

Lecture 5:  
Transformations for Simple Linear Regression  
STAT 632, Spring 2020

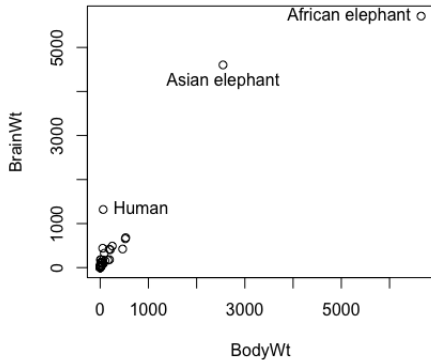
Transformations can be used to

- ▶ Linearize the relationship between the explanatory ( $X$ ) and response ( $Y$ ) variables
- ▶ Overcome problems due to nonconstant variance

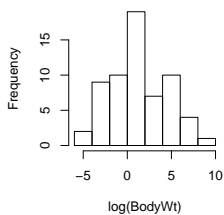
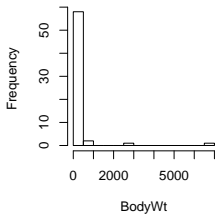
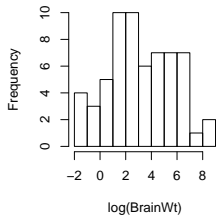
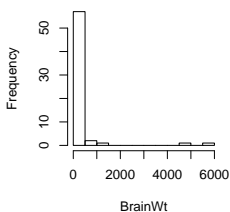
## Example: Modeling Brain Weight

- ▶ We consider a data set called `brains` from the `alr4` package. The data set is on the the brain weight (in grams) and body weight (in kg) for 62 species of mammals.
- ▶ A scatter plot of the data (next slide) shows that the variables are extremely skewed. Three points (humans and two species of elephants) stand out from the rest of the data.

```
> library(alr4)
> plot(BrainWt ~ BodyWt, data=brains)
```

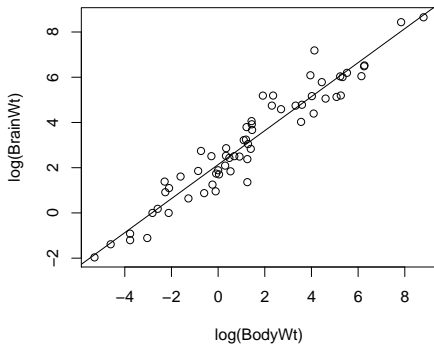


The histograms illustrate that the log transformation can reduce the positive skew in the data.



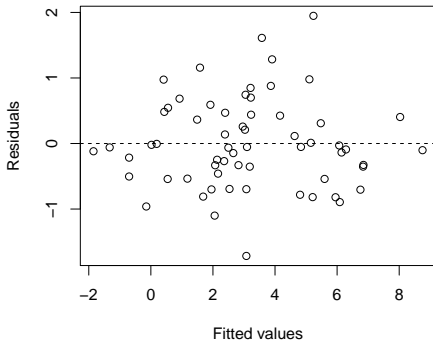
The log transformation linearizes the relationship between the variables.

```
> lm1 <- lm(log(BrainWt) ~ log(BodyWt), data=brains)
> plot(log(BrainWt) ~ log(BodyWt), data=brains)
> abline(lm1)
```



After taking the log transformation, the residual plot shows no discernible patterns (random scatter of points around 0). The assumptions of linearity and constant variance appear to be well satisfied.

```
> plot(predict(lm1), resid(lm1), xlab='Fitted values', ylab='Residuals')  
> abline(h=0, lty=2)
```



```
> summary(lm1)
```

```
Coefficients:
```

	Estimate	Std. Error	t value	Pr(> t )
(Intercept)	2.13479	0.09604	22.23	<2e-16 ***
log(BodyWt)	0.75169	0.02846	26.41	<2e-16 ***

```
---
```

```
Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

