

Parametric inference on proportions

Overview

Example of a final project

Quick review of Normal distributions and hypothesis tests/CI using normal distributions

Parametric inference on proportions

- Distribution of a sample proportion
- Confidence interval for a single proportion
- Tests for a single proportion

Final project

Final project: analyze your own data set

Final project report: a 5-8 page R Markdown document that contains:

1. Background information:
 - What question you will answer and why it is interesting
 - Where you got the data, and any prior analyses
2. Descriptive plots
3. A hypothesis tests using resampling and parametric methods
4. A confidence interval using the bootstrap and parametric methods
5. A conclusion and reflection
6. Optional: an appendix with extra code (appendix can go over the 8 page limit)

A list of a few data sets you can use are on Canvas

There is also an R Markdown template for the final project on Canvas

Question: do beavers have the same body temperature as humans?



Motivation and data



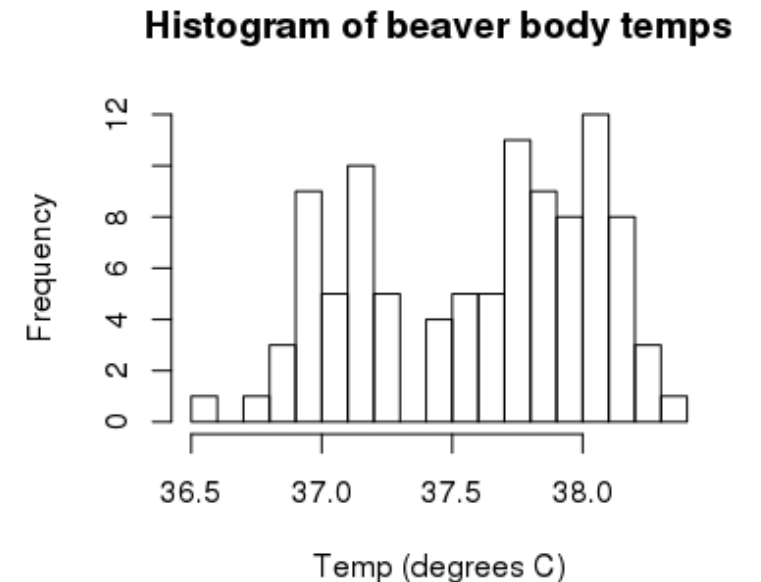
Motivation: There is a labor shortage in the construction industry

- Beavers are a hard working species of animals
- If beavers have the same body temperature as humans (37°C), perhaps they can be employed in the construction industry

The data:

- Body temperatures collected from 400 beavers*
- Data from:
 - Lange et al (1994). In time-series analyses of beaver body temperatures. <https://vincentarelbundock.github.io/Rdatasets/doc/boot/beaver.html>

*not the real data



Results

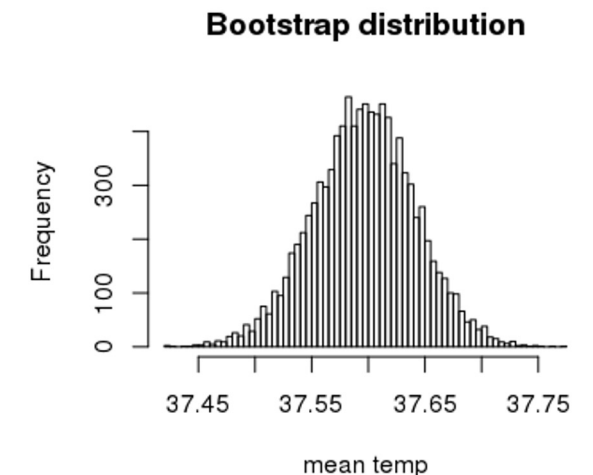
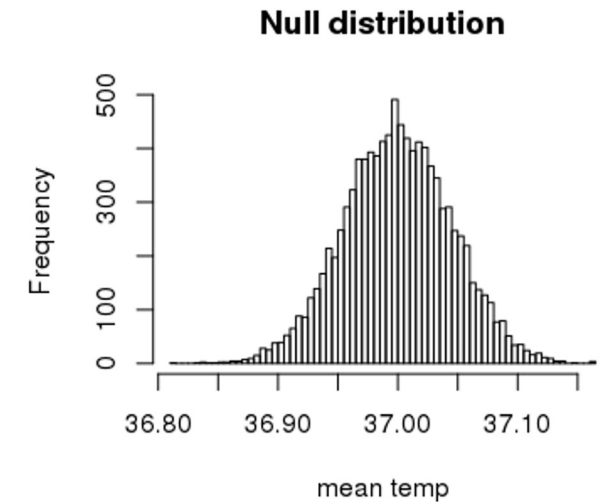
The average human body temperatures is $\mu = 37^{\circ}\text{C}$

Hypothesis test

- $H_0: \mu = 37$ $H_A: \mu \neq 37$
- p-value based on a permutation test: $\bar{x} = 37.6$, p-value = 0
- p-value based on a t-test: $t = 13.35$, $df = 99$, p-value = 0

95% confidence interval for the mean beaver body temp

- Bootstrap: [37.51 37.68]
- t-distribution: [37.51 37.68]



Conclusions

Conclusion: Beavers do not seem to have the same body temperatures as humans

37°C humans vs. 37.6°C beavers

Implications: Due to their higher body temperatures, if beavers join the construction industry they might be too good at their jobs leading to job loss of human workers

