

# **Copulas**

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# Generic Copula

- Copula=Distribution sur  $[0, 1]^n$  such that the marginals are uniform
- definition of distributions:

$$F(X_1, X_2, \dots, X_n) = \text{Prob}[x_1 \leq X_1, x_2 \leq X_2, \dots, x_n \leq X_n]$$

- “Sklar Theorem” : modification of the marginals:

$$F(X_1, X_2, \dots, X_n) = C(G_1(X_1), G_2(X_2), \dots, G_n(X_n))$$

- $X_i$  are variables in the “marginal space”
- $G_i(X_i) \in [0, 1]$  is in the “probability space” ,
- $G_i$  is a cumulative distribution: it is a bijective mapping from the marginal space to the probability space

# Usual Copulas

- density:  $pdf(X_1, X_2, \dots, X_n) = \left( \frac{\partial}{\partial X_1} \cdot \frac{\partial}{\partial X_2} \cdot \dots \cdot \frac{\partial}{\partial X_n} \right) F(X_1, X_2, \dots, X_n)$
- Gaussian density:  $pdf(X_1, X_2, \dots, X_n) = \frac{1}{(2\pi)^{n/2} \sqrt{\det(\Sigma)}} e^{-\frac{1}{2} \left( (X - \mu)^T \cdot \frac{1}{\Sigma} \cdot (X - \mu) \right)}$
- Student density  $pdf(X_1, X_2, \dots, X_n) = \frac{\Gamma\left(\frac{\nu + n}{2}\right)}{\Gamma\left(\frac{\nu}{2}\right) (\pi\nu)^{n/2} \sqrt{\det(\Sigma)}} \left( 1 + \frac{(X - \mu)^T \cdot \frac{1}{\Sigma} \cdot (X - \mu)}{\nu} \right)^{-\left(\frac{\nu + n}{2}\right)}$
- Elliptical Copula: the characteristic function  $\varphi_{X - \mu}(t) = \phi(t^T \Sigma t)$  (like gaussian, student,)
- Archimedian copula: generated by a decreasing function  $\varphi(x)$  from  $[0, 1]$  to  $[0, \infty]$ 
  - $C(u, v) = \varphi^{-1}(\varphi(u) + \varphi(v))$
  - exemple the Gumbel :  $\varphi(u) = (-\log(u))^\theta$  or the Clayton  $\varphi(u) = (u^{-\theta} - 1)/\theta$

# Tail Dependence

- Upper tail dependence

$$\lim_{u \rightarrow 1} \text{Prob}[Y > F_Y^{-1}(u) | (X > F_X^{-1}(u))] = \lambda_{up}$$

- with a copula  $\lambda_{up} = \lim_{u \rightarrow 1} \frac{(1 - 2u + C(u, u))}{1 - u}$

- exemple Gumbel :  $\lambda_{up} = 2 - 2^{\frac{1}{\theta}}$

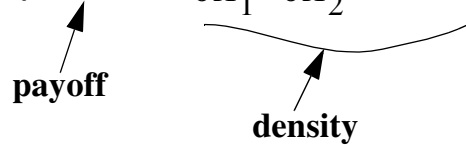
- Gaussian copula :  $\lambda_{up} = 0$

- Lower tail dependence with a copula  $\lambda_{down} = \lim_{u \rightarrow 1} \frac{C(u, u)}{u}$

- Student bivariate :  $\lambda_{up} = \lambda_{down} = \frac{\sqrt{v+1}}{t_{v+1}} (\sqrt{v+1} \sqrt{1-R_{12}} / \sqrt{1+R_{12}})$

