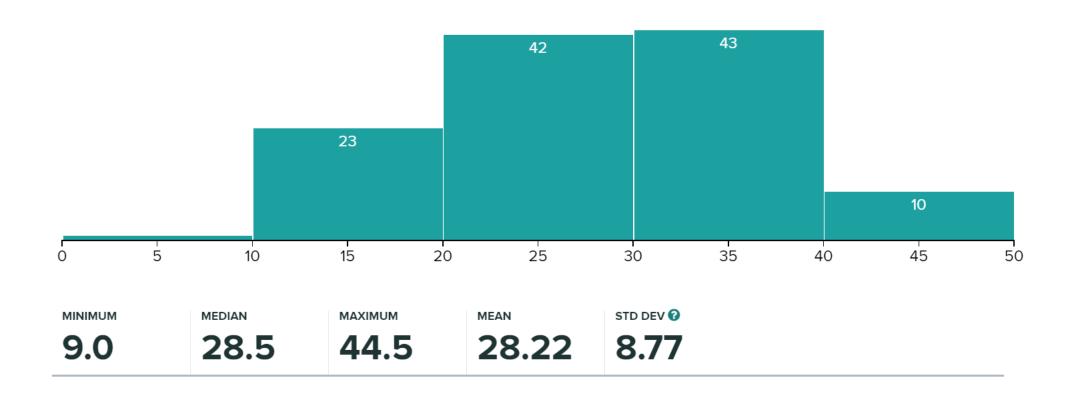
### Midterm



Request regrading on Gradescope by this Saturday

#### Announcement

Homework 4

■ Due: Nov. 18, 11:59pm

Programming Assignment 4

■ Due: Nov. 25, 11:59pm

## Probabilistic Reasoning over Time

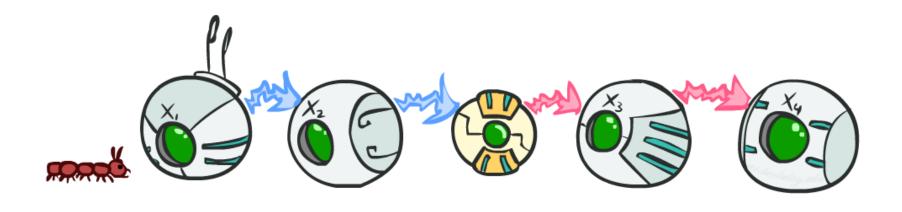


AIMA Chapter 15

## **Uncertainty and Time**

- Often, we want to reason about a sequence of observations
  - Speech recognition
  - Robot localization
  - Medical monitoring
  - User attention
- Need to introduce time into our models

## Markov Models



## Markov Models (aka Markov chain/process)

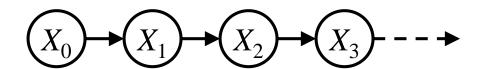
- Assume discrete variables that share the same finite domain
  - Values in the domain is called the states

$$(X_0) \rightarrow (X_1) \rightarrow (X_2) \rightarrow (X_3) \rightarrow (X_3) \rightarrow P(X_0)$$

$$P(X_0) \qquad P(X_t \mid X_{t-1})$$

- The *transition model*  $P(X_t \mid X_{t-1})$  specifies how the state evolves over time
- Stationarity assumption: same transition probabilities at all time steps
- Joint distribution  $P(X_0,...,X_T) = P(X_0) \prod_t P(X_t \mid X_{t-1})$

## Quiz: are Markov models a special case of Bayes nets?

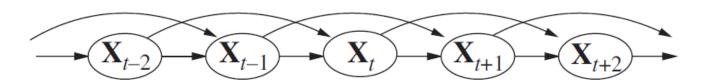


- Yes and no!
- Yes:
  - Directed acyclic graph, joint = product of conditionals
- No:
  - Infinitely many variables (unless we truncate)
  - Repetition of transition model not part of standard Bayes net syntax

## Markov Assumption: Conditional Independence



- Markov assumption:  $X_{t+1}$ , ... are independent of  $X_0$ , ...,  $X_{t-1}$  given  $X_t$ 
  - Past and future independent given the present
  - Each time step only depends on the previous
- This is a first-order Markov model
- A kth-order model allows dependencies on k earlier steps



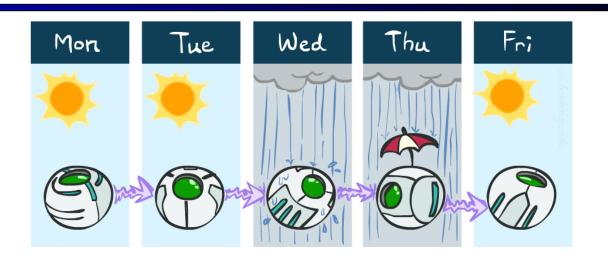
## Example: Weather

- States {rain, sun}
- Initial distribution  $P(X_0)$

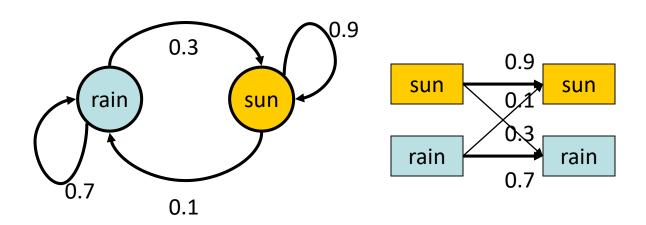
P(X <sub>0</sub> )	
sun	rain
0.5	0.5

• Transition model  $P(X_t \mid X_{t-1})$ 

X <sub>t-1</sub>	$P(X_{t} X_{t-1})$	
	sun	rain
sun	0.9	0.1
rain	0.3	0.7



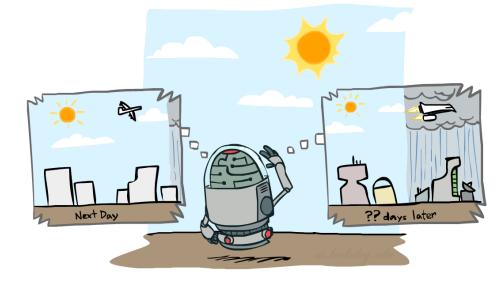
Two new ways of representing the same CPT



## Weather prediction

■ Time 0: <0.5,0.5>

X <sub>t-1</sub>	$P(X_{t} X_{t-1})$	
	sun	rain
sun	0.9	0.1
rain	0.3	0.7

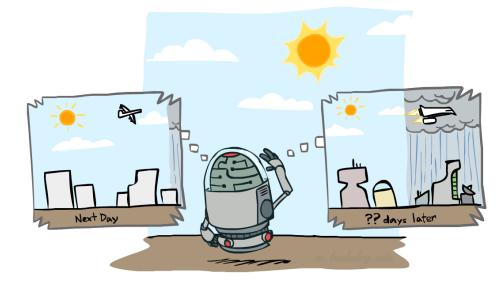


- What is the weather like at time 1?
  - $P(X_1) = \sum_{X_0} P(X_1, X_0 = X_0)$
  - $= \sum_{x_0} P(X_0 = x_0) P(X_1 \mid X_0 = x_0)$
  - **=** 0.5<0.9,0.1> + 0.5<0.3,0.7> = <0.6,0.4>

