# Same as always: I suck at titles

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9 Abstract

Whether an individual is perceived as knowledgeable by others can be biased by several, and 10 potentially irrelevant factors, ranging from the ephemeral (e.g., how speech is produced) to 11 the situational (e.g., who is speaking). The presence of disfluencies, such as filled pauses 12 (e.g., uh or um in English), elicit evaluations of the speaker as deceitful, less intelligent, or 13 less confident in their knowledge. On the other hand, listeners consider alternative 14 explanations for the speaker to be underinformative: Specifically, non-native speakers are 15 more likely to be forgiven when they fail to be informative. However, whether and how 16 speaker's language profiency affects the interpretation of hesitations and listeners' 17 subsequent behaviour is unclear. In a horse-race betting paradigm, we show that listeners 18 are less likely to follow advice from a speaker who is disfluent, regardless of whether they are disfluent for reasons other than low knowledge. This suggests that previous reported perceptions of knowledgeability elicited explicitly have an impact on how individuals use the 21 information they are given, even in situations when it is not advantageous. Although there is 22 evidence for pragmatic leniance towards non-native speakers, it may be the case that this 23 only applies to failures in language skills that are believed to require high proficiency (e.g., irony), whilst hesitation phenomena may not be subject to such forgiveness. Overall, our 25 results align with a broader body of literature suggesting that the interpretation of 26 hesitation phenomena is not context-dependent. 27

Keywords: filled pauses, disfluencies, non-native-accented speech, feeling of another's knowing

30 Word count: X

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Impressions of other people arise naturally and automatically (Uleman, Adil Saribay, 32 & Gonzalez, 2008). Speakers can be evaluated not only by what they say but how they say 33 it. For example, vocal features of speech can affect perceptions of confidence and persuasion (Guyer, Fabrigar, & Vaughan-Johnston, 2019), status and solidarity (Pittam & Gallois, 1986), and even attractiveness (Feinberg, Jones, Little, Burt, & Perrett, 2005). One repeatedly reported feature of speech said to bias listeners' evaluations is its fluency. 37 Disfluencies, such as uh or um in English can lead to poorer evaluations of the speaker in terms of intelligence (Christenfeld, 1995), competence (Norton-Ford & Hogan, 1980), and certainty of their own knowledge (Brennan & Williams, 1995). Interestingly, speech produced with an accent different from one's own - particularly a foreign accent - can also elicit negative perceptions of the speaker, including worse evaluations in terms of status, solidarity, and credibility (Gluszek & Dovidio, 2010; Gluszek, Newheiser, & Dovidio, 2011; Lev-Ari & Keysar, 2010; Rakić, Steffens, & Mummendey, 2011). In this study, however, we explore a rather counterintuitive idea: That the co-presence of these two factors (i.e., (dis)fluency and foreign accents) can diminish the negative impact they have individually. It is important to note that it is not our interested whether these cues correlate with a speaker's knowledge or whether listeners believe the speaker to be intentionally misleading: Rather, we are interested in whether, in scenarios where the speaker is expected to be cooperative and knowledgeable, listeners are still biased by irrelevant cues.

There is a growing body of literature showing that listeners speech qualities are taken as reflective of a speaker's knowledgeability. Voices with a slow speech rate, low amplitude and a large F0 range are more likely to be rated as less confident, and produce distinct neural responses (Jiang & Pell, 2015, 2016b, 2016a). Similarly, prosody guides listeners' evaluations of certainty and honesty cross-linguistically (Goupil, Ponsot, Richardson, Reyes, & Aucouturier, 2021). These non-verbal qualities of speech may be accompanied by verbal

markers of hesitation, such as disfluencies (Grosjean & Deschamps, 1975; Jiang & Pell, 2017; Shriberg, Bates, & Stolcke, 1997; Shriberg & Lickley, 1993). Brennan and Williams (1995) demonstrated that filled pauses (e.g., uh, mm, or um) can be perceived as indices of speakers' 59 confidence in their knowledge. In their study, participants listened to previously recorded answers to trivia questions (without hearing the questions) and were asked to rate how likely 61 each speaker would be to recognise the correct answer to the question i.e., feeling of 62 another's knowing, FOAK. Amongst the cues that biased participants' assessments were filled pauses: Answers containing a disfluency were more likely to receive lower FOAK ratings, suggesting that filled pauses were taken as reflective of speakers' reduced certainty about their knowledge. In contrast, non-answers (i.e., 'I don't know') were more likely to receive a higher FOAK rating if preceded by a filled pause. Brennan and Williams (1995) took this as evidence that listeners are sensitive to the surface form of delivery, and in particular, to the cues displayed by speakers when they do not know (or cannot remember at the moment of being asked) the answer to a question (see also Smith & Clark, 1993).

However, certain speech qualities can yield negative evaluations without being 71 explicitly associated with those. One consistent finding is that speakers with an accent 72 different from that of the listeners are more likely to be negatively evaluated (e.g., 73 (Dragojevic & Giles, 2016; Gluszek & Dovidio, 2010) and statements produced with a foreign-accent are more likely to receive lower ratings of credibility compared to statements produced with a native accent (Lev-Ari & Keysar, 2010). This divide in evaluations as a function of a speaker's accent has been accounted for in terms of processing fluency (Lev-Ari & Keysar, 2010), group membership, and a combination of both (Mai & Hoffmann, 2013). Alternatively, but not exclusively, it is possible that non-native speakers are evaluated differently from native speakers due to the expectations listeners have about their competence: Non-native speakers' accents may invoke stereotypes whereby these speakers are believed to be less competent than native speakers (Fairchild & Papafragou, 2018; Lev-Ari, 2015). This expectation-based account proposes that stereotypes of non-native

speakers' low linguistic competence affect both how their speech is comprehended and interpreted. Specifically, signs that can be attributed to low linguistic competence would be interpreted as such when produced by a non-native speaker and be attributed to other factors when produced by a native speaker. In line with this hypothesis, a speaker's nativeness affects how ironic a statement is perceived to be (Bazzi, Brouwer, & Foucart, 2022; Caffarra, Michell, & Martin, 2018), why a statement is believed to be underinformative (Fairchild, Mathis, & Papafragou, 2020; Fairchild & Papafragou, 2018; Ip & Papafragou, 2022; Lorenzoni, Pagliarini, Vespignani, & Navarrete, 2022) and why non-native speech is processed shallower (Hanuíková, Alphen, Goch, & Weber, 2012; Lev-Ari & Keysar, 2012).

