Same as always: I suck at titles

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Abstract

Whether an individual is perceived as knowledgeable by others can be biased by several, and 10 potentially irrelevant factors, ranging from the ephemeral (e.g., how speech is produced) to 11 the situational (e.g., who is speaking). The presence of disfluencies, such as filled pauses 12 (e.g., uh or um in English), elicit evaluations of the speaker as deceitful, less intelligent, or 13 less confident in their knowledge. On the other hand, listeners consider alternative 14 explanations for the speaker to be underinformative: Specifically, non-native speakers are 15 more likely to be forgiven when they fail to be informative. However, whether and how 16 speaker's language profiency affects the interpretation of hesitations and listeners' 17 subsequent behaviour is unclear. In a horse-race betting paradigm, we show that listeners 18 are less likely to follow advice from a speaker who is disfluent, regardless of whether they are disfluent for reasons other than low knowledge. This suggests that previous reported perceptions of knowledgeability elicited explicitly have an impact on how individuals use the 21 information they are given, even in situations when it is not advantageous. Although there is 22 evidence for pragmatic leniance towards non-native speakers, it may be the case that this 23 only applies to failures in language skills that are believed to require high proficiency (e.g., irony), whilst hesitation phenomena may not be subject to such forgiveness. Overall, our 25 results align with a broader body of literature suggesting that the interpretation of 26 hesitation phenomena is not context-dependent. 27

Keywords: filled pauses, disfluencies, non-native-accented speech, feeling of another's knowing

Word count: 5173

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Impressions of other people arise naturally and automatically (Uleman, Adil Saribay, 32 & Gonzalez, 2008). Speakers can be evaluated not only by what they say but also by how 33 they say it. For example, vocal features of speech can affect perceptions of confidence and persuasion (Guyer, Fabrigar, & Vaughan-Johnston, 2019), status and solidarity (Pittam & 35 Gallois, 1986), or even attractiveness (Feinberg, Jones, Little, Burt, & Perrett, 2005). One speech feature consistently found to bias evaluations of the speaker is speech fluency. Speech 37 which includes disfluencies, such as uh or um in English, can lead to poorer evaluations of the speaker in terms of intelligence (Christenfeld, 1995), competence (Norton-Ford & Hogan, 1980), and certainty of their own knowledge (Brennan & Williams, 1995). Interestingly, speech produced with an accent different from one's own - particularly a foreign accent - can also elicit negative perceptions of the speaker, including worse evaluations in terms of status, solidarity, and credibility (Gluszek & Dovidio, 2010; Gluszek, Newheiser, & Dovidio, 2011; Lev-Ari & Keysar, 2010; Rakić, Steffens, & Mummendey, 2011). In this study, however, we explore a rather counterintuitive idea: That the co-presence of these two factors (i.e., (dis)fluency and foreign accents) may diminish the negative impact they have individually. It is important to note that we are not interested whether these cues correlate with a speaker's knowledge or whether listeners believe the speaker to be intentionally misleading: Rather, we are interested in whether, in scenarios where the speaker is expected to be cooperative and knowledgeable, listeners are still biased by irrelevant cues.

There is a growing body of literature showing that certain speech features impact listeners' perceptions of a speaker's knowledgeability. Voices with a slow speech rate, low amplitude, and a large F0 range are less likely to be rated as confident, and are associated with distinct neural responses in the listener (Jiang & Pell, 2015, 2016b, 2016a). Similarly, prosody guides listeners' evaluations of certainty and honesty cross-linguistically (Goupil, Ponsot, Richardson, Reyes, & Aucouturier, 2021). These non-verbal qualities of speech may

be accompanied by verbal markers of hesitation, such as disfluencies (Grosjean & Deschamps, 1975; Jiang & Pell, 2017; Shriberg, Bates, & Stolcke, 1997; Shriberg & Lickley, 1993). Brennan and Williams (1995) demonstrated that filled pauses (e.g., uh, mm, or um) impact 59 the perception of speakers' confidence in their knowledge. In their study, participants 60 listened to previously recorded answers to trivia questions (without hearing the questions) and were asked to rate how likely each speaker would be to recognise the correct answer to the question i.e., feeling of another's knowing, FOAK. Among the cues that biased participants' assessments were filled pauses: Answers containing a disfluency were more likely to receive lower FOAK ratings, suggesting that filled pauses were taken as reflective of speakers' reduced certainty about their knowledge. In contrast, non-answers (i.e., 'I don't know') were more likely to receive a higher FOAK rating if preceded by a filled pause. Brennan and Williams (1995) took this as evidence that listeners are sensitive to the surface form of delivery, and in particular, to the cues displayed by speakers when they do not know (or cannot remember at the moment of being asked) the answer to a question (see also Smith & Clark, 1993).

