DCP 1206: Probability Lecture 24 — Strong Law of Large Numbers and Central Limit Theorem

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Announcements

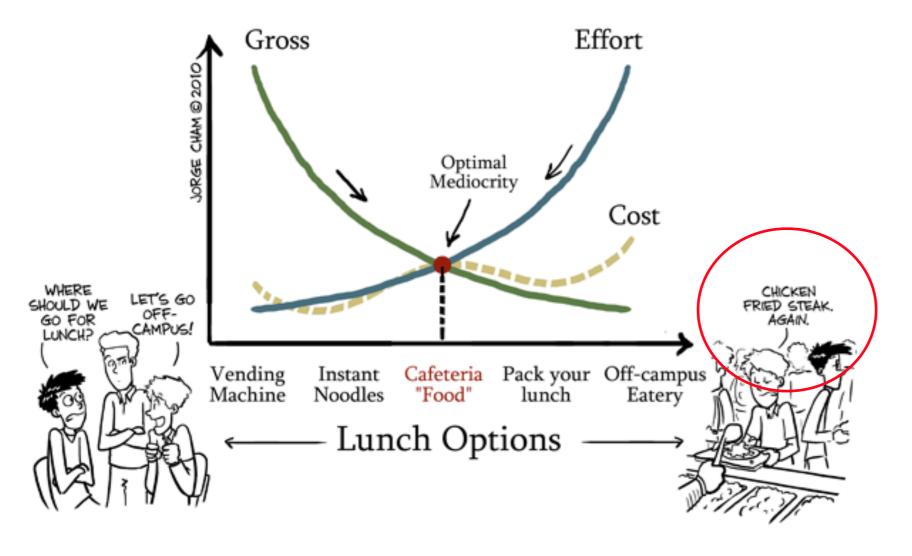
- Class on 12/25 (Wed): 10:10am-11:40am
- Class on 12/13 (Fri): 3:30pm-4:30pm (10min extension)
- Class on 12/20 (Fri): 3:30pm-4:30pm (10min extension)

PHD Comics

WWW.PHDCOMICS.COM

The Cafeteria Potential Well

Why you end up eating there almost every day.





Jorge Cham



Quick Review

Convergence in Probability, But Not Almost Surely

- Example: Let X be a continuous uniform r.v. on (0,1)
 - Consider a sequence of r.v.s X_1, X_2, \cdots as follows:

$$X_{1} = \mathbb{I}\{X \in [0,1]\}$$

$$X_{2} = \mathbb{I}\{X \in [0,\frac{1}{2}]\}$$

$$X_{3} = \mathbb{I}\{X \in [\frac{1}{2},1]\}$$

$$X_{4} = \mathbb{I}\{X \in [0,\frac{1}{3}]\}$$

$$X_{5} = \mathbb{I}\{X \in [\frac{1}{3},\frac{2}{3}]\}$$

$$X_{6} = \mathbb{I}\{X \in [\frac{1}{3},1]\}$$

$$X_{7} = \mathbb{I}\{X \in [\frac{1}{3},\frac{2}{3}]\}$$

$$X_{8} = \mathbb{I}\{X \in [\frac{1}{3},1]\}$$

$$X_{9} = \mathbb{I}\{X \in [\frac{1}{3},\frac{2}{3}]\}$$

$$X_{1} = \mathbb{I}\{X \in [0,\frac{1}{3}]\}$$

$$X_{2} = \mathbb{I}\{X \in [\frac{1}{3},\frac{2}{3}]\}$$

$$X_{3} = \mathbb{I}\{X \in [\frac{1}{3},\frac{2}{3}]\}$$

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$$X_{6} = \mathbb{I}\{X \in [\frac{1}{3},\frac{2}{3}]\}$$

$$X_{7} = \mathbb{I}\{X \in [\frac{1}{3},\frac{2}{3}]\}$$

$$X_{8} = \mathbb{I}\{X \in [\frac{1}{3},\frac{2}{3}]\}$$

Question: Do we have $\lim_{n\to\infty} P(\{\omega: |X_n(\omega)-0|>\varepsilon\})=0$?

Convergence in Probability, But Not Almost Surely (Cont.)

