

4.3

Weak Law of Large Numbers

Let $\{X_n\}$ be a sequence of r.vs and let $\bar{X}_n = \frac{1}{n} \sum_{i=1}^n X_i$ be the mean of first n r.vs. The

weak laws deal with *limits of probabilities involving* \bar{X}_n . The strong laws deal with *probabilities involving limits of* \bar{X}_n .

Definition of Weak Law of Large Numbers

A sequence $\{X_n\}$ of r.vs is said to satisfy the **Weak Law of Large Numbers (WLLN)** if

$$\lim_{n \rightarrow \infty} P \left[\left| \frac{S_n}{n} - E \left(\frac{S_n}{n} \right) \right| < \epsilon \right] = 1$$

for any $\epsilon > 0$, where $S_n = \sum_{i=1}^n X_i$, i. e., $\frac{S_n}{n} \xrightarrow{P} E \left(\frac{S_n}{n} \right)$

Theorem1: Let $\{X_n\}$ be a sequence of r.vs and let $S_n = X_1 + \dots + X_n$ with $B_n = V(S_n) < \infty$. If $\frac{B_n}{n^2} \rightarrow 0$ as $n \rightarrow \infty$, then for any $\epsilon > 0$,

$$\lim_{n \rightarrow \infty} P \left[\left| \frac{S_n}{n} - E \left(\frac{S_n}{n} \right) \right| < \epsilon \right] = 1$$

i. e., $\{X_n\}$ satisfies WLLN.

Proof: On applying Chebychev's inequality to the variable $\frac{S_n}{n}$, we have

$$P \left[\left| \frac{S_n}{n} - E \left(\frac{S_n}{n} \right) \right| \geq \epsilon \right] \leq \frac{V \left(\frac{S_n}{n} \right)}{\epsilon^2} = \frac{V(S_n)}{n^2 \epsilon^2} = \frac{B_n}{n^2 \epsilon^2} \rightarrow 0$$

as $n \rightarrow \infty$. Thus,

$$\lim_{n \rightarrow \infty} P \left[\left| \frac{S_n}{n} - E \left(\frac{S_n}{n} \right) \right| \geq \epsilon \right] = 0 \Rightarrow \lim_{n \rightarrow \infty} P \left[\left| \frac{S_n}{n} - E \left(\frac{S_n}{n} \right) \right| < \epsilon \right] = 1$$

$\Rightarrow \{X_n\}$ satisfies WLLN.

Corollary 1: Let $\{X_n\}$ be a sequence of r.v.s, $\bar{X}_n = \frac{S_n}{n}$ and $\mu = E\left(\frac{S_n}{n}\right)$.

If $\frac{B_n}{n^2} \rightarrow 0$ as $n \rightarrow \infty$, then

$$\lim_{n \rightarrow \infty} P[\bar{X}_n \leq k] = \begin{cases} 0 & , \text{if } k < \mu \\ 1 & , \text{if } k > \mu \end{cases}$$

Proof: Since WLLN holds for $\{X_n\}$, we have

$$\lim_{n \rightarrow \infty} P[|\bar{X}_n - \mu| < \epsilon] = 1 \Rightarrow \lim_{n \rightarrow \infty} P[|\bar{X}_n - \mu| \geq \epsilon] = 0 \quad \dots (1)$$

Since $\{\bar{X}_n \leq \mu - \epsilon\} \subset \{|\bar{X}_n - \mu| \geq \epsilon\}$, we have

$$\begin{aligned} P(\bar{X}_n \leq \mu - \epsilon) &\leq P(|\bar{X}_n - \mu| \geq \epsilon) \\ \Rightarrow \lim_{n \rightarrow \infty} P(\bar{X}_n \leq \mu - \epsilon) &\leq \lim_{n \rightarrow \infty} P(|\bar{X}_n - \mu| \geq \epsilon) \\ \Rightarrow \lim_{n \rightarrow \infty} P(\bar{X}_n \leq \mu - \epsilon) &= 0 \\ \Rightarrow \lim_{n \rightarrow \infty} P(\bar{X}_n \leq k) &= 0, \text{ where } k = \mu - \epsilon, \text{ i.e., } k < \mu \text{ since } \epsilon > 0 \\ \Rightarrow \lim_{n \rightarrow \infty} P(\bar{X}_n \leq k) &= 0 \text{ if } k < \mu \end{aligned}$$

