

Parametric inference on proportions

Overview

Review and continuation of using normal distributions for inference

Parametric inference on proportions

- Normal sampling distributions for proportions and formulas for the standard error
- Parametric confidence intervals for proportions
- Parametric hypothesis tests for proportions

Review and continuation of using normal
distributions for inference

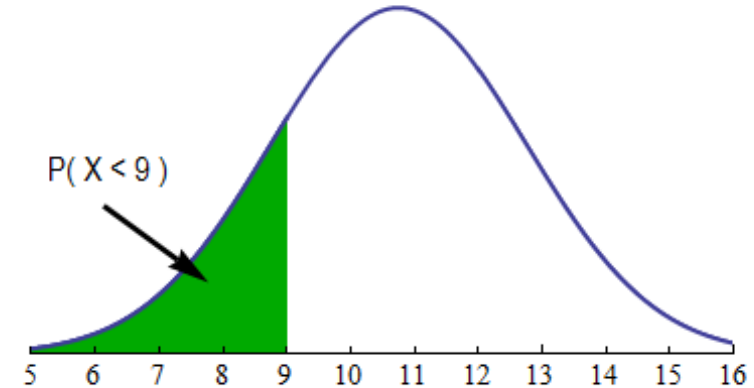
Review: Normal probability functions

Generate random data

- `rnorm(m, mean, sd)`

Plot the density curve

- `dnorm(x_vec, mean, sd)`



Get the probability that we would get a random value less than x: $P(X < x)$

- `pnorm(x_vec, mean, sd)`

Get the quantile value for a given proportion of the distribution

- `qnorm(area, mean, sd)`

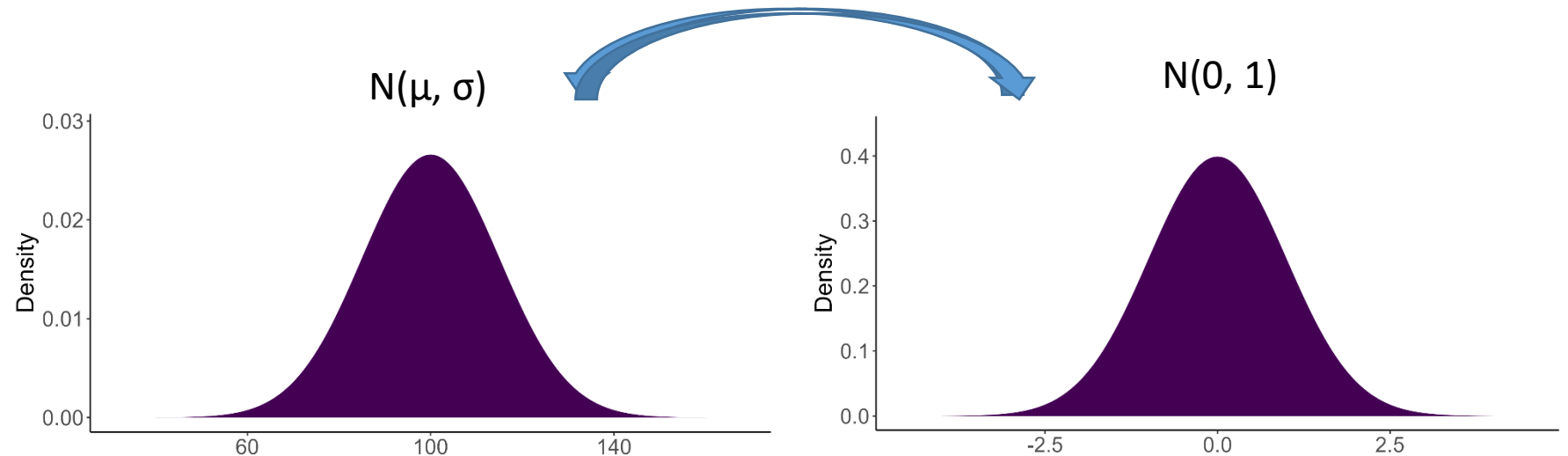
Review: Converting to the standard normal distribution

We can apply a z-score transformation to any normally distributed random variable $X \sim N(\mu, \sigma)$ to convert it to the standard normal distribution $Z \sim N(0, 1)$:

$$Z = (X - \mu) / \sigma$$

To convert from $Z \sim N(0, 1)$ to any $X \sim N(\mu, \sigma)$, we reverse the standardization with:

$$X = \mu + Z \cdot \sigma$$



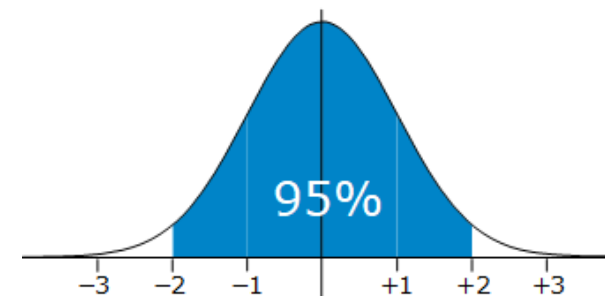
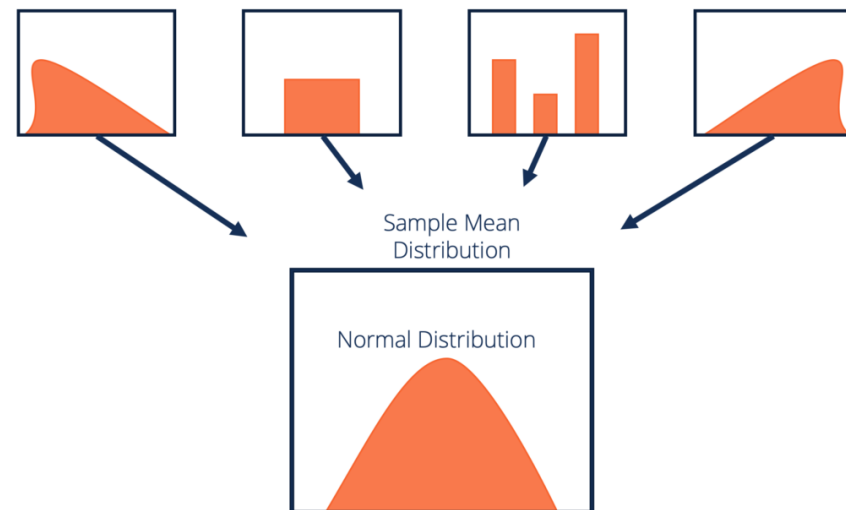
Review: Central limit theorem