- \triangleright Example: Let X be a continuous uniform r.v. on (0,1)
 - Consider a sequence of r.v.s X_1, X_2, \cdots as follows:

$$X_1(\omega) = \mathbb{I}\{X(\omega) \in [0,1]\}$$

$$X_2(\omega) = \mathbb{I}\{X(\omega) \in [0,\frac{1}{2}]\} \quad X_3(\omega) = \mathbb{I}\{X(\omega) \in [\frac{1}{2},1]\}$$

$$X_4(\omega) = \mathbb{I}\{X(\omega) \in [0,\frac{1}{3}]\} \quad X_5(\omega) = \mathbb{I}\{X(\omega) \in [\frac{1}{3},\frac{2}{3}]\} \quad X_6(\omega) = \mathbb{I}\{X(\omega) \in [\frac{2}{3},1]\}$$

$$\dots$$

$$Question: Do we have $P(\{\omega: \lim_{n\to\infty} X_n(\omega) = 0\}) = \mathbb{X}$?$$

 $\chi_1(\omega) = 1$

$$\chi_{2}(\omega) = 1 \qquad \chi_{3}(\omega) = 0$$

$$\chi_{4}(\omega) = 1 \qquad \chi_{5}(\omega) = 0 \qquad \chi_{6}(\omega) = 0$$

This Lecture

1. Strong Law of Large Numbers (SLLN)

2. Central Limit Theorem (CLT)

Reading material: Chapter 11.4-11.5

1. Strong Law of Large Numbers (SLLN)

WLLN vs SLLN

The Weak Law of Large Numbers (Khinchin's Law): Let X_1, \dots, X_n be a sequence of independent and identically distributed (i.i.d.) random variables with mean μ . Define $S_n = (X_1 + \cdots + X_n)$. Then, for every $\varepsilon > 0$, we have

$$\lim_{n\to\infty} P\left(\{\omega: \left|\frac{S_n(\omega)}{n} - \mu\right| \ge \varepsilon\right) = 0$$

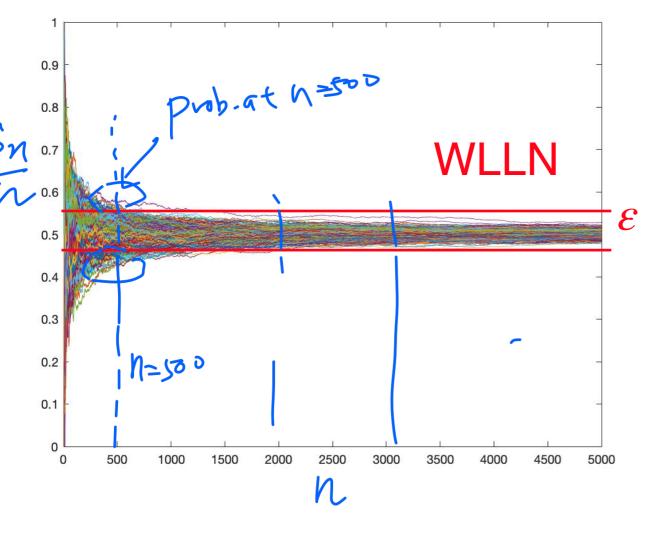
- The Strong Law of Large Numbers: Let X_1, \dots, X_n be a sequence of independent and identically distributed (i.i.d.) random variables with mean μ . Define $S_n = (X_1 + \cdots + X_n)$.

Then, we have

Since
$$S_n(\omega) = \mu$$
 $P(\{\omega : \lim_{n \to \infty} \frac{S_n(\omega)}{n} = \mu\}) = 1$ Convergence)

Visualization of WLLN and SLLN

Example: $X_i \sim \text{Bernoulli}(0.5)$ and $S_n = X_1 + \cdots + X_n$



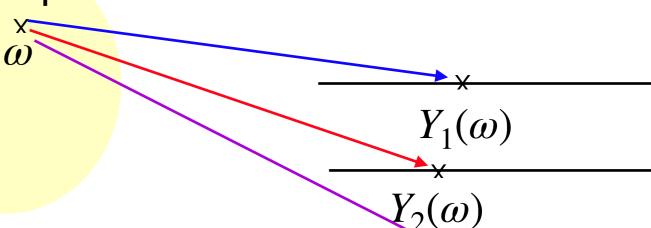
$$\lim_{n\to\infty} P\left(\left\{\omega: \left|\frac{S_n(\omega)}{n} - \mu\right| \ge \varepsilon\right) = 0 \qquad P\left(\left\{\omega: \lim_{n\to\infty} \frac{S_n(\omega)}{n} = \mu\right\}\right) = 1$$

$$P\Big(\Big\{\omega: \lim_{n\to\infty} \frac{S_n(\omega)}{n} = \mu\Big\}\Big) = 1$$

How to Interpret SLLN?