Caveats: human body temperatures might not be exactly 37°C

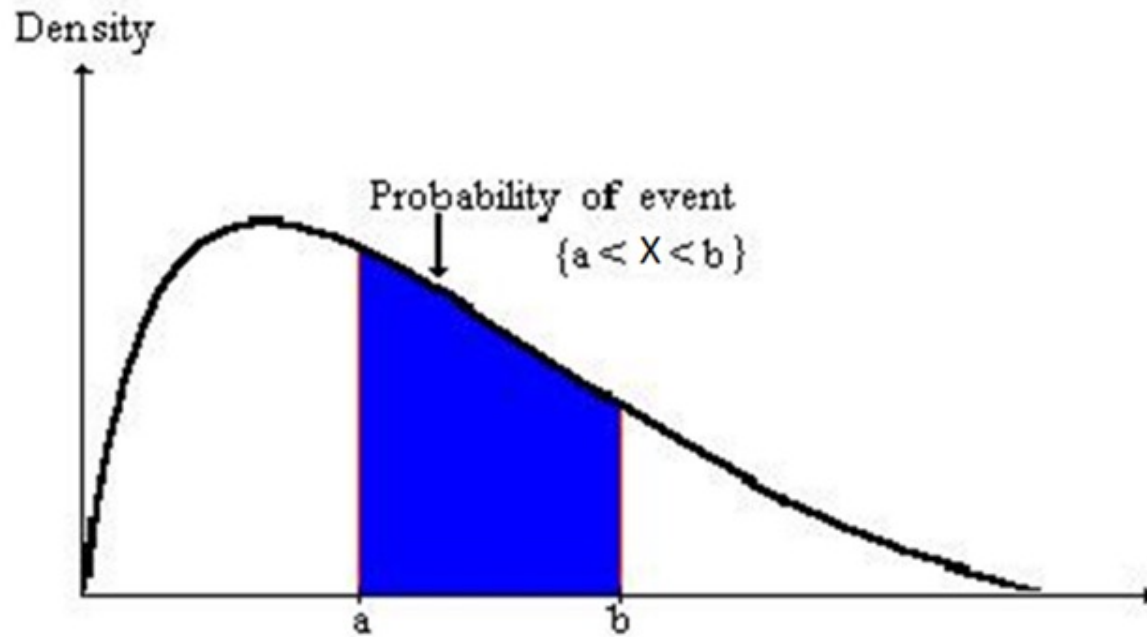


Quick review of Normal distributions

Density Curves

The probability that a random number X will be in the interval $[a, b]$ can be modeled using the area under a density curve

$\Pr(a < X < b)$ is the area under the curve from a to b



Density curve are functions $f(x)$ that have two key properties:

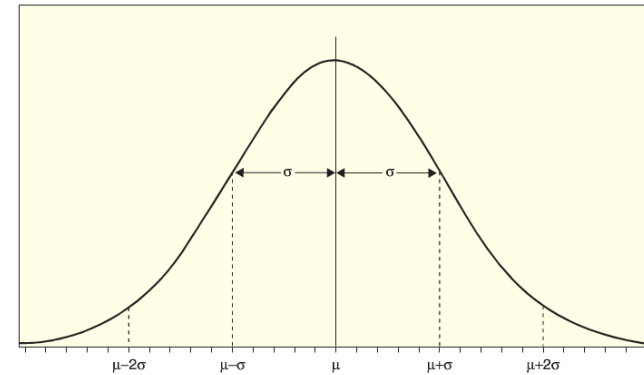
1. The total area under the curve $f(x)$ is equal to 1
2. The curve is always ≥ 0

The Normal Density Curve

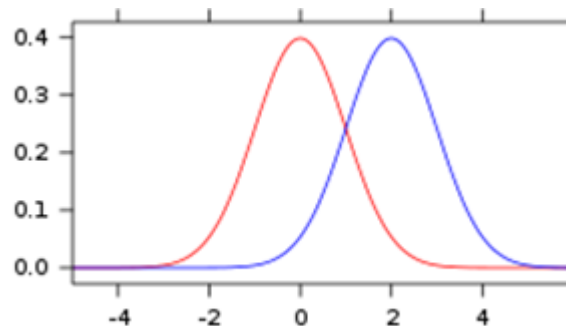
Normal distributions are a family of bell-shaped curves with two parameters

- The mean: μ
- The standard deviation: σ

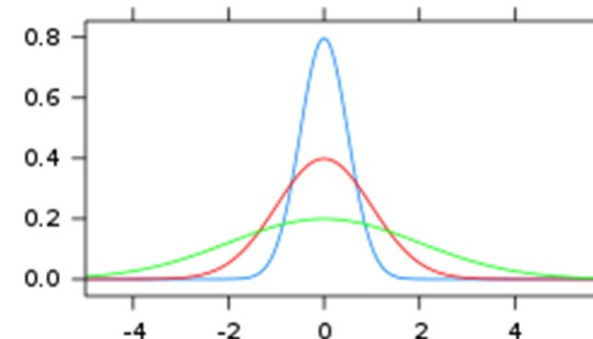
Notation: $X \sim N(\mu, \sigma)$



Changing μ



Changing σ



Densities, probabilities and quantiles from normal distributions

We can plot the density curve using:

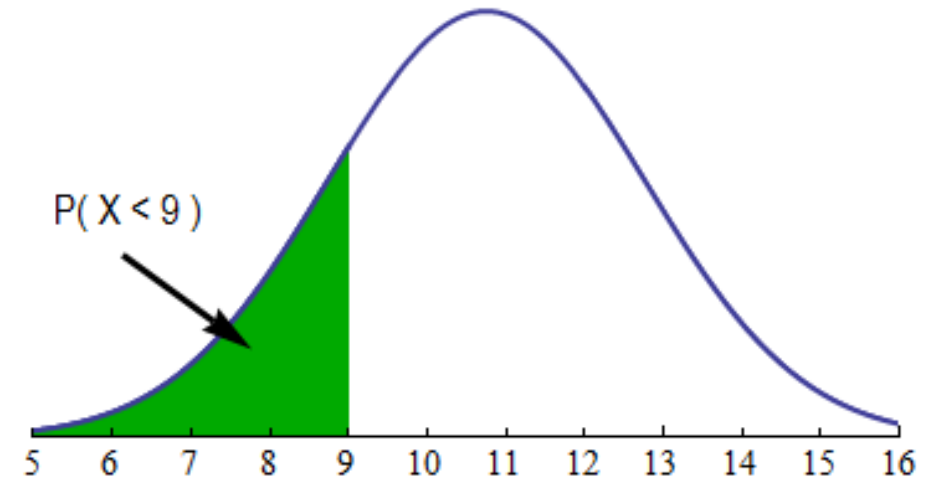
`dnorm(x_vec, mu, sigma)`

We can get the probability that we would get a random value less than x using:

`pnorm(x_vec, mu, sigma)`

We can get the quantile values using:

`qnorm(area, mu, sigma)`



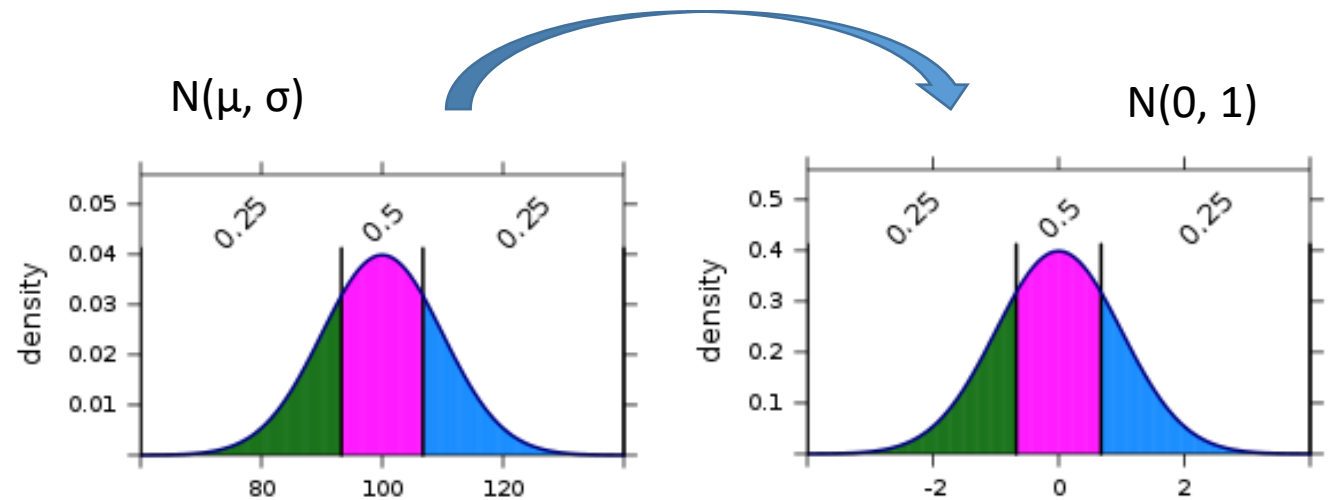
Standard Normal $N(0, 1)$

It is convenient to work with the **standard normal** distribution:

$$Z \sim N(0, 1) \quad \text{i.e., } \mu = 0, \sigma = 1$$

We transform any normally distributed random variable $X \sim N(\mu, \sigma)$ to the standard normal distribution $Z \sim N(0, 1)$ using:

$$Z = (X - \mu) / \sigma$$



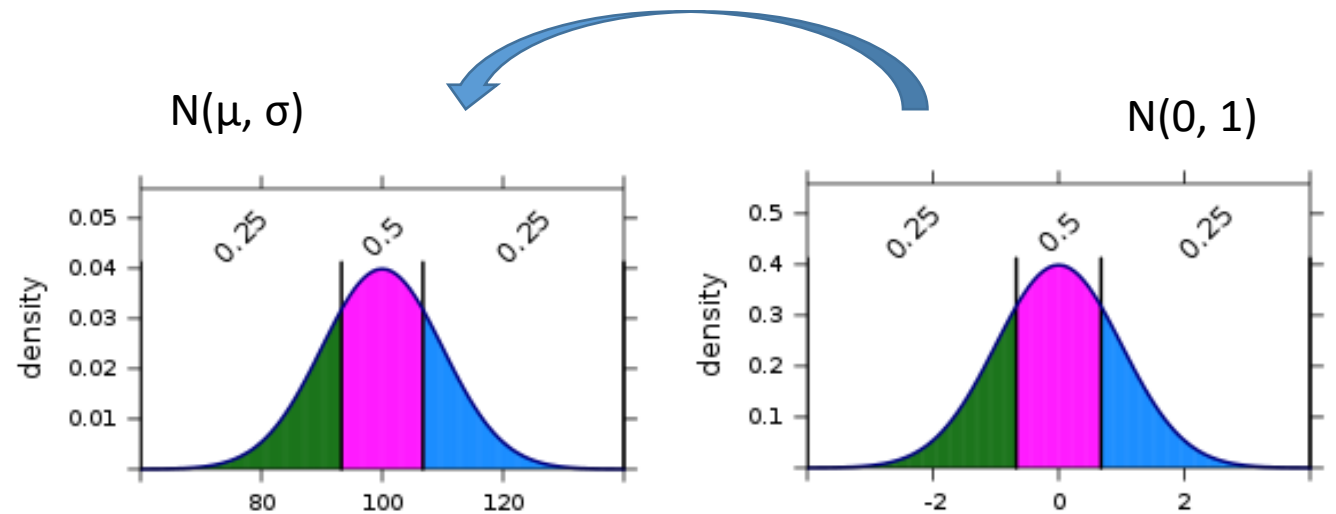
Standard Normal $N(0, 1)$

It is convenient to work with the **standard normal** distribution:

$$Z \sim N(0, 1) \quad \text{i.e., } \mu = 0, \sigma = 1$$

To convert from $Z \sim N(0, 1)$ to any $X \sim N(\mu, \sigma)$, we reverse the standardization with:

$$X = \mu + Z \cdot \sigma$$

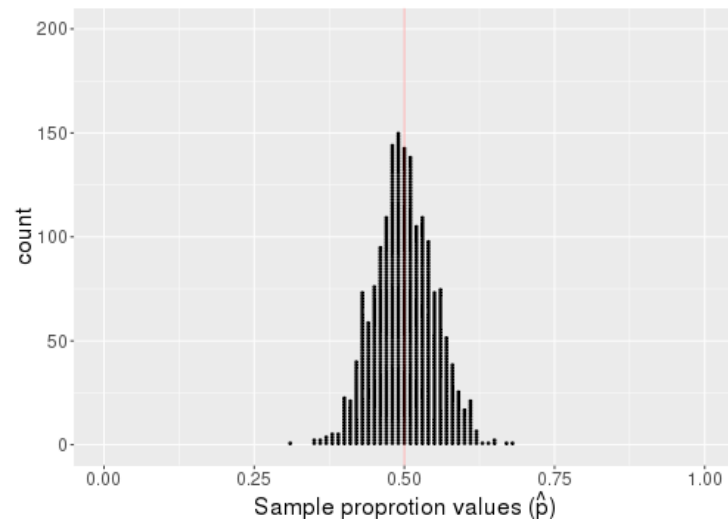


Central limit theorem

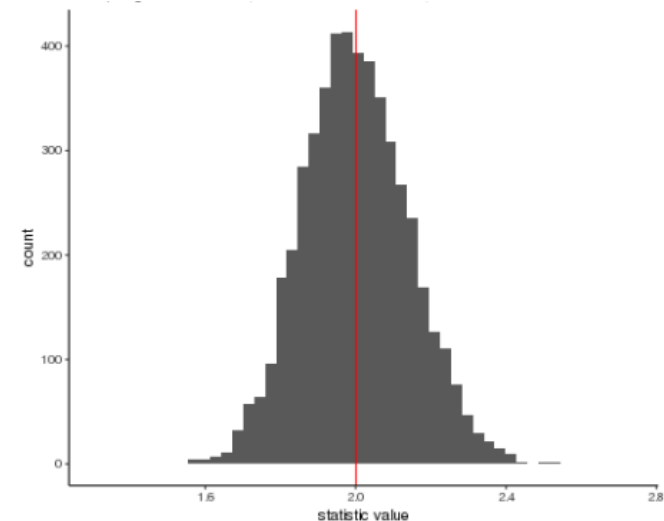
For random samples with a sufficiently large sample size (n), the distribution of sample statistics for a **mean (\bar{x})** or a **proportion (\hat{p})** is:

- normally distributed
- centered at the value of the population parameter

proportion (\hat{p})



mean (\bar{x})

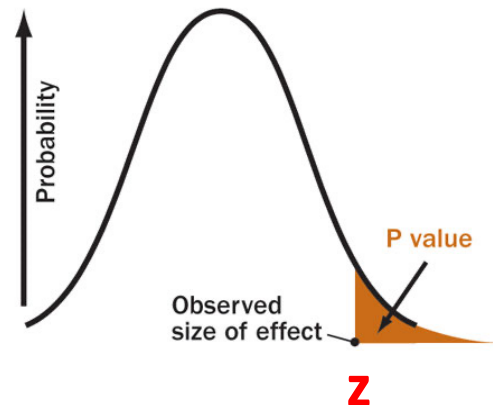


Hypothesis tests based on a Normal Distribution

When the null distribution is normal, it is often convenient to use a standard normal test statistic using:

$$z = \frac{\text{Sample Statistic} - \text{Null Parameter}}{SE}$$

The p-value for the test is the probability a standard normal value is beyond this standardized test statistic



$$\Pr(Z \geq z_{\text{obs}} ; \mu = 0, \sigma = 1)$$

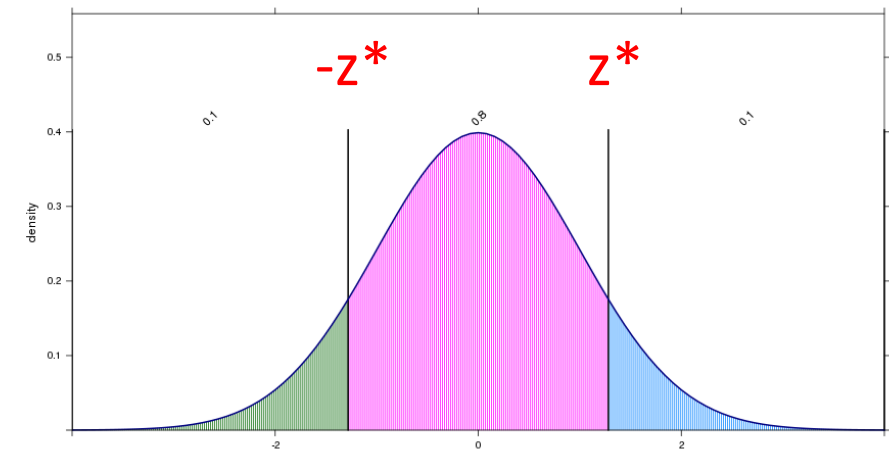
`pnorm(z, 0, 1, lower.tail = FALSE)`

Confidence intervals based on a Normal Distribution

If the distribution for a statistic is normal with a standard error SE, we can find a confidence interval for the parameter using:

$$\text{sample statistic} \pm z^* \times \text{SE}$$

where z^* is chosen so that the area between $-z^*$ and $+z^*$ in the standard normal distribution is the desired confidence level



Confidence level	80%	90%	95%	98%	99%
z^*	1.282	1.645	1.960	2.326	2.576

```
z_stars <- qnorm(c(.90, .95, .975, .99, .995), 0, 1)
```

Parametric inference on proportions

Review: questions about proportions

Q1: What symbols have we been using for the parameter and statistic for proportions?

- Parameter: π
- Statistic: \hat{p}

Q2: What are examples of confidence intervals and hypotheses tests we've run for proportions?

- Hypothesis tests: Doris and Buzz, Paul, Joy, etc.
- Confidence intervals: proportion of red sprinkles, etc.

Review: questions about proportions

Q3: What does the shape of a sampling distribution for a proportions \hat{p} look like?

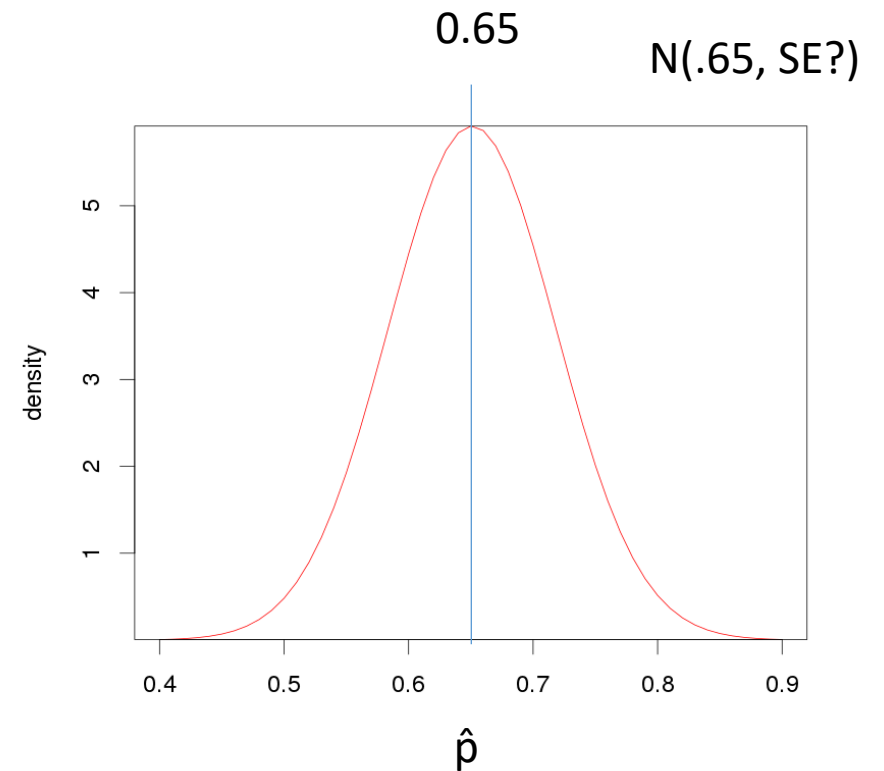
- A: normal!
 - (If the sample size n is larger enough)

Q4: Suppose $\pi = .65$, and $n = 50$, could you draw the sampling distribution for \hat{p} ?

