Inflation Expectations and Macro Dynamics under Finite Horizon Planning

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 $^{^1}$ This does not express the view of the Federal Reserve Board or the Federal Reserve System.

- Macroeconomists have increasingly begun to incorporate behavioral elements into their models as an alternative to rational expectations.
- Evidence from (conensus) survey data on expectations:
 - Coibion & Gorodnichenko (2015): forecasts underreact to new information.
 - Angeletos, Huo, and Sastry (2020): underreaction is followed by overreaction.
- ➤ Gust, Herbst, Lopez-Salido (2021): the finite horizon planning (FHP) model of Woodford (2018) is an attractive alternative to RE (and some imperfect expectations) models for fitting key aggregate macro time series.
- ► This paper: Can the FHP model fit key facts on expectations? Yes!

Summary of Results

- 1. We derive analytical results for a simplified FHP model for the CG and AHS conditions for inflation expectations.
 - We show that these moments depend both on the parameters governing expectations formation and other structural parameters.

- 2. We estimate an FHP DSGE model:
 - Using data on inflation expectations, the FHP model remains preferable to alternatives in terms of statistical fit.
 - ► The FHP model is also consistent with CG and AHS conditions for inflation expectations. Other models are not.

I. Theory

A Simplified FHP Model

▶ k-level planner making a decision at date t only looks forward through period t+k. For any endogenous variable in periods t+k-j, Z_{t+k-j} , with j=0,1,2,...,k:

$$\mathbb{E}_t^k Z_{t+k-j} = E_t Z_{t+k-j}^j, \tag{1}$$

As shown in Woodford (2018), under these assumptions, firms' price-setting behavior implies a log-linearized relationship for each period of the plan given by:

$$\pi_{\tau}^{j} = \beta E_{\tau} \pi_{\tau+1}^{j-1} + \kappa y_{\tau}, \tag{2}$$

where $\tau = t + k - j$ denotes the planning period, and $1 < j \le k$.

Assume the output gap y_t follows an AR(1) process:

$$y_t = \rho y_{t-1} + e_t \tag{3}$$



▶ Iterating forward, we have:

$$\pi_t^k = \kappa E_t \sum_{i=0}^{k-1} \beta^i y_{t+i} + \beta^k E_t \pi_{t+k}^0$$
 (4)

Aggregate inflation depends on expected of output gap, and the expected inflation rate at the end of the planning horizon.

Firms use continuation value functions to assign value to events outside of their planning horizons (i.e., the longer-run from their viewpoint).

$$\pi_{t+k}^0 = \kappa y_{t+k} + \beta (1 - \theta) v_{pt}, \tag{5}$$

where v_{pt} is the (log-linearized) continuation value to the plans of firms.

Firms to learn and update their beliefs based on past experience. In this case, the value function v_{pt} evolves according to:

$$v_{pt+1} = (1 - \gamma_p)v_{pt} + \gamma_p v_{pt}^e$$
, with $v_{pt}^e = (1 - \theta)^{-1} \pi_t^k$. (6)

Taking stock

One can show that dynamic of inflation under FHP can be written as:

$$\pi_t^k = [1 - (\beta \rho)^{k+1}] \pi_t^{RE} + \beta^{k+1} (1 - \theta) \nu_{\rho t}$$
 (7)

- Since $0 < 1 (\beta \rho)^{k+1} \le 1$, inflation in the FHP model is less responsive to fluctuations in the output gap. This muted responsiveness of inflation is a function of ρ and k.
- ► The second deviation from the RE solution is that firm's longer-run beliefs about inflation, as discussed above, depend on past inflation:

$$v_{\rho t} = \frac{\gamma_{\rho}}{1 - \theta} \sum_{i=0}^{t-1} (1 - \gamma_{\rho})^{i} \pi_{t-1-i}^{k}.$$
 (8)

▶ Inflation in the FHP model displays an excess sensitivity to past inflation relative to the RE solution.



The AHS property

Let
$$\mathbb{F}_{t+1}^k = \pi_{t+1}^k - \mathbb{E}_t^k \pi_{t+1}$$
 .

Proposition (Dynamic Overshooting)

(IRFs of Inflation Forecasts and Forecast Errors). Let $\frac{\partial \mathbb{E}^k_{t+i} \pi_{t+1+i}}{\partial e_t}$ and $\frac{\partial \mathbb{F}^k_{t+1+i}}{\partial e_t}$ for $i \geq 0$ be the impulse response functions for a firm's one-step ahead inflation forecast and forecast error, respectively.