Further, $P(\bar{X}_n \leq \mu + \epsilon) + P(|\bar{X}_n - \mu| > \epsilon) \geq 1$, since the region is larger than sample space covered.

$$\begin{aligned} \Rightarrow \lim_{n \rightarrow \infty} P(\bar{X}_n \leq \mu + \epsilon) &\geq 1 \quad \left(\because \lim_{n \rightarrow \infty} P(|\bar{X}_n - \mu| > \epsilon) = 0 \right) \\ \Rightarrow \lim_{n \rightarrow \infty} P(\bar{X}_n \leq \mu + \epsilon) &= 1 \\ \Rightarrow \lim_{n \rightarrow \infty} P(\bar{X}_n \leq k) &= 1 \text{ where } k = \mu + \epsilon \text{ i.e., } k > \mu \text{ since } \epsilon > 0 \\ \Rightarrow \lim_{n \rightarrow \infty} P(\bar{X}_n \leq k) &= 1 \text{ if } k > \mu \end{aligned}$$

Thus, $\lim_{n \rightarrow \infty} P(\bar{X}_n \leq k) = \begin{cases} 0, & k < \mu \\ 1, & k > \mu \end{cases}$

Variations of the WLLN

The following are some special cases of Theorem 1 which are stated without proof.

Theorem 2: (Bernoulli's WLLN)

Let $\{X_n\}$ be a sequence of Bernoulli trials with probability of success equal to p . If S_n is the number of successes in n trials, then

$$\lim_{n \rightarrow \infty} P \left[\left| \frac{S_n - np}{n} \right| < \epsilon \right] = 1, \quad \forall \epsilon > 0$$

Theorem 3: (Khinchine's WLLN)

Let $\{X_n\}$ be a sequence of i.i.d.r.vs with $E(X_i) = \mu < \infty, i = 1, 2, \dots$, then the WLLN holds i.e.,

$$\lim_{n \rightarrow \infty} P \left[\left| \frac{S_n}{n} - \mu \right| > \epsilon \right] = 0$$

Theorem 4: (Bernstein's WLLN)

Let $\{X_n\}$ be a sequence of random variables for which $\text{var}(X_n) = \sigma_n^2 < k, \forall i$, where k is independent of n . If $\sigma_{ij} = \text{cov}(X_i, X_j) \rightarrow 0$ as $|i - j| \rightarrow \infty$ (Asymptotic uncorrelatedness) then the WLLN holds.

Example 1: Let $\{X_n\}$ be i.i.d.r.vs with mean μ and variance σ^2 , if

$$\frac{X_1^2 + X_2^2 + \dots + X_n^2}{n} \xrightarrow{P} c$$

as $n \rightarrow \infty$ for some constant $c(0 \leq c < \infty)$, then find c .

Solution: Here $E(X_i) = \mu$ and $V(X_i) = \sigma^2 \forall i$.

Let $S_n = X_1^2 + X_2^2 + \dots + X_n^2$. Then

$$E(S_n) = nE(X_1^2) \quad (\because X\text{s are i.i.d.r.vs})$$

$$= n \left[V(X_1) + (E(X_1))^2 \right]$$

$$\Rightarrow E(S_n) = n(\sigma^2 + \mu^2)$$

Since $E(X^2) = V(X) + (E(X))^2 = \sigma^2 + \mu^2$ exists for each X^2 in S_n , by Khinchine's WLLN, we have

$$\frac{S_n}{n} = \frac{X_1^2 + X_2^2 + \dots + X_n^2}{n} \quad E(X_1^2) = \mu^2 + \sigma^2$$

Thus, $c = \mu^2 + \sigma^2$.

Example 2: If the i.i.d. r.vs $X_k (k = 1, 2, \dots)$ assume the value $2^{r-2 \ln r}$ with probability $\frac{1}{2^r}$, examine if the WLLN holds for the sequence $\{X_k\}$.