Thus far, the expectations-based account has explained findings pertaining how the 93 content of speech is interpreted. However, stereotypes on linguistic competence may also affect how speech is delivered. In the case of confidence and knowledgeability, nativeness has 95 been shown to affect how confident a speaker is perceived (Caballero & Pell, 2020; Jiang, Gossack-Keenan, & Pell, 2019). Given that speakers producing speech in their second language are more disfluent (Bergmann, Sprenger, & Schmid, 2015; Davies, 2003; Gkalitsiou & Werle, 2023) and is perceived as such by native listenes (Pinget, Bosker, Quené, & De Jong, 2014), it could be possible that a speaker's nativeness leads to different interpretations 100 of disfluencies. This would align with findings in the disfluency literature whereby the effects 101 of those in speech comprehension are dependent on who produces it (Arnold, Kam, & 102 Tanenhaus, 2007; Barr & Sevfeddinipur, 2010; Heller, Arnold, Klein, & Tanenhaus, 2015), 103 including speaker's nativeness (Bosker, Quené, Sanders, & De Jong, 2014), allegedly because 104 listeners are sensitive to the speaker's mental state and the reasons for them to be disfluent. 105

Recently, Matzinger, Pleyer, and Żywiczyński (2023) explored whether listeners'
perceptions of why a speaker was disfluent differed for native and non-native speakers. In
their study, participants listened to native and non-native speakers answer trivial questions
and requests, and were explicitly asked to rate the speaker's knowledge and confidence (for

FOAK) and their willingness to grant the request. Crucially, Matzinger et al. (2023) manipulated speakers' fluency by having answers prefaced with either short (200 ms) or long 111 (1200 ms) pauses. For requests, long pauses were less likely to be associated with 112 unwillingness for non-native compared to native speakers. However, FOAK ratings did not 113 differ between speakers: Long pauses produced by either speaker were likely to be taken as 114 reflecting low confidence and low knowledge. Matzinger et al. (2023) attribute this pattern 115 to different conversational contexts: Requests tap into speakers' cooperativeness, and thus in 116 this context, tuning to the interlocutor's mental state and stereotyping might be more 117 relevant than evaluating the speaker's competence (i.e., knowledge). 118

To date, most experimental evaluations of manner of speech have been explicit. 119 Participants are asked to rate particular traits of speakers on a scale. Further, these ratings 120 are elicited in non-social contexts, where participants do not have anything at stake. 121 Although these experiments demonstrate that listeners show sensitivity to aspects of speech 122 such as fluency and accent when asked to make explicit judgements about the speaker, they 123 may not explain how listeners evaluate speech implicitly, when to do so is consequential. Here, 124 we propose an implicit measurement of listeners' assessments of the speaker's certainty, using 125 a horse-race paradigm. In this task, participants listen to a set of speakers provide 126 descriptions of horses and are asked to distribute virtual tokens as 'bets' on each horse's 127 likelihood of winning a putative race. This approach presents two advantages over previous 128 experiments. First, in the horse-race paradigm, participants are not explicitly asked to 129 evaluate a certain trait of the speaker (in this case, how knowledgeable they are). Instead, 130 we take participants' allocations of 'betting money' as an indirect measurement of their 131 perceptions of speakers' knowledge. Indeed, pilot studies have shown that individuals are 132 sensitive to this manipulation and that disfluent information leads to smaller bets 133 (Butterworth, 2019). Second, horse races provide a scenario where individuals can make 134 decisions based on what they are told, but the content of speech itself may not be informative 135 for many individuals (in that participants are less familiar with the world of horse racing).

In a pre-registered study, we set up to explore whether and how perceptions of 137 certainty are biased by manner of delivery in the form of fluency and the speaker's identity 138 as conveyed by their accent. We presented participants with recordings of a native and a 139 non-native speaker, each describing two horses, with one description produced fluently and 140 one description produced disfluently. If listeners are sensitive to both local and global causes 141 of hesitations when making judgements about certainty, then disfluent descriptions provided 142 by a native speaker should result in less money bet, reflecting listeners' lower FOAK for the 143 speaker. However, disfluent descriptions provided by a non-native speaker may not impact listeners' betting behaviours, to the extent that they consider the possibility of difficulties in 145 production when assessing the speaker's knowledge. To further control for the potential 146 effects of (non)-nativeness on certainty on its own, we measured participants' language 147 attitudes towards each speaker (see Dragojevic & Giles, 2016), perceived fluency, accentedness, and comprehensibility of the native and the non-native speaker, as well as the perceived reliability of each speaker. We additionally measured participants' familiarity with 150 and exposure to native and non-native-accented English on a daily basis, to account for the 151 fact that exposure to non-native accents can reduce their negative effects on listeners' 152 judgements (Boduch-Grabka & Lev-Ari, 2021). 153

154 Methods

All experimental stimuli can be found at https://osf.io/zsut7/.

# Participants

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We conducted a sensitivity power analysis via data simulation following DeBruine and Barr (2021). We explored the required number of participants for a study with a power higher than .8 for the interaction between fluency and speaker identity. We conducted a 1000 simulations per different combinations of the effect size (ranging from small to medium) and the standard deviation of the residuals. In this analysis, we assumed a medium effect size of fluency and no effect of speaker identity. This analysis showed that a sample size of 360 participants ensured enough power to detect a medium or greater effect size.

In our pre-registration, we set that only participants born and raised, and currently residing in the United Kingdom, with English as their first and only language, and with no auditory disorders could partake in the study. Further, we would exclude from analysis data from participants who reported the experiment's aim or manipulation, rated the naturalness of the auditory stimuli (defined as how likely they believed the audios to have been recorded in one go) lower than four, or considered themselves experts in horse races. This meant that we recruited 641 participants for a sample of 360. Participants were recruited via the online platform Prolific, and were reimbursed £1.50 for a 10-minutes experiment.

