

1 Introduction

The dataset we will be using for our group analysis is “Exploring Relationships in Body Dimensions”. This dataset contains information on 21 body dimension measurements (including skeletal measurements and girth measurements) as well as age, weight, height and gender from 507 physically active, young men and women within the normal weight range (a total of 25 variables).

2 Group Analysis

2.1 Variables of the Dataset

Skeletal measurements include: biacromial diameter, pelvic breadth, bitrochanteric diameter, chest depth, chest diameter, elbow diameter, wrist diameter, knee and ankle diameter. Note that elbow, wrist, knee and ankle measurements are the sum of the two ankle diameters. Girth measurements include: shoulder, chest, waist, navel, hip, thigh, flexed bicep, extended forearm, knee, calf maximum, ankle minimum and wrist minimum. Note, for thigh, bicep, forearm, knee, calf, ankle and wrist girth measurements the average of both body parts was taken. Other measurements include: age, weight, height and gender (male= 1 and femal = 0).

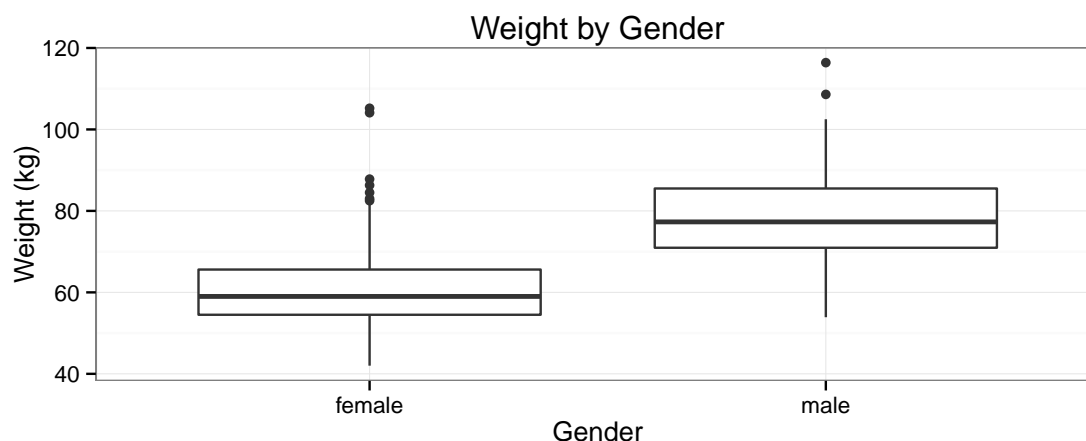
Table 1

Summary table(left) Boxplot(right)					
variable	Min.	Median	Mean	Max.	
weight	42.00	68.20	69.15	116.40	
chest.diam	22.20	27.80	27.97	35.60	
chest.dep	14.30	19.00	19.23	27.50	
bitro.diam	24.70	32.00	31.98	38.00	
wrist.min	13.00	16.10	16.10	19.60	
ankle.min	16.40	22.00	22.16	29.30	
height	147.20	170.30	171.10	198.10	
age	18.00	27.00	30.18	67.00	
shoulder	85.90	108.20	108.20	134.80	
navel	64.00	84.60	85.65	121.10	
hip	78.80	96.00	96.68	128.30	

These are the 11 out of 25 variables selected from the original dataset. Most of these 11 variables are symmetrically distributed except age. The mean of weight, height and age are 69.15kg, 171.10cm and 30.18 yr, respectively, which would suggest that the average subjects are middle and healthy.

2.2 Outcome of Interest: Weight

The initial reason the data was collected was to determine how well weight could be predicted between body build, weight, and girths. Using weight as the response variable, these data can be used to investigate the correspondence between frame size, girths, and weight of physically active young men and women that are within a normal weight range.



