for example, because they are unskilled and thus unable to use this productive technology), while good types will choose  $a_t = 1$  if it is in their best interests to do so. We denote the firm's beliefs about the share of bad types after observing  $t$  realisations in which the goods were high quality as  $\mu_t$ , with initial (correct) beliefs  $\mu_0 \in (0, 1)$ . The firm updates this belief each period using Bayes' Rule. We assume that  $\lambda$  is sufficiently low that, if the firm knew that the supplier was a bad type, they would prefer to exit the relationship and take their outside option,  $\bar{U}$ .

**Moral Hazard:** Good types will choose  $a_t = 1$  if it is in their best interests to do so. Since this is a repeated game, it may be possible to induce them to do this by the promise of future rewards tied to repeated high quality realisations. In particular, any equilibrium in grim trigger strategies in which good type suppliers choose  $a_t = 1$  must satisfy the following standard Dynamic Incentive Compatibility Constraint (DICC),

$$\delta(1 - \lambda)V_{t+1} \geq \xi c q_t, \quad (\text{DICC})$$

where  $V_{t+1} \equiv \sum_{n=1}^{\infty} \delta^n (\tau_{t+1+n} - c q_{t+1+n})$  is the discounted sum of future profits from the relationship. Intuitively, the supplier has already received  $\tau_t$ . If they choose  $a_t = 1$ , they earn  $\delta V_{t+1} - c q_t$ ; if they choose a one-shot deviation to  $a_t = 0$ , they earn  $\lambda \delta V_{t+1} - (1 - \xi) c q_t$ . Choosing  $a_t = 1$  thus increases the probability of maintaining the relationship by  $1 - \lambda$ , but comes at the additional cost of  $\xi c q_t$ . This constraint imposes a (time-varying) ceiling on the contractible quantity. Note that the constraint will have more bite for larger  $\xi$ , and will hold trivially for  $\xi$  near 0.

**Contracts:** In period 0, the firm makes a take-it-or-leave-it offer of a long-term contract to the supplier, which consists of a sequence  $\{q_t, \tau_t\}_0^{\infty}$ . The contract is relational, meaning that it is not enforceable in court: both parties can agree on a long-term plan, but are unable

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<sup>6</sup>One could extend the model to allow for an endogenous share  $\alpha_t$  to be paid upfront, and the remaining  $(1 - \alpha_t)$  to be paid after quality is observed. In our empirical setting, almost 100% of contracts involve full payment upfront, so for simplicity we simply set  $\alpha_t = 1$  for all  $t$  for simplicity.

to commit to it. This implies that both parties face an enforcement constraint that requires that in every period  $t$  it must be optimal for them to take the contracted action. For the foreign supplier, this is already implied by their limited liability constraint (see below), so this is only relevant for the firm, for whom it implies the additional constraint  $U_t \geq \bar{U}$  for all  $t$ . The enforcement constraint prevents the firm from (for example) promising large bonus payments to the supplier in the distant future, because when the distant future arrives the firm will prefer to renege on these contracted payments and instead simply take their outside option. We focus on trigger strategies in which both parties take their outside option if any player has ever deviated from the contract, because this provides maximal incentives.

**Limited Liability:** The contract must satisfy  $\tau_t \geq c q_t$  for all  $t$ ; that is, the supplier must make weakly positive profits period-by-period. We impose this restriction to prevent large period-0 rent extractions that we rarely see in our empirical setting. Instead, in our discussions with suppliers, we often heard of selling at cost early in the relationship, which this restriction permits (and which will occur in equilibrium). Note that this constraint means we can ignore the supplier's participation constraint and enforcement constraint.

**Equilibrium:** We focus on Perfect Public Equilibria in which the good type supplier chooses  $a_t = 1$ . We also restrict attention to pooling equilibria, which we show in Proposition 1 in Appendix B is without loss. Intuitively, the inability to commit combined with the parameter restriction implying that the bad type is inefficient ensures that the firm will immediately terminate the contract if they ever learn that the supplier is a bad type. The bad type thus always earns 0 upon revealing their type, and limited liability combined with the DICC ensures they can always earn a strictly positive expected payoff by mimicking the good type.<sup>7</sup>

Before stating the full dynamic program, it is convenient to denote the firm's expected stage payoff as  $y(q_t, \tau_t, \mu_t) \equiv (1 - \mu_t(1 - \lambda))r(q_t) - \tau_t$ , and the supplier's stage payoff as  $\pi(q_t, \tau_t) \equiv \tau_t - c q_t$ .<sup>8</sup> We can then write the continuation values of the firm,  $U_t$ , and the

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<sup>7</sup>This result obviously relies on the parameter restriction that ensures that the firm would never voluntarily trade with bad types. However, this is not important for either the intuition or the bulk of the results: the separating equilibrium that may obtain in absence of this restriction is qualitatively very similar to the pooling equilibrium. This is because dynamic moral hazard is sufficient to generate an upward sloping quantity profile due to the well-known backloading intuition. Nonetheless, we impose the restriction here because it seems closer in spirit to our empirical setting, in which the majority of firms take their outside option (purchasing from a local wholesaler).

<sup>8</sup>Bad type suppliers behave entirely mechanically, so we do not write out their payoff and in general whenever we refer to the payoff or decision of a foreign supplier it is implied that it is a good type.

supplier,  $V_t$ , recursively as

$$U_t \equiv y(q_t, \tau_t, \mu_t) + \delta \left[ (1 - \mu_t(1 - \lambda))U_{t+1} + \mu_t(1 - \lambda)\bar{U} \right], \\ V_t \equiv \pi(q_t, \tau_t) + \delta V_{t+1}.$$

Standard arguments in the dynamic moral hazard and endogenously incomplete markets literatures establish that it is equivalent—and substantially easier—to solve this problem by explicitly letting the firm choose these continuation values and encoding them as state variables.<sup>9</sup> This implies that the firm solves the following recursive dynamic program in each  $t$ :

$$W_t(U_t, V_t, \mu_t) = \max_{q_t, \tau_t, U_{t+1}, V_{t+1}} y(q_t, \tau_t, \mu_t) + \delta(1 - \mu_t(1 - \lambda))W_{t+1}(U_{t+1}, V_{t+1}, \mu_{t+1})$$

subject to the following constraints:

$$\begin{aligned} \delta(1 - \lambda)V_{t+1} &\geq \xi c q_t & (DICC) \\ U_{t+1} &\geq \bar{U} & (DEC) \\ \tau_t &\geq c q_t & (LL) \\ y(q_t, \tau_t, \mu_t) + \delta \left[ (1 - \mu_t(1 - \lambda))U_{t+1} + \mu_t(1 - \lambda)\bar{U} \right] &\geq U_t & (PK_f) \\ \tau_t - c q_t + \delta V_{t+1} &\geq V_t, & (PK_s) \end{aligned}$$

with  $\mu_{t+1} = \mu_t \lambda / (1 - \mu_t + \mu_t \lambda) < \mu_t$  if the good is high quality in  $t$  and  $\mu_{t+1} = 1$  if the good is low quality in  $t$ , and with  $U_0 = \bar{U}, V_0 = 0, \mu_0$  given. We have already introduced the first three constraints. The final two constraints are known as promise-keeping constraints as they ensure that the continuation utilities promised in the previous period,  $U_t$  and  $V_t$ , are actually delivered through a combination of stage payoffs and future promises.

The solution to this program is generally not available in closed form. In Proposition 2 in Appendix B, we provide a detailed derivation of several properties of the solution, which we summarise briefly here. The optimal contract looks similar to a dynamic version of the typical “sell the firm to the agent” solution in static models without risk aversion. In particular, there exists some finite  $T^*$  such that the supplier will earn zero stage profits for periods  $t = 0, 1, 2, \dots, T^* - 1$  and then earn the entire surplus (net of the firm’s outside option) for all  $t > T^*$ . Intuitively, the firm makes the supplier the residual claimant for most of the relationship—which is the most efficient way to provide incentives—and ex-

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<sup>9</sup>This approach was originally developed somewhat independently in different theoretical contexts by Spear and Srivastava (1987), Abreu, Pearce, and Stacchetti (1990), and Thomas and Worrall (1988). Golosov, Tsivinski, and Werquin (2016) provide an excellent review in the context of incomplete markets models.

tracts surplus in the early periods as these minimise incentive distortions. This is stark, but not unreasonable: the supplier sells at cost at the beginning of the relationship while its reputation is being established, and reaps the benefits of its reputation later on. Moreover, backloading of incentives of this form is a very general prediction of models of reputation and dynamic moral hazard (Shapiro (1983), Banerjee and Duflo (2000)).

We now state two results relevant to intuition and the interpretation of the experiment.

**Result 1** (Intensive Margin). *Quantity and value purchased from the foreign supplier are distorted downwards by both adverse selection and moral hazard.*

The FOCs of this program yield the following equation relating marginal revenue and marginal cost,

$$r'(q_t^*) = \underbrace{\frac{1}{1 - \mu_t + \mu_t \lambda}}_{\text{Adverse Selection}} \underbrace{(1 + \xi \rho_t^*)}_{\text{Moral Hazard}} c,$$

where  $\rho_t^*$  is a weakly positive function of Lagrange multipliers. Both wedge terms are weakly greater than 1, and so distort  $q_t^*$  downwards relative to the level that equates marginal revenue and marginal cost. We prove in Proposition 2 in Appendix B that this distortion will decrease over time, both because of learning ( $\mu_t$  decreasing) and because the optimal contract involves the joint surplus increasing over time to incentivise the agent. A similar equation can be derived for value,  $\tau_t$ . Treatments that alleviate these frictions should therefore increase quantity and value ordered, at least among firms that would have imported directly anyway.

**Result 2** (Extensive Margin). *The period-0 value of the relationship is decreasing in  $\mu_0$  and  $\xi$ .*

This follows from a straightforward application of the Envelope Theorem, and highlights that the extent of adverse selection and moral hazard will limit the possible gains from trade, perhaps to the point where no trade occurs. Treatments that alleviate these frictions should therefore increase the propensity of firms to import directly.

### 3.3 Search

Matching with a foreign supplier is costly. A firm can pay a one-time search cost  $s > 0$  to match with a random foreign supplier. Upon matching, the firm immediately observes a realisation of a match-specific productivity term,  $\psi \sim G$ , that is fixed over time, implemented by replacing  $r(q_t)$  with  $r(\psi q_t)$ . We will denote the period-0 value of a relationship with a foreign supplier with match-specific productivity  $\psi$  as  $U_0(\mu_0, \xi, \psi)$ . Search is sequential, meaning that if a firm's current best option delivers discounted utility  $\tilde{U}$ , then

the firm will search if  $E_\psi[\max\{U_0(\mu_0, \xi, \psi), \tilde{U}\}] - s \geq \tilde{U}$ . Standard arguments then imply that there exists a cutoff value  $\bar{U}$  such that the firm will search if and only if their current best option is less than  $\bar{U}$ , and that their expected return to doing so is exactly  $\bar{U}$ .<sup>10</sup> This reservation value thus defines their outside option.

Finally, all firms can purchase instead from a local supplier, the cost of which is not heterogeneous and does not involve any frictions. If the value of purchasing from the local supplier is lower than  $\bar{U}$ , then they will search; otherwise, they will remain with their local supplier.

### 3.4 Implications for the Experiment

The model highlights the three frictions that we will study in the experiment: search, adverse selection, and moral hazard. Both adverse selection and moral hazard reduce the intensive and extensive margins of transacting with foreign suppliers, while the search friction reduces the extensive margin only. To relax the search friction, we need a treatment that either lowers the cost of matching with foreign suppliers or improves the likelihood of finding a suitable match. To relax adverse selection, we need a treatment that improves beliefs (or improves the ability to learn) about a supplier that a firm has been matched with. To relax moral hazard, we need a treatment that either improves the supplier's incentives directly or changes firms' perceptions about the cost to the supplier of not honouring the contract.

Importantly, the ideal experiment targeting adverse selection should have no effect if  $\mu_0 \approx 0$ ; they should not affect the relationship if there is only one type of supplier. Similarly, the ideal experiment targeting moral hazard should have no effect if  $\xi \approx 0$ ; they should not affect the relationship if the strategic type always chooses  $a_t = 1$ . We describe our experimental design that aims to achieve this in the next section.

## 4 Experimental Design

### 4.1 Treatment Conditions

The goal of the experiment is to generate variation that identifies the three frictions: search, adverse selection, and moral hazard (where we refer to the latter two jointly as trust frictions).

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<sup>10</sup>The cutoff value is defined implicitly by equating marginal benefit of searching with marginal cost, that is,  $\int_{\bar{\psi}(\bar{U})}^{\infty} (U_0(\mu_0, \xi, \psi) - \bar{U}) f(\psi) d\psi = s$ , where  $\bar{\psi}(\bar{U})$  is the value of  $\psi$  for which  $U_0(\mu_0, \xi, \psi) = \bar{U}$ . This cutoff  $\bar{U}$  is decreasing in  $s$ .

**Search** 80% of firms receive the Search treatment. The purpose of this treatment is to generate exogenous variation in the cost of finding a supplier of Turkish-made goods. We add treated firms to the supplier WhatsApp groups of 3 different suppliers. The suppliers to match with are selected at random, subject to being a match to the merchant's chosen sector. We do not give firms any information about the suppliers, except to say that they were recruited by a team in Turkey in a similar manner to how the firm itself was recruited. We communicate to the control group that unfortunately we cannot add them to any supplier groups at this time, but that we might do so at the conclusion of the study.

**Adverse Selection** 50% of firms in the Search treatment condition are treated with the Adverse Selection treatment. As the model highlights, identifying adverse selection requires either improving the ability to learn or providing information directly. This treatment consists of doing both. We add firms in this treatment condition to a fourth WhatsApp group. This group does not contain any suppliers, but instead contains other firms in the study that were matched with the same suppliers. We explain that all members of this fourth group have been matched with the same three suppliers and that the purpose of the group is to share information about them. A member of the study team moderates discussion and encourages firms to share information from time to time. We do not tell suppliers about the existence of these groups, so only the information of the firms—and not the incentives of the suppliers—are affected.<sup>11</sup>

Since no firms have experience with the supplier at this point, we seed the groups with initial information. Treated firms receive a phone call 2-3 days after recruitment from a recommender.<sup>12</sup> The recommender is part of a team of firms—who are not subjects in the study—that we hired prior to the study to make mystery orders from all of the suppliers. The recommender describes their experience ordering from one of the suppliers that the firm was matched with and sends a photo of the item that they ordered. They also explain that they, too, are in the information-sharing WhatsApp group, and post a similar message there.

Despite the fact that treated firms did not know the recommender personally, they generally took this information seriously for two reasons. First, at the end of the baseline survey, we ask all firms if they would be willing to call a few other firms to discuss their experiences working with the study suppliers. They are therefore not surprised when they

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<sup>11</sup>We do not think that firms would expect to be able to credibly communicate the role of these groups to suppliers to try to improve incentives as the firm would have no easy way to convince the supplier that this is not cheap talk, especially since we make clear that we have not told the suppliers about these groups.

<sup>12</sup>All firms not in this treatment condition—both the other 50% of firms treated with Search and all firms not treated with Search—instead receive a “placebo” phone call from a surveyor, asking them for their opinion about supplier WhatsApp groups in general.

receive this call. Second, one of the reasons that social media is so ubiquitously used for commerce is precisely the social nature: even if they don't know the recommender, they can ask questions and assess the preferences and knowledge of the recommender.

**Moral Hazard** 50% of firms in the Search treatment are treated with the Moral Hazard treatment, cross-randomised with the Adverse Selection treatment. As the model highlights, identifying moral hazard requires either shifting the suppliers' incentives or shifting firms' perceptions of suppliers' incentives. To do this, we read the following information to treated firms at the end of the baseline survey:

I have one last piece of information to give you. As you know, you have been added to WhatsApp groups of Senegalese suppliers in Turkey.

We work with many suppliers in our study. We want to assure you that they are motivated.

We would like to collect feedback on these suppliers so that we can recommend the best ones in the future. To do this, we will ask the merchants in the study [such as yourself] to rate your experience with the suppliers we have presented to you on a scale of 1 to 5 on product arrival and quality. These reviews help identify the best suppliers, which is beneficial to them and allows us to continue recommending them to others. They are therefore motivated.

If a supplier gets bad ratings, we will investigate and remove them from the study if they did not do a good job. They will therefore lose access to around 150 merchants if they do not do a good job.

I will give you a phone number that you can use to give your rating or report a problem.

Lastly, I want to emphasise that the suppliers are aware that they are being rated and that, if they receive bad ratings, they will be removed from the study. We can thus assure you that they are motivated.

After delivering this message, the surveyor provides a business card to the firm. The business card has a phone number to call, and prominently highlights that this number should be used to rate the suppliers and/or to signal any problems. Untreated firms receive a similar card, but without any mention of ratings or suppliers—instead saying that the phone number is for questions about the study. Both cards can be seen in Appendix Figure A2. All suppliers are told a similar message about how the ratings will work.

Note that the experiment does not randomise the incentives provided to suppliers. In-

stead, it provides high-powered incentives to all suppliers and randomises whether we tell this fact to firms. We make clear to firms not in this treatment condition that we in no way vouch for or provide guarantees about the suppliers in the study—our only role is to make connections. This treatment should not have an impact in a model of adverse selection only because it does not provide any information to the firms about the supplier—crucially, we do not share the ratings with the firms.<sup>13</sup>

**Sub-Treatments** We cross-randomised two additional sub-treatments within the pure control group. The first sub-treatment aims to test whether the binding constraint behind the lack of traditional B2B e-commerce platform usage is that firm owners do not understand how to use them. We thus provide a short training on Alibaba that covers how to install the app, how to search for products, how to contact suppliers, and how to make purchases and arrange delivery. The second sub-treatment is a placebo check for the fact that, in the Adverse Selection treatment condition, we have connected firms to each other. To ensure that results are not driven by grouping firms per se, we thus also create similar groups here, where none of the firms have been connected with any suppliers.

## 4.2 Randomisation and Balance Check

Overall, there are five equally likely groups: Pure Control, Search Only, Search + Adverse Selection, Search + Moral Hazard, and Search + Adverse Selection + Moral Hazard. We randomly assigned firms to one of these five groups, stratifying on product group (men’s clothing, women’s clothing, or shoes & bags), an indicator whether the firm has a physical store, and an indicator for whether the firm had prior direct importing experience. Any misfits, due to integer indivisibility or other issues, were unconditionally randomised across the five cells.

Since this is an RCT, treatment is independent of pre-randomisation covariates by construction, absent errors in the randomisation protocol. To check that the randomisation protocol operated as expected, we report a balance check in Appendix Table A1. The differences in means across treatment groups are all small and insignificant, and a joint test across variables has  $p$ -value 0.724. We therefore conclude that our analysis consistently estimates average treatment effects.

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<sup>13</sup>In principle, since bad types will eventually draw a low quality realisation, firms could wait and attempt to infer types by observing whether the supplier is still around after a given number of periods. We do not think this happens for two reasons. First, our observation from pilots was that firms typically decide relatively quickly whether to pursue a relationship with the study suppliers. Second, the script does not imply that the enforcement process is particularly fast.

## 5 Outcomes, Data, and Empirical Methodology

### 5.1 Data and Outcomes

**Consumer Survey** In March 2024, we conducted a 15-20 minute survey with 400 households. We use this to calculate two sets of summary statistics. First, we use it to measure the relationship between consumer willingness to pay and various important variables in our analysis. Second, we use it to calculate statistics on household clothing expenditures.

**Baseline Survey** Upon recruiting a firm into the study, between November 2023 and January 2024, we conducted a 30 minute baseline survey. The survey contained questions on their supplier relationships, social media usage, e-commerce usage, and profit and sales. For profit, we use the summary survey question from De Mel, McKenzie, and Woodruff (2009), and a similar question for sales.

**Access to Foreign Goods** To test whether reducing search and trust frictions via social media improves firms' access to foreign goods, we need a measure of access to foreign goods. We constructed a novel measure of this using a detailed mystery shopping exercise. The goal was to measure access to foreign goods on a horizontal dimension (access to more differentiated varieties) and on a vertical dimension (access to higher quality varieties), as well as any effects on price. Around two weeks after recruitment, firms are contacted over WhatsApp by a mystery shopper, played by a trained surveyor. Firms are not aware that the customer is employed by the survey team, but are expecting to be contacted by customers, as we explain to them at the end of the baseline survey that we will put them in touch with customers who often buy high-quality goods.

The mystery shopper explains that they would like to purchase a certain high quality product for an event. Each product is defined by five horizontal criteria that are largely unrelated to quality, such as colour, sleeve style, and presence of a graphic (see Appendix Figure A3 for two examples). The mystery shopper proceeds with the purchase—including asking about price and delivery—if the firm has a good with at least three of the five criteria. The primary outcome for this horizontal component, pre-specified in our PAP, is an indicator for whether the firm had a good with at least three criteria.

If the firm has such a good, the mystery shopper buys it in a random 80% of cases.<sup>14</sup> Then, once the good arrives in our office, two tailors and an expert shoemaker assess its quality

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<sup>14</sup>If the random draw indicates to not buy the good, the mystery shopper explains that they have had a change of plans. They offer a nominal payment of 2.5 USD as a gesture of gratitude for the firm's time. We piloted different ways of doing this, and found that this procedure was natural and largely avoided upset.

according to a 50-point scorecard that we developed.<sup>15</sup> To validate the quality measure, we also gave the surveyors conducting the consumer survey a subset of these goods to present and elicit willingness to pay (WTP). We show a binscatter of the relationship between the quality score and WTP in Panel (a) of Figure 5. There is a clear positive relationship, although it becomes flat in the left tail, reflecting the fact that beyond a certain point consumers simply view goods as “low quality”. In Panel (b), we classify goods as “high quality” or “low quality” (defined as whether a good is above the median quality score of its product type), and plot the CDF of consumer WTP separately for these two sets of goods. The high quality CDF is shifted rightwards of the low quality CDF, with an average premium of 35%.