# Pricing with Copula (bivariate)

- $Option = \int h(X_1, X_2) \frac{\partial}{\partial X_1} \cdot \frac{\partial}{\partial X_2} F(X_1, X_2) dX_1 dX_2$


  
 payoff
   
 density

- we separate the copula from the marginals

$$\begin{aligned}
 \gamma_i(X_i) &= p_i && \text{pricing marginals} \\
 \eta(Y) &= p && \text{copula marginals} \\
 Option &= \int h(X_1, X_2) \frac{\partial}{\partial X_1} \cdot \frac{\partial}{\partial X_2} C(\gamma_1(X_1), \gamma_2(X_2)) dX_1 dX_2 && \text{pricing marginal space} \\
 &= \int h(X_1, X_2) \partial_1 \partial_2 C(\gamma_1(X_1), \gamma_2(X_2)) \gamma_1'(X_1) \gamma_2'(X_2) dX_1 dX_2 \\
 &= \int h(\gamma_1^{-1}(p_1), \gamma_2^{-1}(p_2)) \partial_1 \partial_2 C(p_1, p_2) dp_1 dp_2 && \text{probability space} \\
 &= \int h(\gamma_1^{-1} \cdot \eta(p_1), \gamma_2^{-1} \cdot \eta(p_2)) \partial_1 \partial_2 C(\eta(Y_1), \eta(Y_2)) \eta'(Y_1) \eta'(Y_2) dY_1 dY_2 \\
 &= \int h(\gamma_1^{-1} \cdot \eta(Y_1), \gamma_2^{-1} \cdot \eta(Y_2)) \frac{\partial}{\partial X_1} \cdot \frac{\partial}{\partial X_2} C(\eta(Y_1), \eta(Y_2)) dY_1 dY_2 && \text{copula marginal space} \\
 &&& \text{pricing quantile} \quad \text{copula marginals} \quad \text{total density associated with the copula}
 \end{aligned}$$

## Additional simplification for gaussian copulas:

$$\begin{aligned}
 &= \int h\left(\gamma_1^{-1} \bullet N(Y_1), \gamma_2^{-1} N(Y_2)\right) \frac{e^{-\frac{1}{2(1-\rho^2)}(Y_1^2 + Y_2^2 - 2\rho Y_1 Y_2)}}{2\pi\sqrt{1-\rho^2}} dY_1 dY_2 \\
 &= \int h\left(\gamma_1^{-1} \bullet N(Y_1), \gamma_2^{-1} N(Y_2)\right) \frac{e^{-\frac{1}{2}\left(\left(Y_1\sqrt{1-\rho^2} + \rho Y_2\right)^2 + Y_2^2\right)}}{2\pi\sqrt{1-\rho^2}} dY_1 dY_2 \\
 &= \int h\left(\gamma_1^{-1} \bullet N\left(\frac{Y_1' - \rho Y_2}{\sqrt{1-\rho^2}}\right), \gamma_2^{-1} N(Y_2)\right) \frac{e^{-\frac{1}{2}(Y_1'^2 + Y_2^2)}}{2\pi} dY_1' dY_2
 \end{aligned}$$

