# **Analytical Risks and Natural Sensitivities**

by Olivier Croissant

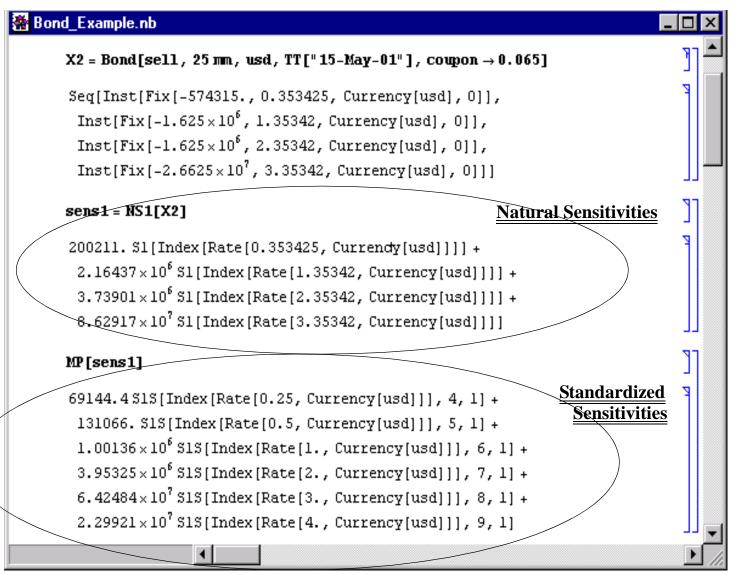
## What we are Looking for

- Fast Risk and Hedge Ratios Computation
- Comprehensive Risk Factor Handling
- Liquidity Risk Factor Handling
- Non Linear Instruments
- Auditability of the Risks
- Engineering of Hedge Ratios

#### How we will Achieve these Goals

- Define and compute sensitivities --> natural sensitivities
- Define standardized sensitivities --> Mappings
- Define non linear risk measure --> cumulants
- Define liquidity risk with cumulants and sensitvities
- Generalized hedge ratios : geometrization of the hedge ratios

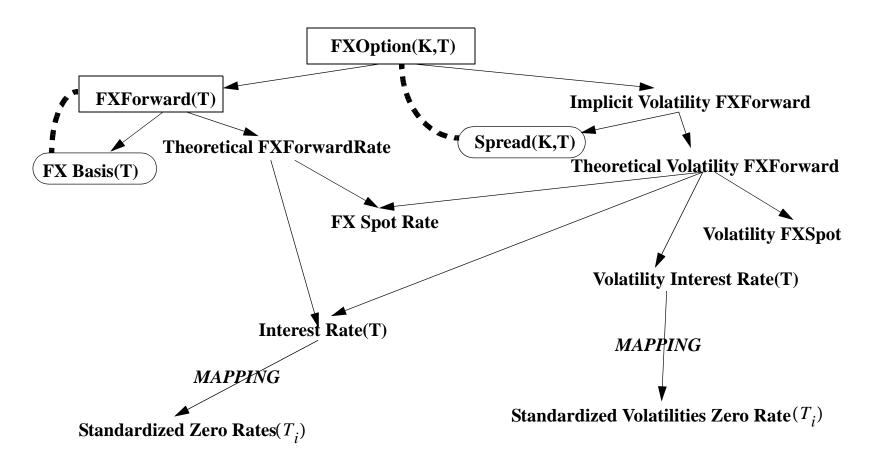
#### **Sensitivities**



### A More Complexe Example

```
👺 FXOption.nb
ln[12]:= inst =
       FxOption[buy, call, 100 mm, usd, TT["30-Mar-98"], currencybase \rightarrow dem, strike \rightarrow 1.5]
ln[21]:= ns1 = NS1a[inst]
                                                                               Top level Natural Sensitivities
Out[24 1.56921 x 106 S1 [Index [ExchangeRate [Currency[dem], Currency[usd]]]] -
        49669.7 S1 [Index [Rate [0, 0.0657534, Currency [dem]]]] +
        2.40549 x 107 S1 [Index [ExchangeRate [0.0657534, Currency [usd], Currency [dem]]], 0] +
        4.48115×107 S1[Volatility][
           Index[ExchangeRate[0.0657534, Currency[usd], Currency[dem]]], 0.0657534, 0.1
          0]
ln[15]:= MP[ns1]
Out[15]= -4.47051 \times 10^{-3}
        515[Index[ExchangeRate[Currency[dem], Currency[usd]]], 2] + 2.61083 × 10 Standardized Sensitivities
         S1S[Volatility[Index[ExchangeRate[Currency[dem], Currency[usd]]]], 2]
        1.48702 x 106 S1S[Index[Rate[0.0194444, Currency[dem]]], 2, 2] +
        1.45159 × 106 S1S[Index[Rate[0.0194444, Currency[usd]]], 2, 1] -
        745993. S1S[Index[Rate[0.0833333, Currency[dem]]], 3, 2] +
        732435. S1S[Index[Rate[0.0833333, Currency[usd]]], 3, 1] +
        354.266 S1S[Volatility[Index[Rate[0.00277778, Currency[dem]]]], 1, 2] -
        516.842 S1S[Volatility[Index[Rate[0.00277778, Currency[usd]]]], 1, 1] +
        12.6911 S1S[Volatility[Index[Rate[1, Currency[dem]]]], 2, 2] -
        16.3948 SIS[Volatility[Index[Rate[1, Currency[usd]]]], 2, 1]
```

### **Hierachy of Natural Market Descriptors**



- Node = Pricing Function = F(X1,X2,X3,...Xn)
- MAPPING = Interpolation function = Discretisation

### **Natural Sensitivities (First Idea)**

• Compounding of Pricing Function -> Modularity of the differentials

$$Z = f(y_1, y_2)$$
  $y_1 = g(x_1)$   $y_2 = h(x_2)$   $y_1$   $y_2$   $y_2$   $y_3$   $y_4$   $y_2$   $y_4$   $y_2$ 