- A: It is centered at 0.65, but what is the spread (SE)?

We could use the bootstrap to estimate the SE with SE^*

Alternatively, we can use a math/theory



Standard Error for Sampling Proportions

When choosing random samples of size n from a population with proportion π , the standard error (SE) of the sample proportions is given by:

$$SE = \sqrt{\frac{\pi(1-\pi)}{n}}$$

The larger the sample size (n) the smaller the standard error (SE)



SE for percentage of houses owned

65.1% of all houses are owned ($\pi = .651$)

If we randomly selected 50 houses...

- a) What would the standard error (SE) of sampling distribution for the proportion of owned houses (\hat{p}) be?
- b) What would this sampling distribution look like?

What if we randomly selected 200 houses?

$$SE = \sqrt{\frac{\pi(1-\pi)}{n}}$$

SE for percentage of houses owned

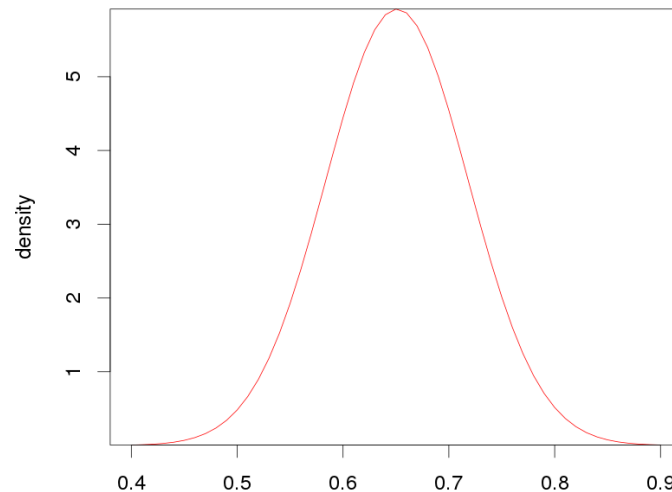
65.1% of all houses are owned

- $\pi = .651$
- When $n = 50$: $SE = .0674$
- When $n = 200$: $SE = .0337$

$$SE = \sqrt{\frac{\pi(1-\pi)}{n}}$$

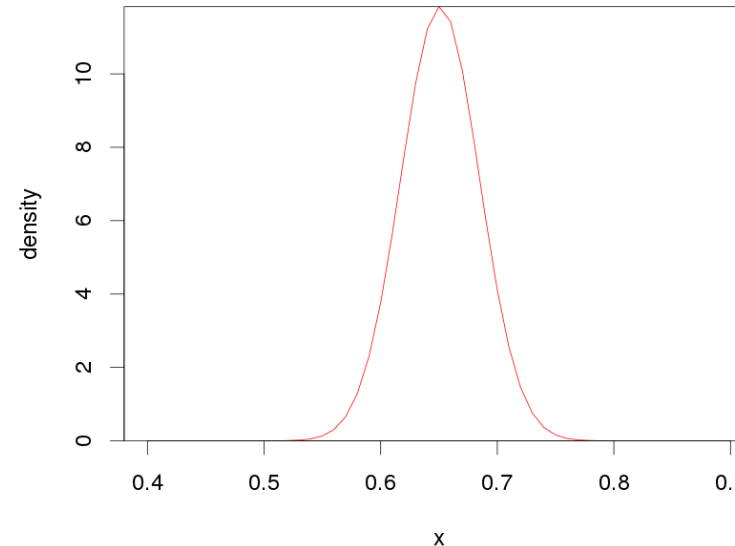
N(.651, .0671)

n = 50



n = 200

N(.651, .0337)



```
y_vals <- dnorm(x_vals, .651, .0674)
```

How large of a sample size n is needed for the sampling distribution of \hat{p} to be normal?