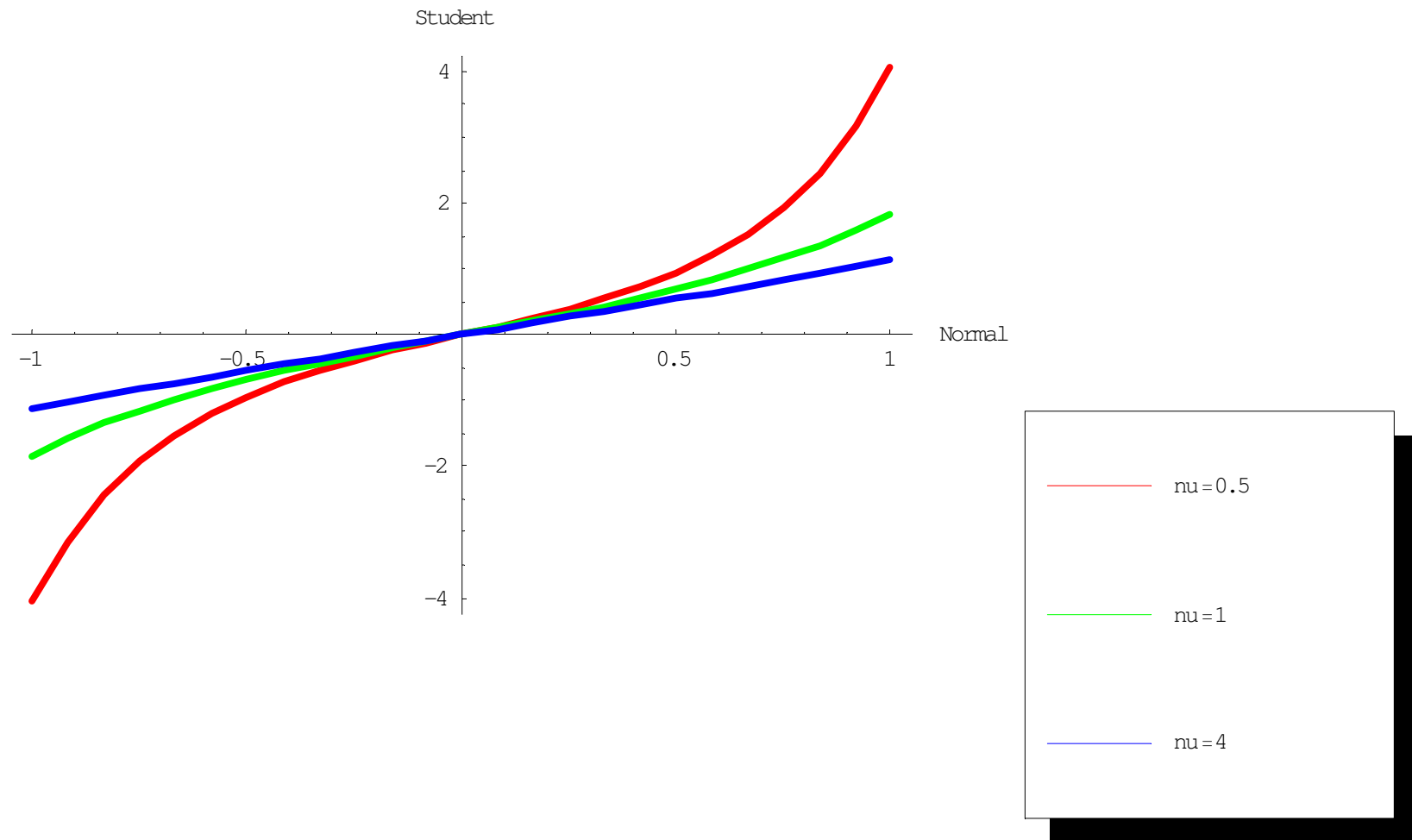
$$Y_1\sqrt{1-\rho^2} + \rho Y_2 = Y_1'$$

