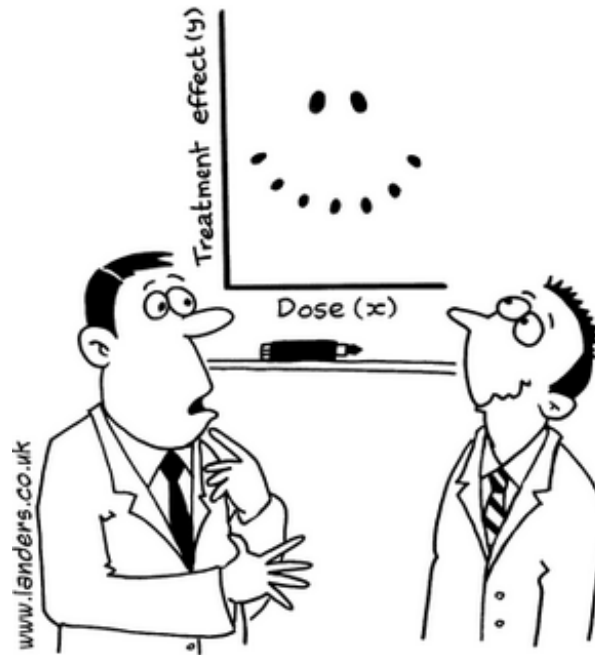


Inference for linear regression



"It's a non-linear pattern with outliers.....but for some reason I'm very happy with the data."

Overview

Quick review of regression models

Inference on regression models

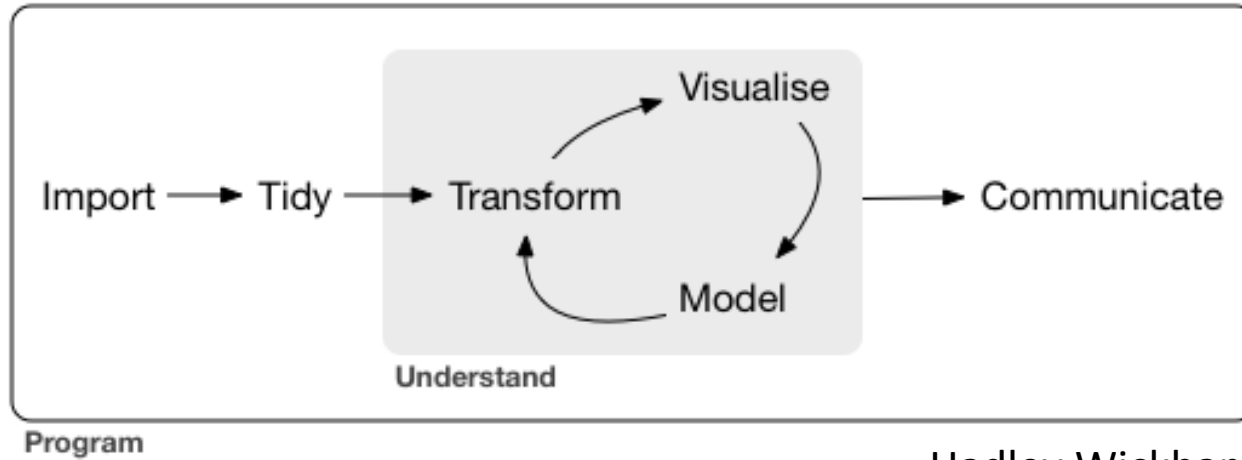
- Hypothesis tests on regression coefficients
- Confidence intervals and predictions intervals

Regression diagnostics

If there is time: statistics for identifying unusual observations

Linear regression continued...

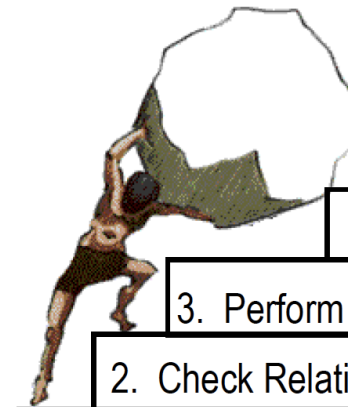
The process of building regression models



Hadley Wickham



Sisyphus' Five Steps for Simple Linear Regression



1. Identify Variables : response and predictor

2. Check Relationships (plots) : make transformations

3. Perform Regression

4. Identify Significant Predictors

5. Check Model Assumptions

Jonathan Reuning-Scherer

"All models are wrong, **but some are useful**"
- George Box

The process of building regression models

Choose the form of the model

- Identify the response variable (y) and explanatory variables (x 's)
- For exploratory analyses, graphical displays can help suggest the model form

Fit the model to the data

- Estimate model parameters, usually using least squares (minimize the SSRs)

Assess how well the model describes the data

- Analyze the residuals, compare to other models, etc.
- If model doesn't fit well, go to step 1.
 - This is as much an art as a science

Use the model to address questions of interest

- 1. Make predictions
- 2. Explore relationships between response variable (y) and explanatory variables (x)
- Keep in mind limitations of the model
 - e.g., can be difficult/impossible to make the claim that changes in x *cause* changes in y from *observational data*

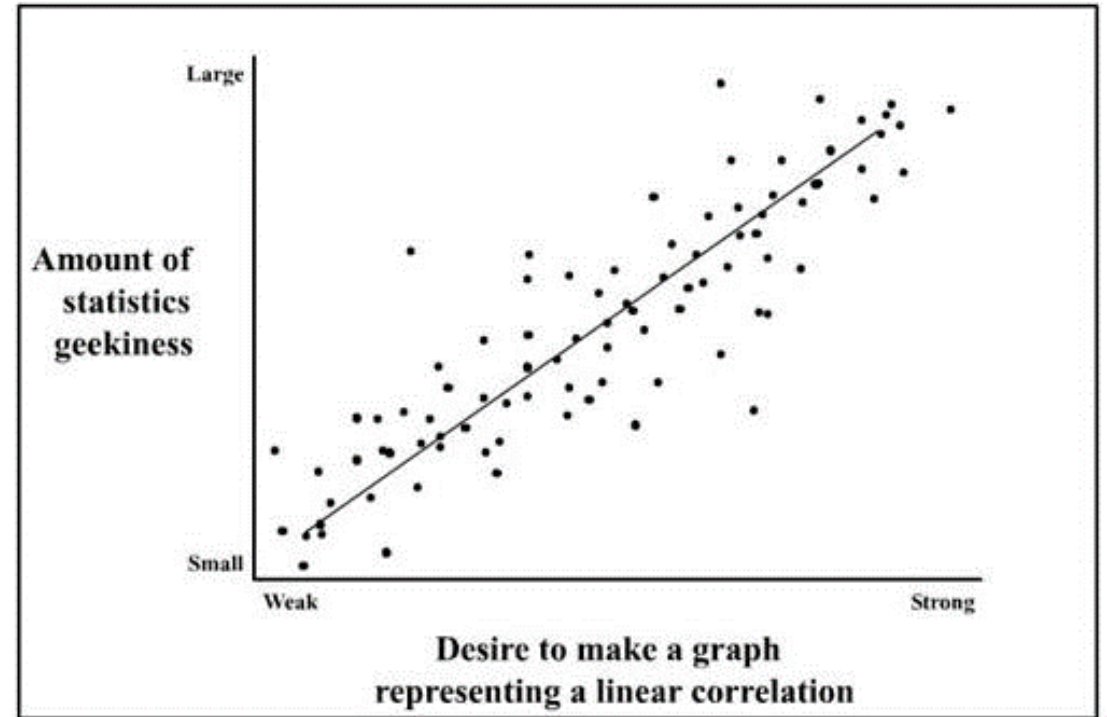


Review of underlying models and inference

Review: Linear regression

In **linear regression** we fit a regression line to the predict a variable y , from other variables x