## Weather prediction, contd.

■ Time 1: <0.6,0.4>

<b>X</b> <sub>t-1</sub>	$P(X_{t} X_{t-1})$	
	sun	rain
sun	0.9	0.1
rain	0.3	0.7

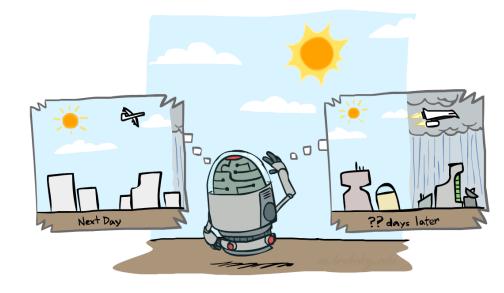


- What is the weather like at time 2?
  - $P(X_2) = \sum_{X_1} P(X_2, X_1 = X_1)$
  - $= \sum_{X_1} P(X_1 = X_1) P(X_2 \mid X_1 = X_1)$
  - = 0.6 < 0.9, 0.1 > + 0.4 < 0.3, 0.7 > = < 0.66, 0.34 >

## Weather prediction, contd.

■ Time 2: <0.66,0.34>

X <sub>t-1</sub>	$P(X_{t}   X_{t-1})$	
	sun	rain
sun	0.9	0.1
rain	0.3	0.7



- What is the weather like at time 3?
  - $P(X_3) = \sum_{X_2} P(X_3, X_2 = X_2)$
  - $= \sum_{X_2} P(X_2 = X_2) P(X_3 \mid X_2 = X_2)$
  - = 0.66 < 0.9, 0.1 > + 0.34 < 0.3, 0.7 > = < 0.696, 0.304 >

## Forward algorithm (simple form)

• What is the state at time t (given an initial distribution  $P(X_0)$ )?

$$P(X_{t}) = \sum_{X_{t-1}} P(X_{t}, X_{t-1} = X_{t-1})$$

$$= \sum_{X_{t-1}} P(X_{t-1} = X_{t-1}) P(X_{t} \mid X_{t-1} = X_{t-1})$$
Probability from previous iteration

Transition model

Iterate this update starting at t=0

### Example Run of Mini-Forward Algorithm

From initial observation of sun

X <sub>t-1</sub>	X <sub>t</sub>	$P(X_{t}   X_{t-1})$
sun	sun	0.9
sun	rain	0.1
rain	sun	0.3
rain	rain	0.7

From initial observation of rain

• From yet another initial distribution  $P(X_0)$ :

$$\left\langle \begin{array}{c} p \\ 1-p \\ P(X_0) \end{array} \right\rangle \qquad \cdots \qquad \left\langle \begin{array}{c} 0.75 \\ 0.25 \\ P(X_{\infty}) \end{array} \right\rangle$$

## **Stationary Distributions**

#### For most chains:

- Influence of the initial distribution gets less and less over time.
- The distribution we end up in is independent of the initial distribution

#### Stationary distribution:

- The distribution we end up with is called the stationary distribution  $P_{\infty}$  of the chain
- It satisfies

$$P_{\infty}(X) = P_{\infty+1}(X) = \sum_{x} P(X|x)P_{\infty}(x)$$

## **Example: Stationary Distributions**

Computing the stationary distribution

$$X_0 \rightarrow X_1 \rightarrow X_2 \rightarrow X_3 \rightarrow X_3$$

$$P_{\infty}(sun) = P(sun|sun)P_{\infty}(sun) + P(sun|rain)P_{\infty}(rain)$$

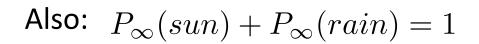
$$P_{\infty}(rain) = P(rain|sun)P_{\infty}(sun) + P(rain|rain)P_{\infty}(rain)$$

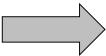
$$P_{\infty}(sun) = 0.9P_{\infty}(sun) + 0.3P_{\infty}(rain)$$

$$P_{\infty}(rain) = 0.1P_{\infty}(sun) + 0.7P_{\infty}(rain)$$

$$P_{\infty}(sun) = 3P_{\infty}(rain)$$

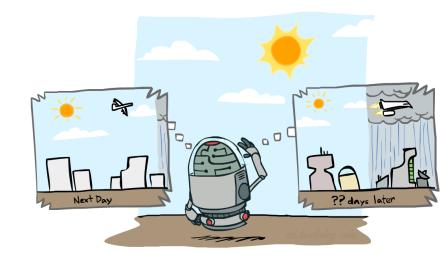
$$P_{\infty}(rain) = 1/3P_{\infty}(sun)$$





$$P_{\infty}(sun) = 3/4$$

$$P_{\infty}(rain) = 1/4$$



X <sub>t-1</sub>	X <sub>t</sub>	$P(X_{t}   X_{t-1})$
sun	sun	0.9
sun	rain	0.1
rain	sun	0.3
rain	rain	0.7

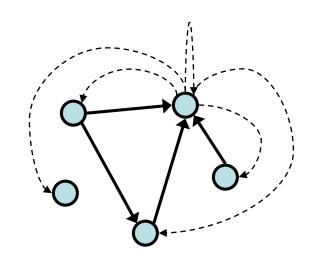
### Application of Stationary Distribution: Web Link Analysis

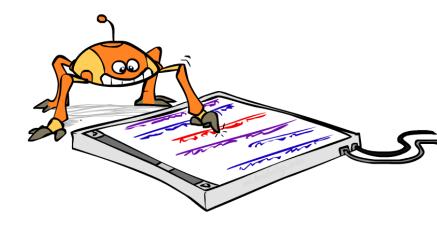
#### Web browsing

- Each web page is a state
- Initial distribution: uniform over pages
- Transitions:
  - With prob. c, uniform jump to a random page
  - With prob. 1-c, follow a random outlink

#### Stationary distribution: PageRank

- Will spend more time on highly reachable pages
- Google 1.0 returned the set of pages containing all your keywords in decreasing rank
- Now: use link analysis along with many other factors (rank actually getting less important)





## Application of Stationary Distributions: Gibbs Sampling

- Each joint instantiation over all hidden and query variables is a state:  $\{X_1, ..., X_n\} = H \cup Q$
- Transitions:
  - Pick a variable and resample its value conditioned on its Markov blanket
- Stationary distribution:
  - Conditional distribution  $P(X_1, X_2, ..., X_n | e_{1_i}, ..., e_m)$
  - When running Gibbs sampling long enough, we get a sample from the desired distribution