- 1. Without learning: $\frac{\partial \mathbb{E}_{t+i}^{k}\pi_{t+1+i}}{\partial e_{t}} \geq 0$ and $\frac{\partial \mathbb{F}_{t+1+i}}{\partial e_{t}} \geq 0$, $\forall i \geq 0$ and k > 0.
- 2. With learning: If $\gamma_p \leq \frac{1-\rho}{1-\beta^{k+1}}$, there is a threshold forecast horizon, i^* , such that:
 - 2.1 $\frac{\partial \mathbb{E}_{t+i}^k \pi_{t+1+i}}{\partial e_t} \geq 0$ for $i \geq 0$,
 - $2.2 \ \frac{\partial \mathbb{F}_{t+1}}{\partial e_t} > 0 \ \text{and} \ \frac{\partial \mathbb{F}_{t+1+i}}{\partial e_t} < 0 \ \text{for} \ i \geq i^\star,$

The CG property

In the FHP model, a firm's inflation forecast revision can be defined as:

$$\mathbb{R}_{t}^{k} = \left[\mathbb{E}_{t}^{k} - \mathbb{E}_{t-1}^{k}\right] \pi_{t+1}.$$

Proposition (Forecast Error and Revision Correlation)

Let $\beta_{CG} = \frac{cov(\mathbb{R}^k_t, \mathbb{R}^k_{t+1})}{var(\mathbb{R}^k_t)}$ denote the univariate regression coefficient from regressing the one-step ahead forecast error on the forecast revision in the FHP model.

- 1. Without learning: If $\rho > 0$, then $\beta_{CG} > 0$, for any finite planning horizon k > 0.
- 2. With learning: If $\rho = 0$ and $\gamma_p < \frac{1-\beta}{1-\beta^{k+1}}$, then $\beta_{CG} > 0$.

Table: Predictability Regression Results For FHP Model

	Æ	eta c $_{G}$		
	ho = 0	$\rho = 0.8$		
$\gamma_p = 0.3$	-0.01	0.66		
$\gamma_{p}=0.9$	-0.03	0.73		

NOTE: Population coefficient from a regression of one-step ahead inflation forecast errors on forecast revision. For the remaining parameters, k=6, $\beta=0.99$, and $\kappa=0.05$.

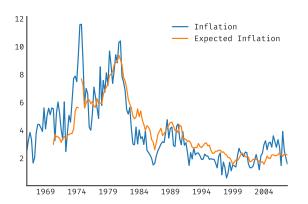
II. Empirics

Why we need system-based estimation

- ▶ We extend to the estimated DSGE of Gust, Herbst, Lopez-Salido (2021) to include data on inflation expectations.
- ➤ The is a methodological departure from CG (2015) and AHS (2020).
- unlike (some) other models of imperfect expectations formation, predictability properties depend on a broader set of structural parameters
- our analysis also allows to understand the extent to which matching the predictability of inflation expectations is also consistent with overall time series fit of inflation, output, short-term interest rates and inflation expectations.

Inflation Expectations

We use the mean forecast for four-quarter-ahead (GDP deflator) inflation expectations from the Survey of Professional Forecasters (SPF).



Assume that this is a noisy measure of model consistent inflation expectations.

- ▶ We also use data on output growth, (GDP deflator) inflation, and the federal funds rate.
- ▶ Our sample is from 1966-2007.
- ▶ Denote the standard macro data as Y and expected inflation data as F.
- Measure of overall fit log marginal data density:

$$\log p(Y, F) = \log p(Y) + \log p(F|Y)$$

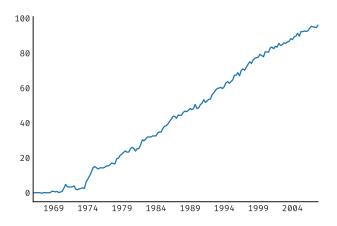
Some Estimation Results

Even with inflation expectations observed, a lot of evidence for low k.