The outcomes are these two measures: the high quality indicator and the raw 50-point quality score. The rationale for the binary outcome is that it is not vulnerable to a long left tail of quality scores that, as we saw in Figure 5, are not meaningful in terms of WTP.

We also attempt to infer whether the good was manufactured in Turkey. As we showed in Figure 4, there is a large premium for Turkish-made goods since it is a strong signal of quality. Thus, while the other two vertical outcomes measure quality directly, in practice quality is not fully observable to consumers and so product origin plays an important role in consumer WTP.<sup>16</sup> For most goods, we record this information from the label, and the outcome is 1 if the label says “Made in Turkey” and 0 if it says it was made elsewhere.<sup>17</sup>

Finally, as noted in our Pre-Analysis Plan, the mystery shopping exercise also had a secondary goal of providing treated firms with an opportunity to experiment with the study suppliers. Of course, nothing in the procedure makes this explicit, but, if a firm was considering making an order, then the mystery shopper reduces the risk that they will be unable to find a buyer.

**Followup Survey** We conducted a 30-minute followup survey with similar questions to the baseline survey between February and April 2024, around 3 months after a firm is recruited to the study. We successfully surveyed 1671 firms, or 90% of the sample. The followup rate is very similar and not significantly different across the four treated groups,

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<sup>15</sup>We designed this scorecard together with the these hired experts specifically for this study. Vitali (2024), who studies the relationship between consumer search costs and firm location choices in Kampala, takes a similar approach to measure the quality of garments. Although the details of the scorecards are quite different, we benefited greatly from showing her scorecard to our hired experts as an example of what we had in mind.

<sup>16</sup>We pre-specified this outcome, but did not attach it to either the horizontal or vertical dimensions. Since the consumer preference for Turkish-made goods reflects a preference for quality, it seems more fitting to include it under the vertical dimension.

<sup>17</sup>For the small share of goods for which the label does not indicate the origin, we ask the hired experts to (independently) give their opinions as to whether the good was made in Turkey (based on sewing patterns, product style, etc.), and set the outcome to 1 if they both opine that it was made in Turkey and 0 otherwise.

but is 5 percentage points higher and statistically significantly different in the pure control group. The main outcomes are questions about the number and location of the firms' suppliers, their profits and sales, and their e-commerce use.

**Mobile Money** To go beyond survey data, we also use real-time, transaction-level data from the largest mobile money provider in Senegal, Wave Mobile Money, made available for this study. This data contains the universe of transactions between the phone numbers of firms in the study and the phone numbers of study suppliers. This data complements survey-based measures and has several advantages: (1) it is dynamic, so we can see transaction profiles over time, (2) it continues long after the followup survey, (3) it is not self-reported.

While we cannot know the exact share of transactions taking place through this medium, we expect that it is relatively large, at least for retailers, for a few reasons. First, we asked the non-study firms that we hired to mystery order from all suppliers prior to the study (mentioned in Section 4.1) to record how the supplier asked them to pay, and in 100% of cases they were asked to pay with this particular mobile money provider. Second, in the baseline survey, 86% of firms reported that they often use this provider to pay suppliers when making payments at distance. We thus expect that we see most small-to-medium sized orders, but likely miss larger orders as—anecdotally—these are more likely to take place with more formal methods such as bank transfers or international transfer services (such as Western Union and Moneygram). Since wholesalers tend to make larger orders and have significantly more experience with formal methods, we expect that this dataset is more representative of retailers than wholesalers.

## 5.2 Empirical Methodology

Our primary empirical method, specified in our Pre-Analysis Plan (PAP), is to estimate the following OLS specification

$$y_i = \alpha + \sum_{j=1}^4 \beta_j T_{ji} + \delta y_{0i} + \gamma_s + \rho' X_i + \varepsilon_i, \quad (1)$$

where  $y_i$  is the outcome for firm  $i$  and  $T_{ji}$  for  $j = \{1, 2, 3, 4\}$  are indicators for treatment arms Search Only, Search + Adverse Selection, Search + Moral Hazard, and Search + Adverse Selection + Moral Hazard.  $y_{0i}$  is the outcome measured at baseline, if available.  $\gamma_s$  are stratum fixed effects.  $X_i$  are firm-level covariates, selected by Double Lasso, following

the method of Belloni, Chernozhukov, and Hansen (2014).<sup>18</sup> We also report a version of the same regression where we pool the four treated groups.

**Inference** Our primary method of inference is randomisation inference, as recommended by Athey and Imbens (2017) and Young (2019). In particular, we compute two-sided  $p$ -values for the sharp null of zero treatment effect using 5,000 permutations of the  $t$ -statistic. We report conventional formula-based heteroskedasticity-robust standard errors in parentheses, but we do not use these for inference directly. As each regression involves four main coefficients, we also calculate and report Romano-Wolf (RW) multiple-testing adjusted  $p$ -values in square brackets (Romano and Wolf, 2005). Since there are three combinations of trust treatments, we report the  $p$ -value of a joint test that all four coefficients are equal (computed by permuting the  $F$ -statistic), which will be the case if the trust treatments have no effect.

**Indexes** To account for multiple hypothesis testing across outcomes, for any table that presents more than one outcome corresponding to the same family of outcomes, we also report the results on an index that aggregates the outcomes using the standardised inverse-variance weighted method of Anderson (2008). Since the disaggregated regressions may include different covariates  $y_{0i}$  and  $X_i$ , before indexing we first residualise each outcome using the covariates that were included in their respective regressions.

**Quantile Regression** Because some of our outcomes, notably profit and sales, may have thick tails and/or exhibit non-uniform distributional treatment effects (see, e.g., Meager, 2022), we also included in our PAP that we may use quantile regression to examine distributional treatment effects. For these, we follow the same specification as above, except that we omit the stratum fixed effects ( $\gamma_s$ ) and the vector of covariates ( $X_i$ ) as quantile regressions are much more demanding and the covariate selection procedure in Belloni, Chernozhukov, and Hansen (2014) is designed for linear treatment effect models.

## 6 Results

In Section 6.1, we present results on access to foreign goods, as measured by our mystery shopping activity. In Section 6.2, we present results on supplier relationships. In Section 6.3, we present results on profit and sales. In almost all tables, we show the pooled

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<sup>18</sup>This means that, prior to each regression, we run lasso to predict  $y_i$  and each  $T_i$  and include the union of selected covariates. In practice, and as Cilliers, Elashmawy, and McKenzie (2024) note is generally the case, this tends to select few covariates and thus makes little difference. In Appendix C, we report versions of the main tables in which we exclude all covariates. The results are very similar.

regression in Panel A, and then disaggregate across the treated groups in Panel B, with standard errors in parentheses, Romano-Wolf multiple-testing adjusted  $p$ -values in square brackets, and the  $p$ -value of a joint test that all coefficients are equal reported at the bottom of the table.

## 6.1 Access to Foreign Goods

As described in Section 5, we designed a mystery shopping exercise to obtain a revealed preference of access to foreign goods. Table 1 reports the outcomes of this exercise.

**Horizontal** In Column 1, we report the main horizontal outcome, which is an indicator for whether the firm had a product with at least three horizontal criteria.<sup>19</sup> Pooling the treatments together, treated firms are 9.5 percentage points more likely to find a suitable good ( $p = 0.002$ ). This is a 26.6% increase from the control mean of 35.7%. In Panel B, we see that the effect is broadly similar in all four treatment groups (and we cannot reject the joint null that they are all equal), which implies that the effect is not larger for the groups with the trust treatments. The results thus show that relaxing the Search friction by connecting firms with foreign suppliers via social media led to a sizable increase in the set of varieties that they can provide to real customers.

**Vertical** The outcome in Column 2 is an indicator for whether the product is “High Quality”, defined as whether the product’s quality score is greater than its product-group median. Pooling the treatments together, treated firms are 13.1 percentage points more likely to be high quality ( $p = 0.052$ ). This is a 30.4% increase from the control mean of 43.1%. In Panel B, we see that the coefficient is positive and similar in all four treatment groups.

The outcome in Column 3 is the raw quality score out of 50. Here, the effect is both small and insignificant. In fact, the coefficients are negative in both groups with the Moral Hazard treatment. As we noted in Section 5, this outcome is vulnerable to a long left tail having an outsize influence that is not particularly meaningful. This is indeed what happens: in Appendix Figure A4 we plot the CDF of quality score by treatment status, and we see that, while the treatment CDF is to the right of the control CDF from the 30th percentile onwards, it also contains a handful of very low scoring goods.

The outcome in Column 4 is an indicator for whether the product was made in Turkey.

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<sup>19</sup>Our PAP specified this indicator variable as the main outcome. Nonetheless, as a robustness check, we use the raw number of criteria as an outcome in Appendix Table A2. The pattern of the results is the same. Our PAP also noted that we planned to disaggregate this outcome into an extensive margin effect (agreeing to sell a Turkish-made product at all) versus an intensive margin effect (how suitable was the product provided). We thus also show this in Appendix Table A2, and find that the effect comes mostly through the intensive margin.

Pooling the treatments together, treated firms are 16.2 percentage points more likely to supply a good saying “Made in Turkey”, a 34.0% increase from the control mean of 47.7% ( $p = 0.017$ ).<sup>20</sup> In Panel B, we see that the effect is positive and similar for all four treatment groups.

Finally, to account for multiple hypothesis testing across outcomes, we aggregate these three outcomes into a vertical index using the standardised inverse-variance weighting method recommended in Anderson (2008). The pooled coefficient is 0.412 standard deviations ( $p = 0.002$ ), and we cannot reject the null that the coefficient is the same for all four treatment groups. The results thus show that relaxing the Search friction by connecting firms with foreign suppliers via WhatsApp groups led to a sizeable increase in their access to higher quality goods.

**Price** In Column 6 of Table 1, the outcome is the unit price charged by the firm. The effect is positive, but small, insignificant, and precisely estimated, so we can rule out modest price increases.

**Summary** Putting together these results, we find that all of the treated groups saw large and significant increases in access to foreign goods. In particular, treated firms are able to sell a wider set of varieties and higher quality varieties. The fact that we find no large effects on the price, while there are gains from variety and quality, suggests that consumer surplus has increased.

Across all of these outcomes, the pattern is consistent: the results are driven by relaxing the Search friction. We therefore conclude that finding a supplier of Turkish-made goods is costly, and that WhatsApp can play an important role in alleviating this friction. This does not necessarily mean that the trust frictions do not exist: these are small orders, and so for many firms the risk may be sufficiently low that relaxing the trust frictions does not have a large effect. Nonetheless, we can at least conclude that trust frictions cannot be so large as to prevent firms from experimenting with new suppliers.

## 6.2 Supplier Relationships

The mystery shopping exercise shows that relaxing the search friction (through social media) improves firms’ access to foreign goods on both horizontal and vertical dimensions. However, to realise these gains in practice, firms need to overcome the trust frictions (if

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<sup>20</sup>As a robustness check, we report various alternative outcome definitions in Appendix Table A3, including specifications where we only use the information on the label and where we only use the expert tailors’ judgements. The pattern of results is the same.

any) and develop these connections into relationships. This section examines whether the treatments caused new relationships to develop, as well as what happened to previous relationships.

### 6.2.1 Survey Data

In our follow-up survey, conducted after 3 months, we asked firms how many regular suppliers they had, and where those regular suppliers were based. We defined a regular supplier as any supplier from whom the firm had made at least two orders, and intended to continue the relationship. We analyse these outcomes in Table 2. In Columns 1, the outcome is an indicator for whether the firm has a regular supplier in Turkey. In Column 2, the outcome is the number of regular suppliers in Turkey. In both cases, pooling all four treatment groups together shows that treatment caused firms to develop new relationships with suppliers in Turkey. The pooled coefficients are 3.7 percentage points (an increase of 22.2% relative to control mean of 16.7%) with  $p = 0.083$ , and 0.083 suppliers (an increase of 37.4% relative to the control mean of 0.222) with  $p = 0.024$ .

Disaggregating by treatment arm, for both outcomes, the coefficient for Search + Adverse Selection + Moral Hazard is substantially larger: this group sees a 7.5 percentage point increase in the likelihood of having a supplier in Turkey, and a 0.188 increase in the number of suppliers in Turkey, both of which are highly significant (including after adjusting for multiple hypothesis testing). The effect is also larger for the Search + Adverse Selection group. Formally, we can reject the null that all four coefficients are equal. This suggests that relaxing trust frictions, and in particular relaxing them together, increased the likelihood that these new connections developed into regular relationships.

Having seen evidence that treated firms developed new supplier relationships, it is therefore natural to ask whether these new relationships complement or substitute for existing relationships. We provide some evidence to address this in columns 4, 5, and 6. In column 4, the outcome is the total number of suppliers; in column 5, the outcome is the number of suppliers in Senegal. The general direction looks closer to a world of substitutes: the coefficients on the total number of suppliers are close to zero, and the coefficients on the number of suppliers in Senegal are of similar magnitude (but opposite sign) to the coefficients on the number of suppliers in Turkey (although they are noisy). Finally, column 6 shows an indicator for whether the firm said that they have ended a relationship with a regular supplier in the past 3 months. The coefficients are generally negative, which is also suggestive of substitutes.

### 6.2.2 Mobile Money Data

As discussed in Section 5, we use data made available for this research from the largest mobile money provider in Senegal to directly observe transactions between study firms and study suppliers. Before turning to formal regression results, we first present broad patterns in the raw data. In Figure 6, we plot cumulative order value over time from study suppliers aggregated across the four treatment groups.<sup>21</sup> The dashed line shows when we finished our mystery shopping activities. The figure shows two striking patterns. First, over the mystery shopping period, the total value ordered is very similar across the four treatment groups. This suggests that trust treatments were not necessary to make small experiments with the study suppliers where there is little demand risk (as a customer is already present). The total value ordered is much larger than the total value purchased by our mystery shoppers, so this is not simply coming from buying and re-selling to us, but—as discussed in Section 5—interpreting this is a little challenging because a secondary goal of the mystery shopping was to lower the cost for firms to experiment. Second, almost immediately after the mystery shopping ends, the Search Only line flattens, suggesting that most of these relationships were not lasting. In contrast, in all three of the trust treatments, the total value ordered continues to increase well beyond when the mystery shopping ended, suggestive of continuing relationships.

We formally test these patterns in Table 3. In this table, the omitted category is Search Only, as this outcome is not defined for pure control. Column 1 shows the effect of the trust treatments on the probability of making any order from the study suppliers. We find no significant effect, either in the pooled coefficient or the disaggregated coefficients. In Columns 2 and 3, we test the observation from Figure 6 that only the trust treatments appear to continue ordering after the mystery shopping ends. The outcome in both cases is the total value ordered after the mystery shopping ends, aggregated over weeks to the firm level. We did not pre-specify this outcome, but rather included it after observing the pattern in Figure 6. The coefficient pooling all three trust treatments together in the Poisson regression implies an  $e^{1.396} - 1 \approx 303.9\%$  increase ( $p = 0.051$ ). The pooled coefficient in the OLS regression is similar, with  $p = 0.040$ . When we disaggregate the treatments, we find that the coefficient is positive in all three trust treatment groups. While it is larger in the Moral Hazard groups and largest in the group with both trust treatments, the standard errors are too large to meaningfully distinguish between the trust groups. Finally, in columns 4 and 5, we report the effects on total value ordered over the entire course of the study. The coefficients are all positive and large, although they are not significant.

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<sup>21</sup>We omit Pure Control from the figure because they were not connected to any study suppliers. Reassuringly, we find very few orders from such firms.

The OLS coefficients in Table 3 are very similar between total value after the mystery shopping and total value, which is in line with the observation from Figure 6 that the differences only open up after the mystery shopping ends. Moreover, the fact that the trust treated groups order more after the mystery shopping ends (and thus more overall) but are not more likely to have ever made an order highlights that these are lasting relationships formed, during the study period rather, than new firms that start to order after the mystery shopping ends.

### 6.2.3 Supplier Relationships: Summary

Putting together the results from Sections 6.2.1 and 6.2.2, we find that trust treatments increased the probability that these new connections developed into relationships. From the survey data, these effects come primarily from the Search + Adverse Selection + Moral Hazard group, and, to some extent, the Search + Adverse Selection group. In the mobile money data, these effects come from all three groups.

## 6.3 Profit and Sales

### 6.3.1 Mean Results

In order to see whether improved access to foreign goods and new relationships flow through to profits, we report the results on profits and sales in Table 4. For profit, we use the summary survey question from De Mel, McKenzie, and Woodruff (2009). For sales, we use a similar summary question.

Columns 1-2 show the results on raw profit and sales. We find large and statistically significant effects. For profit, the pooled coefficient is 82.4 USD ( $p = 0.026$ ), or a 43.8% increase from the control mean. For sales, the pooled coefficient is 245.2 USD ( $p = 0.050$ ), or a 40.2% increase from the control mean. When we disaggregate across the four treated groups separately in Panel B, we see that (for both outcomes), while the effect is positive in all four groups, it is substantially larger and highly significant in the Search + Adverse Selection + Moral Hazard group (including after adjusting for multiple hypothesis testing). Formally, we can reject the null that the four coefficients are equal. The same pattern holds when we combine these two outcomes into an index.

To limit the influence of outliers, in Columns 4-5 we report the results when we winsorize the outcomes at the top 1%. The coefficients decrease in magnitude by around half on average, but the same pattern remains: the coefficient on the Search + Adverse Selection + Moral Hazard group is very large and highly significant, including after adjusting for multiple hypothesis testing.

### 6.3.2 Distributional Results

It is well-known in the literature studying small firms in lower-income countries that profit and sales tend to be thick-tailed, and that these tails can have large effects on the coefficients in OLS regressions (Meager (2022)). Thus, as discussed in Section 5.2 and specified in our PAP, we use quantile regression to examine distributional effects.<sup>22</sup>

**Quantile Treatment Effects** In Figure 7 Panel (a), we show the quantile treatment effects for profit for percentiles 5-95. Across all four groups, the coefficients are small and generally insignificant for percentiles 5-65. However, starting from around the 75th percentile, the Search + Adverse Selection + Moral Hazard group coefficient becomes large and significant. The coefficients for Search Only and Search + Adverse Selection are also large at the 95th percentile, but are very noisy. We report the same analysis for sales in Panel (b). The results are similar: there is little evidence of an effect for percentiles 5-65, but it begins to increase at around the 75th percentile for the Search + Types + Actions group, with some positive but noisy effects for Search Only at the 95th percentile.

The increasing trend in both profit and sales from percentiles 75 to 95 also suggests there may be potentially very large effects in the top 10 percentiles. We thus report the same approach for percentiles 90-99 in Appendix Figure A6. With the caveat that these are very demanding specifications, the coefficients are generally large and positive, further suggestive of treatment effects at the top of the distribution.

**Threshold Regression** An alternative way to analyse distributional treatment effects is to construct a series of indicator variables that are 1 if the outcome is greater than  $t$ , for a range of  $t$ , and run OLS regressions where these indicators are the outcome variable (using the specification in Section 5.2). This has the advantage of being unaffected by high variance in the tails: all that matters is whether the outcome is above the threshold  $t$ . The results, reported in Appendix Figure A5, are similar to the quantile regressions: large and positive treatment effects near the top of the distribution for the Search + Adverse Selection + Moral Hazard group, and some suggestive evidence of positive effects for the other groups at the very top.

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<sup>22</sup>To verify that treatment effects are not driven by measurement error in the tails, we called back all firms whose profit was more than 5 times that at baseline and exceeded a threshold at endline, and we asked them to confirm their previous response. Out of 13 such firms, 12 confirmed that their previous response was correct.

### 6.3.3 Profit and Sales: Summary

In Table 4, we saw large average effects on profit and sales. This suggests that search and trust frictions have quantitatively important effects on firm profits, and that alleviating using social media can unlock large gains. As in Section 6.2, the effects are concentrated in the groups with trust interventions.

We do not think that these effects are simply the result of a few outliers that happen to be in the treatment group, for several reasons. First, the positive distributional effects are coming from at least the top 5% of the distribution, which is considerably more than a few outliers. Second, the threshold regressions use indicators as their outcomes and thus are immune to the risk of a few observations having outsize influence. Third, the  $p$ -values in Table 4 highlight that the patterns we observe are very unlikely under the sharp null, including after adjusting for multiple hypothesis testing. Fourth, as a placebo check, we compute the same quantile figures using baseline profit as the outcome in Appendix Figure A7 (we do not have a measure of sales at baseline), and find no evidence of this pattern.

A positive effect driven by the upper tail of the profit distribution is not unusual in the literature studying firms in lower- and middle-income countries. For example, Meager (2022) aggregates the results of six RCTs on microcredit, and concludes that the evidence suggests precise zero effects on profit throughout most of the distribution, and large but uncertain effects near the top. Another example is De Mel, McKenzie, and Woodruff (2013), who study the effect of formalisation among small firms in Sri Lanka, and find profit results driven by the upper tail. Moreover, an effect concentrated among a relatively small number of firms is consistent with a small subset of firms developing meaningful relationships with the study suppliers, which is what we observe in Section 6.2.

## 6.4 e-Commerce Platforms

Our final set of outcomes relate to the Alibaba training sub-treatment. The goal of this treatment was to test whether the binding constraint explaining the very limited use of formal e-commerce platforms is that firms find these platforms too complex. Thus, we regress outcomes relating to Alibaba use against an indicator for whether the firm was in the Alibaba training group. Since we only randomised this training among firms that received none of the main treatments (i.e., the pure control group), we exclude firms that received any of the main treatments from this regression for ease of interpretation (although the results turn out to be the same if we include them).