- e.g., $\hat{y} = b_0 + b_1 \cdot x$



Review: Linear regression underlying model

True regression line: $\mu_Y = \beta_0 + \beta_1 x$

Intercept Slope } Parameters

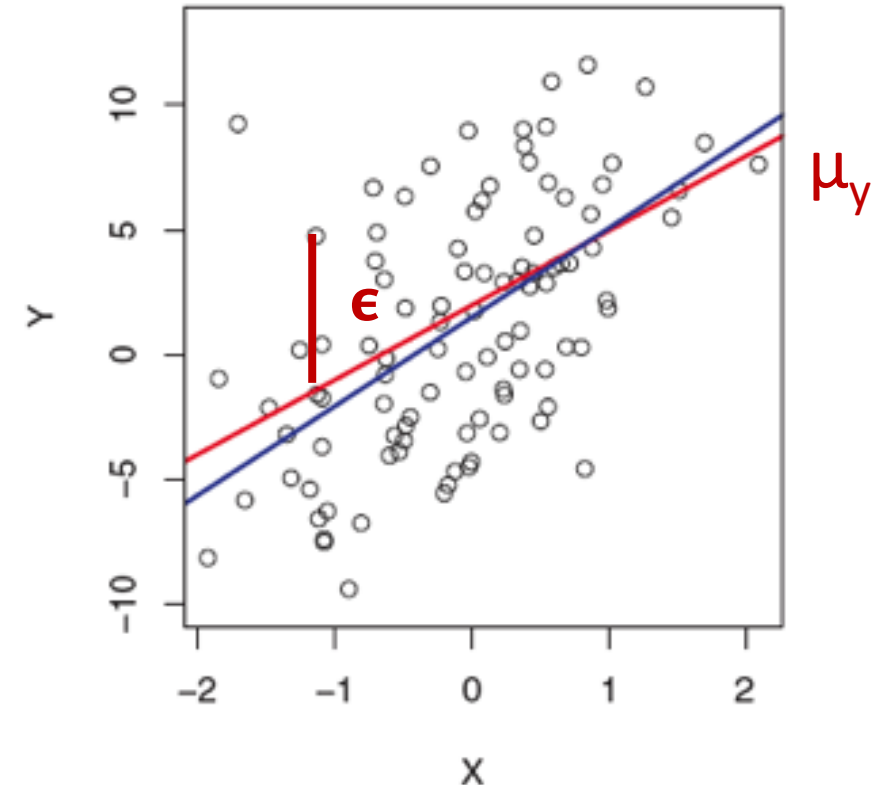
Observed data point: $Y = \beta_0 + \beta_1 x + \epsilon$

Error

$$= \mu_Y + \epsilon$$

Errors ϵ are the difference between the **true regression line** μ_y and observed data points Y

$$\epsilon = Y - \mu_y$$



Review: Linear regression underlying model

True regression line: $\mu_Y = \beta_0 + \beta_1 x$

Intercept Slope } Parameters

Observed data point: $Y = \beta_0 + \beta_1 x + \epsilon$

Error

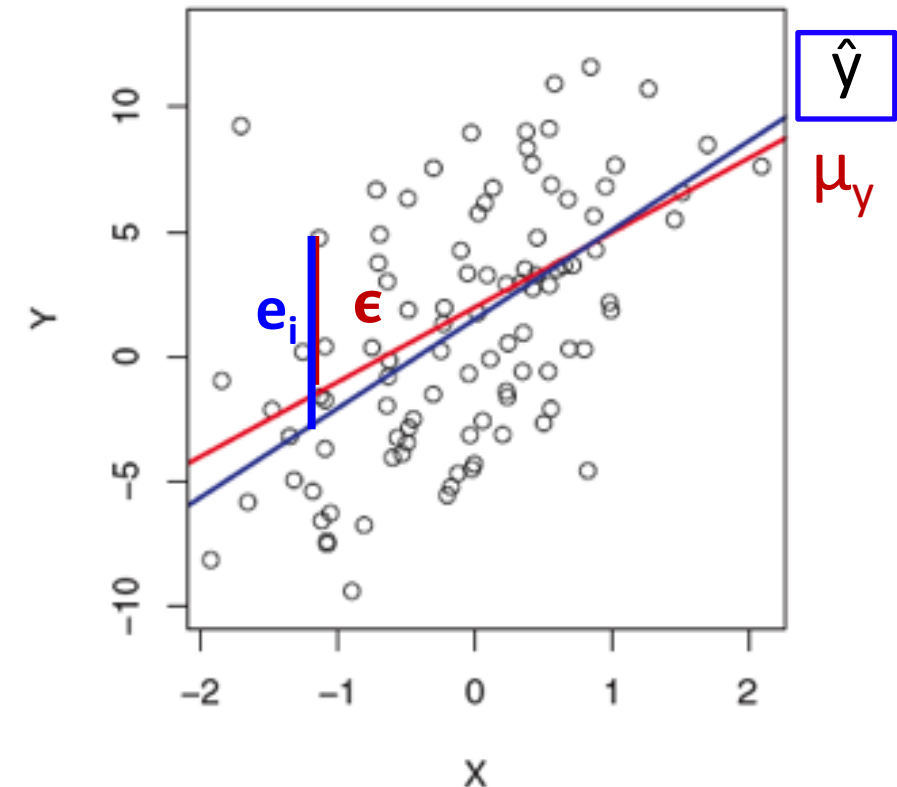
Estimated regression line: $\hat{y} = \hat{\beta}_0 + \hat{\beta}_1 x$

Errors ϵ are the difference between the **true regression line** μ_y and observed data points Y

$$\epsilon = Y - \mu_y$$

Residuals e_i are the difference between the **estimated regression line** \hat{y} and observed data points Y

$$e_i = Y - \hat{y}$$



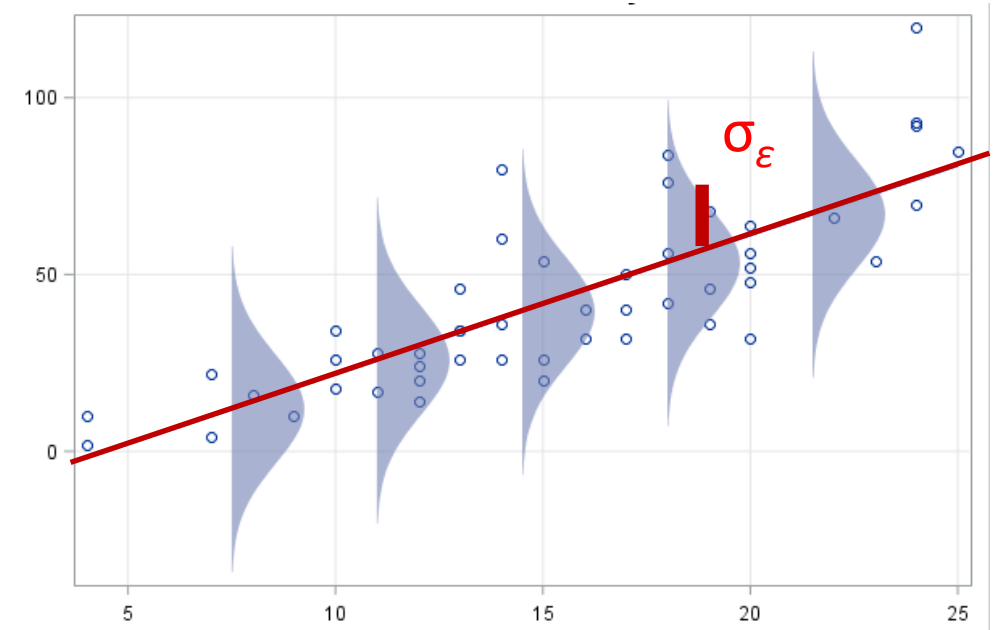
Review: Standard deviation of the errors: σ_ε

The standard deviation of the errors is denoted σ_ε

We can use the **standard deviation of residuals** as an estimate standard deviation of the errors σ_ε .

- $\hat{\sigma}_e$ is often called the "~~residual standard error~~"
- $\hat{\sigma}_e$ we called it the "residual standard deviation"