• the Chain Rule of First Derivatives : we do not need the full space of risk factors

$$dZ = \frac{\partial f}{\partial y_1} dy_1 + \frac{\partial f}{\partial y_2} dy_2 \qquad dy_1 = \frac{dg}{dx_1} dx_1 \qquad dy_2 = \frac{dh}{dx_2} dx_2$$
First Derivatives of h
$$dZ = \frac{\partial f}{\partial y_1} \frac{dg}{dx_1} dx_1 + \frac{\partial f}{\partial y_2} \frac{dh}{dx_2} dx_2 = \begin{bmatrix} \frac{\partial f}{\partial y_1} \\ \frac{\partial f}{\partial y_2} \end{bmatrix} \begin{bmatrix} \frac{dg}{dx_1} dx_1 & \frac{dh}{dx_2} dx_2 \\ \frac{\partial f}{\partial y_2} \end{bmatrix}$$

First Derivatives of f

First Derivatives of g

### **Modularity of the Second Derivatives**

- First differential dZ : Z(world + d(world)) Z(world) as a linear function of d(world)
- Second differential  $d^2Z$ : Z(world + d(world)) Z(world) dZ as a quadratic function of d(world)
- if we have  $Z = f(y_1, y_2, ..., y_n)$  and  $y_i = y_i(x_1, x_2, ..., x_p)$

• -> 
$$\frac{\partial^2}{\partial x_j \partial x_k} Z = \sum_{k=1}^n \sum_{l=1}^n \frac{\partial^2}{\partial y_l \partial y_m} Z \cdot \frac{\partial y_l}{\partial x_j} \cdot \frac{\partial y_m}{\partial x_k} + \sum_{l=1}^n \frac{\partial}{\partial y_l} Z \cdot \frac{\partial^2}{\partial x_j \partial x_k} y_l$$

• another way to say that is :  $d^2Z = \sum_{i,j=1}^{n} \frac{\partial^2Z}{\partial y_i \partial y_j} dy_i \otimes dy_j + \sum_{i=1}^{n} \frac{\partial Z}{\partial y_i} d^2y_i$  Quadratic Forms

Sensitivities (order 1 and 2) of Z

### **Modularity of the Natural Sensitivities**

• By Including the first order :

$$[dZ, d^2Z] - \text{Output of the node}$$

$$= \left[ \sum_{i=1}^{n} \frac{\partial Z}{\partial y_i} d^2y_i, \sum_{i,j=1}^{n} \frac{\partial^2 Z}{\partial y_i \partial y_j} dy_i \otimes dy_j + \sum_{i=1}^{n} \frac{\partial Z}{\partial y_i} d^2y_i \right]$$

$$[dy_i, d^2y_i] - \text{Input of the node}$$

- We compute a relative linear form and a relative quadratic form , <u>not a delta vector</u> <u>nor a gamma matrix</u>
- The difference is the same than between a sparse vector description and a full vector and between a sparse matrix description and a full matrix.
- Naturally every thing extends to higher order (but less interesting)

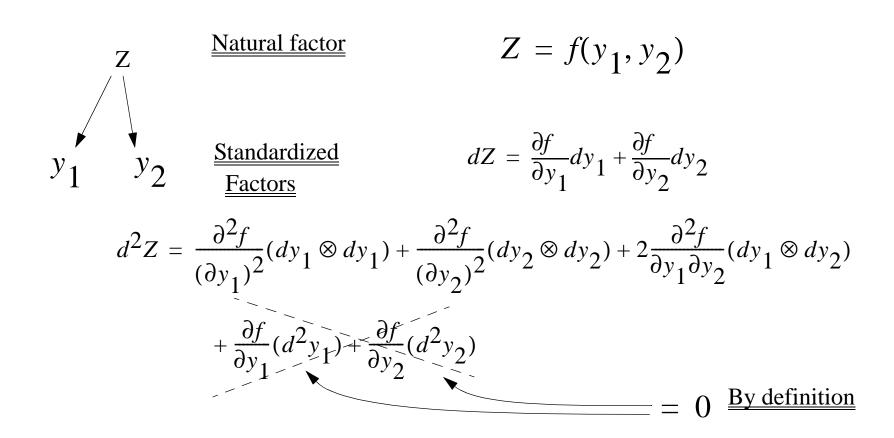
## **Example of a Natural Second Order Sensitivity**

• Top Level Senstivities

```
FXOption.nb *
      inst = FxOption[buy, call, 100 mm, usd, TT["30-Mar-98"], currencybase \rightarrow dem, strike \rightarrow 1.5]
 ln[22]:= NS2a[inst]
Out[22]= 3.46703×107 S1[Index[ExchangeRate[Currency[dem], Currency[usd]]], 0]
         S1[Index[ExchangeRate[0.0657534, Currency[usd], Currency[dem]]], 0] +
        3.94459 \times 10^8 \text{ SI}[Index[ExchangeRate[0.0657534, Currency[usd], Currency[dem]]], 0]^2 -
        103181. S1 [Index [ExchangeRate [Currency [dem], Currency [usd]]], 0]
         S1[Index[Rate[0.0657534, Currency[dem]]], 0] -
        3.16338×10<sup>6</sup> S1[Index[ExchangeRate[0.0657534, Currency[usd], Currency[dem]]], 0]
         S1[Index[Rate[0.0657534, Currency[dem]]], 0] +
        4707.22 S1 [Index [Rate [0.0657534, Currency [dem]]], 0]2 +
        6.45868×107 S1[Index[ExchangeRate[Currency[dem], Currency[usd]]], 0] S1[Volatility1[
           Index[ExchangeRate[0.0657534, Currency[usd], Currency[dem]]], 0.0657534, 0.], 0] +
        7.64823×107
         S1[Index[ExchangeRate[0.0657534, Currency[usd], Currency[dem]]], 0] S1[Volatility]
           Index(ExchangeRate(0.0657534, Currency(usd), Currency(dem()), 0.0657534, 0.1, 0) -
        5.89302 x 106 S1 [Index [Rate [0.0657534, Currency [dem]]], 0] S1 [Volatility] [
           Index[ExchangeRate[0.0657534, Currency[usd], Currency[dem]]], 0.0657534, 0.], 0] +
        4.3603×108 S1[Volatility1[Index[
             ExchangeRate [0.0657534, Currency [usd], Currency [dem]]], 0.0657534, 0.], 0] ^2 +
        2.40549 x 107 S2[Index[ExchangeRate[0.0657534, Currency[usd], Currency[dem]]], 0] -
        49669.7 $2[Index[Rate[0.0657534, Currency[dem]]], 0] - 4.48115 × 107 $2[Volatility][
           Index[ExchangeRate[0.0657534, Currency[usd], Currency[dem]]], 0.0657534, 0.],
          0]
```