We report the results in Table 5. The training has a first-stage: treated firms are 6.5 per-

centage points more likely to have heard of Alibaba, 11.3 percentage points more likely to have searched for goods on Alibaba, and 8.7 percentage points more likely to have compared prices on Alibaba with prices from their regular supplier. However, they are no more likely to have actually made a purchase from Alibaba. The coefficient is 1.4 percentage points and the standard errors are small enough to rule out modest to large effects. These results provide strong evidence against the hypothesis that the binding constraint is that firms struggle to understand how to use the platform. While our experiment was not designed to directly test social media against formal platforms, we can speculate that the fact that firms clearly prefer social media as their main way of doing e-commerce likely reflects something deeper about how social media—in this context—relaxes frictions in a way that formal B2B platforms do not.

## 6.5 Retailers vs Wholesalers

Our study contains both retailers and wholesalers. In particular, for 33% of the sample, some (or all) of their sales are wholesale to other firms. We henceforth refer to such firms as “wholesalers.” These types of firms are quite different, and it was *ex ante* unclear whether we should expect direct connections to foreign suppliers via social media to be more relevant for retailers or wholesalers. Retailers are generally too small to travel, and thus their baseline technology to access foreign goods (in absence of social media) is typically to buy from local wholesalers. Retailers may benefit from direct connections if these allow them to access a greater set of varieties, or to get better prices by shortening the supply chain. Wholesalers, on the other hand, are significantly more likely to travel at baseline. However, travel is expensive, and if social media allows them to partially substitute for travel, then the implications for profit are large. We thus specified that we would examine this dimension of heterogeneity in our PAP, which is what we do in this section.

We report the indexes for the main outcomes in Table 6. We use the same specification as in the rest of the paper, except that we now interact each treatment indicator with an indicator for whether the firm is a wholesaler. Throughout the table, the standard errors on the interaction coefficients are very large, so at best we can speculate on these differences. For the outcomes on access to foreign goods in Columns 1 and 2, the results appear fairly similar across retailers and wholesalers. For the index on suppliers in Turkey, the interaction coefficients are roughly equal in magnitude to the retail coefficients, perhaps suggesting a larger effect for wholesalers. Finally, there are some differences in the profit and sales index: there are essentially no detectable effects for retailers, while for wholesalers the pattern replicates that from the overall sample.

We thus conclude that the treatment does improve access to foreign goods for both retailers

and wholesalers, as evidenced by the increase in both the horizontal and vertical outcomes. The increase in foreign suppliers thus suggests that retailers substitute away from their local suppliers, but, since their scale of operations is small and this process takes time, any increases in profits in the short to medium run are sufficiently small that we are unable to detect them. Wholesalers, on the other hand, see a large increase in profit. This may be because wholesalers buy and sell in large quantities, so a new supplier can have more immediate implications for profit, or it may be because they have saved on travel costs. While we do not know for sure which one is the case, in Appendix Table A6, we show results on indicators for whether the firm has traveled internationally for business in the past 3 months (collected in the followup survey). The effects for wholesalers are negative and very large, with a few significant coefficients, but the standard errors are also large so these effects should be thought of as suggestive at best.

## 7 Model Estimation

In this section, we use the results of the experiment to estimate the model from Section 3. The goal is to estimate the parameters governing the search and trust frictions and to evaluate the gains from trade available if the frictions were to be completely or near-completely removed.

**Modifications to the Model** In order to make the model estimable, we modify the basic model in a few ways. First, we need to implement a functional form for the revenue function,  $r(q)$ , which we do in the following way. The firm faces and internalises a constant elasticity residual demand curve for goods,  $Q^{-1/\sigma} = \nu P$ , where  $\sigma > 1$  is the elasticity of demand and  $\nu$  is a demand shifter. The firm can produce the aggregate good  $Q$  by purchasing inputs from either a foreign supplier that they are matched with,  $q_f$ , or their existing supplier,  $q_e$ , which we let be perfectly substitutable.<sup>23</sup> The firm can purchase from their existing supplier in unlimited quantities at constant price  $p_e$  without frictions. Timing is such that the firm chooses  $(q_f, q_e)$ , observes whether the foreign order  $q_f$  is high quality and therefore whether  $Q = q_f + q_e$  or  $Q = q_e$ , and then sells to their downstream buyer. The firm's stage game payoff is therefore the following,

$$(1 - \mu_t(1 - \lambda))z(\psi q_{ft} + zq_{et})^{\frac{\sigma-1}{\sigma}} + \mu_t(1 - \lambda)zq_{et}^{\frac{\sigma-1}{\sigma}} - \tau_t - p_e q_{et}, \quad (2)$$

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<sup>23</sup>In principle, these could be imperfect substitutes if foreign goods are local and foreign goods represent horizontally differentiated varieties. We focus on the case of perfect substitutes as it allows for analytical solutions, which substantially speeds up the computation as this problem must be solved within an inner loop in the program to solve the dynamic model.

where  $z > 0$  is a general productivity term and  $\psi$  is a match-specific productivity term. The firm will always choose  $q_{et} > 0$ , but may choose  $q_{ft} = 0$  if either  $\mu_t$  or the transfers  $\{\tau_s\}_t^\infty$  required to incentivise the foreign supplier are sufficiently large. Firms only update their beliefs if they order positive quantity, and to prevent arbitrarily small orders to allow near-costless learning we impose a minimum order size of  $\underline{q} > 0$ , which we calibrate later and is in fact a feature of our empirical setting.

Second, we need sufficient heterogeneity to match real data. The model in Section 3 had one dimension of heterogeneity in the form of match-specific productivity  $\psi$ . We set  $\tilde{c} = 1$  and then let  $\psi$  be distributed lognormal with parameters  $(\psi_\mu, \psi_\sigma)$ , which we will estimate. The match-specific productivities are important theoretically as they define firms' beliefs about the value of searching, and empirically as they rationalise the fact that two otherwise-identical firms matched with the same suppliers may order different amounts. We also allow the productivity,  $z$ , to be heterogeneous, drawn from the empirical distribution implied by baseline profits.<sup>24</sup> This is important to rationalise the firm size distribution at baseline.

Third, we need to take a stand on what firms' outside options are and how to account for the fact that some firms already have foreign suppliers at baseline. The sequential search process described in Section 3 implies that there exists a cutoff value  $\bar{z}$  such that, at baseline, all firms with  $z > \bar{z}$  will search until they eventually find a foreign supplier with sufficiently high match-specific productivity. This cutoff is a function of parameters to be estimated. We interpret the baseline equilibrium as the very long-run of the model, meaning that all firms with  $z > \bar{z}$  have found a good-type foreign supplier and fully learned their type. For such firms,  $p_e$  represents the price offered by their foreign supplier, which is a draw of  $\psi$  above the cutoff at which a firm would search again. For all other firms,  $p_e$  represents the cost of buying from a local supplier in Dakar, which we will calibrate below. A firm's outside option,  $\bar{U}$ , is then defined by the value of only purchasing from their existing supplier forever,

$$\bar{U} = \left( \max_{q_e} q_e^{\frac{\sigma-1}{\sigma}} - p_e q_e \right) / (1 - \delta).$$

**Parameters** Our goal is to estimate the parameters governing the three frictions: the search cost,  $s$ , the share of bad types,  $\mu_0$ , and the moral hazard multiplier,  $\xi$ . We also need to estimate the parameters governing the lognormal distribution of match-specific

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<sup>24</sup>Specifically, assuming that the firm faces one input price, the model implies that profits for firm  $i$  are given by  $\pi_i = z^\sigma \frac{(1-\sigma)^{1-\sigma}}{\sigma^\sigma} p_i^{1-\sigma}$ , where  $p_i$  is the input price. Re-arranging allows us to write  $z_i$  as a function of  $\pi_i, p_i, \sigma$ . We use monthly profit from the baseline survey for  $\pi_i$ , average price of the most common input from the baseline survey for  $p_i$ , and we calibrate  $\sigma$  as described below. This gives an empirical distribution of  $z_i$ .

productivity,  $(\psi_\mu, \psi_\sigma)$ , which may also be thought of as characterising the search friction. We thus estimate these five parameters, and calibrate all others. We describe the values and origins of the calibrated parameters,  $(\sigma, \delta, p_l, q, \lambda)$ , in Appendix Table A7.

**Moments** Since the ultimate goal is to use the model to extrapolate from the treatments to consider counterfactuals where we vary the extent of the frictions, we estimate the model by simulating the impact of treatment in the model and then matching the reduced form treatment effects. We select the four treatment effects on winsorized profit, the four treatment effects on likelihood of having a supplier in Turkey, and the three treatment effects on mobile money order value post mystery shopping.

We simulate the three treatments in the model as follows. For search, we implement this as the firm being matched (at zero cost) to a foreign supplier with match-specific productivity equal to the maximum of three draws from the distribution (to capture the idea that the firm is matched to three foreign suppliers in the experiment). For adverse selection, we implement this as the firm receiving one high quality signal realisation without having to purchase anything, meaning that they update as a Bayesian and thus begin the relationship with  $\mu_1 = \mu_0\lambda/(1 - \mu_0(1 - \lambda)) < \mu_0$ . This intends to capture the recommendation call from the adverse selection treatment that explicitly told the firm about one positive order experience. For moral hazard, we implement this as treated firms playing a joint punishment strategy. Specifically, they face a modified DICC of the form  $(1 - \lambda)\delta V_{t+1} \geq \xi c_{qt} - (1 - \lambda)\delta N\tilde{V}_{t+1}$ , where  $\tilde{V}_{t+1}$  is the average relationship value among other firms and  $N = 5$  is the ratio of the number of firms that made orders from study suppliers to the number of study suppliers.

**Solving and Estimating the Model** In order to compute the simulated treatment effects, we need to solve the model. The model is dynamic with a non-stationary optimal contract, meaning that solving it numerically is non-trivial. We use the method of Marcer and Marinon (2019), which involves rewriting the original Lagrangean recursively and then defining a Saddle Point Functional Equation (SPFE), which is analogous to a standard Bellman Equation for saddle point problems. We can then use standard dynamic programming techniques, and in particular we iterate on the value function implied by the SPFE. For estimation, we use Simulated Method of Moments (SMM) with a weighting matrix equal to the inverse of the variance-covariance matrix of the empirical moments, and with 100 randomized initial points for the numerical minimization algorithm. We provide a more detailed description of the procedure to solve and estimate the model in Appendix D.

**Results** We obtain  $[\mu_0, \xi, s, \psi_\mu, \psi_\sigma] = [0.68, 0.59, \$13.1, -3.05, 0.40]$ . The value of  $\mu_0 = 0.68$  suggests a fairly severe adverse selection problem: around two-thirds of suppliers are bad types. The value of  $\xi = 0.59$  is in between a standard model of moral hazard and a model with perfect enforcement (at least among good types), although the overall extent of moral hazard depends on whether and how severely the DICCs bind. The search cost of  $s = 13.1$  USD may not, on the surface, seem especially large. However, the estimated parameters governing the distribution of match-specific productivity imply that the probability of a random foreign supplier being a good enough match to transact with is in the order of 10%. An average firm would thus need to pay the search cost many times before finding a suitable match.

**Counterfactuals** We now use the estimated model to evaluate two sets of counterfactuals. In the first type, we consider a counterfactual where no firms can import directly. This is equivalent to setting  $s = \infty$  and setting  $p_e$  to be the price of the local supplier in Dakar for all firms. With the baseline estimated parameters, the discounted value of profits,  $U_0$ , is 7,953 USD. With the calibrated monthly discount factor of 0.96, this implies a long-run average monthly profit of around 300 USD, which is broadly in line with our baseline profit statistics. When we implement this counterfactual, average lifetime discounted profits decrease to 6,005 USD, a decrease of around one quarter. The ability to trade with foreign suppliers thus quantitatively matters.

For the second type of counterfactual, we set various combinations of the parameters governing the frictions to zero or near zero. We show the results in Table 7. We first cut the search cost in half—reducing it to zero is not well-defined because as  $s \rightarrow 0$  firms will search indefinitely to obtain arbitrarily good matches. This increases lifetime discounted profits by relatively little, only raising it to 8,014 USD. The effect is small for two reasons. First, the likelihood of finding a match that will be worth developing a relationship with, governed by the distribution of match-specific productivity and the trust frictions, is relatively low. Second, the very high value of  $\mu_0$  means that even if a firm finds a good productivity match, there is still a substantial probability that the supplier is a bad type. Together, these two forces mean that the return to searching is low. Setting  $\xi$  to zero (without changing  $\mu_0$ ) has almost no effect. This is because moral hazard only has bite by constraining  $q_t$ . In the early phases of the relationship, the fact that  $\mu_0$  is large means that the firm wants to choose low values of  $q_t$  anyway, so reducing  $\xi$  has little effect. In the latter phases of the relationship, the estimated distribution of productivity implies that there is generally sufficient surplus available to incentivise the supplier, meaning that the DICC does not bind in the long-run.

Setting  $\mu_0$  to 0 has a larger effect, raising lifetime discounted profits to 9,661 USD, or around a 20% increase. This magnitude is not unreasonable when considering the profit effects, and it is worth noting that the theory predicts that the firm realises its profit gains entirely in the early phases of the relationship. Importantly, the effect is large not just because lowering  $\mu_0$  causes new relationships to form, but also because it improves the return to searching. Finally, note also that the impact of setting  $\xi = 0$  increases substantially from almost nothing to around 150 USD when  $\mu_0 = 0$ . This happens because the treatments are complements. As noted before, moral hazard imposes little bite when  $\mu_0$  is large, as the firm already wants to order low quantities. When  $\mu_0 = 0$ , this is no longer the case. Overall, the largest gains from eliminating the frictions come from reducing adverse selection.

## 8 Conclusion

We study the extent to which search and trust frictions limit the ability of small firms to buy inputs from foreign markets, and whether social media can be used to alleviate them. Finding and developing relationships with suppliers is a first order issue, but one potentially rife with these information frictions. Thus, providing rigorous evidence on these issues is important to understand the nature and magnitude of these frictions, to rationalize the descriptive fact that firms use social media ubiquitously to interact with and learn about suppliers, and to suggest policy remedies.

Our results suggest that these frictions social media can meaningfully reduce these frictions. First, we find that connecting firms with foreign suppliers improves their access to foreign inputs, as shown by the results of our mystery shopping exercise. This is not *ex ante* obvious: instead of using social media to interact directly with suppliers abroad, firms could buy from a local wholesaler or travel abroad themselves. Yet, the fact that we find relatively large effects suggests that substantial search frictions are still present, and that direct connections via social media can meaningfully reduce them.

Our mystery shopping results are largely driven by relaxing the search friction. Ultimately, though, whether firms are able to realise this better access to foreign input markets depends on whether these connections develop into lasting relationships. Our findings, using both survey data and mobile money data, suggest that this is more likely to happen among firms in the trust treatments. This suggests that trust frictions are barriers to access, and that using social media to share information and coordinate action can reduce them. Ultimately, we find that our treatments cause some firms to develop meaningful new relationships and to increase profit.

The widespread use of social media suggests that there is substantial demand for e-commerce,

so it may therefore seem puzzling why firms don't use formal platforms that are explicitly designed to alleviate search and trust frictions. The fact that our Alibaba training has no effect on their propensity to buy inputs from it suggests that the reason is more fundamental than simply lack of knowledge. The exact reasons for limited use seem like fruitful topics for future research, but, the fact that firms clearly prefer social media as their main way of doing e-commerce likely reflects something deeper about how social media—in this context—relaxes frictions in a way that formal platforms do not.

Taken together, our results show that both search and trust frictions meaningfully limit the ability of small firms to buy inputs from foreign markets, and that social media can be used to meaningfully reduce them. Moreover, these technologies are changing the structure of supply chains for some firms: retailers can import directly, and wholesalers can better identify target suppliers before traveling or save on travel costs altogether. The rapid growth in access to smartphones and mobile connectivity, as well as the efforts of the companies themselves to introduce e-commerce features, is likely to further these changes. Our results thus suggest that the rapidly developing digital landscape in lower- and middle-income countries is likely to meaningfully benefit small firms and require researchers, policymakers, and organizations to update how they think about how firms find, learn about, and develop relationships with suppliers.

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## Tables

Table 1: Access to Foreign Goods

	Horizontal	Vertical			Price	
	Has Product ≥ 3 Criteria	High Qual Dummy	Qual Score (/50)	Made in Turkey	Index	
		(1)	(2)	(3)	(4)	(5)
<i>Panel A: Pooled</i>						
Treatment	0.095*** (0.030)	0.131* (0.067)	-0.391 (0.562)	0.162** (0.067)	0.412*** (0.135)	0.710 (0.699)
<i>Panel B: Individual Treatments</i>						
Search Only	0.140*** (0.038) [0.001]	0.175** (0.083) [0.110]	0.259 (0.675) [0.899]	0.138* (0.082) [0.164]	0.426** (0.166) [0.037]	1.360 (0.896) [0.364]
Search + Adverse Selection	0.053 (0.037) [0.159]	0.159* (0.087) [0.159]	-0.027 (0.858) [0.974]	0.130 (0.086) [0.164]	0.398** (0.169) [0.037]	0.445 (0.833) [0.807]
Search + Moral Hazard	0.105*** (0.038) [0.016]	0.099 (0.082) [0.388]	-1.179 (0.789) [0.387]	0.183** (0.081) [0.067]	0.399** (0.162) [0.037]	1.064 (0.866) [0.464]
Search + AS + MH	0.072* (0.038) [0.100]	0.083 (0.084) [0.388]	-0.717 (0.814) [0.718]	0.211*** (0.081) [0.035]	0.424** (0.171) [0.037]	-0.025 (0.818) [0.975]
Control Mean	0.357	0.431	43.064	0.477	0.000	19.990
% Increase (Pooled)	26.6%	30.4%	-0.9%	34.0%	N/A	3.6%
All Coefs Equal <i>p</i> -val	0.120	0.615	0.320	0.668	0.996	0.294
Adjusted <i>R</i> <sup>2</sup>	0.09	0.04	0.34	0.09	0.02	0.40
<i>N</i>	1579	359	359	361	359	642

Note: *p*-values are computed using randomisation inference. Specifically, we compute the randomised randomisation-*t* *p*-value from Young (2019) using 5000 reps. \* *p* < 0.1 \*\* *p* < 0.05 \*\*\* *p* < 0.01. We report conventional robust standard errors in parentheses, although we do not use these directly for inference. We also report Romano-Wolf multiple-testing adjusted *p*-values in square brackets (Romano and Wolf, 2005). To test the null that the trust treatments have no effect, we report at the bottom of the table the *p*-value for an *F*-test that all coefficients are equal, computed by permuting the *F*-statistic. Panel A shows the coefficient from a regression on an indicator that pools all treated groups. Panel B shows the coefficients corresponding to treatment indicators for each of the four treatment arms. All regressions include covariates selected by Double Lasso (Belloni, Chernozhukov, and Hansen, 2014), as well as stratum fixed effects and the outcome measured at baseline (if available).

Column 1 is an indicator that is one if the firm has a good that matches at least 3 horizontal criteria, and is missing if the firm never replied to the mystery shopper or was otherwise unreachable. Column 2 is an indicator for whether the good's quality score is above the median product-group quality score. Column 3 is the raw quality score. Column 4 is an indicator for whether the good is made in Turkey, primarily inferred based on the label. See the text for full details of how this outcome is constructed. Column 5 is the Anderson (2008) index combining the vertical outcomes. Column 6 is the price in USD, which is only measured conditional on the firm finding a good matching at least three horizontal criteria.

Table 2: Supplier Relationships (Followup Survey)

	Regular Suppliers in Turkey			Previous Suppliers		
	Any Sup in Turkey (1)	Num Sup in Turkey (2)	Index (3)	Num Sup Total (4)	Num Sup in Senegal (5)	Ended with Sup (6)
<i>Panel A: Pooled</i>						
Treatment	0.037* (0.021)	0.083** (0.035)	0.121* (0.063)	-0.102 (0.164)	-0.130 (0.170)	0.062*** (0.021)
<i>Panel B: Individual Treatments</i>						
Search Only	0.024 (0.027) [0.582]	0.081 (0.051) [0.252]	0.085 (0.081) [0.470]	-0.101 (0.216) [0.891]	-0.111 (0.222) [0.923]	0.050* (0.028) [0.115]
Search + Adverse Selection	0.049* (0.029) [0.203]	0.063 (0.047) [0.306]	0.147* (0.086) [0.209]	-0.171 (0.218) [0.844]	-0.245 (0.226) [0.645]	0.066** (0.028) [0.052]
Search + Moral Hazard	0.003 (0.026) [0.908]	0.003 (0.041) [0.935]	0.009 (0.079) [0.907]	-0.118 (0.201) [0.891]	-0.064 (0.212) [0.923]	0.076** (0.029) [0.035]
Search + AS + MH	0.075*** (0.028) [0.034]	0.188*** (0.057) [0.003]	0.250*** (0.087) [0.018]	-0.014 (0.222) [0.950]	-0.096 (0.225) [0.923]	0.053* (0.028) [0.115]
Control Mean	0.167	0.222	0.000	3.700	3.213	0.135
% Increase (Pooled)	22.2%	37.4%	N/A	-2.8%	-4.0%	45.9%
All Coefs Equal <i>p</i> -val	0.075	0.009	0.048	0.927	0.866	0.835
Adjusted <i>R</i> <sup>2</sup>	0.14	0.12	-0.02	0.32	0.26	0.05
<i>N</i>	1680	1680	1680	1681	1681	1671

Note: *p*-values are computed using randomisation inference. Specifically, we compute the randomised randomisation-*t* *p*-value from Young (2019) using 5000 reps. \* *p* < 0.1 \*\* *p* < 0.05 \*\*\* *p* < 0.01. We report conventional robust standard errors in parentheses, although we do not use these directly for inference. We also report Romano-Wolf multiple-testing adjusted *p*-values in square brackets (Romano and Wolf, 2005). Panel A shows the coefficient from a regression on an indicator that pools all treated groups. Panel B shows the coefficients corresponding to treatment indicators for each of the four treatment arms. All regressions include covariates selected by Double Lasso (Belloni, Chernozhukov, and Hansen, 2014), as well as stratum fixed effects and the outcome measured at baseline (if available).