In a similar vein, speaker identity can itself affect evaluations of both the speaker and
the content of speech. In particular, a speaker's nativeness (whether they are producing
speech in their first or second language) has been consistently reported to elicit different
evaluations. On the one hand, speakers with a foregin accent are more likely to be negatively
evaluated (e.g., Dragojevic & Giles, 2016; Gluszek & Dovidio, 2010). Statements produced
with a foreign accent are more likely to receive lower ratings of credibility compared to
statements produced with a native accent (Boduch-Grabka & Lev-Ari, 2021; Lev-Ari &
Keysar, 2010; Barlow et al., 2024; Foucart, Costa, Morís-Fernández, & Hartsuiker, 2020;
Foucart & Hartsuiker, 2021; Souza & Markman, 2013; but cf., Stocker, 2017; Wetzel,
Zufferey, & Gygax, 2021). These negative effects triggered by a speaker's non-nativeness
have been accounted for in terms of processing fluency (Lev-Ari & Keysar, 2010), group
membership, or a combination of both (Mai & Hoffmann, 2013).

On the other hand, there seem to be instances where being a non-native speaker is 84 more advantageous. Fairchild, Mathis, and Papafragou (2020) examined whether and how 85 underinformativeness is interpreted differently as a function of speakers' nativeness. Fairchild 86 et al. (2020) isolated the processing fluency component by presenting written stimuli, so that 87 any differences in attributions were solely guided by speakers' identities. In a series of four experiments, participants read stories where a speaker (native/non-native) described a new 89 invention. Native speakers who were underinformative were more likely to be rated as unwilling to share the information. The same pragmatic failure by a non-native speaker, however, was taken as a sign of inability to produce the necessary information (Exp. 1), even when participants were not explicitly informed that the non-native speaker could experience language difficulties (Exp. 2). Further, this 'forgiveness' for underinformative statements had consequences for participants' subsequent behaviours: They were more likely to learn new information from a previously encountered underinformative non-native speaker than from a native speaker (Exp. 3 and 4) (for similar findings, see Fairchild & Papafragou, 2018; Lorenzoni, Pagliarini, Vespignani, & Navarrete, 2022). Interestingly, this pattern of results holds even for auditory stimuli (Ip & Papafragou, 2022), suggesting that difficulties in comprehending non-native-accented speech alone do not necessarily lead to negative 100 evaluations of the speaker. 101

One potential explanation is that non-native speakers are evaluated differently from 102 native speakers due to expectations about their competence. Non-native speakers' accents 103 may invoke stereotypes whereby these speakers are believed to be less competent than native 104 speakers (Fairchild & Papafragou, 2018; Lev-Ari, 2015). This expectation-based account 105 proposes that stereotypes of non-native speakers' low linguistic competence affect both how their speech is comprehended and interpreted. Specifically, signs that can be attributed to low linguistic competence would be interpreted as such when produced by a non-native 108 speaker and be attributed to other factors when produced by a native speaker. In line with 100 this hypothesis, a speaker's nativeness affects how ironic a statement is perceived to be 110

(Bazzi, Brouwer, & Foucart, 2022; Caffarra, Michell, & Martin, 2018) or why non-native speech is processed more shallowly (Hanuíková, Alphen, Goch, & Weber, 2012; Lev-Ari & Keysar, 2012).

Interestingly, stereotypes on linguistic competence may also affect how speech is 114 expected to be delivered. For example, a speaker's nativeness has been shown to affect how 115 confident a voice is perceived to be (Caballero & Pell, 2020; Jiang, Gossack-Keenan, & Pell, 116 2019). Given that speakers producing speech in their second language are more disfluent 117 (Bergmann, Sprenger, & Schmid, 2015; Davies, 2003; Gkalitsiou & Werle, 2023) and are 118 perceived as such by native listeners (Pinget, Bosker, Quené, & De Jong, 2014), it is possible that a speaker's nativeness leads to different interpretations of disfluencies. This would align with findings in the disfluency literature whereby the effects of disfluencies in speech 121 comprehension are dependent on who produces them (Arnold, Kam, & Tanenhaus, 2007; 122 Barr & Seyfeddinipur, 2010; Heller, Arnold, Klein, & Tanenhaus, 2015), including non-native 123 versus native speakers (Bosker, Quené, Sanders, & De Jong, 2014), allegedly because 124 listeners are sensitive to the speaker's mental state and the reasons for being disfluent. 125

Recently, Matzinger, Plever, and Żywiczyński (2023) explored whether listeners' 126 perceptions of why a speaker was disfluent differed for native and non-native speakers. In 127 their study, participants listened staged conversations where native and non-native speakers 128 answered trivial questions and requests, and were explicitly asked to rate each speaker's 129 knowledge and confidence (for FOAK) and their willingness to grant the request. Crucially, 130 Matzinger et al. (2023) manipulated speakers' fluency by manipulating inter-turn pauses: Answers were prefaced with either short (200 ms) or long (1200 ms) pauses. For requests, long pauses were less likely to be associated with unwillingness for non-native compared to native speakers. However, FOAK ratings did not differ between speakers: Long pauses 134 produced by either speaker were likely to be taken as reflecting low confidence and low 135 knowledge. Matzinger et al. (2023) attribute this pattern to different conversational 136

contexts: Requests tap into speakers' cooperativeness, and thus in this context, tuning to the interlocutor's mental state and stereotyping might be more relevant than evaluating the speaker's competence (i.e., knowledge).