#### **Mapping (Introduction)**

- The last step in the chain of chain computations of natural senstivities
- Necessary step to do risk and generalized hedge ratio computation



## **Mapping Exemple from an Interpolation**

- Assume that the value of a portfolio PV depends on  $r_{18m}$
- Develop it at the second order  $PV \approx PV_0 + \delta(r_{18m} r_{18m, 0}) + \frac{1}{2}\gamma(r_{18m} r_{18m, 0})^2$
- Assume we have an interpolation function  $r_{18m} = \frac{r_1 + r_2}{2}$
- We replace to get:  $PV \approx PV_0 + \delta\left(\frac{(r_1 r_{1,0}) + (r_2 r_{2,0})}{2}\right) + \frac{\gamma}{2}\left(\frac{(r_1 r_{1,0}) + (r_2 r_{2,0})}{2}\right)^2$
- We say this mapping **derives** from the interpolation

## Do all Mappings Derive from an Interpolation?

- The present value  $PV = C_{18m}e^{-r_{18m}t} = C_{18m}B_{18m} = (dPV = C_{18m}dB_{18m})$
- A Mapping should give birth to another decomposition  $dPV = C_1 dB_1 + C_2 dB_2$
- and therefore we have  $:dB_{18m} = (C_1/C_{18m})dB_1 + (C_2/C_{18m})dB_2$
- From  $dB_t = -tB_t dr_t$  we get  $dr_{18m} = \frac{C_1 B_1}{C_{18m} (3/2) B_{18m}} dr_1 + \frac{C_2 2 B_2}{C_{18m} (3/2) B_{18m}} dr_2$
- Does this derive from an interpolation function  $r_{18m} = f(r_1, r_2)$ ?
- Poincare Lemma tells us :  $<=>\frac{\partial}{\partial r_2} \left[ \frac{C_1 B_1}{C_{18m} (3/2) B_{18m}} \right] = \frac{\partial}{\partial r_1} \left[ \frac{C_2 2 B_2}{C_{18m} (3/2) B_{18m}} \right]$
- If it derives, ==> <u>Implicit Interpolation Function</u>. it does it for RiskMetrics.

## Why should a Mapping Derives from an Interpolation?

- <u>All Sensitivities Mapping</u>: To have a consistent method to derive the mapping of sensitivities of any order.
- <u>Arbitrage</u>: having an interpolation inconsistent with a first order mapping lead to the following risk arbitrage: Any position with a convexity depending on the risk factor whose mapping is wrong will be mispriced.
- <u>Inconsistent Risk Calculation</u>: by having a risk computed for a position at a different market point from its actual Mark to Market, we do an inexact computation. The Risk mesures how the PV varies when the risk factors move. We therefore must have the correct linear relationships for it

## **Risk Metrics Mapping**

• It determine the mapping by enforcing the followings:

$$Var[C_1 + C_2] = \sigma_1^2 Z_1^2 C_1^2 + 2\rho \sigma_1 \sigma_2 Z_1 C_1 Z_2 C_2 + \sigma_2^2 Z_2^2 C_2^2$$

$$Var[C_t] = \sigma_t^2 Z_t^2 C_t^2$$

- And the preservation of the present value which is by no way a needed characteristic of an equivalent risk representation
- An ideal variance property would have been:

$$Var[P+C_1+C_2] = \sigma_1^2 Z_1^2 C_1^2 + 2\rho \sigma_1 \sigma_2 Z_1 C_1 Z_2 C_2 + \sigma_2^2 Z_2^2 C_2^2 + Var[P] + 2Cov[P, C_1+C_2]$$

$$Var[P+C_t] = \sigma_t^2 Z_t^2 C_t^2 + Var[P] + 2Cov[P, C_t]$$
we can try to gess that But we have no ways to determine that !

• The implicit volatility determined by a choosen mapping is by no way arbitrageable which is not the case of the implicit interpolation. Neither It gives a non consistency argument for the risk computation, which is not the case with the implicit interpolation.

#### **Risk Measures for Non Linear Instruments**

• Being able to get the exact sensitivities, we approximate the variation of the MtM by the

quadratic polynomial: 
$$PV \approx PV_0 + \Delta \cdot \begin{bmatrix} dx_1 \\ ... \\ dx_n \end{bmatrix} + \frac{1}{2} \begin{bmatrix} dx_1 & ... & dx_n \end{bmatrix} \Gamma \begin{bmatrix} dx_1 \\ ... \\ dx_n \end{bmatrix}$$

- We assume a jointly normal approximation of the risk factors  $Cov[dx_i, dx_j] = \sum_{i,j} dt$
- The Risk Metrics Risk measure is just  $|Var[PV_t PV_0]| = ({}^t \Delta \Sigma \Delta)t = Trace[t \Sigma(\Delta \otimes \Delta)]$
- If we take into account the gammas:  $Var[PV_t PV_0] = Trace \left[t\left(\Sigma(\Delta \otimes \Delta) + \frac{1}{2}(\Gamma\Sigma)^2\right)\right]$
- skewness and kurtosis:

$$Skew[PV_t - PV_0] = E[(PV_t - PV_0)^3] = Trace[3\Sigma\Gamma\Sigma(\Delta\otimes\Delta) + (\Gamma\Sigma)^3]$$
 
$$Kurtosis[PV_t - PV_0] = E[(PV_t - PV_0)^4 - 3(Var[PV_t - PV_0])^2] = Trace[12\Sigma(\Gamma\Sigma)^2(\Delta\otimes\Delta) + (3(\Gamma\Sigma))^4]$$

#### **More about Cumulants**

• The carateristic function  $\Phi_X(t) = E[e^{itX}]$  is a very interesting object

- It is the moment generating function :  $\mu_k[X] = E[(X)^k] = (-i)^k \frac{d^k}{dt} (\Phi_X(t)|_{t=0})$
- It simply multiply over independent variable :

$$\det \quad \Psi_X(t) = Log[\Phi_X(t)]$$
 X and Y independent  $\Leftrightarrow \Phi_{X+Y}(t) = \Phi_X(t)\Phi_Y(t) \Leftrightarrow \Psi_{X+Y}(t) = \Psi_X(t) + \Psi_Y(t)$ 

- we define the cumulants as  $Cum_k(X) = (-i)^k \frac{d^k}{dt} (\Psi_X(t)|_{t=0})$
- they cumul over independent variables :  $Cum_k(X + Y) = Cum_k(X) + Cum_k(Y)$