#### Visual stimuli 172

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We selected a set of eight images of racehorses from the web. The selected images each featured only one racehorse in the foreground, in motion, ridden by a jockey, and the horses all took up approximately the same proportion of the image. To ensure that the pictures did not bias participants' bets, we recruited ten participants on Prolific, who did not take part in the main study, and asked them to rate on a 10-point scale how likely each horse was to win a hypothetical race and to rank the horses in the order they thought they would cross the finish line, in exchange for £0.45.

A one-way repeated measures ANOVA showed that there were no differences in how 180 likely each horse was thought to be to win a race (F(7) = 0.5, p = .83). An ordinal logistic regression showed that none of the horses were more likely to be ranked differently from the 182 others (all |t| < 2). We therefore selected four out of the eight images as visual stimuli. 183

### Auditory stimuli

We used four descriptions of racing horses from The Racing Post, retrieved on 185 October 2018, originally edited and used by Butterworth (2019). Each passage consisted of 186 three to four sentences describing a horse and its performance in previous races. 30 British 187 English speakers, who did not participate in the final experiment, rated these passages to 188 ensure that all descriptions were perceived as equally likely to describe a winning horse. 189

Participants rated on a 10-point scale how likely they thought each horse was to win a race individually, and then ranked all the horses, in exchange for £0.50.

Based on these results, we further edited the descriptions as some descriptions were more likely to be rated as winning horses. A new sample fo 30 British English speaker rated the edited passages. A one-way repeated measures ANOVA showed no differences in each horse's rated likelihood of winning a race (F(3) = 0.76, p = .52), nor in the order in which they were ranked (all |t| < 2). The final set of descriptions for the experiment can be found in Table 1. Each description was paired with one of the four visual stimuli.

We then recorded a British English native speaker and a non-native (L1: Italian) 198 English speaker to create the auditory stimuli. Both speakers were female. Passages were 199 recorded one at a time. To elicit naturally disfluent recordings, both speakers were 200 instructed to read the passages silently and then were recorded as they tried to recall the 201 passage from memory. To avoid differences between the descriptions provided by the 202 speakers, they were allowed to look at the descriptions as they spoke if they could not 203 remember the continuation. We edited these recordings using Audacity to ensure that the 204 recordings of each speaker had similar numbers of disfluencies in similar locations, by 205 cross-splicing different recordings (see Table 2). To create the fluent counterpart of each description, filled and mid-utterance silent pauses were excised, and elongations and 207 between-clause silent pauses were reduced using the 'Tempo' function. The final auditory experimental stimuli consisted, for each speaker, of two descriptions of each horse (one disfluent, one fluent), resulting in sixteen recordings. Description, fluency, speaker, and order 210 of presentation were counterbalanced in a Latin Square design, resulting in 24 lists. 211

To ensure that the resulting descriptions were perceived to be natural (i.e., our edited 'fluent' audios were not clearly edited) and that they were perceived as differing in fluency (i.e., disfluent and 'fluent' versions were distinguishable), we validated them in a sample of 48 British English participants, who did not take further part in the study. Participants were

Table 1

Original description of each horse that speakers were asked to memorise and reproduce.

Horse	Description
	Fire Walker is looking strong thanks to his come-from-behind success
	in the Acomb Stakes. The impression given in both runs is that Fire
Fire	Walker should handle the demands of the extra furlong and Charlie Hills
Walker	is looking forward to the test. The trainer said "He's done really well for
	a little break, his work's been good and I couldn't be more pleased with
	him".
	Silver Sky, a runner-up of a seven-furlong maiden at Naas on his debut
Silver Sky	last month, the son of Invincible Spirit ran crack French colt Persian King
	to a neck in the Group 3 Autumn Stakes over today's trip at Newmarket
	two weeks ago and his trainer believes he has done well since. O'Brien
	said "Silver Sky is a fine big colt and a talented one".
	Apocalypse has put in a string of consistent performances, most recently
	finishing third to Norway in the Zetland. "He's had a very solid year"
Apocalypse	said trainer Archie Watson. "He ran a good race in the Zetland, beaten
	only a length and a quarter, and I think the field here is of a similar level
	so I'm more than happy for him to take his chance".
	Black Blade proved the market all wrong as the complexion of the 6.5-
DI I	furlong novice race changed dramatically in the final two furlongs, with
Black	the Rebel Racing premier-owned newcomer under Tom Queally collaring
Blade	long-time leaded Monsieur Noir. Spencer said: "He did it well. He's a
	nice horse. We always thought had a bright future".

Table 2

Transcription of the disfluent horse descriptions.

Horse	Description
	FP-um Fire Walker is looking SP strong thanks to his (elongation) SP
	come-from-behind success SP FP-er in the Acomb Stakes. FP-um The
Fire	(elongation) impression given SP in both runs FP-uh is that Fire Walker
	should SP handle the demands of SP the extra furlong and FP-uh Charlie
Walker	Hills is looking forward to the test. The trainer said "He's done (elonga-
	tion) FP-um really well for a little break, FP-uh his work's been good
	and I couldn't be more pleased with him".
	Silver Sky, FP-uh a runner-up of -a seven-furlong maiden at Naas on his
	(elongation) debut last month. FP-uh The son of Invincible Spirit SP
	ran (elongation) crack French colt Persian King to a neck in the FP-uh
Silver Sky	Group 3 Autumn Stakes. over today's trip at Newmarket two weeks ago
	FP-uh and his trainer believes he has done well since. O'Brien said FP-uh
	"Silver Sky is a fine big colt and a talented one".
	FP Apocalypse has put in a (elongation-ish) string SP of consistent
	performances, FP most recently finishing third to SP Norway in the
	Zetland. FP "He's had a very solid year" said trainer SP FP Archie
Apocalypse	Watson. "He ran a SP good race in the Zetland, SP FP SP beaten only a
	length and a quarter, and I think the field here SP is of a similar level so
	I'm- I'm more than happy for him to (elongation-ish) take his chance".
	FP-um Black SP Blade proved SP FP-uh the market all wrong as the
Black Blade	complexion SP of the 6.5 furlong SP novice SP race FP-um changed
	dramatically in SP the final two furlong. SP FP-uh with the Rebel
	Racing SP premier-owned SP newcomer under Tom SP Queally FP-um
	SP collaring long-time leaded Monsieur Noir. FP-uh Spencer said: "He
	did it well. FP-uh He's a nice horse. FP-uh We always thought he had a
	bright future".

accent respectively.

