

# Modeling Titanic Survival

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## ARTICLE HISTORY

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## ABSTRACT

This case study showcases the development of a binary logistic model to predict the probability of survival in the loss of Titanic. I demonstrate the overall modeling process, including preprocessing, exploratory analysis, model fitting, adjustment, bootstrap internal validation and interpretation as well as other relevant techniques such as redundancy analysis and multiple imputation for missing data. The motivation and justification behind critical statistical decisions are explained, touching on key issues such as the choice of a statistical model or a machine learning model, using bootstrap to alleviate selection bias, disadvantages of the holdout sample approach in validation, and more. This analysis is fully reproducible with all source R code and text.

<http://www.crema-research.ch/papers/2009-03.pdf>

Who Survived Titanic? A Logistic Regression Analysis: <https://sci-hub.do/https://journals.sagepub.com/doi/pdf/10.1177/084387140401600205>

<https://www.insider.com/titanic-secrets-facts-2018-4#>

[at-the-memorial-of-frederick-fleet-one-of-the-lookouts-a-prankster-left-a-pair-of-binoculars](#)

<https://www.newscientist.com/article/dn22119-sinking-the-titanic-women-and-children-first>

## 1. Introduction

The sinking of RMS Titanic brought to numerous machine learning competitions a quintessential dataset among others. After the “unsinkable” British passenger liner struck an iceberg in her maiden voyage on 15 April 1912 and was eventually wrecked, more than 1500 people perished. Decades of effort has been devoted to the study of the historic event, in which one major interest for statistical inquiries is to model and predict survival given a number of characteristics, since there was clear account that some people were allowed to get on the lifeboat first.

There are numerous variants of Titanic data existed on the web, with primary source based on [Encyclopedia Titanica](#) (1999), a site started in 1996 as an attempt to tell the story of every person that traveled the Titanic as a passenger or crew member. This project is based on the most recent version as of October 2020, with following columns available (table 1). Source data and steps of data cleaning are elaborated in the [data](#) section in the appendix.

After appropriate formatting and cleaning, the data underlying the whole analysis recorded the survival status 2208 Titanic travelers alongside his/her gender, age, com-

**Table 1.** Cleaned data with 2208 rows and 11 columns

Variable	Definition	Note
survived	survival Status	0 = Lost, 1 = Saved
age	age	In years, some infants had fractional values
gender	gender	
joined	port of embarkation	Belfast, Cherbourg, Queenstown, Southampton
class	cabin class	1st, 2nd, 3rd or Crew
nationality	motherland	from wiki passenger list
title	title	Extracted from name
spouse	# of spouse on board	
sibling	# of siblings on board	
parent	# of parents on board	
children	# of children on board	

panions on board, title, nationality, etc. There were 1496 victims and 712 survivors in total.

It is essential for every fruitful data analysis to first identify key questions of investigation that facilitates exploration and interpretation, however vague at the beginning. Then we can approach the core problem, filtering out trivialities, with statistical expression by abstraction. For our purposes and building on many outstanding works of study (2004; 2009; 2018), we establish the following research questions.

*Question 1:* To which degree is *Women and children first* policy respected? After the collision, the captain explicitly issued an order for women and children to be saved first.<sup>1</sup> Thus we should expect significantly higher proportion of females and children rescued than that in males and adults. If the opposite is true, that Titanic subjects behave more in line with the selfish *homo oeconomicus*, where everybody looked out for himself or herself and possibly even puts other people’s lives in danger, then people in their prime with physical superiority would see higher probability of survival. This requires us to study gender and age effect.

*Question 2:* Did socio-economic advantages mean better chance of survival? If this is the case, passengers with higher financial means, i.e. who live in the first class are more likely to survive. Similarly, passengers from second class will have a higher change of survival than third class people. Cabin class’s impact on survival status needs special notice here.

*Question 3:* What is the crew member’s position in the rescue effort? On the one hand, the 908 crew of men and women are expected to be more experienced, skilled and better informed about the location lifeboats and the incoming danger. On the other hand, their obligation is to care for the safety of passengers first, and only abandon the ship after the task has been fulfilled. We want to see if self interest dominates in the life-and-death situation and crew tend to look out for themselves.

*Question 4:* For those who traveled alone with no companions (spouse, sibling, parent, children) on the vessel, is their survival probability greater or less? People without companions could be in shortage of psychological and physical support during the sinking. While they might also be able to reach a life-saving decision faster without transaction cost and negotiation. Specifically, the effects of having parents or children on the ship has been widely studied. Parental investment suggest that women on average invest more in caring for their offspring than males. In times of a disaster, higher opportunity cost will alert females with offspring more than others, and make them

<sup>1</sup>Though there is no international maritime law enforcing this kind of chivalry.

seek more aggressively for changes to secure the children as well as themselves. In statistical terms, gender-children interaction may exist.

*Question 5:* Did English subjects receive any special care or given priority to aboard lifeboats? After all, Titanic was operated by British crew, and managed by British captain, masters and officers. Conversely, British nobility and elite

This case study has been greatly inspired by Dr. Frank Harrell’s similar example in his *Regression Modeling Strategies* (2015, Chapter 12) book, here I attempt to propose my understanding and interpretation of model development that is as original as possible. To ensure reproducibility, all the analysis is done in R (R Core Team 2020) with code and text made public in this [repo](#). A brief summary of each section is listed below

- **Exploration.** With descriptive statistics, we examine data distribution characteristics, data missing patterns and relative effects, followed by redundancy analysis to study dependencies among predictors. Finish with nonparametric loess regression exploring nonlinear trends.
- **Model development.** The key section in specifying, developing, validating and describing a binary logistic model, split into
  - **Specification** We fit a saturated model with each predictor allowed maximum complexity and nonlinear interactions. This is used to guide development of the final model, based on hypothesis testing and predictor importance ranking with bootstrap.
  - **Multiple imputation:** We use predictive mean matching to impute subject’s age, resulting in 30 complete dataset.
  - **Model fitting and validation.** We fit the final pooled model. Bootstrap internal validation (the “.632” method) is used to study optimism-corrected index measuring discrimination and calibration.
- **Interpretation and discussion.** We describe the model with both graphical methods such as partial effect plots and statistical testing. This provides model-based explanation to address former research problems.
- **Conclusion.** Conclusion and further study.

## 2. Exploration

### 2.1. Descriptive statistics and data processing

A graphical summary of of the data is given by the `Hmisc::describe` function. For numerical variables, a inline histogram is produced alongside summary measures such as the number of missing values and the mean. For discrete variables, we focus on the number of categories and their relative frequency.

11 Variables					2208 Observations									
survived														
n	missing	distinct	Info	Sum	Mean	Gmd								
2208	0	2	0.655	712	0.3225	0.4372								
age														
n	missing	distinct	Info	Mean	Gmd	.05	.10	.25	.50	.75	.90	.95		
1497	711	71	0.999	30.18	14.31	8	17	22	29	38	47	54		
lowest : 0.8 1.0 2.0 3.0 4.0, highest: 67.0 69.0 70.0 71.0 74.0														

### gender

n	missing	distinct
2208	0	2

Value	Female	Male
Frequency	489	1719
Proportion	0.221	0.779

### joined

n	missing	distinct
2208	0	4

Value	Belfast	Cherbourg	Queenstown	Southampton
Frequency	200	271	123	1614
Proportion	0.091	0.123	0.056	0.731

### nationality

n	missing	distinct
2208	0	7

lowest : American English Finnish Irish Other , highest: Finnish Irish Other Swedish Syrian

Value	American	English	Finnish	Irish	Other	Swedish	Syrian
Frequency	246	1002	58	168	549	99	86
Proportion	0.111	0.454	0.026	0.076	0.249	0.045	0.039

### class

n	missing	distinct
2208	0	4

Value	1st	2nd	3rd	crew
Frequency	321	270	709	908
Proportion	0.145	0.122	0.321	0.411

### title

n	missing	distinct
2208	0	4

Value	Miss	Mr	Mrs	other
Frequency	267	1590	212	139
Proportion	0.121	0.720	0.096	0.063