For random samples with a sufficiently large sample size (n), the distribution of sample statistics for a **mean** (\bar{x}) or a **proportion** (\hat{p}) is:

- Normally distributed
- Centered at the value of the population parameter

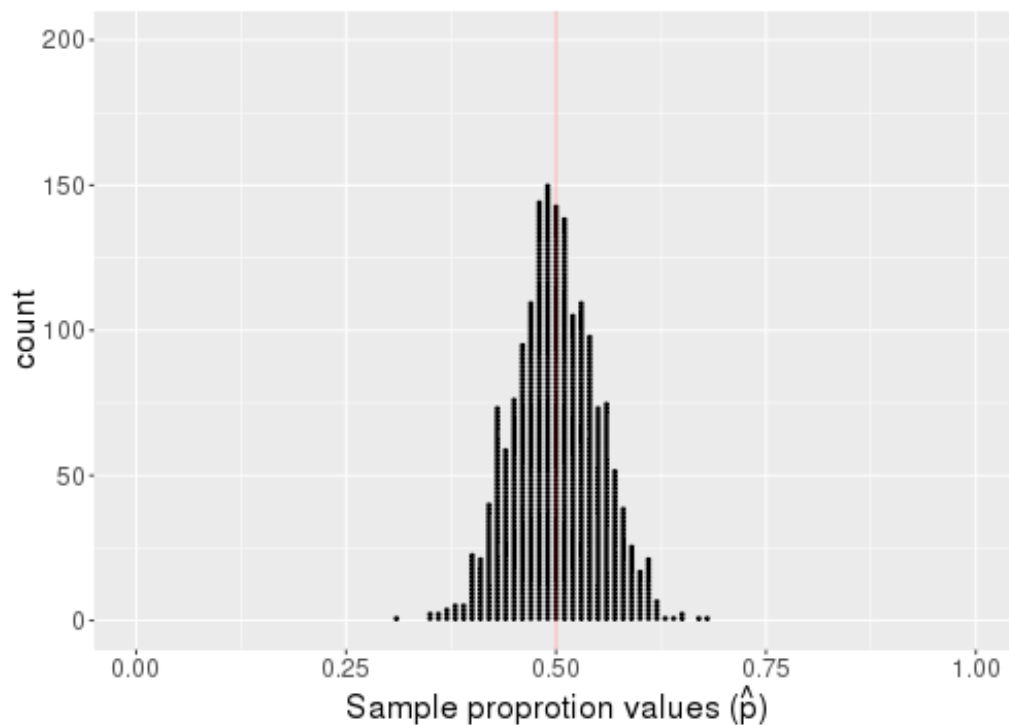
Upshot: We can create confidence intervals and run hypothesis tests using the normal distribution

- Rather than using computational methods like the bootstrap or a randomization methods to create a null distribution!



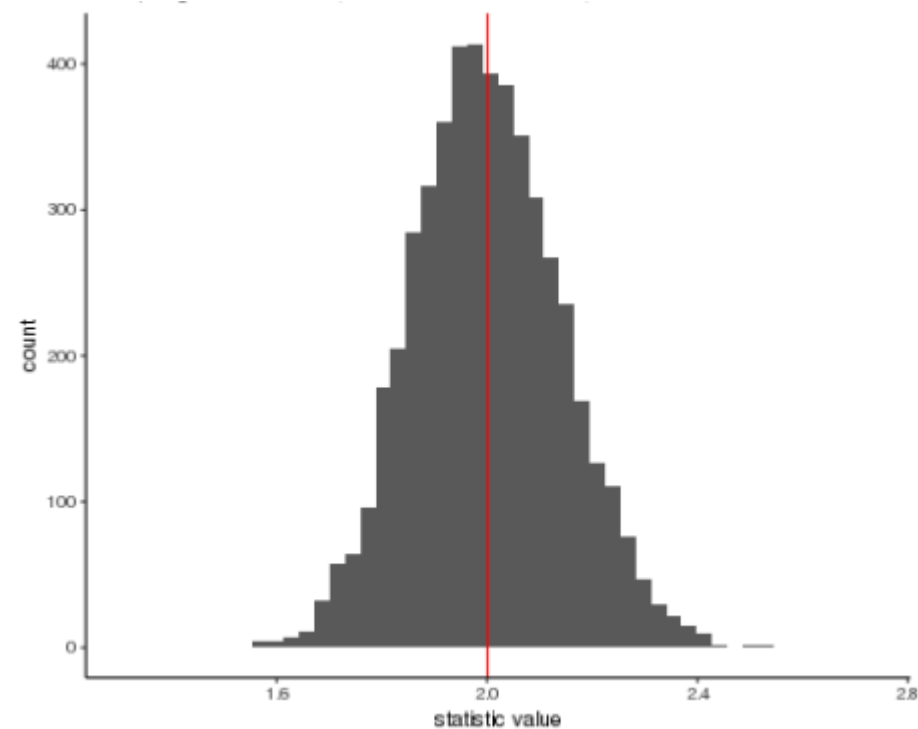
Review: Central limit theorem

proportion (\hat{p})



