

EMSE 6035: Marketing of Technology

Modeling Heterogeneous Preferences

John Paul Helveston, Ph.D.
Assistant Professor
Engineering Management & Systems Engineering
The George Washington University

Background: Homogeneous Utility Model

$$\tilde{u}_j = \boldsymbol{\beta}' \mathbf{x}_j + \tilde{\varepsilon}_j$$
$$= \boxed{\beta_1} x_{j1} + \boxed{\beta_2} x_{j2} + \dots + \tilde{\varepsilon}_j$$

Weights that denote the
relative value of attributes
 x_{j1}, x_{j2}, \dots

This model is a *homogeneous* model:
everyone has the same coefficients

Est.	Std. Err.
$\hat{\beta}_1$	σ_1
$\hat{\beta}_2$	σ_2
\vdots	\vdots
$\hat{\beta}_m$	σ_m

Heterogeneous Models

Heterogeneous models allow for *different* people to have *different* coefficients

We will cover two approaches (there are others):

Interaction Models

Group 1		Group 2	
Estimate	Std. Err.	Estimate	Std. Err.
$\hat{\beta}_1$	σ_1	$\hat{\beta}_1$	σ_1
$\hat{\beta}_2$	σ_2	$\hat{\beta}_2$	σ_2
\vdots	\vdots	\vdots	\vdots
$\hat{\beta}_m$	σ_m	$\hat{\beta}_m$	σ_m

Estimate a “mixed” logit (a.k.a. hierarchical) model

Estimate
$\hat{\beta}_1 \sim N(\hat{\mu}_1, \hat{\sigma}_1)$
$\hat{\beta}_2 \sim N(\hat{\mu}_2, \hat{\sigma}_2)$
\vdots
$\hat{\beta}_m \sim N(\hat{\mu}_m, \hat{\sigma}_m)$

Interaction models

Homogeneous model:

$$\tilde{u}_j = \beta_1 x_{j1} + \tilde{\varepsilon}_j$$

Two groups: A & B

$$\tilde{u}_j = \beta_1 x_{j1} + \underline{\beta_2 x_{j1} \delta^B} + \tilde{\varepsilon}_j$$

Where

$\delta^B = 1$ if the person is in group B

$\delta^B = 0$ if the person is in group A

Interaction models

$$\begin{aligned}\tilde{u}_j &= \beta_1 x_{j1} + \beta_2 x_{j1} \delta^B + \tilde{\varepsilon}_j \\ &= (\beta_1 + \beta_2 \delta^B) x_{j1} + \tilde{\varepsilon}_j\end{aligned}$$

Par.	Interpretation
$\hat{\beta}_1$	Effect of x_{j1} for group A
$\hat{\beta}_2$	<i>Difference</i> in effect of x_{j1} between groups A and B

Effect of x_{j1}	
Group A	Group B
$\hat{\beta}_1$	$\hat{\beta}_1 + \hat{\beta}_2$

The scale parameter

$$\tilde{u}_j = \beta_1 x_{j1} + \beta_2 x_{j1} \delta^B + \tilde{\varepsilon}_j$$

$$\tilde{\varepsilon}_j \sim \text{Gumbel} \left(0, \underset{\substack{\uparrow \\ \text{Scale parameter}}}{\sigma^2} \frac{\pi^2}{6} \right)$$

Scale parameter

$$\frac{\tilde{u}_j}{\cancel{\sigma}} = \frac{1}{\cancel{\sigma}} (\beta_1 x_{j1} + \beta_2 x_{j1} \delta^B + \tilde{\varepsilon}_j)$$

Assume $\sigma = 1$

$$\tilde{\varepsilon}_j \sim \text{Gumbel} \left(0, \frac{\pi^2}{6} \right)$$

What if we split the data and separately estimate two models:

$$\tilde{u}_j^A = \beta_1 x_{j1} + \tilde{\varepsilon}_j^A$$

$$\tilde{u}_j^B = \beta_1 x_{j1} + \tilde{\varepsilon}_j^B$$

$$\frac{\tilde{u}_j^A}{\sigma^A} = \frac{1}{\sigma^A} (\beta_1 x_{j1} + \tilde{\varepsilon}_j^A)$$

$$\frac{\tilde{u}_j^B}{\sigma^B} = \frac{1}{\sigma^B} (\beta_1 x_{j1} + \tilde{\varepsilon}_j^B)$$

Practice Question 1

Suppose we use the following utility model to describe preferences for cars:

$$\tilde{u}_j = \beta_1 x_j^{\text{price}} + \beta_2 x_j^{\text{mpg}} + \beta_3 \delta_j^{\text{elec}} + \varepsilon_j$$

- a) Using interactions, write out a model that accounts for differences in the effects of x_j^{price} , x_j^{mpg} , and δ_j^{elec} between two groups: A and B
- b) Write out the effects of x_j^{price} , x_j^{mpg} , and δ_j^{elec} for each group.