2.3 Initial Multiple Linear Regression Model

The initial model we are interested in fitting is of the form:

$$\text{weight}_i = \beta_0 + \beta_1 \text{chest.diam}_i + \beta_2 \text{chest.dep}_i + \beta_3 \text{bitro.diam}_i + \beta_4 \text{wrist.min}_i + \beta_5 \text{ankle.min}_i + \beta_6 \text{height}_i$$

The paper attached with the dataset mentioned multiple models one could use, we chose this as our base model, due to the fact that the model had variables across the entire body, included measurements of depth, girth and diameter and contained a reasonable amount of variables. The variables are: chest diameter (at mid-expiration level), chest depth (between spine and sternum at mid-expiration), bitrochanteric diameter (distance between both trochanters), wrist minimum girth (average of right and left girths), ankle minimum girth (average), and height.

```
## (Intercept) chest.diam chest.dep bitro.diam wrist.min ankle.min
## -109.890      1.340      1.537      1.196      1.113      1.152
## height
## 0.177
```

Our model indicates that the expected weight when all variables are 0 is -110 (which does not make sense in this context). Furthermore the expected change in weight for a 1 unit change in chest.diam, holding all other variables constant, is 1.34 kgs. The expected change in weight for a 1 unit change in chest.dep, holding all other variables constant, is 1.54 kgs., etc. Note that chest depth has the largest impact on weight. In addition to the coefficients, the model has an R-squared value of 0.8882 which implies that our model explains 88.82% of the variation in weight. This model seems like a good tool to predict weight given these measurements, but can we find a better one?

3 Individual Analyses

3.1 Model Diagnostics (Nick)

3.2 Model Selection (Emily)

3.2.1 Initial Model

As stated previously, the initial model we are interested in fitting is of the form:

$$\text{weight}_i = \beta_0 + \beta_1 \text{ chest.diam}_i + \beta_2 \text{ chest.dep}_i + \beta_3 \text{ bitro.diam}_i + \beta_4 \text{ wrist.min}_i + \beta_5 \text{ ankle.min}_i + \beta_6 \text{ height}_i$$

```
## (Intercept) chest.diam chest.dep bitro.diam wrist.min
## -109.890      1.340      1.537      1.196      1.113
## ankle.min    height
##      1.152      0.177
```

This model indicates that the expected change in weight for a 1 unit change in chest.diam, holding all other variables constant, is 1.34 kgs. The expected change in weight for a 1 unit change in chest.dep, holding all other variables constant, is 1.54 kgs., etc. Note that chest depth has the largest impact on weight. In addition to the coefficients, the R-squared value of 0.8882 implies that our model explains 88.82% of the variation in weight and the *P*-values for each variable and for the model are significant. This model seems like a good tool to predict weight given these measurements, but are there better ones?

3.2.2 Model Selection

We are building a model to predict weight given various body measurements. Before running random models, we need to determine what predictors to use. The predictors needed in our models are age, height and gender. These variables contribute significantly to weight. The predictors we will allow in model selection are the initial predictors: chest diameter, chest depth and bitro diameter. In addition to these variables, pelvic breadth, shoulder, chest, waist, hip and thigh will be used. I chose to allow these predictors in my model since these are directly associated with weight (e.g. waist). However, for the other models we will fit, we will let “R” do it’s work.

3.2.3 Criterion

The Information Criteria we will be using to evaluate our models are Akaike Information Criterion (AIC), Bayes Information Criterion (BIC), adjusted R^2 and Predictive Residual Sum of Squares (PreSS). In short, AIC and BIC measure goodness-of-fit through residual sum of squares (log likelihoods) and penalizes the model size; the smaller the AIC/BIC, the better. Adjusted R^2 adjusts R^2 so that the model is penalized for adding more predictors; the higher the value of the adjusted R^2 the better. Finally, PRESS is a summary measure focused on prediction; the lower the value of PRESS, the better.

$$\begin{aligned} \text{AIC} &= n \log \left(\frac{RSS}{n} \right) + 2(p+1) \\ \text{BIC} &= n \log \left(\frac{RSS}{n} \right) + (p+1) \log(n) \\ \text{adj}R^2 &= 1 - \frac{n-1}{n-p-1} (1 - R^2) \end{aligned}$$

$$\text{PRESS} = \sum \left(\frac{\hat{\epsilon}_i}{1 - h_{ii}} \right)^2$$

3.2.4 Methods in R

There are multiple methods built into different packages in R for Model Selection. To illustrate these, we will use the variables: height, wrist.diam, ankle.diam and chest.