correlated marginal

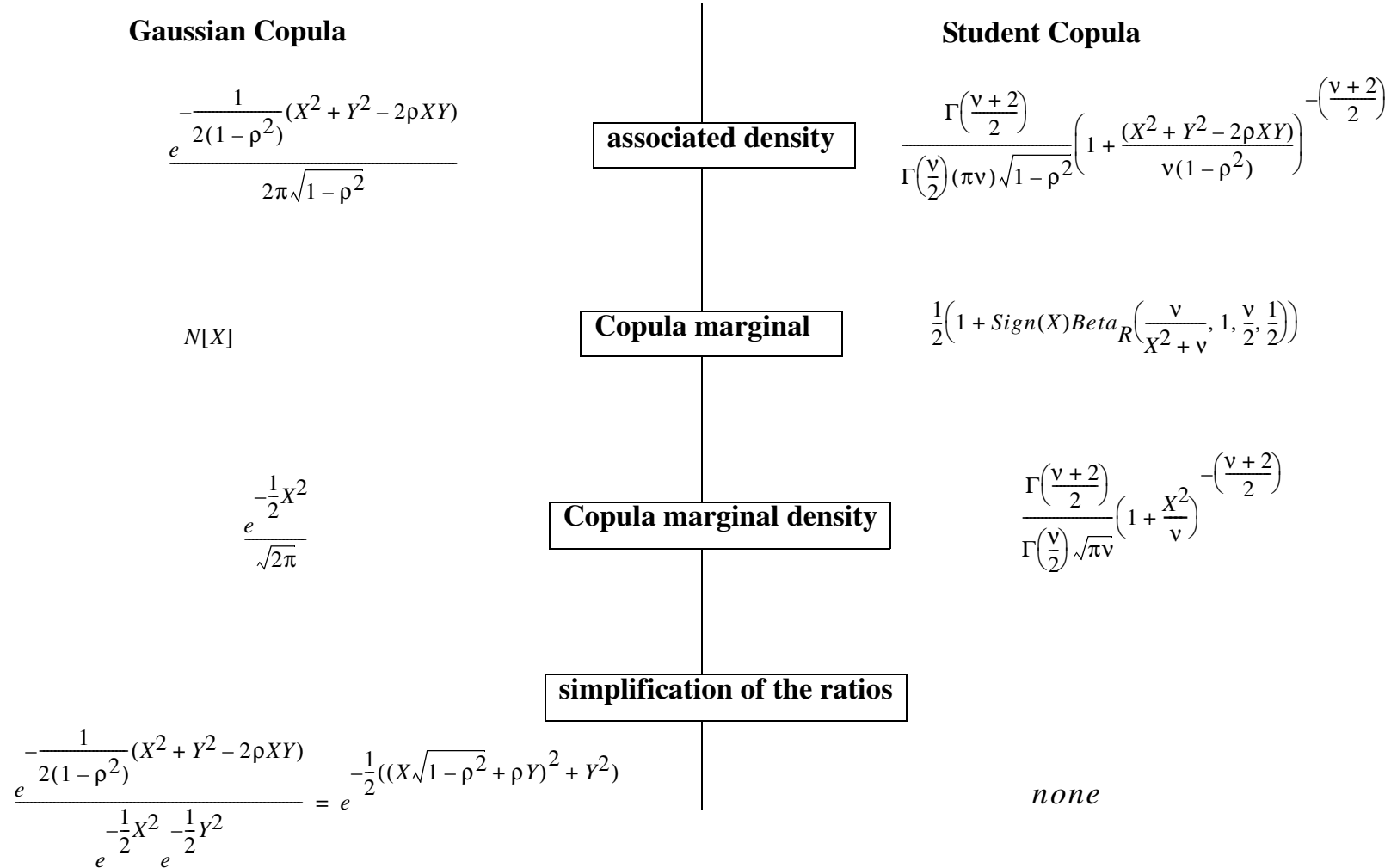
non correlated bivariate density

# Gaussianity of a marginal

- We plot  $\psi_v^{-1} \bullet N(x)$



# Exemple of elliptic copulas





## Other way to get copulas

- Genest and Rivet Transformation : if  $C(u, v)$  is a copule and  $\gamma$  is a one to one function of  $[0,1]$  to  $[0,1]$ , then

$$C_{\gamma}(u, v) = \gamma^{-1}(C(\gamma(u), \gamma(v)))$$

- The Franck copule defined by :

$$C(u, v) = -\frac{1}{\theta} \text{Log} \left( 1 - \frac{(1 - e^{-\theta u})(1 - e^{-\theta v})}{1 - e^{-\theta}} \right)$$

- can be compounded by  $\gamma(x) = x^{1/\alpha}$  with  $\alpha > 1$

# The Symmetric Markov Algebra of copulas

- Product of 2 copulas :

$$(x, y) \rightarrow (C_1 \cdot C_2)(x, y) = \int_0^1 \partial_2 C_1(x, s) \partial_1 C_2(s, x) ds$$