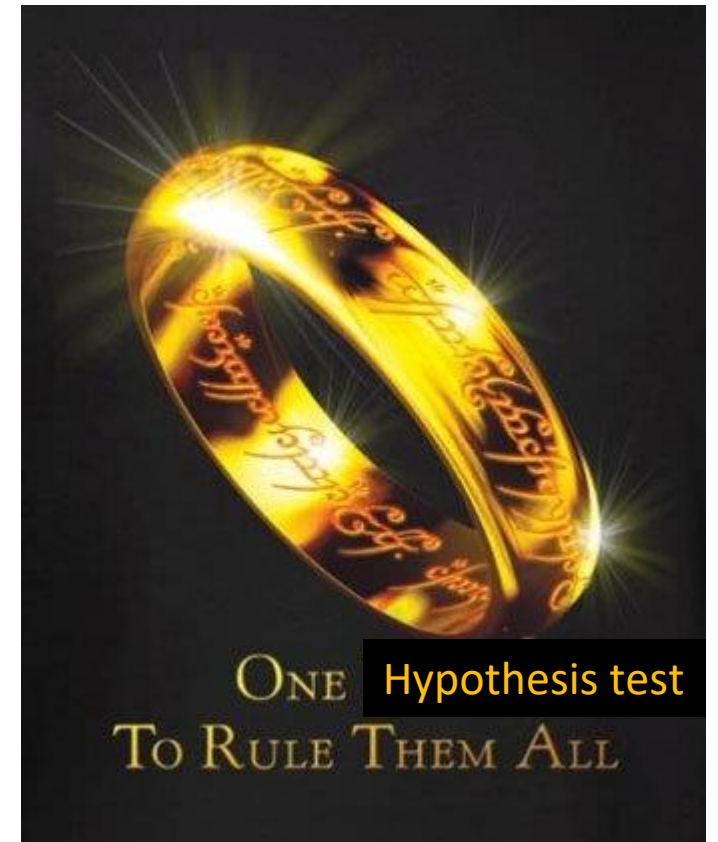
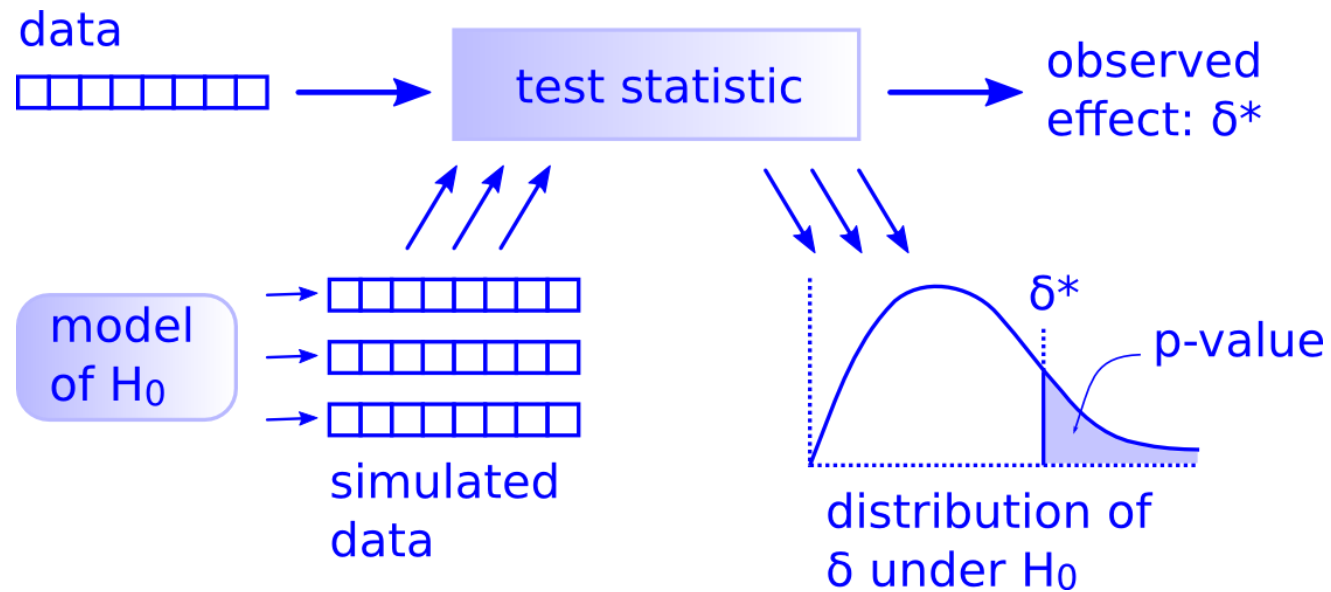
$$\begin{aligned}\hat{\sigma}_e &= \sqrt{\frac{1}{n-2} SSRes} \\ &= \sqrt{\frac{1}{n-2} \sum_{i=1}^n (y_i - \hat{y}_i)^2}\end{aligned}$$



Inference for linear regression: hypothesis tests

Hypothesis test for regression coefficients

There is only one [hypothesis test](#)!



Hypothesis test for regression coefficients

We can run hypothesis tests to assess whether there is a linear relationship between y and x, and calculate p-values

- $H_0: \beta_1 = 0$ (slope is 0, so no linear relationship between x and y)
- $H_A: \beta_1 \neq 0$

One type of hypothesis test we can run is based on a t-statistic: $t = \frac{\hat{\beta}_1 - 0}{\hat{SE}_{\hat{\beta}_1}}$

- The t-statistic comes from a t-distribution with $n - 2$ degrees of freedom

$$\hat{SE}_{\hat{\beta}_1} = \frac{\hat{\sigma}_e}{\sqrt{\sum_{i=1}^n (x_i - \bar{x})^2}}$$

$$\hat{SE}_{\hat{\beta}_0} = \hat{\sigma}_e \sqrt{\frac{1}{n} + \frac{\bar{x}^2}{\sum_{i=1}^n (x_i - \bar{x})^2}}$$

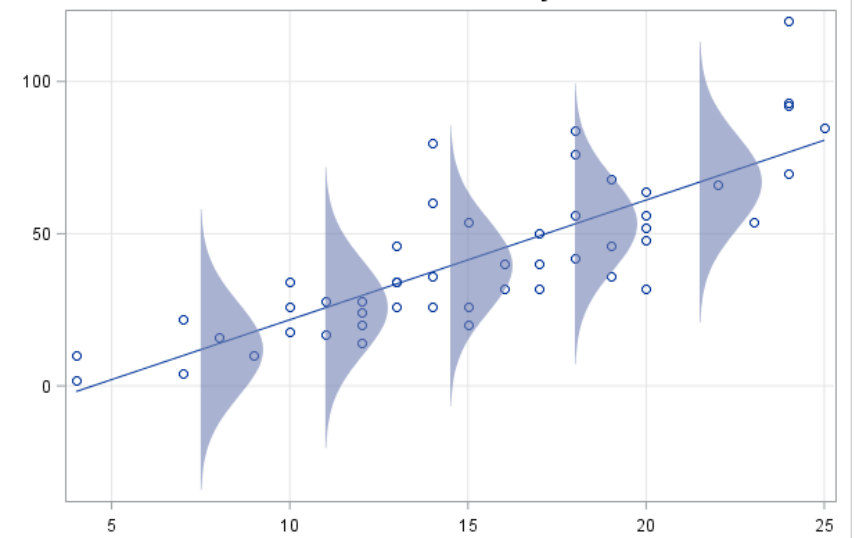
Inference using parametric methods

When using parametric methods, we make the following (LINE) assumptions:

- **Linearity**: A line can describe the relationship between x and y
- **Independence**: each data point is independent from the other points
- **Normality**: errors are normally distributed
- **Equal variance (homoscedasticity)**: constant variance of errors over the whole range of x values

$$Y_i = \beta_0 + \beta_1 x_i + \epsilon_i$$

$$\epsilon_i \sim N(0, \sigma_\epsilon)$$



These assumptions are usually checked after the models are fit using ‘regression diagnostic’ plots.

Let's look at inference for simple linear regression in R

Back to faculty salaries...



Inference for linear regression: confidence intervals

Inference for linear regression: confidence intervals

We can estimate three types of intervals for a regression:

1. Confidence intervals for the regression coefficients: β_0 and β_1
2. Confidence intervals for the full line $\mu_Y(x)$
3. Prediction intervals where most of the data is expected

Confidence intervals for regression coefficients

The confidence interval for the slope coefficient: $\hat{\beta}_1 \pm t^* \cdot \hat{SE}_{\hat{\beta}_1}$

Where: $\hat{SE}_{\hat{\beta}_1} = \frac{\hat{\sigma}_e}{\sqrt{\sum_{i=1}^n (x_i - \bar{x})^2}}$

β_1

t^* is the critical value for the t_{n-2} density curve needed to obtain a desired confidence level

