

Problem 1:

(a) The PDF of a normal r.v. must have the form of

$$\frac{1}{\sqrt{2\pi}\sigma^2} \exp\left(\frac{-(x-\mu)^2}{2\sigma^2}\right), \quad \forall x \in \mathbb{R}$$

Next, we rearrange $f(x)$ as

$$f(x) = \sqrt{K} \exp\left(-K^2\left(x + \frac{1}{K}\right)^2\right)$$

(Note that \sqrt{K} suggests that $K \geq 0$)

$$= \sqrt{K} \exp\left(-\frac{\left(x + \frac{1}{K}\right)^2}{2 \cdot \left(\frac{1}{\sqrt{2}K}\right)^2}\right)$$

$$\Rightarrow \begin{cases} \sqrt{K} = \frac{1}{\sigma\sqrt{2\pi}} \Rightarrow \sigma = \frac{1}{\sqrt{2\pi}K} & \text{--- (1)} \\ \frac{1}{\sqrt{2}K} = \sigma & \text{--- (2)} \end{cases}$$

By (1)-(2), we have $\frac{1}{\sqrt{2}K} = \frac{1}{\sqrt{2\pi}K} \Rightarrow \underline{K = \pi}$ #

(b)

$$P(Z < 0 \mid X = +1) = P(Y < -1 \mid X = +1)$$

$$\begin{aligned} & \rightarrow = P(Y < -1) \\ & = \Phi\left(\frac{-1}{\sigma}\right) = 1 - \Phi\left(\frac{1}{\sigma}\right) \end{aligned}$$

Similarly, $P(Z \geq 0 | X = -1) = P(Y \geq 1 | X = -1)$

$$\begin{aligned} & \rightarrow P(Y \geq 1) \\ & \text{X, Y are independent} = 1 - \Phi\left(\frac{1}{\sigma}\right) \end{aligned}$$

Therefore, the error probability is $1 - \Phi(\frac{1}{\sigma})$ for both $X = +1$ and $X = -1$.

Hence, the overall error probability is also $1 - \Phi(\frac{1}{\sqrt{T}})$

Problem 2: $Y = aX + b$, is a linear transformation of X , and $X \sim \text{Exp}(\lambda)$

CDF:

$$F_Y(t) = P(Y \leq t) = P(aX + b \leq t)$$

Note that the CDF of X is

$$F_X(t) = \begin{cases} 1 - e^{-\lambda t}, & \text{if } t \geq 0 \\ 0, & \text{else.} \end{cases}$$

To find $F_Y(t)$, we need to discuss two cases, namely $a > 0$ and $a < 0$.

① $a > 0$:

$$F_Y(t) = P(aX + b \leq t) = P\left(X \leq \frac{t-b}{a}\right) = \begin{cases} 1 - e^{-\lambda \left(\frac{t-b}{a}\right)}, & \text{if } t \geq b \\ 0, & \text{else} \end{cases}$$

② $a < 0$:

$$F_Y(t) = P(aX + b \leq t) = P\left(X \geq \frac{t-b}{a}\right) = \begin{cases} e^{-\lambda \left(\frac{t-b}{a}\right)}, & \text{if } t \leq b \\ 1, & \text{else} \end{cases}$$

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since $a < 0$

From the above discussion, we know for Y to be an exponential r.v.,

We need to have $a > 0$ and $b = 0$.

(Conti).

PDF:

As it is assumed that the PDF of Y is continuous,

$$\text{we have } f_Y(t) = F'_Y(t).$$

Again, we consider the following two cases:

① $a > 0$:

$$f_Y(t) = F'_Y(t) = \begin{cases} \frac{\lambda}{a} e^{-\lambda(\frac{t-b}{a})} & , \text{ if } t > b \\ 0 & , \text{ else} \end{cases}$$

② $a < 0$:

$$f_Y(t) = F'_Y(t) = \begin{cases} -\frac{\lambda}{a} e^{-\lambda(\frac{t-b}{a})} & , \text{ if } t < b \\ 0 & , \text{ else} \end{cases}$$

Problem 3 :

$$h(x) := - \int_{f(x) > 0} f(x) \cdot \ln(f(x)) dx$$

As $X \sim N(0, \sigma^2)$, we know $f_X(x) = \frac{1}{\sigma\sqrt{2\pi}} \exp\left(-\frac{x^2}{2\sigma^2}\right)$, $\forall x \in \mathbb{R}$.

$$\text{Then, } h(X) = - \int_{-\infty}^{+\infty} f(x) \cdot \left[\ln\left(\frac{1}{\sigma\sqrt{2\pi}}\right) - \frac{x^2}{2\sigma^2} \right] dx$$

$$= - \left[\underbrace{\int_{-\infty}^{+\infty} \ln\left(\frac{1}{\sigma\sqrt{2\pi}}\right) f(x) dx}_{\text{a constant} \quad \parallel \quad \ln \frac{1}{\sigma\sqrt{2\pi}}} - \frac{1}{2\sigma^2} \underbrace{\int_{-\infty}^{+\infty} x^2 f(x) dx}_{\parallel \quad E[X^2] \quad \parallel \quad \text{Var}[X] + (E[X])^2 \quad \parallel \quad \sigma^2} \right]$$

$$= \ln(\sigma\sqrt{2\pi}) + \frac{1}{2}$$

$$= \frac{1}{2} \ln(2\pi e \sigma^2).$$

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Problem 4 :

(a). $X \sim N(0, 1) \Rightarrow$ the PDF of X : $f_X(x) = \frac{1}{\sqrt{2\pi}} \exp\left(-\frac{x^2}{2}\right)$,
 $\forall x \in \mathbb{R}$.