$n = 50$

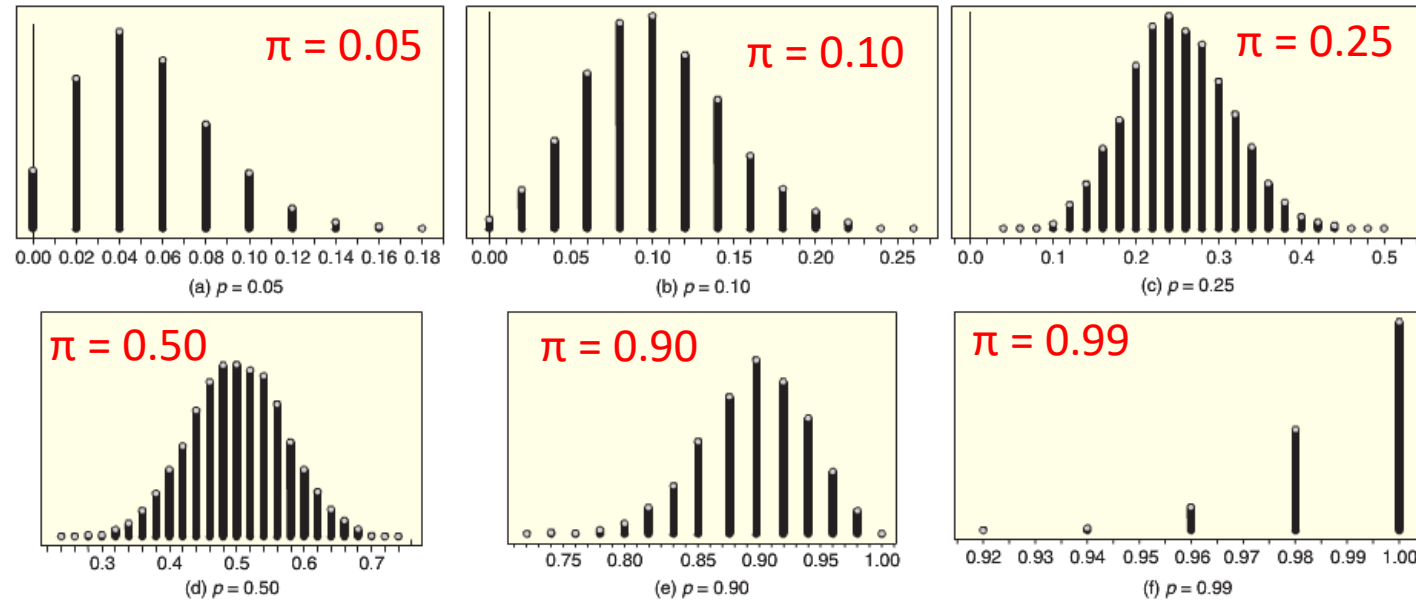


Figure 6.2 Distributions of sample proportions when $n = 50$

$\pi = 0.10$

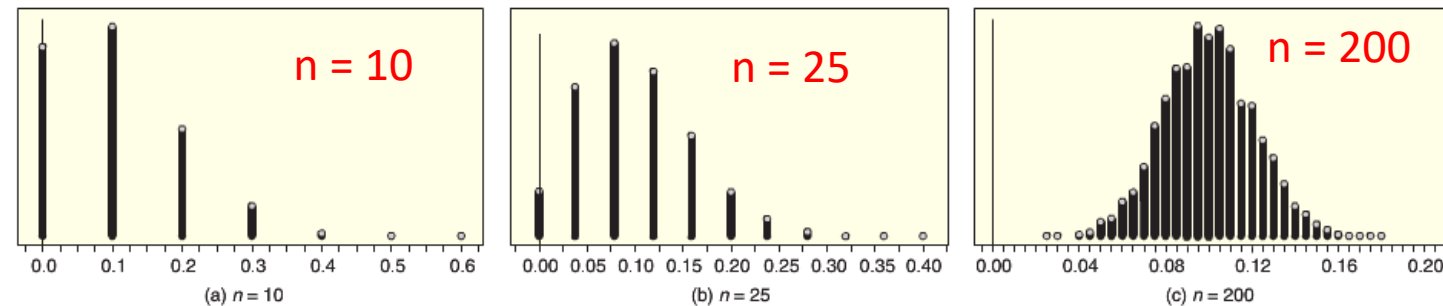


Figure 6.3 Distributions of sample proportions when $p = 0.10$

How large of a sample is needed for the normal approximation?

The normal approximation is reasonable good when we see 10 “positive” outcomes and 10 “negative” outcomes

$$n\pi \geq 10 \quad \text{and} \quad n(1 - \pi) \geq 10$$

Summary: Central Limit Theorem for Sample Proportions

For samples of size n from a population with a proportion π ,
the distribution of the sample proportions has the following characteristics:

Shape: If the sample size is sufficiently large, the distribution is reasonably normal

Center: The mean is equal to the population proportion π

Spread: The standard error is: $SE = \sqrt{\frac{\pi(1-\pi)}{n}}$

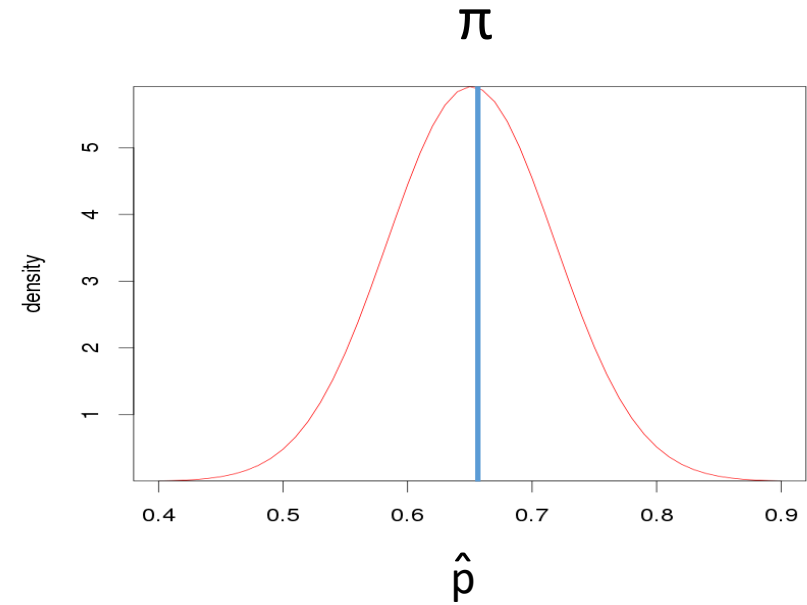
The larger the sample size, the more like a normal distribution it becomes.

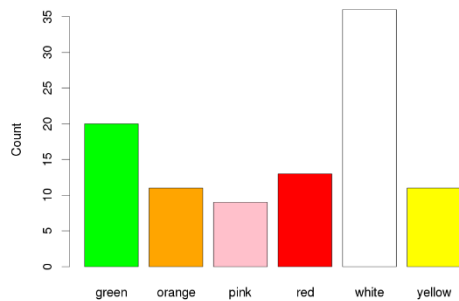
A normal distribution is a good approximation as long as:

$$n\pi \geq 10 \quad \text{and} \quad n(1 - \pi) \geq 10$$

Summary: Central Limit Theorem for Sample Proportions

$$\hat{p} \sim N\left(\pi, \sqrt{\frac{\pi(1-\pi)}{n}}\right)$$



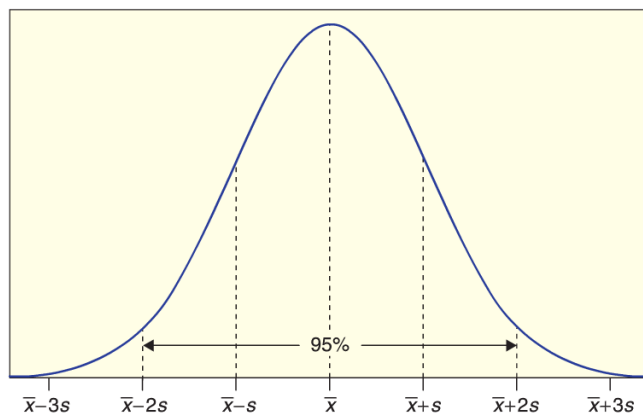


π_{red}

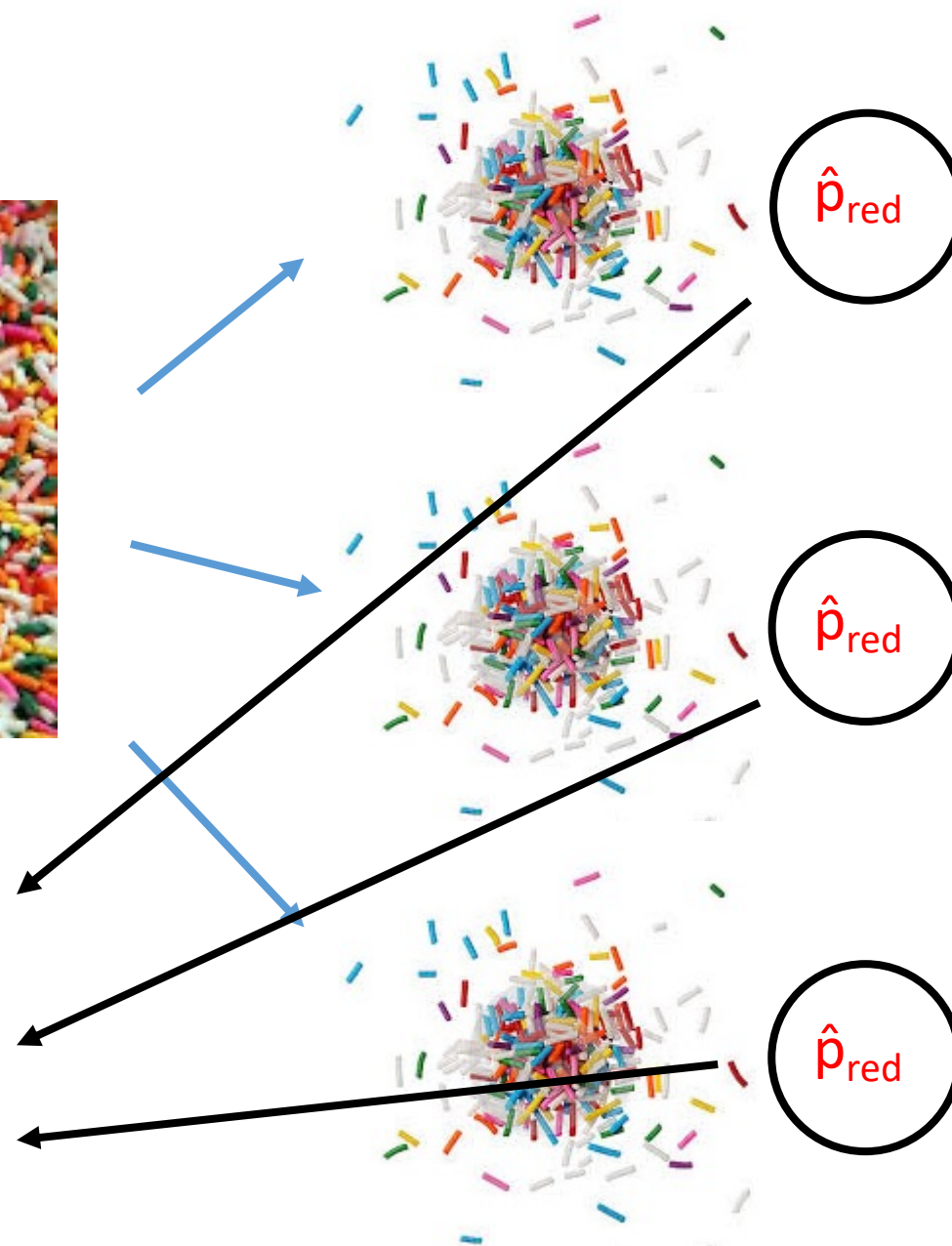


$$\hat{p} \sim N\left(\pi, \sqrt{\frac{\pi(1-\pi)}{n}}\right)$$

$$SE = \sqrt{\frac{\pi(1-\pi)}{n}}$$



Sampling distribution!



Standard Error for Sampling Proportions

Note: we don't usually know π , so we can't compute the standard error exactly using the formula:

$$SE = \sqrt{\frac{\pi(1-\pi)}{n}}$$



However, we can substitute \hat{p} for π and then we can get an estimate of the standard error:

$$\hat{SE} = \sqrt{\frac{\hat{p}(1-\hat{p})}{n}}$$



Comparing formula SE to the bootstrap SE

In previous classes we have used the bootstrap to get an estimate of the standard error SE*

How could we do this for the green sprinkles?

```
bootstrap_dist <- do_it(100000) * {  
  boot_sample <- sample(my_sprinkles, replace = TRUE)  
  sum(boot_sample == 'green')/100  
}
```

```
bootstrap_SE <- sd(bootstrap_dist)
```



Color
White
Red
Red
White
Green
White
.
.
.
White
Green

n = 100 sprinkles

Comparing formula SE to the bootstrap SE

For my green sprinkles I got:

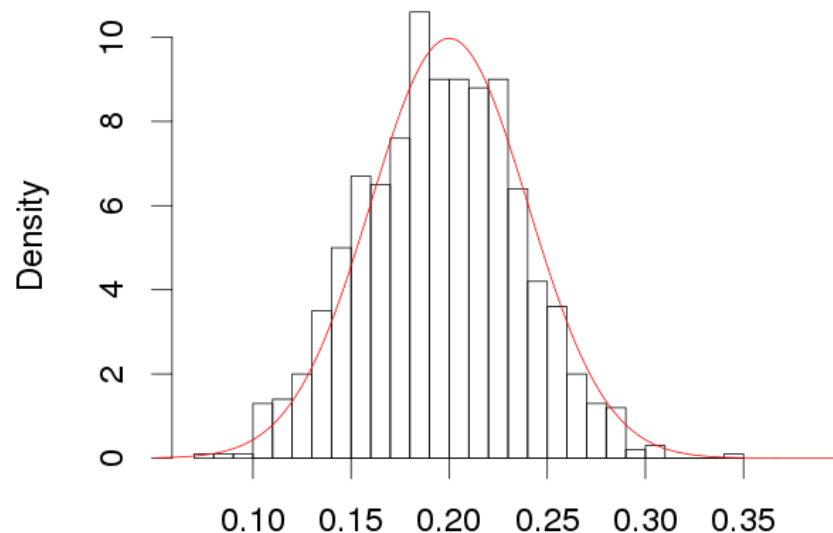
- Bootstrap SE = 0.039959
- Formula SE = 0.04