```
Residual standard error: 0.6943 on 60 degrees of freedom
```

```
Multiple R-squared:  0.9208, Adjusted R-squared:  0.9195
```

```
F-statistic: 697.4 on 1 and 60 DF,  p-value: < 2.2e-16
```



The R output gives the following regression equation:

$$\begin{aligned}\log(\widehat{\text{BrainWt}}) &= \hat{\beta}_0 + \hat{\beta}_1 \log(\text{BodyWt}) \\ &= 2.135 + 0.7517 \log(\text{BodyWt})\end{aligned}$$

- ▶ The interpretation of the estimated slope is that a unit increase in  $\log(\text{BodyWt})$  is associated with an increase in  $\log(\text{BrainWt})$  by 0.7517 (not that useful).
- ▶ Another common interpretation is in terms of percentage effects: A 1% increase in body weight (kg) is associated with an approximate 0.75% increase in brain weight.<sup>1</sup>

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<sup>1</sup>See Sheather, Section 3.3.2, pp.79-80, for a mathematical explanation

# Review of logs

- ▶ Logs are exponents:  $\log_b(x) = y$  (read “the log of  $x$  to the base  $b$  is  $y$ ”) means that  $b^y = x$ . Some examples:

$$\log_{10} 100 = 2 \iff 10^2 = 100$$

$$\log_{10} 0.01 = -2 \iff 10^{-2} = 0.01$$

$$\log_2 8 = 3 \iff 2^3 = 8$$

- ▶ Some useful identities:

$$e^{\log(x)} = x$$

$$\log(x^r) = r \log(x)$$

$$\log(xy) = \log(x) + \log(y)$$

$$\log(x/y) = \log(x) - \log(y)$$

Note:  $\log(x)$  denotes the log base  $e$  here (this is called the natural logarithm, which is also commonly denoted by  $\ln(x)$ )

We can back-transform to write the model in the original scale of the response.

Regression equation for  $\log(\text{BrainWt})$ :

$$\widehat{\log(\text{BrainWt})} = 2.135 + 0.7517 \log(\text{BodyWt})$$

Make a prediction for  $\log(\text{BrainWt})$  when  $\text{BodyWt}$  is 40 kg:

$$\widehat{\log(\text{BrainWt})} = 2.135 + 0.7517 \log(40) = 4.908$$

We can then exponentiate both sides to get the prediction for  $\text{BrainWt}$  (in grams) when  $\text{BodyWt}$  is 40 kg:

$$\widehat{\text{BrainWt}} = e^{4.908} = 135.37$$

```
# use R to make prediction for log(BrainWt) when BodyWt is 40
> pred <- predict(lm1, data.frame(BodyWt = 40),
                        interval="prediction")

> pred
      fit      lwr      upr
1 4.907668 3.501332 6.314003

# exponentiate to make prediction for BrainWt
> exp(pred)
      fit      lwr      upr
1 135.3235 33.15961 552.2514
```

# Summary: Log Transformation

A log transformation might be useful if

- ▶ the distribution of the response or predictor variable is skewed right.
- ▶ the values of the variable range over more than one order of magnitude.
- ▶ there is a fan pattern in the residuals (nonconstant variance).

**Remark:** The transformation  $\log(x_i)$  is only valid for  $x_i > 0$ . For nonpositive data, one workaround is to use the transformation  $\log(x_i + c)$ , where  $c$  is a constant such that  $x_i + c > 0$ , for  $i = 1, \dots, n$ .

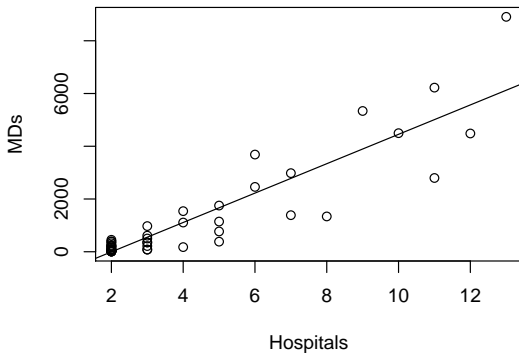
## Example: Doctors and Hospitals

Data set containing counts on the number of medical doctors and number of hospitals in a random sample of 53 counties.