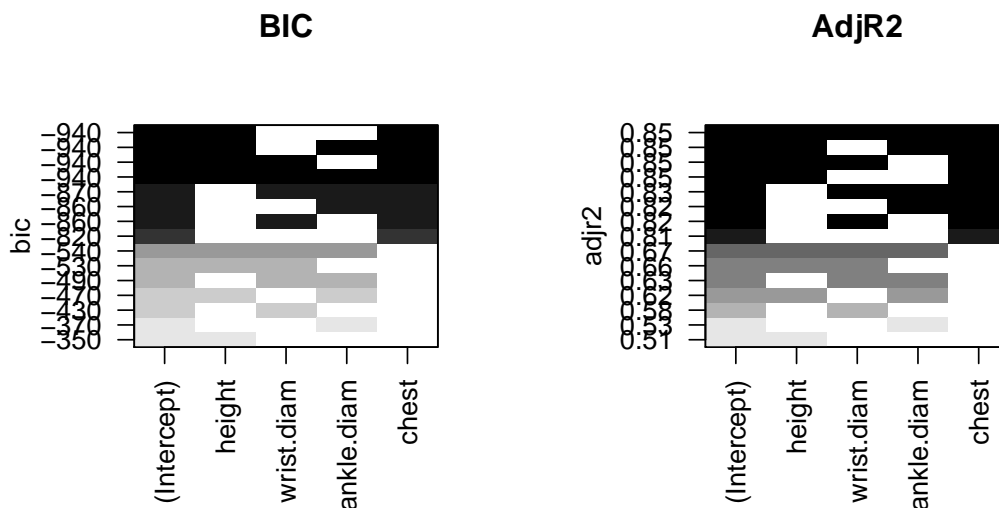
stepAIC()

The R function found in the package “MASS” called “stepAIC()” performs stepwise model selection by AIC. This will output the initial model and the final model (model of best fit determined by this method), and the steps taken. In the output below we can see that this method suggests using a different model that doesn’t contain wrist.diam.

```
Initial Model: weight height + wrist.diam + ankle.diam + chest
Final Model: weight height + ankle.diam + chest
      Step Df Deviance Resid. Df Resid. Dev      AIC
1              502   13294.14 1666.150
2 - wrist.diam    1   47.90427    503   13342.05 1665.974
```

leaps()

The R package “leaps” contains a function “regsubsets()”. This method performs an exhaustive search of models and plots the R^2 criterion by variables and subset size. The class “summary.regsubsets” outputs an object with multiple elements, including adjusted R^2 and BIC. Furthermore, the plots below plot the BIC and Adjusted R^2 values against each subset of variables.



In these plots, for example, the BIC plot is implying the best model is using height and chest as predictors. On the other hand, the AdjR2 plot is saying all variables give the best

fit. Using both of these plots together, one might conclude height, ankle.diam and chest would be the best fit. This conclusion agrees with our analysis using stepAIC.

3.2.5 Model Selection in Action

The Model Selection Lab will explain how to obtain the models summarized in Table 1.

Table 1

MLR#	AIC	BIC	PRESS	Adjusted R^2	Method
all	2216	2326	2384	0.9753	all variables from dataset used
i	2970	3004	10405	0.8869	suggested by paper
1	2256	2319	2560	0.9727	suggested by paper
2	2402	2441	3408	0.9632	my model
3	2206	2282	2329	0.9754	stepAIC
4	2195	2271	2281	0.9759	stepAIC and adjustments
5	2207	2292	2335	0.9755	leaps (adj R^2)
6	2189	2278	2255	0.9764	leaps(adj R^2) and adjustments
7	2213	2272	2353	0.9749	leaps(BIC)
8	2205	2264	2316	0.9753	leaps(BIC) and adjustments