- Let X_1, X_2, \cdots be a sequence of i.i.d. random variables with mean μ
- Define $Y_n = (X_1 + X_2 \dots, + X_n)/n$
- $\sum_{n \to \infty} \frac{S_n(\omega)}{n} = \mu \}) = 1$

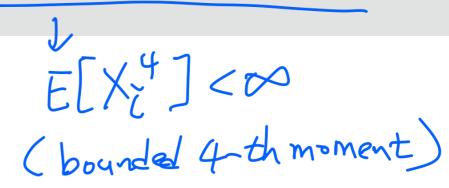
Sample space $\xrightarrow{Y_n(\cdot)}$ Real Numbers



• Question: What is an " ω "?

$$Y_n(\omega)$$

How to Prove SLLN (Under a Mild Condition)?



- 1. Borel-Cantelli Lemma
- 2. A Bound for the 4-th Moment Condition
- 3. Markov's Inequality

1. Borel-Cantelli Lemma

Recall: HW1, Problem 4

Problem 4 (Continuity of Probability Function and Probability Axioms) (8+8=16 points)

Consider an infinite sequence of coin tosses. The probability of having a head at the *i*-th toss is p_i , with $p_i \in [0,1]$ (Note: different tosses may not be independent and can potentially have different head probabilities). We use I to denote the event of having infinite number of heads.

- (a) Suppose $\sum_{i=1}^{\infty} p_i$ is finite. Show that P(I) = 0. (Hint: Define $A_n := \{\text{the } n\text{-th toss is a head}\}$. Then, $B_k := \bigcup_{n=k}^{\infty} A_n$ is the event that there is at least one head after the k-th toss (including the k-th toss). The event I (i.e. we observe infinitely many heads) is equivalent to saying that B_k occurs, for every $k = 1, 2, 3, \cdots$. Therefore, $I = \bigcap_{k=1}^{\infty} \bigcup_{n=k}^{\infty} A_n$. Consider the continuity of probability function for $\bigcap_{k=1}^{\infty} \bigcup_{i=k}^{\infty} A_n$ and then apply the union bound)
- Borel-Cantelli Lemma: Let $\{A_n\}$ be any sequence of events. If

$$\sum_{n=1}^{\infty} P(A_n) \geqslant \infty, \text{ then we have}$$

$$P\Big(A_n \text{ occurs infinitely often}\Big) = P(\bigcap_{n \geq 1} \bigcup_{m \geq n} A_n) = 0$$

Review: Proof of Borel-Cantelli Lemma

ightharpoonup Borel-Cantelli Lemma: Let $\{A_n\}$ be any sequence of events. If

$$\sum_{n=1}^{\infty} P(A_n) < \infty, \text{ then we have}$$

$$P\Big(A_n \text{ occurs infinitely often}\Big) = P(\bigcap \bigcup A_n) = 0$$

Proof:
$$P(\bigcap_{m=1}^{\infty} A_n) = P(\lim_{k \to \infty} A_n)$$

$$=\lim_{k\to\infty} \lim_{m=1}^{\infty} \lim_{n=k}^{\infty} A_n$$

$$=\lim_{k\to\infty} \lim_{n=k}^{\infty} \lim_{n=k}^{\infty} \lim_{n=k}^{\infty} P(A_n) = 0$$

$$=\lim_{k\to\infty} \lim_{n=k}^{\infty} \lim_{n=k}^{\infty} P(A_n) = 0$$

2. A Bound For 4-th Moment of $\frac{5n}{n}$

▶ A Bound on 4-th Moment: Let X_1, \dots, X_n be a sequence of independent and identically distributed (i.i.d.) random variables with mean μ and $E[X_1^4] < \infty$. Define $S_n = (X_1 + \dots + X_n)$. Then, there exists a constant $K < \infty$ such that

$$E[(S_n - n\mu)^4] \le Kn^2$$

Question: How about
$$E[(\frac{S_n}{n} - \mu)^4] \le ?$$
 $\frac{Kn^2}{N^4} = \frac{K}{N^2}$

