

★ Ex 1.1. Uniform distribution. $X \sim U(a, b)$ ★ Ex 1.3

$$f(x) = \begin{cases} \frac{1}{b-a} & x \in [a, b] \\ 0 & \text{otherwise} \end{cases}$$

$$F(x) = \begin{cases} 0 & x \leq a \\ \frac{x-a}{b-a} & a < x < b \\ 1 & x \geq b \end{cases}$$

$$F_{1000}(x) = \frac{1}{1000} \sum_{k=1}^{1000} I(X_k \leq x)$$

$$= \begin{cases} \frac{1}{1000} \times 0 = 0 & x \in (-\infty, 0) \\ \frac{1}{1000} \times 65 = 0.065 & x \in [0, 1) \\ \frac{1}{1000} \times (65 + 246) = 0.311 & x \in [1, 2) \\ \frac{1}{1000} \times (65 + 246 + 358) = 0.669 & x \in [2, 3) \\ \frac{1}{1000} \times (65 + 246 + 358 + 272) = 0.941 & x \in [3, 4) \\ \frac{1}{1000} \times (65 + 246 + 358 + 272 + 59) = 1 & x \in [4, \infty) \end{cases}$$

Binomial $(4, \frac{1}{2})$. $n=4$, $p=\frac{1}{2}$

P.m.f. $f(x) = P(X=x) = \binom{n}{x} p^x (1-p)^{n-x}$

$$f(x) = \begin{cases} 1/16 = 0.0625 & x=0 \\ 4/16 = 0.25 & x=1 \\ 6/16 = 0.375 & x=2 \\ 4/16 = 0.25 & x=3 \\ 1/16 = 0.0625 & x=4 \end{cases}$$

$$F(x) = \begin{cases} 0 & (-\infty, 0) \\ 0.0625 & [0, 1) \\ 0.3125 & [1, 2) \\ 0.6875 & [2, 3) \\ 0.9375 & [3, 4) \\ 1 & [4, \infty) \end{cases}$$

$$f_{1000}(x) = \frac{1}{1000} \sum_{k=1}^{1000} I(X_k = x)$$

★ property 1.3.

$X \sim U(0, 1)$. c.d.f. $F_X(x) = \begin{cases} 0 & x \leq 0 \\ x & 0 < x < 1 \\ 1 & x \geq 1 \end{cases}$

$$\begin{aligned} P(Y \leq y) &= P(F^{-1}(X) \leq y) \\ &= P(F(F^{-1}(X)) \leq F(y)) \\ &= P(X \leq F(y)) = F_X(F(y)) \\ &= F(y). \end{aligned}$$

~~$Y \sim U(0, 1)$~~

★ Property 1.7.

$$(i) E(ax+b) = \sum_{x \in S} (ax+b) f(x) = a \sum_{x \in S} x f(x) + b$$

$$= a E(X) + b.$$

(ii) If $P(X=b)=1$, then

$$E(X) = \sum_{x \in S} x \cdot f(x) = b \cdot 1 = b$$

(iii) $P(a < X \leq b) = 1$.

$$a \leq E(X) = \sum_{x \in S} x \cdot f(x) \leq \sum_{x \in S} b \cdot f(x) = b$$

$$(iv) E[g(x) + h(x)] = \sum_{x \in S} (g(x) + h(x)) f(x) = E[g(x)] + E[h(x)]$$

★ Property 1.8. If $X \geq 0$ takes integer values,

$$\text{then } E(X) = \sum_{n=1}^{\infty} P(X \geq n) = \sum_{n=2}^{\infty} P(X \geq n).$$

proof. ① If $P(X = +\infty) > 0$, then $E(X) = +\infty$

$$P(X \geq n) \geq P(X = +\infty) > 0.$$

$$\text{So } \sum_{n=2}^{\infty} P(X \geq n) \rightarrow +\infty. \text{ Equality Holds!}$$

② If $P(X = +\infty) = 0$, then

$$\sum_{n=2}^{\infty} P(X \geq n) = \sum_{n=2}^{\infty} \sum_{j=n}^{\infty} P(X=j)$$

$$= \sum_{j=2}^{\infty} \sum_{n=2}^{j-1} P(X=j) = \sum_{j=2}^{\infty} j P(X=j) = E(X)$$

★ Property 1.9

(iii) For $a > 0$, $P(h(X) \geq a) \leq E\left(\frac{h(X)}{a}\right)$

$$\text{proof. } E\left(\frac{h(X)}{a}\right) = \frac{1}{a} E(h(X))$$

$$= \frac{1}{a} \left\{ \int_{x \in R} h(x) f(x) dx \right\}$$

$$\geq \frac{1}{a} \left\{ \int_{h(x) \geq a} a \cdot f(x) dx \right\}$$

$$= \int_{h(x) \geq a} f(x) dx$$

$$= P(h(X) \geq a)$$

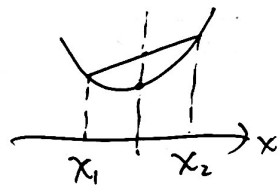
★ Property 1.9.

(iv) If g is convex, then $g(E(x)) \leq E(g(x))$

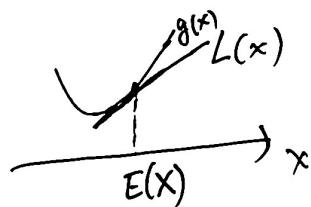
proof: Δ convex function

$$g(tx_1 + (1-t)x_2) \leq tg(x_1) + (1-t)g(x_2)$$

for $(0 \leq t \leq 1)$.