Base on Table 1, we can see that each criteria yields different results. It is up to our discretion to choose a model. Red corresponds to the best value, of all 10 models, for that criteria, blue is the second best. Since AIC and PRESS are lowest in MLR6, the adjusted R^2 is the largest, and the BIC is close to the best and second best values, I would choose MLR6 as the model of best fit. MLR6 is of the form:

$$\text{weight}_i = \beta_0 + \beta_1 \text{pelvic.bredth}_i + \beta_2 \text{bitro.diam}_i + \beta_3 \text{chest.dep}_i + \beta_4 \text{chest.diam}_i + \beta_5 \text{elbow.diam}_i + \beta_6 \text{knee.diam}_i + \beta_7 \text{shoulder}_i + \beta_7 \text{chest}_i + \beta_8 \text{waist}_i + \beta_9 \text{hip}_i + \beta_{10} \text{thigh}_i + \beta_{11} \text{bicep}_i + \beta_{12} \text{forearm}_i + \beta_{13} \text{knee}_i + \beta_{14} \text{calf}_i + \beta_{15} \text{age}_i + \beta_{16} \text{height}_i + \beta_{17} \text{gender}_i + \beta_{18} \text{height}_i^2.$$

Note in Table 1 the first and second lowest values for AIC, BIC and Press, and the first and second highest values for Adjusted R^2 were for stepAIC with adjustments and leaps with adjustments.

3.2.6 Conclusion

Model Selection truly is an art form. R can mechanically run through steps, interactions, combinations, etc. However, R cannot subjectively look at the variables to determine the absolute best model. To achieve the model of best fit, a combination of methods and human adjustment is necessary.

3.3 Logistic regression and Ada-Boosting (Yiding)

3.3.1 Logistic regression

AIC Method: The procedure starts from the full-fit model that include the all variables and then delete the one variable that would decrease the AIC most, and then delete another one.

Step #	Model	AIC
1	$\log i(\text{SEX}) = \text{WT} + \text{CDM} + \text{CDP} + \text{BDM} + \text{WR} + \text{ANK} + \text{HT} + \text{AGE} + \text{SHD} + \text{NAV} + \text{HIP}$	110.35
2	$\log i(\text{SEX}) = \text{WT} + \text{CDP} + \text{BDM} + \text{WR} + \text{ANK} + \text{HT} + \text{AGE} + \text{SHD} + \text{NAV} + \text{HIP}$	108.36
3	$\log i(\text{SEX}) = \text{WT} + \text{CDP} + \text{BDM} + \text{WR} + \text{HT} + \text{AGE} + \text{SHD} + \text{NAV} + \text{HIP}$	106.45
4	$\log i(\text{SEX}) = \text{CDP} + \text{BDM} + \text{WR} + \text{HT} + \text{AGE} + \text{SHD} + \text{NAV} + \text{HIP}$	104.55
5	$\log i(\text{SEX}) = \text{CDP} + \text{WR} + \text{HT} + \text{AGE} + \text{SHD} + \text{NAV} + \text{HIP}$	102.69
6	$\log i(\text{SEX}) = \text{CDP} + \text{WR} + \text{HT} + \text{AGE} + \text{SHD} + \text{HIP}$	101.05

The final model by AIC criteria: $\log i(\text{gender}) = -62.147 + 0.296 \times \text{chest.dep} + 1.307 \times \text{wrist.min} + 0.203 \times \text{height} + 0.054 \times \text{age} + 0.447 \times \text{shoulder} - 0.506 \times \text{hip}$

BIC Method: The method start picking the variables that have highest posterior probabilities individually against null model and then add the secondary highest variables to the model.

