## Math Tests

































So the log likelihood is

p(gij|0i,02)~N(0i,02) p (θ:[72,σ2)~ N(μ, 72σ2)

First we find the form of the posterior:

p(0:, M, 0+,7=|4) < p(M)p(0+)p(7+)p(0|1+,0+), p(4|0,0+)

p(11) ~ 1 P(02) ~ + p(72) ~ IG(生,生)

then set it to 0:

 $\ll \exp\left\{\frac{1}{2\sigma^2}\sum_{i=1}^{p}\sum_{j=1}^{Ni}\left(y_{ij}-\theta_i\right)^2\right\}$ 

 $\varrho(\theta_{1},...,\theta_{p}) = \frac{1}{2\sigma^{2}} \sum_{i=1}^{p} \sum_{j=1}^{N_{i}} (y_{ij} - \theta_{i})^{2}$ 

 $-\frac{1}{\sigma^2}\sum_{j=1}^{N_1}y_{i,j}-\stackrel{\wedge}{\Theta}_{i}^{\dagger}=0 \Rightarrow \stackrel{\wedge}{\Theta}_{i}^{\dagger}=\overline{y_{i}}$ 

b.) The reason extreme \$\overline{y}\_i\$'s occurs when hi is low is due to the sample variance being bigger when the sample size is small.

MVN( [], T202 [p)

 $= (1) \left(\frac{1}{\sigma^{2}}\right) \left(\tau^{2}\right)^{-\left(\frac{1}{2}+1\right)} exp\left[\frac{1}{2\tau^{2}}\right] \left(\underbrace{\det\left(2\pi \tau^{2}\sigma^{2} I_{p}\right)^{\frac{1}{2}}}_{} exp\left[\frac{1}{2}\left(\theta-\mu^{\frac{2}{1}}\right)^{\frac{1}{2}}\left(\tau^{2}\sigma^{2} I_{p}\right)^{\frac{1}{2}}\left(\theta-\mu^{\frac{2}{1}}\right)^{\frac{1}{2}}\right] \left(\underbrace{\prod_{i=1}^{N_{i}} \prod_{j=1}^{N_{i}} \exp\left[\frac{1}{2\sigma^{2}}\left(\psi_{i}-\theta_{i}\right)^{2}\right]}_{}\right) \left(\underbrace{\prod_{i=1}^{N_{i}} \prod_{j=1}^{N_{i}} \left(\psi_{i}-\theta_{i}\right)^{2}}_{}\right) \left(\underbrace{\prod$ 

 $= \left( \left( \sigma^{2} \right)^{-1} \left[ \Upsilon^{2} \right]^{\frac{3}{2}} exp \left\{ \frac{1}{2T^{2}} \right\} \left( \Upsilon^{2} \right)^{\frac{1}{2}} \left( \sigma^{2} \right)^{\frac{1}{2}} exp \left\{ \frac{1}{2T^{2}} \left( \theta - \mu \vec{1} \right)^{1} \left( \theta - \mu \vec{1} \right) \right\} \left( \sigma^{2} \right)^{\frac{1}{2}} exp \left\{ \frac{1}{2\sigma^{2}} \sum_{i=1}^{N} \frac{N_{i}}{2\sigma^{2}} \left( y_{ij} - \theta_{i} \right)^{2} \right\}$ 

we obtain the conditional distributions:  $\rho(\sigma^2) \leftarrow (\sigma^2)^{-\frac{(N_{\frac{1}{2}}^2)}{2}-1} \exp\left\{-\frac{1}{6^2}\left[\frac{1}{2^{\frac{1}{2}}}(\theta-N_{\frac{1}{2}})^{\frac{1}{2}}(\theta-N_{\frac{1}2})^{\frac{1}{2}}(\theta-N_{\frac{1}{2}})^{\frac{1}{2}}(\theta-N_{$ 

 $< (\sigma^{2})^{-\left(\frac{N+p}{2}\right)-1} (\tau^{2})^{-\left(\frac{p+3}{2}\right)} \exp\left\{\frac{-1}{2T^{2}}\right\} \exp\left\{\frac{-1}{2T^{2}\sigma^{2}}\left(\theta-\mu\widehat{\mathbf{1}}\right)^{1}\left(\theta-\mu\widehat{\mathbf{1}}\right)\right\} \exp\left\{\frac{-1}{2\sigma^{2}}\sum_{i=1}^{p}\sum_{j=1}^{p}\left(y_{ij}-\theta_{i}\right)^{2}\right\}$ 