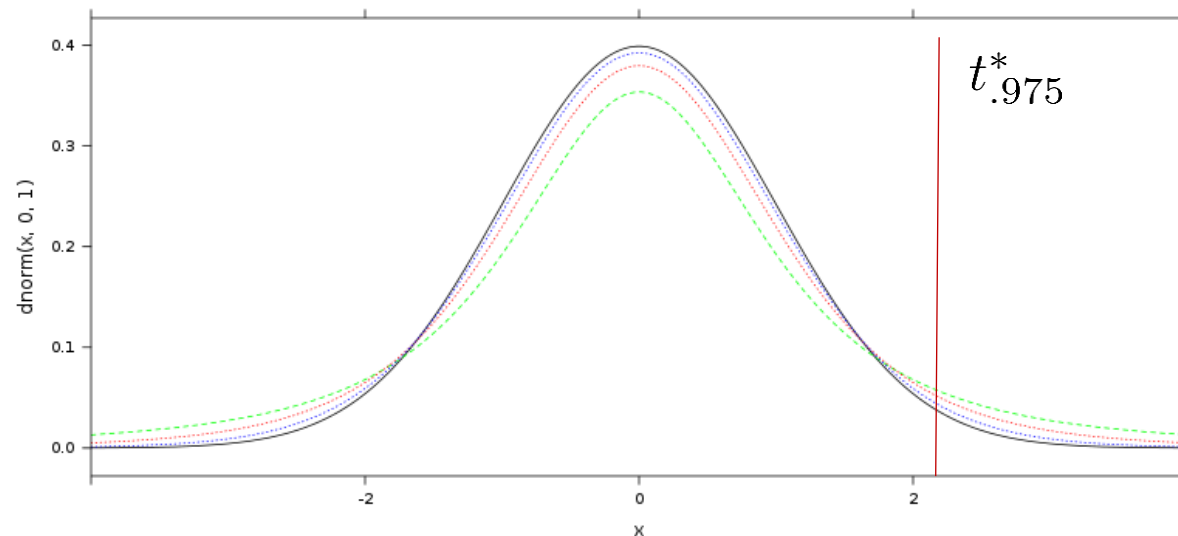
`qt(.975, df)`

N(0, 1)

df = 2

df = 5

df = 15

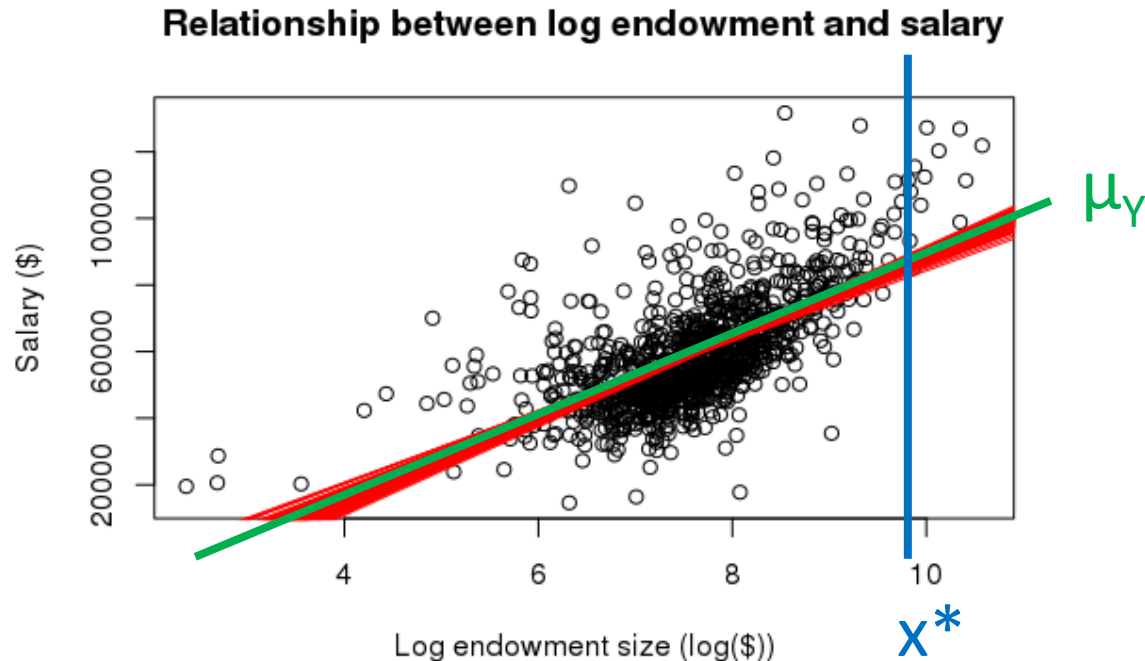


Confidence intervals for the regression line μ_Y

A confidence interval for the mean response for the **true regression line** μ_Y at the value of x^* :

$$\hat{y} \pm t^* \cdot \hat{SE}_{\hat{\mu}} \quad \text{where}$$

$$\hat{SE}_{\hat{\mu}} = \hat{\sigma}_e \sqrt{\frac{1}{n} + \frac{(x^* - \bar{x})^2}{\sum_{i=1}^n (x_i - \bar{x})^2}}$$



Note:

- There is more uncertainty at the ends of the regression line
- The confidence interval for the regression line μ_Y is different than the confidence interval for slope β_1

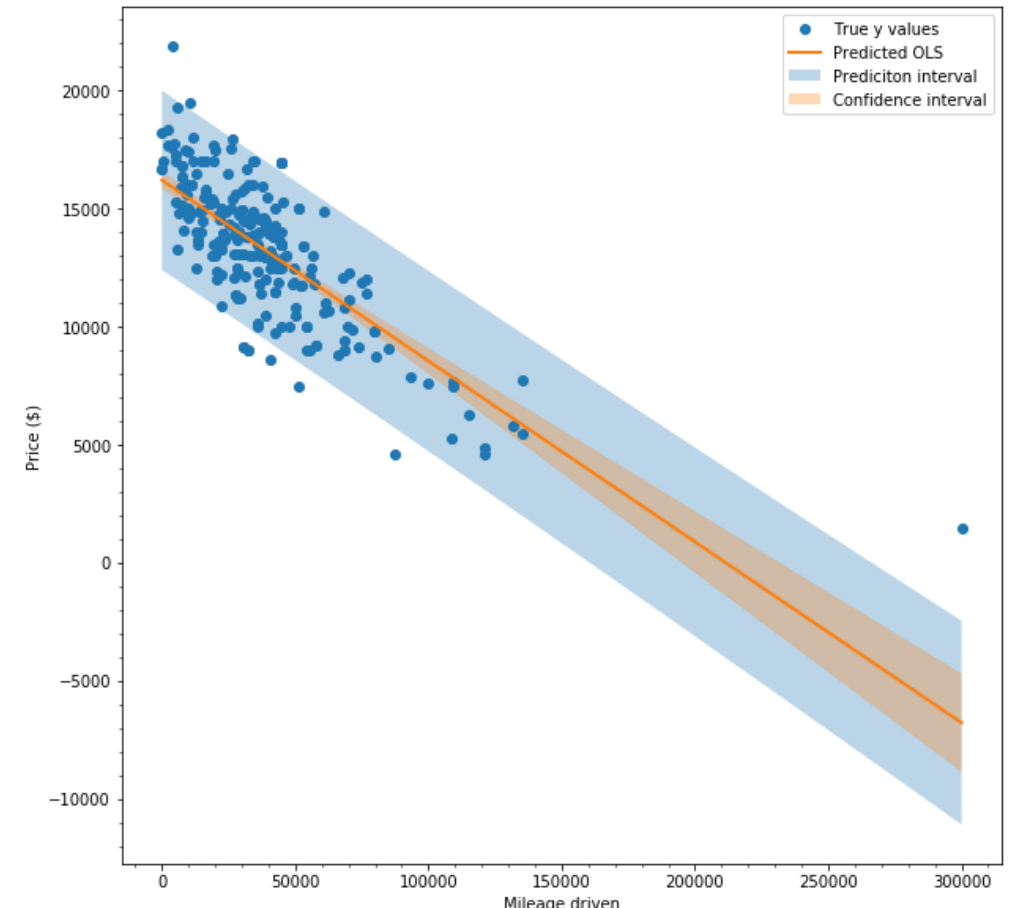
Prediction intervals

Confidence intervals give us a measure of uncertain about our the true relationship between x and y for:

- The true regression slope β_1
- The true regression line μ_y

Prediction intervals give us a range of plausible values for y

- i.e., 95% of our y 's with be within this range



Prediction intervals

A **prediction intervals** for the y can be calculated using:

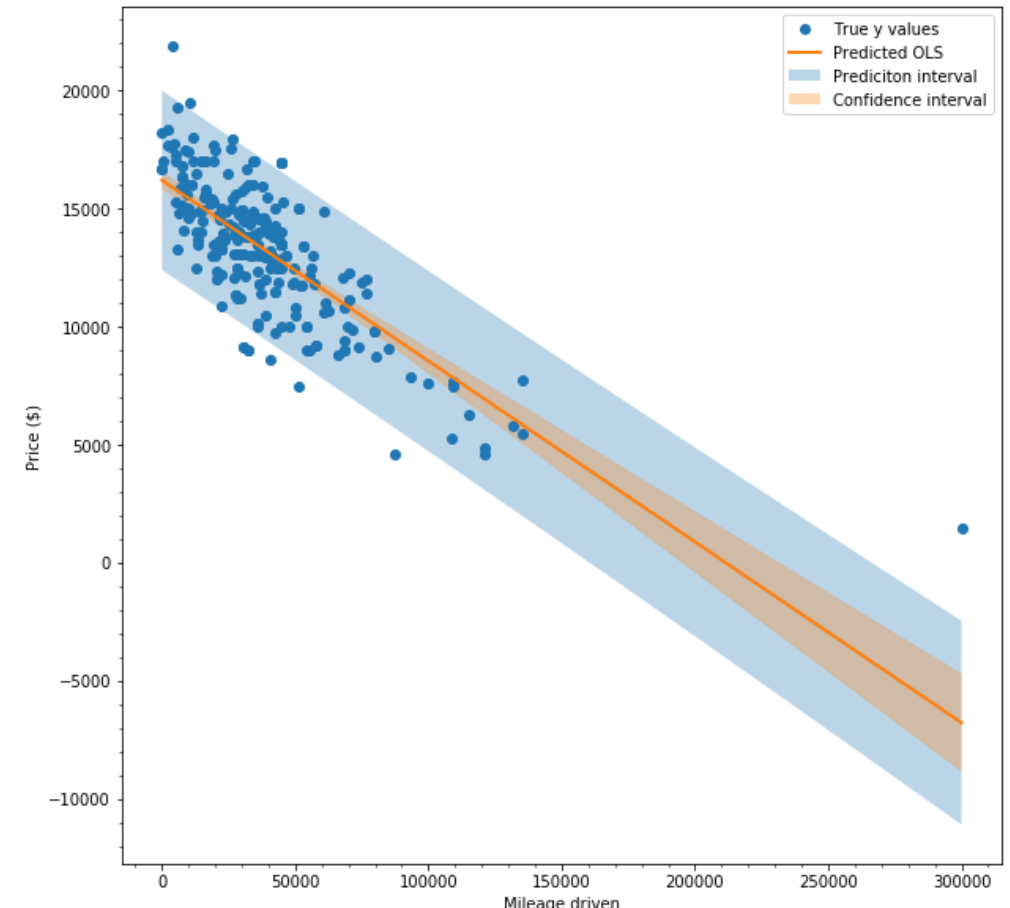
$$\hat{y} \pm t^* \cdot \hat{S}E_{\hat{y}}$$

where

$$\hat{S}E_{\hat{y}} = \hat{\sigma}_e \sqrt{1 + \frac{1}{n} + \frac{(x^* - \bar{x})^2}{\sum_{i=1}^n (x_i - \bar{x})^2}}$$

Due to y 's scattering
around the true
regression line

Due to uncertainty
in where the true
regression line is



Summary of confidence and prediction intervals

1. CI for Slope β

$$\hat{\beta}_1 \pm t^* \cdot \hat{SE}_{\hat{\beta}_1} \quad \hat{SE}_{\hat{\beta}_1} = \hat{\sigma}_e \sqrt{\frac{1}{\sum_{i=1}^n (x_i - \bar{x})^2}}$$

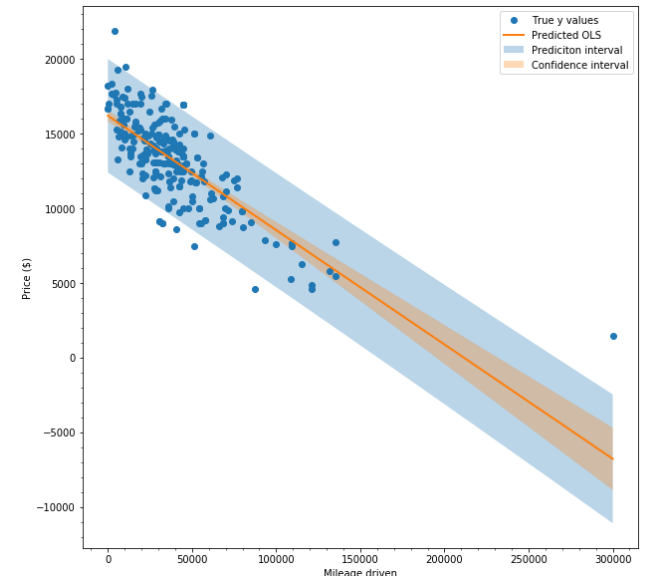


2. CI for regression line μ_y at point x^*

$$\hat{y} \pm t^* \cdot \hat{SE}_{\hat{\mu}} \quad \hat{SE}_{\hat{\mu}} = \hat{\sigma}_e \sqrt{\frac{1}{n} + \frac{(x^* - \bar{x})^2}{\sum_{i=1}^n (x_i - \bar{x})^2}}$$

3. Prediction interval y

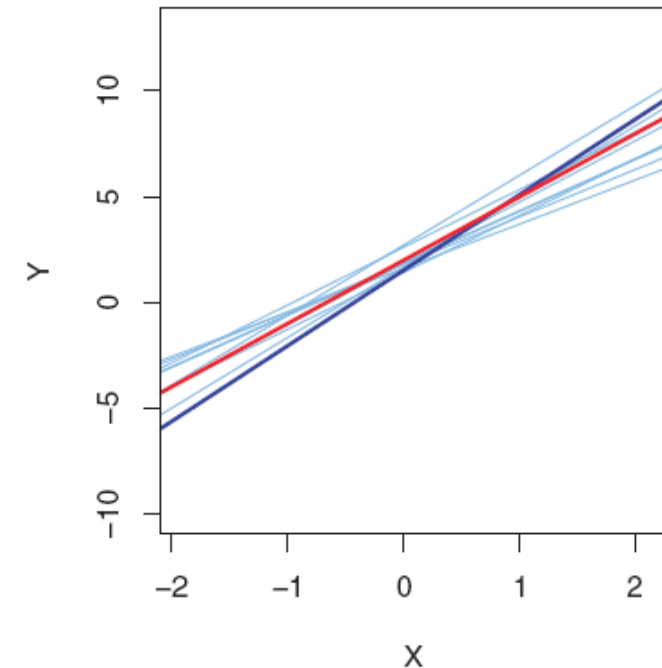
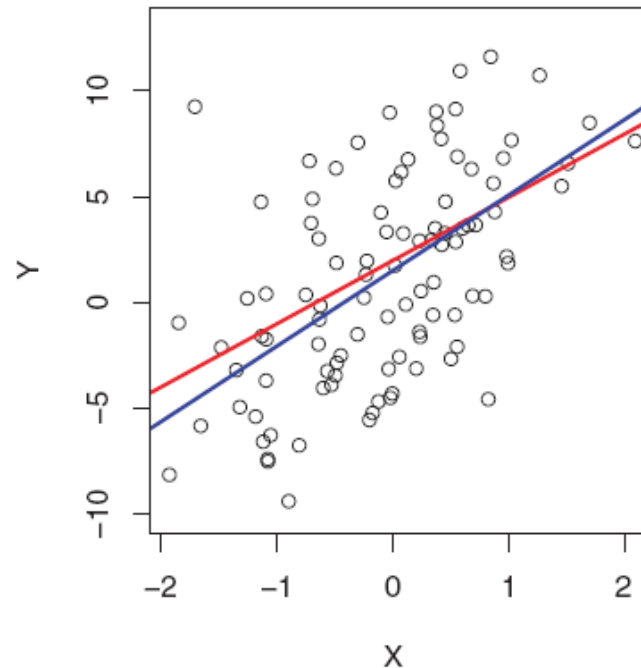
$$\hat{y} \pm t^* \cdot \hat{SE}_{\hat{y}} \quad \hat{SE}_{\hat{y}} = \hat{\sigma}_e \sqrt{1 + \frac{1}{n} + \frac{(x^* - \bar{x})^2}{\sum_{i=1}^n (x_i - \bar{x})^2}}$$



Resampling methods for inference in regression

We can also use resampling methods to estimate run hypothesis tests and create confidence intervals for the regression coefficients

- Bootstrap
- Permutation test



Let's look at inference for simple linear regression in R

Regression diagnostics



Regression diagnostics

We use diagnostics to see if the assumptions/conditions for inference are met

- If they aren't met, we can adjust the model and try again

Choose
Fit
Assess
Use



Regression diagnostics

Let's go through the 4 conditions that should be met when using parametric methods for inference:

- **Linearity**: A line can describe the relationship between x and y
- **Independence**: each data point is independent from the other points
- **Normality**: errors are normally distributed
- **Equal variance (homoscedasticity)**: constant variance of errors over the whole range of x values

Regression diagnostics

Let's go through the 4 conditions that should be met when using parametric methods for inference:

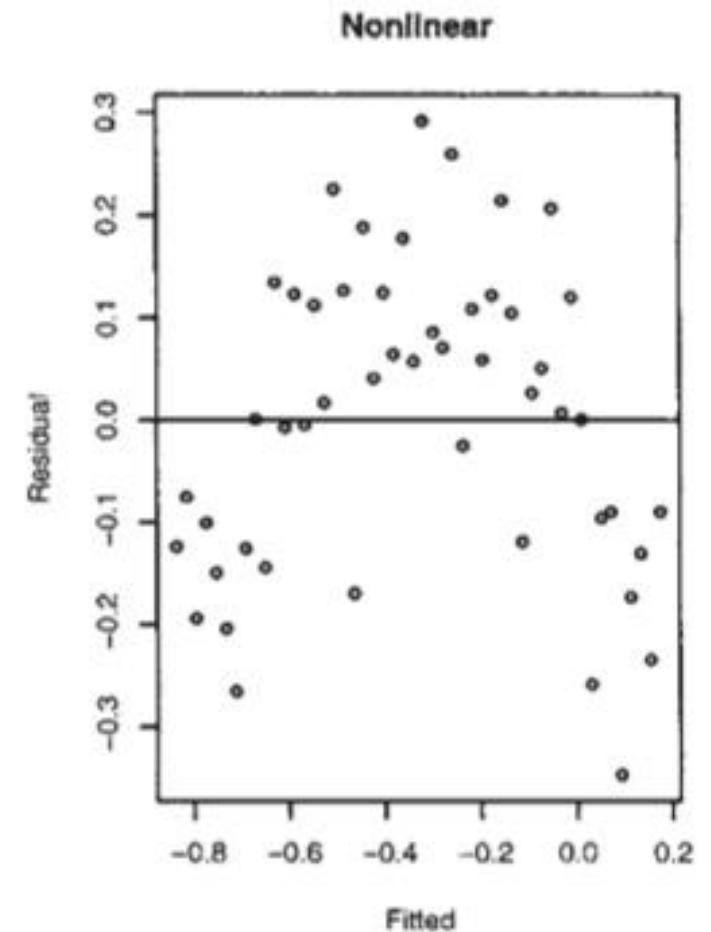
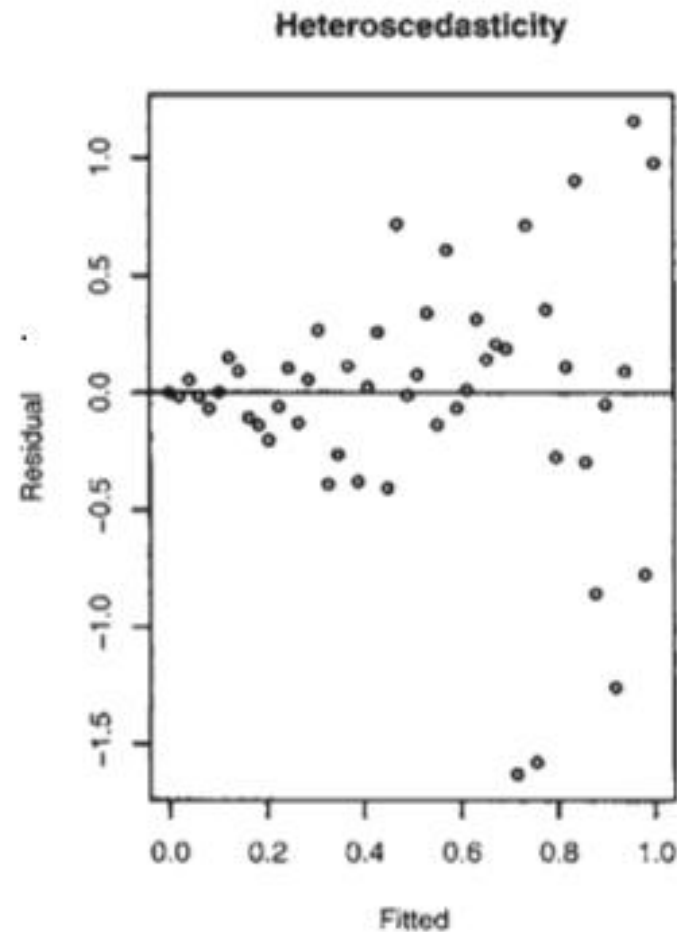
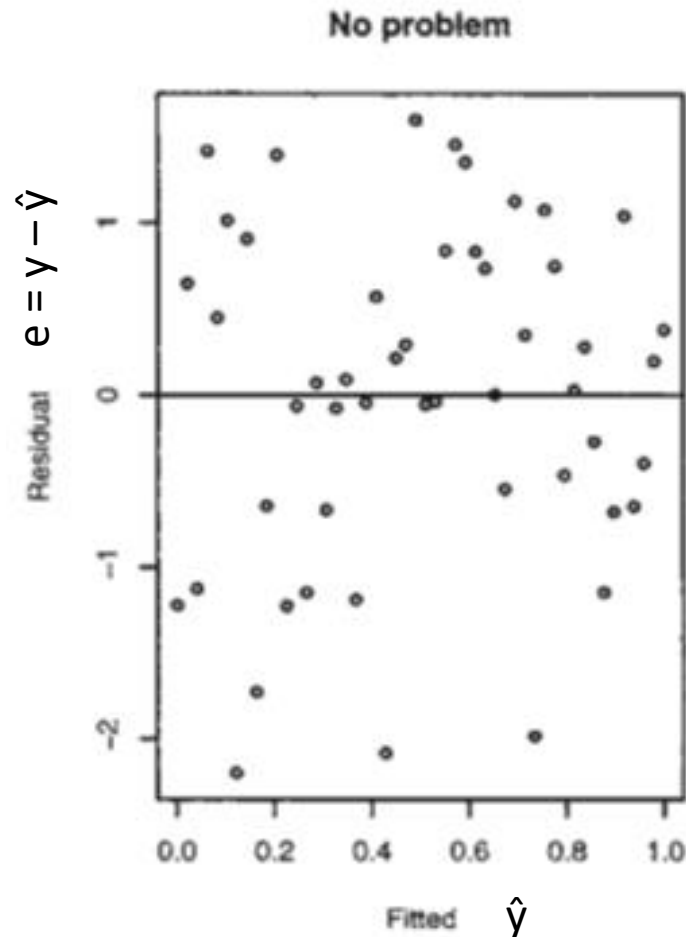
- **Linearity**: A line can describe the relationship between x and y
- **Independence**: each data point is independent from the other points
- **Normality**: errors are normally distributed
- **Equal variance (homoscedasticity)**: constant variance of errors over the whole range of x values

We can check linearity and homoscedasticity by plotting the residuals as a function of the fitted values

