Cheese Let i be the store index and j be the index of the observations of a single store. It is the display indicator.

The model is as follows:

$$log@ij = \alpha_{0,i} + \alpha_{1,i} I_{ij} + \beta_{0,i} logPij + \beta_{1,i} logPij I_{ij}$$

= 
$$\alpha_0$$
;  $+\beta_0$ ;  $\log P_{ij} + I_{ij}(\alpha_1, i + \beta_1, i \log P_{ij})$ 

can also write this in linear regression/matrix form.

Let 
$$g_{ij}$$
 denote  $log Q_{ij}$ , and for each now in the dataset define  $x_{ij} = \begin{bmatrix} 1 & \text{Ii} \\ & \text{Ii} \end{bmatrix} \log P_{ij} = \begin{bmatrix} 1 & \text{Ii} \\ & \text{Ii} \end{bmatrix} \log P_{ij}$ 

Then our model may be written like thus:

Now let's define the full model.

Likelihood: 
$$p(y_{ij} | \beta_i, \sigma_i^2) \sim N(x_{ij} | \beta_i, \sigma_i^2)$$

Priors: 
$$p(\beta) \sim N(m, S^2I)$$

$$p(\sigma_i^2) \sim IG(\frac{a}{2}, \frac{b}{2})$$
 #set  $\alpha = 4, b = 2$ 

$$p(\sigma_i^2) \sim IG(\frac{a}{2}, \frac{b}{2})$$
 #set  $a=4, b=1$ 

Hyperpriors: 
$$p(m) \sim 1$$

$$p(s^2) \sim IG(\frac{1}{2}, \frac{1}{2})$$

Now (on the next page) we compute the postenor.