[Proportion sampling distribution app](#)

mean (\bar{x})



[Sampling/Bootstrap distribution app](#)

The plan

For large n , the sampling distributions of \bar{x} and \hat{p} have normal distributions

We can convert any normal distribution $N(\mu, \sigma)$, into a standard normal distribution $N(0, 1)$

We can then use the standard normal distribution for inference

- I.e., To create confidence intervals and run hypothesis tests

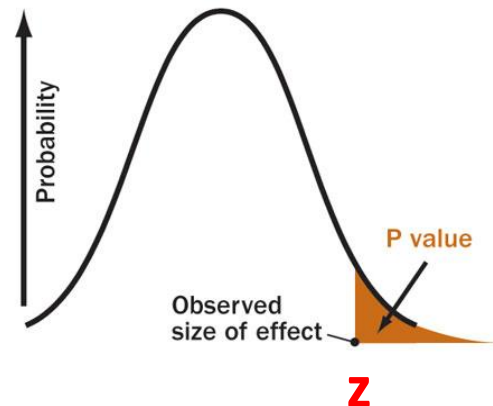
Hypothesis tests and confidence intervals using a normal distribution

Hypothesis tests based on a Normal Distribution

When the null distribution is normal, it is often convenient to use a standard normal test statistic using:

$$z = \frac{\text{Sample Statistic} - \text{Null Parameter}}{SE}$$

The p-value for the test is the probability a standard normal value is beyond this standardized test statistic



$$P(Z \geq z_{\text{obs}}; \mu = 0, \sigma = 1)$$

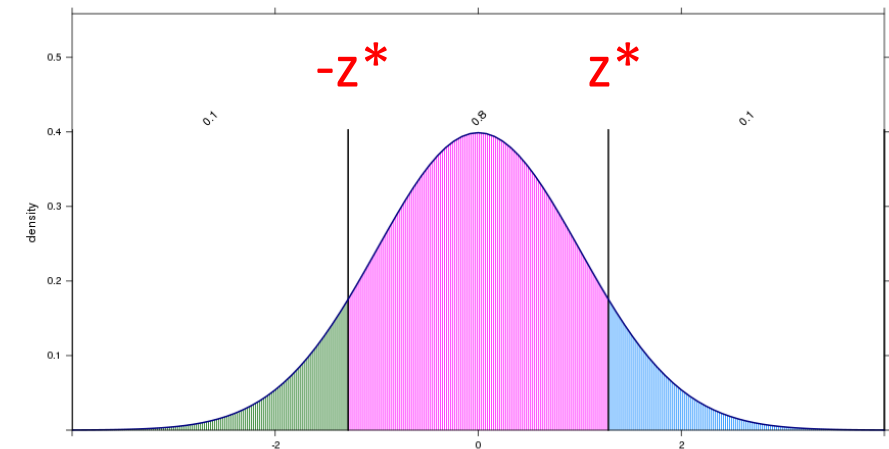
`pnorm(z, 0, 1, lower.tail = FALSE)`

Confidence intervals based on a Normal Distribution

If the distribution for a statistic is normal with a standard error SE, we can find a confidence interval for the parameter using:

$$\text{sample statistic} \pm z^* \times \text{SE}$$

where z^* is chosen so that the area between $-z^*$ and $+z^*$ in the standard normal distribution is the desired confidence level



| Confidence level | 80% | 90% | 95% | 98% | 99% |
|------------------|-------|-------|-------|-------|-------|
| z^* | 1.282 | 1.645 | 1.960 | 2.326 | 2.576 |

```
z_stars <- qnorm(c(.90, .95, .975, .99, .995), 0, 1)
```

```
mosaic::cnorm()
```

Do goalies guess the direction of a penalty shot less than 50% of the time?

1. Start by stating H_0 and H_A

$$H_0: \pi = .5$$

$$H_A: \pi < .5$$

2. Calculate the observed statistic

- Goal keepers correctly guessed the direction 41% of the time out of 128 kicks
- With $SE^* = 0.043$
 - (could you calculate this SE^* ?)

Can you compute a z-statistic?



$$z = \frac{\text{Sample Statistic} - \text{Null Parameter}}{SE}$$

Let's try the rest in R!

Do goalies guess the direction of a penalty shot less than 50% of the time?

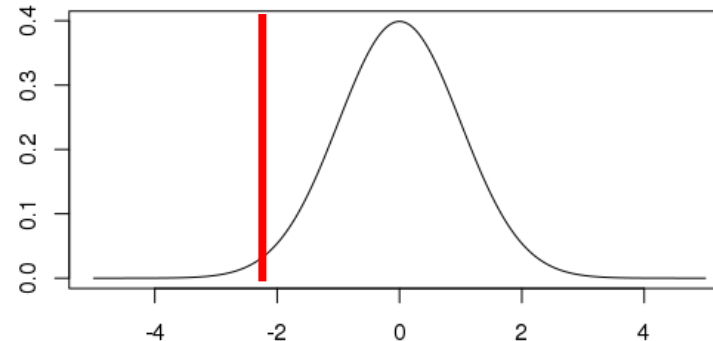
Steps: 3-4. What is the probability one would get a z-statistic as small or smaller than -2.093 from a standard normal distribution?

