

- **Problem 1: [Stationarity]** Let  $(Z_t : t \in \mathbb{Z})$  be a sequence of independent zero mean and normal random variables with variance  $\sigma^2$  and let  $a, b$ , and  $c$  be constants. Which of the following processes are weakly and/or strictly stationary? For each weakly stationary process specify the mean and ACVF.

(a)  $X_t = a + bZ_t + cZ_{t-1}$ .

Answer: Let  $i, j \in \mathbb{Z}$  be any two integers. Then,

$$E[X_i] = E[a + bZ_i + cZ_{i-1}] = a = E[a + bZ_j + cZ_{j-1}] = E[X_j],$$

and

$$\text{Var}[X_i] = \text{Var}[a + bZ_i + cZ_{i-1}] = 0 + b^2\sigma^2 + c^2\sigma^2 = \text{Var}[a + bZ_j + cZ_{j-1}] = \text{Var}[X_j],$$

where we have implicitly used the fact that every pair of  $Z_t$  have zero covariance.

Thus, with knowledge that linear combinations of normally distributed random variables are normal, we have

$$X_i, X_j \sim \mathcal{N}(a, (b^2 + c^2)\sigma^2).$$

Since  $i, j$  were arbitrary, we have that every element  $X_t$  has the same distribution, which implies that  $X_t$  is *strictly stationary*.

Since we have explicitly shown that we have finite variance, then we may also conclude that  $X_t$  is *weakly stationary*.  $E[X_t] = a \forall t$ , and note that

$$\text{Cov}(X_{t+1}, X_t) = \text{Cov}(a + bZ_{t+1} + cZ_t, a + bZ_t + cZ_{t-1}) = \text{Cov}(cZ_t, bZ_t) = bc\sigma^2.$$

Thus,

$$\gamma(h) = \begin{cases} (b^2 + c^2)\sigma^2 & h = 0 \\ bc\sigma^2 & h = \pm 1 \\ 0 & |h| > 1 \end{cases}.$$

(b)  $X_t = Z_t \cos(ct) + Z_{t-1} \sin(ct)$

Answer: Let  $i, j \in \mathbb{Z}$  be any two integers. Then,

$$E[X_i] = E[Z_i \cos(ci) + Z_{i-1} \sin(ci)] = 0 = E[Z_j \cos(cj) + Z_{j-1} \sin(cj)] = E[X_j],$$

and

$$\text{Var}[X_i] = \text{Var}[Z_i \cos(ci) + Z_{i-1} \sin(ci)] = \sigma^2 \cos^2(ci) + \sigma^2 \sin^2(ci) = \sigma^2,$$

$$\text{Var}[X_j] = \text{Var}[Z_j \cos(cj) + Z_{j-1} \sin(cj)] = \sigma^2 \cos^2(cj) + \sigma^2 \sin^2(cj) = \sigma^2,$$

where we have implicitly used the fact that every pair of  $Z_t$  have zero covariance. Thus, with knowledge that linear combinations of normally distributed random variables are normal, we have

$$X_i, X_j \sim \mathcal{N}(0, \sigma^2).$$

Since  $i, j$  were arbitrary, we have that every element  $X_t$  has the same distribution, which implies that  $X_t$  is *strictly stationary*.

Since we have explicitly shown that we have finite variance, then we may also conclude that  $X_t$  is *weakly stationary*.  $E[X_t] = 0 \forall t$ , and note that

$$\begin{aligned} \text{Cov}(X_{t+1}, X_t) &= \text{Cov}(Z_{t+1} \cos(ct + c) + Z_t \sin(ct + c), Z_t \cos(ct) + Z_{t-1} \sin(ct)) \\ &= \text{Cov}(cZ_t, bZ_t) = \sin(ct + c) \cos(ct) \sigma^2 \end{aligned}$$