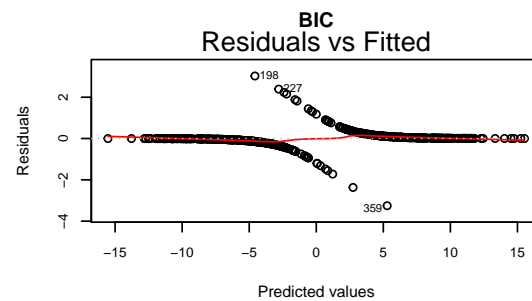
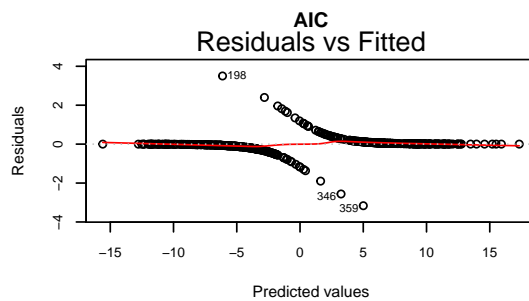
Model #	Model	BIC	Posterior prob
1	$\log i(\text{SEX}) = \text{WR} + \text{HT} + \text{SHD} + \text{HIP}$	-3033	0.453
2	$\log i(\text{SEX}) = \text{WR} + \text{HT} + \text{AGE} + \text{SHD} + \text{HIP}$	-3031	0.157
3	$\log i(\text{SEX}) = \text{CDP} + \text{WR} + \text{HT} + \text{AGE} + \text{SHD} + \text{HIP}$	-3031	0.151
4	$\log i(\text{SEX}) = \text{WR} + \text{HT} + \text{SHD} + \text{NAV} + \text{HIP}$	-3031	0.142
5	$\log i(\text{SEX}) = \text{WR} + \text{HT} + \text{SHD} + \text{HIP}$	-3029	0.047

The model with least BIC and most posterior probability is model 1 and it includes *wrist, height, shoulder, hip*. The final model by BIC criteria: $\log i(\text{gender}) = -61.894 + 1.421 \times \text{wrist} + 0.198 \times \text{height} + 0.461 \times \text{shoulder} - 0.460 \times \text{hip}$

Brief Summary: When the model is small, according to the p-value of each variable, the BIC method perform better. Besides, the final model calculated by AIC includes some implausible variable, such that "age"—meaningless in reality and "chest.dep"—not significant. However, utilize AIC, BIC and p-value to estimate the goodness of a model are not sufficient. What else perspectives should we consider?

Check Goodness of Model

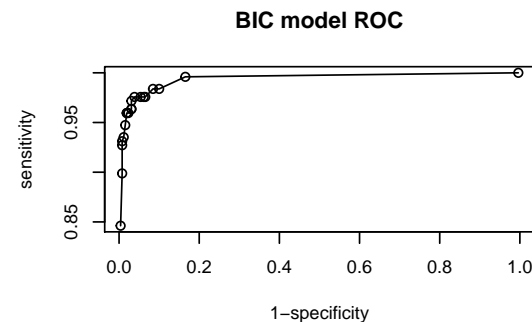
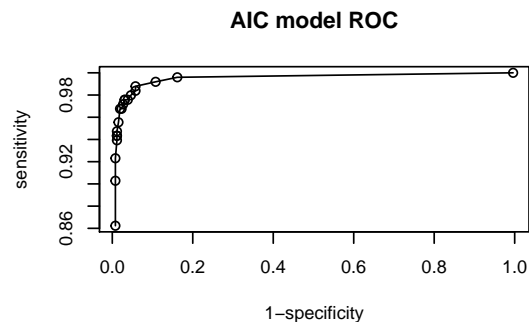
LINEARITY: Checking the residuals can be useful for identifying potential outliers (observations not well fit by the model) as well as misspecified models. The larger the deviance, the poorer the fit.



Both model have a fairly good fit except several subjects in the center of x-axis. These subjects' predicted odds ratio $\simeq 1$, which may be performed as random guess, thus lead higher residuals.

QUANTIFYING PREDICTIVE ABILITY: The Receiver Operating Characteristic (ROC) curve is a very useful method to assess the predictive ability. It is a curve such that plotted based on the sensitivity against 1-specificity of a series cutoff points (such that 0.5, 0.6 ...). The area under the ROC curve can give us insight into the predictive ability of the model. If $c(area) = 0.5$ means model predicting at random and if $c = 1$ indicates a very good predictive ability.