`pnorm(-2.093, 0, 1)`

Normal area app $P(X \leq x)$

p-value = 0.018

Standard normal null distribution



Step 5?



Parametric inference on proportions

Review: questions about proportions

Q₁: What symbols have we been using for the parameter and statistic for proportions?

- Parameter: π
- Statistic: \hat{p}

Q₂: What are examples of confidence intervals and hypotheses tests we've run for proportions?

- Hypothesis tests: Doris and Buzz, Paul the Octopus, etc.
- Confidence intervals: proportion of red sprinkles, etc.

Review: questions about proportions

Q₃: What does the shape of a sampling distribution for a proportions \hat{p} look like?

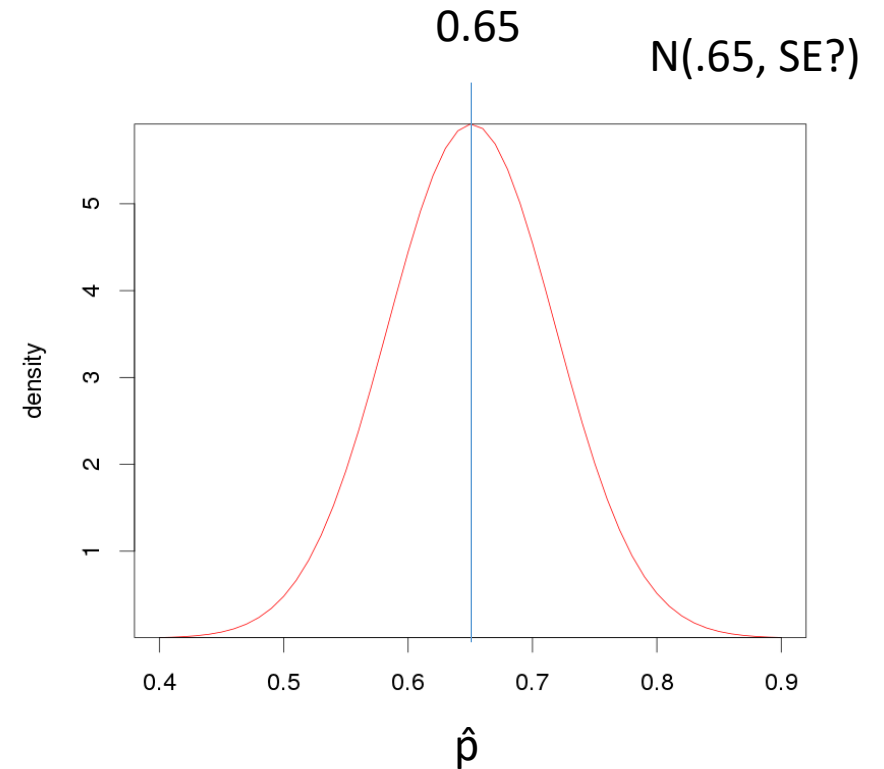
- A: normal!
 - (If the sample size n is larger enough)

Q₄: Suppose $\pi = .65$, and $n = 50$, could you draw the sampling distribution for \hat{p} ?

- A: It is centered at 0.65, but what is the spread (SE)?

We could use the bootstrap to estimate the SE with SE*

Alternatively, we can use a math/theory



Standard Error for Sampling Proportions

When choosing random samples of size n from a population with proportion π , the standard error (SE) of the sample proportions is given by:

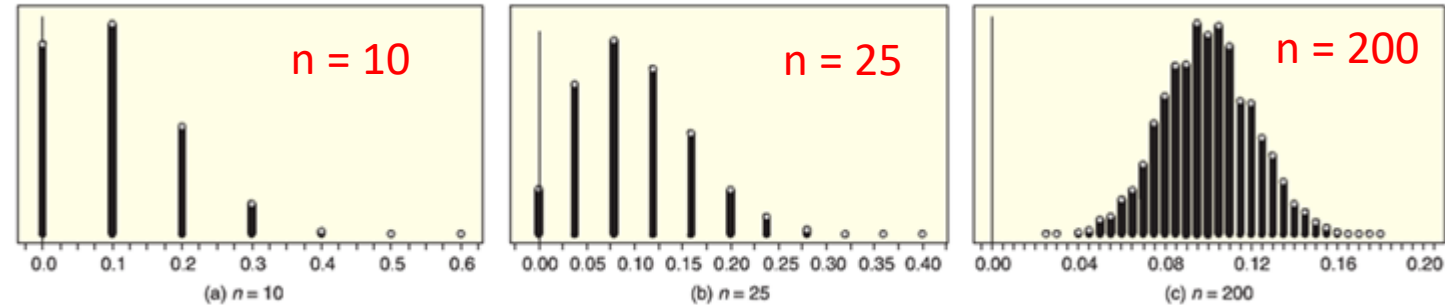
$$SE = \sqrt{\frac{\pi(1-\pi)}{n}}$$

The larger the sample size (n) the smaller the standard error (SE)