- adjoint:  $(x, y) \rightarrow (C^T)(x, y) = C(y, x)$

- The unit element is the product copula  $C^e(x, y) = xy$

- The null element is  $C^{min}(x, y) = \min(x, y)$

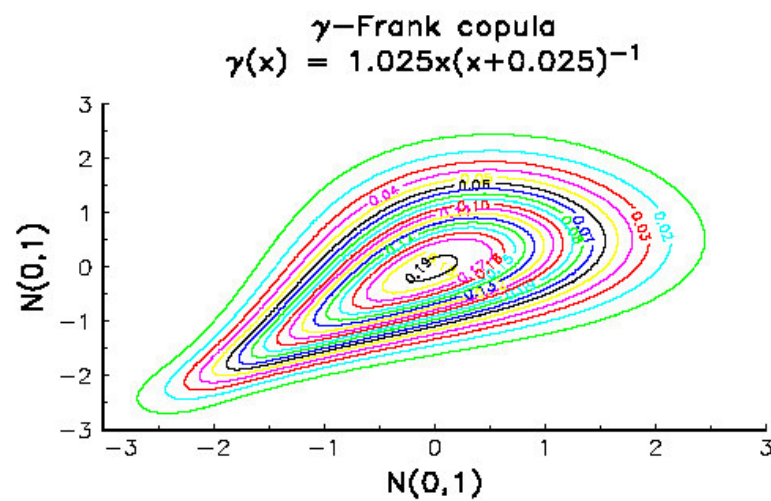
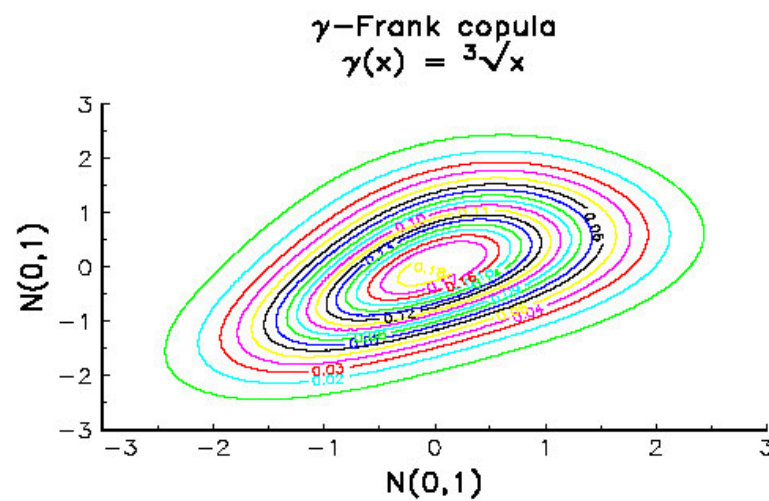
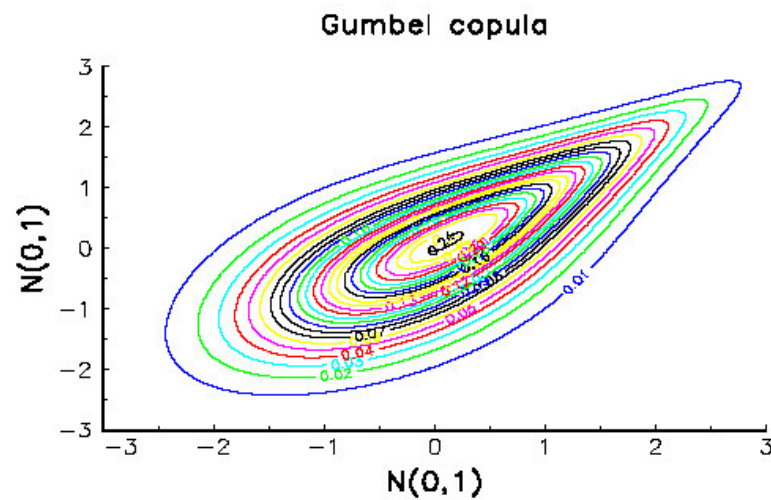
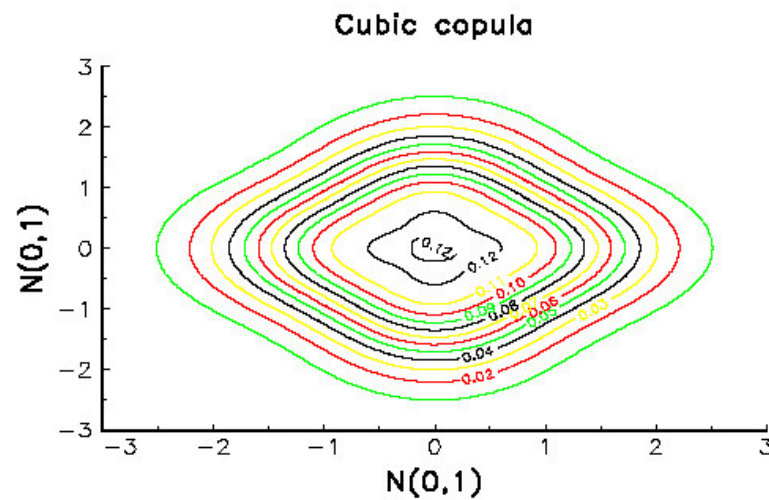
- For all copula  $C$  we have  $C^{min} \leq C \leq C^{max}$  with the stochastic order