Column 1 is 1 if the merchant says that they have a regular supplier in Turkey. Column 2 is the number of regular suppliers in Turkey. Column 3 is the Anderson (2008). Column 4 is the total number of regular suppliers. Column 5 is the number of regular suppliers in Senegal. Column 6 is 1 if the merchant has ended a relationship with a regular supplier in the past 3 months. A regular supplier is defined as a supplier from whom the merchant has made two or more orders with an intention of continuing the relationship.

Table 3: Order Value (Mobile Money Data)

	Any Order (1)	Value Post Mystery Shopping		Total Value	
		Order Value (OLS) (2)	Order Value (Poisson) (3)	Order Value (OLS) (4)	Order Value (Poisson) (5)
<i>Panel A: Pooled</i>					
Trust Treatment	0.001 (0.020)	4.892** (2.546)	1.396* (0.538)	4.349 (3.291)	0.418 (0.310)
<i>Panel B: Individual Treatments</i>					
Search + Adverse Selection	0.004 (0.024) [0.887]	2.151 (1.674) [0.308]	0.804 (0.564) [0.336]	2.215 (2.889) [0.728]	0.209 (0.317) [0.658]
Search + Moral Hazard	0.012 (0.025) [0.877]	6.029*** (2.856) [0.023]	1.573** (0.533) [0.066]	4.840 (3.819) [0.449]	0.453 (0.341) [0.574]
Search + AS + MH	-0.014 (0.024) [0.877]	6.524 (6.506) [0.387]	1.629 (0.889) [0.336]	6.027 (7.419) [0.728]	0.565 (0.556) [0.658]
Control Mean	0.131	1.571	1.571	8.023	8.023
% Increase (Pooled)	0.8%	311.4%	303.9%	54.2%	51.9%
Adjusted $R^2$	0.04	0.00	0.04	0.01	0.02
N	1500	1500	1500	1500	1500

Note:  $p$ -values are computed using randomisation inference. Specifically, we compute the randomised randomisation- $t$   $p$ -value from Young (2019) using 5000 reps. \*  $p < 0.1$  \*\*  $p < 0.05$  \*\*\*  $p < 0.01$ . We report conventional robust standard errors in parentheses, although we do not use these directly for inference. We also report Romano-Wolf multiple-testing adjusted  $p$ -values in square brackets (Romano and Wolf, 2005). To test the null that the trust treatments have no effect, we report at the bottom of the table the  $p$ -value for an  $F$ -test that all coefficients are equal, computed by permuting the  $F$ -statistic. Panel A shows the coefficient from a regression on an indicator that pools all trust treated groups, where Search Only is the omitted category. Panel B shows the coefficients corresponding to treatment indicators for each of the three treatment groups with trust treatments. All regressions include covariates selected by Double Lasso (Belloni, Chernozhukov, and Hansen, 2014), as well as stratum fixed effects and the outcome measured at baseline (if available).

Column 1 is an indicator for whether the firm ever ordered from a study supplier. Column 2 is the total value of orders. Column 3 is the total value of orders, analysed with Poisson regression. Column 4 is the total value of orders. Column 5 is the total value of orders, analysed with Poisson regression. Mystery shopping took place during the first 13 weeks of the study. All values are in USD.

Table 4: Profit and Sales

	Raw			Winsorized (1%)		
	Profit (USD) (1)	Sales (USD) (2)	Index (3)	Profit (USD) (4)	Sales (USD) (5)	Index (6)
<i>Panel A: Pooled</i>						
Treatment	82.4** (31.6)	245.2* (108.4)	0.314** (0.107)	45.5** (21.2)	121.5 (78.6)	0.174** (0.076)
<i>Panel B: Individual Treatments</i>						
Search Only	31.7 (28.8) [0.564]	311* (177) [0.181]	.197* (.114) [0.188]	22.1 (25.4) [0.632]	178 (112) [0.259]	.21** (.108) [0.123]
Search + Adverse Selection	43.5 (37.7) [0.564]	80.8 (114) [0.726]	.193 (.129) [0.242]	30.8 (30.7) [0.632]	73.1 (100) [0.692]	.124 (.102) [0.382]
Search + Moral Hazard	15.8 (25.1) [0.568]	-24 (106) [0.832]	.036 (.085) [0.712]	8.21 (23.2) [0.731]	-25.3 (87) [0.769]	6.2e-03 (.078) [0.939]
Search + AS + MH	254*** (89.3) [0.004]	636** (267) [0.031]	.88*** (.315) [0.005]	128*** (39.2) [0.003]	269** (124) [0.090]	.371*** (.137) [0.019]
Control Mean	188.3	609.5	0.000	188.3	609.5	0.000
% Increase (Pooled)	43.8%	40.2%	N/A	24.2%	19.9%	N/A
All Coefs Equal <i>p</i> -val	0.044	0.046	0.031	0.021	0.053	0.023
Adjusted <i>R</i> <sup>2</sup>	0.13	0.12	-0.01	0.25	0.29	-0.02
<i>N</i>	1351	1378	1298	1351	1378	1298

*Note:* *p*-values are computed using randomisation inference. Specifically, we compute the randomised randomisation-*t* *p*-value from Young (2019) using 5000 reps. \* *p* < 0.1 \*\* *p* < 0.05 \*\*\* *p* < 0.01. We report conventional robust standard errors in parentheses, although we do not use these directly for inference. We also report Romano-Wolf multiple-testing adjusted *p*-values in square brackets (Romano and Wolf, 2005). Panel A shows the coefficient from a regression on an indicator that pools all treated groups. Panel B shows the coefficients corresponding to treatment indicators for each of the four treatment arms. All regressions include covariates selected by Double Lasso (Belloni, Chernozhukov, and Hansen, 2014), as well as stratum fixed effects and the outcome measured at baseline (if available).

Column 1 is total profit from the past 30 days in USD. Column 2 is total sales from the past 30 days in USD. Column 3 is the Anderson (2008) index combining the previous two columns. Column 4 is total profit from the past 30 days in USD, winsorizing the top 1%. Column 5 is total sales from the past 30 days in USD, winsorizing the top 1%. Column 6 is the Anderson (2008) index combining the previous two columns. Profit is measured using the survey question from De Mel, McKenzie, and Woodruff (2009). Sales is measured using a similar survey question.

Table 5: Effect of Alibaba Training

	Heard of Alibaba (1)	Searched on Alibaba (2)	Compared Prices with Supplier (3)	Bought on Alibaba (4)
e-Commerce Treatment	0.065*** (0.025)	0.113** (0.052)	0.087* (0.049)	0.014 (0.035)
Control Mean	0.908	0.423	0.319	0.135
% Increase	7.2%	26.8%	27.4%	10.7%
Adjusted $R^2$	0.07	0.09	0.14	0.22
N	340	340	340	340

Note:  $p$ -values are computed using randomisation inference. Specifically, we compute the randomised randomisation- $t$   $p$ -value from Young (2019) using 5000 reps. \*  $p < 0.1$  \*\*  $p < 0.05$  \*\*\*  $p < 0.01$ . Conventional robust standard errors are reported in parentheses. This table shows the effect of the Alibaba training treatment on Alibaba usage.

Table 6: Heterogeneity by Retailer vs Wholesaler

	Horiz Dummy (1)	Vert Index (2)	Sup Turk Index (3)	Prof Sales Index (1%) (4)
Search Only	0.194*** (0.047)	0.281 (0.190)	0.055 (0.091)	0.004 (0.104)
Search + AS	0.049 (0.047)	0.351* (0.198)	0.144 (0.096)	0.013 (0.121)
Search + MH	0.131*** (0.047)	0.234 (0.199)	-0.017 (0.087)	0.004 (0.091)
Search + AS + MH	0.036 (0.049)	0.291 (0.225)	0.156 (0.101)	0.024 (0.103)
S Only * Wholesaler	-0.158* (0.083)	0.126 (0.408)	0.027 (0.186)	0.469 (0.324)
S + AS * Wholesaler	0.048 (0.083)	-0.079 (0.417)	-0.006 (0.182)	0.255 (0.323)
S + MH * Wholesaler	-0.024 (0.084)	0.044 (0.405)	0.051 (0.177)	-0.144 (0.231)
S + AS + MH * Wholesaler	0.114 (0.083)	0.087 (0.414)	0.105 (0.186)	1.209*** (0.437)
Adjusted $R^2$	0.01	-0.01	0.01	0.05
N	1579	359	1680	1298

Note: This table shows the main results with treatment interacted with an indicator for whether the firm is a wholesaler, defined as having some positive share of sales that are wholesale to other firms.  $p$ -values are computed using randomisation inference. Specifically, we compute the randomised randomisation- $t$   $p$ -value from Young (2019) using 5000 reps. \*  $p < 0.1$  \*\*  $p < 0.05$  \*\*\*  $p < 0.01$ . We report conventional robust standard errors in parentheses, although we do not use these directly for inference. The outcome in Column 1 is the same as Column 1 of Table 1. The outcome in Column 2 is the same as Column 2 of Table 1. The outcome in Column 3 is the same as Column 3 of Table 2. The outcome in Column 4 is the same as Column 6 of Table 4.

Table 7: Counterfactuals

Scenario	Discounted Profits (USD)
Baseline	7,952
$s = 6.5$	8,014
$s = 6.5, \mu_0 = 0$	9,661
$s = 6.5, \xi = 0$	8,016
$s = 6.5, \mu_0 = \xi = 0$	9,813

## Figures

Figure 1: Design Tree

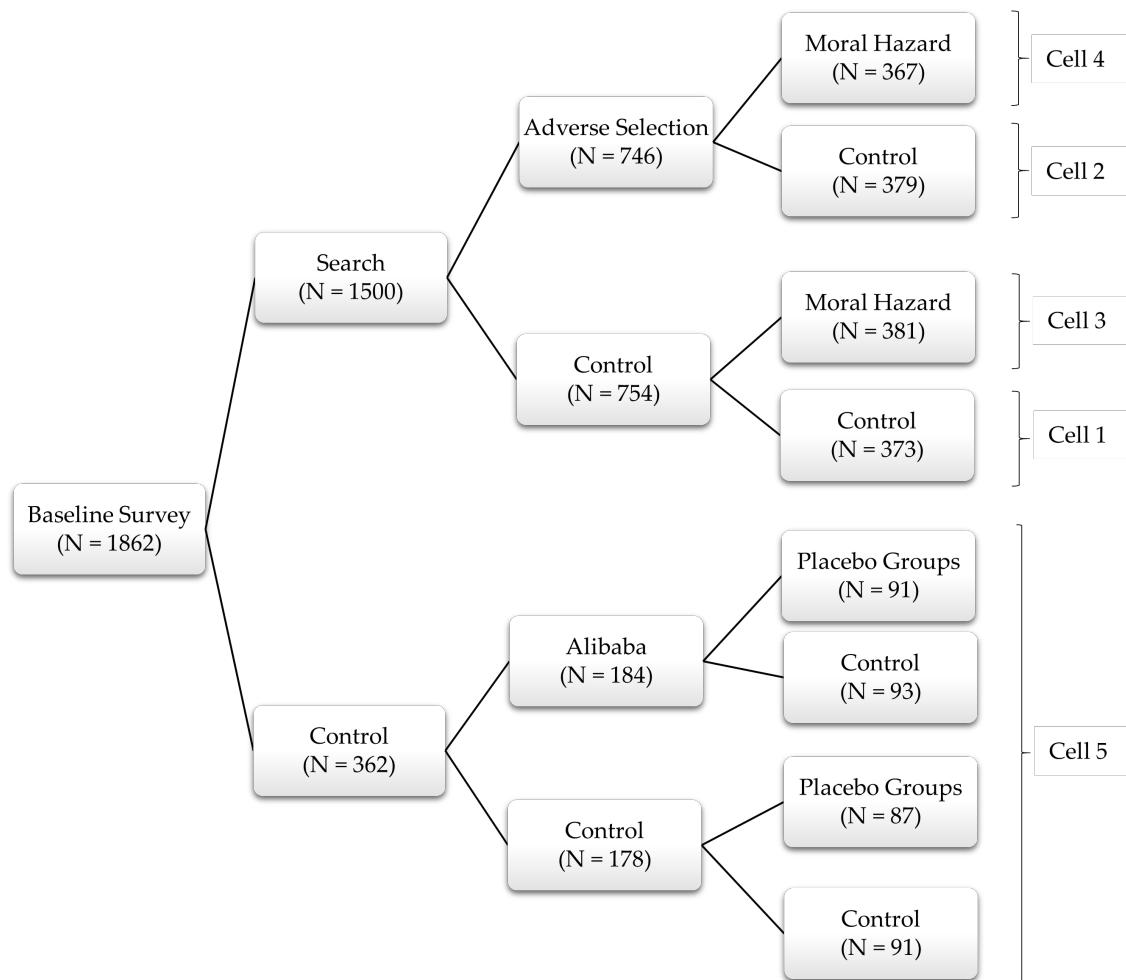


Figure 2: Supplier WhatsApp Groups

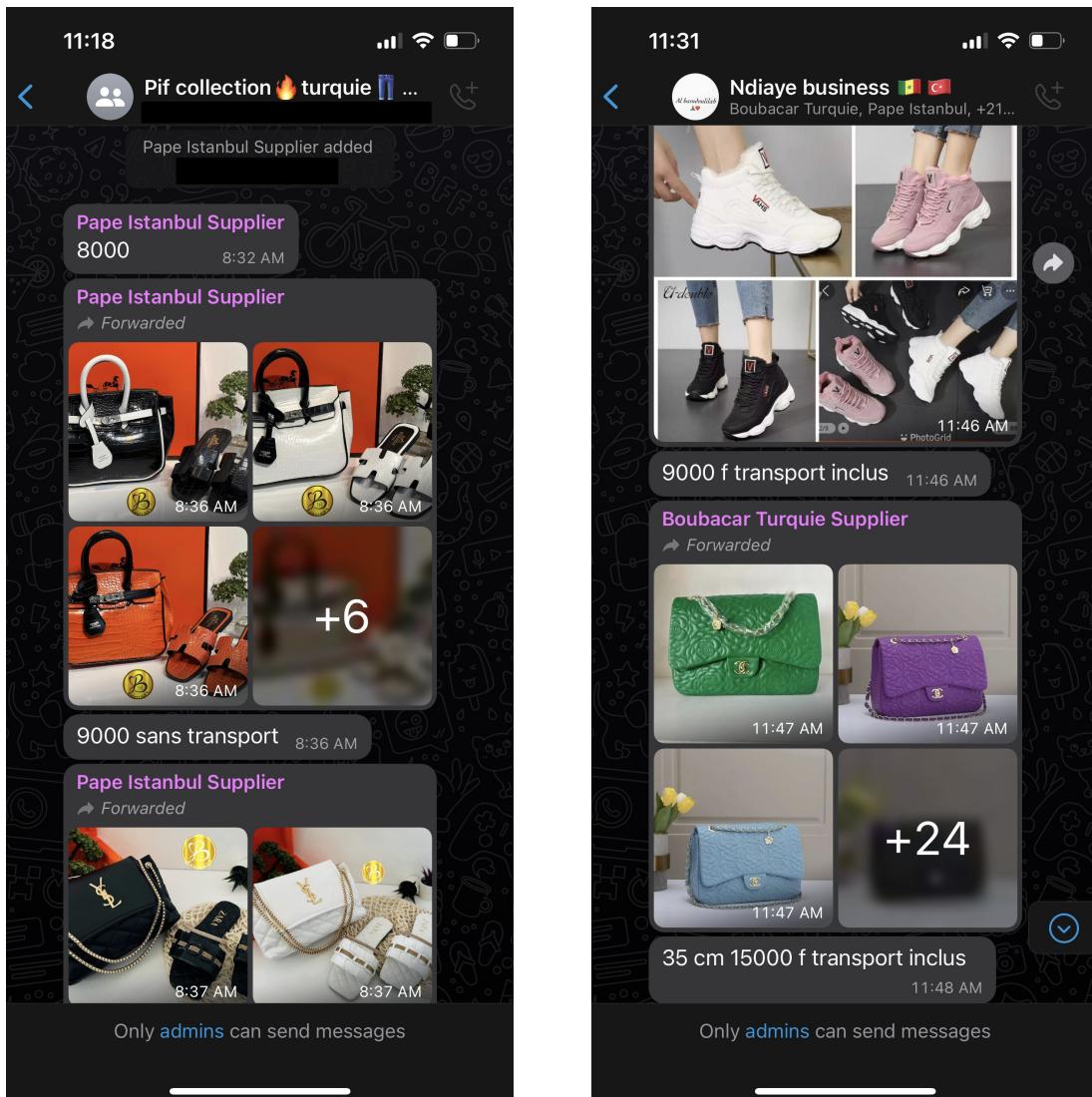
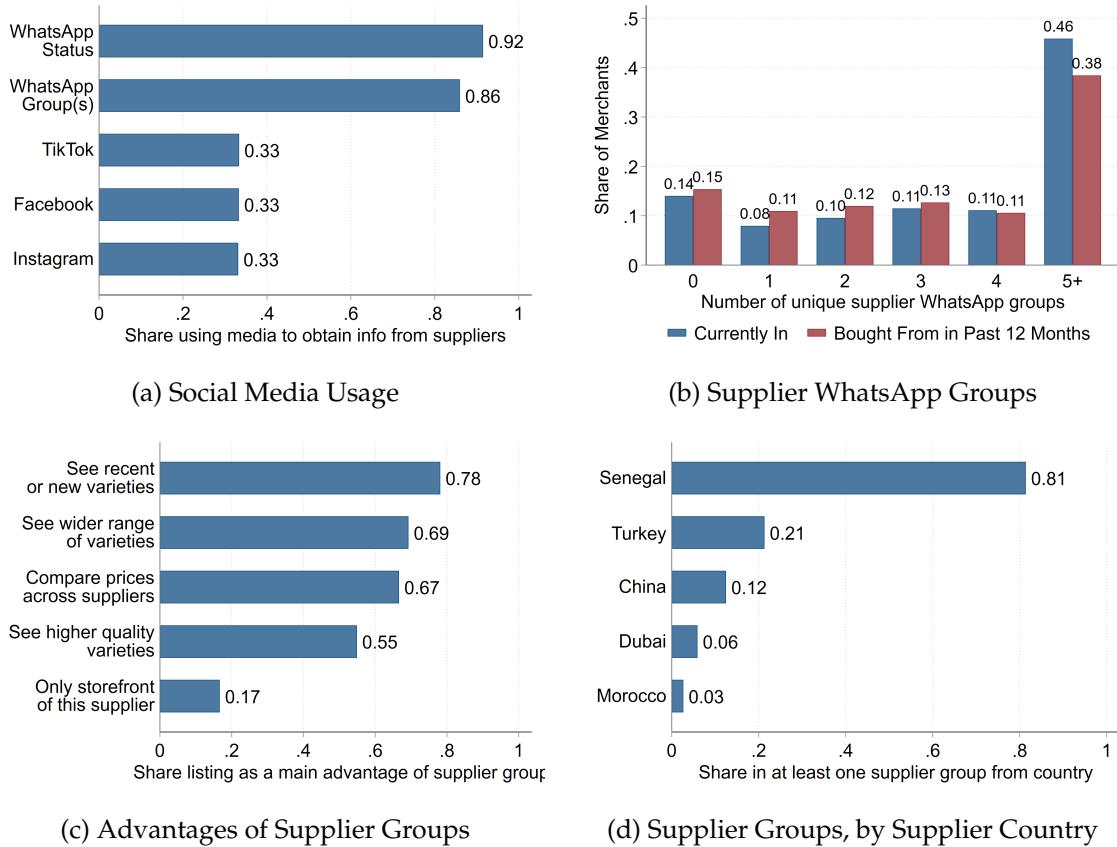
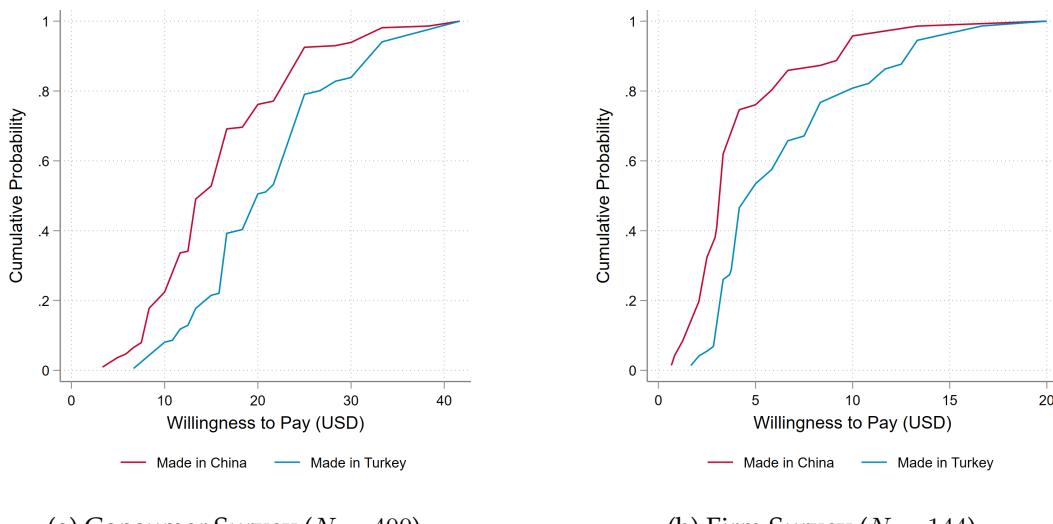


Figure 3: Firm Social Media Usage to obtain Information about Suppliers



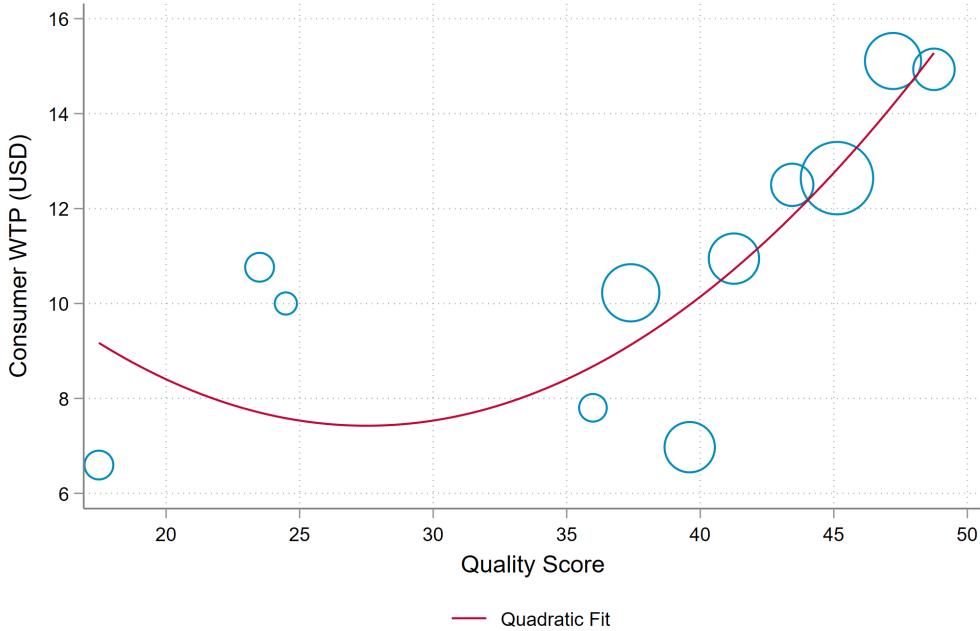
*Note:* This figure shows a number of statistics about how firms in our sample use social media to obtain information about suppliers. All data is from our baseline survey with 1,862 firms. Panel (a) shows the results of a question asking firms to select all social media that they use to obtain information about suppliers for their business. Panel (b) shows the distribution of the number of supplier WhatsApp groups a firm is in at the time of the baseline survey, as well as the distribution of the number of such groups that the firm has directly made at least one purchase from in the past 12 months. Supplier WhatsApp groups are defined as WhatsApp groups in which the primary purpose is for suppliers to advertise their wares to downstream clients. Panel (c) shows the results of a question asking firms that use supplier WhatsApp groups to select all reasons why they find these groups useful. Panel (d) shows the share of firms who are in at least one supplier WhatsApp group where the supplier located and based in the country listed.