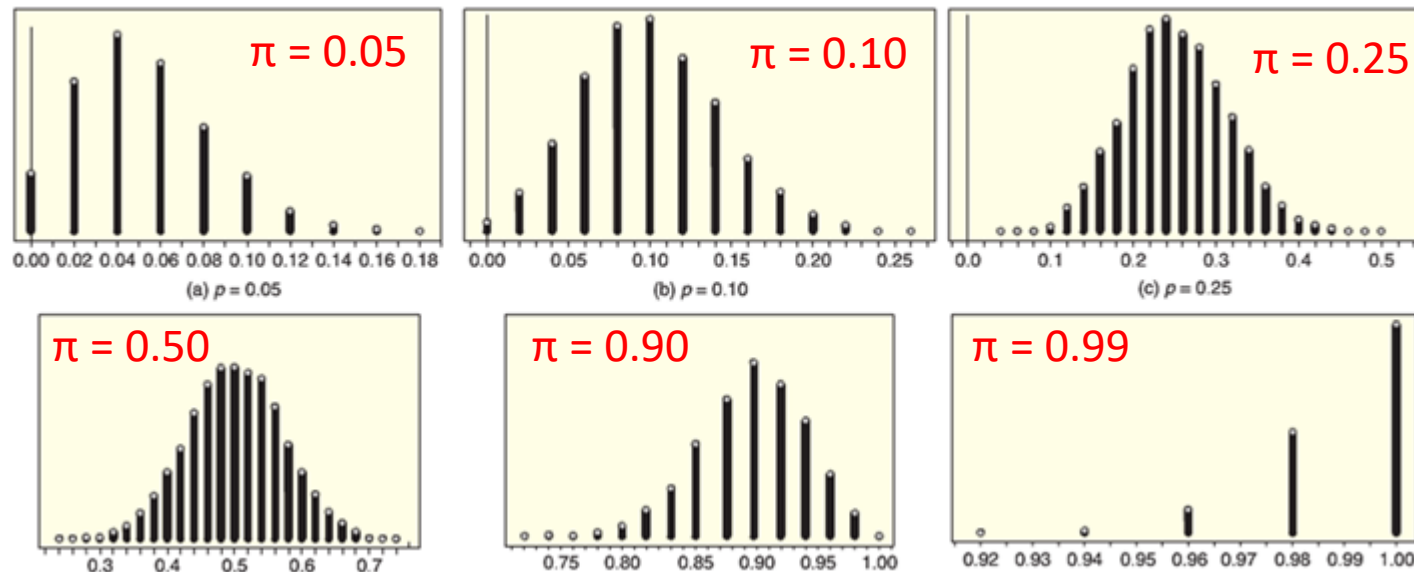


How large of a sample size n is needed for the sampling distribution of \hat{p} to be normal?

Fixed $\pi = 0.10$
Changing n



Fixed $n = 50$
Changing π



How large of a sample is needed for the normal approximation?

The normal approximation is reasonably good when we see 10 “positive” outcomes and 10 “negative” outcomes

$$n\pi \geq 10 \quad \text{and} \quad n(1 - \pi) \geq 10$$

Summary: Central Limit Theorem for Sample Proportions

For samples of size n from a population with a proportion π ,
the distribution of the sample proportions has the following characteristics:

Shape: If the sample size is sufficiently large, the distribution is reasonably normal

Center: The mean is equal to the population proportion π

Spread: The standard error is: $SE = \sqrt{\frac{\pi(1-\pi)}{n}}$

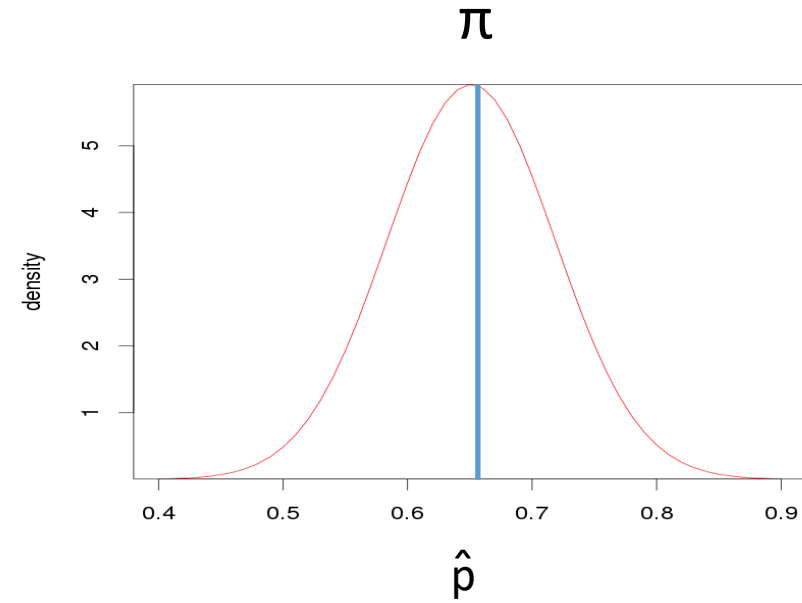
The larger the sample size, the more like a normal distribution it becomes.