Δ Define $L(x) = a + bx$ is the linear function tangential to $g(x)$ at the point $E(x)$.



So $g(x) \geq L(x)$ for all x .

$$\begin{aligned} \text{So } E[g(x)] &\geq E[L(x)] = E(a + bx) \\ &= a + bE(x) = L(E(x)) \end{aligned}$$

$$= g(E(x))$$

★ Property 1.10.

X non-negative

$$E(x) = \int_0^{\infty} x \cdot f(x) dx = \int_0^{\infty} (1 - F(x)) dx$$

proof. $1 - F(x) = 1 - P(X \leq x) = P(X > x)$

$$= \int_x^{\infty} f(t) dt$$

$$\int_0^{\infty} (1 - F(x)) dx = \int_0^{\infty} \int_x^{\infty} f(t) dt dx$$

\uparrow
 $t \geq x$

$$= \int_0^{\infty} \left(\int_0^t 1 dx \right) f(t) dt$$

$$= \int_0^{\infty} t f(t) dt = E(x)$$

★ Property 1.11.

(i) If X is discrete,

$$E(X^2) = 0 \Rightarrow P(X=0) = 1.$$

If $P(X=0) < 1$, then $P(X \neq 0) > 0$.

$$E(X^2) = \sum_{x \neq 0} x^2 P(X=x^2) > 0. \text{ contradicts!}$$

② If X is continuous.

$$0 = E(X^2) = \int_{-\infty}^{\infty} x^2 f(x) dx$$

$$\geq \int_{\varepsilon}^{\infty} x^2 f(x) dx + \int_{-\infty}^{-\varepsilon} x^2 f(x) dx \quad (\varepsilon > 0)$$

$$\geq \varepsilon^2 \left[\int_{\varepsilon}^{\infty} f(x) dx + \int_{-\infty}^{-\varepsilon} f(x) dx \right]$$

$$= \varepsilon^2 [1 - F(\varepsilon) + F(-\varepsilon)]$$

$$= \varepsilon^2 P(X \geq \varepsilon \text{ or } X \leq -\varepsilon) \geq 0$$

So for any $\varepsilon > 0$, $P(|X| \geq \varepsilon) = 0$.

$$\Rightarrow P(X=0) = 1$$

$$(ii) \text{Cov}(ax+b, cY+d) \neq ac \text{Cov}(X, Y)$$

$$(iii) \text{Var}(X+Y) = \text{Var}(X) + \text{Var}(Y) + 2\text{Cov}(X, Y)$$

$$(iv) -1 \leq \rho(X, Y) \leq 1.$$

$$\text{proof. } \rho^2(X, Y) = \frac{\text{Cov}^2(X, Y)}{\text{Var}(X) \text{Var}(Y)}$$

$$\text{Cov}^2(X, Y) = \left\{ E[(X - E(X))(Y - E(Y))] \right\}^2$$

$$\leq E[X - E(X)]^2 E[Y - E(Y)]^2 \quad (\text{property 1.12})$$

$$= \text{Var}(X) \cdot \text{Var}(Y)$$

$$(vi) |P(X, Y)| = 1 \Leftrightarrow P(X = aY + b) = 1.$$

$$\text{proof. } |P(X, Y)| = 1 \Leftrightarrow \text{property 1.12}$$

$$\text{let } X' \equiv X - E(X) \quad Y' \equiv Y - E(Y)$$

$$P(X' = Y' \frac{E(X'Y')}{E(Y'^2)}) = 1$$

$$\Leftrightarrow P(X = aY + b) = 1.$$

★ Property 1.12. $E(XY) \leq \sqrt{E(X^2) E(Y^2)}$.

proof. $0 \leq E[X \cdot E(Y^2) - Y E(XY)]^2$

$$= E\{X^2 [E(Y^2)]^2 + Y^2 E(XY)^2 - 2XY E(Y^2) E(XY)\}$$

$$= E(Y^2) \{E(X^2) E(Y^2) + [E(XY)]^2 - 2[E(XY)]^2\}$$

$$= E(Y^2) \{E(X^2) E(Y^2) - [E(XY)]^2\}$$

$$\Rightarrow E(X^2) E(Y^2) \geq [E(XY)]^2.$$

"=" Holds $\Leftrightarrow P(\textcircled{\times} X E(Y^2) - Y E(XY) = 0) = 1$

$$\Leftrightarrow P\left(X - Y \cdot \frac{E(XY)}{E(Y^2)} = 0\right) = 1.$$

★ Property 1.11.

(vii) $\rho(ax+b, cY+d) = \text{sgn}(ac) \rho(X, Y)$

proof. $\rho(ax+b, cY+d)$

$$= \frac{\text{Cov}(ax+b, cY+d)}{\sqrt{\text{Var}(ax+b) \text{Var}(cY+d)}}$$

$$= \frac{ac \text{Cov}(X, Y)}{\sqrt{\text{Var}(X) \text{Var}(Y) \cdot |ac|}}$$

$$= \text{sgn}(ac) \rho(X, Y)$$

(viii) If X & Y are independent,

then $\text{Cov}(X, Y) = 0$, then $\rho(X, Y) = 0$.