Figure 4: Willingness to Pay by Product Origin

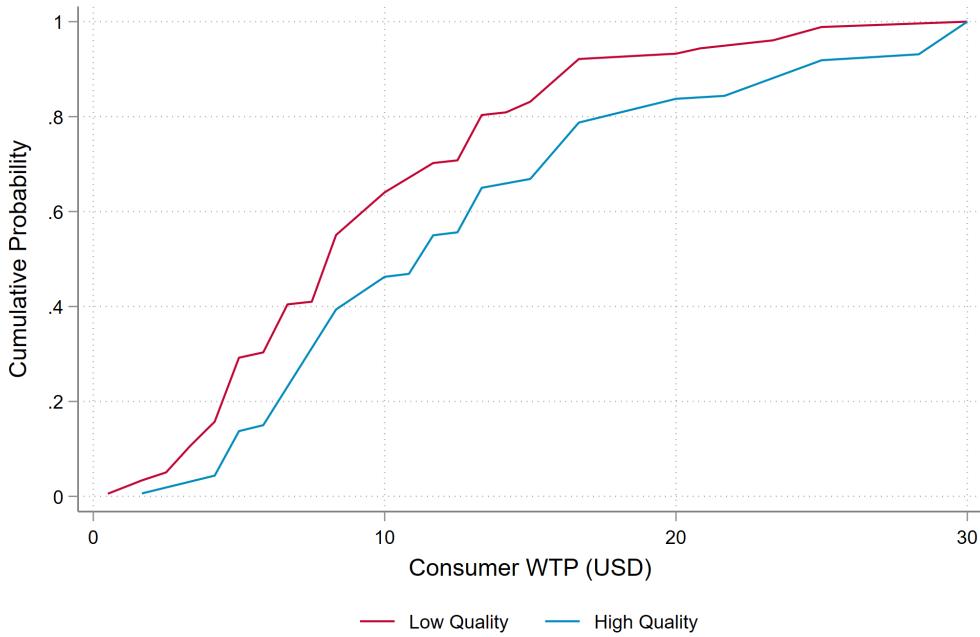


*Note:* This figure shows the results of two exercises in which we showed various images of garments to survey respondents, randomised whether we said the good was made in Turkey or made in China, and elicited willingness to pay for the garments. Panel (a) shows the CDF of WTP in the consumer survey, separately by whether we said the good was made in Turkey or China. The distribution is truncated at 40 USD for ease of readability. Panel (b) shows the CDF of WTP for a small, separate survey of firms (for a different set of goods), with distribution truncated at 20 USD for ease of readability.

Figure 5: Consumer Willingness to Pay for Quality



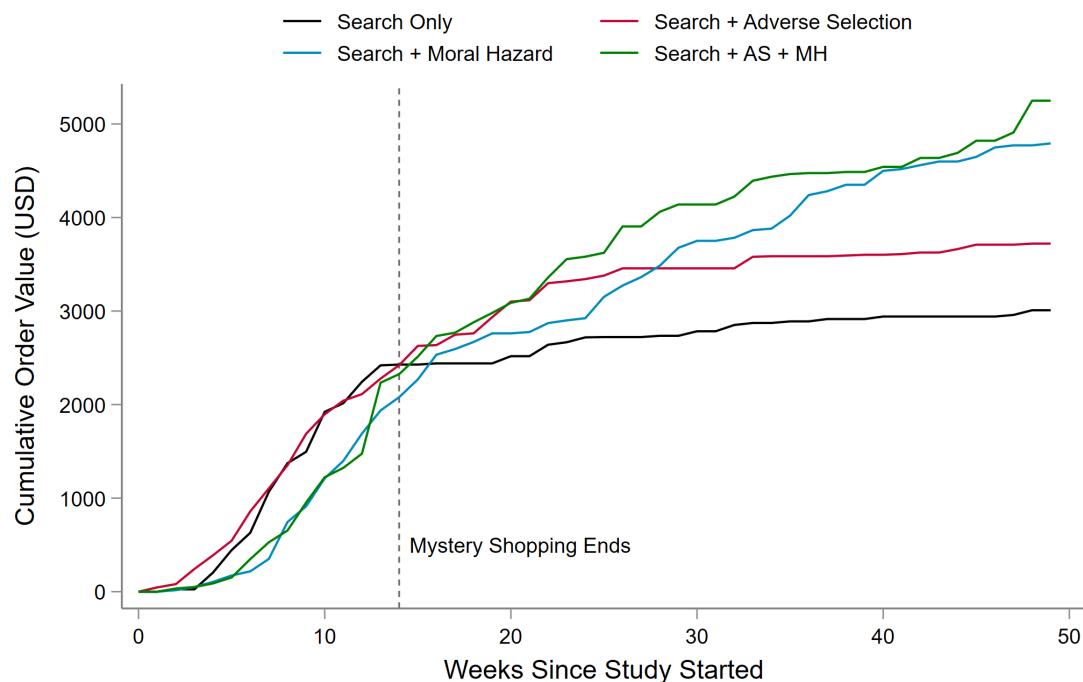
(a) Raw Score



(b) CDF of WTP by High or Low Quality

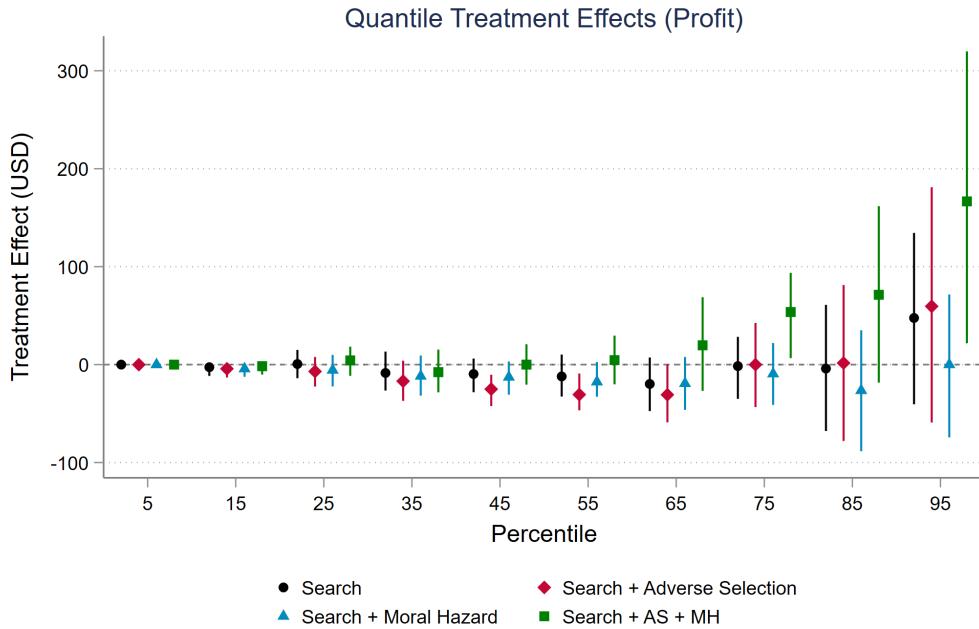
*Note:* Panel (a) shows a binscatter of consumer willingness to pay for garments (as measured by the consumer survey) against the quality score of the garments. The size of each bubble is proportional to the number of observations. Panel (b) shows the CDF of consumer willingness to pay separately based on whether the garment met our definition of high quality. We truncate willingness to pay at 30 USD to avoid unnecessarily stretching the x-axis. See the main text for full details on the consumer survey and variable construction.

Figure 6: Cumulative Order Value (Mobile Money Data)

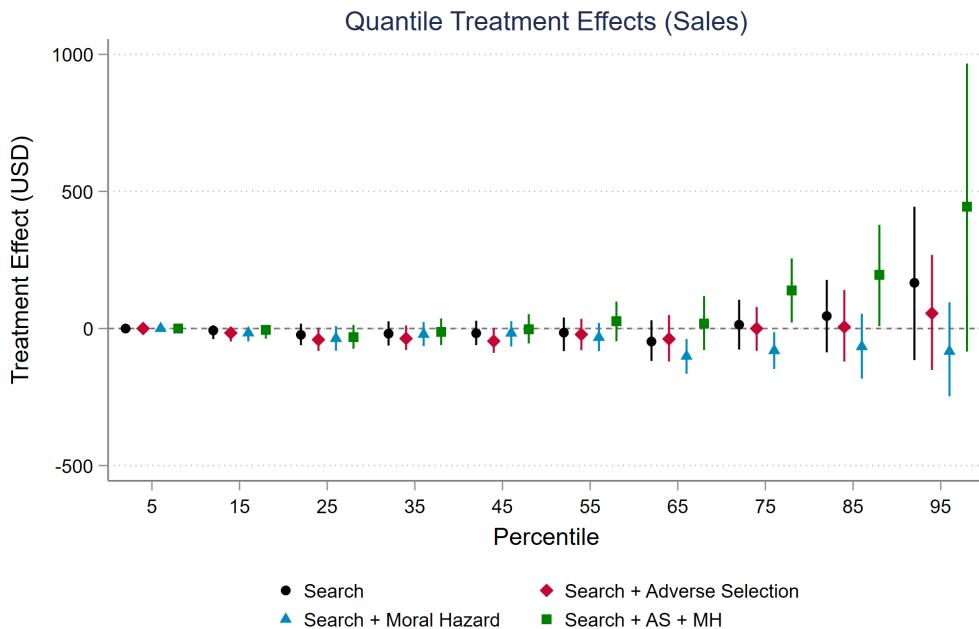


*Note:* This figure shows the total order value from study suppliers, according to the mobile money data, in each treatment group as a function of number of weeks since the study begun (16 November 2023). Pure control is omitted as they were not connected to any study suppliers.

Figure 7: Quantile Treatment Effects



(a) Profit



(b) Sales

*Note:* This figure shows the coefficients from quantile regressions of profit and sales on the four treatment groups. All quantile regressions include the outcome measured at baseline (if available), but otherwise do not include any covariates. We plot 95% confidence intervals constructed using randomisation inference, defined as the set of sharp nulls that do not reject at the 5% level.

## Appendix A – Additional Tables and Figures

### Tables

Table A1: Balance Table

Variable	Control (1)	Search (2)	Search AS (3)	Search MH (4)	Search AS MH (5)	Joint <i>p</i> -value (6)
Female	0.49	0.51	0.49	0.53	0.49	0.74
Online Only	0.67	0.66	0.68	0.7	0.65	0.56
Business Age	4.85	4.74	5.2	4.72	5.21	0.57
Share Cust Turkey	0.43	0.45	0.45	0.44	0.45	0.96
Any Reg Supp Turkey	0.22	0.21	0.24	0.21	0.19	0.67
Travelled Business (5y)	0.09	0.08	0.11	0.09	0.11	0.43
Profit USD (30 Days)	221.35	221.08	262.49	195.36	235.63	0.24
Bought Alibaba Ever	0.16	0.13	0.15	0.18	0.13	0.23
<i>N</i>	362	373	379	381	367	

*Note:* The table shows the mean for each variable in each of the five treatment cells. The final column shows the *p*-value from regressing the variable on indicators for each treatment (where the control group is omitted) and conducting a test that all coefficients are zero. Finally, we run a multinomial logit of treatment group against all of the variables in the table, for which a joint test that all coefficients are zero has a *p*-value of 0.724.

Table A2: Horizontal Outcomes (Detailed)

	Extensive vs Intensive Margin		Number of Criteria	
	Agree Search (1)	Find Product Conditional (2)	Num Criteria Unconditional (3)	Num Criteria Conditional (4)
<i>Panel A: Pooled</i>				
Treatment	0.028 (0.025)	0.101*** (0.035)	0.361*** (0.124)	0.357** (0.140)
<i>Panel B: Individual Treatments</i>				
Search Only	0.029 (0.031) [0.666]	0.144*** (0.043) [0.003]	0.664*** (0.161) [0.000]	0.705*** (0.178) [0.001]
Search + Adverse Selection	0.014 (0.031) [0.666]	0.067 (0.044) [0.113]	0.196 (0.160) [0.315]	0.182 (0.182) [0.316]
Search + Moral Hazard	0.045 (0.030) [0.352]	0.099** (0.043) [0.060]	0.379** (0.158) [0.039]	0.308* (0.175) [0.182]
Search + AS + MH	0.025 (0.032) [0.666]	0.092** (0.044) [0.060]	0.204 (0.158) [0.315]	0.232 (0.179) [0.316]
Control Mean	0.781	0.457	1.650	2.111
% Increase (Pooled)	3.6%	22.1%	21.9%	16.9%
All Coefs Equal <i>p</i> -val	0.750	0.344	0.016	0.016
Adjusted <i>R</i> <sup>2</sup>	0.10	0.04	0.05	0.03
<i>N</i>	1579	1269	1579	1269

Note: *p*-values are computed using randomisation inference. Specifically, we compute the randomised randomisation-*t* *p*-value from Young (2019) using 5000 reps. \* *p* < 0.1 \*\* *p* < 0.05 \*\*\* *p* < 0.01. We report conventional robust standard errors in parentheses, although we do not use these directly for inference. We also report Romano-Wolf multiple-testing adjusted *p*-values in square brackets (Romano and Wolf, 2005). Panel A shows the coefficient from a regression on an indicator that pools all treated groups. Panel B shows the coefficients corresponding to treatment indicators for each of the four treatment arms. All regressions include covariates selected by Double Lasso (Belloni, Chernozhukov, and Hansen, 2014), as well as stratum fixed effects and the outcome measured at baseline (if available).

Column 1 is 1 if the merchant agrees to sell or search for the product, and is missing if the merchant never replied to the mystery shopper or was otherwise unreachable. Column 2 is 1 if the merchant found a suitable product, conditional on agreeing to sell or search for the product. Column 3 is the number of horizontal criteria of the product, and is 0 if the merchant either did not agree to sell or search for a product or agreed but never sent any product. Column 4 is the number of horizontal criteria of the product, conditional on agreeing to sell or search for a product, and is 0 if the merchant agreed but never sent any product.

Table A3: Vertical Outcomes (Detailed)

	From Turkey		
	Made in Turkey (Label) (1)	Made in Turkey (Tailor Judgement) (2)	Made in Turkey (Label + Tailors) (3)
<i>Panel A: Pooled</i>			
Treatment	0.212** (0.082)	0.096 (0.082)	0.158** (0.072)
<i>Panel B: Individual Treatments</i>			
Search Only	0.169* (0.096) [0.138]	0.134 (0.095) [0.425]	0.138 (0.086) [0.187]
Search + Adverse Selection	0.172* (0.100) [0.138]	0.010 (0.102) [0.920]	0.107 (0.089) [0.232]
Search + Moral Hazard	0.241** (0.094) [0.029]	0.132 (0.095) [0.425]	0.192** (0.084) [0.072]
Search + AS + MH	0.271*** (0.097) [0.020]	0.096 (0.097) [0.509]	0.189** (0.084) [0.072]
Control Mean	0.489	0.581	0.544
% Increase (Pooled)	43.4%	16.5%	29.0%
All Coefs Equal <i>p</i> -val	0.523	0.491	0.627
Adjusted <i>R</i> <sup>2</sup>	0.10	-0.00	0.07
<i>N</i>	255	287	330

Note: *p*-values are computed using randomisation inference. Specifically, we compute the randomised randomisation-*t* *p*-value from Young (2019) using 5000 reps. \* *p* < 0.1 \*\* *p* < 0.05 \*\*\* *p* < 0.01. We report conventional robust standard errors in parentheses, although we do not use these directly for inference. We also report Romano-Wolf multiple-testing adjusted *p*-values in square brackets (Romano and Wolf, 2005). Panel A shows the coefficient from a regression on an indicator that pools all treated groups. Panel B shows the coefficients corresponding to treatment indicators for each of the four treatment arms. All regressions include covariates selected by Double Lasso (Belloni, Chernozhukov, and Hansen, 2014), as well as stratum fixed effects and the outcome measured at baseline (if available).

Column 1 is 1 if the label says , 0 if the label says for X other than Turkey, and missing otherwise. Column 2 is 1 if both tailors independently determined that the product was made in Turkey, and is 0 if both tailors independently determined that the product was not made in Turkey. It is missing if the tailors disagreed. For shoes, as there was only one expert shoemaker, we take their opinion directly. Column 3 is an indicator that combines the label and tailor measures of whether the good was made in Turkey. It is equal to the label measure where available, and the tailor measure otherwise.

Table A4: Supplier Relationships (Further Results on Substitution)

	Reg Supp in China		Media for Suppliers			Forward Media	
	Any Supp in China	Num Supp in China	Uses Facebook	Uses TikTok	Uses Instagram	Fwd Photo for Search	Fwd Photo for Price
	(1)	(2)	(3)	(4)	(5)	(6)	(7)
<i>Panel A: Pooled</i>							
Treatment	-0.025 (0.016)	-0.014 (0.035)	-0.084*** (0.027)	-0.026 (0.025)	-0.026 (0.026)	-0.052* (0.027)	-0.095*** (0.030)
<i>Panel B: Individual Treatments</i>							
Search Only	-0.030 (0.019) [0.317]	-0.023 (0.040) [0.923]	-0.090*** (0.033) [0.018]	-0.047 (0.032) [0.371]	-0.037 (0.032) [0.587]	-0.054 (0.036) [0.226]	-0.103*** (0.039) [0.022]
Search + Adverse Selection	-0.029 (0.019) [0.317]	-0.028 (0.045) [0.923]	-0.063* (0.034) [0.070]	-0.006 (0.032) [0.929]	-0.038 (0.033) [0.587]	-0.033 (0.035) [0.334]	-0.088** (0.038) [0.040]
Search + Moral Hazard	-0.018 (0.020) [0.441]	0.009 (0.044) [0.946]	-0.080** (0.033) [0.032]	-0.039 (0.032) [0.458]	-0.016 (0.033) [0.845]	-0.063* (0.035) [0.213]	-0.081** (0.038) [0.041]
Search + AS + MH	-0.023 (0.020) [0.441]	-0.013 (0.045) [0.946]	-0.103*** (0.033) [0.006]	-0.011 (0.033) [0.929]	-0.012 (0.033) [0.845]	-0.058* (0.035) [0.226]	-0.109*** (0.039) [0.022]
Control Mean	0.099	0.167	0.328	0.290	0.279	0.716	0.659
% Increase (Pooled)	-25.3%	-8.4%	-25.6%	-9.0%	-9.3%	-7.3%	-14.4%
All Coefs Equal <i>p</i> -val	0.895	0.831	0.620	0.498	0.779	0.841	0.891
Adjusted <i>R</i> <sup>2</sup>	0.12	0.36	0.09	0.12	0.12	0.04	0.03
<i>N</i>	1680	1680	1671	1671	1671	1671	1565

Note: *p*-values are computed using randomisation inference. Specifically, we compute the randomised randomisation-*t* *p*-value from Young (2019) using 5000 reps. \* *p* < 0.1 \*\* *p* < 0.05 \*\*\* *p* < 0.01. We report conventional robust standard errors in parentheses, although we do not use these directly for inference. We also report Romano-Wolf multiple-testing adjusted *p*-values in square brackets (Romano and Wolf, 2005). Panel A shows the coefficient from a regression on an indicator that pools all treated groups. Panel B shows the coefficients corresponding to treatment indicators for each of the four treatment arms. All regressions include covariates selected by Double Lasso (Belloni, Chernozhukov, and Hansen, 2014), as well as stratum fixed effects and the outcome measured at baseline (if available).