Checking linearity and homoscedasticity

~~E~~qual variance

~~L~~inearity

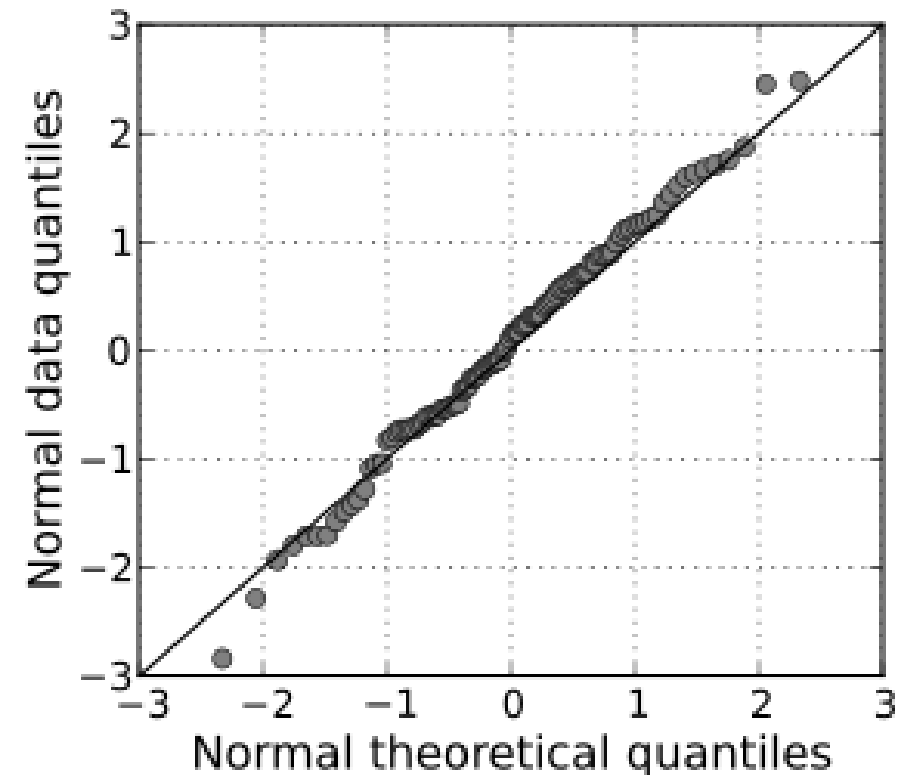


Checking normality

Normality: residuals are normally distributed around the predicted value \hat{y}

We can check this using a Q-Q plot

The 'car' package has a nice function for making qqplots called `qqPlot()`



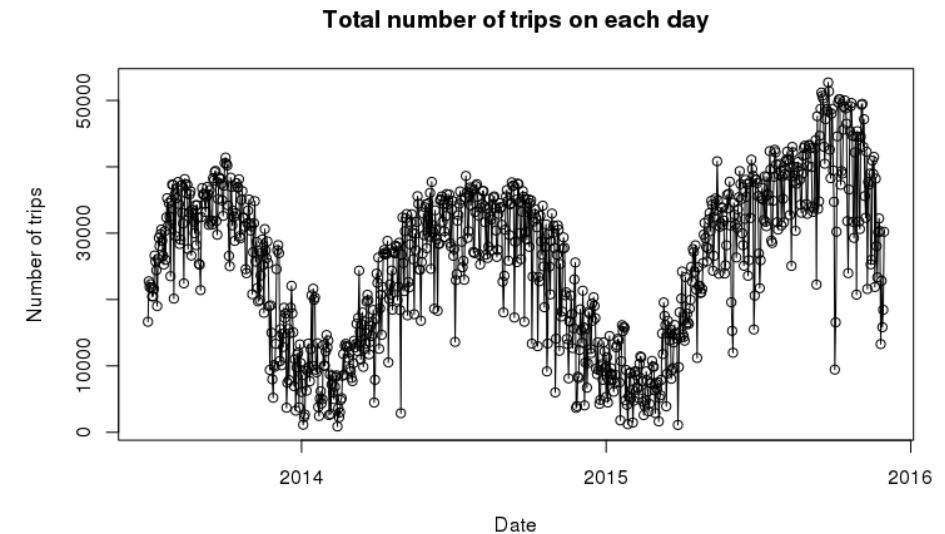
Checking Independence

To check whether each data point is independent requires knowledge of how the data was collected

- Simple random sample from the population is likely independent
- Time series often are not independent

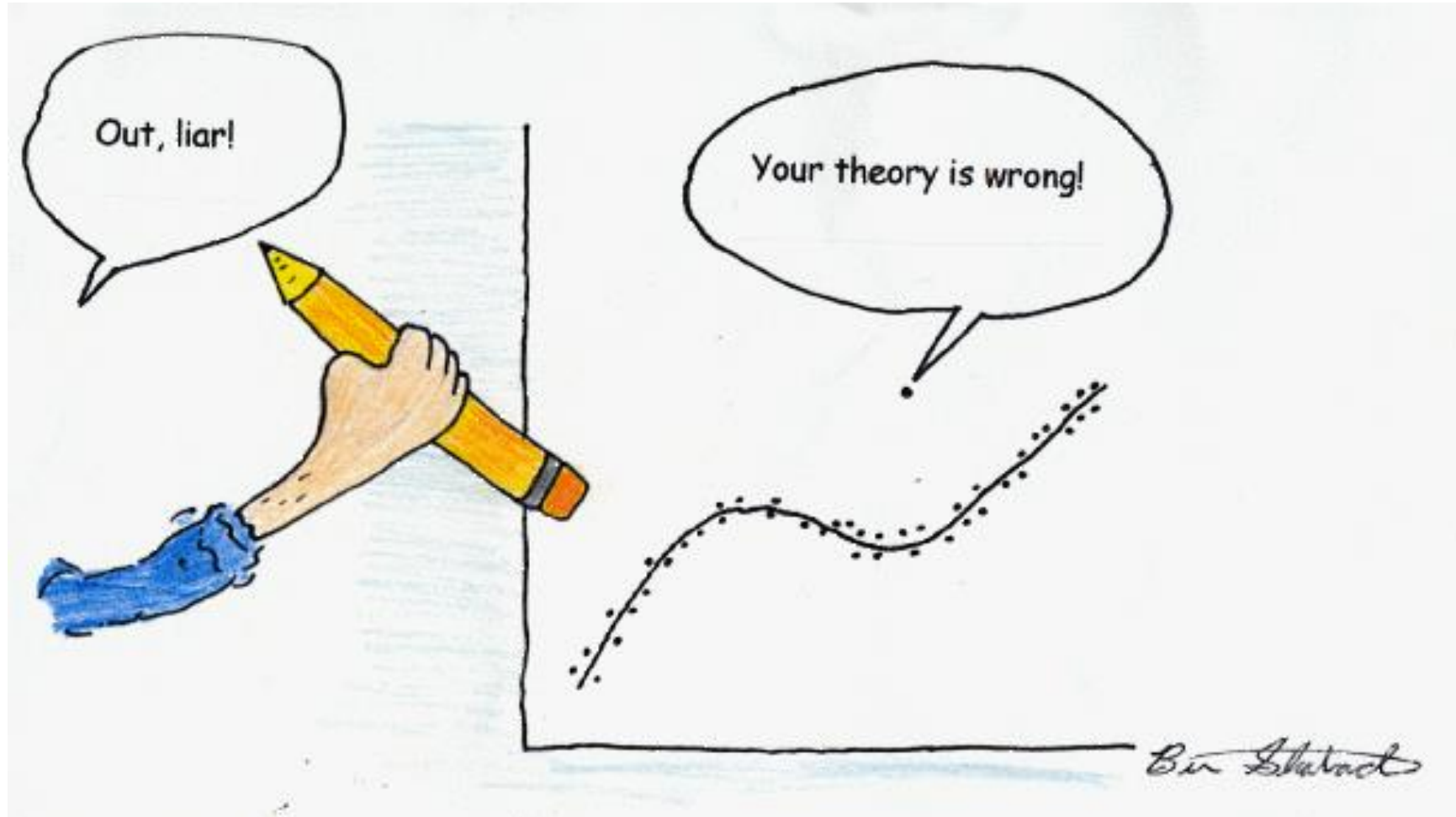
We have basically been assuming independence for everything we have done in this class

- i.i.d. independent and identically distributed



Let's examine these diagnostic plots in R!

Statistics for unusual observations



Statistics for unusual observations

There are statistics that are useful for flagging unusual observations

- **Outliers (large residuals):** unusual y values
- **High leverage points:** unusual x values
- **Influential points:** both an outlier and a high leverage

Unusual observations can indicate:

- An error in data processing
- A need to modify the model
- An interesting phenomenon

Unusual observations **can also have a big effect on the model fit**

- E.g., a big effect on $\hat{\beta}_0$ $\hat{\beta}_1$

Leverage

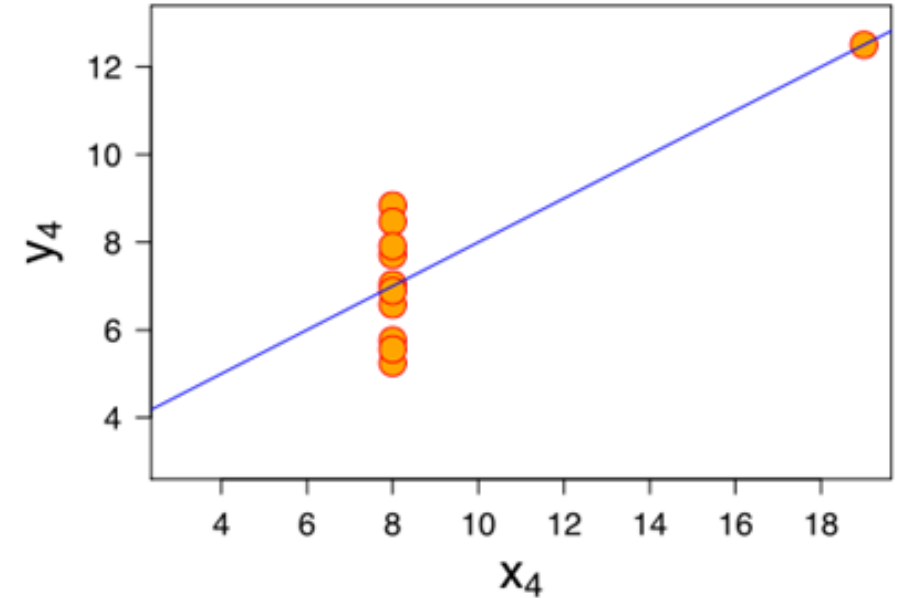
High leverage points are predictors \mathbf{x} that are far from the mean

We can calculate the leverage a data point has using the statistic:

$$h_i = \frac{1}{n} + \frac{(x_i - \bar{x})^2}{\sum_{j=1}^n (x_j - \bar{x})^2}$$

High leverage points can have a big impact on the model that is fit!!!