### spouse

n	missing	distinct	Info	Sum	Mean	Gmd
2208	0	2	0.087	66	0.02989	0.05802

### sibling

n	missing	distinct	Info	Mean	Gmd
2208	0	4	0.138	0.05752	0.1103

Value	0	1	2	3
Frequency	2101	91	12	4
Proportion	0.952	0.041	0.005	0.002

### parent

n	missing	distinct	Info	Mean	Gmd
2208	0	3	0.079	0.03804	0.07441

Value	0	1	2
Frequency	2148	36	24
Proportion	0.973	0.016	0.011

### children

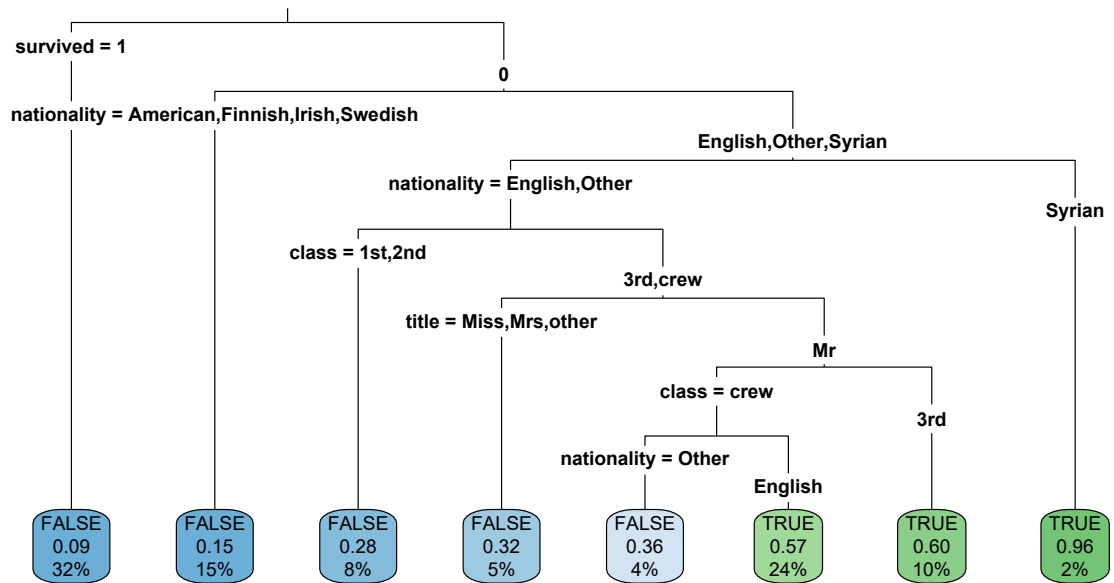
n	missing	distinct	Info	Mean	Gmd
2208	0	5	0.077	0.03895	0.07636

lowest : 0 1 2 3 4, highest: 0 1 2 3 4

Value	0	1	2	3	4
Frequency	2150	37	16	3	2
Proportion	0.974	0.017	0.007	0.001	0.001

There are several noteworthy patterns.<sup>2</sup>

<sup>2</sup>Though this may not be relevant to the model, it is still an surprising discovery that it wasn't until the late 19th century that the idea of women traveling alone gained ground. As a result, there were nearly twice as many males passengers as females on Titanic. In fact, only 40% female passengers have no companion on the



**Figure 1.** The decision tree for predicting `is.na(age)`, which finds strong patterns of missing related to class/department and gender (the Syrian node has very limited samples). Each node shows (top to bottom) the predicted class, the predicted probability of age being missing, the percentage of observations in the node.

Of special importance is the `age` variable, which has roughly 30% missingness. On the other hand, it has a nearly symmetric distribution with 80% known observations falling between 14 and 50. For further examination of patterns of missing data, we could fit a decision tree to predict which type of subject tend to have missing ages. Generally, for some third class male passenger or crew, age is mostly to miss.

```
na_tree <- rpart(factor(is.na(age)) ~ .,
  data = t %>% mutate(survived = as.factor(survived)) ,
  minbucket = 50)
# figure 1
rpart.plot::rpart.plot(na_tree, type = 3, cex = 0.6)
```

We see in figure 1 that survival status, gender and class are essential in determining age missingness. For a 3rd class male passenger who did not survive, age is missing with a probability of 60%. Interestingly, English male crew members are much more likely to have missing age than subjects of other nationality

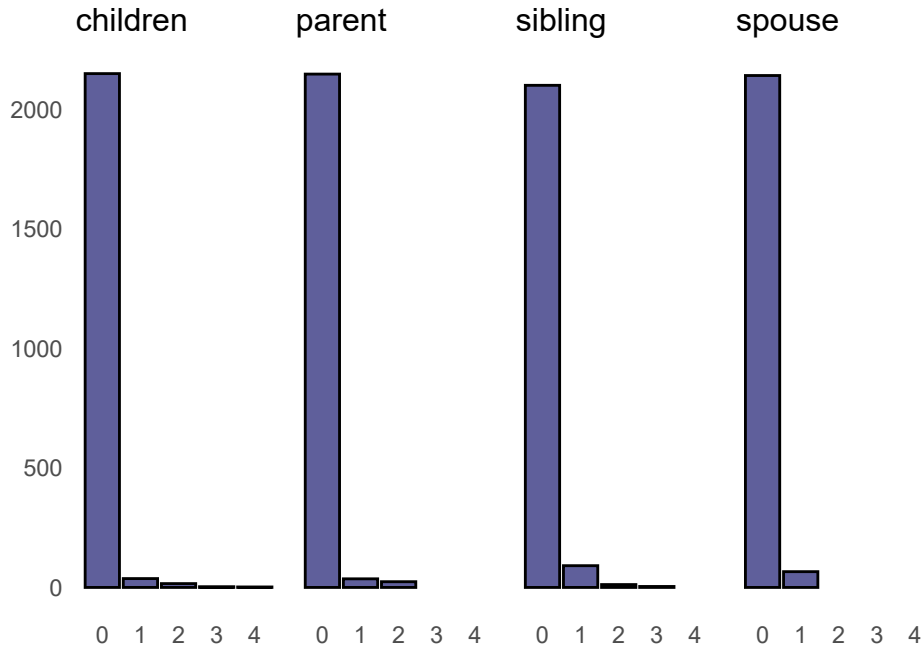
Back to other variables in descriptive statistics. Distributions of subject's companion on Titanic are all too narrow, as illustrated in figure 2. There is too little variation to model them continuously. This motivates categorization since we will not lose too much information. Lastly, nearly half of the subjects are English. And if we focus on crew, the number rise to 85%.

Given this results, the final step in data munging is to dichotomize `spouse`, `parent`, `children` and `sibling` to denote if there is such relation. Thus we no longer have to deal with continuous predictors with poor distribution.

Univariate relationship between each independent variable and survival status is

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ship.



**Figure 2.** Few subjects have more than one companion in any of the 4 relations.

presented in figure 3. For each column, we build a anova-type plot with no control over confounding variables, though it may still assist us in determining how to spend degrees of freedom. If a predictor’s effect on the response is strong, it’s more likely that we need to spend more parameters on it. However, if a variable’s effect appears to be weak, it could either due to a truly flat relationship, or to nonlinearity and predictors among variables that the univariate method fails to detect.

The plot reveals appreciably strong effects of gender and cabin class on survival status. Age effects seem trivial except for the subjects with missing entry, but again, the downside of such kind of univariate display is that it force the audience to think linearly, and only after categorization. For the same reason effects of other variables cannot be determined.

We finish with a redundancy analysis to study if data reduction of removing unnecessary predictors is possible. The checking algorithm expands continuous predictors into cubic splines and categorical predictors into dummy variables, it then use OLS to predict each predictor with all remaining predictors. A predictor is deemed “redundant” if it can be predicted with an  $R^2$  greater than 0.9.

#### Redundancy Analysis

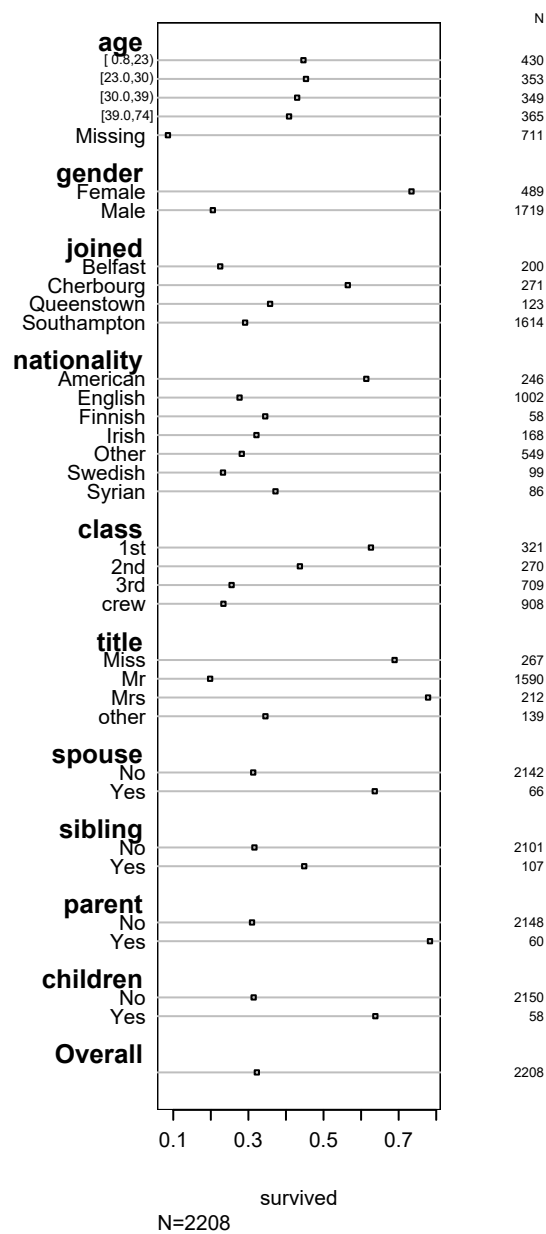
```
redun(formula = ~age + gender + class + nationality + title +
      spouse + sibling + parent + children, data = t)
```

n: 1497      p: 9      nk: 3

Number of NAs: 711

Frequencies of Missing Values Due to Each Variable

age	gender	class	nationality	title	spouse
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**Figure 3.** Univariate summary of relationship between survival and each predictor

711	0	0	0	0	0
sibling	parent	children			
0	0	0			

Transformation of target variables forced to be linear

R-squared cutoff: 0.9    Type: ordinary

R<sup>2</sup> with which each variable can be predicted from all other variables:

age	gender	class	nationality	title	spouse
0.340	0.977	0.551	0.492	0.977	0.341
sibling	parent	children			
0.167	0.289	0.339			

Rendundant variables:

title

Predicted from variables:

age gender class nationality spouse sibling parent children

	Variable Deleted	R <sup>2</sup>	R <sup>2</sup> after later deletions
1	title	0.977	

The redundancy analysis has reported `title` as redundant, with more than 97% of its variation explained by the rest of the predictors. It is not surprising that knowing a person's age could almost determine his/her title in the four categories. This means including `title` contains nearly no predictive value and could be readily deleted.