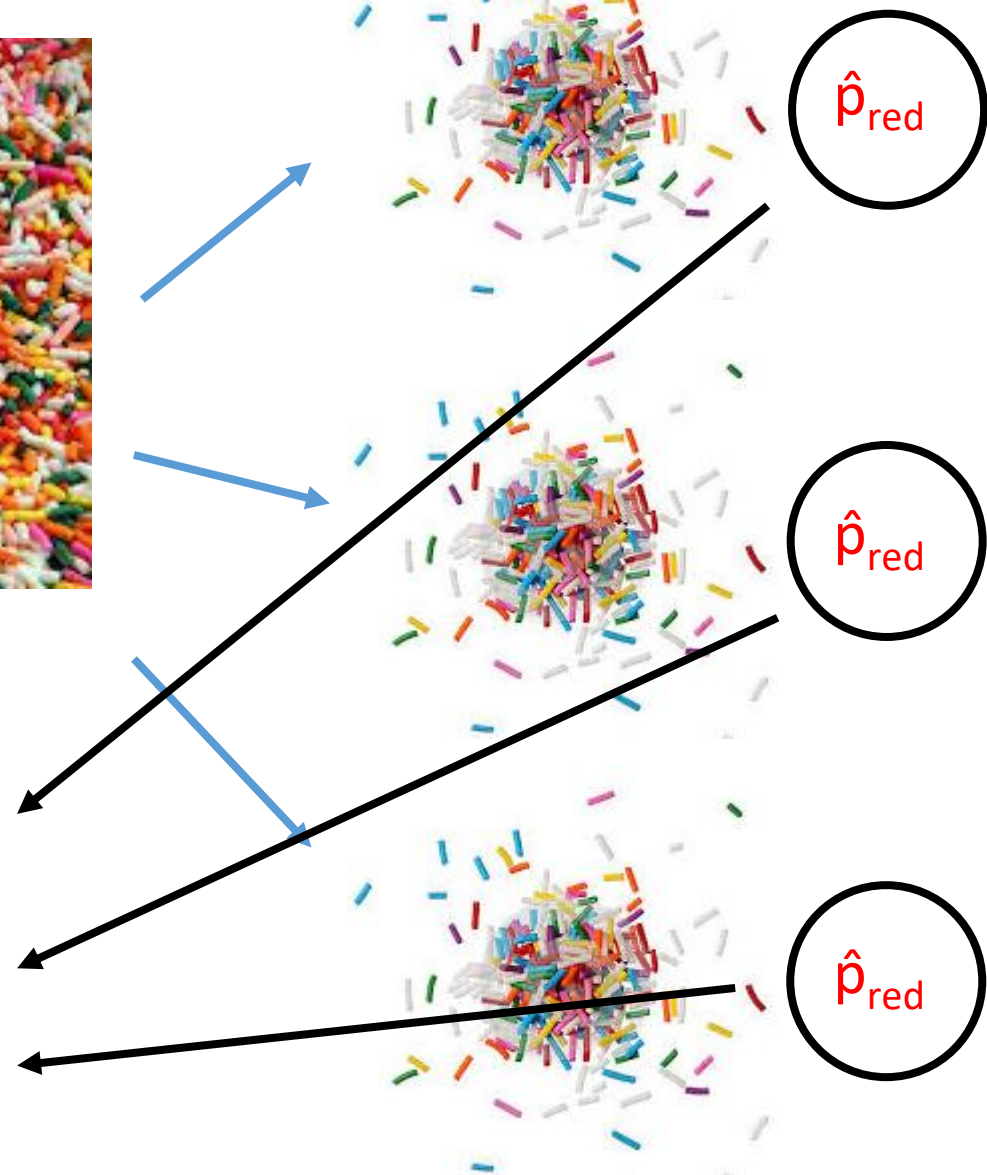
A normal distribution is a good approximation as long as:

$$n\pi \geq 10 \quad \text{and} \quad n(1 - \pi) \geq 10$$

Summary: Central Limit Theorem for Sample Proportions

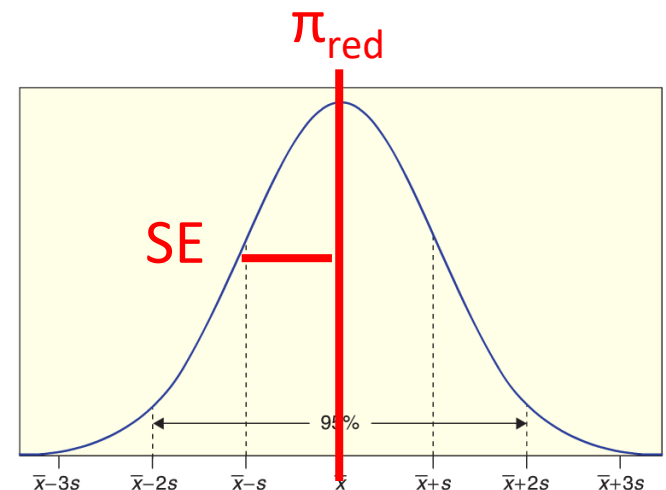
$$\hat{P} \sim N\left(\pi, \sqrt{\frac{\pi(1-\pi)}{n}}\right)$$





$$SE = \sqrt{\frac{\pi(1-\pi)}{n}}$$

$$\hat{P} \sim N(\pi, \sqrt{\frac{\pi(1-\pi)}{n}})$$



Sampling distribution!

SE for percentage of houses owned

65.1% of all houses are owned ($\pi = .651$)

If we randomly selected 50 houses...

- a) What is the standard error (SE) of sampling distribution for the proportion of owned houses (\hat{p}) owned?
- b) What would this sampling distribution look like?

What if we randomly selected 200 houses?

$$SE = \sqrt{\frac{\pi(1-\pi)}{n}}$$

Let's try it in R!

