

Problem 1

(a) Prove by induction:

$$\underline{N=2}: P(A_1 \cup A_2) = P(A_1) + P(A_2) - \underbrace{P(A_1 \cap A_2)}_{\geq 0, \text{ by Axiom 1}} \leq \sum_{n=1}^2 P(A_n) \quad \text{--- (1)}$$

Suppose when $N=K$, the following holds:

$$P\left(\bigcup_{n=1}^K A_n\right) \leq \sum_{n=1}^K P(A_n) \quad \text{--- (2)}$$

For $N=K+1$:

$$\begin{aligned} P\left(\bigcup_{n=1}^{K+1} A_n\right) &= P\left(\left(\bigcup_{n=1}^K A_n\right) \cup A_{K+1}\right) \stackrel{\text{by (1)}}{\leq} P\left(\bigcup_{n=1}^K A_n\right) + P(A_{K+1}) \\ &\stackrel{\text{by (2)}}{\leq} \sum_{n=1}^{K+1} P(A_n) \end{aligned}$$

Therefore, by induction, we know $P\left(\bigcup_{n=1}^N A_n\right) \leq \sum_{n=1}^N P(A_n)$ holds for all $N \in \mathbb{N}$.

(b)

Method 1: It is easy to check that

$$P(\Omega) = P(\{1, 2, 3, 4, 5\}) \leq P(\{1, 5\}) + P(\{1, 2, 4\}) + P(\{3, 4\}) = 0.95$$

Hence, there is no valid assignment.

Method 2: Try to find some other contradiction

$$\text{Since } P(\{1, 5\}) = 0.3 \text{ and } P(\{3, 4\}) = 0.2, \text{ then } P(\{2\}) = 0.5$$

$$\text{However, } P(\{1, 2, 4\}) = P(\{2\}) + P(\{1, 4\}) = 0.45$$

contradiction.

Therefore, there is no valid probability assignment.

Problem 2

(a) (i): $\left(\bigcup_{n=1}^{\infty} S_n\right)^c = \bigcap_{n=1}^{\infty} S_n^c \Rightarrow$ prove this by showing $\left(\bigcup_{n=1}^{\infty} S_n\right)^c \subseteq \bigcap_{n=1}^{\infty} S_n^c$ and $\left(\bigcup_{n=1}^{\infty} S_n\right)^c \supseteq \bigcap_{n=1}^{\infty} S_n^c$

- Suppose that $x \in \left(\bigcup_{n=1}^{\infty} S_n\right)^c$. Then, $x \notin \bigcup_{n=1}^{\infty} S_n$, which implies that we must have $x \notin S_n$, for every n . Hence, $x \in \bigcap_{n=1}^{\infty} S_n^c \Rightarrow \left(\bigcup_{n=1}^{\infty} S_n\right)^c \subseteq \bigcap_{n=1}^{\infty} S_n^c$
- The converse inclusion can be established by reversing the above argument.

(ii) $\left(\bigcap_{n=1}^{\infty} S_n\right)^c = \bigcup_{n=1}^{\infty} S_n^c \Leftrightarrow \bigcap_{n=1}^{\infty} S_n = \left(\bigcup_{n=1}^{\infty} S_n^c\right)^c \Leftrightarrow \bigcap_{n=1}^{\infty} (S_n^c)^c = \left(\bigcup_{n=1}^{\infty} (S_n^c)^c\right)^c$

It is easy to verify that this is true based on (i).

(b). Choose $B_1 = A_1$

$$B_2 = A_2 - A_1 = A_2 \cap A_1^c$$

$$B_n = A_n - \left(\bigcup_{i=1}^{n-1} A_i\right) = A_n \cap \left(\bigcup_{i=1}^{n-1} A_i\right)^c \stackrel{\text{De Morgan's Law}}{=} A_n \cap \left(\bigcap_{i=1}^{n-1} A_i^c\right)$$

Check: (i) For any $j \neq k$, we have $B_j \cap B_k = \emptyset$:

Without loss of generality, let $j < k$.

$$\text{Then, } B_j = A_j \cap \left(\bigcap_{i=1}^{j-1} A_i^c\right) \text{ and } B_k = A_k \cap \left(\bigcap_{i=1}^{k-1} A_i^c\right)$$

It is easy to verify that $B_j \cap B_k = \emptyset$, for all pairs $j \neq k$.