As is the case for Matzinger et al. (2023), most studies evaluating the effects of 140 manner of speech on pereptions of the speaker have used explicit ratings. Participants are asked to rate particular traits of speakers or what they say on a scale. Further, these ratings 142 are elicited in non-social contexts, where participants do not have anything at stake. Although these experiments demonstrate that listeners show sensitivity to aspects of speech such as fluency and accent when asked to make explicit judgements about the speaker, they may not explain how listeners evaluate speech implicitly, when to do so is consequential. Here, we propose an implicit measurement of listeners' assessments of the speaker's certainty, using 147 a horse-race paradigm. In this task, participants listen to a set of speakers provide 148 descriptions of horses, and are asked to distribute virtual tokens as 'bets' on each horse's 149 likelihood of winning a putative race. This approach presents two advantages over previous 150 experiments. First, in the horse-race paradigm, participants are not explicitly asked to 151 evaluate a certain trait of the speaker (in this case, how knowledgeable they are). Instead, 152 we take participants' allocations of 'betting money' as an indirect measurement of their 153 perceptions of speakers' knowledge. Indeed, pilot studies have shown that individuals are 154 sensitive to this manipulation and that disfluent information leads to smaller bets 155 (Butterworth, 2019). Second, horse races provide a scenario where individuals can make 156 decisions based on what they are told, but the content of speech itself may not be 157 informative for many individuals (in that participants are less familiar with the world of 158 horse racing or its technical vocabulary). 159

In a pre-registered study, we set up to explore whether and how perceptions of
certainty are biased by manner of delivery in the form of fluency, and by the speaker's
identity as conveyed by their accent. We presented participants with recordings of a native

and a non-native speaker, each describing two horses, with one description produced fluently 163 and one description produced disfluently. If listeners are sensitive to both local and global 164 causes of hesitations when making judgements about certainty, then disfluent descriptions 165 provided by a native speaker should result in less money bet, reflecting listeners' lower trust 166 in, or lower FOAK, for the speaker. However, disfluent descriptions provided by a non-native 167 speaker may not impact listeners' betting behaviours, to the extent that they consider the 168 possibility of difficulties in production when assessing the speaker's knowledge. To further 169 control for the potential effects of (non)-nativeness on certainty on its own, we measured 170 participants' language attitudes towards each speaker (see Dragojevic & Giles, 2016), 171 perceived fluency, accentedness, and comprehensibility of the native and the non-native 172 speaker, as well as the perceived trustworthiness of each speaker. We additionally measured 173 participants' familiarity with and exposure to native and non-native-accented English on a daily basis, to account for the fact that exposure to non-native accents can reduce their 175 negative effects on listeners' judgements (Boduch-Grabka & Lev-Ari, 2021).

177 Methods

All experimental stimuli can be found at https://osf.io/zsut7/.

179 Participants

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We conducted a sensitivity power analysis via data simulation following DeBruine 180 and Barr (2021). We explored the required number of participants for a study with a power 181 higher than .8 for the interaction between fluency and speaker identity. We conducted a 1000 182 simulations per different combinations of the effect size of the interaction (ranging from 183 small to medium) and the standard deviation of the residuals. In this analysis, we assumed a 184 medium effect size of fluency and no effect of speaker identity. This analysis showed that a 185 sample size of 360 participants ensured enough power to detect a medium or greater effect 186 size. 187

In our pre-registration, we stipulated that only participants born and raised, and

currently residing in the United Kingdom, with English as their first and only language, and 189 with no auditory disorders could partake in the study. Further, we would exclude from 190 analysis data from participants who reported the experiment's aim or manipulation, rated 191 the naturalness of the auditory stimuli (defined as how likely they believed the audios to 192 have been recorded in one go) lower than four, or considered themselves experts in horse 193 races. This meant that we recruited 641 participants for a sample of 360. Participants were 194 aged between 18 and 30, born and raised in the United Kingdom, and reported that they did 195 not consider themselves as experts in horse races. Participants were recruited via the online 196 platform Prolific, and gave their informed consent as approved by the PPLS Ethics 197 Committee (ref no 133-2223/3). Participants were reimbursed £1.50 for a 10-minutes 198 experiment. 199

Visual stimuli 200

We selected a candidate set of eight images of racehorses from the web. The selected 201 images each featured only one racehorse in the foreground, in motion, ridden by a jockey, 202 and the horses all took up approximately the same proportion of the image. To ensure that 203 the pictures did not bias participants' bets, we recruited ten participants on Prolific, who did 204 not take part in the main study, and asked them to rate on a 10-point scale how likely each 205 horse was to win a hypothetical race and to rank the horses in the order they thought they 206 would cross the finish line, in exchange for £0.45. 207

