

# Introduction to Machine Learning, Fall 2023

## Homework 1

(Due Thursday, Oct. 26 at 11:59pm (CST))

October 11, 2023

1. [10 points] [Math review] Suppose  $\{\mathbf{X}_1, \mathbf{X}_2, \dots, \mathbf{X}_n\}$  form a random sample from a multivariate distribution:
- (a) Prove that the covariance of  $\mathbf{X}_i$  is a semi positive definite matrix. [3 points]
  - (b) Assuming  $\mathbf{X}_i \sim \mathcal{N}(\mu, \Sigma)$  which is a multivariate normal distribution, and samples  $X_i$ , derive the the log-likelihood  $l(\mu, \Sigma)$  and MLE of  $\mu$  [4 points]
  - (c) Suppose  $\hat{\theta}$  is an unbiased estimator of  $\theta$  and  $\text{Var}(\hat{\theta}) > 0$ . Prove that  $(\hat{\theta})^2$  is not an unbiased estimator of  $\theta^2$ . [3 points]

(a) Let  $X = (X_1, X_2, \dots, X_n)^T$  be a random  $n$  dimensional vector.  
let  $\mu_i$  and  $\sigma_i^2$  be the mean and variance of  $X_i$ ,  
and  $\sigma_{ij}$  be the covariance between  $X_i$  and  $X_j$  for  $i \neq j$ ,  
So the covariance matrix is a symmetric  $n \times n$  matrix

$$\Sigma = \begin{bmatrix} \sigma_1^2 & \sigma_{12} & \dots & \sigma_{1n} \\ \sigma_{21} & \sigma_2^2 & \dots & \sigma_{2n} \\ \vdots & \vdots & \ddots & \vdots \\ \sigma_{n1} & \sigma_{n2} & \dots & \sigma_n^2 \end{bmatrix} = E[(X-\mu)(X-\mu)^T], \mu = (\mu_1, \mu_2, \dots, \mu_n)^T$$

$$\begin{aligned} \text{For all } a \in \mathbb{R}^n \quad a^T \Sigma a &= E[a^T (X-\mu)(X-\mu)^T a] \\ &= E[(a^T (X-\mu))^2] \geq 0 \end{aligned}$$

So the covariance of  $X_i$  is a semi positive definite matrix.

(b) The PDF of the multivariate normal distribution

$$\text{is } f(X_i | \mu, \Sigma) = \frac{1}{(2\pi)^{\frac{n}{2}} |\Sigma|^{\frac{1}{2}}} \exp\left(-\frac{1}{2} (X_i - \mu)^T \Sigma^{-1} (X_i - \mu)\right)$$

so the log-likelihood

$$l(\mu, \Sigma) = \sum_{i=1}^N \log\left(\frac{1}{(2\pi)^{\frac{n}{2}} |\Sigma|^{\frac{1}{2}}}\right) - \frac{1}{2} (X_i - \mu)^T \Sigma^{-1} (X_i - \mu)$$

$$\frac{\partial l}{\partial \mu} = \sum_{i=1}^N \Sigma^{-1} (X_i - \mu) = 0$$

$$\mu_{MLE} = \frac{1}{N} \sum_{i=1}^N X_i$$

(c) Since  $\hat{\theta}$  is an unbiased estimator of  $\theta$ ,  $E(\hat{\theta}) = \theta$

we need to prove  $E(\hat{\theta}^2) \neq \theta^2$

$$\text{Var}(\hat{\theta}^2) = (\text{Var}(\hat{\theta}))^2 > 0$$

$$E(\hat{\theta}^2) = \text{Var}(\hat{\theta}) + (E(\hat{\theta}))^2 = \text{Var}(\hat{\theta}) + \theta^2$$

since  $\text{Var}(\hat{\theta}) > 0$

$$E(\hat{\theta}^2) > \theta^2$$

so  $\hat{\theta}^2$  is not an unbiased estimator of  $\theta^2$

2. [10 points] Consider real-valued variables  $X$  and  $Y$ , in which  $Y$  is generated conditional on  $X$  according to

$$Y = aX + b + \epsilon, \text{ where } \epsilon \sim \mathcal{N}(0, \sigma^2).$$

Here  $\epsilon$  is an independent variable, called a noise term, which is drawn from a Gaussian distribution with mean 0, and variance  $\sigma^2$ . This is a single variable linear regression model, where  $a$  is the only weight parameter and  $b$  denotes the intercept. The conditional probability of  $Y$  has a distribution  $p(Y|X, a, b) \sim \mathcal{N}(aX + b, \sigma^2)$ , so it can be written as:

$$p(Y|X, a, b) = \frac{1}{\sqrt{2\pi}\sigma} \exp\left(-\frac{1}{2\sigma^2}(Y - aX - b)^2\right).$$

- (a) Assume we have a training dataset of  $n$  i.i.d. pairs  $(x_i, y_i)$ ,  $i = 1, 2, \dots, n$ , and the likelihood function is defined by  $L(a, b) = \prod_{i=1}^n p(y_i|x_i, a, b)$ . Please write the Maximum Likelihood Estimation (MLE) problem for estimating  $a$  and  $b$ . [3 points]
- (b) Estimate the optimal solution of  $a$  and  $b$  by solving the MLE problem in (a). [4 points]
- (c) Based on the result in (b), argue that the learned linear model  $f(X) = aX + b$ , always passes through the point  $(\bar{x}, \bar{y})$ , where  $\bar{x} = \frac{1}{n} \sum_{i=1}^n x_i$  and  $\bar{y} = \frac{1}{n} \sum_{i=1}^n y_i$  denote the sample means. [3 points]