Column 1 is 1 if the merchant says that they have a regular supplier in China. Column 2 is 1 if the merchant says that they have a regular supplier in China. Column 3 is 1 if the merchant says that they use Facebook to learn about suppliers. Column 4 is 1 if the merchant says that they use TikTok to learn about suppliers. Column 5 is 1 if the merchant says that they use Instagram to learn about suppliers. Column 6 is 1 if the merchant says that they have forwarded a photo or video from a supplier group to a regular supplier to try to obtain a similar product in the past 3 months. Column 7 is 1 if the merchant says that they have forwarded a photo or video from a supplier group to a regular supplier to try to obtain a better price in the past 3 months. A regular supplier is defined as a supplier from whom the merchant has made two or more orders with an intention of continuing the relationship.

Table A5: Profit and Sales (Poisson Regression)

	Profit		Sales	
	Profit 30 Days (USD) (1)	Profit 30 Days Winsorized 1% (USD) (2)	Sales 30 Days (USD) (3)	Sales 30 Days Winsorized 1% (USD) (4)
<i>Panel A: Pooled</i>				
Treatment	0.307* (0.132)	0.188 (0.112)	0.184 (0.138)	0.085 (0.123)
<i>Panel B: Individual Treatments</i>				
Search Only	0.088 (0.147) [0.916]	0.074 (0.133) [0.840]	0.315 (0.210) [0.427]	0.198 (0.158) [0.517]
Search + Adverse Selection	0.097 (0.189) [0.916]	0.119 (0.162) [0.840]	-0.041 (0.167) [0.840]	-0.026 (0.159) [0.881]
Search + Moral Hazard	-0.047 (0.138) [0.916]	-0.009 (0.143) [0.951]	-0.151 (0.174) [0.672]	-0.147 (0.143) [0.547]
Search + AS + MH	0.865*** (0.214) [0.002]	0.480*** (0.142) [0.010]	0.475* (0.212) [0.164]	0.260 (0.162) [0.387]
Control Mean	188.3	188.3	609.5	609.5
% Increase (Pooled)	35.9%	20.7%	20.2%	8.9%
All Coefs Equal <i>p</i> -val	0.007	0.020	0.067	0.058
Adjusted <i>R</i> <sup>2</sup>	0.34	0.38	0.46	0.48
<i>N</i>	1351	1351	1378	1378

Note: *p*-values are computed using randomisation inference. Specifically, we compute the randomised randomisation-*t* *p*-value from Young (2019) using 5000 reps. \* *p* < 0.1 \*\* *p* < 0.05 \*\*\* *p* < 0.01. We report conventional robust standard errors in parentheses, although we do not use these directly for inference. We also report Romano-Wolf multiple-testing adjusted *p*-values in square brackets (Romano and Wolf, 2005). Panel A shows the coefficient from a regression on an indicator that pools all treated groups. Panel B shows the coefficients corresponding to treatment indicators for each of the four treatment arms. All regressions include covariates selected by Double Lasso (Belloni, Chernozhukov, and Hansen, 2014), as well as stratum fixed effects and the outcome measured at baseline (if available).

Column 1 is total profit from the past 30 days in USD. Column 2 is total profit from the past 30 days in USD, winsorizing the top 1%. Column 3 is total sales from the past 30 days in USD. Column 4 is total sales from the past 30 days in USD, winsorizing the top 1%. Profit is measured using the survey question from De Mel, McKenzie, and Woodruff (2009). Sales is measured using a similar survey question.

Table A6: Travel

	Wholesalers			Retailers		
	Any Travel (1)	Travel China (2)	Travel Turkey (3)	Any Travel (4)	Travel China (5)	Travel Turkey (6)
<i>Panel A: Pooled</i>						
Treatment	-0.025 (0.035)	-0.029 (0.030)	-0.015 (0.019)	-0.004 (0.014)	-0.021* (0.011)	0.003 (0.005)
<i>Panel B: Individual Treatments</i>						
Search Only	-0.001 (0.049) [0.983]	-0.009 (0.039) [0.933]	-0.020 (0.026) [0.864]	-0.005 (0.017) [0.982]	-0.022* (0.013) [0.247]	-0.004 (0.004) [0.720]
Search + Adverse Selection	0.007 (0.041) [0.983]	-0.011 (0.034) [0.933]	-0.012 (0.022) [0.864]	-0.001 (0.018) [0.996]	-0.025** (0.012) [0.119]	0.004 (0.008) [0.840]
Search + Moral Hazard	-0.051 (0.038) [0.420]	-0.050* (0.030) [0.265]	-0.014 (0.020) [0.864]	-0.008 (0.017) [0.976]	-0.015 (0.014) [0.286]	0.010 (0.008) [0.594]
Search + AS + MH	-0.052 (0.037) [0.420]	-0.043 (0.032) [0.394]	-0.016 (0.020) [0.864]	-0.001 (0.020) [0.996]	-0.022 (0.014) [0.247]	0.001 (0.006) [0.887]
Control Mean	0.130	0.090	0.040	0.041	0.033	0.004
% Increase (Pooled)	-19.2%	-32.2%	-37.5%	-9.8%	-63.6%	75.0%
All Coefs Equal <i>p</i> -val	0.173	0.264	0.985	0.981	0.864	0.173
Adjusted <i>R</i> <sup>2</sup>	0.18	0.28	0.09	0.04	0.07	0.01
<i>N</i>	546	546	546	1125	1125	1125

Note: *p*-values are computed using randomisation inference. Specifically, we compute the randomised randomisation-*t* *p*-value from Young (2019) using 5000 reps. \* *p* < 0.1 \*\* *p* < 0.05 \*\*\* *p* < 0.01. We report conventional robust standard errors in parentheses, although we do not use these directly for inference. We also report Romano-Wolf multiple-testing adjusted *p*-values in square brackets (Romano and Wolf, 2005). Panel A shows the coefficient from a regression on an indicator that pools all treated groups. Panel B shows the coefficients corresponding to treatment indicators for each of the four treatment arms. All regressions include covariates selected by Double Lasso (Belloni, Chernozhukov, and Hansen, 2014), as well as stratum fixed effects and the outcome measured at baseline (if available).

Column 1 is 1 if the firm travelled for business in the past 3 months. Column 2 is 1 if the firm travelled for business to China in the past 3 months. Column 3 is 1 if the firm travelled for business to Turkey in the past 3 months. Column 4 is 1 if the firm travelled for business in the past 3 months. Column 5 is 1 if the firm travelled for business to China in the past 3 months. Column 6 is 1 if the firm travelled for business to Turkey in the past 3 months. Travel is 1 if either the firm owner or someone closely involved with the firm travelled internationally for firm-specific business in the past 3 months.

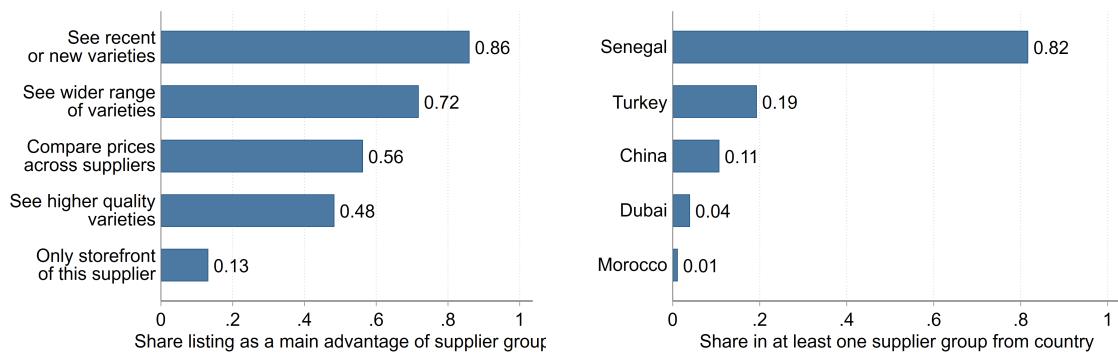
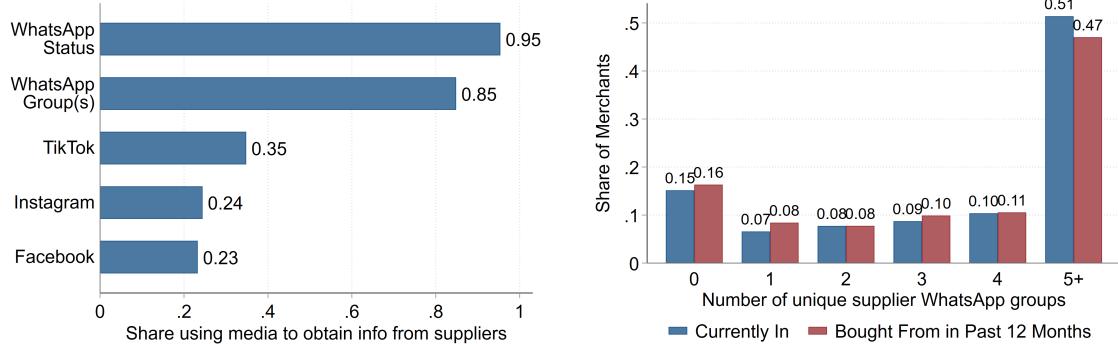
Table A7: Calibrated Parameters

Parameter	Value	Origin
$\sigma$	3.02	Average markup from baseline survey.
$\delta$	0.96	Empirical survival rate from baseline to endline survey.
$p_l$	12.29	Average input price from baseline survey among firms without a foreign supplier.
$\underline{q}$	20	Common price of pack of six articles.
$\lambda$	Varies	See table notes.

*Note:* This table presents the values of the parameters that are calibrated in the model. The value of  $\lambda$  is calculated by inverting the mean response to a baseline survey question asking “Imagine if you were connected to 10 new foreign suppliers and you made test orders from all of them. In how many of these 10 cases would you expect that the goods would arrive with the quality expected?”. The mean is 0.54, so we calculate  $\lambda$  from  $(1 - 0.54) = \mu_0(1 - \lambda)$ , which is the model-implied expected share of low-quality orders from a series of random draws from the population of suppliers. For  $\underline{q}$ , a large share of goods in this market can only be ordered in packs of six, and 20 USD is among the lowest price for a pack of six that suppliers commonly sell.

## Figures

Figure A1: Social Media Usage (Physical Store Only)

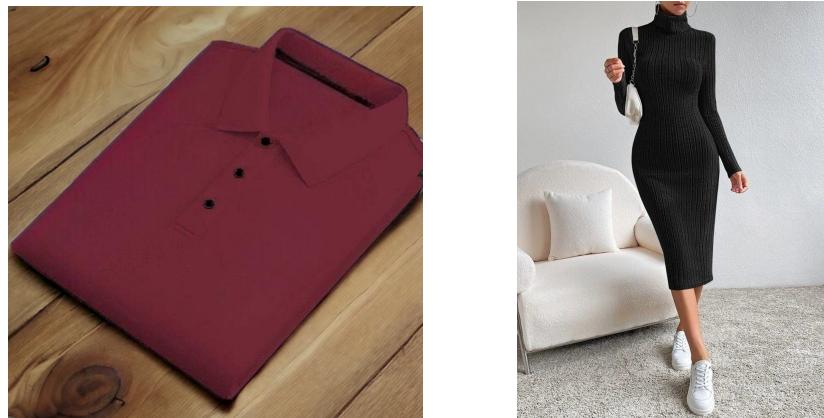


*Note:* This figure shows a number of statistics about how firms in our sample use social media to obtain information about suppliers. It is the same as Figure 3, but instead calculates statistics only for the 607 firms that have physical stores. Panel (a) shows the results of a question asking firms to select all social media that they use to obtain information about suppliers for their business. Panel (b) shows the distribution of the number of supplier WhatsApp groups a firm is in at the time of the baseline survey, as well as the distribution of the number of such groups that the firm has directly made at least one purchase from in the past 12 months. Supplier WhatsApp groups are defined as WhatsApp groups in which the primary purpose is for suppliers to advertise their wares to downstream clients. Panel (c) shows the results of a question asking firms that use supplier WhatsApp groups to select all reasons why they find these groups useful. Panel (d) shows the share of firms who are in at least one supplier WhatsApp group where the supplier located and based in the country listed.

Figure A2: Business Cards for Moral Hazard Treatment



Figure A3: Mystery Shopping Goods (Examples)



1. Red

2. One colour

3. No large logo/print

4. 3 black buttons

5. Size L

1. Black

2. Long sleeves

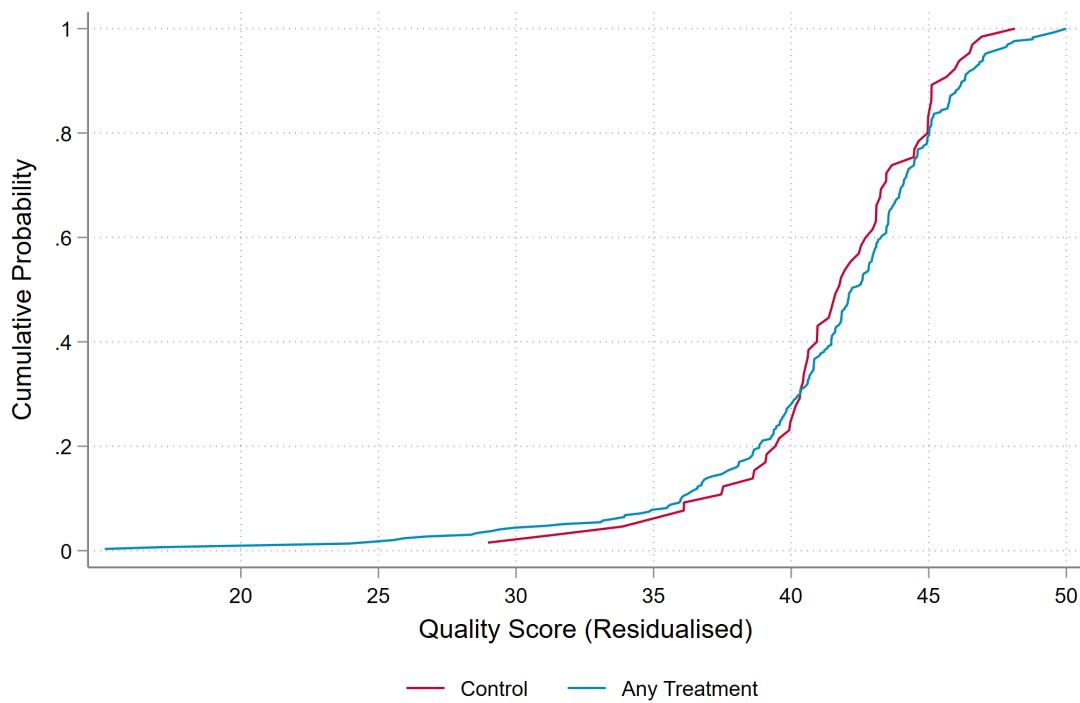
3. Knee length

4. Turtleneck

5. Ribbed Cotton

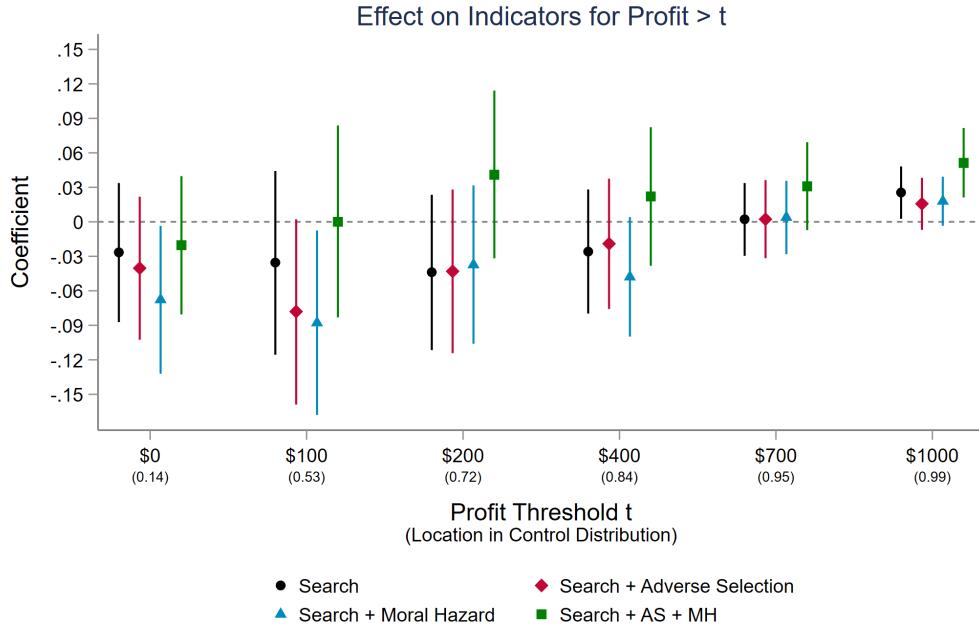
Examples of goods requested in the mystery shopping exercise. In total, there were 28 different goods.

Figure A4: Quality Score Distribution



*Note:* This figure shows CDF of the quality score separately by treatment status, with all four treatment groups (Search Only, Search + Types, Search + Actions, Search + Types + Actions) pooled for ease of readability. To be consistent with the regression in the table, we first residualise quality using stratum fixed effects and the covariates selected in the regression.

Figure A5: Threshold Regressions for Profit and Sales



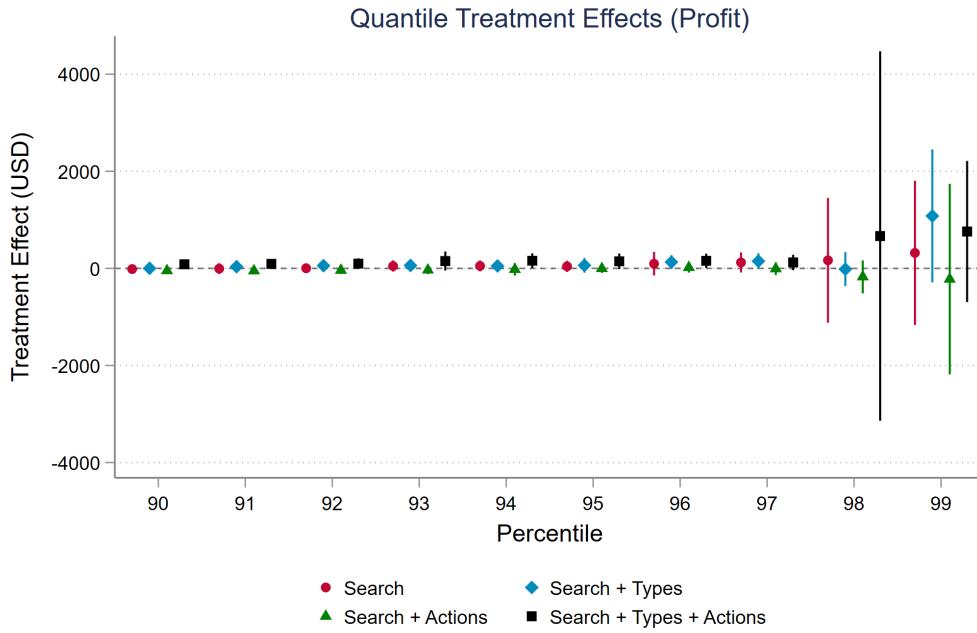
(a) Profit



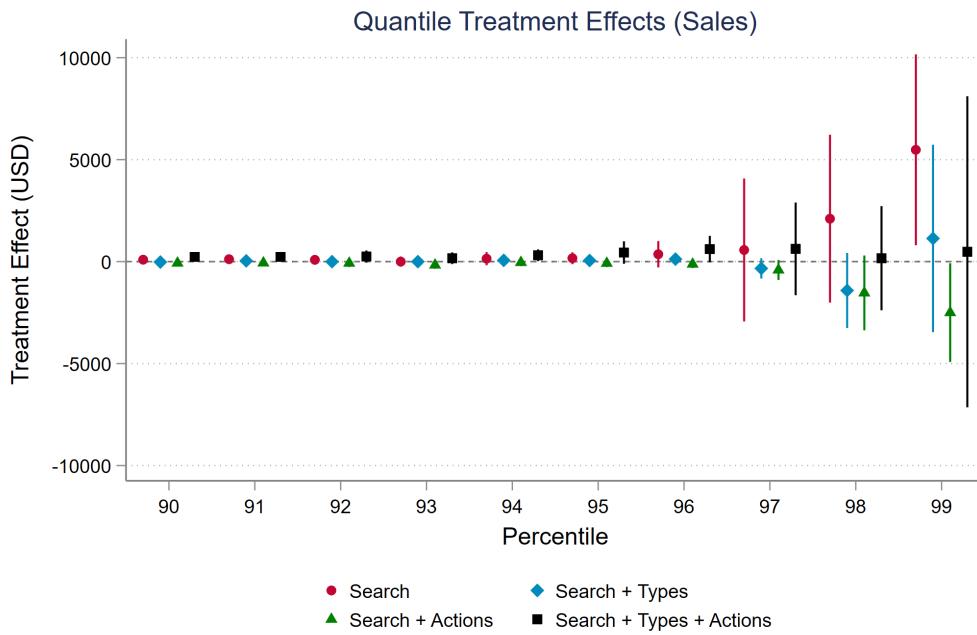
(b) Sales

*Note:* This figure shows the coefficients from regressions of indicators for whether profit and sales are above some threshold  $t$ , for a range of  $t$ . All regressions include covariates selected by Double Lasso (Belloni, Chernožukov, and Hansen, 2014), as well as the outcome measured at baseline (if available). The numbers in parentheses show the percentiles at which  $t$  is located in the distribution of the pure control group. We plot 95% confidence intervals constructed using randomisation inference, defined as the set of sharp nulls that do not reject at the 5% level, using the procedure in Young (2024).