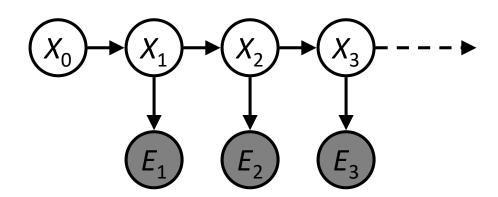


## Hidden Markov Models



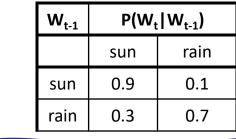
#### Hidden Markov Models

- Usually the true state is not observed directly
  - E.g., you stay indoor and cannot see the weather, but you can see if people come in with umbrella or not.
- Hidden Markov models (HMMs)
  - Underlying Markov chain over states X
  - You observe evidence E at each time step





## Example: Weather HMM

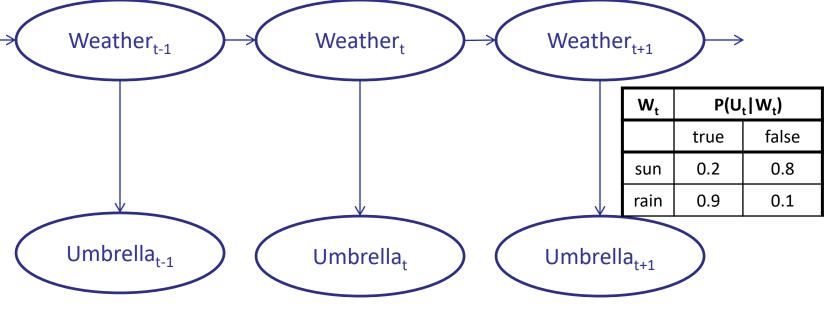


An HMM is defined by:

• Initial distribution:  $P(X_0)$ 

■ Transition model:  $P(X_t | X_{t-1})$ 

■ Emission model:  $P(E_t | X_t)$ 





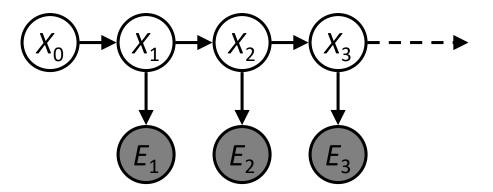


## HMM as probability model

- Joint distribution for Markov model:  $P(X_0,...,X_T) = P(X_0) \prod_{t=1:T} P(X_t \mid X_{t-1})$
- Joint distribution for hidden Markov model:

$$P(X_0, X_1, E_1, ..., X_T, E_T) = P(X_0) \prod_{t=1:T} P(X_t \mid X_{t-1}) P(E_t \mid X_t)$$

- Independence in HMM
  - Future states are independent of the past given the present
  - Current evidence is independent of everything else given the current state



## Real HMM Examples

- Speech recognition HMMs:
  - Observations are acoustic signals (continuous valued)
  - States are specific positions in specific words (so, tens of thousands)
- Machine translation HMMs:
  - Observations are words (tens of thousands)
  - States are translation options
- Robot tracking:
  - Observations are range readings (continuous)
  - States are positions on a map (continuous)
- Molecular biology:
  - Observations are nucleotides ACGT
  - States are coding/non-coding/start/stop/splice-site etc.

### Inference tasks

- Useful notation:  $X_{a:b} = X_a$ ,  $X_{a+1}$ , ...,  $X_b$
- Filtering:  $P(X_t | e_{1:t})$ 
  - belief state posterior distribution over the most recent state given all evidence
  - Ex: robot localization
- **Prediction**:  $P(X_{t+k}|e_{1:t})$  for k > 0
  - posterior distribution over a future state given all evidence
- Smoothing:  $P(X_k | e_{1:t})$  for  $0 \le k < t$ 
  - posterior distribution over a past state given all evidence
- Most likely explanation: arg  $\max_{x_{0:t}} P(x_{0:t} \mid e_{1:t})$ 
  - Ex: speech recognition, decoding with a noisy channel

- Filtering: infer current state given all evidence
- Aim: a recursive filtering algorithm of the form

$$P(X_{t+1}|e_{1:t+1}) = g(e_{t+1}, P(X_t|e_{1:t}))$$

Apply Bayes' rule

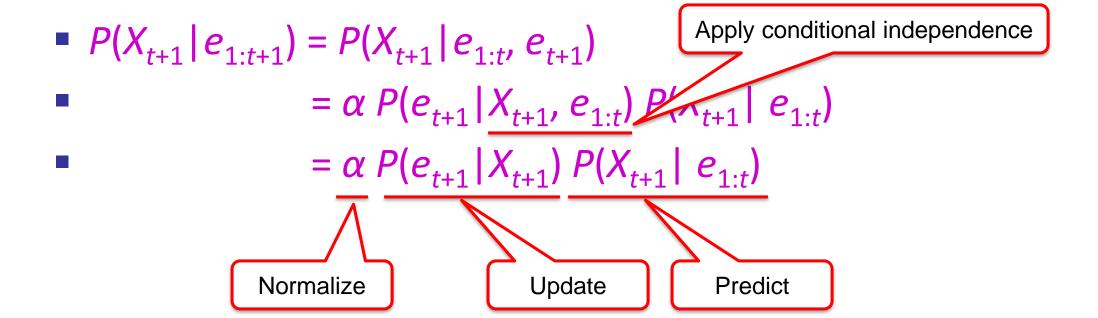
$$P(X_{t+1} | e_{1:t+1}) = P(X_{t+1} | e_{1:t}, e_{t+1})$$

$$= \alpha P(e_{t+1} | X_{t+1}, e_{1:t}) P(X_{t+1} | e_{1:t})$$

$$\alpha = 1 / P(e_{t+1} | e_{1:t})$$

- Filtering: infer current state given all evidence
- Aim: a recursive filtering algorithm of the form

$$P(X_{t+1}|e_{1:t+1}) = g(e_{t+1}, P(X_t|e_{1:t}))$$



- Filtering: infer current state given all evidence
- Aim: a recursive filtering algorithm of the form

$$P(X_{t+1}|e_{1:t+1}) = g(e_{t+1}, P(X_t|e_{1:t}))$$

■ 
$$P(X_{t+1} | e_{1:t+1}) = P(X_{t+1} | e_{1:t}, e_{t+1})$$
  
■  $= \alpha P(e_{t+1} | X_{t+1}, e_{1:t}) P(X_{t+1} | e_{1:t})$  Condition on  $X_t$   
■  $= \alpha P(e_{t+1} | X_{t+1}) P(X_{t+1} | e_{1:t})$   
■  $= \alpha P(e_{t+1} | X_{t+1}) \sum_{X_t} P(X_t | e_{1:t}) P(X_{t+1} | X_t, e_{1:t})$ 