(a) 
$$\operatorname{argmax}_{a, b} \prod_{i=1}^n \frac{1}{\sqrt{2\pi}\sigma} \exp\left(-\frac{1}{2\sigma^2}(y_i - ax_i - b)^2\right)$$

(b) 
$$\hat{a} = \frac{\sum_{i=1}^N (x_i - \bar{x})(y_i - \bar{y})}{\sum_{i=1}^N (x_i - \bar{x})^2} \quad \hat{b} = \bar{y} - \hat{a}\bar{x}$$

where  $\bar{x} = \frac{1}{N} \sum_{i=1}^N x_i$ ,  $\bar{y} = \frac{1}{N} \sum_{i=1}^N y_i$

(c) We know that  $\hat{y} = \hat{a}x_i + \hat{b}$

plug  $(\bar{x}, \bar{y})$  into the equation

$$\bar{y} = \hat{a}\bar{x} + \hat{b} = \hat{a}\bar{x} + \bar{y} - \hat{a}\bar{x} = \bar{y}$$

so the learned linear model  $f(X) = aX + b$  always passes through the point  $(\bar{x}, \bar{y})$ .

3. [10 points] [Regression and Classification]

- (a) When we talk about linear regression, what does 'linear' regard to? [2 points]
- (b) Assume that there are  $n$  given training examples  $\{(x_1, y_1), (x_2, y_2), \dots, (x_n, y_n)\}$ , where each input data point  $x_i$  has  $m$  real valued features. When  $m > n$ , the linear regression model is equivalent to solving an under-determined system of linear equations  $\mathbf{y} = \mathbf{X}\beta$ . One popular way to estimate  $\beta$  is to consider the so-called ridge regression:

$$\underset{\beta}{\operatorname{argmin}} \|\mathbf{y} - \mathbf{X}\beta\|_2^2 + \lambda \|\beta\|_2^2$$

for some  $\lambda > 0$ . This is also known as Tikhonov regularization.

Show that the optimal solution  $\beta_*$  to the above optimization problem is given by

$$\beta_* = (\mathbf{X}^T \mathbf{X} + \lambda \mathbf{I})^{-1} \mathbf{X}^T \mathbf{y}$$

Hint: You need to prove that given  $\lambda > 0$ ,  $\mathbf{X}^T \mathbf{X} + \lambda \mathbf{I}$  is invertible. [5 points]

- (c) Is the given data set linear separable? If yes, construct a linear hypothesis function to separate the given data set. If no, explain the reason. [3 points]

Data	(1,3)	(4,4)	(3,-6)	(-2,1)	(-3,5)	(-6,-4)
Label	+1	-1	-1	+1	-1	-1

(a) It refers to the relationship between independent and dependent variables.

A linear relationship means that a change in the independent variables is associated with a constant change in the dependent variables.

(b) when  $\lambda > 0$ , let  $v$  be a non zero vector

$$\text{the } v^T (\mathbf{X}^T \mathbf{X} + \lambda \mathbf{I}) v = v^T \mathbf{X}^T \mathbf{X} v + \lambda v^T v = \|\mathbf{X}v\|_2^2 + \lambda \|v\|_2^2 > 0$$

so  $\mathbf{X}^T \mathbf{X} + \lambda \mathbf{I}$  is invertible when  $\lambda > 0$

$$\text{define } f(\beta) = (\mathbf{y} - \mathbf{X}\beta)^T (\mathbf{y} - \mathbf{X}\beta) + \lambda \beta^T \beta$$

$$\text{then } f'(\beta) = -2\mathbf{X}^T (\mathbf{y} - \mathbf{X}\beta) + 2\lambda \beta,$$

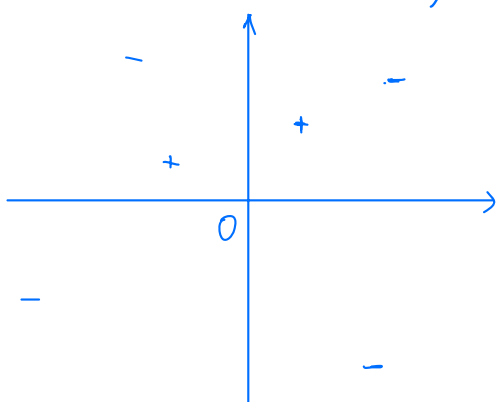
$$f''(\beta) = 2\mathbf{X}^T \mathbf{X} + 2\lambda \mathbf{I} > 0$$

thus the optimal solution  $\beta_*$  is  $f'(\beta) = 0$

$$\beta (2\mathbf{X}^T \mathbf{X} + 2\lambda) - 2\mathbf{X}^T \mathbf{y} = 0$$

$$\beta = (\mathbf{X}^T \mathbf{X} + \lambda \mathbf{I})^{-1} \mathbf{X}^T \mathbf{y}$$

(c)



As the plot shows, it doesn't exist a single straight line to separate the +1 and -1 labelled points. So the given data set is not linear separable.