Lecture 11: Inference for One Proportion STAT 630, Fall 2021

Bernoulli random variables and properties of the sample proportion

Let X be a binary random variable following a **Bernoulli distribution** with probability p. We can write this as $X \sim \text{Bern}(p)$. Find the probability mass function, expectation, and variance of X.

Let X_1, X_2, \dots, X_n be independent random variables (i.e., a random sample) from Bern(p). For example, the parameter p is the **population proportion** of individuals that support a certain political candidate; $X_i = 1$ if the i-th randomly sampled person votes for the candidate, and $X_i = 0$ otherwise. We can estimate the population proportion with the **sample proportion** defined as:

$$\hat{p} = \frac{1}{n} \sum_{i=1}^{n} X_i$$

Find the expectation and variance of the sample proportion:

Central limit theorem for sample proportion

Since the sample proportion $\hat{p} = \frac{1}{n} \sum_{i=1}^{n} X_i$ is the sample mean of Bernoulli random variables, the central limit theorem gives the following approximation for the sampling distribution of \hat{p} :

$$\hat{p} \sim N\left(p, \sqrt{\frac{p(1-p)}{n}}\right)$$

The approximation should only be used if $np \ge 10$ and $n(1-p) \ge 10$. This is often called the "success-failure" condition.

Remark: Let $Y = n\hat{p} = \sum_{i=1}^{n} X_i$, which is the sum of independent Bernoulli random variables. Then the exact distribution Y follows is a binomial distribution with parameters n and p. However, as long as the conditions are satisfied, it is often easier to use the normal approximation.

1- α confidence interval for population proportion p

$$\hat{p} \pm z_{\alpha/2} \sqrt{\frac{\hat{p}(1-\hat{p})}{n}}$$

Conditions: $n\hat{p} \ge 10$ and $n(1-\hat{p}) \ge 10$, and the sample observations are independent. Generally, the independence condition is satisfied if the data come from a simple random sample.

Derivation:

By the central limit theorem, $Z = \frac{\hat{p}-p}{\sqrt{p(1-p)/n}} \sim N(0,1)$ approximately. Therefore,

$$P\left(-z_{\alpha/2} < \frac{\hat{p} - p}{\sqrt{p(1-p)/n}} < z_{\alpha/2}\right) = 1 - \alpha$$

After rearranging terms:

$$P\left(\hat{p} - z_{\alpha/2}\sqrt{p(1-p)/n}$$

The issue with this interval is the endpoints are in terms of the unknown population proportion parameter p. As an approximation, we replace $\sqrt{p(1-p)/n}$ with $\sqrt{\hat{p}(1-\hat{p})/n}$ when calculating the interval.

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\mathbf{Sampl}	e size	deter	$\mathbf{m}\mathbf{m}$	ation

Determine the sample size needed so that the confidence interval will have a margin of error $\pm E$ with confidence level $1-\alpha$.

Ex1: At a survey poll before the elections, candidate A receives the support of 650 voters in a random sample of 1200 voters.

- (a) Construct a 95% confidence interval for the population proportion p of voters that support candidate A. Check the conditions for the interval.
- (b) Find the sample size needed so that the margin of error will be ± 0.01 with confidence level 0.95.

Hypothesis test for population proportion p

Null and alternative hypotheses:

 $H_0: p = p_0$

 H_A : $p > p_0$ or $p < p_0$ or $p \neq p_0$

Test statistic: Assuming H_0 is true, $\hat{p} \sim N(p_0, \sqrt{p_0(1-p_0)/n})$. Standardize to get the following z-test statistic:

$$z = \frac{\hat{p} - p_0}{\sqrt{p_0(1 - p_0)/n}}$$

Once the z-test statistic is computed, the p-value can be computed to decide whether to reject or not reject the null hypothesis.

Conditions: $np_0 \ge 10$ and $n(1-p_0) \ge 10$, and independence.

Ex2: A simple random sample of 1,028 US adults in March 2013 found that 56% support nuclear arms reduction. Does this provide convincing evidence that a majority of Americans support nuclear arms reduction?¹

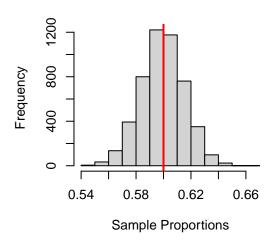
 $^{^{1} \}texttt{https://news.gallup.com/poll/161198/favor-russian-nuclear-arms-reductions.aspx}$

Simulation Study: Simulating the sampling distribution of the sample proportion when:

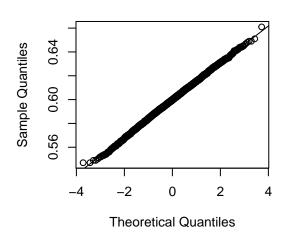
- the population size is 1 million
- the sample size is n = 1000
- the population proportion is p = 0.6

```
set.seed(999)
pop_size <- 10^6
n <- 1000
population <- c(rep(0, 0.4*pop_size), rep(1, 0.6*pop_size))
phats <- c() # initialize vector for sample proportions
for(i in 1:5000) {
   samp <- sample(population, size = n)
   phats[i] <- sum(samp) / n
}</pre>
```

```
par(mfrow = c(1, 2), cex=0.8)
hist(phats, xlab = "Sample Proportions", main = "")
abline(v = 0.6, col = "red", lwd = 2)
qqnorm(phats)
qqline(phats)
```



Normal Q-Q Plot



```
sd(phats)
## [1] 0.01578796

sqrt(0.6 * 0.4 / 1000)
## [1] 0.01549193
```