- Filtering: infer current state given all evidence
- Aim: a recursive filtering algorithm of the form

$$P(X_{t+1}|e_{1:t+1}) = g(e_{t+1}, P(X_t|e_{1:t}))$$

$$P(X_{t+1} | e_{1:t+1}) = P(X_{t+1} | e_{1:t}, e_{t+1})$$

$$= \alpha P(e_{t+1} | X_{t+1}, e_{1:t}) P(X_{t+1} | e_{1:t})$$

$$= \alpha P(e_{t+1} | X_{t+1}) P(X_{t+1} | e_{1:t})$$

$$= \alpha P(e_{t+1} | X_{t+1}) \sum_{X_t} P(X_t | e_{1:t}) P(X_{t+1} | X_t, e_{1:t})$$

 $= \alpha P(e_{t+1}|X_{t+1}) \sum_{x_t} P(x_t | e_{1:t}) P(X_{t+1}| x_t)$ 

Apply conditional independence

$$P(X_{t+1} | e_{1:t+1}) = \alpha P(e_{t+1} | X_{t+1}) \sum_{X_t} P(X_t | e_{1:t}) P(X_{t+1} | X_t)$$
Normalize Update Predict
$$X_{t+1}$$

$$X_{t+1}$$

$$X_{t+1}$$

## Forward algorithm

• 
$$P(X_{t+1} | e_{1:t+1}) = \alpha P(e_{t+1} | X_{t+1}) \sum_{X_t} P(X_t | e_{1:t}) P(X_{t+1} | X_t)$$

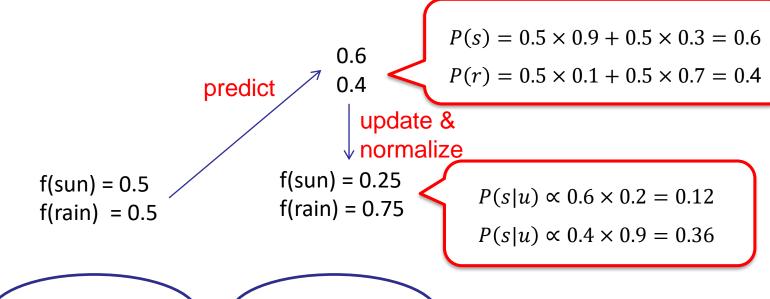
Normalize Update Predict

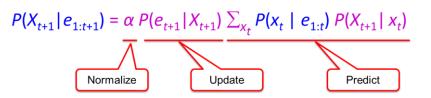
- $f_{1:t+1} = FORWARD(f_{1:t}, e_{t+1})$
- We start with  $f_{1:0} = P(X_0)$  and then iterate
- Cost per time step:  $O(|X|^2)$  where |X| is the number of states

## Example: Weather HMM

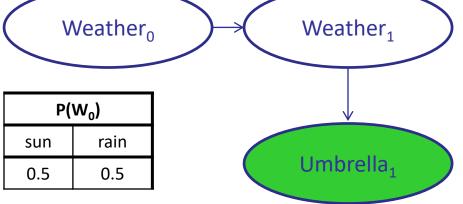








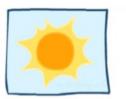
$W_{t-1}$	$P(W_t W_{t-1})$	
	sun	rain
sun	0.9	0.1
rain	0.3	0.7



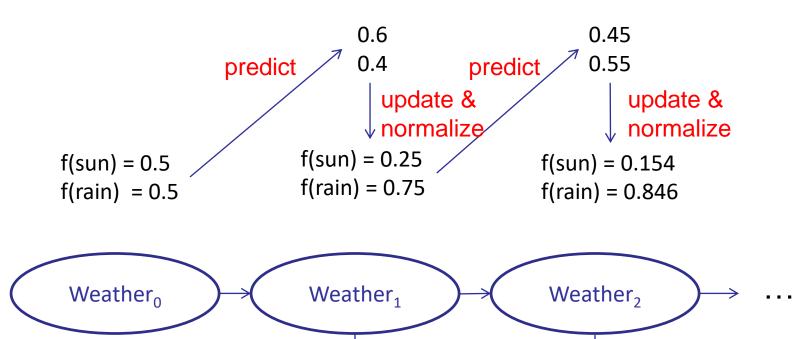
W <sub>t</sub>	$P(U_t W_t)$	
	true	false
sun	0.2	0.8
rain	0.9	0.1

## Example: Weather HMM

Umbrella<sub>2</sub>







Umbrella<sub>1</sub>

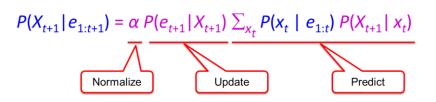
 $P(W_0)$ 

sun

0.5

rain

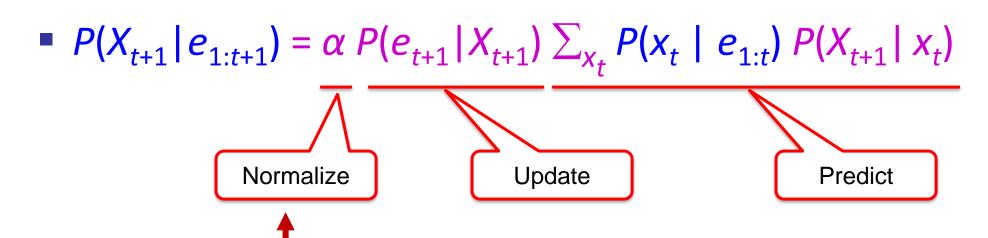
0.5



$W_{t-1}$	$P(W_t W_{t-1})$	
	sun	rain
sun	0.9	0.1
rain	0.3	0.7

W <sub>t</sub>	P(U <sub>t</sub>  W <sub>t</sub> )	
	true	false
sun	0.2	0.8
rain	0.9	0.1

## Forward algorithm

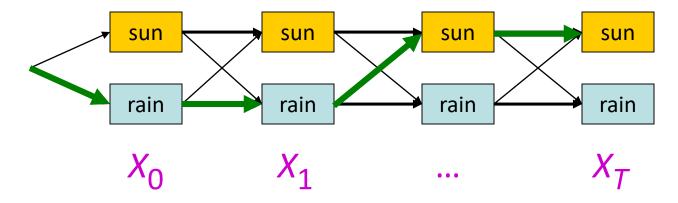


 $\alpha$  is a constant. So if we only want to compute  $P(x_t \mid e_{1:t})$ , then we can skip normalization when computing  $P(x_1 \mid e_1)$ ,  $P(x_2 \mid e_{1:2})$ , ...,  $P(x_{t-1} \mid e_{1:t-1})$ 

Q: How is the algorithm related to variable elimination?