A one-way repeated measures ANOVA showed that there were no differences in how 208 likely each horse was thought to be to win a race (F(7) = 0.5, p = .83). An ordinal logistic 209 regression showed that none of the horses were more likely to be ranked differently from the 210 others (all |t| < 2). We therefore selected four out of the eight images as visual stimuli. 211

Auditory stimuli

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We used four descriptions of racing horses from *The Racing Post*, retrieved on 213 October 2018, originally edited and used by Butterworth (2019). Each passage consisted of

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three to four sentences describing a horse and its performance in previous races. 30 British English speakers, who did not participate in the final experiment, rated these passages to ensure that all descriptions were perceived as equally likely to describe a winning horse.

Participants rated on a 10-point scale how likely they thought each horse was to win a race individually, and then ranked all the horses, in exchange for £0.50.

Based on these results, we further edited the descriptions as some descriptions were more likely to be rated as winning horses. A new sample fo 30 British English speaker rated the edited passages. A one-way repeated measures ANOVA showed no differences in each horse's rated likelihood of winning a race (F(3) = 0.76, p = .52), nor in the order in which they were ranked (all |t| < 2). The final set of descriptions for the experiment can be found in Table 1. Each description was paired with one of the four visual stimuli.

We recorded a British English native speaker and a non-native (L1: Italian) English 226 speaker to create the auditory stimuli. Both speakers were female. Passages were recorded 227 one at a time. To elicit naturally disfluent recordings, both speakers were instructed to read 228 the passages silently and then were recorded as they tried to recall the passage from memory. 229 To avoid differences between the descriptions provided by the speakers, they were allowed to 230 look at the descriptions as they spoke if they could not remember the continuation. We 231 edited these recordings using Audacity to ensure that the recordings of each speaker had 232 similar numbers of disfluencies in similar locations, by cross-splicing different recordings (see 233 Table 2). To create the fluent counterpart of each description, filled and mid-utterance silent pauses were excised, and elongations and between-clause silent pauses were reduced using 235 the 'Tempo' function. The final auditory experimental stimuli consisted, for each speaker, of 236 two descriptions of each horse (one disfluent, one fluent), resulting in sixteen recordings. 237 Description, fluency, speaker, and order of presentation were counterbalanced in a Latin 238 Square design, resulting in 24 lists. 239

To ensure that the resulting descriptions were perceived to be natural (i.e., edits to

Table 1

Original description of each horse that speakers were asked to memorise and reproduce.

Horse	Description				
	Fire Walker is looking strong thanks to his come-from-behind success				
	in the Acomb Stakes. The impression given in both runs is that Fire				
Fire	Walker should handle the demands of the extra furlong and Charlie Hills				
Walker	is looking forward to the test. The trainer said "He's done really well for				
	a little break, his work's been good and I couldn't be more pleased with				
	him".				
Silver Sky	Silver Sky, a runner-up of a seven-furlong maiden at Naas on his debut				
	last month, the son of Invincible Spirit ran crack French colt Persian King				
	to a neck in the Group 3 Autumn Stakes over today's trip at Newmarket				
	two weeks ago and his trainer believes he has done well since. O'Brien				
	said "Silver Sky is a fine big colt and a talented one".				
	Apocalypse has put in a string of consistent performances, most recently				
	finishing third to Norway in the Zetland. "He's had a very solid year"				
Apocalypse	said trainer Archie Watson. "He ran a good race in the Zetland, beaten				
	only a length and a quarter, and I think the field here is of a similar level				
	so I'm more than happy for him to take his chance".				
D. I	Black Blade proved the market all wrong as the complexion of the 6.5-				
	furlong novice race changed dramatically in the final two furlongs, with				
Black	the Rebel Racing premier-owned newcomer under Tom Queally collaring				
Blade	long-time leaded Monsieur Noir. Spencer said: "He did it well. He's a				
	nice horse. We always thought had a bright future".				

 ${\bf Table~2} \\ {\it Transcriptions~of~the~disfluent~horse~descriptions}. \\$

Horse	Description
Fire	FP-um Fire Walker is looking SP strong thanks to his (elongation) SP
	come-from-behind success SP FP-er in the Acomb Stakes. FP-um The
	(elongation) impression given SP in both runs FP-uh is that Fire Walker
	should SP handle the demands of SP the extra furlong and FP-uh Charlie
Walker	Hills is looking forward to the test. The trainer said "He's done (elonga-
	tion) FP-um really well for a little break, FP-uh his work's been good
	and I couldn't be more pleased with him".
	Silver Sky, FP-uh a runner-up of -a seven-furlong maiden at Naas on his
	(elongation) debut last month. FP-uh The son of Invincible Spirit SP
	ran (elongation) crack French colt Persian King to a neck in the FP-uh
Silver Sky	Group 3 Autumn Stakes. over today's trip at Newmarket two weeks ago
	FP-uh and his trainer believes he has done well since. O'Brien said FP-uh
	"Silver Sky is a fine big colt and a talented one".
	FP Apocalypse has put in a (elongation-ish) string SP of consistent
	performances, FP most recently finishing third to SP Norway in the
	Zetland. FP "He's had a very solid year" said trainer SP FP Archie
Apocalypse	Watson. "He ran a SP good race in the Zetland, SP FP SP beaten only a
	length and a quarter, and I think the field here SP is of a similar level so
	I'm- I'm more than happy for him to (elongation-ish) take his chance".
	FP-um Black SP Blade proved SP FP-uh the market all wrong as the
	complexion SP of the 6.5 furlong SP novice SP race FP-um changed
	dramatically in SP the final two furlong. SP FP-uh with the Rebel
Black	Racing SP premier-owned SP newcomer under Tom SP Queally FP-um
Blade	SP collaring long-time leaded Monsieur Noir. FP-uh Spencer said: "He
	did it well. FP-uh He's a nice horse. FP-uh We always thought he had a
	bright future".

accent respectively.

Table 3

Mean (standard deviation) ratings of native and non-native speakers' fluent and disfluent recordings for fluency, naturalness, and accentedness, on a 9-point scale where lower values indicate less fluent, less natural, and less

Speaker	Delivery	Fluency	Naturalness	Accentedness
Native Speaker	Fluent	7.62 (1.59)	5.44 (2.62)	4.38 (2.22)
	Disfluent	6.44 (2.36)	5.73 (2.57)	4 (1.97)
Non-native Speaker	Fluent	6.25 (1.73)	5.94 (1.83)	7.02 (1.59)
	Disfluent	5.21 (1.75)	5.83 (1.99)	7.17 (1.51)

the 'fluent' audios were not obvious) and that they were perceived as differing in fluency (i.e., 241 disfluent and 'fluent' versions were distinguishable), we validated them in a sample of 48 242 British English participants, who did not take further part in the study. Participants were 243 allocated to one of the 24 experimental lists, to ensure that our validation procedure was 244 similar to how participants encountered stimuli in the actual experiment. Following Bosker, 245 Quené, Sanders, and De Jong (2014)'s procedure, participants were asked to rate each 246 audio's fluency on a scale from 1 to 9 (1: not fluent at all, 9: very fluent). We instructed 247 participants to rate fluency by considering silent and filled pauses, speed of speech, and 248 repairs, and to ignore speakers' accents and the content of their speech. Participants 249 additionally rated on a 9-point scale each recording's naturalness (defined as how likely it 250 was that the audio had been recorded in one go; 1: not unlikely at all; 9: very likely), and accentedness (while ignoring the perceived speaker's proficiency in the language; 1: not 252 accented at all, 9: very accented). We additionally asked participants to guess the speakers' 253 country of origin. At the end of the task, participants were further asked how often they interacted with native and non-native English speakers (on a 9-point scale, 1: never, 9: 255 always) and were allowed to report if they noticed anything odd in the auditory stimuli.

Table 3 shows the means (and standard deviations) of participants' ratings of fluency, 257 naturalness, and accentedness. A linear mixed model for fluency ratings with fixed effects of 258 fluency (treatment-coded, reference: fluent), speaker's linguistic background 259 (treatment-coded, reference: native speaker), and their interaction, with random intercepts 260 by participant and by horse description, showed that fluency ratings differed significantly for 261 the fluent and disfluent conditions ($\hat{\beta} = -1.19, 95\%$ CI [-1.79, -0.58], t = -3.84). The 262 non-native speaker was perceived as more disfluent than the native speaker ($\hat{\beta} = -1.37, 95\%$ 263 CI [-1.97, -0.78], t = -4.51), in line with previous findings (Bosker, Quené, Sanders, & 264 Jong, 2014; Pinget et al., 2014), but the interaction between the two variables was not 265 significant ($\hat{\beta} = 0.15, 95\%$ CI [-0.49, 0.78], t = 0.45). An identical model for naturalness 266 ratings showed no significant differences by fluency ($\hat{\beta} = 0.29, 95\%$ CI [-0.52, 1.11], 267 t=0.70), speaker's linguistic background ($\hat{\beta}=0.50, 95\%$ CI [-0.27, 1.27], t=1.27) or their interaction ($\hat{\beta} = -0.40, 95\%$ CI [-1.49, 0.70], t = -0.71).

270 Procedure

Stimuli were presented using JsPsych (Leeuw, Gilbert, & Luchterhandt, 2023), hosted 271 on MindProbe (via JATOS, Lange, Kühn, & Filevich, 2015). The task began with a cover 272 story introducing two horse racing tipsters. Participants were told the tipsters would provide 273 information about the four most popular horses competing in an upcoming race. The cover 274 story explained that the two tipsters were well-known experts in the field, and added that 275 one of the speakers was a non-native English speaker (without specifying the nationality of 276 either speaker), introducing the element of the speakers' linguistic backgrounds as well as the 277 factor of competence. 278

At the beginning of the experiment, participants were shown four pictures of the
horses that they had been told would take part in the race, alongside their names.

Participants were instructed to distribute one-hundred pounds in betting money across the
four horses based on the likelihood they thought each horse had of winning: They could split

the bets as they wished, and they did not have to spend all the money. Each participant was randomly assigned to one out of 24 groups, so that they would listen to each speaker twice, 284 one in each fluency level. The order in which horses were presented was randomised. In each 285 trial, participants listened as one speaker described a given horse's performance. Once the 286 playback stopped, participants were asked to place a bet by typing a number on a web form. 287 Participants could only move to the next horse's description once they had placed a bet. 288 Participants were allowed to modify their previous bets every time they heard a new 280 description. If the sum of bets made at any point was more than the allotted maximum, 290 they were asked to re-distribute their bets until the total was below one-hundred pounds. 291

After the betting round was complete, participants completed a questionnaire similar 292 to that of Foucart et al. (2020) to measure their language attitudes towards the native and 293 the non-native speaker. For each speaker, participants answered six questions measuring 294 affect (three questions for negative affect and three questions for positive affect), five 295 questions measuring solidarity, five questions measuring status, and one question each for 296 comprehensibility, accentedness, fluency, and trustworthiness. Each question used a 9-point 297 scale. Participants first answered questions about, at random, the native or non-native 298 speaker, with the order of presentation of dimensions being randomised. They then answered 290 the same questions for the remaining speaker. We also asked participants to guess the 300 countries of origin of our native and non-native speakers, and from which speaker they would 301 like to learn about horse races in the future. Additional questions included ratings on a 302 9-point scale of how natural the audio sounded (1: unnatural, edited; 9: natural, unedited), 303 and two questions measuring participants' exposure to native and non-native accented English (1: never; 9: always). Likewise, we measured participants' previous experience with betting and their perceived knowledge of horse races (two questions: Whether they had bet on horse races in the past, and to rate on a 5-point scale, from 'Strongly disagree' to 307 'Strongly agree', how closely they identified with the statement 'I am an expert on horse 308 races'). Finally, we included an open-ended question for participants to report what guided 309

Table 4

Mean (standard deviation) of money bet
by manner of delivery and speaker's
linguistic background.

Speaker	Delivery	Money bet	
Native	Fluent	24.24 (13.92)	
	Disfluent	20.66 (12.23)	
Non-native	Fluent	24.15 (13.61)	
	Disfluent	22.31 (13.81)	

their decision-making, as well as their perception of the experiment's aim.

311 Results

All data pre-processing and analyses were carried out in R version 4.4.1 (R Core
Team, 2024), using the packages *tidyverse* version 2.0.0 (Wickham et al., 2019), *ggplot*version 3.5.1 (Wickham, 2016), and *wesanderson* version 0.3.7 for data wrangling and
visualization. *lme4* version 1.1-35.5 (Bates, Mächler, Bolker, & Walker, 2015) was used for
data analysis, and *papaja* version 0.1.2 (Aust & Barth, 2023) for manuscript write-up.
Scripts can be found at https://osf.io/zsut7/

318 Pre-registered analyses

319 Betting behaviour

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Table 4 shows the mean amount of money bet per speaker and fluency condition. On average, participants bet £25 on when information was provided fluently. This distribution suggests that participants followed a rational behaviour: Given their lack of expertise, they distributed the one hundred pounds equally. However, disfluent instructions seem to lower the amount of money bet.

We modelled participants' betting behaviour in a linear mixed model. We modelled

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money bet on each horse, taken from the final amounts submitted in the experiment, after 326 all four descriptions had been heard and valid responses (summing to at most 100 pounds) 327 had been recorded. The model included fixed effects of fluency (sum-coded; fluent coded as 328 -0.5, disfluent as +0.5), speaker's linguistic background (sum-coded, native coded as -0.5; 329 non-native coded as +0.5), and their interaction. The maximal model (Barr, Levy, 330 Scheepers, & Tily, 2013), with random intercepts by-participant and by-item, with random 331 slopes for fluency and speaker's linguistic background by-participant, and for fluency by-item, 332 failed to converge. We first dropped the random intercept by-participant, as most 333 participants used all the money and thus there was no variance in their intercept. The final 334 model included a random intercept by-item. 335

Our model showed a main effect of fluency, whereby participants placed lower bets following disfluent descriptions compared to their fluent counterparts ($\hat{\beta} = -2.71$, 95% CI [-4.04, -1.38], t(1433) = -3.99, p < .001). There was no main effect of speaker ($\hat{\beta} = 0.78$, 95% CI [-0.55, 2.11], t(1433) = 1.15, p = .251) and importantly, no interaction between manner of delivery and speaker ($\hat{\beta} = 1.74$, 95% CI [-0.92, 4.40], t(1433) = 1.28, p = .200)¹.

