## Problem 1 (3 points)

Suppose we have two functions:  $f(n) = n^{\sqrt{n}}, g(n) = 2^{n^{2/3}}$ . Choose the correct option(s).

Explain why.

(a) 
$$f(n) = o(g(n))$$
, (b)  $f(n) = \omega(g(n))$ , (c)  $f(n) = \Theta(g(n))$ , (d)  $f(n) = O(g(n))$ ,

(e) 
$$f(n) = \Omega(g(n))$$
.

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(a) 
$$f(n) = o(g(n))$$
 and (d)  $f(n) = O(g(n))$  are both correct.

From a theorem in class, to show 
$$f(n) = c(g(n))$$
, it is sufficient to show  $\log |f(n)| = o(\log(g(n)))$ . Observe

$$\frac{\log f(n)}{\log g(n)} = \frac{\log n^{\sqrt{n}}}{\log 2^{n^{2/3}}} = \frac{\sqrt{\log n}}{\sqrt{2}}$$

$$= \frac{n^{1/2} \log n}{\sqrt{2} n^{1/6}} = \frac{\log n}{\sqrt{6}}$$

$$\lim_{n\to\infty} \frac{|cg(f(n))|}{|cg(g(n))|} = \lim_{n\to\infty} \frac{|cg(g(n))|}{|cg(g(n))|} = \lim_{n\to\infty} \frac{|cg(g(n))|}{|cg(g($$

# Problem 2 (3 points)

What is the reminder when you divide  $3^{132}$  by 26?

Hence 
$$3^{132} \mod 26 = (3^3)^{44} \mod 26$$
  
 $= (3^3 \mod 26)^{44}$   
 $= (1 \mod 26)^{44}$   
 $= 1 \mod 26$ 

So the remainder is 1.

### Problem 3 (4 points)

Let  $a_n$  be a sequence defined by  $a_1 = 1$ ,  $a_2 = 8$ ,  $a_n = a_{n-1} + 2a_{n-2}$   $(n \ge 3)$ . Prove that  $a_n = 3 \cdot 2^{n-1} + 2(-1)^n$  for all  $n \ge 1$ .

Let 
$$P(n)$$
 be the predicate that

 $a_n = 3 \cdot 2^{n-1} + 2(-1)^n$ . We first prove

the base cases  $n=1$  and  $n=2$ .

We have

$$P(1)$$
:  $3-2^{1-1}+2(-1)^1=3-2=1=a_1$ .  
 $P(2)$ :  $3-2^{2-1}+2(-1)^2=3-2+2=8=a_2$ .  
So  $P(1)$  and  $P(2)$  are time.

Now, for n=3 we assume P(n) and P(n-1) are true a will show P(n+1) is true. We then have

$$a_{n+1} = a_n + 2a_{n-1}$$

$$= 3 \cdot 2^{n-1} + 2(-1)^n + 2(3 \cdot 2^{n-2} + 2(-1)^{n-1})$$

$$= 3 \cdot 2^{n-1} + 2(-1)^n + 3 \cdot 2^{n-1} + 2(-1)^{n-1} + 2(-1)^{n-1}$$

$$= 2(3 \cdot 2^{n-1}) + 2((-1)^n + (-1)^{n-1}) + 2(-1)^{n-1}$$

$$= 3 \cdot 2^n + 2(-1)^{n-1}$$

$$= 3 \cdot 2^n + 2(-1)^{n+1}$$

There fore P(n+1) is true. This completes the proof by induction

#### Problem 4 (5 points)

Consider a random ordering (permutation) of the numbers 1, 2, ..., n (each of the n! orderings are equally likely). A pair i, j is out of order if i < j but i occurs after j in the random ordering.

- (i) Given a pair i, j (such that i < j), what is the probability that i, j is out of order? (2 points)
- (ii) Suppose  $E_{ij}$  is the event that pair i, j is out of order. Are the events  $\{E_{ij}|1 \le i < j \le n\}$  mutually independent? (You only need to give a short justification). (1 point)
- (iii) What is the expected number of pairs that are out of order? Suppose  $p^*$  is the answer for part (i). The answer can be expressed in term of  $p^*$ . (2 points)

  Hint: Define appropriate Bernoulli random variables  $X_{ij}$  for each pair i, j, and express the random variable X that captures the number of pairs out of order in terms of  $\{X_{ij}\}$ .

(i) The probability is 2.

(ii) No they are not mutually independent.

For example, if  $E_{12}$ ,  $E_{23}$ , ...,  $E_{n-1,n}$  all occur, then the permutation must be 1/2,3,...,n-1, n hence all other

Ei occur in this case

(iii) Define  $X_{ij} = \begin{cases} 1 & \text{if } E_{ij} \text{ occurs} \\ 0 & \text{otherwise} \end{cases}$ 

Then  $X = \sum_{1 \leq i < j \leq n} X_{ij}$ .