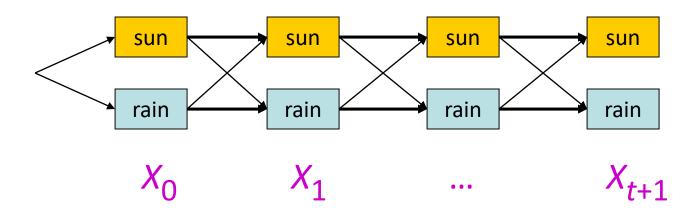
## Another view of the algorithm

State trellis: graph of states and transitions over time



- Each arc represents some transition  $x_{t-1} \rightarrow x_t$
- Each arc has weight  $P(x_t \mid x_{t-1}) P(e_t \mid x_t)$  (arcs to initial states have weight  $P(x_0)$ )
- Each path is a sequence of states
- The **product** of weights on a path is proportional to that state sequence's probability  $P(x_0) \prod_t P(x_t \mid x_{t-1}) P(e_t \mid x_t) = P(x_{1:t}, e_{1:t}) \propto P(x_{1:t} \mid e_{1:t})$

## Another view of the algorithm



Forward algorithm computes sum over all possible paths

$$P(x_{t+1} | e_{1:t+1}) = \sum_{x_{0:t}} P(x_{0:t+1} | e_{1:t+1})$$

- It uses dynamic programming to sum over all paths
  - For each state at time t, keep track of the total probability of all paths to it

$$f_{1:t+1} = \text{FORWARD}(f_{1:t}, e_{t+1})$$

$$= \alpha P(e_{t+1}|X_{t+1}) \sum_{X_t} P(X_{t+1}|X_t) f_{1:t}[X_t]$$

# **Most Likely Explanation**

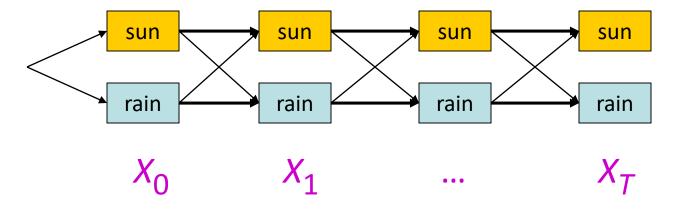


### Inference tasks

- Filtering:  $P(X_t|e_{1:t})$ 
  - belief state—input to the decision process of a rational agent
- **Prediction**:  $P(X_{t+k}|e_{1:t})$  for k > 0
  - evaluation of possible action sequences; like filtering without the evidence
- Smoothing:  $P(X_k | e_{1:t})$  for  $0 \le k < t$ 
  - better estimate of past states, essential for learning
- Most likely explanation:  $arg max_{x_{0:t}} P(x_{0:t} | e_{1:t})$ 
  - speech recognition, decoding with a noisy channel

## Most likely explanation = most probable path

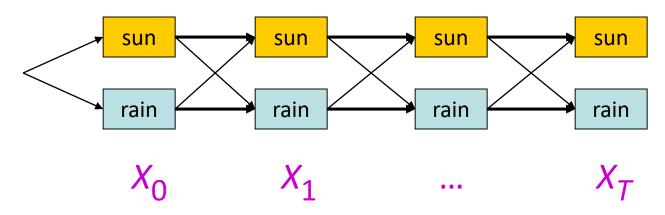
State trellis: graph of states and transitions over time



- The **product** of weights on a path is proportional to that state sequence's probability  $P(x_0) \prod_t P(x_t \mid x_{t-1}) P(e_t \mid x_t) = P(x_{0:t}, e_{1:t}) \propto P(x_{0:t} \mid e_{1:t})$
- Viterbi algorithm computes best paths

$$arg max_{x_{0:t}} P(x_{0:t} | e_{1:t})$$

## Forward / Viterbi algorithms



### Viterbi Algorithm (max)

For each state at time *t*, keep track of the (unnormalized) *maximum probability of any path* to it:

$$\mathbf{m}_{1:t+1}(\mathbf{x}_{t+1}) = \max_{\mathbf{x}_{1:t}} P(\mathbf{x}_{1:t+1} | \mathbf{e}_{1:t+1})$$

$$m_{1:t+1} = VITERBI(m_{1:t}, e_{t+1})$$
  
=  $P(e_{t+1}|X_{t+1}) \max_{X_t} P(X_{t+1}|X_t) m_{1:t}[X_t]$ 

### Forward Algorithm (sum)

For each state at time *t*, keep track of the *total probability of all paths* to it:

$$f_{1:t+1}(x_{t+1}) = P(x_{t+1} | e_{1:t+1})$$
  
=  $\sum_{x_{t+1}} P(x_{1:t+1} | e_{1:t+1})$ 

$$f_{1:t+1} = \text{FORWARD}(f_{1:t}, e_{t+1})$$

$$= \alpha P(e_{t+1}|X_{t+1}) \sum_{X_t} P(X_{t+1}|X_t) f_{1:t}[X_t]$$

## Viterbi algorithm contd.

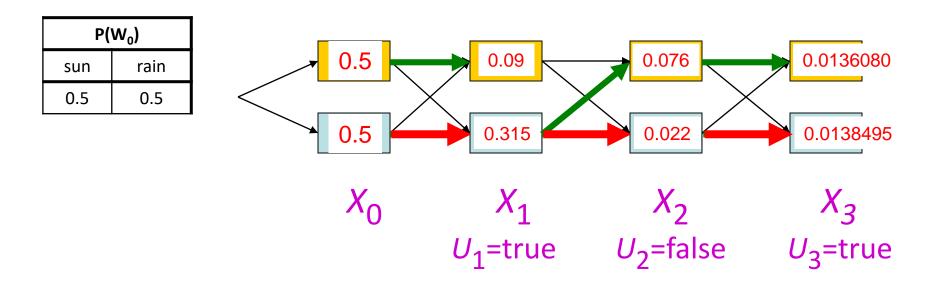
P(W <sub>0</sub> )			•		
sun	rain	0.5	0.09	sun	sun
0.5	0.5		$\times$ $>$	$\times$ $\times$	
		0.5	rain	rain	rain
		$X_{\Omega}$	$X_1$	$X_2$	$X_3$
			$U_1$ =true	$U_2$ =false	U <sub>3</sub> =true

$W_{t-1}$	$P(W_t W_{t-1})$			
	sun	rain		
sun	0.9	0.1		
rain	0.3	0.7		

$W_{t}$	P(U <sub>t</sub>  W <sub>t</sub> )			
	true	false		
sun	0.2	0.8		
rain	0.9	0.1		

$$m_{1:1}(\text{sun}) = 0.2 \times \text{max}(0.9 \times 0.5, 0.3 \times 0.5) = 0.09$$