When Somers' D_{xy} rank correlation = 0, that is $D_{xy} = 2(c - 0.5) = 0$, the model is making random prediction; when = 1, the model discriminates outcome perfectly.



a.Check criteria

Model	Area ROC curve	Somers' D_{xy}
AIC	0.9944721	0.9889443
BIC	0.9941607	0.9883214

b.BIC Logistic model predictive ability

Cutoff point	Sensitivity	Specificity	Accuracy
0.00	1.000	0.004	0.489
0.05	0.996	0.835	0.913
...
0.60	0.960	0.981	0.970
...
0.90	0.899	0.992	0.947
0.95	0.846	0.996	0.923

Both of them have a very good predictive ability, AIC model performs a little higher than BIC model. But from the perspective of efficiency, the BIC model is preferable. Because BIC model use fewer variables to achieve the as high as AIC model's predictive ability. In reality, the performance depends on how the cutoff point been selected.

3.3.2 Ada-Boosting

Ada-boosting is a weak learning based method to find a good classification strategy to classify the data, also called exponential loss minimization algorithm. Its key idea is to give more weight to the subjects that wrongly classified during the last round of classification and change the previous distribution. This algorithm does not make any assumption on the model.

The algorithm of AdaBoost: Yoav.F, Robert E.S, *A Short Introduction to Boosting*, 1999.

Given: $(x_1, y_1), \dots, (x_m, y_m)$ where $x_i \in X, y_i \in Y = \{-1, +1\}$ Initialize $D_1(i) = 1/m$.

For $t = 1, \dots, T$:

- Train weak learner using distribution D_t .
- Get weak hypothesis $h_t : X \rightarrow \{-1, +1\}$ with error $\epsilon_t = Pr_{i \sim D_t} [h_t(x_i) \neq y_i]$
- Choose $\alpha_t = \frac{1}{2} \ln \left(\frac{1-\epsilon_t}{\epsilon_t} \right)$

$$D_{t+1}(i) = \frac{D_t(i)}{Z_t} \times \begin{cases} e^{-\alpha_t} & \text{if } h_t(x_i) = y_i \\ e^{\alpha_t} & \text{if } h_t(x_i) \neq y_i \end{cases}$$

where Z_t is a normalization factor.

Output the final hypothesis: $H(x) = \text{sign} \left(\sum_{t=1}^T \alpha_t h_t(x) \right)$

Conduct 5 rounds ada-boosting training on original data, AIC selected data and BIC selected data, using 1/2 of the data as training data and the rest 1/2 of data as test data. Besides the results will be compared with the best cutoff point of BIC logistic model. The

Round #	Original data	AIC data	BIC data	BIC LR
	SNS; SPC; ACC	SNS; SPC; ACC	SNS; SPC; ACC	CUT; SNS; SPC; ACC
1	0.975; 0.925; 0.949	0.967; 0.962; 0.965	0.967; 0.947; 0.957	0.75; 0.934; 0.962; 0.949
2	0.953; 0.936; 0.945	0.953; 0.920; 0.937	0.961; 0.936; 0.949	0.80; 0.938; 0.984; 0.961
3	0.949; 0.951; 0.949	0.954; 0.959; 0.957	0.954; 0.959; 0.957	0.70; 0.939; 1.000; 0.969
4	0.983; 0.932; 0.957	0.983; 0.955; 0.961	0.983; 0.962; 0.972	0.65; 0.943; 1.000; 0.972
5	0.983; 0.924; 0.953	0.983; 0.931; 0.957	0.983; 0.947; 0.965	0.75; 0.967; 0.969; 0.969
Average	0.969; 0.934; 0.951	0.968; 0.945; 0.955	0.970; 0.950; 0.960	0.73; 0.944; 0.983; 0.964

performance of Ada-boosting improved by subsetting data according to the AIC and BIC model data and a little better in BIC model data. Compared with BIC logistic regression model, Ada-boosting has better performance in sensitivity and has more stable performance in SNS, SPC and ACC through all 5 rounds. BIC logistic regression model performs better in average than Ada-boosting, but its performance depends on the cutoff point.

