

(discuss with roommate) .

1. Given $f(x) = \frac{1}{(2\pi)^k |\Sigma|^{1/2}} e^{-\frac{1}{2}(x-\mu)^T \Sigma^{-1}(x-\mu)}$, where $x, \mu \in \mathbb{R}^k$,

Σ is a k -by- k positive definite matrix and $|\Sigma|$ is its determinant.

Show that $\int_{\mathbb{R}^k} f(x) dx = 1$.

Proof.

Since the hypothesis given, we can rewrite $\int_{\mathbb{R}^k} f(x) dx$

$$\text{as } \int_{\mathbb{R}^k} (2\pi)^k |\Sigma|^{-\frac{1}{2}} \cdot e^{-\frac{1}{2}(x-\mu)^T \Sigma^{-1}(x-\mu)} dx.$$

Let $u = \Sigma^{-\frac{1}{2}}(x-\mu)$. we get $(x-\mu)^T \Sigma^{-1}(x-\mu) = u^T u$ and $dx = \sqrt{|\Sigma|} du$.

$$\text{Then we have } \int_{\mathbb{R}^k} (2\pi)^k |\Sigma|^{-\frac{1}{2}} \cdot e^{-\frac{1}{2}(x-\mu)^T \Sigma^{-1}(x-\mu)} dx = (2\pi)^{\frac{k}{2}} \cdot \int_{\mathbb{R}^k} e^{-\frac{u^T u}{2}} du.$$

$$= (2\pi)^{\frac{k}{2}} \cdot \int_{\mathbb{R}^k} e^{-\frac{1}{2}(\sum_{i=1}^k u_i^2)} du.$$

$$= (2\pi)^{\frac{k}{2}} \cdot \prod_{i=1}^k \int_{\mathbb{R}} e^{-\frac{1}{2} \cdot u_i^2} du_i$$

$$= (2\pi)^{\frac{k}{2}} \cdot \prod_{i=1}^k \sqrt{2\pi} = (2\pi)^{\frac{k}{2}} \cdot (2\pi)^{\frac{k}{2}} = 1,$$

where u_i , for $i=1, 2, \dots, k$ are components of u .

2. Let A, B be n -by- n matrices and x be a n -by-1 vector.

(a) Show that $\frac{\partial}{\partial A} \text{trace}(AB) = B^T$.

$$\text{Since } \text{trace}(AB) = \sum_{i=1}^n (A, B)_{i,i}$$

$$= \sum_{i=1}^n \sum_{j=1}^n A_{i,j} B_{j,i},$$

$$\text{we have } \frac{\partial}{\partial A_{k,l}} \text{trace}(AB) = \frac{\partial}{\partial A_{k,l}} \sum_{i=1}^n \sum_{j=1}^n A_{i,j} B_{j,i}$$

$$= B_{l,k}.$$

Hence $\frac{\partial}{\partial A} \text{trace}(AB) = B^T$.

(b) Show that $x^T A x = \text{trace}(x x^T A)$.

We prove this by expanding both sides.

(Left)

$$\begin{aligned} x^T A x &= \sum_{i=1}^n x_i (Ax)_i \\ &= \sum_{i=1}^n x_i \sum_{j=1}^n A_{i,j} x_j \\ &= \sum_{i=1}^n \sum_{j=1}^n x_i A_{i,j} x_j = \sum_{i=1}^n \sum_{j=1}^n A_{i,j} x_i x_j. \end{aligned}$$

(Right)

Since $\text{trace}(M_1 M_2) = \text{trace}(M_2 M_1)$ for any M_1, M_2 are n -by- n matrices, for all $n \in \mathbb{N}$.

we have $\text{trace}(x x^T A) = \text{trace}(A x x^T)$

$$= \sum_{i=1}^n \sum_{j=1}^n A_{i,j} x_i x_j.$$

Therefore, the left side is equal to right side.

(c) *discuss with ChatGPT and roommate.*

Derive the maximum likelihood estimators for a multivariate Gaussian.

Given x_1, x_2, \dots, x_m are i.i.d. from $N(\mu, \Sigma)$ where $\mu \in \mathbb{R}^n$ the mean vector

$\Sigma \in \mathbb{R}^{n \times n}$ is the covariance matrix.

The log-likelihood function

$$l(\mu, \Sigma) = -\frac{mn}{2} \log(2\pi) - \frac{m}{2} \log |\Sigma| - \frac{1}{2} \sum_{i=1}^m (x_i - \mu)^T \Sigma^{-1} (x_i - \mu).$$

$$\frac{\partial l}{\partial \mu} = \sum_{i=1}^m (x_i - \mu).$$

Setting to zero.

we have $\sum_{i=1}^m (x_i - \mu) = 0$ such that $\hat{\mu} = \frac{1}{m} \sum_{i=1}^m x_i$.

By using matrix identities,

$$\text{we have } \frac{\partial}{\partial \Sigma} \log |\Sigma| = (\Sigma^{-1})^T = \Sigma^{-1}$$

$$\frac{\partial}{\partial \Sigma} (x_i - \mu)^T \Sigma^{-1} (x_i - \mu) = -\Sigma^{-1} (x_i - \mu) (x_i - \mu)^T \Sigma^{-1}.$$

$$\text{Then } \frac{\partial l}{\partial \Sigma} = -\frac{m}{2} \Sigma^{-1} + \frac{1}{2} \Sigma^{-1} \left(\sum_{i=1}^m (x_i - \mu) (x_i - \mu)^T \right) \Sigma^{-1}$$

$$\text{Setting to zero, we get } -m \Sigma^{-1} + \Sigma^{-1} \left(\sum_{i=1}^m (x_i - \mu) (x_i - \mu)^T \right) \Sigma^{-1} = 0.$$

$$\text{Therefore, } \sum_{i=1}^m (x_i - \mu) (x_i - \mu)^T = m \Sigma,$$

$$\hat{\Sigma} = \frac{1}{m} \sum_{i=1}^m (x_i - \hat{\mu}) (x_i - \hat{\mu})^T.$$