### Viterbi algorithm contd.

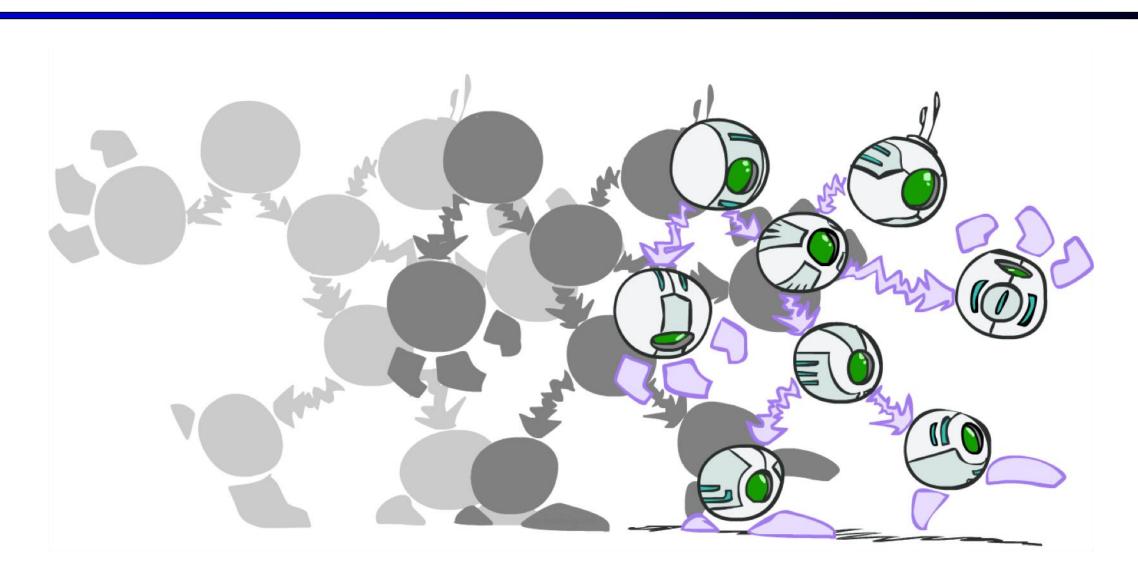


$W_{t-1}$	$P(W_t W_{t-1})$			
	sun	rain		
sun	0.9	0.1		
rain	0.3	0.7		

$W_{t}$	P(U <sub>t</sub>  W <sub>t</sub> )				
	true	false			
sun	0.2	0.8			
rain	0.9	0.1			

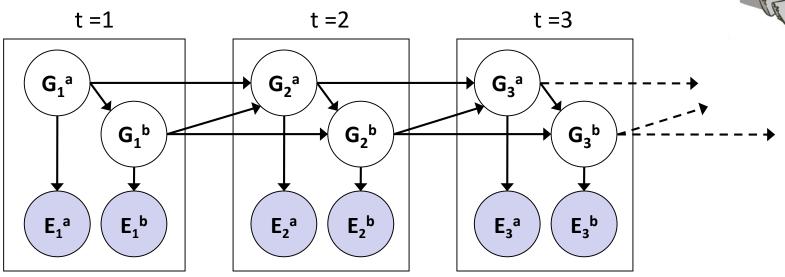
- $\mathbf{m}_{1:t+1} = P(e_{t+1}|X_{t+1}) \max_{X_t} P(X_{t+1}|X_t) \mathbf{m}_{1:t}[X_t]$
- Time complexity: O(|X|<sup>2</sup> T)
- Space complexity: O(|X| T)

# **Dynamic Bayes Nets**



# Dynamic Bayes Nets (DBNs)

- We want to track multiple variables over time, using multiple sources of evidence
- Idea: Repeat a fixed Bayes net structure at each time
- Variables from time t can condition on those from t-1



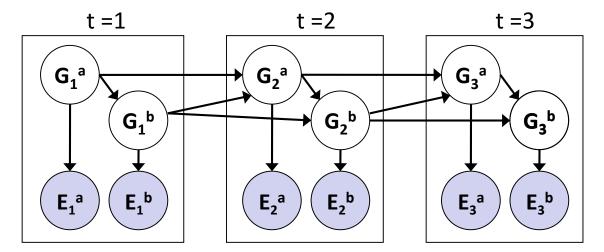


### **DBNs** and HMMs

- Every HMM is a DBN
- Every discrete DBN can be represented by a HMM
  - Each HMM state is Cartesian product of DBN state variables
    - E.g., 3 binary state variables => one state variable with 2<sup>3</sup> possible values
  - Advantage of DBN vs. HMM?
    - Sparse dependencies => exponentially fewer parameters
    - E.g., 20 binary state variables, 2 parents each; DBN has  $20 \times 2^{2+1} = 160$  parameters, HMM has  $2^{20} \times 2^{20} = 10^{12}$  parameters

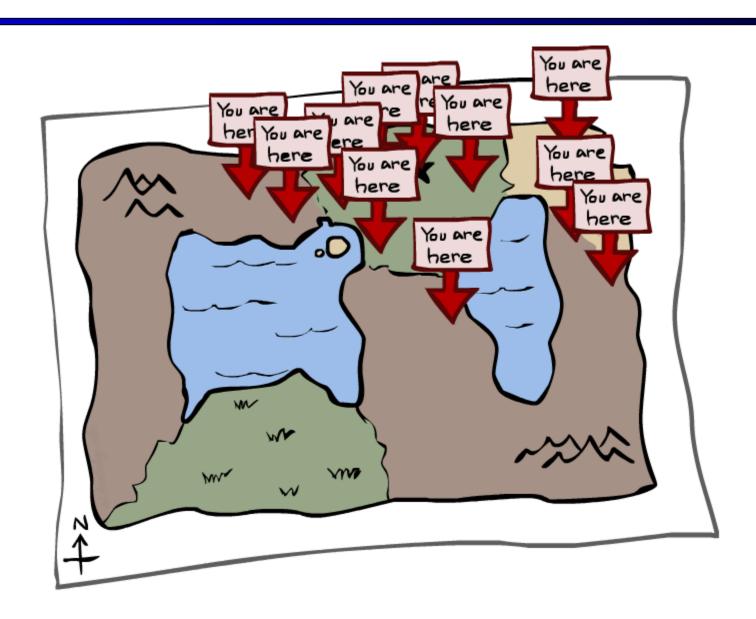
### **Exact Inference in DBNs**

- Variable elimination applies to dynamic Bayes nets
- Offline: "unroll" the network for T time steps, then eliminate variables to find  $P(X_T | e_{1:T})$ 
  - Problem: results in very large BN



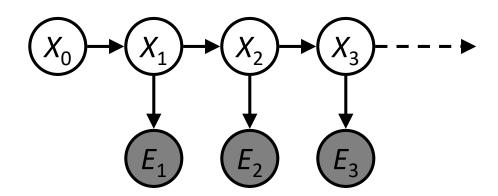
- Can we do better?
  - Do we need to unroll for many steps? What is the best variable order of elimination?
- Online: unroll as we go, eliminate all variables from the previous time step
  - A generalization of the Forward algorithm

# Particle Filtering



### Large state space

- When |X| is huge (e.g., position in a building), exact inference becomes infeasible
- Can we use approximate inference, e.g., likelihood weighting?
  - Evidences are "downstream"
  - By ignoring the evidence: with more states sampled over time, the weight drops quickly (going into low-probability region)
  - Hence: too few "reasonable" samples



## Particle Filtering

- Represent belief state at each step by a set of samples
  - Samples are called particles
- Our representation of P(X) is now a list of N particles (samples)
  - P(x) approximated by number of particles with value x
    - So, many x may have P(x) = 0
  - Generally, N << |X|</li>
    - More particles, more accuracy; but a large N would defeat the point.