3.3.3 Conclusion

As we see, both Ada-boosting and logistic regression model can make good prediction in this case. Ada-boosting is more easy and convenient to use; however, it cannot make any inferences about variables. If one needs to determine and make a strong inference about certain variables, the logistic regression model is necessary.

It is worth keeping in mind that logistic regression model could be better than Ada-boosting when the model been fitted so well and the best cutoff point been selected, but it not always be. Similarly, Ada-boosting could be better when logistic model is too complex to get , but it will also lose some efficiency when there are too many noise variables.

3.4 Regression Trees: Differences in Males and Females (Liza)

As we have shown in previous sections, males and females differ significantly in terms of weight and body measurements. However, the regression presented for predicting weight did not include a term for gender, citing that doing so would not add significantly to the model. My hypothesis is that separate models for males and females would be more appropriate based on the systematic differences in body shape and size between the genders. Using recursive partitioning, or regression tree analysis, I plan to explore what the most important variables are in predicting weight for the two genders both together and separately to see if there are differences in the way the data is divided based on the different body measurements sampled.

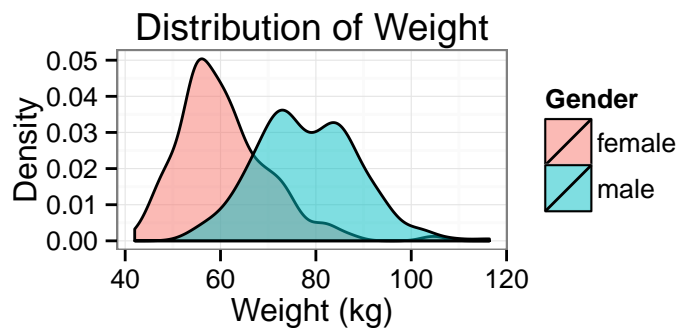
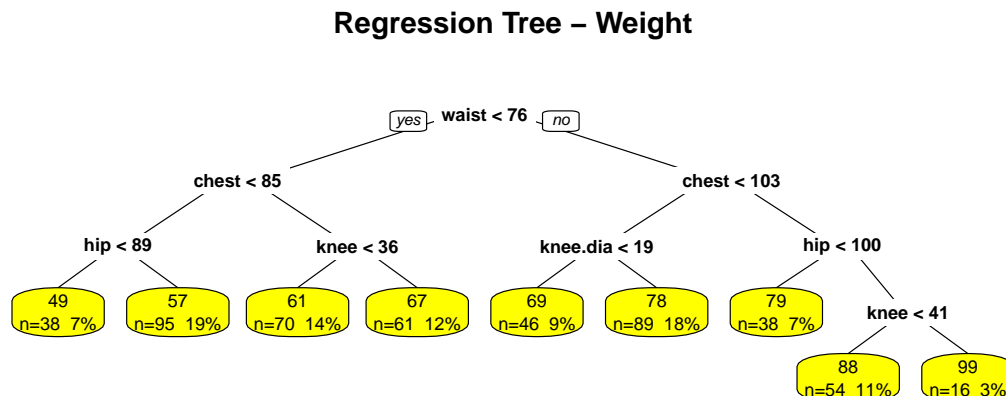


Figure 1: Desnsity distribution of Weight for males and females

3.4.1 Tree 1: Males and Females

To explore whether gender is an important dividing variable for the entire dataset, I first grew a regression tree with all of the data. As shown in the regression tree below, a split was not produced based on gender of the subjects. Variables used in this tree are waist girth, chest girth, hip girth, knee girth and knee diameter.