- with the gaussian copula we have  $(C^{max} = C_{-1}) < (C^e = C_0) < (C^{min} = C_1)$

- Any element that posses a right or lft inverse is extreme

- exemple  $C^{max} \cdot C^{max} = C^{min}$

# Exemples of equiprobability curves



4 Distributions with Gaussian Margins

## Pricing formulas for

- Double Digital bivariate : payoff = 1 if  $s_1 \leq K_1$  and  $s_2 \leq K_2$  is :

$$OptFwdVal = C^Q(F_1(K_1), F_2(K_2))$$

- Digital Spreadoption: payoff =1 if  $s_2 - s_1 > K$  is

$$: \quad OptFwdVal = \int_{-\infty}^{\infty} (\partial_1 C^Q)(F_1(x), F_2(x+K)) dx$$

- Spreadoption: payoff =  $(s_2 - s_1 - K)^+$  is

$$: \quad OptFwdVal = S_2 - S_1 - K + \int_{-\infty}^{\infty} \int_{-x}^K f_1(x) (\partial_1 C^Q)(F_1(x), F_2(x+y)) dx dy$$

- Basketoption: payoff =  $(s_2 + s_1 - K)^+$  is

$$: \quad OptFwdVal = S_2 + S_1 - K + \int_{-\infty}^{\infty} \int_x^K f_1(x) (\partial_1 C^Q)(F_1(x), F_2(y-x)) dx dy$$

## More Option Formulas

- Max Digital option: payoff = 1 if  $\text{Max}(S_2, S_1) > 1$  is

$$: \quad \text{OptFwdVal} = \int_{-\infty}^K C^Q(F_1(x), F_2(x)) dx$$

- BestOf option 1 : payoff =  $\text{Max}((K_1 - S_1)^+, (K_2 - S_2)^+)$  is

$$: \quad \text{OptFwdVal} = \text{Put}(K_2) + \int_{-\infty}^{K_1} \int_{-\infty}^{\infty} f_2(x) (\partial_2 C^Q)(F_1(\text{Max}(x + y - K_2, x)), F_2(x)) dx dy$$

- BestOf option 2 : payoff =  $\text{Max}((K_1 - S_1)^+, (S_2 - K_2)^+)$  is

$$: \quad \text{OptFwdVal} = \text{Put}(K_2) + \int_{-\infty}^{K_1} \int_{-\infty}^{\infty} f_2(x) (\partial_2 C^Q)(F_1(\text{Max}(x + y - K_2, x)), F_2(x)) dx dy$$

# Stochastic stability of copulas and time dependent copulas

- We look at the following model:

$$dX = (\mu(X, t) - \lambda(X, t))dt + \Sigma(X, t)dW \quad dW \cdot dW = \rho dt$$

- Multivariate constant coefficients  $\Rightarrow$  Girsanov Do not change the copula
- Constant girsanov is not sufficient:

$$\begin{aligned} dX_1 &= \alpha_1(A_1 - X_1)dt + \Sigma_1 dW_1 \\ dX_2 &= \alpha_2(A_2 - X_2)dt + \Sigma_2 dW_2 \end{aligned} \quad dW_1 \cdot dW_2 = \rho dt$$

- Correlation of  $\langle X_1(t), X_2(t) \rangle$  depends on time:

$$\rho(t) = \rho \frac{2\sqrt{\alpha_1\alpha_2}}{\alpha_1 + \alpha_2} \frac{(1 - e^{-(\alpha_1 + \alpha_2)t})}{\sqrt{1 - e^{-\alpha_1 t}} \sqrt{1 - e^{-\alpha_2 t}}}$$

- Conditional copula, pseudo-copula and others