Use samples of $\hat{\boldsymbol{\beta}}$ to include uncertainty

Effect of x_{j1}	
Group A	Group B
$\hat{\beta}_1$	$\hat{\beta}_1 + \hat{\beta}_2$

$$\boldsymbol{\beta} \sim N(\hat{\boldsymbol{\beta}}, \Sigma)$$

$$\begin{bmatrix} \hat{\beta}_1 \\ \hat{\beta}_2 \\ \vdots \\ \hat{\beta}_n \end{bmatrix} \quad \underbrace{\begin{bmatrix} \sigma_{11}^2 & \cdots & \sigma_{1n}^2 \\ \vdots & \ddots & \vdots \\ \sigma_{1n}^2 & \cdots & \sigma_{jn}^2 \end{bmatrix}}$$

$$-\underbrace{\left[\nabla_{\boldsymbol{\beta}}^2 \ln(\mathcal{L}) \right]^{-1}}$$

Hessian

Example in R:

```
> library(MASS)
> beta = c(b1 = -0.7, b2 = 0.1)
> hessian = matrix(c(
+   -6000,  50,
+   50,   -6000),
+   ncol=2, byrow=T)
> covariance = -1*(solve(hessian))
> draws = as.data.frame(mvrnorm(10^5, beta,
+   covariance))
> b1_A = draws$b1
> b1_B = draws$b1 + draws$b2
> mean(b1_A)
[1] -0.7000618
> mean(b1_B)
[1] -0.6000408
> quantile(b1_A, c(0.025, 0.975))
      2.5%      97.5%
-0.7253724 -0.6748339
> quantile(b1_B, c(0.025, 0.975))
      2.5%      97.5%
-0.6358317 -0.5642215
```


Practice Question 2

Suppose we estimate the following utility model describing preferences for chocolate bars between two groups: A & B

$$\tilde{u}_j = \beta_1 x_j^{\text{price}} + \beta_2 x_j^{\text{cacao}} + \beta_3 x_j^{\text{price}} \delta_j^B + \beta_4 x_j^{\text{cacao}} \delta_j^B + \varepsilon_j$$

The estimated model produces the following coefficients:

Parameter	Coef.	Hessian			
β_1	-0.7	-6000	50	60	70
β_2	0.1	50	-700	50	100
β_3	0.2	60	50	-300	20
β_4	0.8	70	100	20	-6000

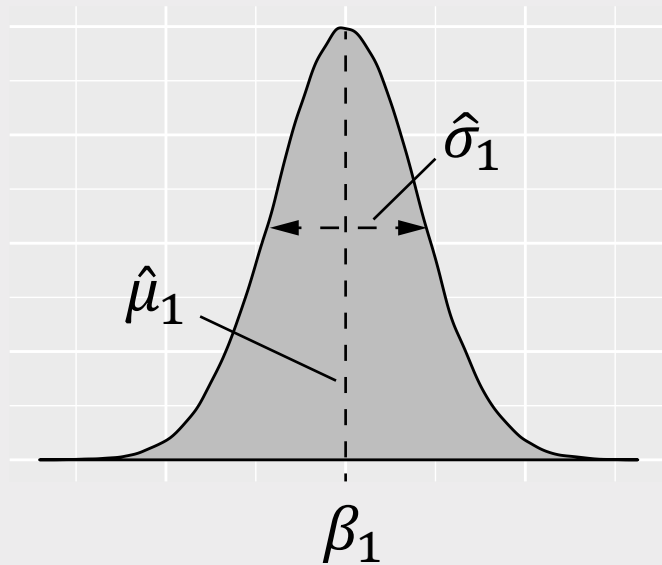
- Use the `mvrnorm()` function from the MASS library to generate 10,000 draws of the model coefficients.
- Use the draws to compute the mean and 95% confidence intervals of the effects of x_j^{price} and x_j^{cacao} for each group (A & B).

Mixed logit models

Homogeneous model:

$$\tilde{u}_j = \beta_1 x_{j1} + \tilde{\varepsilon}_j$$

In mixed logit, we assume that β_1 is distributed across the population:



Parameter	Estimate
β_1	$\hat{\mu}_1$
	$\hat{\sigma}_1$

The mlogit package includes mixed logit functionality

Standard Logit

$$P_j = \frac{e^{v_j}}{\sum_{k=1}^J e^{v_k}}$$

Mixed Logit

$$P_j = \int \left(\frac{e^{v_j}}{\sum_{k=1}^J e^{v_k}} \right) f(\beta) d\beta$$

Example in R:

$$\tilde{u}_j = \beta_1 x_j^{\text{price}} + \beta_2 x_j^{\text{cacao}} + \varepsilon_j$$

$$\beta_1 \sim N(\mu_1, \sigma_1), \quad \beta_2 \sim N(\mu_2, \sigma_2)$$

To estimate mixed logit models, we use simulation:

1. Draw a value of β from $f(\beta)$, label it β^r
2. Calculate the standard logit fraction using β^r :

$$P_j^r = \frac{e^{v_j^r}}{\sum_{k=1}^J e^{v_k^r}}$$

3. Repeat steps 1 & 2 R times and average the results:

$$\hat{P}_j = \frac{1}{R} \sum_{r=1}^R P_j^r$$

```
library(mlogit)
```

```
data_mlogit = mlogit.data(  
  data      = data,  
  choice    = 'choice',  
  alt.var   = 'alt',  
  id.var    = 'id')
```

```
model_mixed = mlogit(data_mlogit,  
  formula = choice ~ price + cacao | 0,  
  rpar = c(price = 'n', cacao = 'n'))
```

Practice Question 3

- a) Use the `read_csv()` function from the `tidyverse` library and the link in the video description to read in the “yogurt.csv” data frame.
- b) Use the `mlogit` library to estimate the following homogeneous model:

$$\tilde{u}_j = \beta_1 x_j^{\text{price}} + \beta_2 \delta_j^{\text{feat}} + \beta_3 \delta_j^{\text{dannon}} + \beta_4 \delta_j^{\text{highland}} + \beta_5 \delta_j^{\text{weight}} + \varepsilon_j$$

where the three δ coefficients are dummies for Dannon, Highland, and Weight Watchers brands and Yoplait is the baseline brand.

(Hint: Don’t forget to first use the `mlogit.data` function to format your data!)

- c) Use the `mlogit` library to estimate the same model but with the following mixing distributions:

$$\beta_1 \sim N(\mu_1, \sigma_1)$$

$$\beta_2 \sim N(\mu_2, \sigma_2)$$