#### **More about Cumulants (2)**

• The first cumulants are:

$$Cum_1(X) = E[X]$$

$$Cum_2(X) = E[X^2] - (E[X])^2 = Var[X]$$

$$Cum_3(X) = E[(X - E[X])^3] = Skewness[X]$$

$$Cum_4(X) = E[(X - E[X])^4 - 3(E[(X - E[X])^2])^2] = kurtosis[X]$$

- They measure how far from the normal distribution X is :  $Cum_k[X] = 0 \quad \forall k > 2$  is equivalent to X Normal
- They caracterise a distribution :

Knowing the cumulants <=> Knowing the distribution

#### **More about Cumulants (3)**

• You can modify the set of cumulants by the superoperator

$$g(x) = \exp\left[\sum_{i=1}^{\infty} \frac{\varepsilon_i (-D)^i}{i!}\right] h(x) \text{ where h(x) and g(x) are densities}$$

Then 
$$Cumulant_k[g] = Cumulant_k[h] + \varepsilon_k$$

- We can find the distribution associated with a set of cumulants  $c_2$ ,  $c_3$ ,  $c_4$
- We can find option values in case where the market is non normal:

$$Option = B_T \int (x - K)^+ g(x)$$

Then by plugging the formula and integrating by parts we correct the option price

## **Computing Long Maturity Risks**

• Using Lognormal Variables for the Risk Factors :

$$\frac{dX_t}{X_t} = a(t)dt + \sigma dW_t \qquad Cov\left[\frac{dX_t}{X_t}, \frac{dY_t}{Y_t}\right] = \sigma_{X, Y}dt$$

=> we just handle Log[X] instead of X

• Using Mean Reverting Variables for the Risk Factors :

$$dX_t = (a(t) - X_t)dt + \sigma dW_t \qquad \frac{dX_t}{X_t} = (a(t) - X_t)dt + \sigma dW_t \qquad \frac{dX_{i,j}(t)}{X_{i,j}(t)} = B_i(a_{i,j}(t) - X_i)dt + \overrightarrow{\sigma_{i,j}}dW_t$$

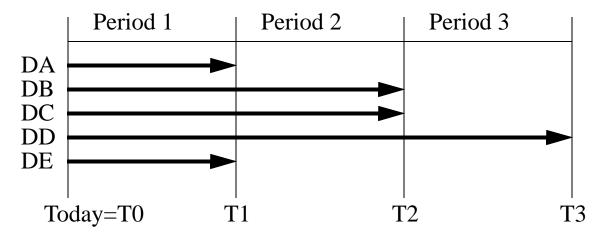
=> we just replace the dependency in the horizon of risks

$$\sigma_t = \sqrt{t} \| \overrightarrow{\sigma_i} \|$$
 by  $\sigma_t = \sqrt{\frac{1 - e^{-2B_i t}}{2B_i}} \| \overrightarrow{\sigma_i} \|$  for independent brownians

-> principal component analysis to identify the blocs of factors i with the same  $B_i$ 

### **Handling of Liquidity Risks**

- Thanks to cumulants we define a distribution of the P&L associated with a complete unwinding of the portfolio, knowing the unwinding period for each deals:
- No interest rates (the futur values have to be corrected by the zero coupon price)



- $PL(3)-PL(0)={PL(3)-PL(2)} + {PL(2)-PL(1)} + {PL(1)-PL(0)}$
- Markovian Markets:  $Var[PL_3 PL_0] = Var[PL_3 PL_2] + Var[PL_2 PL_1] + Var[PL_1 PL_0]$
- More generally:  $Cu_n[PL_3 PL_0] = Cu_n[PL_3 PL_2] + Cu_n[PL_2 PL_1] + Cu_n[PL_1 PL_0]$

## **Computation of Forward cumulants**

• Portfolio:

Computation of the Variance for Period 1 + Period 2

$$Tr\Big[(\Delta \otimes \Delta)(C_1 + C_2) + \frac{1}{2}(\Gamma(C_1 + C_2))^2\Big] \neq Tr\Big[(\Delta \otimes \Delta)C_1 + \frac{1}{2}(\Gamma C_1)^2\Big] + Tr\Big[(\Delta \otimes \Delta)C_2 + \frac{1}{2}(\Gamma C_2)^2\Big]$$
Variance associated with
$$\Delta(t_1) \text{ unknown}$$

$$\Delta(t_1) \text{ known}$$

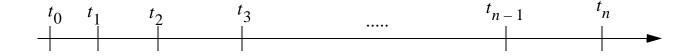
• But  $\Delta(t_1)$  is not known due to  $\Gamma$  effect, so:

$$Var(Period_2) = Tr \left[ (\Delta \otimes \Delta)(C_1 + C_2) + \frac{1}{2}(\Gamma(C_1 + C_2))^2 \right] - Tr \left[ (\Delta \otimes \Delta)C_1 + \frac{1}{2}(\Gamma C_1)^2 \right]$$

• And the same for other cumulants :  $Cu_n(Period_q) = Cu_n \left( \sum_{i \le q} Period_i \right) - Cu_n \left( \sum_{i \le q-1} Period_i \right)$ 

## **Equivalent Unwinding Periods**

• Lets assume that the deals are sorted and regrouped by unwinding periods. Let be the number of classes be n



- Let be the pth forward cumulant associated with the period k:  $[t_{k-1}, t_k]$  be  $Cu_p, k$ . There are normalized to 1 unit of time
- The total p cumulant of the total P&L is therefore  $Cu_p = \sum_{k=1, n} Cu_{p, k} (t_k t_{k-1})$
- The Equivalent unwinding period associated with the order p is

$$T_{p} = \frac{\sum_{k=1, n} Cu_{p, k}(t_{k} - t_{k-1})}{Cu_{p, 1}}$$

## **Modularity of Unwinding Periods**

• Let assume that every subportfolio k has an unwinding period of order  $p:T_{p,\,k}$  with a normalized p cumulant associated with the period 1 (instantaneous cumulant) of the subportfolio k be