Thus,

$$\gamma(h) = \begin{cases} \sigma^2 & h = 0 \\ \sin(ct + c) \cos(ct) \sigma^2 & h = +1 \\ \cos(ct + c) \sin(ct) \sigma^2 & h = -1 \\ 0 & |h| > 1 \end{cases}.$$

(c)  $X_t = a + bZ_0$

Answer: Let  $i, j \in \mathbb{Z}$  be any two integers. Then,

$$E[X_i] = E[a + bZ_0] = a = E[a + bZ_0] = E[X_j],$$

and

$$\text{Var}[X_i] = \text{Var}[a + bZ_0] = 0 + b^2 \sigma^2 = \text{Var}[a + bZ_0] = \text{Var}[X_j].$$

Thus, with knowledge that linear combinations of normally distributed random variables are normal, we have

$$X_i, X_j \sim \mathcal{N}(0, b^2 \sigma^2).$$

Since  $i, j$  were arbitrary, we have that every element  $X_t$  has the same distribution, which implies that  $X_t$  is *strictly stationary*.

Since we have explicitly shown that we have finite variance, then we may also conclude that  $X_t$  is *weakly stationary*.  $E[X_t] = 0 \forall t$ , and

$$\gamma(h) = b^2 \sigma^2, \forall h.$$

(d)  $X_t = Z_t Z_{t-1}$

Answer:

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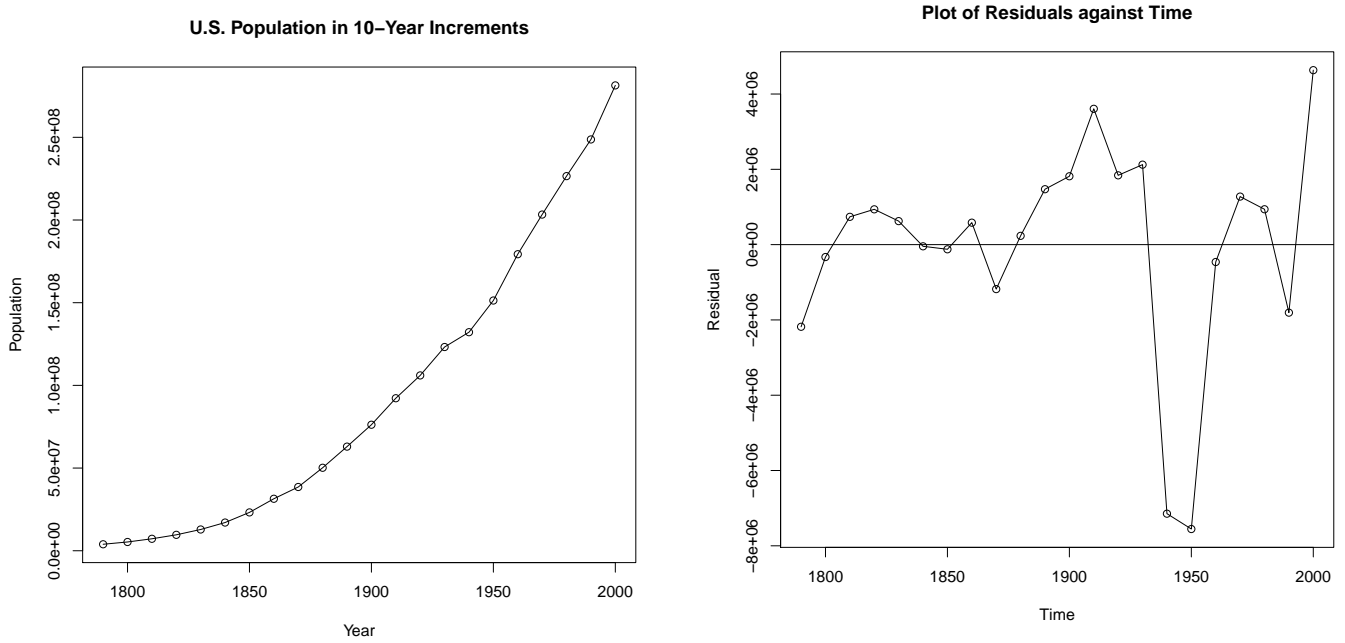


Figure 1: Plot of the population data for Problem 2      Figure 2: Plot of the Residuals against Time for the fitted values in Table 1.