Table 3

Mean (standard deviation) ratings of native and non-native speakers' fluent and disfluent recordings for fluency, naturalness, and accentedness, on a 9-point scale where lower values indicate less fluent, less natural, and less

Speaker	Delivery	Fluency	Naturalness	Accentedness
Native Speaker	Fluent	7.62 (1.59)	5.44 (2.62)	4.38 (2.22)
	Disfluent	6.44 (2.36)	5.73 (2.57)	4 (1.97)
Non-native Speaker	Fluent	6.25 (1.73)	5.94 (1.83)	7.02 (1.59)
	Disfluent	5.21 (1.75)	5.83 (1.99)	7.17 (1.51)

allocated to one of the 24 experimental lists, to ensure that our validation procedure was 216 similar to how participants encountered stimuli in the actual experiment. Following Bosker, Quené, Sanders, and De Jong (2014) procedure, participants were asked to rate each audio's 218 fluency on a scale from 1 to 9 (1: not fluent at all, 9: very fluent). We instructed participants to rate fluency by considering silent and filled pauses, speed of speech, and 220 repairs, and to ignore speakers' accents and the content of their speech. Participants 221 additionally rated on a 9-point scale each recording's naturalness (defined as how likely it 222 was that the audio had been recorded in one go; 1: not unlikely at all; 9: very likely), and 223 accentedness (while ignoring the perceived speaker's proficiency in the language; 1: not 224 accented at all, 9: very accented). We additionally asked participants to guess the speakers' 225 country of origin. At the end of the task, participants were further asked how often they 226 interacted with native and non-native English speakers (on a 9-point scale, 1: never, 9: 227 always) and were allowed to report if they noticed anything odd in the auditory stimuli. 228

Table 3 shows the means (and standard deviations) of participants' ratings of fluency, naturalness, and accentedness. A linear mixed model for fluency ratings with fixed effects of

fluency (treatment-coded, reference: fluent), speaker's linguistic background 231 (treatment-coded, reference: native speaker), and their interaction, with random intercepts 232 by participant and by horse description showed that fluency ratings differed significantly for 233 the fluent and disfluent conditions ( $\hat{\beta} = -1.19$ , 95% CI [-1.79, -0.58], t = -3.84). The 234 non-native speaker was perceived as more disfluent than the native speaker ( $\hat{\beta} = -1.37, 95\%$ 235 CI [-1.97, -0.78], t = -4.51), in line with previous findings (Bosker, Quené, Sanders, & 236 Jong, 2014; Pinget et al., 2014), but the interaction between the two variables was not 237 significant ( $\hat{\beta} = 0.15, 95\%$  CI [-0.49, 0.78], t = 0.45). An identical model for naturalness 238 ratings showed no significant differences by fluency ( $\hat{\beta} = 0.29, 95\%$  CI [-0.52, 1.11], 239 t=0.70), speaker's linguistic background ( $\hat{\beta}=0.50, 95\%$  CI [-0.27, 1.27], t=1.27) or their 240 interaction ( $\hat{\beta} = -0.40, 95\%$  CI [-1.49, 0.70], t = -0.71).

242 Procedure

Stimuli were presented using JsPsych (Leeuw, Gilbert, & Luchterhandt, 2023), hosted 243 on MindProbe (via JATOS, Lange, Kühn, & Filevich, 2015). The task began with a cover 244 story introducing two horse racing tipsters. Participants were told the tipsters would provide 245 information about the four most popular horses competing in an upcoming race at 246 Musselburgh Racecourse (Edinburgh). The cover story explained that the two tipsters were 247 well-known experts in the field, and added that one of the speakers was a non-native English 248 speaker (without specifying the nationality of either speaker), introducing the element of the 249 speakers' linguistic backgrounds as well as the factor of competence. 250

At the beginning of the experiment, participants were shown four pictures of the
horses that they had been told would take part in the race, alongside their names.

Participants were instructed to distribute one-hundred pounds in betting money across the
four horses based on the likelihood they thought each horse had of winning: They could split
the bets as they wished, and they did not have to spend all the money. Each participant was
randomly assigned to one out of 24 groups, so that they would listen to each speaker twice,

one in each fluency level. The order in which horses were presented was randomised. In each trial, participants listened to one speaker described a given horse's performance. Once the playback stopped, participants were asked to place a bet by typing a number on a web form. Participants could only move to the next horse's description once they had placed a bet. Participants were allowed to modify their previous bets every time they heard a new description. If the sum of bets made at any point was more than the allotted maximum, they were asked to re-distribute their bets until the total was below one-hundred pounds.

After the betting round was complete, participants completed a questionnaire similar 264 to that of Foucart, Costa, Morís-Fernández, and Hartsuiker (2020) to measure their language 265 attitudes towards the native and the non-native speaker. For each speaker, participants 266 answered six questions measuring affect (three questions for negative affect and three 267 questions for positive affect), five questions measuring solidarity, five questions measuring 268 status, and one question each for comprehensibility, accentedness, fluency, and 269 trustworthiness. Each question used a 9-point scale. Participants first answered questions 270 about, at random, the native or non-native speaker, with the order of presentation of 271 dimensions being randomised. They then answered the same questions for the remaining 272 speaker. We also asked participants to guess the countries of origin of our native and 273 non-native speakers, and from which speaker they would like to learn about horse races in 274 the future. Additional questions included ratings on a 9-point scale of how natural the audio 275 sounded (1: unnatural, edited; 9: natural, unedited), and two questions measuring 276 participants' exposure to native and non-native accented English (1: never; 9: always). 277 Likewise, we measured participants' previous experience with betting and their perceived knowledge of horse races (two questions: Whether they had bet on horse races in the past, and to rate on a 5-point scale, from 'Strongly disagree' to 'Strongly agree', how closely they identified with the statement 'I am an expert on horse races'). Finally, we included an 281 open-ended question for participants to report what guided their decision-making, as well as 282 their perception of the experiment's aim. 283

Table 4

Mean (standard deviation) of money bet

by manner of delivery and speaker's

linguistic background.