Table: Log MDDs as function of k.

horizon	log p(Y)	log p(F given Y)
k = 0	-720.42	-56.02
k = 1	-716.54	-49.37
k = 2	-718.91	-49.53
k = 3	-721.46	-48.95
k = 4	-723.52	-48.46

$\log p(F|Y)$ relative to the Hybrid New Keynesian Model

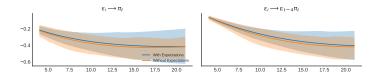


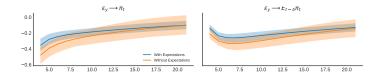
Parameters are remarkably stable

Table: FHP Model k = 1: Key Parameter Estimates

Wit	h Exp. Data	With	out Exp. Data		
Horizon and Learning Parameters					
0.48	[0.32, 0.65]	0.45	[0.30, 0.62]		
0.20	[0.16, 0.24]	0.22	[0.14, 0.30]		
Endogenous Propagation					
0.02	[0.01, 0.04]	0.03	[0.01, 0.05]		
2.67	[1.86, 3.60]	2.75	[1.93, 3.69]		
Persistence of Exogenous Process					
0.87	[0.79, 0.94]	0.87	[0.79, 0.93]		
0.95	[0.91, 0.99]	0.95	[0.91, 0.99]		
0.36	[0.30, 0.43]	0.46	[0.31, 0.61]		
0.93	[0.89, 0.98]				
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And so are impulse responses





Posterior predictive checks

- Let \tilde{y} denote a random variable as distinguished from the realized value y.
- The posterior predictive distribution is given by:

$$p(\tilde{y}|y) = \int p(\tilde{y}|\theta)p(\theta|y)d\theta. \tag{9}$$

► A natural "test" is to reject the model if some important feature of the data lies far in a tail of the predictive distribution.

CG and AHS as Posterior Predictive Checks

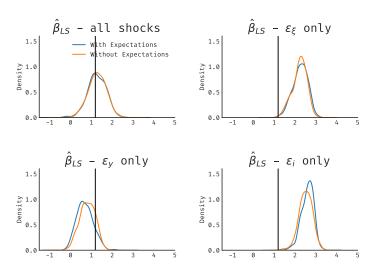
Posterior Predictive Checks For i = 1, ..., N:

- 1. Construct \tilde{y} . Draw $\theta^i \sim p(\theta|Y,F)$, and simulate a single trajectory of $\tilde{y} = \{\Delta y_t, \pi_t, i_t, \mathbb{E}^k_t[\bar{\pi}_{t+4}], \mathbb{E}^k_{t-1}[\bar{\pi}_{t+4}]\}_{t=1}^T$, where T = 168, the length of the actual observables.
- 2. Construct $S_{AHS}(\tilde{y})$ as the point estimates of the impulse response coefficients of $\bar{\pi}_t$ and $\mathbb{E}^k_{t-4}[\bar{\pi}_t]$ to a "supply" shock in a VAR(4) model for $[\Delta y_t, \pi_t, i_t, \mathbb{E}^k_t[\bar{\pi}_{t+4}]]$. The "supply" shock is identified as the shock that maximizes the forecast error variance in inflation over the medium term.
- 3. Construct $S_{CG}(\tilde{y})$. Estimate the regression model,

$$\bar{\pi}_t - \mathbb{E}_{t-4}^k[\bar{\pi}_t] = \alpha + \beta \left(\mathbb{E}_{t-4}^k[\bar{\pi}_t] - \mathbb{E}_{t-5}^k[\bar{\pi}_t] \right) + u_t.$$
 (10)

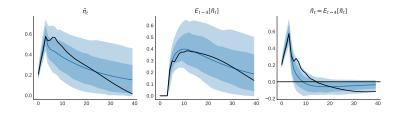
Store the OLS point esitmate of β .

CG Coefficients



AHS

► AHS construct their shock using a SVAR identification: by maximizing the FEVD at "business cycle frequencies." We do the same:



Note, however, we can also do this conditional on our (DSGE) identified shocks.

AHS

We can also count the proportion of trajectories from our posterior predictive checks which satisfy the *dynamic* overshooting property.

Shock	$\mathbb{P}(Dynamic\;Overshooting)$	mean(i*)	std(i*)
AHS	0.86	10.09	5.08
Supply	0.93	9.52	3.50
Monetary Policy	0.52	30.26	8.43
Demand	0.68	22.83	5.44

- ► Like the CG property, the significant heterogeneity across shocks.
- ▶ But in general, the FHP model is consistent with the AHS evidence.

Conclusion

- ► FHP model does a good job at matching key moments related into inflation expectation (revision) predictability while fitting the aggregate time series of inflation expectations well.
- In ongoing work, we show that this performance is very difficult to achieve under alternative models of expectations formation, like diagnostic expectations.
- ► Thanks!