(ii) Check $\bigcup_{i=1}^n B_i = \bigcup_{i=1}^n A_i$, for all n

Prove by induction: We already know $B_1 = A_1$

Suppose $\bigcup_{i=1}^k B_i = \bigcup_{i=1}^k A_i$.

$$\text{Then } \bigcup_{i=1}^{k+1} B_i = B_{k+1} \cup \bigcup_{i=1}^k B_i$$

$$= B_{k+1} \cup \bigcup_{i=1}^k A_i$$

$$= (A_{k+1} \cap (\bigcap_{i=1}^k A_i^c)) \cup \bigcup_{i=1}^k A_i$$

$$= (A_{k+1} \cap (\bigcup_{i=1}^k A_i)^c) \cup \bigcup_{i=1}^k A_i$$

$$= \bigcup_{i=1}^{k+1} A_i$$

Therefore, we conclude that $\bigcup_{i=1}^n B_i = \bigcup_{i=1}^n A_i$, for all $n \in \mathbb{N}$.

Problem 3

(a). Prove by contradiction =

Suppose there are countably infinite real numbers in $(0,1)$.

We denote these numbers as x_1, x_2, x_3, \dots

For each x_i , we express the number in decimal expansion, i.e.

$$x_i = 0. a_i^{(1)} a_i^{(2)} a_i^{(3)} a_i^{(4)} \dots$$

Now we construct a real number $y = 0. b^{(1)} b^{(2)} b^{(3)} b^{(4)} \dots$,

$$\text{where } b^{(k)} = \begin{cases} 1, & \text{if } a_i^{(k)} = 2 \\ 2, & \text{otherwise} \end{cases} \Rightarrow y \in (0,1)$$

Therefore, we know $y \neq x_i$, for all $i \Rightarrow$ Contradiction!

Hence, there are uncountably infinite real numbers in $(0,1)$.

(b). We know all irrational numbers in $(0,1)$ $= \mathbb{R}_{(0,1)} - \mathbb{Q}_{(0,1)}$
 \Downarrow denoted by $\mathbb{I}_{(0,1)}$

We show that $\mathbb{I}_{(0,1)}$ is an uncountably infinite set by contradiction =

Suppose $\mathbb{I}_{(0,1)}$ is countably infinite and denote $\mathbb{I}_{(0,1)} = \{y_1, y_2, y_3, \dots\}$

We already know $\mathbb{Q}_{(0,1)}$ is countable, i.e. $\mathbb{Q}_{(0,1)} = \{z_1, z_2, z_3, \dots\}$

Then, we can write $\mathbb{R}_{(0,1)} = \{y_1, z_1, y_2, z_2, y_3, z_3, \dots\}$ and $\mathbb{R}_{(0,1)}$ is also countable
 \Rightarrow contradiction!

Problem 4:

(a) Define A_n to be the event that the n -th toss is a head.

It is easy to verify that $I = \bigcap_{k=1}^{\infty} \bigcup_{n=k}^{\infty} A_n$.

$$\text{Then, } P(I) = P\left(\bigcap_{k=1}^{\infty} \bigcup_{n=k}^{\infty} A_n\right) = P\left(\lim_{m \rightarrow \infty} \bigcap_{k=1}^m \bigcup_{n=k}^{\infty} A_n\right)$$

$$\begin{array}{c} \text{by continuity} \searrow \\ = \lim_{m \rightarrow \infty} P\left(\bigcap_{k=1}^m \bigcup_{n=k}^{\infty} A_n\right) \end{array}$$

$$\begin{array}{c} \text{union bound} \searrow \\ = \lim_{m \rightarrow \infty} P\left(\bigcup_{n=m}^{\infty} A_n\right) \end{array}$$

$$\begin{array}{c} \searrow \\ \leq \lim_{m \rightarrow \infty} \sum_{n=m}^{\infty} P(A_n) \\ \sum_{i=1}^{\infty} P_i \text{ is finite} \searrow \\ = 0 \end{array} \quad \square$$

(b) Define N to be the event that there is no head at all

Also, define E_n to be the event that there are only a finite number of heads and the last head occurs at the n -th toss.

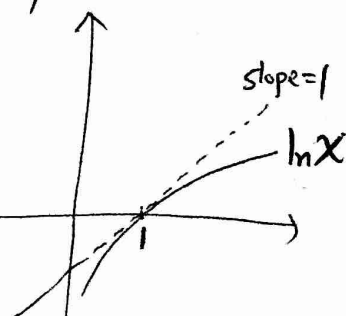
Then, it is easy to verify that $I^c = \left(\bigcup_{n \geq 1} E_n\right) \cup N$ and that E_n 's and N are mutually exclusive.