```
> library(Stat2Data)
> data("CountyHealth")
> head(CountyHealth)
```

County	MDs	Hospitals	Beds
1 Bay, FL	351	3	605
2 Beaufort, NC	95	2	134
3 Beaver, PA	260	2	567
4 Bernalillo, NM	2797	11	1435
5 Bibb, GA	769	5	976
6 Clinton, PA	42	2	245

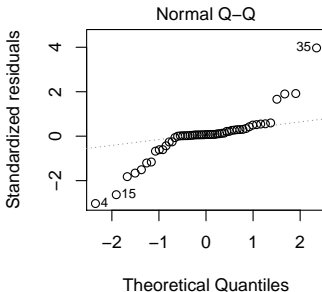
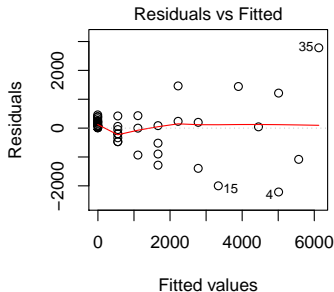
```
> lm1 <- lm(MDs ~ Hospitals, data = CountyHealth)
> plot(MDs ~ Hospitals, data = CountyHealth)
> abline(lm1)
```





There residual plot shows nonconstant variance. That is, the variability in the residuals tends to increase with the fitted values. The QQ plot also indicates that the residuals deviate from a normal distribution.

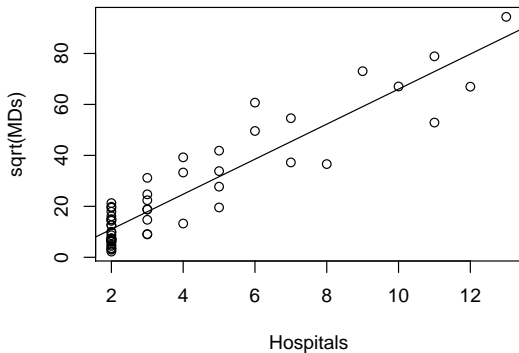
```
> par(mfrow = c(1, 2))  
> plot(lm1, 1:2)
```



# Using Transformations to Stabilize Variance

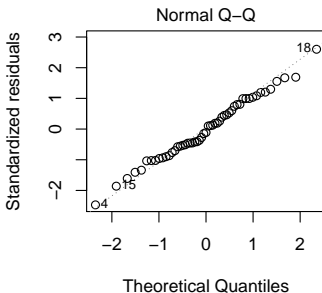
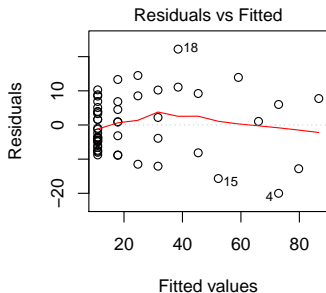
- ▶ Problems with nonconstant variance can be overcome with transformations.
- ▶ Two common variance stabilizing transformations are the log transformation,  $\log(Y)$ , and the square root transformation,  $\sqrt{Y}$ .
- ▶ The square root transformation is often appropriate for count data.
- ▶ Since the data in this example are in the form of counts, we will try a square root transformation of the response.

```
lm2 <- lm(sqrt(MDs) ~ Hospitals, data = CountyHealth)
plot(sqrt(MDs) ~ Hospitals, data = CountyHealth)
abline(lm2)
```



After taking the square root transformation, the residual plot and QQ plot show considerable improvement. The assumptions of constant variability and normality in the residuals appear reasonably satisfied.

```
> par(mfrow = c(1, 2))  
> plot(lm2, 1:2)
```



```
> summary(lm2)
```

Coefficients:

	Estimate	Std. Error	t value	Pr(> t )
(Intercept)	-2.7533	1.9850	-1.387	0.171
Hospitals	6.8764	0.4011	17.144	<2e-16 ***

Regression equation for the transformed model:

$$\widehat{\sqrt{\text{MDs}}} = -2.7533 + 6.8764 \text{Hospitals}$$

We can back-transform to write the model in the original scale of the response.

$$\widehat{\text{MDs}} = (-2.7533 + 6.8764 \text{Hospitals})^2$$

# Summary

- ▶ It can take some trial and error to find a good transformation. Looking at scatterplots of the data and residuals can help determine which transformation best linearizes the relationship or stabilizes the variance.
- ▶ The log transform is commonly applied to skewed data that ranges over several orders of magnitude.
- ▶ The square root transform is commonly applied to count data to stabilize the variance.
- ▶ Transformations can be applied to the response variable, explanatory variable(s), or both.
- ▶ Transforming the response can make the model more difficult to interpret. Predictions and prediction intervals need to be back-transformed so that they can be interpreted on the original scale.