It follows that

$$\mathbb{E}[X] = \mathbb{E}\left[\sum_{1 \leq i < j \leq n} X_{ij}\right] = \sum_{1 \leq i < j \leq n} \mathbb{E}[X_{ij}]$$

$$= \sum_{1 \leq i \leq j \leq n} 1 \cdot Pr(X_{ij} = 1) + 0 \cdot Pr(X_{ij} = 0)$$

$$= \sum_{1 \leq i < j \leq n} 1 \cdot \frac{1}{2} = \sum_{1 \leq i < j \leq n} \frac{1}{2}$$

Note that the total # of pairs
$$1 \le i \le j \le n \text{ is } \binom{n}{2} = \frac{n(n-1)}{2}$$

Thus 
$$\#(X) = \frac{1}{2}$$
,  $\frac{n(n-1)}{2} = \frac{n(n-1)}{4}$ 

#### Problem 5 (4 points)

Let A and B be two events with  $A \subset B$  and 0 < P(A) < P(B) < 1. Which of the following are true statements. You **do not** need to justify your answer.

- (a)  $P(A \cup B) = P(A) + P(B)$
- (b)  $P(A \cup \overline{B}) = P(A) + P(\overline{B})$
- (c)  $P(A|\overline{B}) > P(A)$
- (d)  $P(\overline{B}|A) > P(B)$

Only (b) is true.

Note: While a justification is not needed, one is given since this is a practice final

(a) Note 
$$P(AUB) = P(A) + P(B) - P(ADB)$$
.  
However  $ADB = A$  since  $A \subseteq B$ .  
So  $P(AUB) = P(B) \neq P(A) + P(B)$  unless  $P(A) = 0$ .  
Thus a) is not true.

(b) Since 
$$A \subseteq B$$
 we have  $A \cap \overline{R} = \emptyset$   
thus  $P(A \cup \overline{B}) = P(A) + P(\overline{B}) + P(A \cap \overline{B})$   
 $= P(A) + P(\overline{B}) + P(\emptyset)$   
 $= P(A) + P(\overline{B})$ 

(c),(d) Since 
$$A \cap \overline{B} = \emptyset$$
 we have  $P(A|\overline{B}) = P(\overline{B}|A) = 0$ .  
So (c), (d) are not true

#### Problem 6 (5 points)

Consider a graph G(V, E) on n vertices. A subset  $S \subseteq V$  is called a vertex cover of G if and only if for every edge  $(u, v) \in E$ , either  $u \in S$  or  $v \in S$ , or both  $u, v \in S$  i.e., for every edge, at least one of its endpoints must be in the vertex cover S. The minimum vertex cover of G is the vertex cover of G with the fewest number of vertices.

- (i) Let  $G^c$  represent the graph complement of G. Prove that S is a vertex cover if and only if  $V \setminus S$  is a clique in  $G^c$ . (3 points)
- (ii) Let k be the size of the largest clique in  $G^c$ . Is the size of the vertex cover of smallest size n-k? If yes, prove this. If not, give a counter-example. (2 points)

# Write G=(V,E) and $G^c=(V,E^c)$ .

(i) Suppose S is a vertex cover in Vo
Assume toward a contradiction that
VIS is not a clique. Then there
exists vertices u, v EVIS such
that (u, v) \( \mathbb{E} \) a Thur (u,v) \( \mathbb{E} \) a.
However u,v \( \mathbb{E} \) since u,v \( \mathbb{E} \) \( \mathbb{N} \),
contradicting that S is a cover.
We conclude that If S is a
cover, then G is a Clique.

To prove that G is a clique implies S is a cover, we will proceed by contrapositive.

Assume that S is not a cover.

Then there exists some edge (u,v) EE
such that u,v & S. But then

u,v & V\S and (u,v) & E^c so

V\S is not a clique. This shows
that if S is not a cover, then V\S

is not a clique. We conclude that if VIS is a clique, then S is a cover.

(ii) Yes. To prove this, suppose towards a contradiction that there is a cover S of size S < n-k. Then  $V \setminus S$  is a clique in S = |V| - |S| = n-S > n - (n-k) = k. However,  $|V \setminus S| = |V| - |S| = n-S > n - (n-k) = k$ . Hence S = k has a clique of size larger than k. However this contradicts the assumption that the largest clique in S = k has size k. Thus S = k cannot have a cover of size less than n-k.

To see that S has a cover of size

n-k, let S' be a clique of k in

G, and set S=V\S'. Then

S'=V\S so it follows from part (i)

that S is a cover in G. Furthermore

[S]= |V\S'| = |V|-|S'| = n-k

#### Problem 7 (9 points)

Identify whether the follows proofs are correct, and if there are any mistakes, identify them.

 Consider a drunken and that follows the following random process to go along a straightline path from a point A to a point B. In each second, it takes one step in the forward direction with probability 4/5, and with probability 1/5 it takes three steps in the backward direction.

Claim: Pr[ The ant has covered 3n/5 steps after n seconds]  $\leq 1/3$ .