Proof: A Bound For 4-th Moment

- Given: $S_n = (X_1 + \dots + X_n)$ and $E[X_1^4] < \infty$
- $\operatorname{Want:} E[(S_n n\mu)^4] \le Kn^2$
- Proof: For simplicity, let $Z_i = X_i \mu$

Put Everything Together: Proof of SLLN

$$P\left(\frac{|S_{n}-\mu| \geq n^{-1}/|E|}{n}\right) = P\left(\frac{|S_{n}-\mu|^{4}}{n}\right)^{\frac{1}{5}} = \frac{|S_{n}-\mu|^{4}}{|S_{n}-\mu|} = \frac{|S_{n}-\mu|^{4}}{|S_{n}-\mu|^{4}} = \frac{|S_$$

2. Central Limit Theorem (CLT)

Beyond SLLN

- ► The Strong Law of Large Numbers: Let X_1, \dots, X_n be a sequence of independent and identically distributed (i.i.d.) random variables with mean μ . Define $S_n = (X_1 + \dots + X_n)$. Then, we have $P\Big(\Big\{\omega: \lim_{n\to\infty} \frac{S_n(\omega)}{n} = \mu\Big\}\Big) = 1$
- Question: What does SLLN say about $S_n(\omega)$?

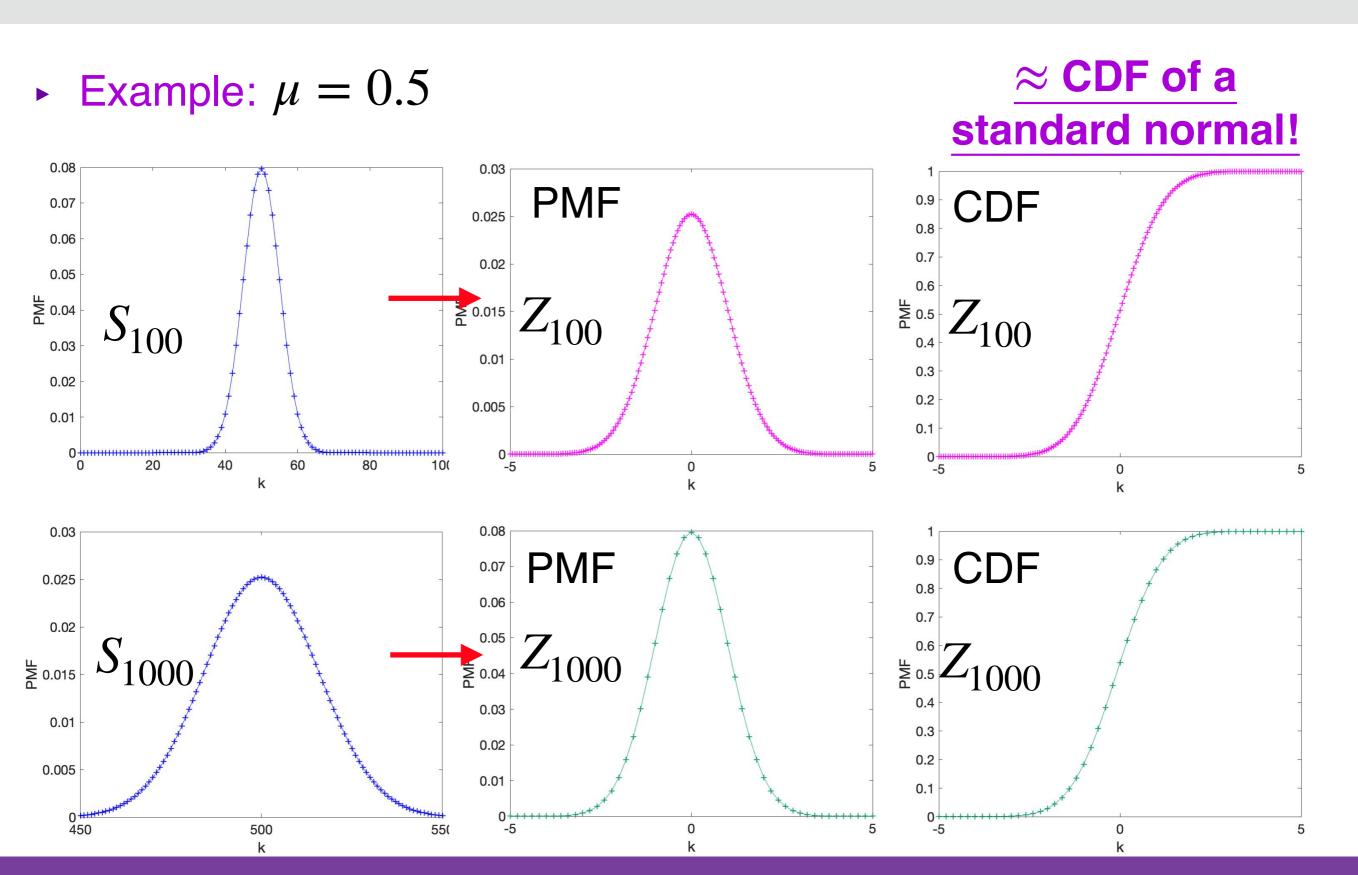
• Question: Do we have $S_n(\omega) = n\mu + o(n)$?