Solution:

$$\begin{aligned} E(X_k) &= \sum_{r=1}^{\infty} 2^{r-2 \ln r} \cdot \frac{1}{2^r} = \sum_{r=1}^{\infty} (2^{-2})^{\ln r} = \sum_{r=1}^{\infty} \left(\frac{1}{4}\right)^{\ln r} \\ &= \sum_{r=1}^{\infty} (r)^{\ln\left(\frac{1}{4}\right)} \left(\because a^{\ln n} = n^{\ln a}\right) \\ &= \sum_{r=1}^{\infty} \left(\frac{1}{r}\right)^{\ln 4} \text{ converges since } \ln 4 = 1.39 > 1 \quad \left(\sum_{n=1}^{\infty} \frac{1}{n^p} \text{ converges if } p > 1 \right) \end{aligned}$$

Thus $E(X_k) < \infty$

Since $\{X_k\}$ are i.i.d.r.vs with $E(X_k) < \infty$, the WLLN holds for the sequence, by Khinchine's theorem.

Example 3: Let $\{X_n\}$ be a sequence of i.i.d $U(0, 1)$ r.vs. For the geometric mean $G_n = (X_1 \cdot X_2 \cdot \dots \cdot X_n)^{\frac{1}{n}}$, show that $G_n \xrightarrow{P} c$ where c is some constant. Find c .

Solution: Let $Y = -\ln X$ where $X \sim U(0,1)$. The c.d.f. of Y is given by

$$F_Y(y) = P(Y \leq y) = P(-\ln X \leq y) = P(X \geq e^{-y}) = \int_{e^{-y}}^1 1 \, dx = 1 - e^{-y}$$

$\Rightarrow F_Y(y) = 1 - e^{-y}$ and the p.d.f of Y is given by

$$f_Y(y) = \frac{d}{dx}(F_Y(y)) = e^{-y} \text{ for } y > 0.$$

Then $E(Y) = V(Y) = 1$.

Thus, the sequence $\{Y_n\}$ is i.i.d with finite mean $E(Y_n) = 1$. Hence, by Khinchine's WLLN

$$\sum_{i=1}^n \frac{Y_i}{n} \xrightarrow{P} E(Y_1) = 1 \quad \dots (1)$$

$$\text{But } \ln G_n = \sum_{i=1}^n \ln \frac{X_i}{n} = - \sum_{i=1}^n \frac{Y_i}{n}$$

$$\Rightarrow \sum_{i=1}^n \frac{Y_i}{n} = - \ln G_n \quad \dots (2)$$

From (1) and (2), we have

$$- \ln G_n \xrightarrow{P} 1 \text{ i.e., } G_n \xrightarrow{P} e^{-1}$$

Thus, $c = \frac{1}{e}$.

Example 4: Let X_i can have only two values i^α and $-i^\alpha$ with equal probabilities.

If $\{X_i\}$ is a sequence of independent r.vs, then show that WLLN holds if $\alpha < \frac{1}{2}$.

Solution: Here $E(X_i) = i^\alpha \frac{1}{2} - i^\alpha \frac{1}{2} = 0$ and

$$V(X_i) = E(X_i^2) = i^{2\alpha} \frac{1}{2} + i^{2\alpha} \frac{1}{2} = i^{2\alpha}$$

Let $S_n = \sum_{k=1}^n X_k$. Then

$$B_n = V(S_n) = \sum_{i=1}^n V(X_i) \quad (\because X_i \text{ s are independent})$$

$$= \sum_{i=1}^n i^{2\alpha} = 1^{2\alpha} + 2^{2\alpha} + \dots + n^{2\alpha}$$

$$= \int_0^n x^{2\alpha} dx \quad (\text{Euler - Maclaurion formula})$$

$$\Rightarrow B_n = \frac{n^{2\alpha+1}}{2\alpha+1} \Rightarrow \frac{B_n}{n^2} = \frac{n^{2\alpha+1}}{2\alpha+1} \rightarrow 0 \text{ as } n \rightarrow \infty \text{ if } \alpha < \frac{1}{2}$$

Thus, $\frac{B_n}{n^2} \rightarrow 0$ as $n \rightarrow \infty$ when $\alpha < \frac{1}{2}$

Therefore, $\{X_n\}$ holds WLLN when $\alpha < \frac{1}{2}$.