# Probability Theory I Assignment 3

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## Exercise 1

We are proving (i)  $\Rightarrow$  (ii)  $\Rightarrow$  (ii)  $\Rightarrow$  (i) and (ii)  $\Rightarrow$  (i).

(i)  $\Rightarrow$  (ii) Suppose events  $\{A_{\alpha}\}_{\alpha\in I}$  are independent. Note that  $\sigma(A_{\alpha}) := \{A_{\alpha}, A_{\alpha}^{c}, \varnothing, \Omega\}$ . Also, for any set  $A, A \cap \varnothing = \varnothing$  and  $A \cap \Omega = A$  so the property of independence works on these types of sets. Moreover, for  $\alpha_{1}, \ldots, \alpha_{n} \in I$   $P(\bigcap_{i=1}^{n} A_{\alpha_{i}}) = \prod_{i=1}^{n} P(A_{\alpha_{i}})$  by hypothesis. Therefore, to show that our  $\{\sigma(A_{\alpha})\}_{\alpha\in I}$  are independent, it suffices to show that for  $A_{1}, A_{2}$  in our events collection,  $P(A_{1} \cap A_{2}^{c}) = P(A_{1})P(A_{2}^{c})$ . Will follow that for any  $\alpha_{0}, \ldots, \alpha_{n} \in I$ ,  $P(\bigcap_{j=1}^{n} A_{\alpha_{j}} \cap A_{\alpha_{0}}^{c}) = \prod_{j=1}^{n} P(A_{\alpha_{j}})P(A_{\alpha_{0}}^{c})$ .

$$P(A_1 \cap A_2^c) = P(A_1 \setminus (A_2 \cap A_1))$$
  
=  $P(A_1) - P(A_2 \cap A_1)$  by independence of events  
=  $P(A_1) - P(A_2)P(A_1)$   
=  $P(A_1)(1 - P(A_2)) = P(A_1)P(A_2^c)$  as required.

- (ii)  $\Rightarrow$  (i) Is straightforward using that  $\sigma(A_{\alpha}) := \{A_{\alpha}, A_{\alpha}^{c}, \emptyset, \Omega\}.$
- (ii)  $\Rightarrow$  (iii) Suppose  $\sigma(\{A_{\alpha}\})_{\alpha \in I}$  is independent. Let  $\alpha_1, \ldots, \alpha_n \in I$ , note the following, for  $u_i \in \mathbb{R}$

$$\{\mathbf{1}_{A_{\alpha_i}} = u_i\} = \begin{cases} A_{\alpha_i} & \text{if } u_i = 1\\ A_{\alpha_i}^c & \text{if } u_i = 0\\ \varnothing & \text{otherwise} \end{cases}$$

So for any  $u_i \in \mathbb{R}$ ,  $\mathbf{1}_{A_{\alpha_i}}^{-1}(u_i) \in \sigma(A_{\alpha_i})$ . By assumption, since  $\sigma(\{A_{\alpha}\})_{\alpha \in I}$  is independent, the random variables  $(\mathbf{1}_{A_{\alpha_i}})_{i \in I}$  are independent.

(iii)  $\Rightarrow$  (i) Assume events  $\{\mathbf{1}_{A_{\alpha}}\}_{\alpha\in I}$  are independent. Note that  $\mathbf{1}_{A_{\alpha}}$  is in  $L^1$  since  $1\in L^1$  and  $0\leq \mathbf{1}_{A_{\alpha}}\leq 1$ , for  $\alpha_1,\ldots,\alpha_n\in I$ 

$$\begin{split} P(\bigcap_{i=1}^n A_{\alpha_i}) &= E(\mathbf{1}_{\bigcap_{i=1}^n A_{\alpha_i}}) = E(\mathbf{1}_{A\alpha_1} \cdots \mathbf{1}_{A_{\alpha_n}}) \\ &= E(\mathbf{1}_{A_{\alpha_1}}) \cdots E(\mathbf{1}_{A_{\alpha_n}}) \text{ using independence of } \mathbf{1}_{A_{\alpha_i}} \\ &= P(A_{\alpha_1}) \cdots P(A_{\alpha_n}) \\ &= \prod_{i=1}^n P(A_{\alpha_i}) \text{ as required} \end{split}$$

## Exercise 2

 $(\Rightarrow)$  Suppose  $\sigma(Y) \subset \sigma(X)$ .