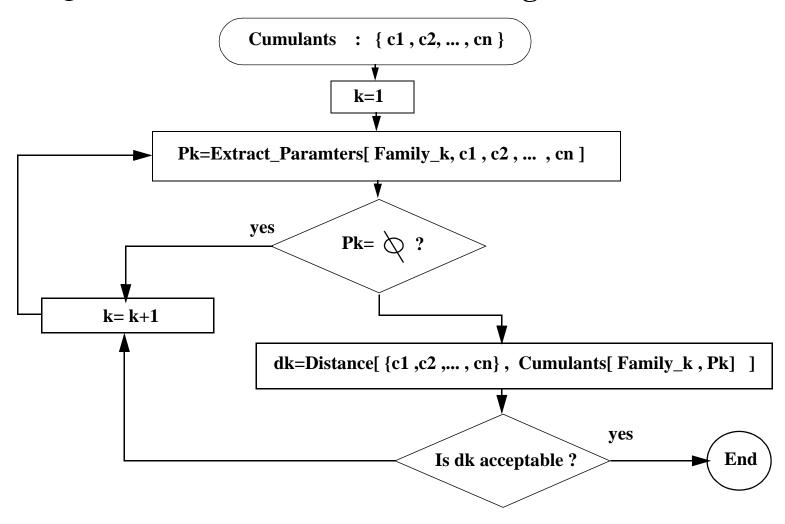
$$Cu_{p, 1, k}$$

• The unwinding period associated with the total portfolio of order p is therefore :

$$T_{p} = \frac{\sum_{k=1, n}^{Cu_{p, 1, k}} x^{T_{p, k}}}{Cu_{p, 1}}$$

• Because of central limit theorem, we observe that this compounding tend to make  $T_p \to 0$  for p > 2 Which is an example of ergodicity property

## **Quadratic Risk Calculation: Fitting the Cumulants**



### **Quadratic Risks: The Different Families of Distributions**

- Extract\_ Parameters [ Cumulants ]
- Compute\_Cumulants [ Parameters ]

Distance [ cumulants-1 , cumulants-2 ]

$$d[a, b] = \sum_{k=2}^{N} \alpha_k (a_k - b_k)^2$$
 scalar product : easy optimization

• Example of families :

$$Z = \frac{1}{2}gX^2 + dX + hY^2 - \frac{1}{2}g - \frac{1}{2}h$$

$$X \text{ and } Y \text{ are independant normal variables}$$

$$Z = \frac{1}{2}gX^2 + dX + cY - \frac{1}{2}g$$

$$Z = \gamma + \delta \log \left[\frac{Y}{1 - Y}\right] \quad Z = \gamma + \delta Sinh^{-1}[Y] \quad Z = \gamma + \delta \log[Y] \quad \text{Johnson Family of Distribution}$$

$$\frac{1}{p}\frac{dp}{dx} = -\frac{a + x}{c_0 + c_1 x + c_2 x^2}$$
Pearson Family of Distribution

## **Example of a Fitting**

• let s study the following family of distribution :  $U = \frac{1}{2}aX^2 + bX + cY$  with X and Y independant Normalized Normal variables

• but the 4th degre equation has a solution only if  $c_4 < 3(c_3)^{4/3}$ 

## **Exemple of a Fitting: Suite**

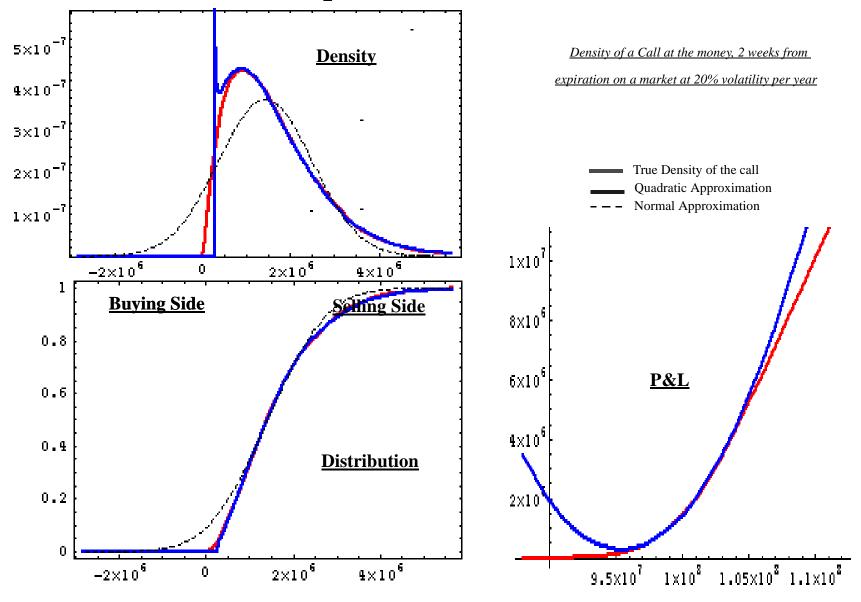
• Computation of the Distribution :

$$\alpha = \begin{pmatrix} \int_{-\frac{b^2}{2a}}^{\infty} \frac{\phi'(\frac{R-y}{c})}{c} \{\phi(x_2(y)) - \phi(x_1(y))\} dy & (g>0) \\ -\frac{b^2}{2a} \frac{\phi'(\frac{R-y}{c})}{c} \{\phi(x_2(y)) - \phi(x_1(y))\} dy & (g<0) \end{pmatrix}$$

• with the convention:

$$\frac{x_2(y)}{x_1(y)} = \frac{-b}{a} \pm \frac{\sqrt{b^2 + 2ay}}{a}$$
  $\phi'(x) = \frac{e^{-\frac{x^2}{2}}}{\sqrt{2\pi}}$ 

## **Exemples of Quadratic Risks**



#### **Exact Results**

#### • For the 2 weeks option at the money

Expected value: 1.41/524	Risks	level	Real Risk	Quadratic	relat. err.	Linear	relat. err.	Improv./Lin.
	Buyer	5%	0.988	1.056	-3.29%	1.747	60.1%	94.7%
(in \$ millions)		1%	1.2721	1.1415	-11.4%	2.472	94.3%	89.1%
	Seller	5% 1%	2.223 3.401	2.274 3.472	2.25% 4.79%	1.747 2.472	-21.4% -27.3	89.2% 81.6%