$$\begin{split} \rho(\beta;l\dots) & \sim \rho(\beta;) \cdot \rho(y;l|\beta;,\sigma;^2) \\ & \propto \exp\left\{\frac{1}{2s^2}(m-\beta)^l(s^2I)^{-l}(m-\beta)^l\right\} \exp\left\{\left(y;-X_i^l\beta;\right)^l(\sigma_{i^2}I)^{-l}(y;-X_i^l\beta;)\right\} \text{ where } y;=\begin{bmatrix}y_{11}\\ \vdots\\ y_{1Ni}\end{bmatrix} \\ & (\text{Normal-Normal Conjugate model done previously}) \\ & \rightarrow \text{Normal}\left(\mu^*, \Xi^{1*}\right) \text{ with } \Xi^{1*}_{1} = \left(\frac{1}{\sigma_{i^2}}\chi^{i}/X_i + \frac{1}{s^2}I\right)^{-l} \text{ and } \mu^* = \Xi^{1*-l}\left(\frac{1}{\sigma_{i^2}}\chi^{i}/Y_i + \frac{1}{s^2}m\right) \\ & \rho(\sigma_{i^2}) \sim \rho(a) \rho(b) \rho(\sigma_{i^2}^{-l} | \alpha_{r}b) \rho(y_i | \beta_i, \sigma_{i^2}^{-l}) \\ & \propto \alpha \exp\left\{-\alpha^{\frac{1}{2}}b \exp\left\{-b^{\frac{1}{2}}(\sigma_{i^2})^{\frac{\alpha}{2}-l} \exp\left\{-\frac{1}{\sigma_{i^2}}\frac{b}{2}\right\} \cdot \frac{1}{\sqrt{\det(\sigma_{i^2}I)}} \exp\left\{-\frac{1}{2}\left(y_i - \chi_i^l\beta_i\right)(\sigma_{i^2}I)^{-l}(y_i - \chi_i^l\beta_i)\right\} \\ & \sim \rho(a) \rho(b) \rho(\sigma_{i^2}^{-l} | \alpha_{r}b) \rho(y_i | \beta_i, \sigma_{i^2}^{-l}) \\ & \sim \alpha \exp\left\{-\alpha^{\frac{1}{2}}b \exp\left\{-b^{\frac{1}{2}}(\sigma_{i^2})^{\frac{\alpha}{2}-l} \exp\left\{-\frac{1}{\sigma_{i^2}}\frac{b}{2}\right\} \cdot \frac{1}{\sqrt{\det(\sigma_{i^2}I)}} \exp\left\{-\frac{1}{2}\left(y_i - \chi_i^l\beta_i\right)(\sigma_{i^2}I)^{-l}(y_i - \chi_i^l\beta_i)\right\} \\ & \sim \rho(a) \rho(b) \rho(\sigma_{i^2}^{-l} | \alpha_{r}b) \rho(y_i | \beta_i, \sigma_{i^2}^{-l}) \\ & \sim \rho(a) \rho(b) \rho(\sigma_{i^2}^{-l} | \alpha_{r}b) \rho(y_i | \beta_i, \sigma_{i^2}^{-l}) \\ & \sim \rho(a) \rho(b) \rho(\sigma_{i^2}^{-l} | \alpha_{r}b) \rho(y_i | \beta_i, \sigma_{i^2}^{-l}) \\ & \sim \rho(a) \rho(b) \rho(\sigma_{i^2}^{-l} | \alpha_{r}b) \rho(y_i | \beta_i, \sigma_{i^2}^{-l}) \\ & \sim \rho(a) \rho(b) \rho(\sigma_{i^2}^{-l} | \alpha_{r}b) \rho(y_i | \beta_i, \sigma_{i^2}^{-l}) \\ & \sim \rho(a) \rho(b) \rho(\sigma_{i^2}^{-l} | \alpha_{r}b) \rho(y_i | \beta_i, \sigma_{i^2}^{-l}) \\ & \sim \rho(a) \rho(b) \rho(\sigma_{i^2}^{-l} | \alpha_{r}b) \rho(y_i | \beta_i, \sigma_{i^2}^{-l}) \\ & \sim \rho(a) \rho(b) \rho(\sigma_{i^2}^{-l} | \alpha_{r}b) \rho(y_i | \beta_i, \sigma_{i^2}^{-l}) \\ & \sim \rho(a) \rho(b) \rho(\sigma_{i^2}^{-l} | \alpha_{r}b) \rho(y_i | \beta_i, \sigma_{i^2}^{-l}) \\ & \sim \rho(a) \rho(b) \rho(\sigma_{i^2}^{-l} | \alpha_{r}b) \rho(y_i | \beta_i, \sigma_{i^2}^{-l}) \\ & \sim \rho(a) \rho(b) \rho(\sigma_{i^2}^{-l} | \alpha_{r}b) \rho(y_i | \beta_i, \sigma_{i^2}^{-l}) \\ & \sim \rho(a) \rho(b) \rho(\sigma_{i^2}^{-l} | \alpha_{r}b) \rho(y_i | \beta_i, \sigma_{i^2}^{-l}) \\ & \sim \rho(a) \rho(b) \rho(\sigma_{i^2}^{-l} | \alpha_{r}b) \rho(y_i | \beta_i, \sigma_{i^2}^{-l}) \\ & \sim \rho(a) \rho(b) \rho(\sigma_{i^2}^{-l} | \alpha_{r}b) \rho(y_i | \beta_i, \sigma_{i^2}^{-l}) \\ & \sim \rho(a) \rho(b) \rho(\sigma_{i^2}^{-l} | \alpha_{r}b) \rho(\sigma_{i^2}^{-l} | \alpha_$$

$$\longrightarrow$$
 Normal  $\left(\frac{1}{n}\sum_{i}^{n}\beta_{i},\frac{1}{nS^{2}}I\right)$ 

$$p(s^{2}|...) \propto p(s^{2}) \cdot p(\beta;|s;)$$

$$\propto (s^{2})^{\frac{1}{2}-1} \exp\left\{-\frac{1}{s^{2}} \cdot \frac{1}{2}\right\} \cdot \prod_{i=1}^{l} \frac{1}{\sqrt{\det(s^{2}I)}} \exp\left\{-\frac{1}{2}(\beta;-m)^{l}(s^{2}I)^{-l}(\beta;-m)\right\}$$

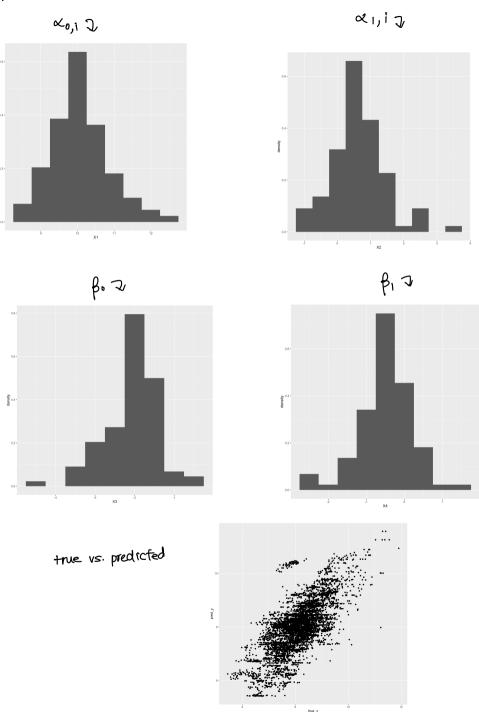
$$\propto (s^{2})^{\frac{-1}{2}-\frac{n}{2}-1} \exp\left\{-\frac{1}{s^{2}} \cdot \frac{1}{2} - \frac{1}{2s^{2}} \cdot \sum_{i=1}^{l} (\beta;-m)^{l}(\beta;-m)\right\}$$

$$= (s^{2})^{-(n+1)/2-1} \exp\left\{-\frac{1}{s^{2}} \cdot \frac{1}{2} \left(1 + \sum_{i=1}^{l} (\beta;-m)^{l}(\beta;-m)\right)\right\}$$

Now, code a Gibbs sampler for it-

 $\rightarrow IG(\frac{n+1}{2},\frac{1}{2}(1+\frac{1}{2}(\beta_{i}-m)^{i}(\beta_{i}-m)))$ 

Posterior mean estimates for:



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Polls

Model using

P(Y;j=1) = \Phi(X;j^{1}\beta;)
\beta; \sim N(\mu;, \Sigma;)
\mu; \sim N(m,S^{2}I_{p}) \text{ here take } m=0, S^{2}=10^{6}
\Sigma_{i}^{1} \sim Inv-Wishard(p-1, I_{p})
Its wields the following posteriors:

N(x,y) = (\sum_{i=1}^{j-1} + \sum_{i=1}^{j-1} y_{i})^{-1} m = V^{-1}(\sum_{i=1}^{j-1} y_{i} + \sum_{i=1}^{j-1} y_{i})^{-1} m = V^{-1}(\sum_{i=1}^{j-1} y_
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Besag (1974) lattice systems

This yields the following posteriors: 
$$(\beta_i|...) \times N(\theta, V) = \left(\sum_{i=1}^{j-1} + \chi_i^*|\chi_i^*\right)^{-1}, \ m = V^{-1}\left(\sum_{i=1}^{j-1} \mu_i + \chi_i^*|\chi_i^*\right)^{-1}$$
 
$$(\Xi_{ij}^*|...) \times \left\{ N(\chi_{ij}^*|\beta_i, 1) \middle|_{[0, \infty)} \text{ for } Y_{ij} = 1 \right.$$
 
$$\left[ N(\chi_{ij}^*|\beta_i, 1) \middle|_{[-\infty, 0]} \text{ for } Y_{ij} = 0 \right.$$
 
$$\left( \mu_i^*|... \right) \times N(\theta, V) \text{ w/ } V = \left(\sum_{i=1}^{j-1} + s^2 I\right)^{-1}, \theta = \beta_i^* \Xi_{i}^{j-1}$$
 
$$\left( \Sigma_{ij}^*|... \right) \times Inv-W_i \left( n+p-1, I_p + \sum_{i=1}^{j-1} (\beta_i - \mu_i^*) (\beta_i - \mu_i^*)^T \right)$$