#### — Step 1

Let's consider as the hint suggests that we take Y to be an indicator. Then, it follows that  $Y = \mathbf{1}_A$  where  $A \in \sigma(X)$ , hence there exists  $B \in \mathcal{B}(\mathbb{R})$  such that  $A = X^{-1}(B)$ . Then, for any  $\omega$ ,

$$Y(\omega) = \mathbf{1}_A(\omega) = \mathbf{1}_{X^{-1}(B)}(\omega) = \mathbf{1}_B(X(\omega))$$

Hence  $Y = \mathbf{1}_B \circ X$  where  $\mathbf{1}_B$  is Borel by construction.

#### — Step 2

Now let's take Y to be a simple function, we can write  $Y = \sum_{k=1}^{n} a_k \mathbf{1}_{A_k}$  where  $(A_k)$  events on  $(\Omega, \mathcal{F})$ , and  $(a_k) \in \mathbb{R}^n$ , by the previous construction, can see that there exists for each  $A_k$  a  $B_k \in \mathcal{B}(\mathbb{R})$  such that  $A_k = X^{-1}(B_k)$ . And we see that for  $\omega \in \Omega$ , setting  $h = \sum_{k=1}^{n} a_k \mathbf{1}_{B_k}$  which is Borel,

$$Y(\omega) = \sum_{k=1}^{n} a_k \mathbf{1}_{A_k}(\omega) = \sum_{k=1}^{n} a_k \mathbf{1}_{X^{-1}(B_k)}(\omega) = \sum_{k=1}^{n} a_k \mathbf{1}_{B_k}(X(\omega)) = h(X(\omega))$$

So  $Y = h \circ X$  for Y simple function.

#### — Step 3

Now we can take any Y real valued non negative random variable, by simple approximation theorem we can find a sequence of simple functions  $Y_n \uparrow Y$  as taken as before and denoting each Borel function  $h_n$  for  $Y_n$ , so  $Y_n = h_n \circ X$ . And by taking  $Y_n$  increasing, we have  $h_n$  increasing and Borel. Take  $h = \limsup_n h_n$  so h is also Borel. For any  $\omega \in \Omega$ ,

$$Y(\omega) = \lim_{n} Y_n(\omega) = \lim_{n} h_n(X(\omega)) = h(X(\omega))$$

Hence,  $Y = h \circ X$ 

#### — Step 4

Now taking any real valued Y, we can write  $Y = Y^+ - Y^-$  where  $Y^+$  and  $Y^-$  are non negative. So we can write  $Y^+ = h_+ \circ X$  and  $Y^- = h_- \circ X$  and setting  $h = h_+ - h_-$ ,  $Y = h \circ X$  as required.

 $(\Leftarrow)$  Now suppose that  $Y = h \circ X$  with h Borel. Let's show that  $\sigma(Y) \subset \sigma(X)$ . Take  $A \in \sigma(Y)$ , there exists  $B \in \mathcal{B}(\mathbb{R})$  such that  $A = Y^{-1}(B) = (h \circ X)^{-1}(B) = X^{-1}(h^{-1}(B))$  since h is Borel,  $h^{-1}(B) \in \mathcal{B}(\mathbb{R})$  as well, and setting  $B' = h^{-1}(B)$ ,  $A = X^{-1}(B')$  hence  $A \in \sigma(X)$  as required.

# Exercise 3

Let  $X: \Omega \to [0, \infty)$  be a random variable with E(X) = 1 on  $(\Omega, \mathcal{F}, P)$  and the mapping  $Q: \mathcal{F} \to [0, 1]$  defined by  $Q(A) := \int_A X dP = E(X \mathbf{1}_A)$  for  $A \in \mathcal{F}$ .

Let's show it is a probability measure.

1. 
$$Q(\emptyset) = E(X\mathbf{1}_{\emptyset}) = E(0) = 0$$

2. For  $(A_n)_{n\in\mathbb{N}}$  pairwise disjoint sets,

$$\begin{split} Q(\bigcup_{n\in\mathbb{N}}A_n) &= E(X\mathbf{1}_{\bigcup_{n\in\mathbb{N}}A_n}) = E(X(\sum_{n\in\mathbb{N}}\mathbf{1}_{A_n})) \\ &= E(\sum_{n\geq 1}X\mathbf{1}_{A_n}) \text{ then using exercise 5} \\ &= \sum_{n\geq 1}E(X\mathbf{1}_{A_n}) \\ &= \sum_{n\geq 1}Q(A_n) \end{split}$$

3. 
$$Q(\Omega) = E(X) = 1$$

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Hence Q is a probability measure.