$$\hat{p} = 0.20$$

$$n = 100$$

$$\hat{SE} = \sqrt{\frac{\hat{p}(1-\hat{p})}{n}}$$

Bootstrap Distribution



```
SE <- sqrt( (.2 * (1 - .2) ) / 100 )
```

Parametric confidence intervals for proportions

Confidence intervals for a single proportion

Suppose we have a sample of size n of categorical data

Suppose that n is large enough so that $n\pi \geq 10$ and $n(1 - \pi) \geq 10$

A confidence interval for a population proportion π can be computed from our random sample of size n using:

$$\hat{p} \pm z^* \cdot \sqrt{\frac{\hat{p}(1 - \hat{p})}{n}}$$

Equation for SE



Where \hat{p} is the sample proportion and z^* is a standard normal endpoint to give the desired confidence level

Sprinkle example

To create a confidence interval for proportion of green sprinkles π_{green} we take a sample of size $n = 100$



1	orange
2	red
3	green
4	white
5	white
6	white
7	white
8	white
9	red

My green sprinkles

20 of the 100 sprinkles were green

What is a 95% confidence interval for the population proportion π of green sprinkles?

$$\hat{p} \pm z^* \cdot \sqrt{\frac{\hat{p}(1-\hat{p})}{n}}$$

My green sprinkles

$$\hat{p} = 20/100 = .20$$

$$n = 100$$

$$SE = .04$$

$$z^* = 1.96 \text{ (for 95\% CI)}$$

$$CI = 0.1216 \text{ to } 0.2784$$

$$\hat{p} \pm z^* \cdot \sqrt{\frac{\hat{p}(1-\hat{p})}{n}}$$

$$.20 \pm 1.96 \cdot \sqrt{\frac{.2 \cdot (1-.2)}{100}}$$

Parametric hypothesis tests for proportions

Test for single proportions

To compute p-values when the null distribution is normal we use:

$$z = \frac{\text{Sample Statistic} - \text{Null Parameter}}{SE}$$

In the context of proportions our null hypothesis is of the form $H_0: \pi = \pi_0$

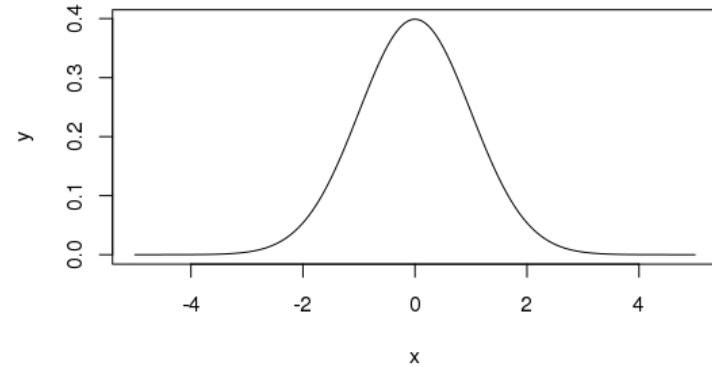
Our formula for z then becomes:

$$z = \frac{\hat{p} - \pi_0}{SE} \qquad SE = \sqrt{\frac{\pi_0(1 - \pi_0)}{n}}$$

Test for single proportions

To test for $H_0: \pi = \pi_0$ vs $H_A: \pi \neq \pi_0$ (or the one-tail alternative), we use the standardized test statistic:

$$z = \frac{\hat{p} - \pi_0}{\sqrt{\frac{\pi_0(1 - \pi_0)}{n}}}$$



Where \hat{p} is the proportion in a random sample of size n

Provided the sample size is reasonable large (usual conditions), the p-value of the test is computed using the standard normal distribution

Do more than 25% of US adults believe in ghosts?

A telephone survey of 1000 randomly selected US adults found that 31% of them say they believe in ghosts. Does this provide evidence that more than 1 in 4 US adults believe in ghosts?

1. State the null and alternative hypothesis
2. Calculate the statistic of interest
- 3-4. Calculate the p-value
Hint: the `pnorm()` function will be useful
5. What do you conclude?

$$z = \frac{\hat{p} - \pi_0}{\sqrt{\frac{\pi_0(1 - \pi_0)}{n}}}$$

Do more than 25% of US adults believe in ghosts?

Step 1:


$$H_0: \pi = .25$$

$$H_A: \pi > .25$$

Step 2:

$$\hat{p} = .31$$

$$n = 1000$$

$$z = \frac{\hat{p} - \pi_0}{\sqrt{\frac{\pi_0(1 - \pi_0)}{n}}}$$


```
SE <- sqrt( (.25 * (1 - .25))/1000)
```

```
z_val <- (.31 - .25)/SE
```

z_val is 4.38

Do more than 25% of US adults believe in ghosts?

Step 1:

$$H_0: \pi = .25$$

$$H_A: \pi > .25$$

Step 2:

$$z_val \leftarrow 4.38$$

Step 3-4:

$$p\text{-value} = \text{pnorm}(z_val, 0, 1, \text{lower.tail} = \text{FALSE})$$

Step 5:

Indeed, very strong evidence!



Sinister lawyers

10% of American population is left-handed

A study found that out of a random sample of 105 lawyers, 16 were left-handed

Test whether the proportion of left-handed lawyers is greater than the proportion found in the American population.

1. State the null and alternative hypothesis
- 2-4. Calculate the statistic of interest and calculate the p-value
 - Hint: the `pnorm()` function will be useful
5. What do you conclude?

$$z = \frac{\hat{p} - \pi_0}{\sqrt{\frac{\pi_0(1 - \pi_0)}{n}}}$$

Sinister lawyers

1. State the null and alternative hypothesis

- $H_0: \pi = .10$
- $H_A: \pi > .10$

2-4. Calculate the statistic of interest and the p-value

- $\hat{p} = 16/105 = .152$
- $SE = \sqrt{(.10 * (1 - .10))/105} = .029$
- $z = (.152 - .10)/.029 = 1.79$
- $\text{pnorm}(z, 0, 1, \text{lower.tail} = \text{FALSE}) = .037$

$$z = \frac{\hat{p} - \pi_0}{\sqrt{\frac{\pi_0(1 - \pi_0)}{n}}}$$

5. What do you conclude?