Speaker	Delivery	Money bet	
Native	Fluent	24.24 (13.92)	
	Disfluent	20.66 (12.23)	
Non-native	Fluent	24.15 (13.61)	
	Disfluent	22.31 (13.81)	

284 Results

All data pre-processing and analyses were carried out in R version 4.4.1 (R Core
Team, 2024), using the packages tidyverse version 2.0.0 (Wickham et al., 2019), ggplot
version 3.5.1 (Wickham, 2016), and wesanderson version for data wrangling and
visualization. lme4 version 1.1-35.5 (Bates, Mächler, Bolker, & Walker, 2015) was used for
data analysis, and papaja version 0.1.2 (Aust & Barth, 2023) for manuscript write-up.
Scripts can be found at https://osf.io/zsut7/

## Pre-registered analyses

# 292 Betting behaviour

Table 4 and Figure 1 show the mean amount of money bet per speaker and fluency condition. On average, participants bet £25 on each horse. This distribution suggests that participants followed a rational behaviour: Given their lack of expertise, they distributed the one hundred pounds equally. However, disfluent instructions seem to lower the amount of money bet.

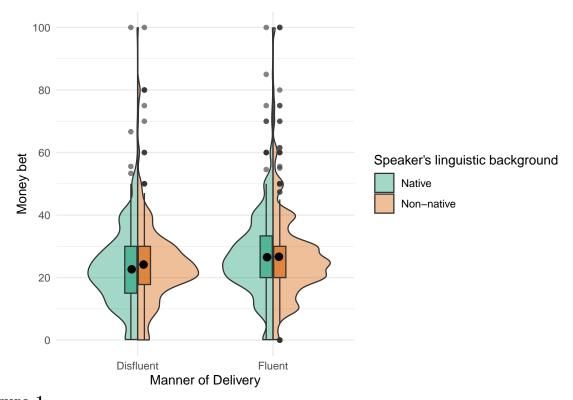


Figure 1
Money distribution by manner of delivery (fluent/disfluent) and speaker (green: native/orange: non-native).

We modelled participants' betting behaviour in a linear mixed model. We modelled 298 money bet on each horse, taken from the final amounts submitted in the experiment, after 299 all four descriptions had been heard and valid responses (summing to at most 100 pounds) 300 had been recorded. The model included fixed effects of fluency (sum-coded; fluent coded as 301 -0.5, disfluent as +0.5), speaker's linguistic background (sum-coded, native coded as -0.5; non-native coded as +0.5), and their interaction. The maximal model (Barr, Levy, Scheepers, 303 & Tily, 2013), with random intercepts by-participant and by-item, with random slopes for fluency and speaker's linguistic background by-participant, and for fluency by-item, failed to 305 converge. We first dropped the random intercept by-participant, as most participants used 306 all the money and thus there was no variance in their intercept. The final model included a 307

random intercept by-item. Results were deemed significant at |t| > 2 (Baayen, 2008).

Our model showed a main effect of fluency, whereby participants placed lower bets following disfluent descriptions compared to their fluent counterparts ( $\hat{\beta} = -2.71$ , 95% CI [-4.04, -1.38], t = -3.99). There was no main effect of speaker ( $\hat{\beta} = 0.78$ , 95% CI [-0.55, 2.11], t = 1.15) and importantly, no interaction between delivery and speaker ( $\hat{\beta} = 1.74$ , 95% CI [-0.92, 4.40], t = 1.28)<sup>1</sup>.

# 314 Language attitudes

Table 5 depicts the mean (and standard deviation) of ratings in each of the measures 315 of interest by speaker's linguistic background. Constructs measured with more than one 316 question (affect, status, solidarity) were obtained by calculating the average score. For the 317 Affect dimension, we reverse-scored items measuring negative affect. Cronbach's alpha 318 showed that the scores for these attitudes were reliable ( $\alpha_{affect} = 0.80$ ,  $\alpha_{status} = 0.93$ , 319  $\alpha_{solidarity} = 0.87$ ). We explored differences between the native and the non-native speaker's 320 evaluations in these three social dimensions as well as on Comprehensibility, Accentedness, 321 and Trustworthiness via paired t-test using Bonferroni correction for p values. 322

Analyses showed that speakers were rared differently across all six dimensions (see
Table 6). The largest differences were, unsurprisingly, in comprehensibility and accentedness,
where the non-native speaker received lower ratings.

Following our pre-registered analysis, we included these six variables in our previous model to explore whether speakers' evaluations could further explain participants' betting behaviours. This second model improved model fit ( $\chi^2(6) = 19.46$ , p < .01). Besides a main effect of manner of delivery ( $\hat{\beta} = -2.71$ , 95% CI [-4.03, -1.39], t = -4.02), the model

<sup>&</sup>lt;sup>1</sup> An identical model including excluded participants showed a main effect of fluency ( $\hat{\beta} = -2.71$ , 95% CI [-4.04, -1.38], t = -3.99), and no other significances (for speaker,  $\hat{\beta} = 0.78$ , 95% CI [-0.55, 2.11], t = 1.15; for the interaction  $\hat{\beta} = 1.74$ , 95% CI [-0.92, 4.40], t = 1.28)

Table 5

Average score (and standard deviation) in each dimension by speaker.

Dimension	Native Speaker	Non-native Speaker	
Comprehensibility	7.62 (1.57)	5.51 (2.14)	
Accentedness	5.46 (2.1)	6.84 (1.37)	
Affect	5.77 (1.49)	6.05 (1.32)	
Status	5.95 (1.52)	6.43 (1.31)	
Solidarity	5.67 (1.48)	6.41 (1.24)	
Trustworthy	6 (1.61)	6.36 (1.41)	

Table 6  $Paired\ t\ test\ for\ each\ dimension\ between\ speakers.$ 

Dimension	t(359)	95% CI	d
Comprehensibility	15.99	[1.85, 2.36]	2.11
Accent	-11.51	[-1.62, -1.15]	-1.39
Affect	-3.54	[-0.44, -0.13]	-0.29
Status	-6.55	[-0.62, -0.33]	-0.48
Solidarity	-9.63	[-0.89, -0.59]	-0.74
Trustworthy	-4.12	[-0.53, -0.19]	-0.36

showed an effect of affect, whereby higher ratings of affect were more likely to yield higher bettings ( $\hat{\beta} = 0.81, 95\%$  CI [0.14, 1.49], t = 2.35)

# 332 Exploratory analysis

We additionally explored participants' preferences to learn from either speaker in the future. 206 participants reported they would prefer to learn from the native spaker, and 154 preferred the non-native speaker. A  $\chi^2$  test of goodness of fit showed that there was a significant difference in participants' preferences ( $\chi^2(1) = 7.51$ , p < .01).

Discussion

Listeners can infer speaker's knowledgeability from a range of cues, such as voice 338 pitch or amplitude. A speaker's speech fluency i.e., whether they produces disfluencies as 339 they speak, has been previously shown to impact how confident in their knowledge a speaker is judged to be (Brennan & Williams, 1995). However, speakers can be disfluent for reasons 341 other than (un)confidence: For example, speaking in own's second language is also associated 342 with an increase in disfluencies (De Jong, Groenhout, Schoonen, & Hulstijn, 2015; Derwing, 343 Munro, Thomson, & Rossiter, 2009). In fact, it has been proposed that second language 344 speakers are stereotyped and thus expected to display low linguistic performance (Lev-Ari, 345 2015), which consequently leads native listeners to "forgive" what otherwise would be lead to 346 a negative evaluation (Fairchild et al., 2020; Lorenzoni et al., 2022). In the present 347 experiment, we explored whether attributions of (un)knowledgeability as a function of the 348 speaker's nativeness, and if those affected listeners' subsequent behaviours in a task where 349 participants had to place bets on horse. Our findings suggest that manner of delivery, in the 350 form of fluency, was the sole factor that guided participants' behavior: Disfluent descriptions 351 yielded lower bets, regardless of who was speaking. 352

The pattern found aligns with the idea that listeners are sensitive to how an utterance is produced (Brennan & Williams, 1995). Descriptions that included hesitation phenomena, in the form of filled pauses, led to smaller bets compared to their fluent

counterparts. This aligns with Brennan and Williams (1995) findings wherein listeners were 356 less likely to attribute confidence in their knowledge to a speaker when their answers 357 included hesitation phenomena (i.e., Feeling of Another's Knowing). Brennan and Williams 358 (1995) attributed participants' ratings to inferences made about the speaker's mental state -359 specifically, inferences about the degree of confidence a speaker has in their knowledge. In 360 our experiment, since participants have little a priori knowledge that can guide their betting 361 behaviour, and because the utterance is semantically and acoustically identical apart from 362 the excision of disfluencies, an inference about a speaker's confidence based on those 363 disfluencies is an important potential determinant of participants' decisions. 364

This inference explanation would predict that listeners would forgive foreign 365 disfluencies, in line with the expectations-based account. While previous research has 366 reported that whether a filled pause affects processes involved in speech comprehension 367 depending on speaker identity (e.g., Bosker, Quené, Sanders, & De Jong, 2014), recent 368 studies have failed to find this effect when it comes to speaker's attributions (Matzinger et 369 al., 2023). In our experiment, participants distributed money similarly when either speaker 370 provided disfluent descriptions, suggesting that listeners did not weigh in our (non-native) 371 speaker's identity to interpret the disfluency. This is particularly remarkable given that in 372 our post-experimental questionnaire, the native speaker was more likely to be chosen as 373 someone participants would like to learn from about horse races in the future. One 374 possibility for this lack of interaction has to do with the fact that our non-native speaker was 375 introduced as a knowledgeable tipster. This introduction of the speaker as an authoritative 376 figure may have overridden any other features of their identity, including their non-native speaker identity. Indeed, beliefs about a speaker's expertise have been shown to guide 378 perceptions of their certainty (Mol, Kuhlen, Van der Steen, & Obbens, 2013). Considering the language attitudes triggered by both accents, the difference in speaker preference may 380 more likely reflect the ease of comprehending the native speaker, rather than an implicit 381 negative bias towards the non-native speaker, and particularly, a diminished perception of 382

competence for non-native speakers.

The experiment here also introduced a novel approach to measuring how certain the 384 speaker is perceived. While previous studies had participants explicitly rate a speaker on 385 different dimensions (e.g., knowledgeability, trust) in non-social contexts, the present study offers a new venue to explore how different factors bias individuals' evaluations indirectly, as well as the implications of those evaluations. However, a potential shortcoming of our design 388 has to do with how people approached the betting system. Because of the nature of the task, 389 in that we tested participants with no prior knowledge of horse races, our participants 390 behaved rationally: The vast majority employed the hundred pounds allocated and 391 distributed them fairly rationally in a way that maximizes their chances of winning (i.e., 392 around £25 per horse). It is, therefore, possible that in situations of less uncertainty (e.g., a 393 scenario where individuals have more knowledge) or where they are rewarded for allocating 394 money to the winning horse, how the speaker is perceived may yield a larger effect on 395 participants' behaviours. 396