0.0	0.1	0.0
0.0	0.0	0.2
0.0	0.2	0.5

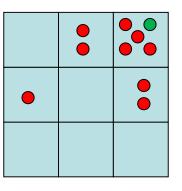


	• •

### Representation: Particles

#### Initialization

- sample N particles from the initial distribution  $P(X_0)$
- All particles have a weight of 1



#### Particles:

(3,3)

(2,3)

(3,3)

(3,2)

(3,3)

(3,2)

(1,2)

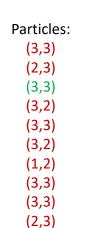
(3,3)

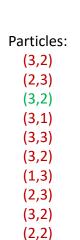
(3,3)

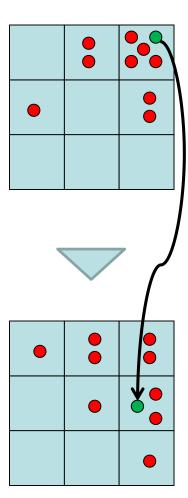
(2,3)

## Particle Filtering: Propagate forward

- Each particle is moved by sampling its next position from the transition model:
  - $x_{t+1} \sim P(X_{t+1} \mid x_t)$
- This captures the passage of time
  - If enough samples, close to exact probabilities (consistent)







### Particle Filtering: Observe

- Similar to likelihood weighting, weight samples based on the evidence
  - $W = P(e_t | X_t)$
  - Particles that fit the evidence better get higher weights, others get lower weights
- What happens if we repeat the Propagate-Observe procedure over time?
  - It is exactly likelihood weighting (if we multiply the weights)
  - Weights drop quickly...

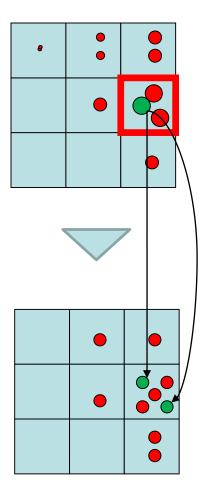
Particles:				
(3,2)				
(2,3)		0		•
(3,2)	)			
(3,1)				
(3,3)				
(3,2) (1,3)				
(2,3)				
(3,2)				
(2,2)				
Particles:			7	
(3,2) w=.9				
(2,3) w=.2	•	•		
(3,2) w=.9				
(3,1) w=.4 (3,3) w=.4			y	
(3,2) w=.9		•		
(1,3) w=.1			L	
(2,3) w=.2				

## Particle Filtering: Resample

- Rather than tracking weighted samples, we *resample*
  - Generate N new samples from our weighted samples
  - Each new sample is selected from the current population of samples; the probability is proportional to its weight.
  - The new samples have weight of 1
- Now the update is complete for this time step, continue with the next one

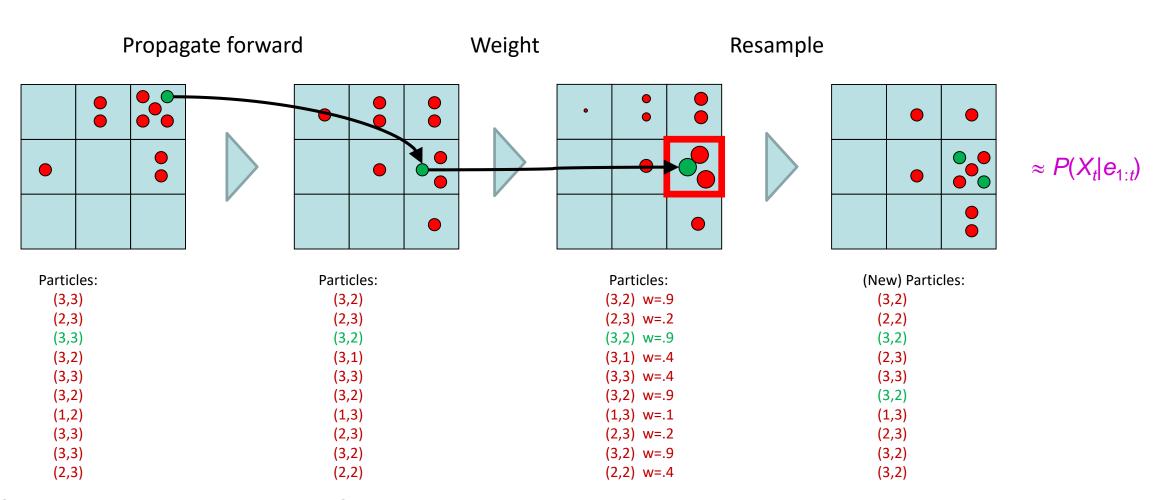
Particles:
(3,2) w=.9
(2,3) w=.2
(3,2) w=.9
(3,1) w=.4
(3,3) w=.4
(3,2) w=.9
(1,3) w=.1
(2,3) w=.2
(3,2) w=.9
(2,2) w=.4

(New) Particles:
(3,2)
(2,2)
(3,2)
(2,3)
(3,3)
(3,2)
(1,3)
(2,3)
(3,2)
(3,2)
(3,2)



## Summary: Particle Filtering

Particles: track samples of states rather than an explicit distribution

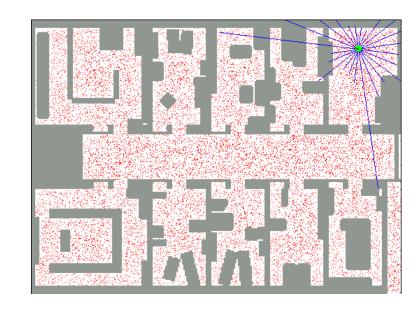


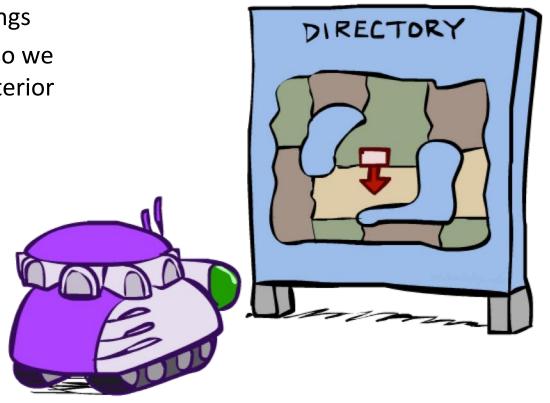
Consistency: see proof in AIMA Ch. 15

### **Robot Localization**

#### In robot localization:

- We know the map, but not the robot's position
- Observations may be vectors of range finder readings
- State space and readings are typically continuous so we cannot usually represent or compute an exact posterior
- Particle filtering is a main technique