Figure A6: Quantile Treatment Effects (90-99)



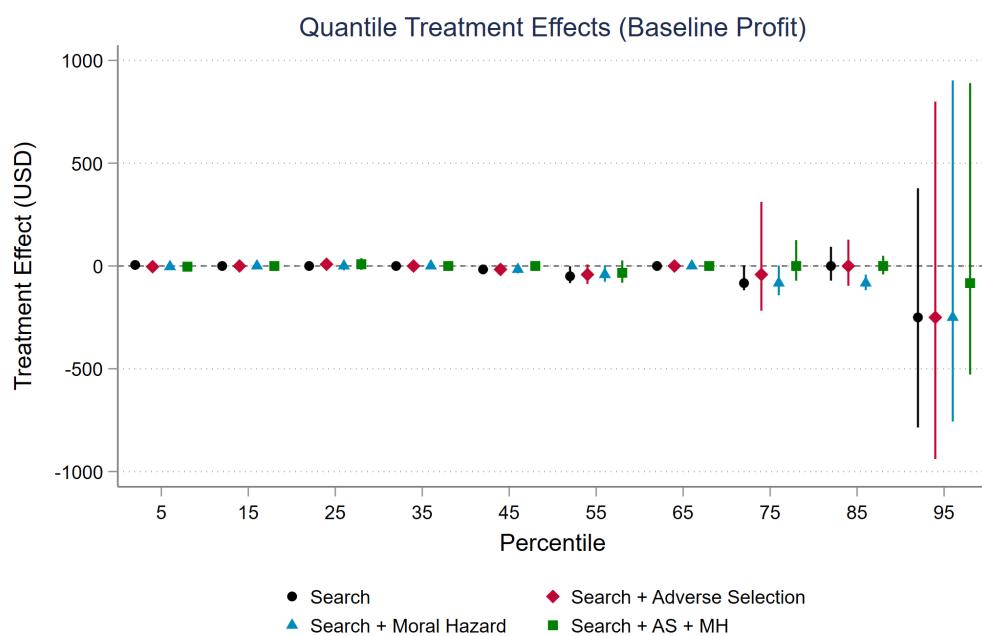
(a) Profit



(b) Sales

*Note:* This figure shows the coefficients from quantile regressions of profit and sales on the four treatment groups, for quantiles 90-99. All quantile regressions include the outcome measured at baseline (if available), but otherwise do not include any covariates. We plot 95% confidence intervals constructed using randomisation inference, defined as the set of sharp nulls that do not reject at the 5% level.

Figure A7: Placebo Check: Quantile Treatment Effects on Baseline Profit



*Note:* This figure shows the coefficients from quantile regressions of baseline profit on the four treatment groups, intended as a placebo test. As the only covariate included in the main quantile regressions is the outcome measured at baseline, which is itself the outcome here, we do not include any covariates. We plot 95% confidence intervals constructed using randomisation inference, defined as the set of sharp nulls that do not reject at the 5% level.

## Appendix B – Mathematical Appendix

In this Appendix, we prove further detail on the model. In particular, we formally state and prove various properties of the optimal contract.

**Assumption 1.** (*No trade with bad types*)

$$\max_q \lambda r(q) - (1 - \xi)cq < \max_q r(q) - p_l q$$

This assumption states that if a firm knew that the supplier was a bad type, they would prefer to order from the local supplier.

**Proposition 1.** *It is not optimal to offer a menu of contracts that fully separates good and bad suppliers.*

*Proof.* Suppose the contrary, and consider the state of the world where the supplier is a bad type. Since the menu fully separates the types, the firm's posterior is then that the supplier is a bad type with probability 1. Because the supplier has limited liability, the maximum that the firm can earn under any such contract is the full surplus, i.e.,  $(\max_q \lambda r(q) - (1 - \xi)q)/(1 - \delta)$ . Assumption 1 implies that the firm can always do better than this, because at the very least they can buy from a local supplier in every period. Since the contract is relational, the firm would thus renege before sending the first transfer. Thus, the expected payoff to the bad type from accepting the revealing contract is 0. So long as the contract recommended to the good type involves positive quantity, the bad type can always earn a positive expected payoff by accepting the good type's contract, because limited liability ensures that  $\tau_t \geq cq_t > (1 - \xi)cq_t$  for all  $t$ . Therefore, the bad type would not accept the contract that reveals their type, which is a contradiction.  $\square$

The original program is as follows:

$$\begin{aligned} L = & \min_{\{\rho_t\}, \{\eta_t\}, \{\gamma_t\}} \max_{\{q_t\}, \{\tau_t\}} \sum_{t=0}^{\infty} \delta^t (1 - \mu_0(1 - \lambda^t)) ((1 - \mu_t(1 - \lambda)r(q_t) - \tau_t + \delta\mu_t(1 - \lambda)\bar{U}) \\ & + \sum_{t=0}^{\infty} \delta^t \rho_t \left[ \sum_{\tau=t+1}^{\infty} \delta^{\tau-t} (R_{\tau} - cq_{\tau}) - \xi cq_t \right] \\ & + \sum_{t=0}^{\infty} \delta^t (1 - \mu_0(1 - \lambda^t)) \eta_t \left[ \sum_{\tau=t}^{\infty} \delta^{\tau-t} (1 - \mu_{\tau}(1 - \lambda^{\tau-t})) ((1 - \mu_{\tau}(1 - \lambda)r(q_t) - \tau_t + \delta\mu_t(1 - \lambda)\bar{U}) - \bar{U}) \right] \\ & + \sum_{t=0}^{\infty} \delta^t \gamma_t [\tau_t - cq_t] \end{aligned}$$

The modified program is as follows.

$$\begin{aligned}
W_t(U_t, V_t, \mu_t) = & y(q_t, \tau_t, \mu_t) + \rho_t (\delta(1 - \lambda)V_{t+1} - \xi c q_{t+1}) \\
& + \eta_t (U_{t+1} - \bar{U}) \\
& + \gamma_t (\tau_t - cq_t) \\
& + \nu_t^b (y(q_t, \tau_t, \mu_t) + \delta(1 - \mu_t(1 - \lambda))U_t + \delta\mu_t(1 - \lambda)\bar{U} - U_t) \\
& + \nu_t^s (\tau_t - cq_t + \delta V_{t+1} - V_t) \\
& + \delta(1 - \mu_t(1 - \lambda))W_{t+1}(U_{t+1}, V_{t+1}, \mu_{t+1})
\end{aligned}$$

The FOCs are as follows

$$(1 - \mu_t(1 - \lambda))r'(q_t)(1 + \nu_t^b) = \rho_t \xi(c - c_0) + (\gamma_t + \nu_t^s)c = 0 \quad (q_t)$$

$$1 + \nu_t^b = \gamma_t + \nu_t^s \quad (\tau_t)$$

$$\rho_t(1 - \lambda) + \nu_t^s = -(1 - \mu_t(1 - \lambda)) \frac{\partial W_{t+1}}{\partial V_{t+1}} \quad (V_{t+1})$$

$$\eta_{t+1} + \delta(1 - \mu_t(1 - \lambda))\nu_t^b = -\delta(1 - \mu_t(1 - \lambda)) \frac{\partial W_{t+1}}{\partial U_{t+1}} \quad (U_{t+1})$$

Substituting the FOC for  $\tau_t$  into the FOC for  $q_t$  and re-arranging gives

$$r'(q_t) = \frac{1}{1 - \mu_t(1 - \lambda)} \left( 1 + \xi \frac{\rho_t}{1 + \nu_t^b} \right) c.$$

The Envelope Condition implies that  $\frac{\partial W_{t+1}}{\partial V_{t+1}} = -\nu_{t+1}^s$  and  $\frac{\partial W_{t+1}}{\partial U_{t+1}} = -\nu_{t+1}^b$ . Combining FOCs 2-4 then gives the following equation relating  $\gamma_t$  and  $\gamma_{t+1}$  that we will make extensive use of in the following proofs.

$$\rho_t(1 - \lambda) + 1 - \gamma_t + \mu_t(1 - \lambda)\nu_t^b = (1 - \gamma_{t+1})(1 - \mu_t(1 - \lambda)) + \frac{\eta_{t+1}}{\delta}. \quad (3)$$

**Remark 1.**  $\nu_{t+1}^b > \nu_t^b \iff \eta_{t+1} > 0$

*Proof.* This follows immediately from the FOC for  $U_{t+1}$  after substituting in the Envelope Condition.  $\square$

The optimal contract is generally not available in closed form, but we state and prove some properties in the following proposition.

**Proposition 2.** *There exists finite  $T^*$  such that:*

1. *The agent earns zero stage profits for all  $t < T^*$  (i.e., LL binds for all  $t < T^*$ ).*

2. *The principal earns  $(1 - \delta)\bar{U}$ , that is, zero stage profits net of their outside option, for all  $t > T^*$  (i.e., DEC binds for all  $t > T^*$ ).*
3.  *$q_t$  is strictly increasing for  $t < T^*$ .*
4. *ICC binds for at least some  $t < T^*$ .*

*Proof.* We prove this through a series of Lemmas, which we state and prove below. For parts 1 and 2, see Lemma 3. For part 3, see Lemma 1. For part 4, see Lemma 5.  $\square$

Before formally stating the Lemmas, we first provide an intuitive sketch of the approach. First, we show that  $q_t$  must be strictly increasing whenever LL binds. Intuitively, a (weakly) decreasing  $q_t$  despite beliefs improving would imply that ICC strongly becomes “more binding” over time. But when LL is binding, the supplier is earning zero stage profits, so the ICC must be getting less binding over time.

Second, we show that if LL in  $t$  binds, then DEC in  $t$  cannot bind. Intuitively, both parties cannot be earning their outside option at the same time, as belief improving and  $q_t$  growing would imply that in other periods one of them must be making a loss.

Third, we show that the problem can be divided into two phases: LL will bind for all early periods and be slack for all late periods. “Backloading” results of this kind are standard in the dynamic moral hazard literature. Backloading happens for two reasons. The first reason is that incentives must be given to the agent at some point, and backloading incentives is efficient because it improves both early and late ICCs (whereas frontloading or even-loading still improves early ICCs but improves late ICCs less). The second reason is that adverse selection means that the “good type” agent is more patient than the principal, as the good type knows their own type. This means that it is always cheaper for the principal to backload payments. As a corollary, the DEC must bind for all late periods, as the principal wants to backload as much as possible, and will continue to do so until the DEC binds.

Finally, we show that ICC must bind for some  $t$  in the early section. Intuitively, if it didn’t, then the principal would just extend the early period—where they earn all the surplus—for longer. The only reason to ever end this early phase is precisely because an ICC eventually binds (the principal has to pay the agent eventually, and in absence of the ICC would always prefer not to).

**Lemma 1.** *If LL binds in  $t + 1$ , then  $q_{t+1} > q_t$ .*

*Proof.* Suppose that LL binds in  $t + 1$  and  $q_{t+1} \leq q_t$ . The FOC for  $q_t$  is

$$r'(q_t) = \frac{1}{1 - \mu_t(1 - \lambda)} \left( 1 + \xi \frac{\rho_t}{1 + \nu_t^b} \right) c.$$

We already know that  $\nu_{t+1}^b \geq \nu_t^b$ , and that  $\mu_{t+1} \leq \mu_t$ . Then, the hypothesis  $q_{t+1} \leq q_t$  implies that  $\rho_{t+1} > \rho_t$ . Then, we can write

$$\delta(1 - \lambda)V_{t+1} < (1 - \lambda)V_{t+1} = (1 - \lambda)(R_{t+1} - cq_{t+1} + \delta V_{t+2}) = \delta(1 - \lambda)V_{t+2} = \xi cq_{t+1} \leq \xi cq_t.$$

The first equality is a definition, the second equality follows from the fact that LL binds in  $t + 1$ , the third equality follows from the fact that  $\rho_{t+1} > \rho_t$  implies that  $\rho_{t+1} > 0$ , which implies that ICC binds in  $t + 1$ . The final inequality follows from the hypothesis that  $q_{t+1} \leq q_t$ . Thus, we have shown that

$$\delta(1 - \lambda)V_{t+1} < \xi cq_t.$$

But this implies that ICC in  $t$  is violated, which is a contradiction.  $\square$

**Lemma 2.** *If LL in  $t + 1$  binds, then DEC in  $t + 1$  is slack.*

*Proof.* Suppose that DEC in  $t + 1$  binds. Then, we can write

$$\begin{aligned} U_t &= (1 - \mu_t(1 - \lambda))r(q_t) - \tau_t + \delta(1 - \mu_t(1 - \lambda))U_{t+1} + \delta\mu_t(1 - \lambda)\bar{U} \\ &= (1 - \mu_t(1 - \lambda))r(q_t) - \tau_t + \delta\bar{U} \\ &\leq (1 - \mu_t(1 - \lambda))r(q_t) - cq_t + \delta\bar{U} \\ &< (1 - \mu_{t+1}(1 - \lambda))r(q_t) - cq_t + \delta\bar{U} \\ &\leq (1 - \mu_{t+1}(1 - \lambda))r(q_{t+1}) - cq_{t+1} + \delta\bar{U} \\ &\leq (1 - \mu_{t+1}(1 - \lambda))r(q_{t+1}) - cq_{t+1} + \delta(1 - \mu_{t+1}(1 - \lambda))U_{t+2} + \delta\mu_{t+1}(1 - \lambda)\bar{U} \\ &= (1 - \mu_{t+1}(1 - \lambda))r(q_{t+1}) - R_{t+1} + \delta(1 - \mu_{t+1}(1 - \lambda))U_{t+2} + \delta\mu_{t+1}(1 - \lambda)\bar{U} \\ &= U_{t+1} \\ &= \bar{U}. \end{aligned}$$

The first line is the definition of  $U_t$ . The second line follows from DEC binding in  $t + 1$ . The third line follows from LL in  $t$ . The fourth line follows from  $\mu_{t+1} < \mu_t$ . The fifth line follows from Lemma 1. The sixth line follows from DEC in period  $t + 2$ , i.e.,  $U_{t+2} \geq \bar{U}$ . The seventh line follows from the hypothesis that LL binds in  $t + 1$ , i.e.,  $\tau_t = cq_t$ . The eighth line is the definition of  $U_{t+1}$ . The ninth line follows from DEC binding in  $t + 1$ .

The above thus establishes that  $U_t < \bar{U}$ . But this is a violation of DEC in period  $t$ , which is a contradiction.  $\square$

**Corollary 1.** *If LL is slack in  $t$ , then it is also slack in  $t + 1$ .*

*Proof.* LL slack in  $t$  means  $\gamma_t = 0$ . Suppose it binds in  $t + 1$ , which means  $\gamma_{t+1} > 0$ . Lemma 2 then implies that  $\eta_{t+1} = 0$ . But, Equation (3) gives

$$\gamma_{t+1}(1 - \mu_t(1 - \lambda)) = \frac{\eta_{t+1}}{\delta} - \rho_t(1 - \lambda) - \mu_t(1 - \lambda)(1 + \nu_t^b).$$

It must then be that  $\eta_{t+1} > 0$ , which is a contradiction.  $\square$

**Corollary 2.** *If LL is slack in  $t$ , then DEC binds in  $t + 1$ .*

*Proof.* Corollary 1 implies that  $\gamma_t = \gamma_{t+1} = 0$ . Then, the FOC implies

$$\frac{\eta_{t+1}}{\delta} = \rho_t(1 - \lambda) + \mu_t(1 - \lambda)(1 + \nu_t^b) \geq \mu_t(1 - \lambda)(1 + \nu_t^b).$$

The final term is strictly positive, which implies that  $\eta_{t+1} > 0$ .  $\square$

**Lemma 3.** *If any trade occurs, then there exists finite  $T^* \geq 1$  such that (i) LL binds for all  $t < T^*$  and is slack for all  $t \geq T^*$ , and (ii) DEC is slack for all  $t < T^*$  and binds for all  $t > T^*$ .*

*Proof.* For (i): Corollary 1 shows that  $\gamma_t = 0 \implies \gamma_{t+1} = 0$ . Thus, if there exists  $T^*$  such that  $\gamma_t = 0$ , then  $\gamma_s = 0$  for all  $s \geq T^*$ . We already know that  $\gamma_0 = 1$ , so this is not the case for  $t = 0$ . Suppose that  $\gamma_t > 0$  for all  $t$ . Then, the supplier earns zero profit, which implies that all ICCs will fail unless  $q_t = 0$  for all  $t$ , which violates the supposition that trade occurs at some point. Thus, there must be at least one  $t \geq 1$  such that LL is slack in  $t$ . If there are multiple, define  $T^*$  as the earliest such  $t$ .

For (ii): Since LL is slack for all  $t \geq T^*$ , Corollary 2 implies that DEC binds for all  $t > T^*$ .  $\square$

**Corollary 3.**  $\nu_t^b = 0$  for all  $t \leq T^*$ , and  $\nu_{t+1}^b > 0$  for all  $t > T^*$ .

*Proof.* This follows from Remark 1 and Lemma 3.  $\square$

**Lemma 4.**  $\gamma_{t+1} \leq \gamma_t$ , with inequality strict if  $\gamma_t \in (0, 1)$ .

*Proof.* If  $\gamma_{t+1} = 0$ , then this holds trivially. We thus need to establish the claim for  $\gamma_{t+1} > 0$ . Suppose then that  $\gamma_{t+1} > \gamma_t$  with  $\gamma_{t+1} > 0$ . Then, Equation (3) implies

$$0 > \rho_t(1 - \lambda) - \frac{\eta_{t+1}}{\delta} + \mu_t(1 - \lambda)(1 + \nu_t^b - \gamma_{t+1}).$$

Since  $\gamma_{t+1} > 0$ , Lemma 2 implies that  $\eta_{t+1} = 0$ . We are thus left with

$$\begin{aligned} 0 &> \rho_t(1 - \lambda) + \mu_t(1 - \lambda)(1 + \nu_t^b - \gamma_{t+1}) \\ &= \rho_t(1 - \lambda) + \mu_t(1 - \lambda)(1 + \nu_{t+1}^b - \gamma_{t+1}) \\ &= \rho_t(1 - \lambda) + \mu_t(1 - \lambda)\nu_{t+1}^s \end{aligned}$$

where the first equality follows from the fact that the FOC for  $U_{t+1}$  implies that  $\eta_{t+1} = 0 \implies \nu_t^b = \nu_{t+1}^b$ , and the second equality follows from the FOC for  $R_{t+1}$ . But the RHS is weakly positive, so this is a contradiction, which establishes that  $\gamma_{t+1} \leq \gamma_t$ .

Then, to establish the claim about strict inequality, suppose instead that  $\gamma_{t+1} = \gamma_t$ . We instead have

$$0 = \rho_t(1 - \lambda) + \mu_t(1 - \lambda)\nu_{t+1}^s,$$

which is only possible if  $\rho_t = \nu_{t+1}^s = 0$ . But this implies that  $\gamma_{t+1} = 1 + \nu_t^b \geq 1$ . If  $\gamma_t \in (0, 1)$ , this implies  $\gamma_{t+1} > \gamma_t$ , which is a contradiction.  $\square$

**Lemma 5.** *ICC binds for some  $t < T^*$ .*

*Proof.* We prove this by iterating forward Equation (3). Since  $\eta_t = \nu_t^b = 0$  for all  $t < T^*$ , the equation can be written

$$\gamma_t = \mu_t(1 - \lambda) + \rho_t(1 - \lambda) + \gamma_{t+1}(1 - \mu_t(1 - \lambda)).$$

Starting with  $\gamma_0 = 1$  and iterating this until  $T^* - 1$ , for which  $\gamma_{T^*} = 0$ , we get

$$1 + (1 - \mu_0(1 - \lambda^{T^*}))\frac{\eta_{T^*}}{\delta} = \sum_{t=0}^{T^*-1} (1 - \mu_0(1 - \lambda^t))[\mu_t(1 - \lambda) + \rho_t(1 - \lambda)].$$

Note that  $LHS \geq 1$ . The first term on the RHS simplifies to

$$\sum_{t=0}^{T^*-1} (1 - \mu_0(1 - \lambda^t))\mu_t(1 - \lambda) = (1 - \lambda)\mu_0 \sum_{t=0}^{T^*-1} \lambda^t = \mu_0(1 - \lambda^{T^*}) < 1.$$

Thus, it cannot be that  $\rho_t = 0$  for all  $t \leq T^* - 1$ , as otherwise  $LHS > RHS$ . So there must

be some  $t \leq T^*$  for which the ICC binds.