*Proof.* Let  $X_i$  be the number of steps taken by the ant in the *i*th second, and let X be the random variable representing the total number of steps taken after n seconds. From the problem description,

$$X_i = \begin{cases} 1 & \text{with probability } 4/5 \\ -3 & \text{with probability } 1/5 \end{cases}.$$

Hence, for each  $i \in [n]$ ,  $\mathrm{E}[X_i] = \frac{4}{5} - \frac{3}{5} = 1/5$ . Hence by linearity of expectation,  $\mathrm{E}[X] = \sum_{i=1}^{n} \mathrm{E}[X_i] = n/5$ . Since  $\mathrm{E}[X] \ge 0$ , by applying Markov's inequality, we have

$$\Pr[X \ge 3n/5] \le \frac{\mathrm{E}[X]}{3n/5} = \frac{n/5}{3n/5} \le 1/3.$$

Is the above proof correct? If not, identify the incorrect step(s). (3 points)

The proof is incorrect. One cannot apply Markov since X is not non-negative.

Instructor's Note: The above answer is sufficient.

However to highlight the importance
of nonnegativity to Markov, observe

then if n=1, then X=X1 and Pr(X, >3/5) = 4/5

so the claimed inequality doesn't holdo

2. There is a murder in the town of Braavos which has a population of N=40000, and anyone could be the culprit with equal probability. The forensics team has a finger-print test that for a given person says there is a match with probability 0.4 if the person is guilty, and with probability  $10^{-4}$  if the person is not guilty. The finger-print test outputs a match with a certain Mr. Jack in Evanstown. The prosecutor argues that Mr. Jack is guilty since the probability that Mr. Jack is guilty is at least 0.9975 using the following proof

*Proof.* Let I be the event that Jack is innocent, and G be the event that Jack is guilty. Let T be the event that the test returns positive, and F be the event that the test returns negative.

$$\Pr[\text{ Jack is innocent }] = \Pr[I] = \frac{\Pr[I|T]}{\Pr[I|T] + \Pr[G|T]} = \frac{10^{-4}}{0.4 + 10^{-4}} = \frac{0.0001}{0.4001} = \frac{1}{4000} \leq 0.0025$$

Hence the probability that Jack is guilty is at least 0.9975.

Is the above proof correct? If not, how what is the correct calculation for Pr[ Jack is innocent] (you don't need to simplify the numeric expressions)? (3 points)

The proof is incorrect o Baye's rule is not used correctly

3. Let M be an  $n \times n$  matrix.

Claim. We have  $\sum_{i,j=1}^{n} M(i,j)^2 = \sum_{\ell} \lambda_{\ell}^2$ , where the  $\lambda_1, \lambda_2, \ldots$  are the eigenvalues of the matrix M, and M(i,j) is the (i,j)th entry of M.

*Proof.* From the spectral theorem, we know that

$$M = \sum_{\ell=1}^{n} \lambda_{\ell} e_{\ell} e_{\ell}^{T}$$
, and  $M(i, j) = \sum_{\ell} \lambda_{\ell} e_{\ell}(i) e_{\ell}(j)$ .

Hence,

$$\begin{split} \sum_{i=1}^{n} \sum_{j=1}^{n} M(i,j)^{2} &= \sum_{i=1}^{n} \sum_{j=1}^{n} \left( \sum_{\ell} \lambda_{\ell} e_{\ell}(i) e_{\ell}(j) \right)^{2} \\ &= \sum_{i} \sum_{j} \sum_{\ell_{1}} \sum_{\ell_{2}} \lambda_{\ell_{1}} \lambda_{\ell_{2}} e_{\ell_{1}}(i) e_{\ell_{1}}(j) e_{\ell_{2}}(i) e_{\ell_{2}}(j) \\ &= \sum_{\ell_{1}} \sum_{\ell_{2}} \lambda_{\ell_{1}} \lambda_{\ell_{2}} \sum_{i} e_{\ell_{1}}(i) e_{\ell_{2}}(i) \sum_{j} e_{\ell_{1}}(j) e_{\ell_{2}}(j) \\ &= \sum_{\ell_{1}} \sum_{\ell_{2}} \lambda_{\ell_{1}} \lambda_{\ell_{2}} (\langle e_{\ell_{1}}, e_{\ell_{2}} \rangle)^{2}. \end{split}$$

But note that by the orthogonality of eigenvectors,  $\langle e_{\ell_1}, e_{\ell_2} \rangle = \sum_i e_{\ell_1}(i)e_{\ell_2}(i)$  is 0 when  $\ell_1 \neq \ell_2$  and 1 when  $\ell_1 = \ell_2$ . Hence,

$$\sum_{i=1}^{n} \sum_{j=1}^{n} M(i,j)^{2} = \sum_{\ell_{1}} \lambda_{\ell_{1}}^{2} \langle e_{\ell_{1}}, e_{\ell_{1}} \rangle^{2} = \sum_{\ell_{1}} \lambda_{\ell_{1}}^{2}.$$

The proof is not correct.

We cannot apply the spectral

theorem since M is not necessarily
symmetric.