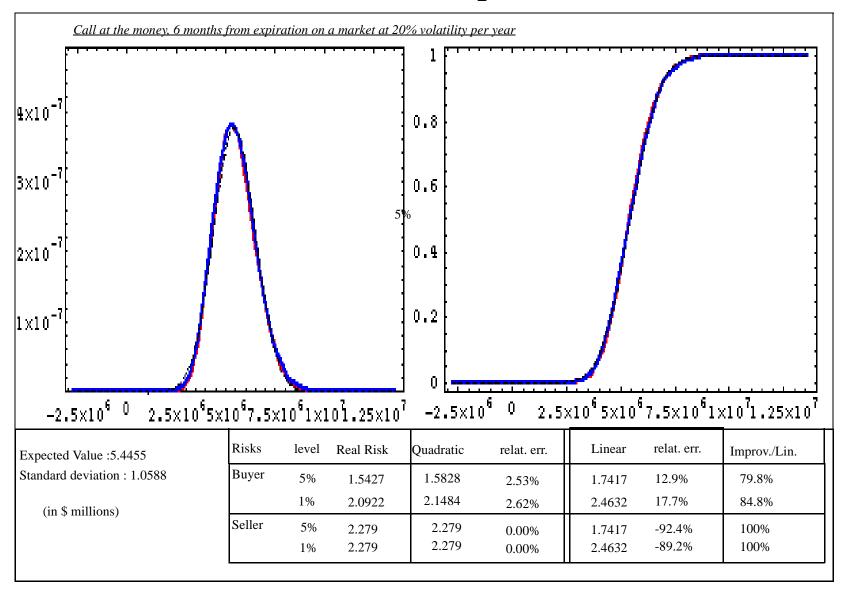
#### • A "Normal Option" (6 month)

Expected value :5.4455	Risks	level	Real Risk	Quadratic	relat. err.	Linear	relat. err.	Improv./Lin.
	Buyer	5%	1.5427	1.5828	2.53%	1.7417	12.9%	79.8%
(in \$ millions)		1%	2.0922	2.1484	2.62%	2.4632	17.7%	84.8%
	Seller	5% 1%	2.279 2.279	2.279 2.279	0.00% 0.00%	1.7417 2.4632	-92.4% -89.2%	100% 100%

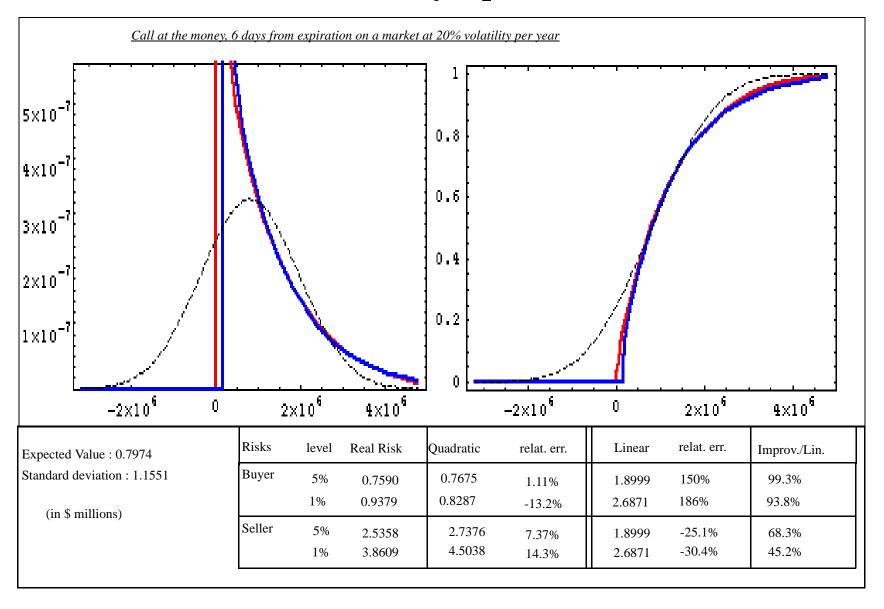
#### • A Nasty Option (6 days)

Expected Value : 0.7974 Standard deviation : 1.1551 (in \$ millions)	Risks	level	Real Risk	Quadratic	relat. err.	Linear	relat. err.	Improv./Lin.
	Buyer	5% 1%	0.7590 0.9379	0.7675 0.8287	1.11% -13.2%	1.8999 2.6871	150% 186%	99.3% 93.8%
	Seller	5% 1%	2.5358 3.8609	2.7376 4.5038	7.37% 14.3%	1.8999 2.6871	-25.1% -30.4%	68.3% 45.2%

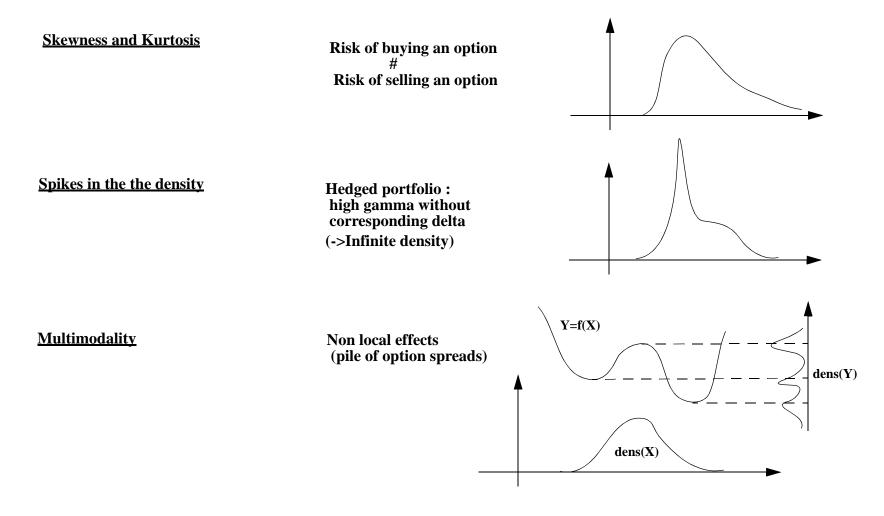
## A "Normal Option"



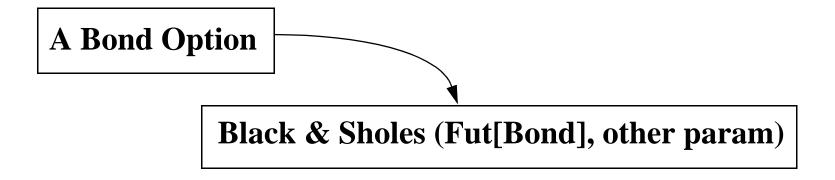
## **A Nasty Option**



#### **Difficulties associated with Risk Calculations**



## **Simple Hedge Ratio Computation**



$$\begin{aligned} \text{Hedge Ratio (type I)} &= \frac{\partial Black \ (Fut(Bond) \ , \ other \ param)}{\partial \ Fut \ (Bond)} \end{aligned}$$