- **Problem 2: [U.S. Population]** Download the file `population.xls` from the course website. It contains the size of the population in the U.S.A. at ten-year intervals from 1790 to 2000.

(a) Plot the data.

*See Figure 1.*

(b) Assuming the model  $X_t = m_t + Z_t$ ,  $E[Z_t] = 0$ , fit a polynomial trend  $\hat{m}_t$  to the data.

*See Table 1 for the fitted values for a 2<sup>nd</sup> degree polynomial*

$\hat{m}_t$
6110720.45
5637937.68
6501266.64
8700707.34
12236259.78
17107923.95
23315699.86
30859587.50
39739586.88
49955698.00
61507920.85
74396255.44
88620701.76
104181259.82
121077929.62
139310711.15
158879604.42
179784609.42
202025726.16
225602954.63
250516294.84
276765746.79

Table 1: Fitted values for a 2<sup>nd</sup> degree polynomial fit

- (c) Plot the residuals  $\hat{Z}_t = X_t - \hat{m}_t$ . Comment on the quality of the fitted model.

*See Figure 2.*

- (d) Use the fitted model to predict the population size in 2010 and 2020 (using predicted noise values of zero).

*See Table 2.*

	2010	2020
$\hat{m}$	304351310.47	333272985.89

Table 2: Predicted population values for the years 2010 and 2020.

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- **Problem 3: [Projection Theorem]** If  $\mathcal{M}$  is a closed subspace of a Hilbert Space  $\mathcal{H}$  and  $x \in \mathcal{H}$ , prove that

$$\min_{y \in \mathcal{M}} \|x - y\| = \max \left\{ |\langle x, z \rangle| : z \in \mathcal{M}^\perp, \|z\| = 1 \right\},$$

where  $\mathcal{M}^\perp$  is the orthogonal complement of  $\mathcal{M}$ .

Answer: We show the desired result in two cases:

Case 1: Suppose that  $x \in \mathcal{M}$ . Then  $\min_{y \in \mathcal{M}} \|x - y\| = 0$ , where the minimum is obtained at  $y = x$ . Moreover,  $\langle x, z \rangle = 0 \forall z \in \mathcal{M}^\perp$ , due to orthogonality.

Case 2: Suppose that  $x \in \mathcal{M}^\perp$ . Then, for  $y \in \mathcal{M}, z \in \mathcal{M}^\perp$ , we can write

$$\langle x, z \rangle = \langle x - y, z \rangle + \langle y, z \rangle = \langle x - y, z \rangle,$$

and by the triangle inequality,

$$|\langle x - y, z \rangle| \leq \|x - y\| \|z\| = \|x - y\|.$$

Thus, we have

$$|\langle x, z \rangle| \leq \|x - y\|,$$

and it follows naturally that in order for equality to hold we must have the maximum on the left-hand-side, and the minimum on the right-hand-side. That is,

$$\max |\langle x, z \rangle| = \min \|x - y\|.$$

- **Problem 4: [Prediction Equations]** If  $X_t = Z_t - \theta Z_{t-1}$ , where  $|\theta| < 1$  and  $(Z_t : t \in \mathbb{Z})$  is a sequence of uncorrelated random variables, each with mean 0 and variance  $\sigma^2$ , show by checking the prediction equations that the best mean square predictor of  $X_{n+1}$  in  $\overline{\text{sp}}(X_j : j \leq n)$  is

$$\hat{X}_{n+1} = - \sum_{j=1}^{\infty} \theta^j X_{n+1-j}.$$

What is the mean squared error of  $\hat{X}_{n+1}$ ?

Answer: