Homework 4

```
In [1]: import cvxpy as cp
import numpy as np
from scipy.linalg import expm, norm
from scipy.stats import norm as norm_dist
import matplotlib.pyplot as plt
```

Ouestion 1

(a) A One-Parameter Stressing Method for an Equicorrelation Matrix

An equicorrelation matrix for n assets has the form $C = (1 - \rho)I_n + \rho J_{n'}$ where I_n is the identity, J_n is an $n \times n$ matrix of ones, and ρ is the common off-diagonal correlation. A key property is that to be a valid correlation matrix, one must have

$$\rho \in \left[-\frac{1}{n-1}, 1\right]$$

A simple "stress" method is to introduce a single parameter $\theta \in [-1,1]$ and map it linearly into the admissible range for ρ . For example, define

$$\rho(\theta) = \left(\frac{1+\theta}{2}\right) \left(1 + \frac{1}{N-1}\right) - \frac{1}{N-1}.$$

In this mapping:

- When $\theta = -1$, $\rho(-1) = -1/(N-1)$ (the minimum allowable value).
- When $\theta = 1$, $\rho(1) = 1$.

The stressed correlation matrix is then given by

$$C(\theta) = (1 - \rho(\theta))I_n + \rho(\theta)J_n$$

This method is appealing because it depends on just a single parameter, and the resulting matrix is still a valid correlation matrix by construction. Moreover, the same idea extends directly to Gaussian or t copula models since you would simply plug the stressed $C(\theta)$ into the copula's dependence structure while keeping marginal distributions unchanged.

Using This Stress to Estimate Derivative Sensitivity:

Once you have such a stressed matrix:

- 1. Base Valuation: Start by pricing your derivative instrument using the "base" correlation (say, corresponding to $\theta = 0$).
- 2. Stress Testing: Reprice the derivative for a range of θ values (e.g., θ shifted slightly up and down from 0). In a copula model (Gaussian or t), the stressed $C(\theta)$ is used to generate joint scenarios.
- 3. Sensitivity Estimation: Compute the sensitivity by approximating the derivative of the instrument's price P with respect to θ (or indirectly with respect to ρ). For example, a finite difference estimate could be:

$$\frac{\partial P}{\partial \theta} \approx \frac{P(\theta + \Delta \theta) - P(\theta - \Delta \theta)}{2\Delta \theta}.$$

This procedure tells you how the derivative's price responds to changes in correlation.

(b) The Archakov-Hansen Parametrization Algorithm

```
In [2]: # Archakov-Hansen correlation matrix construction
def archakov_hansen_corr(A_off, tol=1e-6, max_iter=1000):
    """
    Given an n x n symmetric matrix A_off containing the off-diagonal elements,
    this function finds a vector x (diagonals) such that if A = A_off + diag(x),
    then C = expm(A) is a correlation matrix (i.e. diag(C) = 1).

Parameters:
    A_off: np.ndarray (n x n)
    A symmetric matrix with zeros on the diagonal (or preset off-diagonals).
    tol: float
        Convergence tolerance.
    max_iter: int
        Maximum number of iterations.

Returns:
```

```
C : np.ndarray (n x n)
        The correlation matrix.
n = A_off.shape[0]
x = np.zeros(n) # initial guess for diagonal adjustments
for it in range(max_iter):
    A = A_off + np.diag(x)
    expA = expm(A)
    diag_expA = np.diag(expA)
    # Update x to enforce diag(expA)==1: x_new = x - log(diag(expA))
    x \text{ new} = x - np.log(diag expA)
    if np.linalg.norm(x_new - x) < tol:</pre>
        x = x_new
        break
    x = x_new
# Compute final correlation matrix
A_{final} = A_{off} + np.diag(x)
C = expm(A_final)
```

Estimating Sensitivities of a Derivative to (d)

Assuming you have an estimate for the off-diagonal vector d (or you treat it as your "base" parameterization for the log-correlation), you can assess the sensitivity of a derivative instrument as follows:

- 1. Baseline Price: Compute the derivative price using the base correlation matrix *C* (obtained from your current *d* via the Archakov–Hansen algorithm).
- 2. Perturbation: Slightly perturb one or several elements of d (or a direction in the space of d). For example, for each component d_i compute $d_i^{\epsilon} = d_i + \epsilon$ while keeping other elements constant.
- 3. Recompute *C* and Price: For each perturbed *d* recompute the correlation matrix using the algorithm, then reprice the derivative instrument.
- 4. Finite Differences: Estimate the sensitivity (or "Greek") with respect to each component of d by a finite difference approximation:

$$\frac{\partial P}{\partial d_i} \approx \frac{P(d_i + \epsilon) - P(d_i)}{\epsilon}.$$

This procedure quantifies how small changes in the "low" off-diagonal entries (and thus in the implied log-correlation) affect the derivative's price.

Question 2

First let's restate subadditivity and positive homogeneity.

A risk measure ρ is subadditive if for all positions (or losses) L_1, L_2 in the domain M we have $\rho(L_1 + L_2) \le \rho(L_1) + \rho(L_2)$.

A risk measure ρ is positively homogeneous if for every position L in M and every scalar $\lambda \geq 0$ we have $\rho(\lambda L) = \lambda \rho(L)$.

Monotonicity axiom is stated as follows. If $L_1, L_2 \in M$ satisfy $L_1 \ge L_2$ then $\rho(L_1) \ge \rho(L_2)$.

The assumption we can make is that $L \le 0$ we have $\rho(L) \le 0$. We want to show that if L_1 and L_2 are two positions with $L_1 \ge L_{2'}$ then it must be that $\rho(L_1) \ge \rho(L_2)$.

Define the difference $M:=L_1-L_2$. Since $_1\geq L_2$, it follows that $M\geq 0$. Equivalently, $-M\leq 0$. By our assumption, we then have $\rho(-M)\leq 0$.

Now, express L_2 as $L_2 = L_1 - M$. Using the subadditivity property of ρ , we get $\rho(L_2) = \rho(L_1 - M) \le \rho(L_1) + \rho(-M)$.

Since $\rho(-M) \leq 0$, it follows that $\rho(L_2) \leq \rho(L_1)$.

This is exactly the monotonicity property (a larger risk measure corresponds to a riskier position).

Question 3

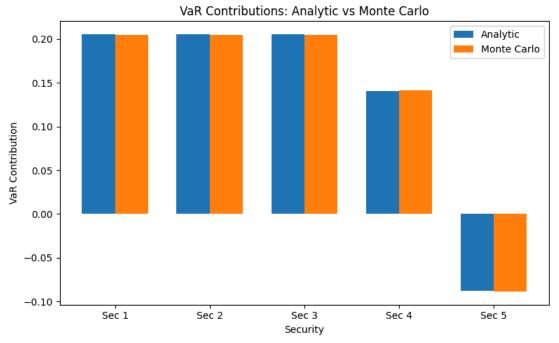
```
In [3]: def construct_A_off(n, rho_p, rho_n):
    """
    Construct an n x n matrix of off-diagonals for use in the Archakov-Hansen algorithm.
    We assume:
        - Securities 0,1,2 (first three) have target positive correlation rho_p.
        - Security 3 has zero correlation with securities 0,1,2.
```

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- Security 4 (last one) has target negative correlation rho n with securities 0,1,2 and security 3.
   The hints suggest using:
     gamma_p = - (1/n)*log((1 - rho_p)/(1 + rho_p*(n-1)))
     Parameters:
       n : int
          Dimension (here, 5).
       rho_p : float
          Target positive correlation (> 0).
       rho n : float
          Target negative correlation (< 0).
   Returns:
       A_off : np.ndarray (n x n)
          Symmetric matrix with the prescribed off-diagonals.
   A_{off} = np.zeros((n, n))
   for i in range(n):
       for j in range(i+1, n):
           # Securities 0,1,2 (first three) have positive correlation among themselves.
           if i in [0, 1, 2] and j in [0, 1, 2]:
              A_off[i, j] = gamma_p
           # Security 3 (index 3) has zero correlation with securities 0,1,2.
           elif (i in [0, 1, 2] and j == 3) or (i == 3 and j in [0, 1, 2]):
              A_{off[i, j]} = 0.0
           # Security 4 (index 4) has negative correlation with securities 0,1,2.
           elif (i in [0, 1, 2] and j == 4) or (i == 4 and j in [0, 1, 2]):
              A_off[i, j] = gamma_n
           # For the pair between security 3 and security 4, also use gamma_n.
           elif (i == 3 and j == 4) or (i == 4 and j == 3):
              A_off[i, j] = gamma_n
           else:
              A_{off[i, j]} = 0.0
   A_off = A_off + A_off.T # make symmetric
   return A off
def analytic_var_contributions(w, Sigma, alpha):
   Compute analytic VaR contributions for a portfolio with weights w
   under the assumption that losses are normally distributed.
   VaR = z_alpha * sigma_p, with sigma_p = sqrt(w' Sigma w)
   and the Euler contributions are:
     c_i = w_i * (Sigma * w)_i / sigma_p * z_alpha
   Parameters:
      w : np.ndarray (n,)
          Portfolio weights.
       Sigma : np.ndarray (n x n)
          Covariance (or correlation) matrix.
       alpha : float
          Confidence level (e.g., 0.95).
   Returns:
       var total : float
          Total portfolio VaR.
       contributions : np.ndarray (n,)
          Contributions for each security.
   sigma_p = np.sqrt(w @ Sigma @ w)
   z_alpha = norm_dist.ppf(alpha)
   var_total = z_alpha * sigma_p
   contributions = w * (Sigma @ w) / sigma_p * z_alpha
   return var_total, contributions
def monte_carlo_var_contributions(w, Sigma, alph, n_sim, seed):
   Estimate VaR contributions via Monte Carlo simulation.
   We simulate losses L \sim N(0, Sigma). With portfolio loss p = w' L,
   note that analytically, Cov(L_i, p) = (Sigma * w)_i.
   We estimate these covariances from the simulated data.
   Parameters:
       w : np.ndarray (n,)
```

```
Portfolio weights.
    Sigma : np.ndarray (n x n)
        Covariance (or correlation) matrix.
    alpha : float
        Confidence level.
    n sim : int
       Number of simulations.
    seed : int
        Random seed for reproducibility.
Returns:
    var_total : float
        Estimated portfolio VaR.
    contributions : np.ndarray (n,)
       Estimated contributions for each security.
np.random.seed(seed)
n = len(w)
# simulate losses: shape (n_sim, n)
L = np.random.multivariate_normal(mean=np.zeros(n), cov=Sigma, size=n_sim)
# portfolio loss for each simulation:
p = L @ w
sigma_p_mc = np.std(p)
z alpha = norm dist.ppf(alpha)
# Estimated VaR from simulation (should be close to analytic value)
var_total = np.quantile(p, alpha)
# Estimate covariances: since E[L_i] = 0, we have Cov(L_i, p) \sim mean(L_i * p)
cov_est = np.array([np.mean(L[:, i] * p) for i in range(n)])
# Monte Carlo marginal contributions:
contributions = w * cov_est / sigma_p_mc * z_alpha
return var_total, contributions
```

```
In [4]: # Setup portfolio and compute contributions
        # Portfolio size
        n = 5
        # Define target correlations:
        rho_p = 0.3  # target positive correlation for group 1 (securities 0,1,2)
        rho_n = -0.2 # target negative correlation for security 5 with others
        n_sim=100000
        seed = 44
        alpha = 0.95
        # Construct the off-diagonal matrix A_off using our specification
        A_off = construct_A_off(n, rho_p, rho_n)
        # Obtain the correlation matrix via the Archakov-Hansen algorithm
        C = archakov_hansen_corr(A_off)
        print("Correlation matrix C:")
        print(np.round(C, 3))
        # Assume unit volatilities so that covariance Sigma = C
        Sigma = C.copy()
        # Define portfolio weights (e.g., equal weights)
        w = np.ones(n) / n
        # Compute analytic VaR contributions
        var_analytic, contrib_analytic = analytic_var_contributions(w, Sigma, alpha)
        print("\nAnalytic portfolio VaR: {:.4f}".format(var_analytic))
        print("Analytic VaR contributions:")
        for i, c in enumerate(contrib_analytic):
            print(" Security {}: {:.4f}".format(i+1, c))
        # Compute Monte Carlo VaR contributions
        var_mc, contrib_mc = monte_carlo_var_contributions(w, Sigma, alpha, n_sim=n_sim, seed=seed)
        print("\nMonte Carlo portfolio VaR: {:.4f}".format(var_mc))
        print("Monte Carlo VaR contributions:")
        for i, c in enumerate(contrib_mc):
            print(" Security {}: {:.4f}".format(i+1, c))
        # PLot
        indices = np.arange(n)
        width = 0.35
        fig, ax = plt.subplots(figsize=(8, 5))
        bar1 = ax.bar(indices - width/2, contrib_analytic, width, label='Analytic')
        bar2 = ax.bar(indices + width/2, contrib_mc, width, label='Monte Carlo')
        ax.set xlabel('Security')
```

```
ax.set ylabel('VaR Contribution')
 ax.set_title('VaR Contributions: Analytic vs Monte Carlo')
 ax.set xticks(indices)
 ax.set_xticklabels([f'Sec {i+1}' for i in indices])
 ax.legend()
 plt.tight_layout()
 plt.show()
Correlation matrix C:
[[ 1. 0.305 0.305 0.067 -0.404]
[ 0.305 1.
                0.305 0.067 -0.404]
 [ 0.305 0.305 1. 0.067 -0.404]
 [ 0.067 0.067 0.067 1. -0.332]
[-0.404 -0.404 -0.404 -0.332 1. ]]
Analytic portfolio VaR: 0.6700
Analytic VaR contributions:
 Security 1: 0.2057
 Security 2: 0.2057
 Security 3: 0.2057
 Security 4: 0.1405
 Security 5: -0.0877
Monte Carlo portfolio VaR: 0.6655
Monte Carlo VaR contributions:
 Security 1: 0.2046
 Security 2: 0.2046
 Security 3: 0.2044
 Security 4: 0.1413
 Security 5: -0.0889
```



Question 4

The function optimize_portfolio simulates M return scenarios and then sets up a CVaR constraint using the Rockafellar-Uryasev linear formulation. The variable tau represents the VaR and z the slack for excess losses. The optimization maximizes the portfolio's expected return

$$w^{\mathsf{T}}\mu + \left(1 - \sum_{i=1}^{N} w_i\right) r_T,$$

subject to the CVaR constraint being no more than 50%. In part a+b we impose $w \ge 0$ and $sumw \le 1$ (no short sales and no borrowing), while in part (c) we remove the no-short-sale constraint. For various values of M (the number of simulated scenarios used in the optimization), we solve for the portfolio and then use a large simulation sample (here, $M_{\text{sim_eval}} = 100000$) to re-estimate the portfolio's '95%' CVaR. Comparing the estimated CVaR to the '50%' target shows the bias as a function of M.

```
Simulate M sample vectors R ~ MVN(mu, Sigma)
    return np.random.multivariate normal(mu, Sigma, size=M)
def optimize_portfolio(M_sample, allow_short, CVaR_target, alpha_level):
    Simulate M_sample scenarios for the risky returns,
    then solve for the portfolio (risky weights w) that maximizes expected portfolio return
    subject to CVaR_{alpha_level} (on losses) <= CVaR_target.</pre>
    The overall portfolio return is
      r_p = w^T R + (1 - sum(w)) * r_T.
    In part (a)/(b) we require w \ge 0 and 1-sum(w) \ge 0 (i.e. no short-selling and no borrowing).
    In part (c) we allow negative w (risky short-selling) and allow borrowing (risk-free weight can be negative).
      w_opt: optimal risky weights (vector of length N)
      obj_value: achieved expected portfolio return (using sample average, computed as
                 w^T (sample mean of R) + (1 - sum(w))*r_T)
     cvx_status: status of the optimization.
    problem: the cvxpy Problem object.
    # Simulate M sample scenarios for risky returns
    R_samples = simulate_returns(M_sample, mu, Sigma) # shape: (M_sample, N)
    # Decision variable: weights on risky assets
    w = cp.Variable(N)
    # The weight on the risk-free asset is 1 - sum(w)
    # For each scenario m, portfolio return = w^T R_m + (1-sum(w))*r_T.
    portfolio returns = R samples @ w + (1 - cp.sum(w)) * r T
    # Define loss as negative portfolio return.
    L = - portfolio_returns
    # CVaR formulation: introduce auxiliary variable tau and slack variables z_m for each scenario.
    tau = cp.Variable()
    z = cp.Variable(M_sample)
    # Constraints for slack variables: z m >= L m - tau, z m >= 0.
    constraints = [z >= L - tau,
                   z >= 01
     \# \ CVaR \ constraint: \ tau \ + \ (1/(1-alpha\_level))*(1/M\_sample)*sum(z) \ <= \ CVaR\_target. 
    constraints += [tau + (1/(1 - alpha_level))*(cp.sum(z)/M_sample) <= CVaR_target]</pre>
    # Weight constraints:
    if not allow short:
        # No short-sales in risky assets and no borrowing: w \ge 0 and risk-free weight = 1 - sum(w) \ge 0.
        constraints += [w >= 0, cp.sum(w) <= 1]
    # Objective: maximize expected portfolio return (using sample average or using mu).
    # Here we use the known parameters: expected return = w^T mu + (1 - sum(w)) * T.
    objective = cp.Maximize(w@mu + (1 - cp.sum(w)) * r_T)
    prob = cp.Problem(objective, constraints)
    prob.solve(solver=cp.SCS)
    \textbf{return} \ \textbf{w.value,} \ (\textbf{w.value} \ \underline{\textbf{0}} \ \textbf{mu} \ + \ (\textbf{1 - np.sum}(\textbf{w.value})) * \textbf{r_T}), \ \textbf{prob.status,} \ \textbf{R\_samples}
# Function to estimate portfolio CVaR using a large simulation sample
def estimate_portfolio_CVaR(w, R_samples_eval, r_T, alpha_level):
    Given portfolio weights \mbox{w} (on risky assets), and a matrix of simulated risky returns,
    compute the portfolio return r_p = w^T R + (1-sum(w))*r_T and estimate CVaR_{alpha_level}
    using the sample average of the worst (1-alpha_level) fraction of losses.
    Returns the estimated CVaR.
    port_returns = R_samples_eval @ w + (1 - np.sum(w)) * r_T
    losses = -port_returns
    # Compute VaR (the alpha-level quantile of losses)
    VaR = np.quantile(losses, alpha_level)
    # CVaR is the average loss in the tail (losses above VaR)
    CVaR_est = np.mean(losses[losses >= VaR])
    return CVaR_est
# Run optimization and bias investigation for a given M_sample
def run_experiment(M_sample, allow_short, CVaR_target, alpha_level):
```

```
Solve for the optimal portfolio with M sample simulated scenarios.
   Then simulate a large evaluation sample to estimate the actual CVaR.
   w_opt, expected return, estimated CVaR.
   w_opt, exp_ret, status, _ = optimize_portfolio(M_sample, allow_short=allow_short, CVaR_target=CVaR_target, alpha_level=alp
   # Use a large number of simulations to evaluate the actual CVaR.
   R_samples_eval = simulate_returns(M_sim_eval, mu, Sigma)
   CVaR_est = estimate_portfolio_CVaR(w_opt, R_samples_eval, r_T, alpha_level=alpha_level)
   return w_opt, exp_ret, CVaR_est
# Parameters
N = 10
                         # number of risky securities
T = 0.5
                        # time horizon in vears
                        # annual continuously compounded risk-free rate
r_{annual} = 0.03
r_T = r_annual * T
                         # Log risk-free return over T (approximation)
                        # number of simulations for evaluating CVaR
M_sim_eval = 100000
CVaR_target = 0.50
                        # target CVaR Level
alpha_level = 0.95
                         # confidence level for CVaR
# Construct mean vector and covariance matrix for the risky assets.
mu = np.array([0.10]*(N//2) + [0.20]*(N - N//2)) # first half 10%, second half 20%
sigma = 0.30
var = sigma**2
# Off-diagonals: correlation 0.3 so covariance = 0.3*0.3*0.3 = 0.027
corr = 0.3
Sigma = np.full((N, N), corr * sigma * sigma)
np.fill_diagonal(Sigma, var)
# Part (a) and (b): No short-selling on risky assets.
np.random.seed(123)
M_values = [500, 1000, 5000, 10000] # different numbers of scenarios for the optimization sample
results_no_short = {}
print("=== Optimization with No Short Sales on Risky Assets ===")
for M sample in M values:
   w_opt, exp_ret, CVaR_est = run_experiment(M_sample, allow_short=False, CVaR_target=CVaR_target, alpha_level=alpha_level)
   results_no_short[M_sample] = (w_opt, exp_ret, CVaR_est)
   print(f"M_sample = {M_sample}:")
   print(" Optimal risky weights:", np.round(w_opt, 4))
   print(" Expected portfolio return:", round(exp_ret, 4))
   print(" Estimated CVaR (95%):", round(CVaR_est, 4))
   print(" CVaR Target: 0.50")
   print(" CVaR Bias: {:.4f}\n".format(CVaR_est - 0.50))
# Part (c): Allow short-sales (no nonnegativity on risky weights)
results_short = {}
print("=== Optimization Allowing Short Sales on Risky Assets ===")
for M sample in M values:
   w_opt, exp_ret, CVaR_est = run_experiment(M_sample, allow_short=True, CVaR_target=CVaR_target, alpha_level=alpha_level)
   results_short[M_sample] = (w_opt, exp_ret, CVaR_est)
   print(f"M_sample = {M_sample}:")
   print(" Optimal risky weights:", np.round(w_opt, 4))
   print(" Expected portfolio return:", round(exp_ret, 4))
   print(" Estimated CVaR (95%):", round(CVaR_est, 4))
   print(" CVaR Target: 0.50")
   print(" CVaR Bias: {:.4f}\n".format(CVaR_est - 0.50))
# For one chosen M sample, compare optimal weights for the two cases.
w_no_short, exp_ret_no_short, CVaR_no_short = results_no_short[M_plot]
w_short, exp_ret_short, CVaR_short = results_short[M_plot]
indices = np.arange(N)
width = 0.35
plt.figure(figsize=(8,5))
plt.bar(indices - width/2, w_no_short, width, label='No Short Sales')
plt.bar(indices + width/2, w_short, width, label='Allow Short Sales')
plt.xlabel('Risky Security (Index)')
plt.ylabel('Optimal Weight')
plt.title(f'Optimal Risky Weights (M_sample = {M_plot})')
plt.xticks(indices, [f'Sec {i+1}' for i in indices])
plt.legend()
plt.tight_layout()
plt.show()
```

```
=== Optimization with No Short Sales on Risky Assets ===
M_sample = 500:
 Optimal risky weights: [-0.
                                       0. 0.
                                                     0.
                                                             0.2302 0.2043 0.1204 0.2382
                               0.
 0.20691
 Expected portfolio return: 0.2
 Estimated CVaR (95%): 0.2131
 CVaR Target: 0.50
 CVaR Bias: -0.2869
M_sample = 1000:
 Optimal risky weights: [0.
                                                       0.2279 0.2057 0.2103 0.176 0.18 ]
                              0.
                                      0.
                                          0.
 Expected portfolio return: 0.2
 Estimated CVaR (95%): 0.2081
 CVaR Target: 0.50
 CVaR Bias: -0.2919
M_sample = 5000:
 Optimal risky weights: [-0.
                                -0.
                                         0.
                                                 0.
                                                       -0.
                                                                0.1961 0.1833 0.1991 0.2245
 0.1969]
 Expected portfolio return: 0.2
 Estimated CVaR (95%): 0.2102
 CVaR Target: 0.50
 CVaR Bias: -0.2898
M_sample = 10000:
 Optimal risky weights: [-0.
                                        0. -0.
                                                      0. 0.218 0.1892 0.2022 0.2004
 0.1901]
 Expected portfolio return: 0.2
 Estimated CVaR (95%): 0.2121
 CVaR Target: 0.50
 CVaR Bias: -0.2879
=== Optimization Allowing Short Sales on Risky Assets ===
M_sample = 500:
 Optimal risky weights: [-0.1783 -0.3339 -0.4108 -0.3343 -0.1621 0.5207 0.9179 0.6743 0.6664
 0.3228]
 Expected portfolio return: 0.4682
 Estimated CVaR (95%): 0.5473
 CVaR Target: 0.50
 CVaR Bias: 0.0473
M_sample = 1000:
 Optimal risky weights: [-0.3294 -0.2164 -0.3517 -0.0133 -0.3259 0.4047 0.6255 0.7014 0.6067
 0.6776]
 Expected portfolio return: 0.4678
 Estimated CVaR (95%): 0.5216
 CVaR Target: 0.50
 CVaR Bias: 0.0216
M_sample = 5000:
 Optimal risky weights: [-0.1561 -0.1643 -0.2315 -0.1842 -0.1802 0.4888 0.5647 0.6225 0.5724
 0.6218]
 Expected portfolio return: 0.4681
 Estimated CVaR (95%): 0.4953
 CVaR Target: 0.50
 CVaR Bias: -0.0047
M_sample = 10000:
 Optimal risky weights: [-0.1391 -0.1608 -0.1624 -0.2169 -0.1995 0.5204 0.5632 0.6161 0.5939
 0.5661]
 Expected portfolio return: 0.4693
 Estimated CVaR (95%): 0.4995
 CVaR Target: 0.50
 CVaR Bias: -0.0005
```

Optimal Risky Weights (M_sample = 5000) No Short Sales Allow Short Sales 0.4 Optimal Weight 0.2 0.0 -0.2 Sec 1 Sec 2 Sec 3 Sec 4 Sec 5 Sec 6 Sec 7 Sec 8 Sec 9 Sec 10 Risky Security (Index)

b: CVaR as function of M

d: Estimation Error Impact

Estimation error in the simulated return distribution can have a significant impact on the portfolio chosen in both cases. When the sample size M is small, the optimized portfolio may exploit noise in the CVaR constraint, leading to portfolios that in out-of-sample evaluation violate the intended CVaR threshold (i.e. the estimated CVaR is biased). In the case with short-sales allowed, the optimizer may take extreme positions (long/short) to artificially lower the sample CVaR. In contrast, imposing no-short-sale constraints tends to provide more stable and robust portfolios. Overall, estimation error may result in "portfolios that are overly aggressive in their expected return but underestimate tail risk.