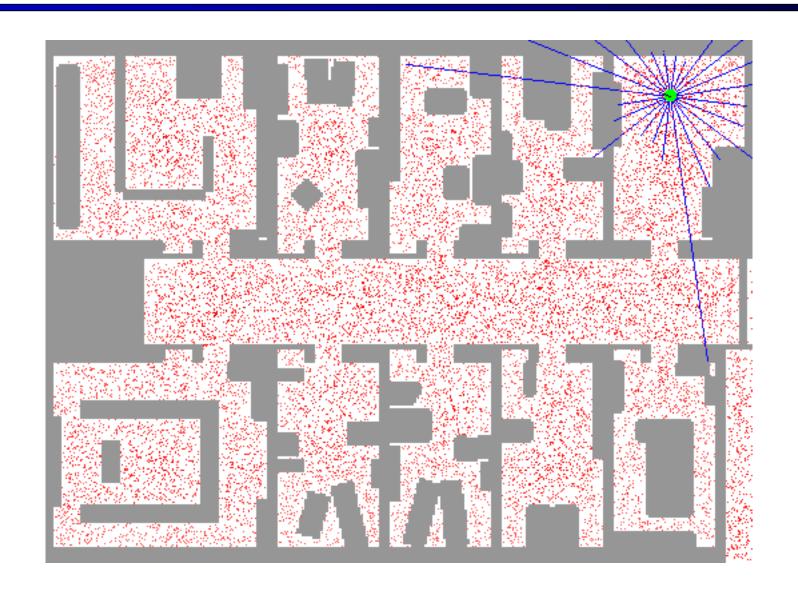


# Particle Filter Localization (Sonar)



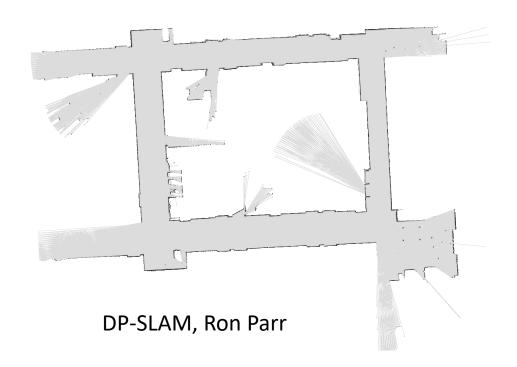
[Dieter Fox, et al.]

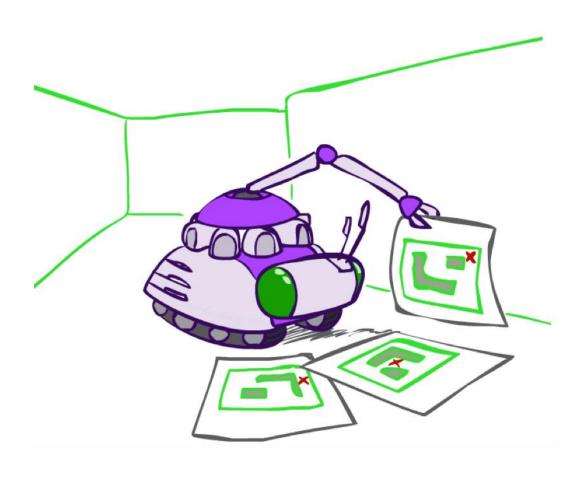
# Particle Filter Localization (Laser)



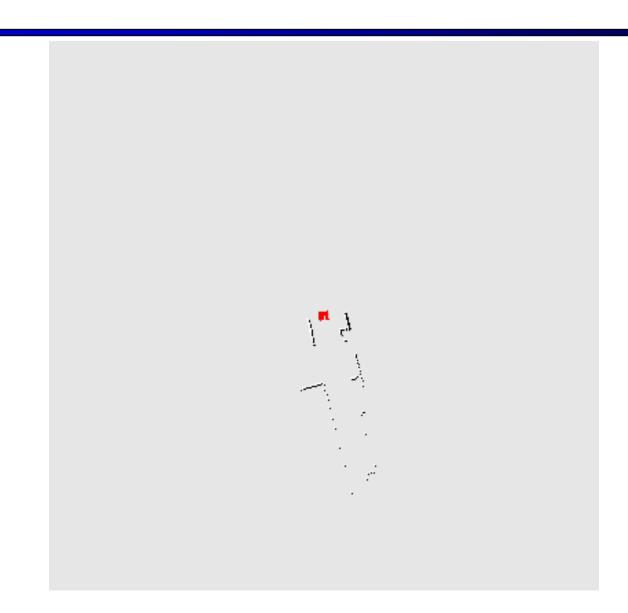
# Robot Mapping

- SLAM: Simultaneous Localization And Mapping
  - We do not know the map or our location
  - State consists of position AND map!
  - Main techniques: Kalman filtering (Gaussian HMMs) and particle methods



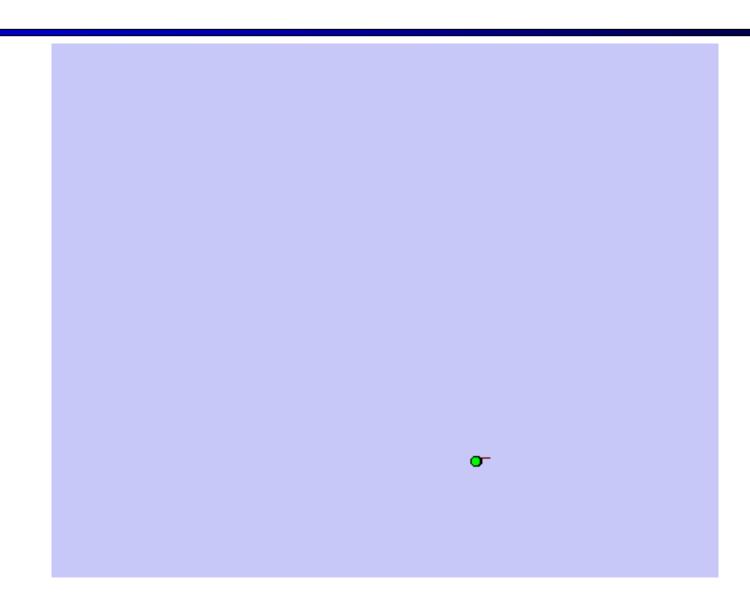


### Particle Filter SLAM – Video 1



[Sebastian Thrun, et al.]

### Particle Filter SLAM – Video 2



### Summary

- Probabilistic temporal models
  - Markov model
  - Hidden Markov model
    - Filtering: forward algorithm
    - MLE: Viterbi algorithm
  - Dynamic Bayesian network
  - Approximate inference by particle filtering