Now let  $A \in \mathcal{F}$  such that P(A) = 0. Since X is non negative, there exists a sequence of simple functions  $X_n$  such that  $X_n \uparrow X$  and  $X_n = \sum_{i=1}^n a_i \mathbf{1}_{A_i}$ 

$$Q(A) = \int_{A} X dP = \int_{A} \lim_{n} X_{n} dP \underset{MCT}{=} \lim_{n} \int \sum_{i=1}^{n} a_{i} \mathbf{1}_{A_{i} \cap A} dP = \lim_{n} \sum_{i=1}^{n} a_{i} P(A_{i} \cap A) \underset{P(A)=0}{=} 0$$

Now let Y be a non negative random variable. Using our simple approximation theorem, have  $Y_n \uparrow Y$  sequence of simple functions such that  $Y_n = \sum_{i=1}^n b_i \mathbf{1}_{B_i}$  and  $Y = \lim_n Y_n$ . Now define as usual

$$\int_{\Omega} Y dQ = \lim_{n} \sum_{i=1}^{n} b_i Q(B_i)$$

Using the observation that X is non negative and the simple sequence used previously, for  $i = 1, \ldots, n$ 

$$Q(B_i) = \int_{B_i} XdP = \lim_n \sum_{j=1}^n a_j P(A_j \cap B_i)$$

$$\int_{\Omega} YdQ = \lim_n \sum_{i=1}^n b_i \left(\lim_n \sum_{j=1}^n a_j P(A_j \cap B_i)\right)$$

$$= \lim_n \sum_{i=1}^n b_i \left(\int_{B_i} \lim_n \sum_{j=1}^n a_j \mathbf{1}_{A_j} dP\right) \text{ by MCT}$$

$$= \lim_n \sum_{i=1}^n b_i \left(\int_{B_i} \lim_n X_n dP\right) = \lim_n \sum_{i=1}^n b_i \int_{\Omega} \mathbf{1}_{B_i} XdP$$

$$= \int_{\Omega} X \lim_n \sum_{i=1}^n b_i \mathbf{1}_{B_i} dP = \int_{\Omega} X \lim_n Y_n dP \text{ using MCT}$$

$$= \int_{\Omega} XYdP \text{ as required}$$

## Exercise 4

Take  $(\Omega, \mathcal{F})$  to be  $(\mathbb{R}, \mathcal{B}(\mathbb{R}))$  and let  $\mu$  the Lebesgue measure. Let  $X_n(\omega) := \frac{1}{n} \mathbf{1}_{[0,n]}(\omega)$  for any  $\omega \in \Omega$ . Note that  $X_n(\omega) \to 0$  for all  $\omega \in \Omega$  and  $E(X_n) = \frac{1}{n} \mu([0,n]) = 1 < \infty$ . Now, see that  $E(X_n) = \int_{\mathbb{R}} X_n d\mu = 1 \to 1 \neq 0$ .

## Exercise 5

- (a) Let  $(X_n)_{n\in\mathbb{N}}$  be a sequence of non negative rv. Note that  $(\sum_{k=1}^n X_k)_{n\in\mathbb{N}}$  is a non negative non decreasing sequence of random variables. Also see that  $(\sum_{k=1}^n X_k) \uparrow X =: \sum_{n\geq 1} X_n$ . Using monotone convergence,  $E(\sum_{n\geq 1} X_n) = \sum_{n\geq 1} E(X_n)$ .
  - (b) Now assume X takes values in N. Note that  $X_n := \mathbf{1}(X \ge n)$  is a non negative sequence of random

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variables and  $X = \sum_{n \ge 1} \mathbf{1}(X \ge n)$ . Using (a),

$$E(X) = E(\sum_{n\geq 1} \mathbf{1}(X \geq n))$$
$$= \sum_{n\geq 1} E(\mathbf{1}(X \geq n))$$
$$= \sum_{n\geq 1} P(X \geq n)$$

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