## **Hedge Ratios using the First Order Differentials**

$$\Delta_{Option} = \begin{bmatrix} a_1 \\ a_2 \\ \dots \\ a_n \end{bmatrix} \quad \Delta_{Futur} = \begin{bmatrix} b_1 \\ b_2 \\ \dots \\ b_n \end{bmatrix} \qquad \begin{array}{c} \textbf{Geometry determined by } \Sigma \\ X \cdot Y = {}^t X \Sigma Y \\ \|X\|^2 = {}^t X \Sigma X \end{array}$$

Geometry determined by 
$$\Sigma$$

$$X \cdot Y = {}^{t}X\Sigma Y$$

$$\|X\|^{2} = {}^{t}X\Sigma X$$

Hedge ratio (type II) = 
$$\frac{\Delta_{Option} \cdot \Delta_{Futur}}{\|\Delta_{Futur}\|^2}$$

If n=1, 
$$\Delta_{Option}$$
 //  $\Delta_{Futur}$  => Hedge ratio (type II)=Hedge ratio (type I)

## **Hedge Ratio as a Distance Minimizer**

$$D_{Option} = \begin{bmatrix} \frac{\partial}{\partial X_i} Option \\ \frac{\partial^n}{\partial X_{i_1} \partial X_{i_2} ... \partial X_{i_n}} Option \\ ... \\ Option(X_k + a_k) \\ Option(h_k(X)) \\ \frac{\partial^n}{\partial X_{i_1} \partial X_{i_2} ... \partial X_{i_n}} Option(h_k(X)) \end{bmatrix}$$

**Risk Descriptor:** 

**One constraint: being linear** 

**Risk Function:** 
$$D_{Option} \longrightarrow ||D_{Option}|| \in R^+$$

**Risk Function :**  $D_{Option} \longrightarrow \|D_{Option}\| \in R^+$  **Hedge Ratio (type III)** =  $\{h_i\}$  **such Minimum**  $D_{Option} - \sum_{j=1}^{p} h_j D_{hedge(j)}$ 

If p=1, Risk descriptor =  $\Delta$ , Risk function = Var[.] => Hedge ratio (type III)=Hedge ratio (type II)

## Conclusion

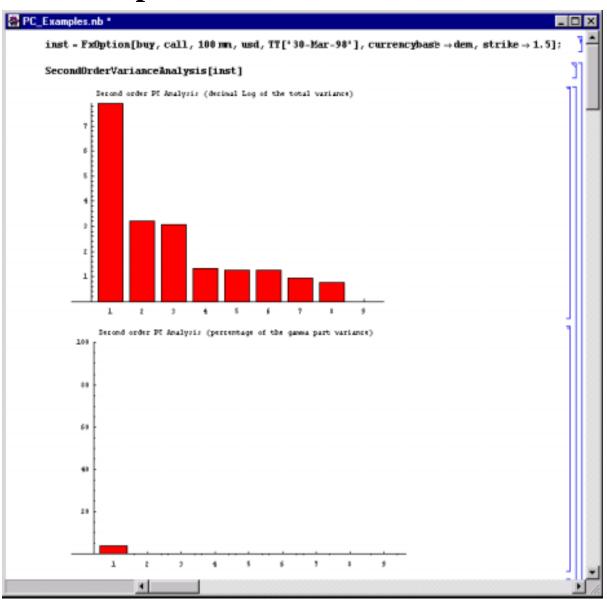
## Second Order principal component analysis

- We have the universal formula :  $Var[PV_t PV_0] = Trace \left[t \left(C(\Delta \otimes \Delta) + \frac{1}{2}(\Gamma C)^2\right)\right]$
- If we set in the basis where C=1 and gamma is diagonal, this formula can decompose into:

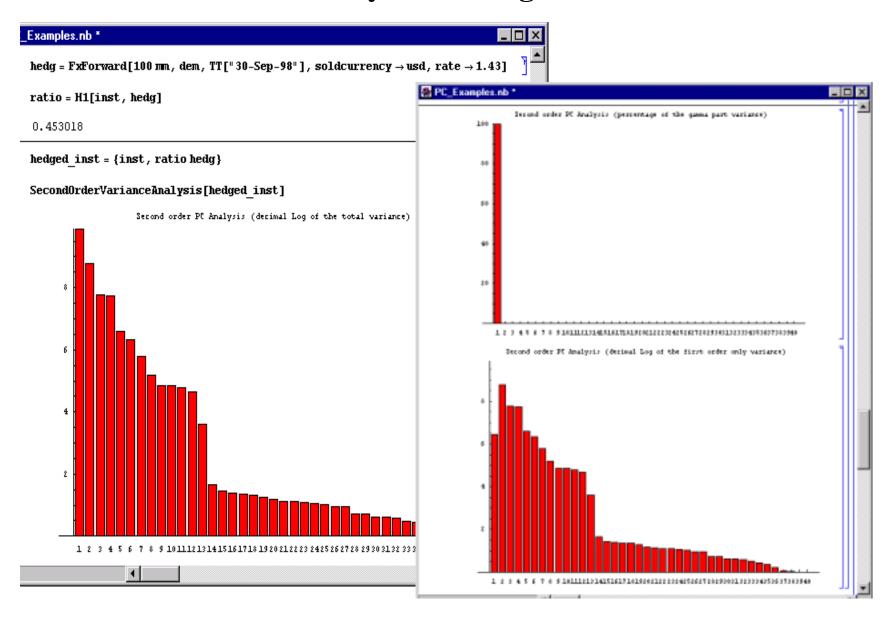
$$Var[PV_t - PV_0] = t \sum_{i=1}^{n} \left( \delta_i^2 + \frac{1}{2} \gamma_i^2 \right)$$

- Therefore we can analyse the origin of the variance, Sort the dimensions by importances
- Other Cumulants are also diagonal, allowing us to refine our discretisation depending on the severity of the nonlinearity :
- There is no third order principal component analysis because we cannot diagonalize the third order tensors  $\frac{\partial^3 PV}{\partial x_i \partial x_j \partial x_k}$