□

## **Appendix C – Main Tables without Covariates**

We pre-specified that we would use the specification in Equation (1). Nonetheless, in this section, we replicate all of the main tables in the analysis using the following simpler regression specification that does not include any covariates:

$$y_i = \alpha + \sum_{j=1}^4 \beta_j T_{ji} + \varepsilon_i.$$

Table C1: Access to Foreign Goods (No Covariates)

	Horizontal		Vertical			Price
	Find Product ≥ 3 Criteria	High Quality Dummy	Quality Score (/50)	Made in Turkey	Index	Price (USD)
	(1)	(2)	(3)	(4)	(5)	(6)
<i>Panel A: Pooled</i>						
Treatment	0.101*** (0.031)	0.107 (0.068)	-0.636 (0.591)	0.145** (0.068)	0.456*** (0.139)	1.656** (0.840)
<i>Panel B: Individual Treatments</i>						
Search Only	0.141*** (0.039) [0.003]	0.139* (0.084) [0.231]	-0.177 (0.683) [0.951]	0.112 (0.081) [0.315]	0.492*** (0.174) [0.025]	2.444** (0.991) [0.051]
Search + Adverse Selection	0.062 (0.039) [0.116]	0.154* (0.087) [0.229]	-0.026 (0.845) [0.966]	0.108 (0.085) [0.315]	0.360* (0.179) [0.058]	1.409 (1.049) [0.287]
Search + Moral Hazard	0.121*** (0.039) [0.005]	0.076 (0.084) [0.577]	-1.284 (0.843) [0.357]	0.154* (0.080) [0.139]	0.452** (0.170) [0.025]	1.746* (0.976) [0.171]
Search + AS + MH	0.076** (0.039) [0.088]	0.062 (0.085) [0.577]	-0.934 (0.846) [0.548]	0.203** (0.081) [0.053]	0.505*** (0.178) [0.025]	0.932 (0.994) [0.347]
Control Mean	0.357	0.431	43.064	0.477	0.000	19.990
% Increase (Pooled)	28.3%	24.8%	-1.5%	30.4%	N/A	8.3%
Adjusted $R^2$	0.01	0.00	0.32	0.09	0.01	0.23
$N$	1579	359	359	361	361	642

*Note:*  $p$ -values are computed using randomisation inference. Specifically, we compute the randomised randomisation- $t$   $p$ -value from Young (2019) using 2000 reps. \*  $p < 0.1$  \*\*  $p < 0.05$  \*\*\*  $p < 0.01$ . We report conventional robust standard errors in parentheses, although we do not use these directly for inference. We also report Romano-Wolf multiple-testing adjusted  $p$ -values in square brackets (Romano and Wolf, 2005). Panel A shows the coefficient from a regression on an indicator that pools all treated groups. Panel B shows the coefficients corresponding to treatment indicators for each of the four treatment arms.

Column 1 is an indicator that is one if the merchant finds a good that matches at least 3 horizontal criteria, and is missing if the merchant never replied to the mystery shopper or was otherwise unreachable. Column 2 is an indicator for whether the good's quality score is above the median product-group quality score. Column 3 is the raw quality score. Column 4 is an indicator for whether the good is made in Turkey, primarily inferred based on whether the label says . See the text for full details of how this outcome is constructed. Column 5 is the Anderson (2008) index combining the vertical outcomes. Column 6 is the price in USD, which is only measured conditional on the firm finding a good matching at least three horizontal criteria.

Table C2: Supplier Relationships (Followup Survey) (No Covariates)

	Regular Suppliers in Turkey			Previous Suppliers		
	Any Reg Sup in Turkey	Num Reg Sup in Turkey	Index	Num Reg Sup Total	Num Reg Sup in Senegal	Ended with Reg Sup
	(1)	(2)	(3)	(4)	(5)	(6)
<i>Panel A: Pooled</i>						
Treatment	0.035 (0.023)	0.079** (0.038)	0.138** -0.032 (0.062) (0.193)	-0.092 (0.191)	0.070*** (0.022)	
<i>Panel B: Individual Treatments</i>						
Search Only	0.018 (0.029) [0.762]	0.069 (0.055) [0.325]	0.118 0.092 (0.081) (0.269) [0.240] [0.854]	0.088 (0.266) [0.917]	0.062** (0.029) [0.061]	
Search + Adverse Selection	0.056** (0.030) [0.127]	0.077* (0.048) [0.240]	0.172** -0.154 (0.084) (0.253) [0.101] [0.854]	-0.236 (0.249) [0.736]	0.067** (0.029) [0.050]	
Search + Moral Hazard	-0.001 (0.029) [0.974]	-0.009 (0.043) [0.835]	0.023 -0.197 (0.078) (0.248) [0.789] [0.818]	-0.168 (0.244) [0.817]	0.091*** (0.029) [0.008]	
Search + AS + MH	0.069** (0.031) [0.076]	0.183*** (0.062) [0.007]	0.244*** 0.139 (0.087) (0.258) [0.018] [0.854]	-0.049 (0.251) [0.917]	0.057** (0.029) [0.061]	
Control Mean	0.167	0.222	0.000 3.700	3.213	0.135	
% Increase (Pooled)	21.0%	35.6%	N/A -0.9%	-2.9%	51.9%	
Adjusted $R^2$	0.00	0.01	0.00 -0.00	-0.00	0.00	
$N$	1680	1680	1680 1681	1681	1671	

*Note:*  $p$ -values are computed using randomisation inference. Specifically, we compute the randomised randomisation- $t$   $p$ -value from Young (2019) using 2000 reps. \*  $p < 0.1$  \*\*  $p < 0.05$  \*\*\*  $p < 0.01$ . We report conventional robust standard errors in parentheses, although we do not use these directly for inference. We also report Romano-Wolf multiple-testing adjusted  $p$ -values in square brackets (Romano and Wolf, 2005). Panel A shows the coefficient from a regression on an indicator that pools all treated groups. Panel B shows the coefficients corresponding to treatment indicators for each of the four treatment arms.

Column 1 is 1 if the merchant says that they have a regular supplier in Turkey. Column 2 is the number of regular suppliers in Turkey. Column 3 is the Anderson (2008). Column 4 is the total number of regular suppliers. Column 5 is the number of regular suppliers in Senegal. Column 6 is 1 if the merchant has ended a relationship with a regular supplier in the past 3 months. A regular supplier is defined as a supplier from whom the merchant has made two or more orders with an intention of continuing the relationship.

Table C3: Order Value (Mobile Money Data) (No Covariates)

	Any Order (1)	Value Post Mystery Shopping		Total Value	
		Order Value (OLS) (2)	Order Value (Poisson) (3)	Order Value (OLS) (4)	Order Value (Poisson) (5)
<i>Panel A: Pooled</i>					
Trust Treatment	0.000 (0.020)	4.684* (2.402)	1.246* (0.508)	4.382 (3.202)	0.428 (0.300)
<i>Panel B: Individual Treatments</i>					
Search + Adverse Selection	0.003 (0.025) [0.896]	1.713 (1.547) [0.504]	0.645 (0.531) [0.425]	2.173 (2.934) [0.745]	0.235 (0.310) [0.686]
Search + Moral Hazard	0.013 (0.025) [0.837]	6.386*** (2.908) [0.023]	1.476** (0.502) [0.062]	5.335 (3.870) [0.357]	0.501 (0.331) [0.438]
Search + AS + MH	-0.017 (0.024) [0.825]	5.983 (6.303) [0.504]	1.426 (0.876) [0.420]	5.674 (7.263) [0.745]	0.525 (0.549) [0.686]
Control Mean	0.134	1.891	1.891	8.209	8.209
% Increase (Pooled)	0.0%	247.7%	247.6%	53.4%	53.4%
Adjusted $R^2$	-0.00	0.00	0.03	-0.00	0.01
N	1500	1500	1500	1500	1500

Note:  $p$ -values are computed using randomisation inference. Specifically, we compute the randomised randomisation- $t$   $p$ -value from Young (2019) using 5000 reps. \*  $p < 0.1$  \*\*  $p < 0.05$  \*\*\*  $p < 0.01$ . We report conventional robust standard errors in parentheses, although we do not use these directly for inference. We also report Romano-Wolf multiple-testing adjusted  $p$ -values in square brackets (Romano and Wolf, 2005). Panel A shows the coefficient from a regression on an indicator that pools all trust treated groups, where Search Only is the omitted category. Panel B shows the coefficients corresponding to treatment indicators for each of the three treatment groups with trust treatments.

Column 1 is an indicator for whether the firm ever ordered from a study supplier. Column 2 is the total value of orders. Column 3 is the total value of orders, analysed with Poisson regression. Column 4 is the total value of orders. Column 5 is the total value of orders, analysed with Poisson regression. Mystery shopping took place during the first 13 weeks of the study. All values are in USD.

Table C4: Profit and Sales (No Covariates)

	Raw			Winsorized (1%)		
	Profit (USD) (1)	Sales (USD) (2)	Index (3)	Profit (USD) (4)	Sales (USD) (5)	Index (6)
<i>Panel A: Pooled</i>						
Treatment	57.3*	204.0*	0.213**	24.1	83.0	0.109
	(28.8)	(108.1)	(0.081)	(20.9)	(83.6)	(0.069)
<i>Panel B: Individual Treatments</i>						
Search Only	.566 (29.1) [0.988]	254 (187) [0.392]	.228** (.115) [0.109]	-5.26 (26.2) [0.836]	107 (125) [0.570]	.126 (.096) [0.411]
Search + Adverse Selection	30.6 (36.4) [0.630]	56 (119) [0.662]	.069 (.092) [0.482]	18.4 (31.5) [0.746]	38.1 (111) [0.726]	.035 (.092) [0.880]
Search + Moral Hazard	-27.4 (24.8) [0.572]	-103 (103) [0.522]	.106 (.078) [0.325]	-27.4 (24.8) [0.515]	-121 (93.6) [0.396]	.033 (.076) [0.880]
Search + AS + MH	244*** (93.2) [0.012]	640** (282) [0.047]	.463** (.196) [0.043]	120*** (45.2) [0.022]	325** (152) [0.081]	.253** (.112) [0.081]
Control Mean	188.3	609.5	0.000	188.3	609.5	0.000
% Increase (Pooled)	30.4%	33.5%	N/A	12.8%	13.6%	N/A
Adjusted $R^2$	0.01	0.01	0.01	0.01	0.01	0.00
$N$	1351	1378	1431	1351	1378	1431

Note:  $p$ -values are computed using randomisation inference. Specifically, we compute the randomised randomisation- $t$   $p$ -value from Young (2019) using 2000 reps. \*  $p < 0.1$  \*\*  $p < 0.05$  \*\*\*  $p < 0.01$ . We report conventional robust standard errors in parentheses, although we do not use these directly for inference. We also report Romano-Wolf multiple-testing adjusted  $p$ -values in square brackets (Romano and Wolf, 2005). Panel A shows the coefficient from a regression on an indicator that pools all treated groups. Panel B shows the coefficients corresponding to treatment indicators for each of the four treatment arms.

Column 1 is total profit from the past 30 days in USD. Column 2 is total sales from the past 30 days in USD. Column 3 is the Anderson (2008) index combining the previous two columns. Column 4 is total profit from the past 30 days in USD, winsorizing the top 1%. Column 5 is total sales from the past 30 days in USD, winsorizing the top 1%. Column 6 is the Anderson (2008) index combining the previous two columns. Profit is measured using the survey question from De Mel, McKenzie, and Woodruff (2009). Sales is measured using a similar survey question.

## Appendix D – Further Details on Solving and Estimating the Model

In this Appendix, we provide more detail on our algorithm to numerically solve and estimate the model.

### Solving the Model

As is often the case in dynamic optimisation, the original infinite horizon program is very difficult to work with directly. It is much more tractable to find a way to work with a recursive formulation. The challenge is that, unlike standard dynamic problems encountered in macro, we have two constraints (the DEC and DICC) that are forward looking, and, in particular, forward-looking to an infinite horizon.

The literatures on dynamic moral hazard and limited commitment typically deal with this in one of two ways. One way is to define continuation values as state variables, which completely summarise the future and thus allow the constraints to be written recursively. This “promised utility” approach was originally developed somewhat independently in different theoretical contexts by Spear and Srivastava (1987), Abreu, Pearce, and Stacchetti (1990), and Thomas and Worrall (1988), and in fact we take this approach in Appendix B when we derive qualitative properties of the optimal contract.

The other approach, pioneered by Marcket and Marimon (2019), follows the idea that the original Lagrangean can be rewritten recursively as a pseudo planner’s problem, where the Pareto weights are state variables that evolve endogenously to perfectly summarise historical binding constraints. Intuitively, if the DICC is binding in period 0, which implies that the agent must be delivered a certain amount of utils at some point in the future, the Pareto weight on the agent increases over time to ensure that the planner delivers precisely the required amount of utils. This allows the problem to be written recursively because the principal can trade off the benefit of making a constraint “more binding” today against the cost of increasing next period’s Pareto weight on the agent. The recursive formulation delivers a Saddle Point Functional Equation, which is analogous to the familiar Bellman Equation but for saddle point problems, which satisfies a number of familiar properties that permit the use of dynamic programming techniques.

The key advantage of the Marcket and Marimon (2019) approach over the promised utility approach is that the feasible set of Pareto weights is known. This is important, because one needs to know the feasible set in order to numerically solve the model. In contrast, in the promised utility approach, we would need to know the feasible set of continuation values, which are endogenous objects that likely depend in complicated ways upon the model parameters. This is not a problem for qualitatively analysing the model, which

is why we use this approach in Appendix B, but is a problem for numerically solving it. It is also not an insurmountable obstacle: Abreu et al. (1990) provide an algorithm that can be used within an inner loop to compute the feasible set, and many papers in the literature fruitfully pursue this approach. Nonetheless, the Pareto weight approach completely sidesteps this issue, which is why we use it here.

We first rewrite our problem as a recursive Lagrangean and thus derive the Saddle Point Functional Equation. Define  $y(q_t, \tau_t, \mu_t) \equiv (1 - \mu_t(1 - \lambda))r(q_t) - \tau_t$ . Then, the original dynamic program is as follows

$$\begin{aligned} & \max_{\{q_t\}, \{\tau_t\}} y(q_0, \tau_0, \mu_0) + \delta\mu_0(1 - \lambda)\bar{U} \\ & \quad + \delta(1 - \mu_0(1 - \lambda))[y(q_1, \tau_1, \mu_1) + \delta\mu_1(1 - \lambda)\bar{U}] \\ & \quad + \delta(1 - \mu_1(1 - \lambda))[y(q_2, \tau_2, \mu_2) + \delta\mu_2(1 - \lambda)\bar{U}] \\ & \quad + \delta(1 - \mu_2(1 - \lambda))[...] \end{aligned}$$

subject to

$$\begin{aligned} & \sum_{n=1}^{\infty} \delta^n (\tau_{t+n} - cq_{t+n}) \geq \xi cq_t \quad \forall t \\ & U_t \geq \bar{U} \quad \forall t \\ & \tau_t \geq cq_t \quad \forall t, \end{aligned}$$

and with  $\mu_t$  evolving according to Bayes' Rule. Rewriting the objective function as an infinite sum and including the constraints with Lagrange multipliers, the program can be expressed as follows

$$\begin{aligned} L = & \min_{\{\rho_t\}, \{\eta_t\}, \{\gamma_t\}} \max_{\{q_t\}, \{\tau_t\}} \sum_{t=0}^{\infty} \delta^t (1 - \mu_0(1 - \lambda^t)) ((1 - \mu_t(1 - \lambda)r(q_t) - \tau_t + \delta\mu_t(1 - \lambda)\bar{U}) \\ & + \sum_{t=0}^{\infty} \delta^t \rho_t \left[ \sum_{\tau=t+1}^{\infty} \delta^{\tau-t} (R_{\tau} - cq_{\tau}) - \xi cq_t \right] \\ & + \sum_{t=0}^{\infty} \delta^t (1 - \mu_0(1 - \lambda^t)) \eta_t \left[ \sum_{\tau=t}^{\infty} \delta^{\tau-t} (1 - \mu_{\tau}(1 - \lambda^{\tau-t})) ((1 - \mu_{\tau}(1 - \lambda)r(q_t) - \tau_t + \delta\mu_t(1 - \lambda)\bar{U}) - \bar{U} \right] \\ & + \sum_{t=0}^{\infty} \delta^t \gamma_t [\tau_t - cq_t]) \end{aligned}$$

Then, with some algebra, we can collect the Lagrange terms directly inside the first infinite sum to express the Lagrangean as a function of "Pareto weights", Lagrange multipliers,

state variables, and time-invariant functions.

$$L = \min_{\{\rho_t\}, \{\eta_t\}, \{\gamma_t\}} \max_{\{q_t\}, \{\tau_t\}} \sum_{t=0}^{\infty} \delta^t [(\zeta_t^b + \eta_t) \beta_t h_0^b(q_t, \tau_t, \mu_t) + \eta_t \beta_t h_1^b(q_t, \tau_t, \mu_t) \\ + (\zeta_t^s + \gamma_t) h_0^s(q_t, \tau_t) + \rho_t h_1^s(q_t, \tau_t)]$$

where  $\zeta_t^b \equiv \zeta_0^b + \sum_{\tau=1}^{t-1} \eta_\tau$ ,  $\zeta_t^s \equiv \zeta_0^s + \sum_{\tau=1}^{t-1} \rho_\tau$ ,  $\beta_t = \prod_{s=0}^{t-1} (1 - \mu_s(1 - \lambda))$ , and

$$\begin{aligned} h_0^b(q_t, \tau_t, \mu_t) &\equiv [(1 - \mu_t(1 - \lambda))r(q_t) - \tau_t + \delta \mu_t(1 - \lambda)\bar{U}] \\ h_1^b(q_t, \tau_t, \mu_t) &\equiv -\bar{U} \\ h_0^s(q_t, \tau_t) &\equiv \tau_t - cq_t \\ h_1^s(q_t, \tau_t) &\equiv -\xi cq_t \end{aligned}$$

$\zeta_t^b$  and  $\zeta_t^s$  behave like Pareto weights and are equal to the sum of all prior Lagrange multipliers from the forward-looking constraints for the principal (the buyer) and the agent (the seller), respectively. This is the sense in which they fully summarise the shadow cost of constraints from earlier periods of the problem. For example, if DICC is “very binding” in early periods, meaning  $\rho_t$  is large for early  $t$ , then this is reflected in a large value of  $\zeta_t^s$  for later  $t$ , which causes the “planner” to endogenously choose a high value of  $\tau_t$  and thus give the agent utility.

We can then write this recursively as a Saddle Point Functional Equation (SPFE) as follows,

$$W(\zeta_t^b, \zeta_t^s, \mu_t, \beta_t) = \min_{\eta_t, \rho_t, \gamma_t} \max_{q_t, \tau_t} (\zeta_t^b + \eta_t) \beta_t h_0^b(q_t, \tau_t, \mu_t) + \eta_t \beta_t h_1^b(q_t, \tau_t, \mu_t) \\ + (\zeta_t^s + \gamma_t) h_0^s(q_t, \tau_t) + \rho_t h_1^s(q_t, \tau_t) + \delta W(\zeta_{t+1}^b, \zeta_{t+1}^s, \mu_{t+1}, \beta_{t+1}),$$

subject to

$$\begin{aligned} \zeta_{t+1}^b &= \zeta_t^b + \eta_t \\ \zeta_{t+1}^s &= \zeta_t^s + \rho_t \\ \beta_{t+1} &= \beta_t(1 - \mu_t(1 - \lambda)), \end{aligned}$$

and  $\mu_{t+1} = \mu_t \lambda / (1 - \mu_t(1 - \lambda))$  if  $q_t > 0$  and high quality is observed,  $\mu_{t+1} = 1$  if  $q_t > 0$  and low quality is observed, and  $\mu_{t+1} = \mu_t$  if  $q_t = 0$ .

The Saddle Point Functional Equation is analogous to the Bellman Equation for saddle point problems, and Marimon (2019) prove that—under some regularity conditions—it has the usual desirable properties associated with dynamic programming problems. In

particular, this means that we can obtain solutions to the original problem by using dynamic programming techniques to solve for the value function and policy functions associated with the SPFE.

We solve the model using value function iteration. Since  $W$  is homogeneous of degree one in the Pareto weights, note that we can write  $W(\zeta_t^b, \zeta_t^s, \mu_t, \beta_t) = \zeta_t^b W(1, \frac{\zeta_t^s}{\zeta_t^b}, \mu_t, \beta_t)$ . This means that we can eliminate one state variable. We define a discrete grid over the three state variables and interpolate over the grid using a shape-preserving spline. Within each iteration, the FOCs for  $\tau_t$  combined with our results in Appendix B allow us to obtain analytical solutions for  $\eta_t$ ,  $\gamma_t$ , and  $\tau_t$  as a function of the state variables. We can then use the FOCs for  $q_t$  and  $q_{ot}$  to obtain analytically solve for these. Unfortunately, the FOC for  $\rho_t$  involves a derivative of the value function so we cannot obtain analytical solutions. Since  $\rho_t$  is bounded below at zero, we first evaluate the derivative at zero. If it is positive, we set  $\rho_t^* = 0$ ; otherwise, we use a numerical minimiser to solve for  $\rho_t^*$ .

Once we have obtained the value function, we iterate forward from the initial conditions at  $t = 0$  to get the solution path. Since the agent's utility is linear in  $\tau_t$ , the value function has a kink in the neighbourhood of  $\beta_t \zeta_t^b = \zeta_t^s$ , which Marimon and Werner (2021) show results in inconsistent promises. We resolve this by imposing their Envelope Selection Condition.

## Estimation

The above sub-section describes how we solve the model for a given guess of the parameters. In order to estimate the parameters, we need to find the parameters that best match the empirical moments, and thus solve the model for many combinations of parameters. We do this in two steps. First, we solve the model for a grid of  $(\mu_0, \xi, c)$ , and interpolate over the grid using a shape-preserving spline. This gives *functions* for all of the relevant theoretical objects, such as  $y_t(\mu_0, \xi, c)$ . This means that we do not need to further solve the model as we can simply evaluate these functions at a given  $(\mu_0, \xi, c)$ . Second, we use the above functions to compute the theoretical moments and then use SMM to estimate the parameters.