SE for percentage of houses owned

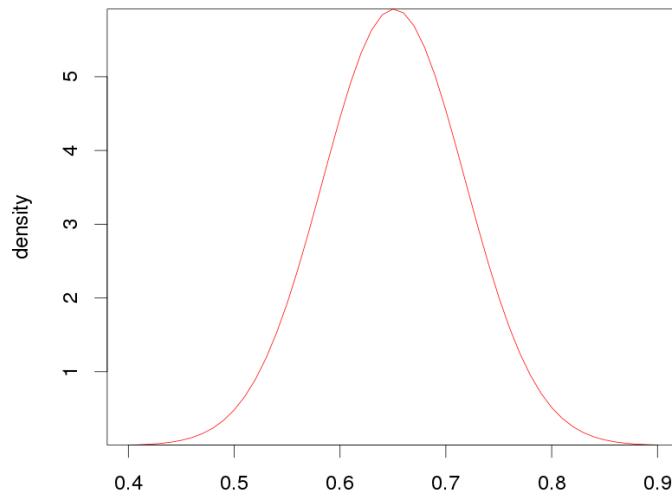
65.1% of all houses are owned

- $\pi = .651$
- When $n = 50$: $SE = .0674$
- When $n = 200$: $SE = .0337$

$$SE = \sqrt{\frac{\pi(1-\pi)}{n}}$$

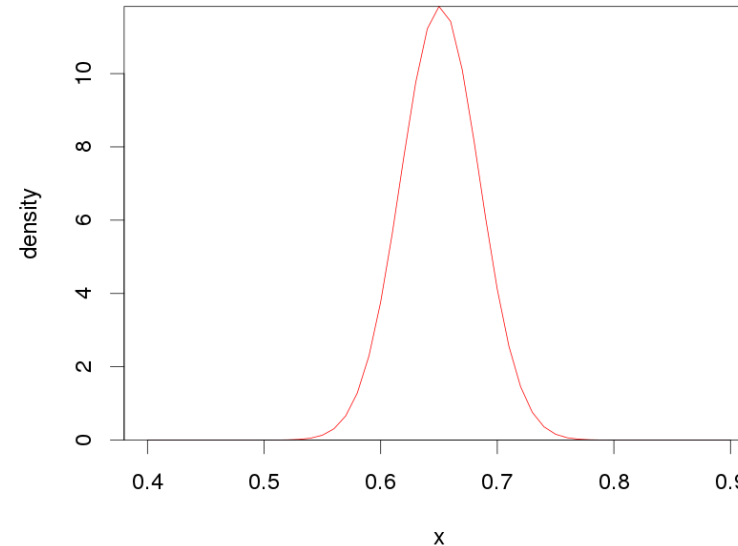
$N(.651, .0671)$

$n = 50$



$n = 200$

$N(.651, .0337)$



```
y_vals <- dnorm(x_vals, .651, .0674)
```

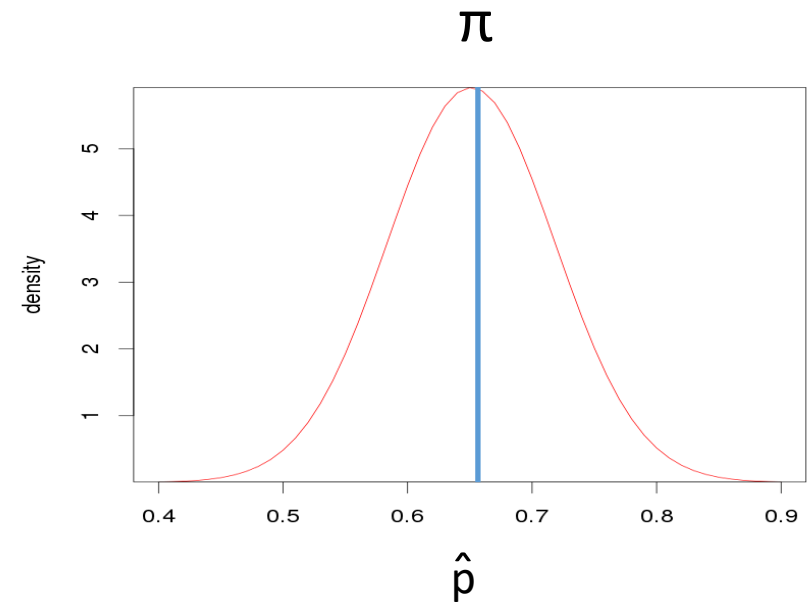

Parametric inference on proportions continued

Summary: Central Limit Theorem for Sample Proportions

We just showed that the sampling distribution of proportions \hat{p} is normal distributed with a standard error of:

$$SE = \sqrt{\frac{\pi(1-\pi)}{n}}$$

$$\hat{P} \sim N\left(\pi, \sqrt{\frac{\pi(1-\pi)}{n}}\right)$$



What would be a problem with using this formula for the SE for inference?

- I.e., what is the problem using this formula for confidence intervals and hypothesis tests?

Standard Error for Sampling Proportions

Note: we don't usually know π , so we can't compute the standard error exactly using the formula:

$$SE = \sqrt{\frac{\pi(1-\pi)}{n}}$$



However, we can substitute \hat{p} for π and then we can get an estimate of the standard error:

$$\hat{SE} = \sqrt{\frac{\hat{p}(1-\hat{p})}{n}}$$



Comparing formula SE to the bootstrap SE

In previous classes we have used the bootstrap to get an estimate of the standard error SE*

How could we do this for the green sprinkles?

```
bootstrap_dist <- do_it(100000) * {  
  boot_sample <- sample(my_sprinkles, replace = TRUE)  
  sum(boot_sample == 'green')/100  
}
```

```
bootstrap_SE <- sd(bootstrap_dist)
```



| Color |
|-------|
| White |
| Red |
| Red |
| White |
| Green |
| White |
| . |
| . |
| . |
| White |
| Green |

n = 100 sprinkles

Comparing formula SE to the bootstrap SE

For my green sprinkles I got:

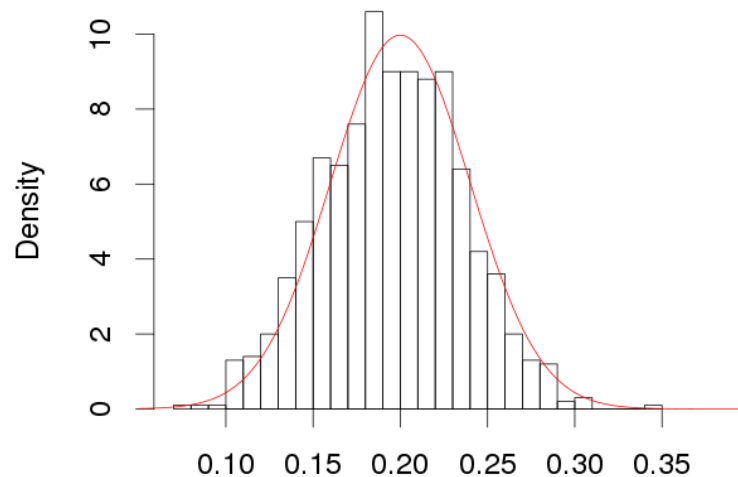
- Bootstrap SE = **0.039959**
- Formula SE = **0.04**

$$\hat{p} = 0.20$$

$$n = 100$$

$$\hat{SE} = \sqrt{\frac{\hat{p}(1-\hat{p})}{n}}$$

Bootstrap Distribution



```
SE <- sqrt( (.2 * (1 - .2) ) / 100 )
```

Parametric confidence intervals for proportions

Confidence intervals for a single proportion

Suppose we have a sample of size n of categorical data

Suppose that n is large enough so that $n\pi \geq 10$ and $n(1 - \pi) \geq 10$

A confidence interval for a population proportion π can be computed from our random sample of size n using:

$$\hat{p} \pm z^* \cdot \sqrt{\frac{\hat{p}(1 - \hat{p})}{n}}$$

Equation for SE



Where \hat{p} is the sample proportion and z^* is a standard normal endpoint to give the desired confidence level

One true love?

A survey asked 2625 people whether they agreed with the statement
“There is only one true love for each person”

1812 of the respondents disagreed

Compute a 90% confidence interval for the proportion who disagreed

$$\hat{p} \pm z^* \cdot \sqrt{\frac{\hat{p}(1-\hat{p})}{n}}$$

Let's try it in R!

One true love?

```
n <- 2625
```

```
p_hat <- 1812/2625 = .69
```

```
SE <- sqrt((p_hat * (1 - p_hat))/n)
```

```
z_star <- qnorm(.95, 0, 1) = 1.64
```

```
ME <- z_star * SE = .032
```

```
CI <- c(p_hat - ME, p_hat + ME) = [.658 .723]
```

$$\hat{p} \pm z^* \cdot \sqrt{\frac{\hat{p}(1-\hat{p})}{n}}$$



Parametric hypothesis tests for proportions

Test for single proportions

To compute p-values when the null distribution is normal we use:

$$z = \frac{\text{Sample Statistic} - \text{Null Parameter}}{SE}$$

In the context of proportions our null hypothesis is of the form $H_0: \pi = \pi_0$

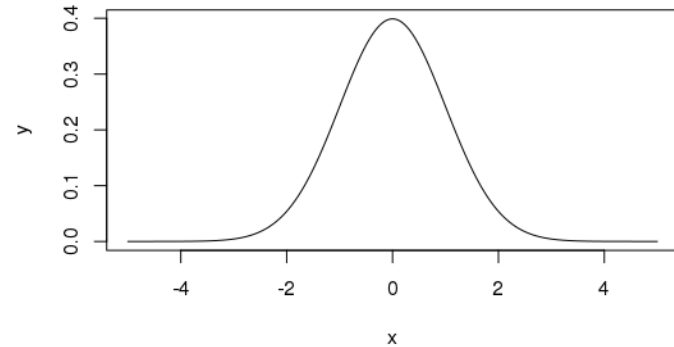
Our formula for z then becomes:

$$z = \frac{\hat{p} - \pi_0}{SE} \qquad SE = \sqrt{\frac{\pi_0(1 - \pi_0)}{n}}$$

Test for single proportions

To test for $H_0: \pi = \pi_0$ vs $H_A: \pi \neq \pi_0$ (or the one-tail alternative), we use the standardized test statistic:

$$z = \frac{\hat{p} - \pi_0}{\sqrt{\frac{\pi_0(1 - \pi_0)}{n}}}$$



Where \hat{p} is the proportion in a random sample of size n

Provided the sample size is reasonable large (usual conditions), the p-value of the test is computed using the standard normal distribution

Do more than 25% of US adults believe in ghosts?

A telephone survey of 1000 randomly selected US adults found that 31% of them say they believe in ghosts. Does this provide evidence that more than 1 in 4 US adults believe in ghosts?

1. State the null and alternative hypothesis
2. Calculate the statistic of interest
- 3-4. Calculate the p-value
Hint: `pnorm()` function will be useful
5. What do you conclude?

$$z = \frac{\hat{p} - \pi_0}{\sqrt{\frac{\pi_0(1 - \pi_0)}{n}}}$$

Do more than 25% of US adults believe in ghosts?

Step 1:


$$H_0: \pi = .25$$

$$H_A: \pi > .25$$

Step 2:

$$\hat{p} = .31$$

$$n = 1000$$

$$z = \frac{\hat{p} - \pi_0}{\sqrt{\frac{\pi_0(1 - \pi_0)}{n}}}$$


```
SE <- sqrt( (.25 * (1 - .25))/1000)
```

```
z_val <- (.31 - .25)/SE
```

z_val is 4.38

Do more than 25% of US adults believe in ghosts?

Step 1:

$$H_0: \pi = .25$$

$$H_A: \pi > .25$$

Step 2:

$$z_val \leftarrow 4.38$$

Step 3-4:

$$p\text{-value} = \text{pnorm}(z_val, 0, 1, \text{lower.tail} = \text{FALSE})$$

Step 5:

Indeed, very strong evidence!