## **Exemple of Second order PCA**



## **PC Analysis : Strong Gamma**



## Reminder of Linear Agebra

• Matrix Multiplication

$$\begin{bmatrix} 2 & 3 \\ 1 & 5 \end{bmatrix} \begin{bmatrix} x \\ y \end{bmatrix} = \begin{bmatrix} 2x + 3y \\ x + 5y \end{bmatrix}$$

Norm of a Vector

$$x = \begin{bmatrix} 1 \\ 2 \\ 3 \end{bmatrix} \qquad ||x|| = \sqrt{1^2 + 2^2 + 3^2} = \sqrt{14}$$

• Quadratic Form

$$C = \begin{bmatrix} 1 & 2 & -1 \\ 2 & 3 & 2 \\ -1 & 2 & 5 \end{bmatrix} \quad x = \begin{bmatrix} x \\ y \\ z \end{bmatrix} \qquad {}^{t}xCx = \begin{bmatrix} x & y & z \end{bmatrix} \begin{bmatrix} 1 & 2 & -1 \\ 2 & 3 & 2 \\ -1 & 2 & 5 \end{bmatrix} \begin{bmatrix} x \\ y \\ z \end{bmatrix} = x^{2} + 3y^{2} + 5z^{2} + 4xy - 2xz + 4yz$$

• Exemple of Quadratic Form

Covariance matrix, Correlation matrix, Gamma matrix

positive when  $t Cx \ge 0$ Diagonal when  $C = \begin{bmatrix} 3 & 0 & 0 \\ 0 & 4 & 0 \\ 0 & 0 & 1 \end{bmatrix}$ 

## Reminder of linear Algebra (2)

• Scalar product

$$x = \begin{bmatrix} 1 \\ 2 \end{bmatrix}$$
  $y = \begin{bmatrix} 3 \\ 5 \end{bmatrix}$   $x \cdot y = 1 \times 3 + 2 \times 5$  (projection)

• Tensorial product

$$x = \begin{bmatrix} 1 \\ 2 \end{bmatrix} \qquad y = \begin{bmatrix} 3 \\ 5 \end{bmatrix} \qquad x \otimes y = \begin{bmatrix} 1 \\ 2 \end{bmatrix} \begin{bmatrix} 3 & 5 \end{bmatrix} = \begin{bmatrix} 1 \times 3 & 1 \times 5 \\ 2 \times 3 & 2 \times 5 \end{bmatrix} = \begin{bmatrix} 3 & 5 \\ 6 & 10 \end{bmatrix}$$

• Trace

$$Trace \begin{bmatrix} 3 & 5 \\ 6 & 10 \end{bmatrix} = 3 + 10 = 13$$
 (non commutative integration)

• Link beween the precedings:

$$Trace[x \otimes y] = x \cdot y$$
  $Trace[(x \otimes y)C] = Trace[C(x \otimes y)] = {}^txCy$ 

## **Reminder of Differential Analysis**

• Derivative

$$f(x)$$
  $\frac{df}{dx} = Lim\left(\frac{f(x+dx)-f(x)}{dx}\right)$  when  $dx \to 0$ 

• Partial Differential

$$f(x, y)$$
  $\frac{\partial f}{\partial x} = Lim\left(\frac{f(x + dx, y) - f(x, y)}{dx}\right)$  when  $dx \to 0$ 

• Total Diffential

$$df = \frac{\partial f}{\partial x}dx + \frac{\partial f}{\partial y}dy$$

## **Differential Analysis (2)**

• Chain rule for the derivative

$$f(x) = g(h(x)) = g \bullet h(x)$$

$$h \quad g$$

$$x ----> f(x)$$

$$\frac{df}{dx} = \frac{dg}{dy} \times \frac{dh}{dx}$$

$$df = \frac{df}{dy} \times dy$$
Notation Abuse

Chain Rule for the differential

$$f(x_{1}, x_{2}) = g(h_{1}(x_{1}, x_{2}), h_{2}(x_{1}, x_{2}))$$

$$x_{1} = h_{1}$$

$$x_{2} = h_{2}$$

$$x_{2} = h_{2}$$

$$x_{3} = h_{2}$$

$$x_{4} = h_{2}$$

$$x_{5} = h_{2}$$

$$x_{6} = h_{1}$$

$$x_{6} = h_{1}$$

$$x_{7} = h_{2}$$

$$x_{8} = h_{1}$$

$$x_{1} = h_{1}$$

$$x_{1} = h_{1}$$

$$x_{2} = h_{1}$$

$$x_{1} = h_{1}$$

$$x_{2} = h_{1}$$

$$x_{1} = h_{1}$$

$$x_{2} = h_{2}$$

$$x_{3} = h_{2}$$

$$x_{4} = h_{2}$$

$$x_{2} = h_{2}$$

$$x_{4} = h_{2$$

## Simultaneous Diagonalization of C and $\Gamma$

• 
$$P = \{V_1, V_2, ..., V_n\}$$
 is an eigen basis for C then 
$$\begin{pmatrix} \Delta' = P^{-1}\Delta \\ C' = P^*CP \text{ because we can check :} \\ \Gamma' = P^*\Gamma P \end{pmatrix}$$

 $-\Delta^{*}C'\Delta' = \Delta^{*}C\Delta$  and  $\Delta^{*}\Gamma'\Delta' = \Delta^{*}\Gamma\Delta$  form all vector  $\Delta$ 

• Now  $C' = \begin{bmatrix} \lambda_1 & \dots & 0 \\ \dots & \dots & \dots \\ 0 & \dots & \lambda_n \end{bmatrix} = \Lambda^2$  because C is definite positive, therfore we can define a new

basis in wich C'' = 1, and we will have :  $\Delta'' = \Lambda \Delta'$ ,  $\Gamma'' = \Lambda^{-1} \Gamma' \Lambda^{-1}$  in order to have :

$$-\Delta$$
"\* $\Delta$ " =  $\Delta$ '\* $C$ ' $\Delta$ ' and  $\Delta$ "\* $\Gamma$ " $\Delta$ " =  $\Delta$ '\* $\Gamma$ ' $\Delta$ ' again

- We observe than  $\Gamma$ " is still symetric because :  $\Gamma$ "\* =  $\Lambda^{-1}\Gamma$ "\*  $\Lambda^1 = \Lambda^{-1}(P^*\Gamma P)^*\Lambda^1 = \Gamma$ "
- Therefore  $\Gamma$ " is diagonalizable and its eigen basis is orthonormal