By the third probability axiom, $P(I^c) = \left(\sum_{n \geq 1} P(E_n)\right) + P(N)$

Note that $P(N) \leq \prod_{i=1}^M (1 - P_i)$ for all M .

$$\Rightarrow \ln P(N) \leq \sum_{i=1}^M \ln(1 - P_i) \leq \sum_{i=1}^M -P_i, \text{ for all } M$$

As $\sum_{i=1}^{\infty} P_i = \infty$, then $\ln P(N) = -\infty$ and equivalently $P(N) = 0$.



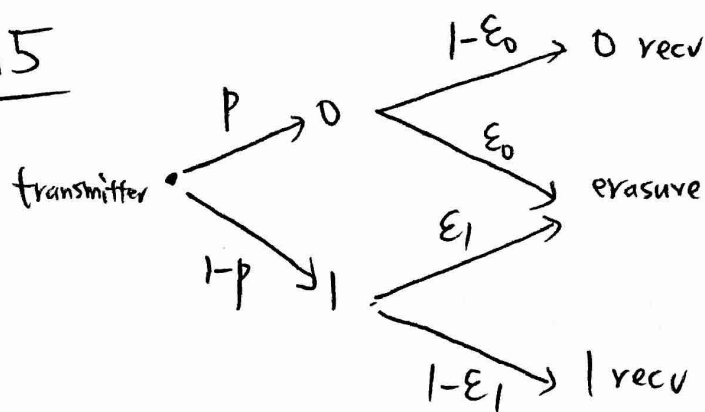
Similarly, We can show that $P(E_n) = 0$, for all n .

$$\text{Hence, } P(I^c) = \left(\sum_{n \geq 1} P(E_n) \right) + P(N) = 0.$$

Therefore, we conclude that $P(I) = 1$.

□

Problem 5



(a) $P(\text{a transmitted bit is erased}) \stackrel{\text{total probability theorem}}{=} P(0 \text{ sent}) \cdot P(\text{erased} | 0 \text{ sent}) + P(1 \text{ sent}) \cdot P(\text{erased} | 1 \text{ sent})$

$$= p \cdot \epsilon_0 + (1-p) \cdot \epsilon_1$$

(b) $P(\text{at most 1 bit erased in 3 sent bits})$

$$= \sum_{\substack{x_1 \in \{0,1\} \\ x_2 \in \{0,1\} \\ x_3 \in \{0,1\}}} P(\text{at most 1 bit erased and } x_1 x_2 x_3 \text{ sent})$$

$$= \sum_{\substack{x_1 \in \{0,1\} \\ x_2 \in \{0,1\} \\ x_3 \in \{0,1\}}} \left\{ P(x_1 \text{ is erased and } x_1 x_2 x_3 \text{ sent}) + P(x_2 \text{ is erased and } x_1 x_2 x_3 \text{ sent}) \right. \\ \left. + P(x_3 \text{ is erased and } x_1 x_2 x_3 \text{ sent}) + P(\text{no erasure and } x_1 x_2 x_3 \text{ sent}) \right\}$$

$$= \sum_{\substack{x_1 \in \{0,1\} \\ x_2 \in \{0,1\} \\ x_3 \in \{0,1\}}} \left\{ \left(\prod_{i=1}^3 (1-p)^{x_i} p^{1-x_i} \right) \left[\epsilon_{x_1} (1-\epsilon_{x_2})(1-\epsilon_{x_3}) + (1-\epsilon_{x_1}) \epsilon_{x_2} (1-\epsilon_{x_3}) \right. \right. \\ \left. \left. + (1-\epsilon_{x_1})(1-\epsilon_{x_2}) \epsilon_{x_3} + (1-\epsilon_{x_1})(1-\epsilon_{x_2})(1-\epsilon_{x_3}) \right] \right\}$$

(c).

$$P(10 \text{ sent} | 2 \text{ erasures recd})$$

Bayes' rule
↓
=

$$P(10 \text{ sent and both are erased})$$

$$\sum_{\substack{x_1 \in \{0,1\} \\ x_2 \in \{0,1\}}} P(x_1 x_2 \text{ sent and both are erased})$$

$$= \frac{(1-p) \cdot p \cdot \epsilon_1 \cdot \epsilon_0}{2(1-p) \cdot p \cdot \epsilon_1 \cdot \epsilon_0 + (1-p)^2 \epsilon_1^2 + p^2 \epsilon_0^2}$$

(d). As the first two received bits are 01, then the first two sent bits must be 01.