Review (Lecture 11): Binomial and Normal

- Example: X_1, X_2, \cdots are i.i.d. Bernoulli r.v.s with mean μ and variance $\sigma^2 = \mu(1-\mu)$
 - Define $S_n = X_1 + X_2 + \dots + X_n$
 - Question: What type of r.v. is S_n ? $E[S_n] = ? Var[S_n] = ?$

Question: How to find the distribution of $\frac{S_n - n\mu}{\sigma\sqrt{n}}$?

Method 1: Plotting $Z_n = (S_n - n\mu)/(\sigma\sqrt{n})$



Method 2: MGF

- Idea: Suppose we find the MGF of $\frac{S_n n\mu}{\sigma\sqrt{n}}$ for $n \to \infty$
 - Question: Can we find its distribution?

Levy Continuity Theorem: Let $V_1, V_2 \cdots$ be a sequence of random variables with CDFs F_1, F_2, \cdots and MGFs $M_{V_1}(t), M_{V_2}(t) \cdots$. Let V be a random variable with CDF F and MGF $M_V(t)$. If for every $t \in \mathbb{R}$, $\lim_{n \to \infty} M_{V_n}(t) = M_V(t)$, then the CDFs F_n converge to F.

Method 2: MGF (Cont.)

- Example: X_1, X_2, \cdots are i.i.d. r.v.s with mean μ and variance σ^2
 - $\qquad \qquad \textbf{Define } S_n = X_1 + X_2 + \cdots + X_n \text{ and } Y_i = X_i \mu$
 - Question: $E[Y_i] = ____? Var[Y_i] = ____?$
 - Question: What is the MGF of $\frac{S_n n\mu}{\sigma\sqrt{n}}$ (in terms of MGF of Y_i)?

Method 2: MGF (Cont.)

Question: When $n \to \infty$, what is the MGF of $\frac{S_n - n\mu}{\sigma \sqrt{n}}$?

Central Limit Theorem (Formally)

• Central Limit Theorem (CLT): Let X_1, \dots, X_n be a sequence of independent and identically distributed (i.i.d.) random variables with mean μ and variance σ^2 . Define $S_n = (X_1 + \dots + X_n)$. Then, we have

$$\lim_{n \to \infty} P\left(\frac{S_n - n\mu}{\sigma\sqrt{n}} \le z\right) = \Phi(z)$$

where $\Phi(z)$ is the CDF of standard normal

Next Lecture

- Parameter Estimation
 - Maximum Likelihood Estimation (MLE)
 - Maximum a Posteriori (MAP)

1-Minute Summary

1. Strong Law of Large Numbers (SLLN)

. SLLN:
$$P\left(\left\{\omega: \lim_{n\to\infty} \frac{S_n(\omega)}{n} = \mu\right\}\right) = 1$$

 Combo of 3 tools: Borel-Cantelli Lemma / Bound for 4th moment / Markov's inequality

2. Central Limit Theorem (CLT)

. CLT:
$$\lim_{n \to \infty} P\left(\frac{S_n - n\mu}{\sigma\sqrt{n}} \le z\right) = \Phi(z)$$

Proof by MGF