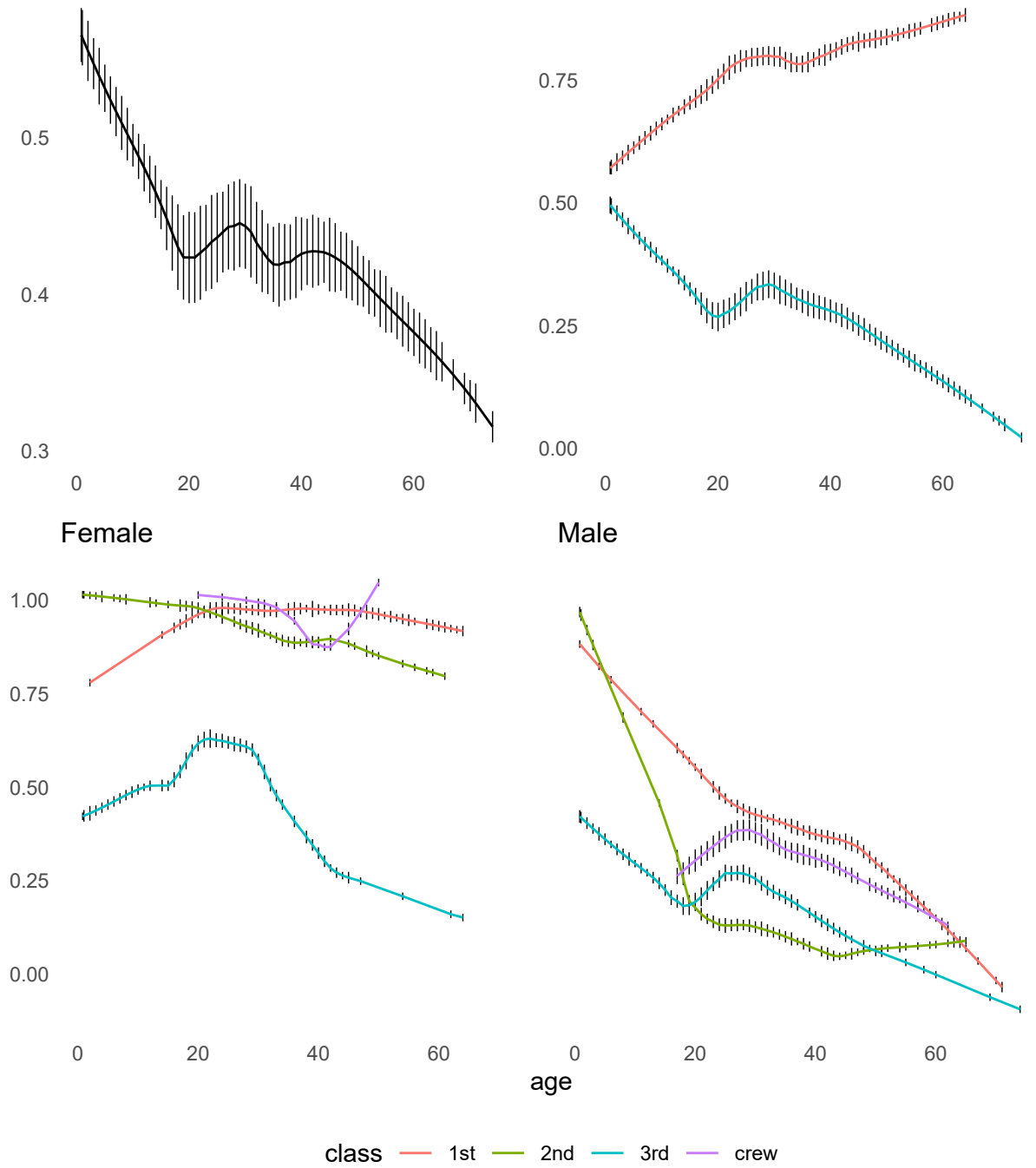
```
t$title <- NULL
```

## 2.2. *Loess regression for nonlinear pattern*

The loess method is a common nonparametric regression model to study nonlinear relationship. In the case of binary response, the fitted value at  $x = x_0$  is the weighted proportion of positive cases near the neighborhood of  $x_0$ . If a loess curve exhibits a reasonable degree of nonmonotonicity, it will often pay to not assume linearity, e.g., modeling the predictor with polynomial transformation or with splines.

It was widely documented that gender interacts directly with age, in other words, age effects for men and women are most likely to be nonparallel in the logit scale. Another prominent joint effects, according to many follow up studies, happened between age and cabin class. Figure 4 displays loess estimates of survival probability given age under certain stratification. Not only in a powerful nonlinear fashion does age affect survival status (top left panel), we also observe it interact with other two factors in a nonlinear way. Despite that the loess estimation here would by no means serve as a strict inference-oriented model, and we only include 3 factors at this point, it inspires a treatment of these effects that is more thorough and cautious.





**Figure 4.** loess estimates of  $P(\text{survived})$ , with tick marks representing frequency counts within equal-width bins. Top left panel indicates nonlinear age effect without controlling other factors. Other plots give estimates under stratification by sex and class. The top right plot reflects women have a higher survival probability in general, and older females (red) and younger males (blue) are most likely to survive in their respective gender group. Two bottom plots depict survival pattern of subjects from different cabin class condition on sex.

### 3. Model development

A typical modeling workflow begins with an choice of a statistical model or a machine learning model. A statistical model often stems from a hypothesized probabilistic data generating mechanism and assumes additivity, whereas machine learning models is algorithmatic in nature, optimized with parameter tuning. We choose to develop a statistical model, a “simple” binary logistic regression, for the following reasons.

We prefer probabilistic predictions to classification with output label 0 and 1, since we are placing emphasis upon the *tendency* of survival. And the value of the model consist not in a dichotomous prediction, but in what characteristics would increase or decrease the possibility of survival. The notion has ruled out most of the machine learning models for classification, say, random forest, support vector machines and neural network, which are not intrinsically probability oriented. Such classifiers can often only yield a forced choice.

Interpretability and inference matters. Many top data science competitions has reported moderately high signal to noise ratio (e.g., 90% prediction accuracy) that might tip the balance towards machine learning models, whereas interpretability is harmed. Specifically, statistical models favours additivity and explicit specification. It follows that there are natural distinctions between main effects and interactions, linearity and nonlinearity. And the inference procedure is well defined provided that the model is correctly specified. While in a multi-layer neural network, everything can interact with one another and it could be daunting to isolate effects and conduct former inference.

Machine learning models are data hungry and sometimes create the need for big data (van der Ploeg, Austin, and Steyerberg 2014). To guard against overfitting, the analyst has to have a sample size that is 10 times larger at least if he chooses a tree model instead of regression. While this case study uses a Titanic dataset that is about 1/3 larger than those only concerned with passengers, it is far less sufficient for a typical data-hungry machine learning model to validate well. The rationale is that a statistical model is a safer approach as Dr. Harrell commented

If n is too small to do something simple, it is too small to do something complex

#### 3.1. Specification

We start by fitting a relatively large model, to decide how model complexity should be properly represented. This includes deciding the number of knots for continuous predictors and the number of categories of categorical predictors, could we remove some term, where should we place interaction, etc. The large model also gives an overall sense of the predictive ability of each subject characteristics on survival status. This strategy as a starting point is also called prespecification of predictor complexity. It avoids creating phantom degrees of freedom when one has subjective judgment according to scatter diagrams or descriptive statistics on how to represent variables in a model. Commonly done, for example, is excluding a quadratic term simply because it is “non-significant”, with p-value on the edge of 0.05. This approach is known to distort coefficient estimates, confidence intervals, p-value and calibration (too optimistic) of the final model (Grambsch and O’Brien 1991).<sup>3</sup> Because it fails to accounts for sampling variability and suffers selection bias. Therefore, it is essential to have extra caution, as demonstrated below using resampling, to do model simplification.

Prespecification of predictor complexity is done first by developing a saturated lo-

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<sup>3</sup>confidence interval too narrow, p-value and standard errors too small and calibration too optimistic

gistic model and then making necessary adjustments and improvements. In this model, we granted age effect maximal flexibility represented as natural splines with 5 knots, and all categorical predictors retain their original categories without pooling. Two way interactions have been specified between age and gender, age and class, and age and parent. Since this is an initial model, observations with missing age are not used. The model equation is

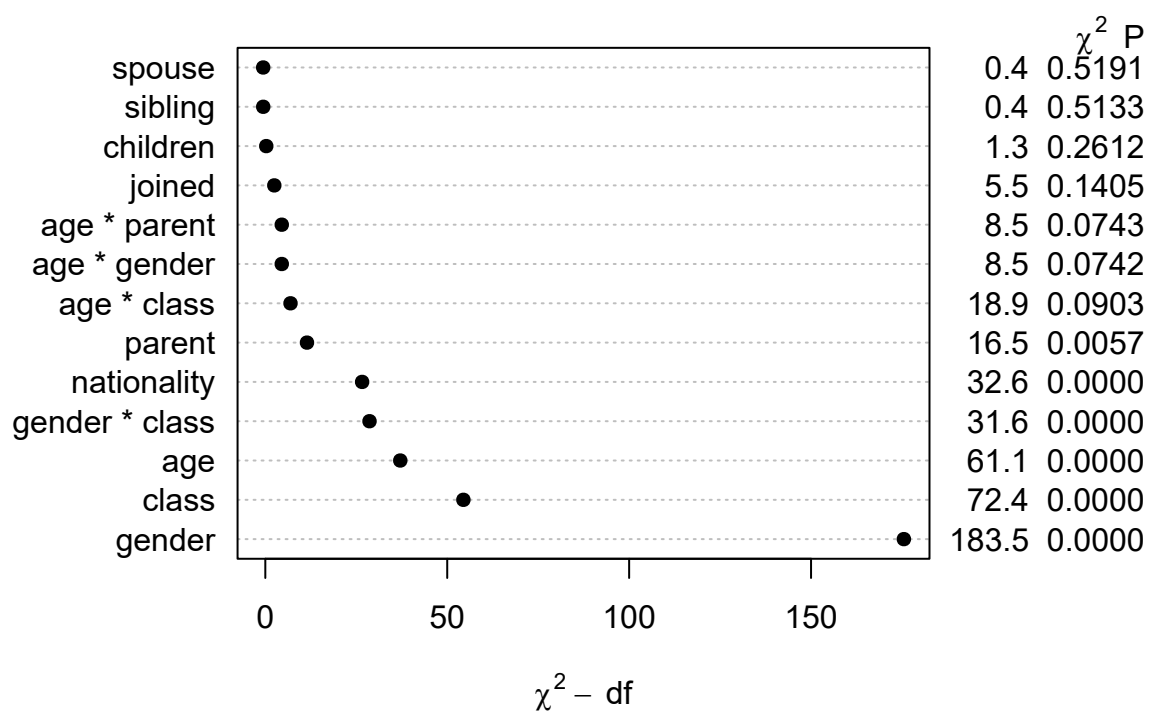
```
survived ~ (rcs(age, 5) + gender + class)^2 + (rcs(age, 5) * parent) +
          joined + spouse + sibling + children + nationality
```

**Table 2.** Hypothesis testing for the saturated model

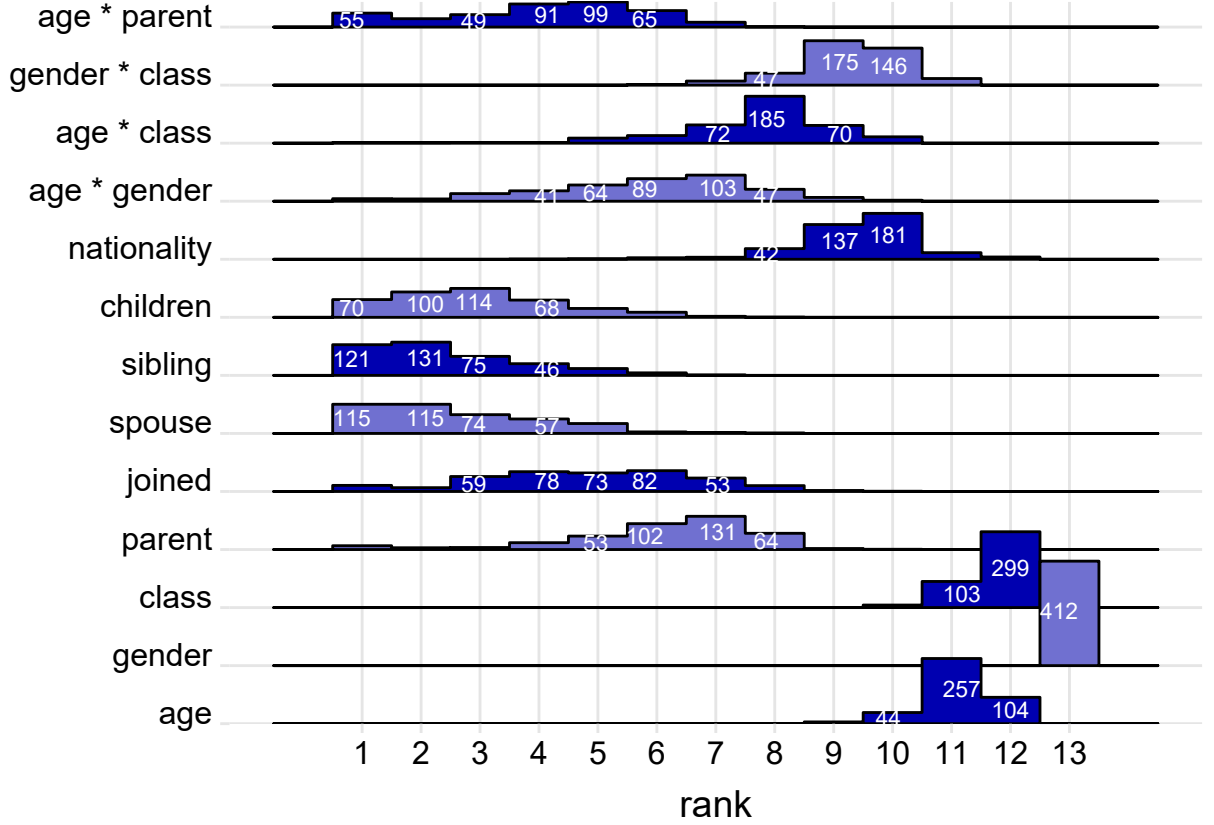
	$\chi^2$	d.f.	$P$
age (Factor+Higher Order Factors)	61.09	24	<0.0001
<i>All Interactions</i>	35.33	20	0.0184
<i>Nonlinear (Factor+Higher Order Factors)</i>	45.96	18	0.0003
gender (Factor+Higher Order Factors)	183.51	8	<0.0001
<i>All Interactions</i>	50.33	7	<0.0001
class (Factor+Higher Order Factors)	72.44	18	<0.0001
<i>All Interactions</i>	55.11	15	<0.0001
parent (Factor+Higher Order Factors)	16.45	5	0.0057
<i>All Interactions</i>	8.52	4	0.0743
spouse	0.42	1	0.5191
sibling	0.43	1	0.5133
children	1.26	1	0.2612
joined	5.47	3	0.1405
nationality	32.61	6	<0.0001
age $\times$ gender (Factor+Higher Order Factors)	8.52	4	0.0742
<i>Nonlinear</i>	8.09	3	0.0441
<i>Nonlinear Interaction : <math>f(A,B)</math> vs. <math>AB</math></i>	8.09	3	0.0441
age $\times$ class (Factor+Higher Order Factors)	18.93	12	0.0903
<i>Nonlinear</i>	17.77	9	0.0380
<i>Nonlinear Interaction : <math>f(A,B)</math> vs. <math>AB</math></i>	17.77	9	0.0380
gender $\times$ class (Factor+Higher Order Factors)	31.64	3	<0.0001
age $\times$ parent (Factor+Higher Order Factors)	8.52	4	0.0743
<i>Nonlinear</i>	1.52	3	0.6768
<i>Nonlinear Interaction : <math>f(A,B)</math> vs. <math>AB</math></i>	1.52	3	0.6768
TOTAL NONLINEAR	45.96	18	0.0003
TOTAL INTERACTION	74.90	23	<0.0001
TOTAL NONLINEAR + INTERACTION	93.56	26	<0.0001
TOTAL	265.18	44	<0.0001

Table 2 selects specific hypothesis to test for general power, linearity and additivity assumptions of individual predictors, as well as their “chunky” version of global effects. We see dominant main effects of gender, age and cabin class, be it linear or nonlinear ( $p < 0.0001$ ) More notably are the strong nonlinear interaction terms between the three. Of all four companion variables only parent manifests clear predictive ability. The impact of embarkation point is somewhat ambiguous ( $p = 0.14$ ). As a graphical illustration, figure 5 plots “adjusted” partial  $\chi^2$  statistic of each predictor in the saturated model, with correction for degrees of freedom allocated to them.<sup>4</sup> This adjustment levels the playing field for comparison of predictive ability. The larger the

<sup>4</sup>The correction is done by subtracting the d.f. from the partial  $\chi^2$  statistic, its expected value under the null hypothesis.



**Figure 5.** Ranking of predictive power in the saturated model based on adjusted  $\chi^2$

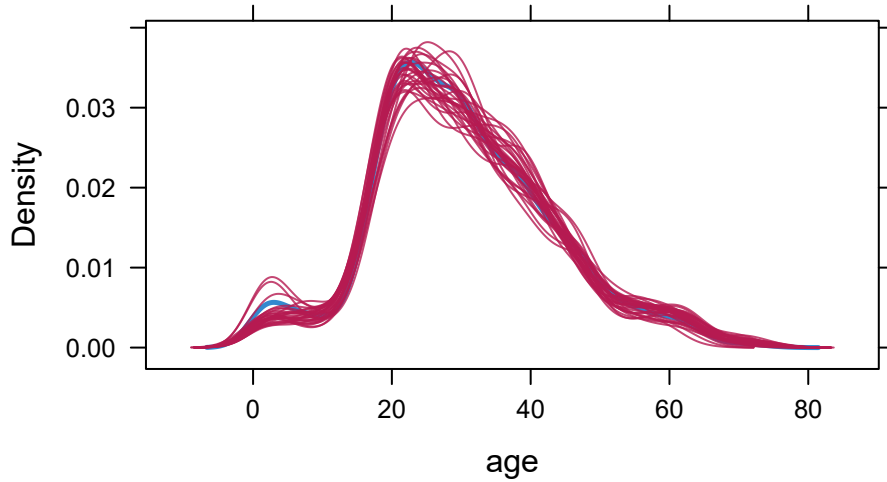


**Figure 6.** Distribution of importance ranking over 500 bootstrap resamples, 412 of which are actually fitted without numerical problems. Text indicates the number of times a term has a specific ranking, when the term ranked more than 40 times at that position. For example, gender ranks 13 (the most important) in all valid resamples.

adjusted  $\chi^2$ , the more likely a variable would have a non-flat impact on survival status.

As mentioned before, the goal of the saturated model is guiding model complexity. More specifically, should we allocate more degrees of freedom to a certain term because some complex effects has been underrepresented? Or is there a term that is highly irrelevant thus could be deleted? The polynomial transformation on age and the resulting nonlinear interaction carry substantial statistical power, while further increasing knots or creating high-order interactions causes numerical problems. Therefore, it is positive advantage to us to keep them as is. There are also not sufficient reasons to collapse levels for nationality and port. Binary variables like spouse and sibling have extremely large p-values ( $p > 0.5$ ), indicating relatively small predictive power. Still, great care should be taken when one attempts to conduct aggressive model simplification based on hypothesis testing and p-values. A reliable way is using bootstrap resampling. Figure 6 studies the importance of all terms including main effects and interaction over 500 bootstrap resamples. In each resample, we fit the saturated model, rank all 13 terms by the adjusted statistic  $\chi^2 - \text{d.f.}$  in ascending order so that 13 is most important and 1 is least important. The height of a bar indicates the number of times a term is ranked at that position.

The importance ranking echoes previous findings that gender, age and cabin class are predominant factors. It also reveals great variability in terms of assessing predic-



**Figure 7.** Density plot of observed and imputed data. In general, the imputed dataset mimic the age distribution seen in the observed data.

tive power. For example, we are only confident that `joined` is not one of the 5 most influential predictors. Nonetheless, rankings of the aforementioned “weak” variables, sibling and spouse, are highly concentrated at 1 to 3. In fact, if we perform backward selection in nearly 500 bootstrap resamples with AIC as stopping rule, none of the 3 binary variables entered the selected model more than 20 times. This results in the final decision to remove them in the final model.

```
t$sibling <- NULL
t$spouse <- NULL
```

### 3.2. Multiple imputation

```
# multiple imputation with predictive mean matching to generate 30 complete dataset
imp <- mice(t, method = "pmm", m = 30, printFlag = FALSE)
```

The last step before fitting the final model is imputing missing values for age. The goal of multiple imputation, in contrast to simple alternatives such as filling in conditional mean, is to provide an accurate estimate of the variance-covariance matrix that not only accounts for sampling variability, but also for the extra variance caused by missing values and finite number of imputations (Van Buuren 2018). Thus tests on individual parameters gain power and bias are reduced. The general idea is to generate multiple complete dataset, fit the model in parallel, and then obtain a pooled final estimate by averaging over all fitted models.

We use predictive mean matching with  $m = 30$ , since approximately 30% of age is missing. The nonparametric method selects a group of Titanic subjects from all complete cases that have predicted values closest to the predicted value for the subject with missing age.<sup>5</sup> One donor is randomly drawn from the candidates, and the observed

<sup>5</sup>The predicted value is generated by fitting a linear main effect model conditional on all other variables.

age of the donor is taken to replace the missing value. We use the default “type 1 matching” and 5 donors (Van Buuren (2018), Section 3.4.2). Advantages of predictive mean matching in the Titanic age setting are manifold. Since imputations are based on values observed elsewhere, they are realistic (e.g., no negative age). For another, it is compatible with non-normality which allows us to have fewer assumptions.

### 3.3. Model fitting, validation and calibration

We now fit the final logistic model across 30 complete dataset. Pooled estimates are obtained by averaging over all pieces. We also get an imputation-corrected variance-covariance matrix based on within- and between-imputation variances. Individual estimates are listed below.

	$\hat{\beta}$	S.E.	Wald Z	Pr(>  Z )
Intercept	-0.1475	1.9879	-0.07	0.9409
age	0.2225	0.1244	1.79	0.0737
age'	-0.9850	0.6877	-1.43	0.1520
age''	4.4844	3.6008	1.25	0.2130
age'''	-5.5980	4.9656	-1.13	0.2596
gender=Male	-1.8842	1.0197	-1.85	0.0646
class=2nd	5.4575	2.8123	1.94	0.0523
class=3rd	-1.0535	1.8583	-0.57	0.5708
class=crew	-2.4149	3.2651	-0.74	0.4595
parent=1	3.4640	1.2035	2.88	0.0040
nationality=English	0.0730	0.2757	0.26	0.7912
nationality=Finnish	0.1843	0.4289	0.43	0.6674
nationality=Irish	0.0116	0.3901	0.03	0.9763
nationality=Other	-0.2328	0.2669	-0.87	0.3832
nationality=Swedish	-0.2673	0.3820	-0.70	0.4840
nationality=Syrrian	-0.0582	0.4429	-0.13	0.8955
joined=Cherbourg	0.6565	0.3321	1.98	0.0480
joined=Queenstown	0.0538	0.4133	0.13	0.8965
joined=Southampton	-0.0632	0.2003	-0.32	0.7523
age × gender=Male	-0.1498	0.0573	-2.61	0.0090
age' × gender=Male	0.8213	0.3849	2.13	0.0329
age'' × gender=Male	-3.9560	2.3149	-1.71	0.0875
age''' × gender=Male	4.8271	3.5665	1.35	0.1759
age × class=2nd	-0.3846	0.1754	-2.19	0.0283
age' × class=2nd	1.4771	0.8999	1.64	0.1007
age'' × class=2nd	-6.7250	4.5517	-1.48	0.1395
age''' × class=2nd	8.5082	6.1731	1.38	0.1681
age × class=3rd	-0.1403	0.1159	-1.21	0.2262
age' × class=3rd	0.7699	0.6331	1.22	0.2240
age'' × class=3rd	-4.9212	3.2961	-1.49	0.1354
age''' × class=3rd	7.7752	4.5292	1.72	0.0860
age × class=crew	0.0589	0.1890	0.31	0.7554
age' × class=crew	-0.0468	0.8023	-0.06	0.9535
age'' × class=crew	-0.5748	3.6057	-0.16	0.8733
age''' × class=crew	1.5477	4.4606	0.35	0.7286
gender=Male × class=2nd	-0.4060	0.7087	-0.57	0.5667
gender=Male × class=3rd	2.1475	0.6288	3.41	0.0006
gender=Male × class=crew	0.5495	0.8726	0.63	0.5289
age × parent=1	-0.1190	0.1118	-1.06	0.2871
age' × parent=1	-0.6885	1.1693	-0.59	0.5560
age'' × parent=1	10.4417	11.6224	0.90	0.3690
age''' × parent=1	-29.7852	35.4739	-0.84	0.4011

Table 3 again constructs meaningful hypothesis testing for the final model. The  $\chi^2$  statistic of age decreased by a minor amount, resulting from using patterns of association with survival status to impute missing age. Remaining predictors generally have larger  $\chi^2$  statistic and smaller p-value compared to the saturated model in table 2, due to larger sample size in model development. Table 4 prints model assessment metrics and details individual parameters. The model exhibits moderate discrimination power, e.g. the ability to separate perished and survived subjects (concordance probability = area under the ROC curve  $\approx 0.81$ ). Brier score, as a proper quadratic scoring rule that incorporate both aspects, is a promising 0.144.

**Table 3.** Hypothesis testing for the final model

	$\chi^2$	d.f.	$P$
age (Factor+Higher Order Factors)	57.01	24	0.0002
<i>All Interactions</i>	34.32	20	0.0240
<i>Nonlinear (Factor+Higher Order Factors)</i>	36.45	18	0.0062
gender (Factor+Higher Order Factors)	252.05	8	<0.0001
<i>All Interactions</i>	50.72	7	<0.0001
class (Factor+Higher Order Factors)	94.17	18	<0.0001
<i>All Interactions</i>	53.61	15	<0.0001
parent (Factor+Higher Order Factors)	13.82	5	0.0168
<i>All Interactions</i>	7.58	4	0.1082
children	5.26	1	0.0218
nationality	4.71	6	0.5815
joined	7.31	3	0.0626
age $\times$ gender (Factor+Higher Order Factors)	6.73	4	0.1506
<i>Nonlinear</i>	5.97	3	0.1129
<i>Nonlinear Interaction : <math>f(A,B)</math> vs. <math>AB</math></i>	5.97	3	0.1129
age $\times$ class (Factor+Higher Order Factors)	16.18	12	0.1831
<i>Nonlinear</i>	13.64	9	0.1359
<i>Nonlinear Interaction : <math>f(A,B)</math> vs. <math>AB</math></i>	13.64	9	0.1359
gender $\times$ class (Factor+Higher Order Factors)	30.97	3	<0.0001
age $\times$ parent (Factor+Higher Order Factors)	7.58	4	0.1082
<i>Nonlinear</i>	1.48	3	0.6870
<i>Nonlinear Interaction : <math>f(A,B)</math> vs. <math>AB</math></i>	1.48	3	0.6870
TOTAL NONLINEAR	36.45	18	0.0062
TOTAL INTERACTION	74.55	23	<0.0001
TOTAL NONLINEAR + INTERACTION	86.92	26	<0.0001
TOTAL	356.20	42	<0.0001

**Table 4.** Model index and estimation

		Model Likelihood Ratio Test		Discrimination Indexes		Rank Discrim. Indexes	
Obs	2208	LR $\chi^2$	783.83	$R^2$	0.418	$C$	0.813
0	1496	d.f.	41	$g$	1.799	$D_{xy}$	0.626
1	712	$\Pr(> \chi^2) < 0.0001$		$g_r$	6.053	$\gamma$	0.628
$\max  \frac{\partial \log L}{\partial \beta} $		0.007		$g_p$	0.274	$\tau_a$	0.274
				Brier	0.144		



**Table 5.** Optimism-corrected metrics

Index	Original Sample	Training Sample	Test Sample	Optimism	Corrected Index	$n$
$D_{xy}$	0.6253	0.6411	0.5642	0.0386	0.5867	492
$R^2$	0.4182	0.4364	0.3139	0.0659	0.3523	492
Intercept	0.0000	0.0000	-0.2479	0.1567	-0.1567	492
Slope	1.0000	1.0000	0.6921	0.1946	0.8054	492
$E_{\max}$	0.0000	0.0000	0.0769	0.0769	0.0769	492
$D$	0.3552	0.3743	0.2581	0.0614	0.2938	492
$U$	-0.0009	-0.0009	<i>Inf</i>	<i>-Inf</i>	<i>Inf</i>	492
$Q$	0.3561	0.3752	<i>-Inf</i>	<i>Inf</i>	<i>-Inf</i>	492
$B$	0.1437	0.1409	0.1506	-0.0044	0.1480	492
$g$	1.7827	1.9959	1.3862	0.2506	1.5321	492
$g_p$	0.2744	0.2815	0.2209	0.0339	0.2406	492

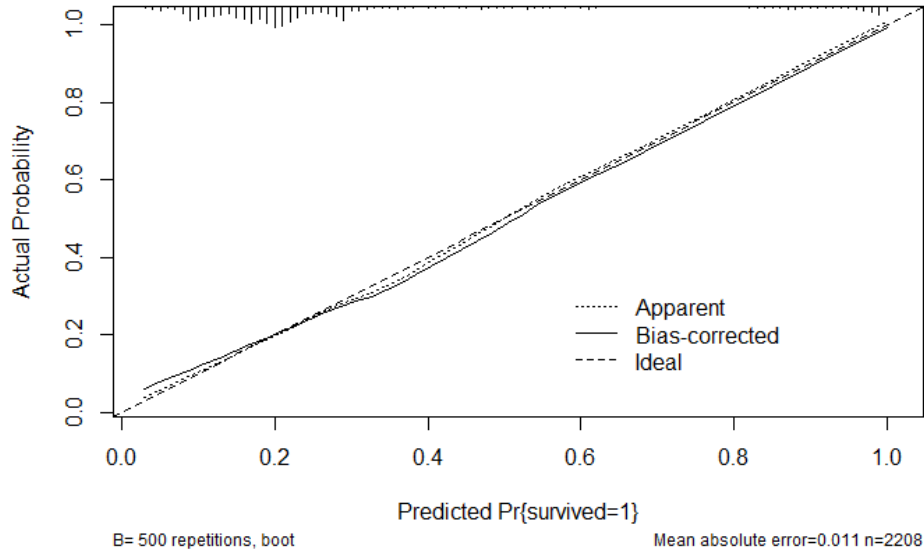
Although there will not be a second Titanic, making prediction a lesser problem, validation can still be used for good purposes. It quantifies the degree of overfitting by presented unbiased, optimism-corrected measures. More accurately, we will be using bootstrap internal validation to study the “future” performance of the model. In an award-winning solution to this legendary dataset submitted by IBM Watson, a holdout test set was used to validate their model. The data-splitting approach is known to require a significantly larger sample size ( $> 20000$ ) than resampling methods on average to work acceptably well (Harrell, Jr. 2020b). Moreover, when the model developed on training sample is validated, the researcher would recombine training and testing set to fit a full model. This model, however, is never validated.

As an improved alternative, we choose Efron’s 0.632 method for bootstrap internal validation. In each of the 494 bootstrap resamples, a logistic model is developed and evaluated on observations omitted from bootstrap samples. Per-bootstrap optimism is then the apparent index of accuracy subtracting that in the test sample formed by omitted observations. An weighted average  $\hat{\varepsilon}_0$  over all 494 bootstrap resamples is computed to estimate the true optimism, while the bias-corrected estimate of predictive accuracy is calculated as  $0.632(\text{apparent accuracy} - \hat{\varepsilon}_0)$ . Table 5 displays the results. It validates two general aspects of model accuracy, discrimination and calibration. Calibration is the ability to make unbiased estimates of survival status, while discrimination is the a measure in how separated predictions are for survivors and victims.

The output does not deviate much from the apparent index. The validated area under the ROC curve as well as the concordance probability is now 0.79, and pseudo  $R^2$  0.35. This indicates certain shortage of discrimination ability in comparison to some published models on the older dataset. Slope = 0.8 signals small amount of overfitting, with coefficients shrinking by 20% on new data.<sup>6</sup> Brier score is now near 0.15.

Discrimination index are often associated with rank correlations between predictions and response, thus may not be sensitive enough to evaluate their closeness. Figure 8 aims to gauge the concordance between predicted values and observed data, or in other words, calibration. The actual probability is estimated with loess regression, and the bias correction is computed in a similar way as in table 4. The 45 degree line indicates the ideal scenario in which prediction perfectly matches observation. The model is well-calibrated by and large, with slight departure from the straight line in central

<sup>6</sup>  $D_{xy} = 2(\text{concordance probability} - 0.5)$



**Figure 8.** Bias-corrected calibration curve from internal bootstrap validation. Predictions output by the final model are plotted against estimated actual probabilities with loess. Bars on the top indicates number of observations in the same prediction bin. The bias corrected line is close to the ideal 45 degree line

region where there were few observations. The mean squared error is 0.00018, and the 0.9 quantile of absolute error is 0.023. Similarly, the unreliability index  $E_{max} = 0.54$  measures maximum error in predicted probabilities. All these metrics reflects strong calibration.

To sum up, the model presents reasonable discrimination power and satisfactory quality of fitting. Excellent performance in calibration lifts overall metrics such as Brier score.

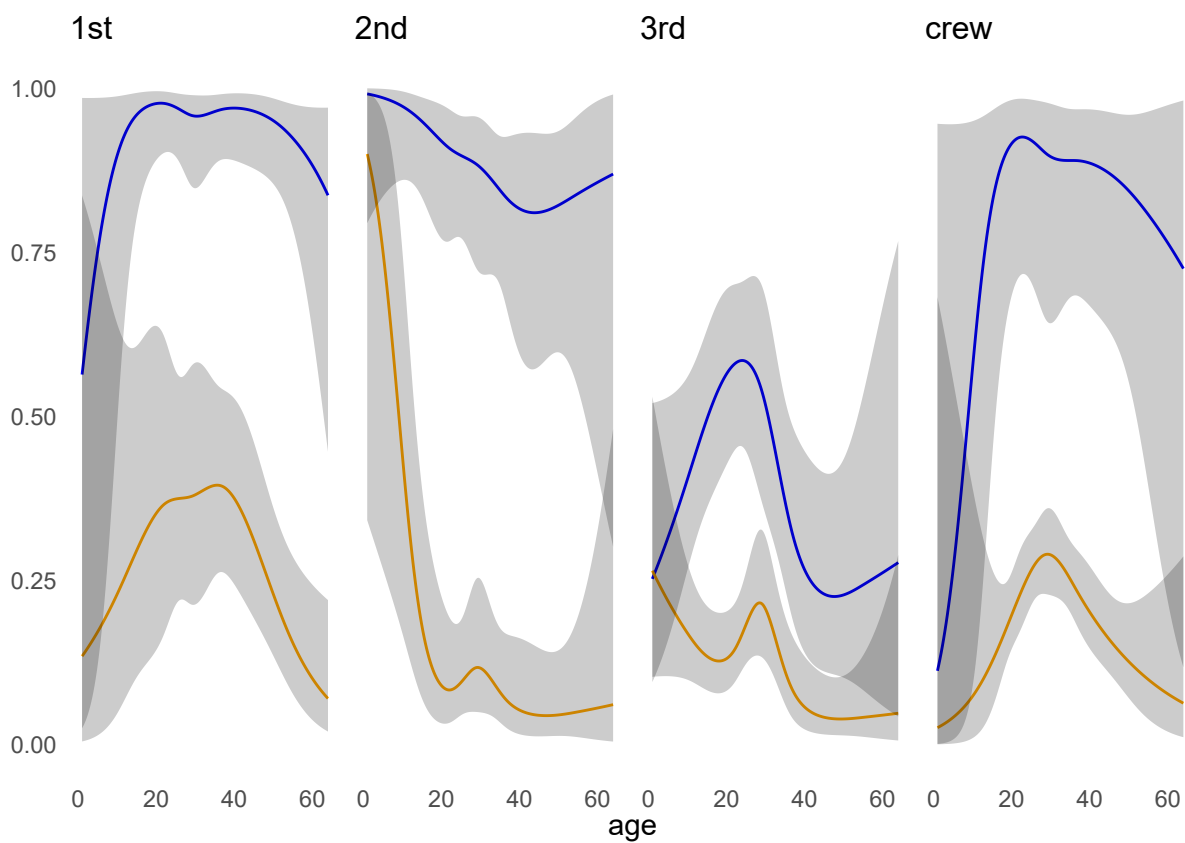
#### 4. Interpretation and discussion

With non-monotonic relations and interactions involved, interpretations built upon parameter estimates or interquartile-range odds ratios are rarely informative. In this case, the model can be effectively described by partial effects plots, where we plot each characteristics against  $\hat{P}(\text{survived})$  while holding other predictors constant at mean or median for continuous variables and mode for categorical variables. When interactions are involved, the survival probability are estimated separately for various levels of the interacting factors. For categorical predictors, level-specific effects are computed comparing to a reference category.

Figure 9 is an example partial effect plot for age under stratification of sex and cabin class.

We could now address question 1, 2 and 3 with figure 9, where we are inspecting all these three variables' influence on survival status. We want to know if women, youth, living in a higher class as passenger or being crew results in greater chance to survive.

Women and children are evacuated first in general, supported both by data and stories. 73.4% of the women and 57.3% of the children (age under 12) were saved, but



**Figure 9.** Age effects on survival status for **female** and **male**, stratified by cabin class while holding other predictors constant at parent = 0, children = 0, nationality = English, and port = Southampton. Shaded region shows 95% confidence interval.

only 20.5% of the men. There was also recollection that at the beginning of the sinking, the Second Officer on the Titanic suggested to Captain Smith, “Hadn’t we better get the women and children into the boats, sir?”, to which the captain responded: “put the women and children in and lower away”. Fifth Officer Harold Lowe in the American Inquiry as reinforced this policy as one regardless of classes

It was simply the first woman, whether first class, second class, third class, or sixty-seventh class. It was all the same; women and children were first.

A closer look at numbers, particularly under stratification suggests another story. To begin with, not all classes of women and children receive the same level of help during the sinking, whether intended or unintended. We observe that first class passengers have the highest predicted survival possibility in general, whereas the third class is the lowest. The chasm between classes is most manifest in adult women. A 30-year-old first class woman is likely to survive with a likelihood between (84.7%, 98.9%), and if she were in the third class, (35.0%, 68.1%). For children, knowing a 12-year-old comes from the first class increases the upper bound of predicted survival probability by nearly 30%. The most decisive explanation is that first-class passengers had better access to information about the imminent danger and were aware that the lifeboats were located close to the first class cabins. Thus, their marginal effort costs to survive were lower. In contrast, most third-class passengers are located at the quarters down the stern, from which the designer deliberately made it difficult to reach the upper decks.

Our second argument against “women and children first”, is formed by digging into each of the 4 groups of people in figure ?? . We see that, even without class effects, this policy is in many cases not the truth.

The second class are chosen as a reference group, where the prediction roughly matches what should happen if children and women received the asserted assistance on the part of the authorities. The likelihood of survival for both gender starts approximately at the same higher region, then the probability for male plunged as he ages, while female only see a minor decrease after leaving the “girl” region.

The first class is the cohort that survive the most in percentages. Women have a mean survival advantage of 40% in terms of predicted probability. Yet there is not much evidence to argue children are saved first. Boys’ survival percentage is only 2% higher than men between 20 and 40 (72% versus 70%), and their prediction interval is much wider. For girls, there might even be a slight advantage for adult female over the young.

The second class has a noteworthy male survival trend,

Next, we will focus on the third class, where this code of conduct is not respected as much. There are we see two abnormal trends. Firstly, people in their prime have higher predicted survival probability. The survival curve displays an inverted U-shape, where people aged between 20 and 30 are mostly likely to survive, for both men and women. Secondly, the gender gap in survival probability narrowed drastically for third class, if someone is a third class miss or elderly woman, it is hard to conclude she has any priority in getting in the lifeboat over a man of similar age. For example,

Interestingly, the crew out of four roles share a similar survival trend with first class passengers.

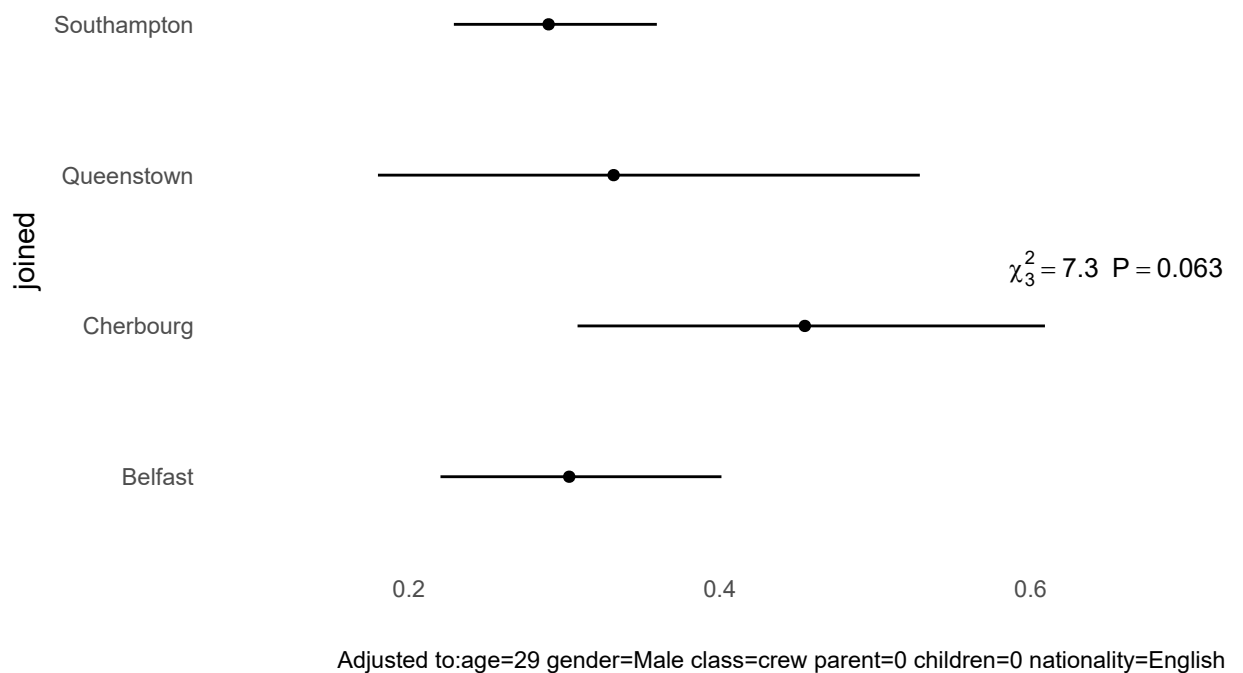
Nationality is highly insignificant ( $p = 0.58$ ).

<https://www.encyclopedia-titanica.org/community/threads/>

[passengers-who-spoke-other-languages.20103/](#)

One explanation is that the crew's instructions (in English) tended to be along the lines of "Wait down here for further orders" a lack of understanding might well have saved many lives. Also many of the immigrants in 3rd class were traveling in family or neighborhood groups which included at least one English-speaker (often an established immigrant returning to the US from a visit back home) who could act as their spokesperson.

One last discovery concerns embarkation point. Passengers boarded at Cherbourg have a surprisingly higher survival rate of over 50%.

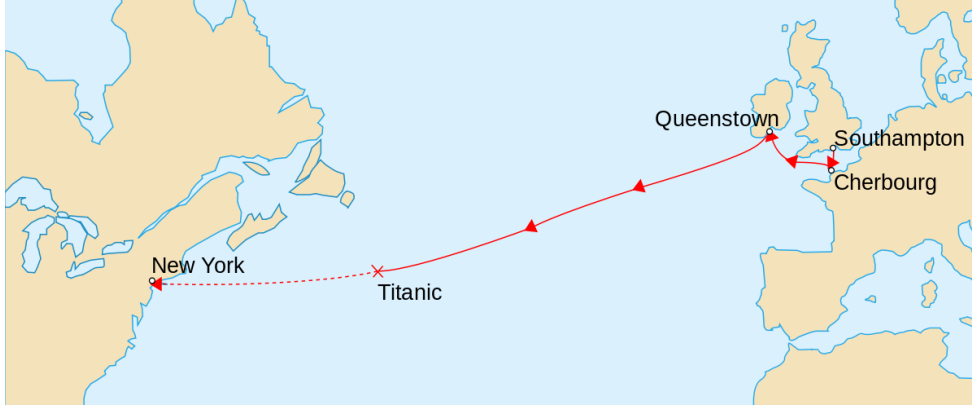


<sup>7</sup>

## 5. Conclusion

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<sup>7</sup>Source: <https://evanexpertsattheroutethird.weebly.com/route>



**Figure 10.** The Titanic's route.

## Appendix A. Data

A variety of other versions and forms of Titanic data sources have been collected due to public's constant interests in the tragedy as well as modern efforts trying to unveil the mystery. A comprehensive overview of several data variants is given by [Symanzik, Friendly, and Onder \(2018\)](#). Data in this case study is accessed on [Encyclopedia Titanica](#), a leading archive on titanic facts. In contrast to the the famous titanic dataset (known as `titanic3`) distributed by [kaggle](#) for introductory level machine learning practices, the case study uses a more up-to-date and complete dataset in the following ways

- **Larger sample size.** Our data includes crew and staff members alongside passengers, while `titanic3` only incorporate passenger information. We do not use a separate test set approach for validation either. As a result, the sample size is about 2.5 times larger.
- **More columns.** Additional variables such as role on the ship, nationality and occupation are added. A major difference is made by separating the travel companion data into four distinct columns: number of parents, children, sibling and spouses that each passenger traveled with. These were combined into two columns before.
- **More accurate.** `titanic3` was an effort to study Titanic in the 20th century, lastly updated and improved by Thomas Cason in 1999. During the recent two decades the data has been constantly revised, many errors corrected, many missing ages filled in, and new variables created. Now it reflects our our most up-to-date understanding of the event, in the digital form, as of 21 October 2020.

The data cleaning process involves using appropriate data types, creating new features, adjusting levels for categorical variable and excluding irrelevant columns. Code can be found at [clean.R](#).

`title` is extracted through each person's name with regular expressions and then collapsed into 4 levels.<sup>8</sup> This is a predictor that has been widely reported to have good predictive ability in many submissions. However, as we see in the redundancy analysis at the end of Section 2.1, it should not even be accepted in the tentative, saturated model.

---

<sup>8</sup>For example, the title for passenger "Abbing, Mr Anthony" is "Mr".

Passengers are classified according to their cabin class. Others on the vessel fall into one of crew and staff members. Crew includes victualling crew<sup>9</sup>, engineering crew, deck crew and officers, substitute crew and guarantee group. Staff members include restaurant staff and orchestra.

Rare nationality (lower than 50 people) is collapsed.

Age information is presented as non-missing on the surface yet there is an indicator column representing when a person's age is only approximate and cannot be fully determined from current facts. These inaccurate age have been assigned NA. There were also ten subjects whose four companion variables were all explicitly missing. For simplicity, the mode 0 is filled in. Therefore, the problem of missing data is reduced to univariate missing of `age`.

Variables we do not utilize in this project includes name, date of birth and death, lifeboat number<sup>10</sup>, fare, and cabin number.<sup>11</sup>

## Appendix B. Model formula

$$\text{Prob}\{\text{survived} = 1\} = \frac{1}{1 + \exp(-X\beta)}, \text{ where}$$

---

<sup>9</sup>crew in charge of food, housekeeping, laundry, room service, etc.

<sup>10</sup>There were 9 recorded passengers who got on the lifeboat yet died before reaching Carpathia, another RMS which spearheaded the rescue of Titanic survivors. There were also 13 passengers who survived with no boat information documented, and this is most likely due to data quality issues after looking up on Encyclopedia Titanica. Even with these exceptions, whether a passenger got on a lifeboat yields perfect prediction on his/her survival. If one fits a logistic regression model on survival based on whether `boat` is missing, the apparent accuracy will be nearly 1. In this sense `boat` is more the result of survival, rather than a cause.

<sup>11</sup>Although some study used this attribute to find cabin locations, its large amount of missingness could be a major source of complexity.

$$\begin{aligned}
X\hat{\beta} = & 0.01051937 \\
& +0.212622\text{age} - 0.0005693763(\text{age} - 10)_+^3 + 0.00243494(\text{age} - 22)_+^3 \\
& -0.00296655(\text{age} - 29)_+^3 + 0.001247244(\text{age} - 37)_+^3 - 0.0001462583(\text{age} - 54.65)_+^3 \\
& -1.976675[\text{Male}] \\
& +4.807962[2\text{nd}] - 1.20667[3\text{rd}] - 2.373247[\text{crew}] \\
& +3.484454[1] \\
& +0.9160294[1] \\
& +0.09633225[\text{English}] + 0.230353[\text{Finnish}] + 0.02481168[\text{Irish}] - 0.2209938[\text{Other}] \\
& -0.2027473[\text{Swedish}] - 0.03927193[\text{Syrian}] \\
& +0.6509122[\text{Cherbourg}] + 0.1320052[\text{Queenstown}] - 0.06351243[\text{Southampton}] \\
& +[\text{Male}][ -0.1405648\text{age} + 0.0004524996(\text{age} - 10)_+^3 - 0.002005646(\text{age} - 22)_+^3 \\
& +0.002351455(\text{age} - 29)_+^3 - 0.0008518177(\text{age} - 37)_+^3 \\
& +0.00005350913(\text{age} - 54.65)_+^3 ] \\
& +[2\text{nd}][ -0.3419093\text{age} + 0.0007375354(\text{age} - 10)_+^3 - 0.003089915(\text{age} - 22)_+^3 \\
& +0.003795193(\text{age} - 29)_+^3 - 0.001665266(\text{age} - 37)_+^3 + 0.0002224522(\text{age} - 54.65)_+^3 ] \\
& +[3\text{rd}][ -0.1350322\text{age} + 0.0004403169(\text{age} - 10)_+^3 - 0.002559262(\text{age} - 22)_+^3 \\
& +0.003927355(\text{age} - 29)_+^3 - 0.002087077(\text{age} - 37)_+^3 + 0.0002786668(\text{age} - 54.65)_+^3 ] \\
& +[\text{crew}][ 0.05036612\text{age} + 0.0000222523(\text{age} - 10)_+^3 - 0.0005447881(\text{age} - 22)_+^3 \\
& +0.001158836(\text{age} - 29)_+^3 - 0.0007325991(\text{age} - 37)_+^3 \\
& +0.00009629903(\text{age} - 54.65)_+^3 ] \\
& +[\text{Male}][ -0.4116912 [2\text{nd}] + 2.182068 [3\text{rd}] + 0.5489062 [\text{crew}] ] \\
& +[1][ -0.1204537\text{age} - 0.0004222734(\text{age} - 10)_+^3 + 0.005086125(\text{age} - 22)_+^3 \\
& -0.01293022(\text{age} - 29)_+^3 + 0.01045057(\text{age} - 37)_+^3 - 0.002184206(\text{age} - 54.65)_+^3 ]
\end{aligned}$$

and  $[c] = 1$  if subject is in group  $c$ , 0 otherwise;  $(x)_+ = x$  if  $x > 0$ , 0 otherwise

## Appendix C. Computing environment

```
sessionInfo()
```

```
R version 4.0.2 (2020-06-22)
Platform: x86_64-w64-mingw32/x64 (64-bit)
Running under: Windows 10 x64 (build 18363)
```

```
Matrix products: default
```

```
locale:
```

```
[1] LC_COLLATE=English_United States.1252
[2] LC_CTYPE=English_United States.1252
```



```
[3] LC_MONETARY=English_United States.1252
[4] LC_NUMERIC=C
[5] LC_TIME=English_United States.1252
system code page: 936
```

attached base packages:

```
[1] stats      graphics  grDevices  utils      datasets  methods   base
```

other attached packages:

```
[1] ggribges_0.5.2  scales_1.1.1    rpart_4.1-15    patchwork_1.0.1
[5] mice_3.11.0     rms_6.0-1       SparseM_1.78    Hmisc_4.4-1
[9] Formula_1.2-4   survival_3.1-12 lattice_0.20-41 ggplot2_3.3.2
[13] dplyr_1.0.2
```

loaded via a namespace (and not attached):

```
[1] tidyr_1.1.2      splines_4.0.2    assertthat_0.2.1
[4] latticeExtra_0.6-29 ymisc_0.0.0.9000 yaml_2.2.1
[7] pillar_1.4.6     backports_1.2.0  quantreg_5.75
[10] glue_1.4.2       digest_0.6.27    RColorBrewer_1.1-2
[13] checkmate_2.0.0   colorspace_1.4-1 sandwich_3.0-0
[16] plyr_1.8.6        htmltools_0.5.0  Matrix_1.2-18
[19] conquer_1.0.2     pkgconfig_2.0.3  broom_0.7.2
[22] bookdown_0.21     purrr_0.3.4      mvtnorm_1.1-1
[25] jpeg_0.1-8.1      MatrixModels_0.4-1 htmlTable_2.1.0
[28] tibble_3.0.4      rtticles_0.17    farver_2.0.3
[31] generics_0.1.0    ellipsis_0.3.1   TH.data_1.0-10
[34] withr_2.3.0       nnet_7.3-14      cli_2.1.0
[37] magrittr_1.5      crayon_1.3.4     polyspline_1.1.19
[40] evaluate_0.14     fansi_0.4.1      nlme_3.1-148
[43] MASS_7.3-51.6     foreign_0.8-80   tools_4.0.2
[46] data.table_1.13.2 hms_0.5.3        lifecycle_0.2.0
[49] matrixStats_0.57.0 multcomp_1.4-14  stringr_1.4.0
[52] rpart.plot_3.0.9  munsell_0.5.0    cluster_2.1.0
[55] compiler_4.0.2    rlang_0.4.8      grid_4.0.2
[58] rstudioapi_0.11   htmlwidgets_1.5.2 labeling_0.4.2
[61] base64enc_0.1-3   rmarkdown_2.5    gtable_0.3.0
[64] codetools_0.2-16  R6_2.5.0         gridExtra_2.3
[67] zoo_1.8-8         knitr_1.30       readr_1.4.0
[70] stringi_1.5.3     Rcpp_1.0.5       vctrs_0.3.4
[73] png_0.1-7         tidyselect_1.1.0 xfun_0.19
```

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