Language attitudes and post-experimental questionnaire

Table 5 depicts the means (and standard deviation) of ratings in each of the measures of interest by speaker's linguistic background. Constructs measured with more than one question (affect, status, solidarity) were obtained by calculating the average score. For the Affect dimension, we reverse-scored items measuring negative affect. Cronbach's alpha showed that the scores for these attitudes were reliable ($\alpha_{affect} = 0.80$, $\alpha_{status} = 0.93$, $\alpha_{solidarity} = 0.87$). We explored differences between evaluations of the native and the non-native speaker in these three social dimensions as well as on Comprehensibility,

¹ An identical model including excluded participants showed a main effect of fluency ($\hat{\beta} = -2.92, 95\%$ CI [-3.91, -1.92], t(2557.08) = -5.75, p < .001), and a main effect for speaker ($\hat{\beta} = 1.21, 95\%$ CI [0.21, 2.20], t(2557.09) = 2.38, p = .017), but no significant interaction ($\hat{\beta} = 0.93, 95\%$ CI [-1.06, 2.92], t(2557.12) = 0.92, p = .358)

Table 5

Average score (and standard deviation) in each dimension by speaker.

Dimension	Native Speaker	Non-native Speaker	
Comprehensibility	7.62 (1.57)	5.51 (2.14)	
Accentedness	5.46 (2.1)	6.84 (1.37)	
Fluency	7.97 (1.43)	6.11 (1.83)	
Affect	5.77 (1.49)	6.05 (1.32)	
Status	5.95 (1.52)	6.43 (1.31)	
Solidarity	5.67 (1.48)	6.41 (1.24)	
Trustworthy	6 (1.61)	6.36 (1.41)	

Accentedness, Fluency, and Trustworthiness via paired t-test using Bonferroni correction for p values.

Analyses showed that speakers were rated differently across all six dimensions (see Table 6). The largest differences were, unsurprisingly, in comprehensibility and accentedness, where the non-native speaker received lower ratings.

Following our pre-registered analysis, we included these six variables in our previous model to explore whether speakers' evaluations could further explain participants' betting behaviours. This second model improved model fit ($\chi^2(7) = 24.77$, p < .01). The model showed a main effect of manner of delivery ($\hat{\beta} = -2.71$, 95% CI [-4.03, -1.39], t(1436.00) = -4.03, p < .001). Although affect ($\hat{\beta} = 0.83$, 95% CI [0.15, 1.50], t(1436.09) = 2.41, p = .016) and fluency ($\hat{\beta} = -0.67$, 95% CI [-1.24, -0.10], t(1436.10) = -2.31, p = .021) were significant, they did not survive Bonferroni correction.

Table 6					
Paired t test	for	each	dimension	between	speakers.

Dimension	t(359)	95% CI	d
Comprehensibility	15.99	[1.85, 2.36]	2.11
Accent	-11.51	[-1.62, -1.15]	-1.39
Fluency	17.15	[1.65, 2.08]	1.87
Affect	-3.54	[-0.44, -0.13]	-0.29
Status	-6.55	[-0.62, -0.33]	-0.48
Solidarity	-9.63	[-0.89, -0.59]	-0.74
Trustworthy	-4.12	[-0.53, -0.19]	-0.36

Exploratory analysis

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We additionally explored participants' preferences to learn from either speaker in the future. 206 participants reported they would prefer to learn from the native spaker, and 154 preferred the non-native speaker. A χ^2 test of goodness of fit showed that this difference in preferences was reliable ($\chi^2(1) = 7.51$, p < .01).

Following previous research suggesting the role of exposure to foreign-accented speech in its role to trust and comprehensibility, we explored whether exposure to native and non-native speakers played a role in participants' betting behaviour beyond that of speaker identity and manner of delivery. A similar linear mixed model to that employed in our main analysis, but with the addition of exposure, showed that it did not affect participants' betting behaviour ()

372 Discussion

Listeners can attribute knowledgeability to a speaker based on a range of cues, such as voice pitch or amplitude. Speech (dis)fluency has been previously shown to impact how confident in their knowledge a speaker is judged to be (Brennan & Williams, 1995). However,

speakers can be disfluent for reasons other than lack of confidence: For example, speaking in 376 one's second language is also associated with an increase in disfluencies (De Jong, 377 Groenhout, Schoonen, & Hulstijn, 2015; Derwing, Munro, Thomson, & Rossiter, 2009). In 378 fact, it has been proposed that second language speakers are stereotyped and thus expected 379 to display low linguistic performance (Lev-Ari, 2015), which consequently leads native 380 listeners to "forgive" what otherwise would be lead to a negative evaluation (Fairchild et al., 381 2020; Lorenzoni et al., 2022). In the present experiment, we explored whether attributions of 382 lack of knowledgeability as a function of the speaker's nativeness, and if those affected 383 listeners' subsequent behaviours in a task where participants had to place bets on horse. Our 384 findings suggest that manner of delivery, in the form of fluency, was the sole factor that 385 guided participants' behavior: Disfluent descriptions yielded lower bets, regardless of who 386 was speaking.