Since $\text{Var}[X] = E[X^2] - (E[X])^2$ and $E[X] = 0$,

all we need is to show that $E[X^2] = 1$.

$$E[X^2] = \int_{-\infty}^{+\infty} x^2 \cdot f_X(x) dx$$

$$= \int_{-\infty}^{+\infty} x \cdot \frac{x}{\sqrt{2\pi}} \exp\left(-\frac{x^2}{2}\right) dx$$

$$= \left. x \cdot -\frac{1}{\sqrt{2\pi}} \exp\left(-\frac{x^2}{2}\right) \right|_{-\infty}^{+\infty}$$

$$- \int_{-\infty}^{+\infty} -\frac{1}{\sqrt{2\pi}} \exp\left(-\frac{x^2}{2}\right) dx$$

$$= \int_{-\infty}^{+\infty} \underbrace{\frac{1}{\sqrt{2\pi}} \exp\left(-\frac{x^2}{2}\right)}_{\text{PDF of } X} dx = 1$$

Recall: Integration by parts

$$\int u dv = uv - \int v du$$

$$\frac{d\left(\frac{-1}{\sqrt{2\pi}} \exp\left(-\frac{x^2}{2}\right)\right)}{dx} = \frac{x}{\sqrt{2\pi}} \exp\left(-\frac{x^2}{2}\right)$$

(b). $X \sim \text{Exp}(\lambda)$, then $f_X(x) = \begin{cases} \lambda e^{-\lambda x}, & \text{if } x > 0 \\ 0, & \text{else} \end{cases}$

$$E[X^2] = \int_0^{\infty} x^2 \lambda e^{-\lambda x} dx$$

$$= \left(x^2 \cdot (-e^{-\lambda x}) \right) \Big|_0^{\infty} = 0$$

$$- \int_0^{\infty} -e^{-\lambda x} \cdot 2x dx$$

$$= \frac{2}{\lambda} \int_0^{\infty} \lambda e^{-\lambda x} \cdot x dx$$

$E[X] = \frac{1}{\lambda}$

$$= \frac{2}{\lambda^2}$$

$$\text{Var}[X] = E[X^2] - (E[X])^2$$

$$= \frac{2}{\lambda^2} - \left(\frac{1}{\lambda}\right)^2$$

$$= \frac{1}{\lambda^2}$$

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Recall: integration by parts

$$\int u dv = uv - \int v du$$

Problem 5:

(a).

$$\int_{-\infty}^{+\infty} f(x) dx = \int_{-\infty}^{+\infty} C \cdot \exp(-|x|) dx$$

$$= 2 \cdot \int_0^{\infty} C \cdot \exp(-x) dx$$

$$= 2C \cdot (-\exp(-x)) \Big|_0^{\infty} = 2C = 1.$$

Hence, $C = \frac{1}{2}$

(b). $E[X^{2n}] = \int_{-\infty}^{+\infty} x^{2n} \cdot \frac{1}{2} \exp(-|x|) dx$

$$= \cancel{2} \cdot \int_0^{\infty} x^{2n} \cdot \cancel{\frac{1}{2}} \exp(-x) dx$$

$$= \int_0^{\infty} x^{2n} \exp(-x) dx = (2n)!$$

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by Lemma 1 in the next page

Lemma 1: $\int_0^{\infty} x^m \exp(-x) dx = m!, \text{ for all } m \in \mathbb{N}$

pf = Prove by induction

$$\begin{aligned} \underline{m=1}: \quad \int_0^{\infty} x \exp(-x) dx &= \underbrace{\left. x \cdot -\exp(-x) \right|_0^{\infty}}_{\substack{\text{integration} \\ \text{by parts}}} - \int_0^{\infty} 1 \cdot -\exp(-x) dx \\ &= \int_0^{\infty} \exp(-x) dx = -\exp(-x) \Big|_0^{\infty} = 1. \end{aligned}$$

Assume for $m=k$, we have $\int_0^{\infty} x^k \exp(-x) dx = k!$

Then, for $m=k+1$,

$$\begin{aligned} \int_0^{\infty} x^{k+1} \exp(-x) dx &= \underbrace{\left. x^{k+1} \cdot -\exp(-x) \right|_0^{\infty}}_{\substack{\text{integration} \\ \text{by parts}}} - \int_0^{\infty} (k+1) \cdot x^k \cdot -\exp(-x) dx \\ &= (k+1) \cdot \int_0^{\infty} x^k \cdot \exp(-x) dx = (k+1) \cdot k! = (k+1)! \end{aligned}$$

(Conti).

As we found that $E[X^{2n}] = (2n)!$, we know the existence of $E[|X^{2n+2}|]$ implies that $E[|X^{2n+1}|]$ also exists, for all $n \in \mathbb{N}$.

Now, we can use the symmetry and show that for any $n \in \mathbb{N}$,

$$E[X^{2n+1}] = \int_{-\infty}^{+\infty} x^{2n+1} \cdot \frac{1}{2} \exp(-|x|) dx$$

$$= \int_0^{\infty} x^{2n+1} \cdot \frac{1}{2} \exp(-x) dx + \int_{-\infty}^0 x^{2n+1} \cdot \frac{1}{2} \cdot \exp(x) dx$$

$$= 0$$

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