References

- Arnold, J. E., Kam, C. L. H., & Tanenhaus, M. K. (2007). If you say thee uh you are
- describing something hard: The on-line attribution of disfluency during reference
- comprehension. Journal of Experimental Psychology: Learning, Memory, and Cognition,
- 33(5), 914–930. https://doi.org/10.1037/0278-7393.33.5.914
- <sup>402</sup> Aust, F., & Barth, M. (2023). papaja: Prepare reproducible APA journal articles with R
- 403 Markdown. Retrieved from https://github.com/crsh/papaja
- Baayen, R. H. (2008). Analyzing linguistic data: A practical introduction to statistics using r.
- 405 Cambridge University Press. https://doi.org/10.1017/CBO9780511801686
- Barr, D. J., Levy, R., Scheepers, C., & Tily, H. J. (2013). Random effects structure for
- confirmatory hypothesis testing: Keep it maximal. Journal of Memory and Language,
- 68(3), 255–278. https://doi.org/10.1016/j.jml.2012.11.001
- Barr, D. J., & Seyfeddinipur, M. (2010). The role of fillers in listener attributions for speaker
- disfluency. Language and Cognitive Processes, 25(4), 441–455.
- Bates, D., Mächler, M., Bolker, B., & Walker, S. (2015). Fitting linear mixed-effects models
- using lme4. Journal of Statistical Software, 67(1), 1–48.
- https://doi.org/10.18637/jss.v067.i01
- 414 Bazzi, L., Brouwer, S., & Foucart, A. (2022). The impact of foreign accent on irony and its
- consequences on social interaction. Journal of Multilingual and Multicultural
- Development, 1-13.
- Bergmann, C., Sprenger, S. A., & Schmid, M. S. (2015). The impact of language
- co-activation on L1 and L2 speech fluency. Acta Psychologica, 161, 25–35.
- Boduch-Grabka, K., & Lev-Ari, S. (2021). Exposing individuals to foreign accent increases
- their trust in what nonnative speakers say. Cognitive Science, 45(11), e13064.
- https://doi.org/10.1111/cogs.13064
- Bosker, H. R., Quené, H., Sanders, T., & De Jong, N. H. (2014). Native 'um's elicit
- prediction of low-frequency referents, but non-native 'um's do not. Journal of Memory

- and Language, 75, 104–116. https://doi.org/10.1016/j.jml.2014.05.004
- Bosker, H. R., Quené, H., Sanders, T., & Jong, N. H. de. (2014). The perception of fluency
- in native and nonnative speech. Lang. Learn., 64(3), 579–614.
- https://doi.org/10.1111/lang.12067
- Brennan, S. E., & Williams, M. (1995). The feeling of another's knowing: Prosody and filled
- pauses as cues to listeners about the metacognitive states of speakers. Journal of
- 430 Memory and Language, 34(3), 383-398.
- Butterworth, H. (2019). How much do they, um, know?: Disfluencies in speech as cues to
- speakers' knowledge. University of Edinburgh.
- <sup>433</sup> Caballero, J. A., & Pell, M. D. (2020). Implicit effects of speaker accents and
- vocally-expressed confidence on decisions to trust. Decision, 7(4), 314.
- https://doi.org/10.1037/dec0000140
- <sup>436</sup> Caffarra, S., Michell, E., & Martin, C. D. (2018). The impact of foreign accent on irony
- interpretation.  $PLoS \ One, \ 13(8), \ e0200939.$
- <sup>438</sup> Christenfeld, N. (1995). Does it hurt to say um? Journal of Nonverbal Behavior, 19, 171–186.
- 439 Davies, A. (2003). The native speaker: Myth and reality (Vol. 38). Multilingual matters.
- https://doi.org/10.21832/9781853596247
- De Jong, N. H., Groenhout, R., Schoonen, R., & Hulstijn, J. H. (2015). Second language
- fluency: Speaking style or proficiency? Correcting measures of second language fluency
- for first language behavior. Applied Psycholinguistics, 36(2), 223–243.
- DeBruine, L. M., & Barr, D. J. (2021). Understanding mixed-effects models through data
- simulation. Advances in Methods and Practices in Psychological Science, 4(1),
- 2515245920965119. https://doi.org/10.1177/251524592096511
- Derwing, T. M., Munro, M. J., Thomson, R. I., & Rossiter, M. J. (2009). The relationship
- between L1 fluency and L2 fluency development. Studies in Second Language Acquisition,
- *31* (4), 533–557.
- 450 Dragojevic, M., & Giles, H. (2016). I don't like you because you're hard to understand: The

- role of processing fluency in the language attitudes process. Human Communication
- 452 Research, 42(3), 396–420. https://doi.org/10.1111/hcre.12079
- Fairchild, S., Mathis, A., & Papafragou, A. (2020). Pragmatics and social meaning:
- Understanding under-informativeness in native and non-native speakers. Cognition, 200,
- 455 104171.
- Fairchild, S., & Papafragou, A. (2018). Sins of omission are more likely to be forgiven in
- non-native speakers. Cognition, 181, 80–92.
- 458 Feinberg, D. R., Jones, B. C., Little, A. C., Burt, D. M., & Perrett, D. I. (2005).
- Manipulations of fundamental and formant frequencies influence the attractiveness of
- human male voices. Animal Behaviour, 69(3), 561–568.
- Foucart, A., Costa, A., Morís-Fernández, L., & Hartsuiker, R. J. (2020). Foreignness or
- processing fluency? On understanding the negative bias toward foreign-accented speakers.
- Language Learning, 70(4), 974-1016. https://doi.org/10.1111/lang.12413
- Gkalitsiou, Z., & Werle, D. (2023). Speech disfluencies in bilingual greek-english young
- adults. Journal of Fluency Disorders, 78, 106001.
- 466 Gluszek, A., & Dovidio, J. F. (2010). The way they speak: A social psychological
- perspective on the stigma of nonnative accents in communication. Personality and Social
- Psychology Review, 14(2), 214–237.
- 469 Gluszek, A., Newheiser, A.-K., & Dovidio, J. F. (2011). Social psychological orientations and
- accent strength. Journal of Language and Social Psychology, 30(1), 28-45.
- Goupil, L., Ponsot, E., Richardson, D., Reyes, G., & Aucouturier, J.-J. (2021). Listeners'
- perceptions of the certainty and honesty of a speaker are associated with a common
- prosodic signature. Nature Communications, 12(1), 861.
- 474 Grosjean, F., & Deschamps, A. (1975). Analyse contrastive des variables temporelles de
- l'anglais et du français: Vitesse de parole et variables composantes, phénoménes
- d'hésitation. *Phonetica*, 31 (3-4), 144–184.
- Guyer, J. J., Fabrigar, L. R., & Vaughan-Johnston, T. I. (2019). Speech rate, intonation,