The pattern found aligns with the idea that listeners are sensitive to how an 388 utterance is produced. Descriptions that included hesitation phenomena, in the form of filled 389 pauses, led to smaller bets compared to their fluent counterparts. This aligns with Brennan 390 and Williams (1995) findings wherein listeners were less likely to attribute confidence in their 391 knowledge to a speaker when their answers included hesitation phenomena (i.e., Feeling of 392 Another's Knowing). Brennan and Williams (1995) attributed participants' ratings to 393 inferences made about the speaker's mental state - specifically, inferences about the degree of 394 confidence a speaker has in their knowledge. In our experiment, since participants have little 395 a priori knowledge that can guide their betting behaviour, and because utterances were 396 semantically identical and acoustically identical apart from the excision of disfluencies and 397 accent, an inference about a speaker's confidence based on those disfluencies is the most 398 likely potential determinant of participants' decisions. 399

The fact that disfluencies could also be accounted for by speaker's nativeness did not impact participants' decisions. This pattern differs from what we predicted given previous

research suggesting that stereotypes about speaker's linguistic abilities have consequences for 402 whether and how disfluencies affect speech comprehension (Arnold et al., 2007; e.g., Bosker, 403 Quené, Sanders, & De Jong, 2014; Heller et al., 2015), and the fact that errors produced by 404 non-native speakers are less likely to trigger otherwise negative interpretations (e.g., 405 Fairchild et al., 2020; Fairchild & Papafragou, 2018; Ip & Papafragou, 2022; Lorenzoni et al., 406 2022) possibly because native listeners expect them to produce those (Lev-Ari, 2015). This is 407 particularly remarkable given that in our post-experimental questionnaire, the native speaker 408 was more likely to be chosen as someone participants would like to learn from about horse 400 races in the future. 410

Nonetheless, our results align with recent studies that have failed to find this an 411 interaction between disfluencies and nativeness when it comes to speaker's attributions 412 (Matzinger et al., 2023). Matzinger et al. (2023) accounted for their findings as a the results 413 of the degree with which a listener needs to model the speaker: Assessing a speaker's 414 knowledge is not as socially engaging as other communicative contexts (e.g., granting a 415 request) and consequently, inferring an interlocutor's cognitive state is not relevant. 416 Interestingly, there are several differences between their paradigm and ours that shed doubts 417 on whether this explanation applies to our results. Specifically, in Matzinger et al. (2023) 418 participants were overhearing a conversation, and given their task (i.e., explicitly rate their 419 perception on speaker's knowledge or confidence), accurately ascribing the cause of 420 disfluency did not have any consequences. In contrast, our participants listened to speakers 421 talking in isolation and, most importantly, assessing speakers' knowledgeability was crucial 422 for performing their task. 423

One possibility for this lack of interaction has to do with the fact that our non-native speaker was introduced as a knowledgeable tipster. This introduction of the speaker as an authoritative figure may have overridden any other features of their identity, including their non-native speaker identity. Indeed, beliefs about a speaker's expertise have been shown to guide perceptions of their certainty (Mol, Kuhlen, Van der Steen, & Obbens, 2013). The
preference to learn in future occasions for the native speaker may be more likely to reflect the
ease of comprehending them, rather than an implicit negative bias towards the non-native
speaker, and particularly, a diminished perception of competence for non-native speakers.

This experiment introduced a novel approach to measuring how certain the speaker is 432 perceived to be. While previous studies had participants explicitly rate a speaker on different 433 dimensions (e.g., knowledgeability, trust) in non-social contexts, the present study offers a 434 new set-up to explore how different factors bias individuals' evaluations indirectly, as well as 435 the implications of those evaluations. However, a potential shortcoming of our design has to 436 do with how people approached the betting system. Because of the nature of the task, in 437 that we tested participants with no prior knowledge of horse races, our participants behaved 438 rationally: The vast majority employed the hundred pounds allocated and distributed them 439 fairly rationally in a way that maximizes their chances of winning (i.e., around £25 per 440 horse). It is, therefore, possible that in situations of less uncertainty (e.g., a scenario where 441 individuals have more knowledge) or where they are rewarded for allocating money to the 442 winning horse, how the speaker is perceived may yield a larger effect on participants' 443 behaviours. 444

Overall, our study shows that disfluency is a cue that listeners take as an index of speaker's knowledgeability. Importantly, in situations of uncertainty where individuals lack information to optimise their behaviour, the presence of disfluency biases their otherwise rational behaviour. This highlights individuals' tuning to the different ways in which a speaker's mental state is signalled in speech and how in turn it can affect their subsequent behaviours.

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