 $P[\Upsilon^{2}] \times [\Upsilon^{2}] \xrightarrow{(\Gamma^{2})^{-1}} \exp\left\{-\frac{1}{7^{2}} \left(\frac{1}{2} + \frac{1}{20^{2}} (\theta - M\vec{1})^{1} (\theta - M\vec{1})^{2}\right) \right\} \rightarrow IG\left(\frac{Df!}{2}, \frac{1}{2} + \frac{1}{20^{2}} (\theta - M\vec{1})^{1} (\theta - M\vec{1})^{2}\right)$ 

 $\frac{d}{d\theta_i} L(\theta) = \frac{-1}{20^{L}} \sum_{i=1}^{N} 2(y_{ij} - \theta_i)$ 

Now to maximize this we take the derivative wint each ti:











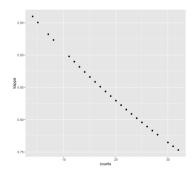
$$\begin{split} &p(\Theta_{i}) < exp\left\{\frac{-1}{2T^{2}\sigma^{2}}\left(\Theta_{i}^{-}\mathcal{M}\right)^{2} - \frac{1}{2\sigma^{2}}\sum_{j=1}^{N_{i}}\left(\Theta_{i}^{-}\mathcal{Y}_{i,j}^{\prime}\right)^{2}\right\} \rightarrow Normal, \, mean = \frac{1}{T^{2}\sigma^{2}}\frac{N_{i}}{\sigma^{2}} \quad \text{and vormance } = \frac{N_{i}}{\sigma^{2}} + \frac{1}{\sigma^{2}}\\ &p(\mathcal{M}) < exp\left\{\frac{-1}{2T^{2}\sigma^{2}}\left(\Theta_{-}\mathcal{M}^{\frac{1}{2}}\right)^{1}\left(\Theta_{-}\mathcal{M}^{\frac{1}{2}}\right)^{\frac{1}{2}} = exp\left\{\frac{-1}{2T^{2}\sigma^{2}}\left(P_{\mathcal{M}^{2}} - 2\mu\sum_{j=1}^{p}\Theta_{i}^{+} + \sum_{j=1}^{p}\Theta_{i}^{+}^{2}\right)^{\frac{1}{2}} \rightarrow N\left(\overline{\Theta}, \frac{T^{2}\sigma^{2}}{P}\right) \end{split}$$

Implement Gibbs sampling...



d) 
$$E(\Theta_{i} | y, \gamma^{2}, \sigma^{2}, \mu) = \frac{\frac{1}{\gamma^{2}\sigma^{2}}A + \frac{N}{\sigma^{2}}\overline{y}_{i}}{\frac{1}{\gamma^{2}\sigma^{2}} + \frac{N!}{\sigma^{2}}} = \left(\frac{\frac{1}{\gamma^{2}\sigma^{2}}}{\frac{1}{\gamma^{2}\sigma^{2}} + \frac{N!}{\sigma^{2}}}\right)A + \left(1 - \frac{\frac{1}{\gamma^{2}\sigma^{2}}}{\frac{1}{\gamma^{2}\sigma^{2}} + \frac{N!}{\sigma^{2}}}\right)\overline{y}_{i}$$

$$k_i = \frac{\frac{1}{\tau^2 \sigma^2}}{\frac{1+\tau^2 N_i}{\tau^2 \sigma^2}} = \boxed{\frac{1}{1+\tau^2 N_i}}$$



$$\begin{aligned} & \text{cov}(y_{ij},y_{ik}) = \text{cov}(\mu + \delta_i + e_{ij}, \mu + \delta_i + e_{ik}) = \text{E}[(\mu + \delta_i + e_{ij} - \mu - \text{E}(\delta_i))(\mu + \delta_i + e_{ik} - \mu - \text{E}(\delta_i))] \\ & = \text{E}[(S_i - \text{E}(S_i) + e_{ij})(S_i - \text{E}(\delta_i) + e_{ik})] = \text{Var}(S_i) = \boxed{\uparrow^2 \sigma^2} \\ & \text{cov}(y_{ij},y_{ik}) = \text{Cov}(\mu + S_i + e_{ij}, \mu + \delta_i + e_{ik}) = \boxed{0} \end{aligned}$$

f.) No—the data has higher variance when Ni is low

Blood Pressure