- and pitch: Investigating the bias and cue effects of vocal confidence on persuasion.
- Personality and Social Psychology Bulletin, 45(3), 389–405.
- Hanuíková, A., Alphen, P. M. van, Goch, M. M. van, & Weber, A. (2012). When one
- person's mistake is another's standard usage: The effect of foreign accent on syntactic
- processing. Journal of Cognitive Neuroscience, 24(4), 878–887.
- https://doi.org/10.1162/jocn\_a\_00103
- Heller, D., Arnold, J. E., Klein, N., & Tanenhaus, M. K. (2015). Inferring difficulty:
- Flexibility in the real-time processing of disfluency. Language and Speech, 58(2), 190–203.
- <sup>486</sup> Ip, M. H. K., & Papafragou, A. (2022). Integrating non-native speaker identity in semantic
- and pragmatic processing. Proceedings of the Annual Meeting of the Cognitive Science
- 488 Society, 44.
- Jiang, X., Gossack-Keenan, K., & Pell, M. D. (2019). To believe or not to believe? How voice
- and accent information in speech alter listener impressions of trust. Quarterly Journal of
- Experimental Psychology, 73(1), 55-79. https://doi.org/10.1177/1747021819865833
- <sup>492</sup> Jiang, X., & Pell, M. D. (2015). On how the brain decodes vocal cues about speaker
- confidence. Cortex, 66, 9–34.
- Jiang, X., & Pell, M. D. (2016a). Neural responses towards a speaker's feeling of (un)
- knowing. Neuropsychologia, 81, 79–93.
- Jiang, X., & Pell, M. D. (2016b). The feeling of another's knowing: How "mixed messages"
- in speech are reconciled. Journal of Experimental Psychology: Human Perception and
- 498 Performance, 42(9), 1412.
- <sup>499</sup> Jiang, X., & Pell, M. D. (2017). The sound of confidence and doubt. Speech Communication,
- 500 *88*, 106–126.
- Lange, K., Kühn, S., & Filevich, E. (2015). just another tool for online studies (JATOS):
- An easy solution for setup and management of web servers supporting online studies.
- 503 PloS One, 10(6), e0130834. https://doi.org/10.1371/journal.pone.0130834
- Leeuw, J. R. de, Gilbert, R. A., & Luchterhandt, B. (2023). jsPsych: Enabling an

- open-source collaborative ecosystem of behavioral experiments. Journal of Open Source
- 506 Software, 8(85), 5351. https://doi.org/10.21105/joss.05351.
- Lev-Ari, S. (2015). Comprehending non-native speakers: Theory and evidence for
- adjustment in manner of processing. Frontiers in Psychology, 5, 1546.
- Lev-Ari, S., & Keysar, B. (2010). Why don't we believe non-native speakers? The influence
- of accent on credibility. Journal of Experimental Social Psychology, 46(6), 1093–1096.
- https://doi.org/10.1016/j.jesp.2010.05.025
- Lev-Ari, S., & Keysar, B. (2012). Less-detailed representation of non-native language: Why
- non-native speakers' stories seem more vague. Discourse Processes, 49(7), 523–538.
- Lorenzoni, A., Pagliarini, E., Vespignani, F., & Navarrete, E. (2022). Pragmatic and
- knowledge range lenience towards foreigners. Acta Psychologica, 226, 103572.
- Mai, R., & Hoffmann, S. (2013). Accents in business communication: An integrative model
- and propositions for future research. Journal of Consumer Psychology, 24(1), 137–158.
- https://doi.org/10.1016/j.jcps.2013.09.004
- Matzinger, T., Pleyer, M., & Żywiczyński, P. (2023). Pause length and differences in
- cognitive state attribution in native and non-native speakers. Languages, 8(1), 26.
- https://doi.org/10.3390/languages8010026
- Mol, L., Kuhlen, A., Van der Steen, R., & Obbens, M. (2013). Beliefs about a speaker affect
- feeling of another's knowing. Proceedings of the Annual Meeting of the Cognitive Science
- 524 Society, 35.
- Norton-Ford, J. D., & Hogan, D. R. (1980). Role of nonverbal behaviors in social judgments
- of peers' assertiveness. Psychological Reports, 46(3 suppl), 1085–1086.
- Pinget, A. F., Bosker, H. R., Quené, H., & De Jong, N. H. (2014). Native speakers'
- perceptions of fluency and accent in L2 speech. Language Testing, 31(3), 349–365.
- https://doi.org/10.1177/026553221452
- 530 Pittam, J., & Gallois, C. (1986). Predicting impressions of speakers from voice quality:
- Acoustic and perceptual measures. Journal of Language and Social Psychology, 5(4),

- <sub>532</sub> 233–247.
- R Core Team. (2024). R: A language and environment for statistical computing. Vienna,
- Austria: R Foundation for Statistical Computing. Retrieved from
- https://www.R-project.org/
- Rakić, T., Steffens, M. C., & Mummendey, A. (2011). When it matters how you pronounce
- it: The influence of regional accents on job interview outcome. British Journal of
- Psychology, 102(4), 868–883.
- Shriberg, E. E., Bates, R., & Stolcke, A. (1997). A prosody only decision-tree model for
- disfluency detection. Fifth European Conference on Speech Communication and
- Technology.
- Shriberg, E. E., & Lickley, R. J. (1993). Intonation of clause-internal filled pauses.
- Phonetica, 50(3), 172–179.
- Smith, V. L., & Clark, H. H. (1993). On the course of answering questions. Journal of
- Memory and Language, 32(1), 25–38.
- Uleman, J. S., Adil Saribay, S., & Gonzalez, C. M. (2008). Spontaneous inferences, implicit
- impressions, and implicit theories. Annu. Rev. Psychol., 59, 329–360.
- Wickham, H. (2016). ggplot2: Elegant graphics for data analysis. Springer-Verlag New York.
- Retrieved from https://ggplot2.tidyverse.org
- Wickham, H., Averick, M., Bryan, J., Chang, W., McGowan, L. D., François, R., ... Yutani,
- H. (2019). Welcome to the tidyverse. Journal of Open Source Software, 4(43), 1686.
- https://doi.org/10.21105/joss.01686