

# ARMA Processes

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## Non Normal residuals

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As it is stated in Introduction to Time Series and Forecasting, by P.J. Brockwell and R.A. Davis 1996, Cap 5.2 pp 157:

The likelihood (5.2.1) of the vector  $\mathbf{X}_n$  therefore reduces to

$$L(\Gamma_n) = \frac{1}{\sqrt{(2\pi)^n v_0 \cdots v_{n-1}}} \exp \left\{ -\frac{1}{2} \sum_{j=1}^n (X_j - \hat{X}_j)^2 / v_{j-1} \right\}. \quad (5.2.6)$$

If  $\Gamma_n$  is expressible in terms of a finite number of unknown parameters  $\beta_1, \dots, \beta_r$  (as is the case when  $\{X_t\}$  is an ARMA( $p, q$ ) process) the **maximum likelihood estimators** of the parameters are those values that maximize  $L$  for the given data set. When  $X_1, X_2, \dots, X_n$  are iid, it is known, under mild assumptions and for  $n$  large, that maximum likelihood estimators are approximately normally distributed with variances that are at least as small as those of other asymptotically normally distributed estimators (see e.g., Lehmann, 1983).

Even if  $\{X_t\}$  is not Gaussian, it still makes sense to regard (5.2.6) as a measure of goodness of fit of the model to the data, and to choose the parameters  $\beta_1, \dots, \beta_r$  in such a way as to maximize (5.2.6). We shall always refer to the estimators,  $\hat{\beta}_1, \dots, \hat{\beta}_r$ , so obtained as “maximum likelihood” estimators, even when  $\{X_t\}$  is not Gaussian. Regardless of the joint distribution of  $X_1, \dots, X_n$ , we shall refer to (5.2.1) and its algebraic equivalent (5.2.6) as the “likelihood” (or “Gaussian likelihood”) of  $X_1, \dots, X_n$ . A justification for using maximum Gaussian likelihood estimators of ARMA coefficients is that the large-sample distribution of the estimators is the same for  $\{Z_t\} \sim \text{IID}(0, \sigma^2)$ , regardless of whether or not  $\{Z_t\}$  is Gaussian (see TSTM, Section 10.8).