For the last two bits, we can follow the similar argument as (c):

$$P(00 \text{ sent} | 2 \text{ erasures}) = \frac{p^2 \epsilon_0^2}{2(1-p) \cdot p \cdot \epsilon_1 \epsilon_0 + (1-p)^2 \epsilon_1^2 + p^2 \epsilon_0^2} = \frac{225}{1521}$$

$$\underbrace{P(01 \text{ sent} | 2 \text{ erasures})}_{//} = \frac{p(1-p) \cdot \epsilon_0 \epsilon_1}{2(1-p) \cdot p \cdot \epsilon_1 \epsilon_0 + (1-p)^2 \epsilon_1^2 + p^2 \epsilon_0^2} = \frac{360}{1521}$$

$$P(10 \text{ sent} | 2 \text{ erasures})$$

$$P(11 \text{ sent} | 2 \text{ erasures}) = \frac{(1-p)^2 \epsilon_1^2}{2(1-p) \cdot p \cdot \epsilon_1 \epsilon_0 + (1-p)^2 \epsilon_1^2 + p^2 \epsilon_0^2} = \frac{576}{1521}$$

Therefore, 0111 is the most probable sequence of transmitted bits.

□

Problem 6

$$\begin{aligned} \text{(a). } P(A_1|B) &= \frac{P(A_1) \cdot P(B|A_1)}{P(B)} = \frac{P(A_1) \cdot P(B|A_1)}{P(A_1) \cdot P(B|A_1) + P(A_2) \cdot P(B|A_2) + P(A_3) \cdot P(B|A_3)} \\ &= \frac{\frac{1}{3} \times 0.3}{\frac{1}{3} \times 0.3 + \frac{1}{3} \times 0.5 + \frac{1}{3} \times 0.7} = \frac{1}{5} \end{aligned}$$

Similarly, $P(A_2|B) = \frac{1}{3}$, $P(A_3|B) = \frac{7}{15}$.

$$\begin{aligned} \text{(b). } P(A_1|C) &= \frac{P(A_1) \cdot P(C|A_1)}{P(C)} = \frac{P(A_1) \cdot P(C|A_1)}{P(A_1) \cdot P(C|A_1) + P(A_2) \cdot P(C|A_2) + P(A_3) \cdot P(C|A_3)} \\ &= \frac{\frac{1}{3} \times (0.3)^8 \times (0.7)^2}{\frac{1}{3} \times (0.3)^8 \times (0.7)^2 + \frac{1}{3} \times (0.5)^8 \times (0.5)^2 + \frac{1}{3} \times (0.7)^8 \times (0.3)^2} \\ &\approx 0.52\% \end{aligned}$$

Similarly, by Bayes' rule, we have:

$$P(A_2|C) \approx 15.76\%$$

\Rightarrow Hence, the most probable value for θ is 0.7

$$P(A_3|C) \approx 83.72\%$$

(c). Suppose $P(A_1) = \frac{2}{5}$, $P(A_2) = \frac{2}{5}$, $P(A_3) = \frac{1}{5}$.

$$P(A_1|C) = \frac{\frac{2}{5} \times (0.3)^8 \times (0.7)^2}{\frac{2}{5} \times (0.3)^8 \times (0.7)^2 + \frac{2}{5} \times (0.5)^8 \times (0.7)^2 + \frac{1}{5} \times (0.7)^8 \times (0.3)^2} \approx 0.89\%$$

$$P(A_2|C) \approx 27.11\%$$

\Rightarrow Hence, the most probable value for θ is 0.7

$$P(A_3|C) \approx 72.00\%$$

Problem 7

$$(a). \quad (1+x)^{2n} = (1+x)^n (1+x)^n$$

The coefficient of the term x^n in $(1+x)^{2n} = C_n^{2n}$.

$$,, \quad ,, \quad x^n \text{ in } (1+x)^n (1+x)^n$$

$$= \sum_{i=0}^n \left(\text{the coefficient of the term } x^i \text{ in } (1+x)^n \right) \cdot \left(\text{the coefficient of the term } x^{n-i} \text{ in } (1+x)^n \right)$$

$$= \sum_{i=0}^n C_i^n \cdot C_{n-i}^n$$

$$= \sum_{i=0}^n (C_i^n)^2$$

□

$$(b). \quad \text{Since } C_i^{n+i} = C_i^{n+i+1} - C_{i-1}^{n+i} \text{ for } i \geq 1,$$

$$\text{then } \sum_{i=0}^r C_i^{n+i} = C_0^n + \sum_{i=1}^r (C_i^{n+i+1} - C_{i-1}^{n+i})$$

$$= \underbrace{C_0^n}_{//} + (C_r^{n+r+1} - \underbrace{C_0^{n+1}}_{//}) = C_r^{n+r+1}$$

□