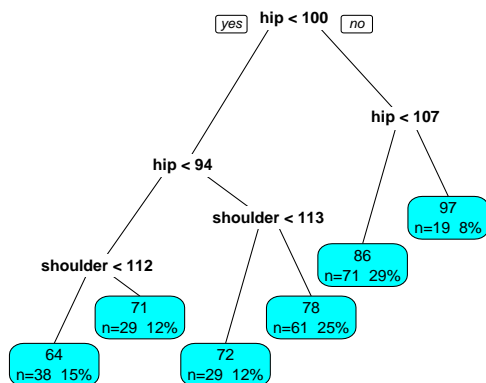


3.4.2 Tree 2: Males

Even though the full dataset did not utilize gender to partition the data, it might be informative to see if we obtain different trees by subsetting the data by gender. Next, I produced a regression tree with weight as the dependent variable for males only using all available body measurement variables. The CP tables showed us that at 5 splits we obtain a minimum relative error, therefore we were able to prune the tree to a reasonable size to avoid overfitting the data (Figure 2a). The final variables used to produce the pruned tree were hip girth and shoulder girth.

```
##
## Regression tree:
## rpart(formula = weight ~ ., data = bodym, method = "anova")
##
## Variables actually used in tree construction:
## [1] hip      shoulder
##
## Root node error: 27188/247 = 110
##
## n= 247
##
##      CP nsplit rel error xerror  xstd
## 1 0.552      0      1.00   1.00 0.094
## 2 0.129      1      0.45   0.47 0.049
## 3 0.069      2      0.32   0.36 0.041
## 4 0.032      3      0.25   0.31 0.030
## 5 0.025      4      0.22   0.32 0.029
## 6 0.014      5      0.19   0.31 0.028
```

Pruned Tree, Weight (Male)



Regression Tree, Weight (Female)

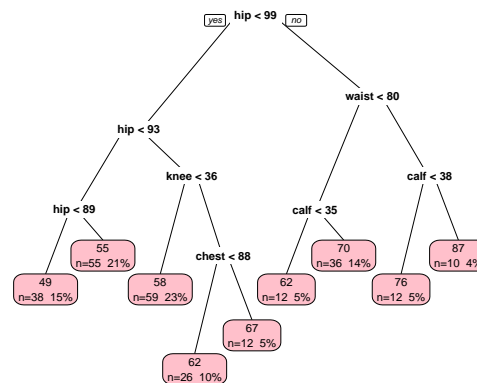


Figure 2: a. Pruned regression tree for male weight. b. Regression tree for female weight.

3.4.3 Tree 3: Females

The regression tree for the female subset (Figure 2b) does not reach a minimum relative error before the final node and therefore did not need pruning. The variables used to produce this tree vary significantly in number and type from the male tree. They are: hip girth, knee girth, calf girth, chest girth, and waist girth. In fact, the only variable the female tree and the pruned male tree have in common is shoulder girth.

```
##
## Regression tree:
## rpart(formula = weight ~ ., data = bodyf, method = "anova")
##
## Variables actually used in tree construction:
## [1] calf chest hip knee waist
##
## Root node error: 23948/260 = 92
##
## n= 260
##
##      CP nsplit rel error xerror  xstd
## 1 0.516     0     1.00  1.01 0.134
## 2 0.129     1     0.48  0.54 0.072
## 3 0.112     2     0.35  0.41 0.069
## 4 0.031     3     0.24  0.30 0.040
## 5 0.028     4     0.21  0.29 0.039
## 6 0.027     5     0.18  0.28 0.039
## 7 0.019     6     0.16  0.26 0.039
## 8 0.010     7     0.14  0.25 0.038
## 9 0.010     8     0.13  0.23 0.038
```

3.4.4 Conclusion

While regression trees are not intended to produce a model for prediction, they can be quite useful in elucidating the most important variables for partitioning data. Here we found that these variables differed significantly between men and women, perhaps suggesting that separate models for predicting weight would be more appropriate and potentially more accurate than one general model for both genders. Further analysis would be needed to test this hypothesis, specifically model fitting and selection for subsets of the data by gender. One could then show whether these separate models have better predictive power than one combined model.

4 Conclusion/ Discussion