R: `hatvalues()`



$$\sum_{i=1}^n h_i = 2$$

Typical: $h_i = 2/n$

High: $h_i = 4/n$

Very high: $h_i = 6/n$

Outliers: standardized residuals

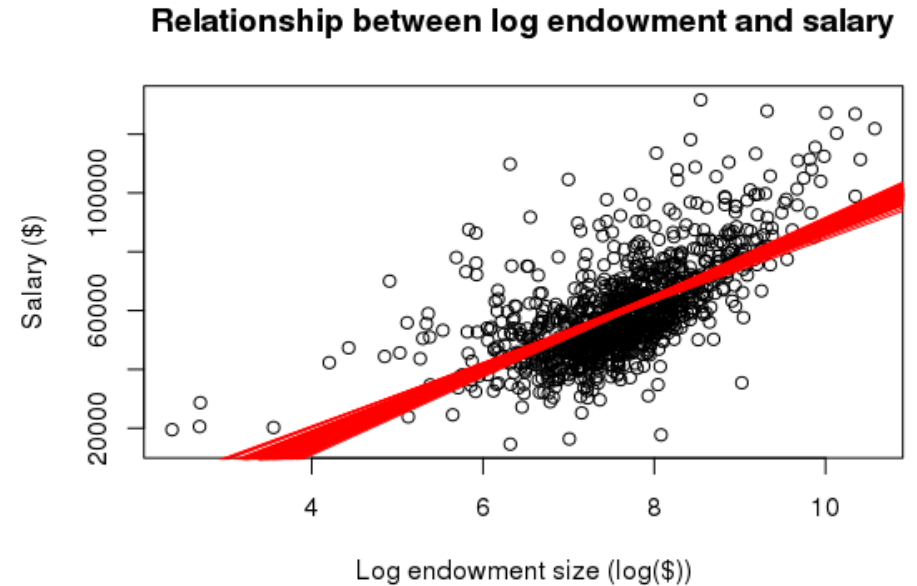
The **standardized residual** for the i^{th} data point in a regression model can be computed using:

$$stdres_i = \frac{y_i - \hat{y}}{\hat{\sigma}_\epsilon \sqrt{1 - h_i}}$$

Puts residuals on a
'normalized' scale

R: `rstandard()`

Makes residuals at the ends a bit larger to deal with the fact that they are 'overfit'



Outliers: studentized residuals

The **studentized residual** for the i^{th} data point in a regression model can be computed using:

$$studres_i = \frac{y_i - \hat{y}}{\hat{\sigma}_{(i)} \sqrt{1 - h_i}}$$

Here $\hat{\sigma}_{(i)}$ is the an estimate of $\hat{\sigma}_{\epsilon}$
with the i^{th} point removed

Q: Why might we want to remove the i^{th} point when calculating $\hat{\sigma}_{\epsilon}$?

A: Outliers could have a big effect on our estimate of $\hat{\sigma}_{\epsilon}$

R: `rstudent ()`


Cook's distance

The amount of influence a point has on a regression line depends on:

- The size of the residual e_i
- The amount of leverage h_i

Cook's distance is a statistic that captures how much influence a point has on a regression line

$$D_i = \frac{(\text{stdres}_i)^2}{k+1} \frac{h_i}{1-h_i}$$



Larger for larger
residuals (outliers)



Larger for high
leverage points

Where k is the number of predictors in the model

R: `cooks.distance ()`

- For simple linear regression $k = 1$ (just a single predictor x)


Cook's distance

The amount of influence a point has on a regression line depends on:

- The size of the residual e_i
- The amount of leverage h_i

Cook's distance is a statistic that captures how much influence a point has on a regression line

$$D_i = \frac{(stdres_i)^2}{k+1} \frac{h_i}{1-h_i}$$



Larger for larger
residuals (outliers)



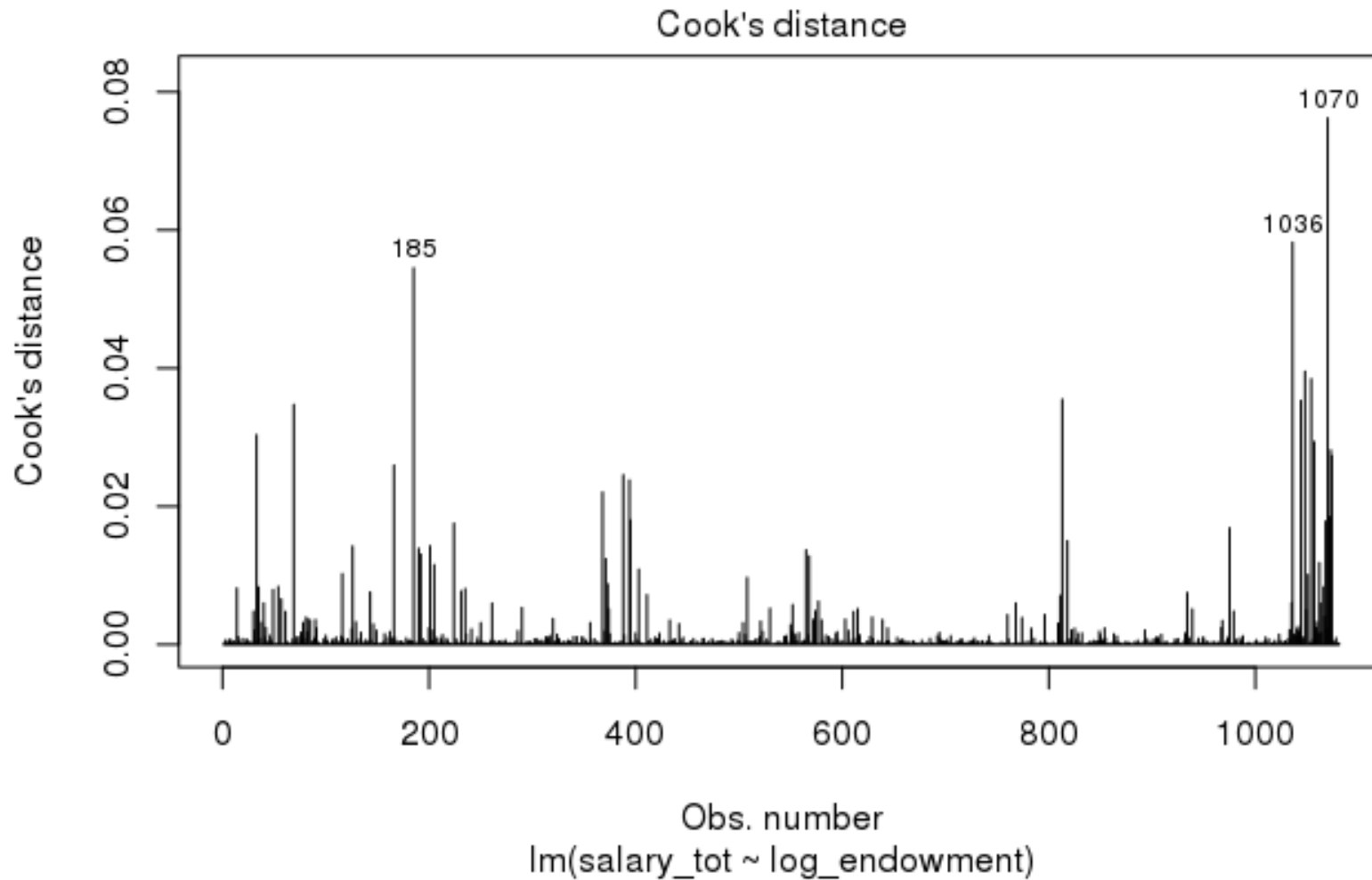
Larger for high
leverage points

Rule of thumb:

- Moderately influential: $D_i > 0.5$
- Very influential: $D_i > 1$

R: `cooks.distance ()`

Cook's distances for $\text{salary} \sim \log_{10}(\text{endowment})$



`plot(lm_fit, 4)`

Unusual points rules of thumb

Statistic	Moderately unusual	Very unusual
Leverage, h_i	Above $2(k + 1)/n$	Above $3(k + 1)/n$
Standardized residual	Beyond ± 2	Beyond ± 3
Studentized residual	Beyond ± 2	Beyond ± 3
Cook's D	Above 0.5	Above 1.0

Where:

- k is the number of explanatory variables
- n is the number of data points

Questions?