Inclusion-Exclusion Principle:

$$P(A_1 \cup A_2 \cup \dots \cup A_n) = \sum_{i=1}^n P(A_i) - \sum_{i \le i_1 < i_2 \le n} P(A_{i_1} A_{i_2}) + \dots + (-1)^{r+1} \sum_{1 \le i_1 < \dots < i_r \le n} P(A_{i_1} \dots A_{i_r}) + \dots + (-1)^{n+1} P(A_1 \dots A_n)$$

General Multiplication Rule: $P(A_1A_2\cdots A_n)=P(A_1)P(A_2|A_1)P(A_3|A_1A_2)\cdots P(A_n|A_1A_2\cdots A_{n-1})$

Total Probability: $P(B) = P(B|A)P(A) + P(B|A^C)P(A^C)$

Bayes' formula: Events A_1, \dots, A_n partitions sample space, assume $P(A_i) > 0$ for $1 \le i \le n$. Let B be any event, then for any $1 \le i \le n$, we have $P(A_i|B) = \frac{P(B|A_i)P(A_i)}{P(B|A_1)P(A_1) + \dots + P(B|A_n)P(A_n)}$

Probability mass function: $p_X(x) = \begin{cases} P(X = x) & \text{if } x = x_1, x_2, \cdots \\ 0 & \text{otherwise} \end{cases}$

Cumulative distribution function: $F_X(x) = P(X \le x)$ for $x \in \mathbb{R}$

Expected Value: $E(X) = \sum_{x} x p_X(x), E[g(x)] = \sum_{i} g(x_i) p_X(x_i) = \sum_{x} g(x) p_X(x)$

Tail Sum Formula: $E(X) = \sum_{k=1}^{\infty} P(X \ge k) = \sum_{k=0}^{\infty} P(X > k)$

Variance: $var(X) = E(X - \mu)^2 = E(X^2) - [E(X)]^2$

Expected Value of Sum of RV: $E\left[\sum_{i=1}^{n} X_i\right] = \sum_{i=1}^{n} E[X_i]$

Bernoulli random variable: Be(p), X = 1 if success, 0 if failure.

$$P(X = 1) = p, P(X = 0) = 1 - p,$$
 $\mathbb{E}(X) = p, \text{var}(X) = p(1 - p)$

Binomial random variable: Bin(n, p), X = # of successes in n Bernoulli(p) trials.

For
$$0 \le k \le n$$
, $P(X = k) = \binom{n}{k} p^k q^{n-k}$ $\mathbb{E}(X) = np, \text{var}(X) = np(1-p)$

Geometric random variable: Geom(p), X = # of Bernoulli(p) trials required to obtain the first success.

For
$$k \ge 1$$
, $P(X = k) = pq^{k-1}$ $\mathbb{E}(X) = \frac{1}{p}$, $var(X) = \frac{1-p}{p^2}$.

OR, X' = # of failures in Bernoulli(p) trials to obtain 1st success. X = X' + 1

For
$$k \ge 0$$
, $P(X' = k) = pq^k$, $\mathbb{E}(X') = \frac{1-p}{p}$, $var(X') = \frac{1-p}{p^2}$

Negative Binomial random variable: NB(r,p), X=# of Bernoulli(p) trials required to obtain r success.

For
$$k \ge r$$
, $P(X = k) = \binom{k-1}{r-1} p^r q^{k-r}$, $\mathbb{E}(X) = \frac{r}{p}$, $var(X) = \frac{r(1-p)}{p^2}$

Note that
$$Geom(p)=NB(1,p),$$
 $\binom{k-1}{r-1}=(-1)^{r-1}\binom{-(k-r+1)}{r-1}$

Poisson Random Variable: $X \sim \text{Poisson}(\lambda)$ For $k \geq 0$, $P(X = k) = \frac{e^{-\lambda} \lambda^k}{k!}$, $\mathbb{E}(X) = \lambda$, $\text{var}(x) = \lambda$

Usually if n > 20 and np < 15, $Bin(n, p) \approx Poisson(np)$.

Hypergeometric Random Variable: H(n, N, m), a set of N balls, of which m are red and N - m are blue. We choose n of these balls without replacement, X = # of red balls in sample.

For
$$0 \le x \le \min(m, n)$$
, $P(X = x) = \frac{\binom{m}{x} \binom{N - m}{n - x}}{\binom{N}{n}}$ $\mathbb{E}(X) = \frac{nm}{N}$, $\text{var}(X) = \frac{nm}{N} \left[\frac{(n - 1)(m - 1)}{N - 1} + 1 - \frac{nm}{N} \right]$

Expectation and Variance of Continuous RV:

$$\mathbb{E}(X) = \int_{-\infty}^{\infty} x f_X(x) dx, \text{ var}(X) = \int_{-\infty}^{\infty} (x - \mathbb{E}(X))^2 f_X(x) dx = \mathbb{E}[(x - \mu_X)^2]$$

Tail sum formula:
$$\mathbb{E}(X) = \int_0^\infty P(X > x) dx = \int_0^\infty P(X \ge x) dx$$

Uniform Distribution:
$$X \sim U(a, b), \ f(x) = \frac{1}{b-a}, \ a < x < b, \ \mathbb{E}(X) = \frac{a+b}{2}, \ \text{var}(x) = \frac{(b-a)^2}{12}$$

 $F_X(x) = 0, \ if \ x < a; \ \frac{x-a}{b-a}, \ if \ a \le x < b; \ 1, \ if \ b \le x$

$$\textbf{Normal distribution:} \quad X \sim N(\mu, \sigma^2), \ f_X(x) = \frac{1}{\sqrt{2\pi}\sigma} e^{-\frac{(x-\mu)^2}{2\sigma^2}}, \ -\infty < x < \infty, \ \mathbb{E}(X) = \mu, \text{var}(X) = \sigma^2$$

$$\begin{array}{ll} \textbf{Standard normal distribution:} & X \sim N(0,1), \ \phi(x) = \frac{1}{\sqrt{2\pi}} e^{-\frac{x^2}{2}}, \ \Phi(x) = \frac{1}{\sqrt{2\pi}} \int_{-\infty}^x e^{-\frac{y^2}{2}} dy \\ P(a < Z < b) = \Phi(b) - \Phi(a), \ P(Z < b) = \Phi(b), \ \Phi(-x) = 1 - \Phi(x) \\ \frac{X - \mu}{\sigma} \sim N(0,1), \ \therefore Y \sim N(\mu,\sigma^2) \Rightarrow P(a < Y \leq b) = \Phi\left(\frac{b - \mu}{\sigma}\right) - \Phi\left(\frac{a - \mu}{\sigma}\right) \end{array}$$

Exponential distribution:
$$X \sim Exp(\lambda), \ f_X(x) = \lambda e^{-\lambda x}, \ x \ge 0$$
, the c.d.f is $F_X(x) = 1 - e^{-\lambda x}, \ x > 0$
 $\mathbb{E}(X) = \frac{1}{\lambda}, \ \text{var}(X) = \frac{1}{\lambda^2}$

memoryless property of exp dist: $P(X > s + t | X > s) = P(X > t), \ s, t > 0$

Gamma distribution:
$$X \sim \Gamma(\alpha, \lambda), \ f_X(x) = \frac{\lambda e^{-\lambda x} (\lambda x)^{\alpha - 1}}{\Gamma(\alpha)}, \ x \ge 0, \ \text{where } \lambda, \alpha > 0, \ \Gamma(\alpha) = \int_0^\infty e^{-y} y^{\alpha - 1} dy$$
 $\mathbb{E}(X) = \frac{\alpha}{\lambda}, \ \text{var}(X) = \frac{\alpha}{\lambda^2}, \ \Gamma(1) = 1, \ \Gamma(\alpha) = (\alpha - 1)\Gamma(\alpha - 1), \ \Gamma(n) = (n - 1)!, \ \Gamma(1, \lambda) = \text{Exp}(\lambda), \ \Gamma(1/2) = \sqrt{\pi}$

Beta distribution:
$$X \sim \text{Beta}(a,b), \ f(x) = \frac{1}{B(a,b)}x^{a-1}(1-x)^{b-1}, \ 0 < x < 1, \text{ where beta function } B(a,b) = \int_0^1 x^{a-1}(1-x)^{b-1} dx, \ \mathbb{E}(X) = \frac{a}{a+b}, \ \text{var}(X) = \frac{ab}{(a+b)^2(a+b+1)}, \ B(a,b) = \frac{\Gamma(a)\Gamma(b)}{\Gamma(a+b)}$$

Cauchy distribution:
$$X \sim \text{Cauchy}(\theta), \ f(x) = \frac{1}{\pi} \frac{1}{1 + (x - \theta)^2}, \ -\infty < x < \infty, \ \mathbb{E}(X) = \infty, \ \text{var}(X) = \infty$$

De Moivre-Laplace Limit Thm: $X \sim \text{Bin}(n, p)$, then for any a < b, $\text{Bin}(n, p) \approx N(np, npq)$

$$P\left(a < \frac{X - np}{\sqrt{npq}} \le b\right) \approx \Phi(b) - \Phi(a)$$

$$\begin{aligned} & \textbf{Continuity Correction:} \quad X \sim \text{Bin}(n,p), \ Z \sim N(0,1), \text{ then} \\ & P(a \leq X \leq b) \approx P\left(\frac{a - 0.5 - np}{\sqrt{npq}} \leq Z \leq \frac{b + 0.5 - np}{\sqrt{npq}}\right), \ P(a < X < b) \approx P\left(\frac{a + 0.5 - np}{\sqrt{npq}} \leq Z \leq \frac{b - 0.5 - np}{\sqrt{npq}}\right) \end{aligned}$$

Dist of a func of a RV: For monotonic Y = g(X), $f_Y(y) = f_X\left(g^{-1}(y)\right) \left| \frac{d}{dy}g^{-1}(y) \right|$ If X is a RV with c.d.f F, then $F(X) \sim U(0,1)$.

Marginal distribution function: $F_X(x) = \lim_{y \to \infty} F_{X,Y}(x,y) = \lim_{y \to \infty} P(X \le x, Y \le y) = P(X \le x)$. (c.d.f of X)

$$\textbf{Marginal p.m.f:} \ p_X(x) = P(X=x) = \sum_{y \in \mathbb{R}} p_{X,Y}(x,y) \qquad \textbf{Marginal p.d.f:} \ f_X(x) = \int_{-\infty}^{\infty} f_{X,Y}(x,y) dy$$

Relation between p.d.f and c.d.f: $f_{X,Y}(x,y) = \frac{\partial^2}{\partial x \partial y} F_{X,Y}(x,y)$.

Independent: $p_{X,Y}(x,y) = p_X(x)p_Y(y), F_{X,Y}(x,y) = F_X(x)F_Y(y)$

Sum of Indep:
$$F_{X+Y}(x) = \int_{-\infty}^{\infty} F_X(x-t) f_Y(t) dt = \int_{-\infty}^{\infty} F_Y(x-t) f_X(t) dt$$
,

$$f_{X+Y}(x) = \int_{-\infty}^{\infty} f_X(x-t)f_Y(t)dt = \int_{-\infty}^{\infty} f_Y(x-t)f_X(t)dt$$

Some conclusions: X_1, \dots, X_n be n independent RV $\sim \text{Exp}(\lambda)$, then $\sum_{i=1}^{n} X_i \sim \text{Gamma}(n, \lambda)$.

 X_1, \dots, X_n be *n* independent RV~ $N(\mu_i, \sigma_i^2)$, then $\sum_{i=1}^n X_i \sim N\left(\sum_{i=1}^n \mu_i, \sum_{i=1}^n \sigma_i^2\right)$. (used to approx Binominal Dist.)

Sum of Discrete RV: $X \sim \text{Poisson}(\lambda), Y \sim \text{Poisson}(\mu)$, then $X + Y \sim \text{Poisson}(\lambda + \mu)$ $X \sim \text{Bin}(n,p), Y \sim \text{Bin}(m,p), \text{ then } X+Y \sim \text{Bin}(n+m,p)$ $X \sim \text{Geom}(p), Y \sim \text{Geom}(p), \text{ then } X+Y \sim NB(2,p)$

Conditional Dist.: (Discrete:) $p_{X|Y}(x|y) = \frac{p_{X,Y}(x,y)}{p_Y(y)}, \ F_{X|Y}(x|y) = P(X \le x|Y=y)$ (Cont.:) $f_{X|Y}(x|y) = \frac{f_{X,Y}(x,y)}{f_{Y}(y)}, \ F_{X|Y}(x|y) = P(X \le x|Y=y) = \int_{-\infty}^{x} f_{X|Y}(t|y)dt$

Joint p.d.f of Func of RV: $J(x,y) = \frac{\partial(u,v)}{\partial(x,y)} = \frac{\partial g}{\partial x} \frac{\partial h}{\partial y} - \frac{\partial g}{\partial y} \frac{\partial h}{\partial x}, \ f_{U,V}(u,v) = f_{X,Y}(x,y)|J(x,y)|^{-1}$

Expectation of Sum of RV: $\mathbb{E}[g(X,Y)] = \sum_{X} \sum_{x} g(x,y) p_{X,Y}(x,y), \ \mathbb{E}[g(X,Y)] = \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} g(x,y) f_{X,Y}(x,y) dxdy$

Covariance: $cov(X,Y) = \mathbb{E}(X - \mu_X)(Y - \mu_Y)$, if $cov(X,Y) \neq 0$, then X,Y are correlated.

$$cov(X,Y) = \mathbb{E}(XY) - \mathbb{E}(X)\mathbb{E}(Y), \ cov\left(\sum_{i=1}^{n} a_i X_i, \sum_{j=1}^{m} b_j Y_j\right) = \sum_{i=1}^{n} \sum_{j=1}^{m} a_i b_j cov(X_i, Y_j),$$

$$\operatorname{var}\left(\sum_{k=1}^{n}X_{k}\right) = \sum_{k=1}^{n}\operatorname{var}(X_{k}) + 2\sum_{1 \leq i < j \leq n}\operatorname{cov}(X_{i}, X_{j}), \quad \text{under indep., } \operatorname{var}\left(\sum_{k=1}^{n}X_{k}\right) = \sum_{k=1}^{n}\operatorname{var}(X_{k})$$

Independent Case: X, Y independent, then $\mathbb{E}[g(X)h(Y)] = \mathbb{E}[g(X)]\mathbb{E}[h(Y)]$, cov(X, Y) = 0 (reverse not true)

correlation coefficient: $\rho(X,Y) = \frac{\text{cov}(X,Y)}{\sqrt{\text{var}(X)\text{var}(Y)}}$

Conditional Expectation: $\mathbb{E}[X|Y=y] = \sum x p_{X|Y}(x|y) = \int_{-\infty}^{\infty} x f_{X|Y}(x|y) dx$, $\mathbb{E}\left[\sum_{k=1}^{n} X_k | Y=y\right] = \sum_{k=1}^{n} \mathbb{E}[X_k | Y=y]$ y

Expectation by Conditioning: $\mathbb{E}[X] = \mathbb{E}[\mathbb{E}(X|Y)] = \sum_{x} \mathbb{E}(X|Y=y)P(Y=y) = \int_{-\infty}^{\infty} \mathbb{E}(X|Y=y)f_Y(y)dy$

Probability by Conditioning: $P(A) = \sum_{y} P(A|Y=y)P(Y=y) = \int_{-\infty}^{\infty} P(A|Y=y)f_Y(y)dy$

conditional variance: $var(X|Y) = \mathbb{E}[(X - \mathbb{E}[X|Y])^2|Y], \quad var(X) = \mathbb{E}[var(X|Y)] + var(\mathbb{E}[X|Y])$

Moment Generating Function: $M_X(t) = \mathbb{E}[e^{tX}] = \sum_{x,y} e^{tx} p_X(x) = \int_{-\infty}^{\infty} e^{tx} f_X(x) dx$, $\mathbb{E}(X^n) = M_X^{(n)}(0)$

If X, Y are independent, $M_{X+Y}(t) = M_X(t)M_Y(t)$

MGF for dist.: $X \sim \text{Be}(p)$, $M(t) = 1 - p + pe^t$, $X \sim \text{Bin}(n, p)$, $M(t) = (1 - p + pe^t)^n$ $X \sim \text{Geom}(p)$, $M(t) = \frac{pe^t}{1 - (1 - p)e^t}$, $X \sim \text{Poisson}(\lambda)$, $M(t) = \exp(\lambda(e^t - 1))$

 $X \sim U(\alpha, \beta), M(t) = \frac{e^{\beta t} - e^{\alpha t}}{(\beta - \alpha)t}, \quad X \sim \text{Exp}(\lambda), M(t) = \frac{\lambda}{\lambda - t} \text{ for } t < \lambda, \quad X \sim N(\mu, \sigma^2), M(t) = e^{\mu t + \frac{\sigma^2 t^2}{2}}$